

# Lecture 23

## Method of Moments & Maximum Likelihood Estimation

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## 2 General Methods for estimating parameters:

1. Method of moments estimation (MoM)
2. Maximum likelihood estimation (MLE)

# Method of Moments (MoM)

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# Method of Moments (MoM)

## Definition:

- The  $k^{th}$  *moment* of a R.V  $X$  is defined as  $\mu_k = E(X^k)$
- The  $k^{th}$  *sample moment* is defined as  $m_k = \frac{1}{n} \sum_{i=1}^n X_i^k$

The *method of moments (MoM)* estimators for parameters are found by equating (known) sample moments to (unknown) population moments, and then solving for the parameters in terms of the data.

- If our model has more than one unknown parameter, we need to make equations with more than one moment.
- In general, need  $k$  equations to derive MoM estimators for  $k$  parameters.

To obtain MoM estimators for  $k$  parameters: Set the sample moments ( $m_k$ ) equal to population moments ( $\mu_k$ ), and solve.

- $m_1 = \mu_1 \rightarrow \frac{1}{n} \sum x_i = E(X)$
- $m_2 = \mu_2 \rightarrow \frac{1}{n} \sum x_i^2 = E(X^2)$
- $\vdots$
- $m_k = \mu_k \rightarrow \frac{1}{n} \sum x_i^k = E(X^k)$

**Note:**

- MoM estimators may be biased
- Sometimes you can get estimates outside of parameter space

## MoM Examples

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Example 1: Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Geo}(p)$

Estimate one parameter  $p \rightarrow$  need the first moment.

- 1<sup>st</sup> (population) moment:  $\mu_1 = E(X) = \frac{1}{p}$ .
- 1<sup>st</sup> sample moment is  $m_1 = \frac{1}{n} \sum_{i=1}^n X_i = \bar{X}$

Set 1<sup>st</sup> moment equal 1<sup>st</sup> sample moment, and solve for  $p$ .

$$\frac{1}{p} = \bar{X} \rightarrow \hat{p}_{MoM} = \frac{1}{\bar{X}}$$

Example 2: Let  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

Estimate two parameters  $\rightarrow$  need first two moments

Set the first two moments equal to the first two sample moments.

1.  $\frac{1}{n} \sum_{i=1}^n X_i = E(X)$
2.  $\frac{1}{n} \sum_{i=1}^n X_i^2 = E(X^2)$

For our random variables,  $E(X) = \mu$  and  $Var(X) = \sigma^2$

From Eq 1, we have  $\frac{1}{n} \sum X_i = E(X) = \mu$

$$\rightarrow \hat{\mu}_{MoM} = \frac{1}{n} \sum X_i$$

$$\rightarrow \hat{\mu}_{MoM} = \bar{X}$$



## MoM Examples Cont.

$$\text{Var}(X) = E(X^2) - E(X)^2 \rightarrow E(X^2) = \text{Var}(X) + E(X)^2 = \sigma^2 + \mu^2$$

From Eq 2. we have:

$$\frac{1}{n} \sum_{i=1}^n X_i^2 = E(X^2) = \sigma^2 + \mu^2$$

$$\rightarrow \sigma^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \mu^2$$

$$\rightarrow \hat{\sigma}_{MoM}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \hat{\mu}_{MoM}^2$$

$$= \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}^2$$

$$= \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$$

# Maximum Likelihood Estimation (MLE)

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# Likelihood Function

We have  $X_1, \dots, X_n \stackrel{iid}{\sim} f_X(x)$ , where  $f_X(x)$  has (unknown) parameter  $\theta$ .

The model for our data is the *joint distribution* of  $X_1, \dots, X_n$

$$f_X(x_1, \dots, x_n) = \prod_{i=1}^n f_X(x_i)$$

When the joint distribution is viewed as a function of the unknown parameter, it is referred to as the *likelihood function*

$$\mathcal{L}(\theta) = \prod_{i=1}^n f_X(x_i)$$

## Likelihood Example

Example 3:  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$

The marginal distribution of each  $X_i$  is

$$f_X(x) = \frac{e^{-\lambda} \lambda^{x_i}}{x_i!}$$

The joint distribution/likelihood function is

$$\begin{aligned}\mathcal{L}(\lambda) = f(x_1, \dots, x_n) &= \prod_{i=1}^n f_X(x_i) \\ &= \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} \\ &= \frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!}\end{aligned}$$

# Maximum Likelihood Estimation (MLE)

## Definition

A *maximum likelihood estimator*  $\hat{\theta}_{MLE}$  of  $\theta$  is the function that “maximizes the likelihood (probability) of the data”

Thus, the MLE maximizes the joint distribution model or likelihood function:

$$\hat{\theta}_{MLE} = \operatorname{argmax}_{\theta \in \Theta} \mathcal{L}(\theta) = \operatorname{argmax}_{\theta \in \Theta} \prod_{i=1}^n f(x_i)$$

Example 4: Flip a coin 10 times. Let  $X$  be the # of heads obtained. A reasonable model for  $X$  is  $\text{Bin}(n = 10, p)$  where  $p$  is our unknown parameter that we would like to estimate.

Suppose we observe the value  $x = 3$ . (only 1 data value).

Since there's only 1 data value, the likelihood/joint distribution is just the marginal distribution  $f(x)$ :

$$\begin{aligned}\mathcal{L}(p) = f(x) &= \binom{10}{x} p^x (1-p)^{10-x} \\ &= \binom{10}{3} p^3 (1-p)^{10-3} \\ &= 120 p^3 (1-p)^7\end{aligned}$$

What value of  $p$  maximizes the likelihood?

# General Calculation of MLE

- Maximizing the likelihood from  $L(\theta)$  when there are multiple observed values becomes difficult.
- The common trick is to use the *log-likelihood function* instead:

$$\ell(\theta) = \log \mathcal{L}(\theta)$$

where  $\ell(\cdot)$  is the natural-log

→ Since  $\ell(\cdot)$  is increasing, the same  $\theta$  that maximizes log-likelihood  $\ell(\cdot)$  also maximizes the likelihood  $\mathcal{L}(\theta)$

- Use calculus to find the maximum of  $\ell(\theta)$



### Finding MLE:

1. Find the likelihood function:  $\mathcal{L}(\theta) = \prod_{i=1}^n f(x_i)$
2. Find the log-likelihood function:  $\ell(\theta) = \log \mathcal{L}(\theta)$
3. Take the first derivative:  $\ell'(\theta) = \frac{d}{d\theta} \ell(\theta)$
4. Set  $\ell'(\theta) = 0$  and solve for  $\theta$   
→ this is your  $\hat{\theta}_{MLE}$
5. Check if second derivative  $\ell''(\theta) < 0$  to make sure  $\hat{\theta}_{MLE}$  is maximum

## MLE Examples

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# MLE Examples

Example 5: Roll a (6-sided) die until you get a 6, and record the number of rolls. Repeat for 100 trials. For  $i = 1, \dots, 100$ ,

$X_i = \#$  of rolls until you obtain a 6 in the  $i^{th}$  trial

$$X_i \stackrel{iid}{\sim} \text{Geo}(p) \text{ and } f(x_i) = p(1-p)^{x_i-1}$$

Data:

x	1	2	3	4	5	6	7	8	9
#	18	20	8	9	9	5	8	3	5
x	11	14	15	16	17	20	21	27	29
#	3	3	3	1	1	1	1	1	1

## MLE Examples Cont.

1. Find the likelihood function  $\mathcal{L}(p)$ :

$$\mathcal{L}(p) = \prod_{i=1}^n f(x_i) = \prod_{i=1}^n p(1-p)^{x_i-1} = p^n(1-p)^{\sum_{i=1}^n x_i - n}$$

2. Find the log-likelihood function  $\ell(p) = \log \mathcal{L}(p)$ :

$$\ell(p) = \log \mathcal{L}(p)$$

3. Take the 1<sup>st</sup> derivative w.r.t  $p$ :  $\ell'(p)$ :

$$\ell'(p) = \frac{d}{dp}\ell(p) = \frac{d}{dp}n\log(p) + \left(\sum_{i=1}^n x_i - n\right)\log(1-p)$$

4. Set  $\ell'(p) = 0$  and solve for  $p$ :

$$\frac{d}{dp}\ell(p) \stackrel{\text{set}}{=} 0$$

5. 2<sup>nd</sup> derivative test to confirm we have maximum:

$$\frac{d^2}{dp^2} \ell(p)$$

Plug in the data into our MLE: