

10/1

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HW 2

1) Hoff's book 3.3

$$y_A = (12, 9, 12, 14, 13, 13, 15, 8, 15, 6)$$

$$y_B = (11, 11, 10, 9, 9, 8, 7, 10, 6, 8, 8, 9, 7)$$

a) Find Posterior dist, Means, Variances, and 95% quantile-based CI for  $\theta_A$  and  $\theta_B$ ...

We know that...

$$y_A \sim \text{Poisson}(\theta_A)$$

$$y_B \sim \text{Poisson}(\theta_B)$$

given we have priors...

$$\theta_A \sim \text{Gamma}(120, 10)$$

$$\theta_B \sim \text{Gamma}(12, 1)$$

\* Independent...

$$P(\theta_A, \theta_B) = P(\theta_A) \times P(\theta_B)$$

$$\frac{\alpha}{\lambda} = 12 \quad \star$$

$$\hookrightarrow \frac{\alpha}{10} = 12 \Rightarrow \alpha = 120$$

Now, establishing Likelihood func for Poisson...

$$P(y|\theta) \propto \theta^{\sum y_i} \times e^{-n\theta}, \text{ where } n =$$

Therefore,

posterior for  $\theta_A$ ... using  $n_A = 10$  and  $\sum y_A = 117$ 

now, combining prior w/ post...

$$\boxed{\theta_A \sim \text{Gamma}(237, 20)}$$

$$\text{Mean } \theta_A: E(\theta_A|y_A) = 237/20 \Rightarrow \boxed{E(\theta_A|y_A) = 11.85}$$

$$\text{Variance } \theta_A: \text{Var}(\theta_A|y_A) = 237/20^2 \Rightarrow \boxed{\text{Var}(\theta_A|y_A) = 0.5925}$$

$$\text{95\% CI for } \theta_A: L_{\theta_A} = q_{\text{gamma}}(0.025, 237, 20) = 10.38924$$

$$\hookrightarrow \boxed{(10.39, 13.41)} \quad U_{\theta_A} = q_{\text{gamma}}(0.975, 237, 20) = 13.40545$$

1) 3.3, a cont'd)

Now, for  $\theta_B \dots$

Posterior  $\theta_B \dots$  using  $n_0 = 13$ ,  $\Sigma y_0 = 118$

$$\theta_B \sim \text{Gamma}(130, 14)$$

$\downarrow$                        $\downarrow$   
 $12+18$                        $1+13$

Mean  $\theta_B \dots$

$$\frac{130}{14} \Rightarrow E(\theta_B | Y_0) = 9.29$$

Variance  $\theta_B \dots$

$$\frac{130}{14^2} \Rightarrow \text{Var}(\theta_B | Y_0) = 0.663$$

95% CI for  $\theta_B$ :

$$L_{\theta_B} = q\text{gamma}(0.025, 130, 14) = 7.75819$$

$$U_{\theta_B} = q\text{gamma}(0.975, 130, 14) = 10.94047$$

$$(7.76, 10.95)$$

b) Compute/plot posterior expectation of  $\theta_B$  under the prior dist...  $\theta_B \sim \text{Gamma}(12 \times n_0, n_0)$

for  $n_0 \in \{1, 2, \dots, 50\}$

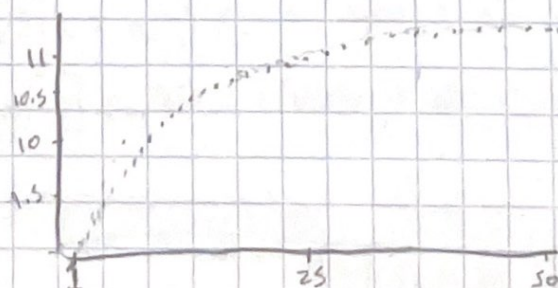
Therefore we would believe that when

Values of " $n_0$ " increases 1 to 50

↳ Smaller  $n_0$ 's have less influence

While bigger ones have greater influence

Plot...



\* Larger  $n_0$ 's have higher  $\theta$ 's but asymptote at 11.5 showing not too much influence



1) 3.3)

c.) Does it make sense to have  $P(\theta_A, \theta_B) = P(\theta_A) \times P(\theta_B)$

↳ Since Both population B and A are related in a way it does make sense combining them in some way, as both share similar pdfs (prior + post) where they use poisson and gamma.

3.4) mix of Beta priors...

prob  $\theta$  of teen recidivism w/

$n=43$  released

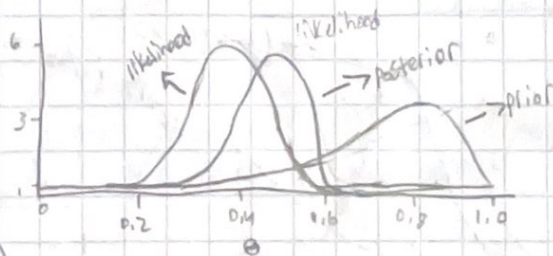
$y=15$  re-offenders  $\rightarrow$  w/in 36 months

a) using  $\text{beta}(2, 8)$  prior for  $\theta$ , plot  $P(\theta)$ ,  $P(y|\theta)$ , and  $P(\theta|y)$  as functions of  $\theta$ . Then find Posterior Mean, Mode, and SD of  $\theta$ . And 95% CI for  $\theta$

w/ a  $P(\theta) = \text{beta}(2, 8)$ , then

$P(y|\theta) = \text{binomial}(n, \theta) \sim y$

$P(\theta|y) = \text{Beta}(17, 36)$



beta(2, 8)

Mean of  $\theta$ ...

$$E(\theta) = \frac{17}{17+36} \Rightarrow \boxed{E(\theta) = 0.321}$$

Mode of  $\theta$ ...

$$\text{Mode} = \frac{17-1}{17+36-2} \Rightarrow \boxed{\text{Mode} = 0.314}$$

SD of  $\theta$ ...

$$\sqrt{\frac{17 \cdot 36}{(17+36)^2 \cdot (17+36+1)}} \Rightarrow \boxed{SD = 0.064}$$

95% CI

$$\begin{aligned} L_0 &= q\text{beta}(0.025, 17, 36) = 0.2032972 \\ U_0 &= q\text{beta}(0.975, 17, 36) = 0.4510240 \end{aligned} \Rightarrow \boxed{(0.203, 0.451)}$$

2) 3.4)

b) Beta(8,2)

using  $\theta|y \sim \text{Beta}(23, 30)$

Mean of  $\theta$

$$E[\theta|y] = \frac{\alpha}{\alpha+\beta} = \frac{23}{23+30} \Rightarrow E[\theta|y] = 0.434$$

Mode of  $\theta$

$$\text{mode} = \frac{\alpha-1}{\alpha+\beta-2} = \frac{23-1}{23+30-2} \Rightarrow \text{mode} = 0.431$$

Standard Dev of  $\theta$

$$SD = \sqrt{\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}} = \sqrt{\frac{23 \times 30}{(53)^2(54)}} \Rightarrow SD_{\theta} = 0.0674$$

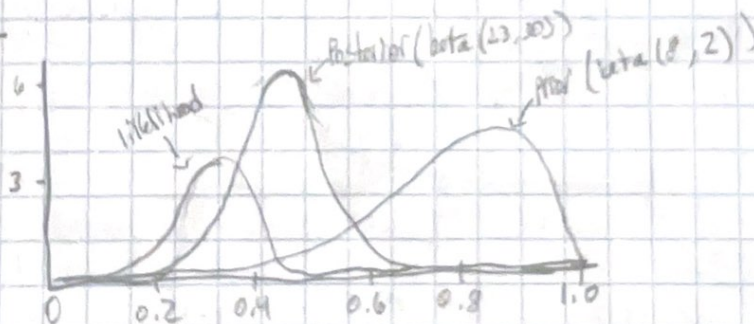
95% CI of  $\theta$

$$L_{\theta} = q_{\text{beta}}(0.025, 23, 30) = 0.3046956$$

$$U_{\theta} = q_{\text{beta}}(0.975, 23, 30) = 0.5679528$$

$$(0.305, 0.568)$$

Plot



Part C



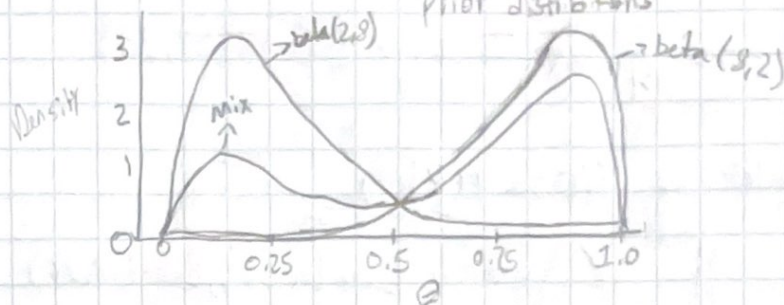
2) 34) C.)

Consider...

$$P(\theta) = \frac{1}{4} \frac{\Gamma(10)}{\Gamma(2)\Gamma(8)} [3\theta(1-\theta)^7 + \theta^7(1-\theta)]$$

↳ mix of beta(2,8) and beta(8,2)

prior distributions



Looking at the plot of the 3 priors, we can conclude that w/ the 75% weight on beta(8,2) and 25% on beta(2,8) it represents the prior for teen recidivism is more likely to skew w/ higher values, along w/ some consideration for lower values. Overall is emphasizing a strong belief in high recidivism.

d) For prior in C)...

i) write out  $P(\theta) \times P(y|\theta)$  and simplify

$$P(\theta) = \frac{1}{4} \text{Beta}(2,8) + \frac{3}{4} \text{Beta}(8,2)$$

using  $y=15$  and  $n=43$

$$P(y|\theta) \propto \theta^y(1-\theta)^{n-y} = \theta^{15}(1-\theta)^{28}$$

now...

$$P(\theta) * P(y|\theta) \propto \left( \frac{1}{4}\theta(1-\theta)^7 + \frac{3}{4}\theta^7(1-\theta) \right) * \theta^{15}(1-\theta)^{28}$$

↳ simplifies to...

$$P(\theta|y) \propto \frac{1}{4}\theta^{16}(1-\theta)^{35} + \frac{3}{4}\theta^{22}(1-\theta)^{19}$$



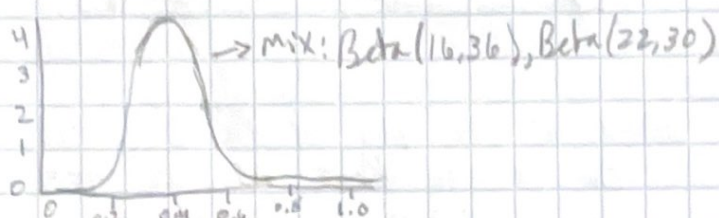
d) ii) ID two Posterior dist

1) Beta(16, 36)  $\rightarrow$  beta(28) w/ likelihood

2) Beta(22, 30)  $\rightarrow$  beta(8, 2) w/ likelihood

$$\Rightarrow \frac{1}{4} \text{Beta}(16, 36) + \frac{3}{4} \text{Beta}(22, 30)$$

iii) plot of  $P(\theta) \times P(y|\theta)$



$$\text{mode}_{B(16, 36)} = 0.30$$

$$\text{mode}_{B(22, 30)} = 0.41$$

The mix ~~prior~~ posterior mode compares more w/ beta(8, 2) ~~prior~~ posterior mode as mix prior is weighted more heavily w/ beta(8, 2) prior.

e) i) General weights formula for mix prior/posterior

$$P(\theta|y) \propto P(y|\theta) \times \text{prior}$$

$$\text{Where } i=1, 2 \quad W_i = \frac{P(y|\text{prior}_i) \times w_i}{P(y|\text{prior}_1) \times w_1 + P(y|\text{prior}_2) \times w_2}$$

$\Rightarrow$  Where  $w_1$  and  $w_2$  represent weights of priors, and

$w_1$  and prior<sub>1</sub>  $\rightarrow$  goes w/ Beta(28)

$w_2$  and prior<sub>2</sub>  $\rightarrow$  goes w/ Beta(8, 2)



### 3) Jeffery's Prior

a) If  $\theta \sim \text{Beta}(\alpha, \beta)$ , show that  $\lambda = \log(\theta/(1-\theta)) \rightarrow \text{logit}(\theta)$

$$P(\theta) \propto \theta^{\alpha-1} (1-\theta)^{\beta-1}$$

$$\text{w/ } \lambda = \log\left(\frac{\theta}{1-\theta}\right)$$

$$\theta = \frac{e^\lambda}{1+e^\lambda} \rightarrow \frac{d\theta}{d\lambda} = \frac{e^\lambda}{(1+e^\lambda)^2}$$

Combine...

$$P(\lambda) \propto e^{(\alpha-1)\lambda} (1+e^\lambda)^{-\alpha-\beta+2} \propto \frac{e^\lambda}{(1+e^\lambda)^2}$$

$$\propto \frac{e^{\lambda/2}}{(1+e^\lambda)^{3/2}}$$

b) Jeffery's prior w/  $\lambda$ : is  $\text{Beta-logit}(\frac{1}{2}, \frac{1}{2})$

Log-likelihood of  $x|\lambda \sim \text{Binom}(n, \text{logit}^{-1}(\lambda))$  is...

$$\log P(X|\lambda) = n * \lambda * X - n \log(1+e^\lambda)$$

Jefferys prior should be proportionally square root to fisher criteria...

$$I(\lambda) \propto \frac{e^\lambda}{(1+e^\lambda)^2}$$

So,

$$P(\lambda) \propto \sqrt{I(\lambda)} \propto \frac{e^{\lambda/2}}{(1+e^\lambda)^{3/2}}$$

which is exact form of beta-logit  
w/ parameters  $\alpha = \frac{1}{2}, \beta = \frac{1}{2}$

$$\therefore \theta \sim \text{Beta}(\frac{1}{2}, \frac{1}{2})$$

since  $\lambda = \log\left(\frac{\theta}{1-\theta}\right)$ , transformed  $\theta = \text{logit}^{-1}(\lambda)$ , meaning  
if  $\lambda$  has a beta-logit dist w/  $\alpha = \frac{1}{2}$  and  $\beta = \frac{1}{2}$   
then dist will be  $\text{Beta}(\frac{1}{2}, \frac{1}{2})$