

Precise EM-based LCDG-approximation of a Probability Distribution

I. SEQUENTIAL INITIALIZATION OF THE LCDG MODEL

The initial LCDG model closely approximating a given marginal gray level distribution \mathbf{F} is built using the conventional EM-algorithm [1]–[3] adapted to the DGs. The approximation involves the following steps:

- 1) The distribution \mathbf{F} is approximated with a mixture \mathbf{P}_K of K positive DGs, relating each to a dominant mode.
- 2) Deviations between \mathbf{F} and \mathbf{P}_K are approximated with the scaled sign-alternate “subordinate” DGs as follows:
 - a) The positive and negative deviations are separated and scaled up to form two temporary “probability distributions” \mathbf{D}^p and \mathbf{D}^n .
 - b) The same conventional EM algorithm is used iteratively to find a subordinate mixture of positive or negative DGs that approximates best \mathbf{D}^p or \mathbf{D}^n , respectively (the sizes $C_p - K$ and C_n of these mixtures are found by minimizing sequentially the total absolute error between each “distribution” \mathbf{D}^p or \mathbf{D}^n and its mixture model by the number of the components).
 - c) The obtained positive and negative subordinate mixtures are scaled down and then added to the dominant mixture yielding the initial LCDG model of the size $C = C_p + C_n$.

The resulting initial LCDG has K dominant non-negative weights, $w_{p,1}, \dots, w_{p,K}$, such that $\sum_{r=1}^K w_{p,r} = 1$, and a number of subordinate smaller non-negative weights such that $\sum_{r=K+1}^{C_p} w_{p,r} - \sum_{l=1}^{C_n} w_{n,l} = 0$.

II. FINAL REFINEMENT OF THE LCDG MODEL

The initial LCDG is refined by approaching the closest local maximum of the log-likelihood

$$L(\mathbf{w}, \Theta) = \sum_{q \in \mathbf{Q}} f(q) \log p_{\mathbf{w}, \Theta}(q) \quad (1)$$

with the EM process adapting that in [4] to the DGs. The latter process extends in turn the conventional EM algorithm in [2] onto the sign-alternate linear combinations.

Let $p_{\mathbf{w}, \Theta}^{[m]}(q) = \sum_{r=1}^{C_p} w_{p,r}^{[m]} \psi(q|\theta_{p,r}^{[m]}) - \sum_{l=1}^{C_n} w_{n,l}^{[m]} \psi(q|\theta_{n,l}^{[m]})$ denote the current LCDG at iteration m . Relative contributions of each signal $q \in \mathbf{Q}$ to each positive and negative DG at iteration m are specified by the respective conditional weights

$$\pi_p^{[m]}(r|q) = \frac{w_{p,r}^{[m]} \psi(q|\theta_{p,r}^{[m]})}{p_{\mathbf{w}, \Theta}^{[m]}(q)}; \quad \pi_n^{[m]}(l|q) = \frac{w_{n,l}^{[m]} \psi(q|\theta_{n,l}^{[m]})}{p_{\mathbf{w}, \Theta}^{[m]}(q)} \quad (2)$$

such that the following constraints hold:

$$\sum_{r=1}^{C_p} \pi_p^{[m]}(r|q) - \sum_{l=1}^{C_n} \pi_n^{[m]}(l|q) = 1; \quad q = 0, \dots, Q-1 \quad (3)$$

The following two steps iterate until the log-likelihood is increasing and its changes become small:

- E– step^[m]:** Find the weights of Eq. (2) under the fixed parameters $\mathbf{w}^{[m-1]}, \Theta^{[m-1]}$ from the previous iteration $m-1$, and
- M– step^[m]:** Find the conditional MLEs $\mathbf{w}^{[m]}, \Theta^{[m]}$ by maximizing $L(\mathbf{w}, \Theta)$ under the fixed weights of Eq. (2).

Considerations closely similar to those in [1]–[3] show this process converges to a local log-likelihood maximum. As shown in a very general way in [3] and for the LCDG model in [4], it is actually a block relaxation MM-process. Let the log-likelihood of Eq. (1) be rewritten in the equivalent form with the constraints of Eq. (3) as unit factors:

$$L(\mathbf{w}^{[m]}, \Theta^{[m]}) = \sum_{q=0}^Q f(q) \left[\sum_{r=1}^{C_p} \pi_p^{[m]}(r|q) \log p^{[m]}(q) \right] - \sum_{q=0}^Q f(q) \left[\sum_{l=1}^{C_n} \pi_n^{[m]}(l|q) \log p^{[m]}(q) \right] \quad (4)$$

Let the terms $\log p^{[m]}(q)$ in the first and second brackets be replaced with the equal terms $\log w_{p,r}^{[m]} + \log \psi(q|\theta_{p,r}^{[m]}) - \log \pi_p^{[m]}(r|q)$ and $\log w_{n,l}^{[m]} + \log \psi(q|\theta_{n,l}^{[m]}) - \log \pi_n^{[m]}(l|q)$, respectively, which follow from Eq. (2).

At the E-step, the constrained Lagrange maximization of the log-likelihood of Eq. (4) under the Q constraints of Eq. (3) results just in the weights $\pi_p^{[m+1]}(r|q)$ and $\pi_n^{[m+1]}(l|q)$ of Eq. (2) for all $r = 1, \dots, C_p$; $l = 1, \dots, C_n$ and $q \in \mathbf{Q}$. At the M-step, the DG weights $w_{p,r}^{[m+1]} = \sum_{q \in \mathbf{Q}} f(q) \pi_p^{[m+1]}(r|q)$ and $w_{n,l}^{[m+1]} = \sum_{q \in \mathbf{Q}} f(q) \pi_n^{[m+1]}(l|q)$ follow from the constrained Lagrange maximization of the log-likelihood in Eq. (4) under the constraints $\sum_{r=1}^{C_p} w_{p,r} - \sum_{l=1}^{C_n} w_{n,l} = 1$ and the fixed conditional weights of Eq. (2). Under these constraints, the conventional MLEs of the parameters of each DG stem from maximizing the log-likelihood after each difference of the cumulative Gaussians is replaced with its close Gaussian density approximation (below “c” stands for “p” or “n”, respectively):

$$\mu_{c,r}^{[m+1]} = \frac{1}{w_{c,r}^{[m+1]}} \sum_{q \in \mathbf{Q}} q \cdot f(q) \pi_c^{[m+1]}(r|q)$$

$$(\sigma_{c,r}^{[m+1]})^2 = \frac{1}{w_{c,r}^{[m+1]}} \sum_{q \in \mathbf{Q}} (q - \mu_{c,i}^{[m+1]})^2 \cdot f(q) \pi_c^{[m+1]}(r|q)$$

This modified EM-algorithm is valid until the weights \mathbf{w} are strictly positive. The iterations should be terminated when the log-likelihood of Eq. (1) almost does not change or begins to decrease due to accumulation of rounding errors.

The final mixed LCDG-model $p_C(q)$ is partitioned into the K LCDG-submodels $\mathbf{P}_{[k]} = [p(q|k) : q \in \mathbf{Q}]$, one per dominant mode, or class $k = 1, \dots, K$, by associating

the subordinate DGs with the dominant terms so that the misclassification rate is minimal.

REFERENCES

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