ISYE 6414 - HW7

Manvitha Kalicheti

gtID: 903838438

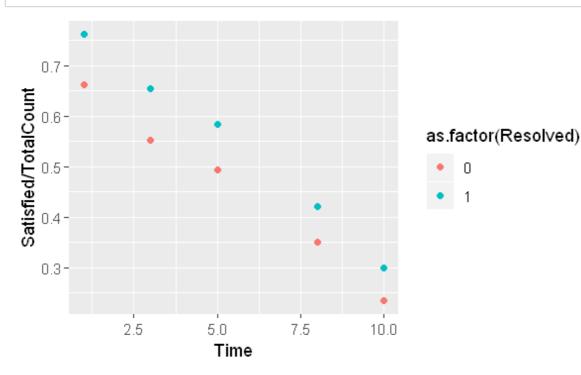
```
In [1]:
```

```
library(ggplot2)
data1 = read.csv('homework07data01.csv')
TotalCount = data1[, 1]
Time = data1[, 2]
Resolved = data1[, 3]
Satisfied = data1[, 4]
Registered S3 methods overwritten by 'ggplot2':
```

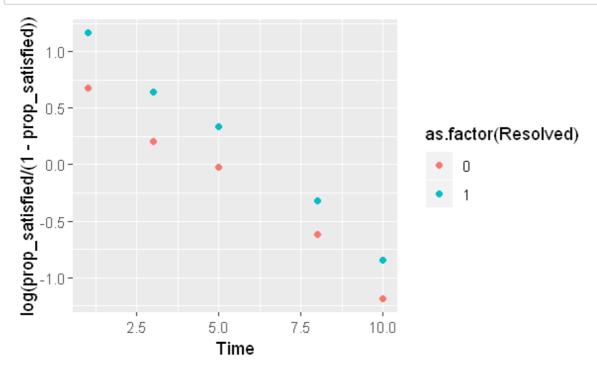
method from
[.quosures rlang
c.quosures rlang
print.quosures rlang

Q1

In [2]:



In [3]:



The trend of the logit function vs time looks almost linear, seems like logistic regression would be a good fit.

Q2

In [4]:

```
In [5]:
```

```
summary(model1)
```

```
Call:
glm(formula = Satisfied/TotalCount ~ Resolved + Time, family = binomial,
    weights = TotalCount)
Deviance Residuals:
     Min
         10
                      Median
                                    3Q
                                             Max
-0.46754 -0.28478 -0.03754 0.34630
                                         0.44079
Coefficients:
            Estimate Std. Error z value Pr(>|z|)
(Intercept) 0.87899 0.12414 7.080 1.44e-12 ***
             0.41527
                        0.12510
                                 3.319 0.000902 ***
Resolved
                        0.02327 -8.577 < 2e-16 ***
Time
            -0.19964
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 95.7521 on 9 degrees of freedom
Residual deviance: 1.0727 on 7 degrees of freedom
AIC: 55.886
Number of Fisher Scoring iterations: 3
\hat{\beta}_0 = 0.87899
\hat{\beta}_{Resolved} = 0.41527
\hat{\beta}_{Time} = -0.19964
```

Q3

The log odds of customer satisfaction seem to increase by 0.41527 for Resolved vs. Not Resolved (odds ratio is $e^{0.41527} = 1.5148$) holding Time fixed. Thus the odds of satisfaction are $51.48\% (= e^{0.41527} - 1)$ higher for customers whose issues were resolved holding Time fixed.

Q4

The log odds of customer satisfaction seem to decrease by -0.19964 for a unit (1 minute) increase in time (odds ratio is $e^{-0.19964} = 0.8190$) holding all other variables fixed. Thus the odds of satisfaction are $18.09\% (= e^{-0.19964} - 1)$ lower for 1 minute increase in time.

Q5

In [6]:

```
new1 = data.frame(Time = 3, Resolved = 1)
new2 = data.frame(Time = 3, Resolved = 0)
```

In [7]:

```
predict(model1, new1)
```

1: 0.695353982258555

In [8]:

```
predict(model1, new2)
```

1: 0.280079539245874

The mean log odds that a customer will be satisfied when their call took 3 minutes and their issue was resolved is 0.6953. Odds ratio is $e^{0.6953} = 2.0043$. The odds that the customer will be satisfied are $100 \times \frac{2.0043}{1+2.0043}\% = 66.71\%$.

The mean log odds that a customer will be satisfied when their call took 3 minutes and their issue was not resolved is 0.2801. Odds ratio is $e^{0.2801}=1.3233$. The odds that the customer will be satisfied are $100\times\frac{1.3233}{1+1.3233}\%=56.96\%$.

In [9]:

```
data2 = read.csv('homework07data02.csv')
PriorWeekPurchase = data2[, 1]
LastWeekCompPurchase = data2[, 2]
Age = data2[, 3]
Region = data2[, 4]
Count = data2[, 5]
Purchase = data2[, 6]
PurchaseCount = data2[, 7]
```

Q6

In [10]:

In [11]:

summary(model2)

```
Call:
```

glm(formula = Purchase/Count ~ PriorWeekPurchase + LastWeekCompPurchase +
 Age + Region, family = binomial, weights = Count)

Deviance Residuals:

Min 1Q Median 3Q Max -2.6513 -0.8410 0.2038 0.7031 2.2835

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-0.33961	0.12612	-2.693	0.00709	**
PriorWeekPurchase	0.38392	0.08650	4.438	9.06e-06	***
LastWeekCompPurchase	-0.26885	0.08655	-3.106	0.00189	**
Age25 to 34	0.06858	0.12053	0.569	0.56935	
Age35 to 44	-0.14704	0.12190	-1.206	0.22773	
Age45+	-0.24521	0.12294	-1.995	0.04608	*
Regionb	0.19027	0.12205	1.559	0.11902	
Regionc	0.24966	0.12165	2.052	0.04014	*
Regiond	0.26563	0.12273	2.164	0.03044	*
Signif. codes: 0 '*	**' 0.001	'**' 0.01 [']	'*' 0.05	'.' 0.1	' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 118.705 on 63 degrees of freedom Residual deviance: 74.991 on 55 degrees of freedom

AIC: 344.32

Number of Fisher Scoring iterations: 4

$$\hat{\beta}_0 = -0.33961$$

$$\hat{\beta}_{PriorWeekPurchase} = 0.38392$$

$$\hat{\beta}_{LastWeekCompPurchase} = -0.26885$$

$$\hat{\beta}_{Age:25-34} = 0.06858$$

$$\hat{\beta}_{Age:35-44} = -0.14704$$

$$\hat{\beta}_{Age:45+} = -0.24521$$

$$\hat{\beta}_{Region-b} = 0.19027$$

$$\hat{\beta}_{Region-c} = 0.24966$$

$$\hat{\beta}_{Region-d} = 0.26563$$

Null deviance - Residual Deviance = 43.714

```
In [12]:
```

```
1 - pchisq(model2$null.deviance-model2$deviance, 63-55)
```

6.44101047608814e-07

 H_0 : all β s are 0.

 H_a : atleast one β is non zero.

P-value = $Pr(\mathcal{X}_1^2 > 43.714) \sim 0$

p-value << $\alpha=0.05 \implies$ we reject $H_0 \implies$ We accept H_a : atleast one β is non zero.

Model is statistically significant.

Q8

In [13]:

```
c(deviance(model2), 1-pchisq(deviance(model2), 55))
```

74.9907728616975 0.0378463616235688

In [14]:

```
pearres2 = residuals(model2,type="pearson")
pearson.tvalue = sum(pearres2^2)
c(pearson.tvalue, 1-pchisq(pearson.tvalue,55))
```

73.5225716343887 0.0483089734645745

 H_0 : the logistic model fits the data.

 H_a : the logistic model does not fit the data.

With both deviance and Pearson residuals, p-value < α = 0.05, thus, we reject the null hypothesis of good fit with α = 0.05.

With both deviance and Pearson residuals, p-value > α = 0.01, thus, we accept the null hypothesis of good fit with α = 0.01.

Q9

In [15]:

In [16]:

summary(model3)

```
Call:
glm(formula = PurchaseCount ~ PriorWeekPurchase + LastWeekCompPurchase +
    Age + Region + offset(log(Count)), family = "poisson")
Deviance Residuals:
     Min
                1Q
                       Median
                                      3Q
                                               Max
-2.48516 -0.65035 -0.06216
                                0.63731
                                           2.77040
Coefficients:
                      Estimate Std. Error z value Pr(>|z|)
(Intercept)
                                   0.04021 18.665 < 2e-16 ***
                       0.75055
PriorWeekPurchase
                       0.39280
                                   0.02678 14.669 < 2e-16 ***
                                   0.02663 -8.099 5.54e-16 ***
LastWeekCompPurchase -0.21569
Age25 to 34
                       0.06285
                                   0.03586 1.753 0.07967 .
Age35 to 44
                      -0.04930
                                   0.03686 -1.337 0.18110
                                   0.03929 -6.352 2.12e-10 ***
Age45+
                      -0.24960
Regionb
                       0.11974
                                   0.03914
                                            3.060 0.00222 **
                                             7.472 7.92e-14 ***
                                  0.03760
Regionc
                       0.28095
                       0.17495
                                   0.03876 4.513 6.38e-06 ***
Regiond
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for poisson family taken to be 1)
    Null deviance: 492.205 on 63 degrees of freedom
Residual deviance: 70.202 on 55 degrees of freedom
AIC: 490.66
Number of Fisher Scoring iterations: 4
\hat{\beta}_0 = 0.75055
\beta_{PriorWeekPurchase} = 0.39280
\beta_{LastWeekCompPurchase} = -0.21569
\hat{\beta}_{Age:25-34} = 0.06285
\hat{\beta}_{Age:35-44} = -0.04930
\hat{\beta}_{Age:45+} = -0.24960
\hat{\beta}_{Region-b} = 0.11974
```

 $\beta_{Region-c} = 0.28095$

 $\hat{\beta}_{Region-d} = 0.17495$

In [17]:

1 - pchisq(model3\$null.deviance-model2\$deviance, 63-55)

0

 H_0 : all β s are 0.

 H_a : atleast one β is non zero.

P-value = $Pr(\mathcal{X}_k^2 > Deviance) = 0$

p-value $< \alpha = 0.05 \implies$ we reject $H_0 \implies$ We accept H_a : atleast one β is non zero.

Model is statistically significant.

Q11

In [18]:

```
c(deviance(model3), 1-pchisq(deviance(model3), 55))
```

In [19]:

```
pearres3 = residuals(model3,type="pearson")
pearson.tvalue = sum(pearres3^2)
c(pearson.tvalue, 1-pchisq(pearson.tvalue,55))
```

 H_0 : the poisson model fits the data.

 H_a : the poisson model does not fit the data.

In both cases, p-value > α = 0.05 and 0.01, thus, we accept the null hypothesis of good fit.

Q12

$$H_0: \hat{\beta}_{Age:25-34} = \hat{\beta}_{Age:35-44} = \hat{\beta}_{Age:45+} = 0$$

 H_a : Atleast one of $\hat{eta}_{Age:25-34},\hat{eta}_{Age:35-44},\hat{eta}_{Age:45+}
eq 0$

For the model made in 6, the p-value for $\beta_{Age:45+}$ is < $\alpha=0.05$. Thus, we reject H_0 .

- \implies Atleast one of $\hat{\beta}_{Age:25-34}, \hat{\beta}_{Age:35-44}, \hat{\beta}_{Age:45+} \neq 0$
- ⇒ categorical variable 'Age' is statistically significant.

Q13

$$H_0: \hat{\beta}_{Age:25-34} = \hat{\beta}_{Age:35-44} = \hat{\beta}_{Age:45+} = 0$$

$$H_a$$
 : Atleast one of $\hat{eta}_{Age:25-34},\hat{eta}_{Age:35-44},\hat{eta}_{Age:45+}
eq 0$

For the model made in 9, the p-values for all $\hat{\beta}_{Age:25-34}$, $\hat{\beta}_{Age:35-44}$ are > $\alpha=0.05$. But the p-value for $\hat{\beta}_{Age:45+}$ << $\alpha=0.05$. Thus, we reject H_0 .

- \implies Atleast one of $\hat{\beta}_{Age:25-34}, \hat{\beta}_{Age:35-44}, \hat{\beta}_{Age:45+} \neq 0$
- ⇒ Coefficients associated with Age cannot be 0.

Q14

$$\hat{\beta}_{PriorWeekPurchase} = 0.38392$$

The log odds of someone purchasing the new product seem to increase by 0.38392 if the customer bought one of the company's products in the previous week holding all other variables fixed. Odds ratio is $e^{0.38392} = 1.4680$. Thus, the odds of the customer purchasing the new product are $46.80\% (= e^{0.38392} - 1)$ higher for customers who bought one of the company's products in the previous week.

Q15

$$\hat{\beta}_{PriorWeekPurchase} = 0.39280$$

If the customer bought one of the company's products in the previous week, holding all other variables fixed, the rate ratio of new item purchased would be expected to increase by a factor of $e^{0.39280} = 1.4811$.