3.4 Log-linear regression

1. Introduction

The term log-linear models is used to describe any model of the form

$$\log \mathbb{E}(\mathrm{Y}_i) = c + \mathbf{x}_i^{ op} \boldsymbol{\beta}$$
 (3.4.1)

In this section, we will analyse the introduction of interaction terms and the analogous of the ANOVA for log-linear models.

2. Modelling example

Example:

Let's consider a cross-sectional study of patients with a form of skin cancer called "malignant melanoma". We have 400 patients, with information about the site of the tumour and its histological type. The data are collected into the **melanoma** dataset available in the **dobson** package in **R**.

	Site			
	Head	Trunk	Extrem	Total
Tumor type	& neck		-ities	
Hutchinson's melanotic freckle	22	2	10	34
Superficial spreading melanoma	16	54	115	185
Nodular	19	33	73	125
Indeterminate	11	17	28	56
Total	68	106	226	400

The table represents a **contingency table**. Let's call the probability of being in cell (j, k) θ_{jk} . Then, in case of no association

$$\theta_{jk} = \theta_{j.}\theta_{.k}$$
 $j = 1, \dots, J$ and $k = 1, \dots, K$ (3.4.2)

i.e, in the case of independence

$$\log \mathbb{E}(Y)_{jk} = \log n + \log \theta_{j.} + \log \theta_{.k} \tag{3.4.3}$$

which can be compared with the **dependent** model, i.e.

$$\log \mathbb{E}(Y)_{jk} = \log n + \log \theta_{jk} \tag{3.4.4}$$

Analogously to the ANOVA model, we should *introduce the factors relative to the single predictors/factors*

$$\log \mathbb{E}(Y)_{ik} = \mu + \alpha_i + \beta_k + (\alpha \beta)_{ik}$$
(3.4.5)

where $(\alpha\beta)_{jk}$ represents a coefficient relative to the *interaction term*; therefore, to test for independence we can compare (3.4.5) with

$$\log \mathbb{E}(\mathbf{Y})_{ik} = \mu + \alpha_i + \beta_k \tag{3.4.6}$$

or with the minimal model

$$\log \mathbb{E}(Y)_{jk} = \mu \tag{3.4.7}$$

The specification of log-linear models is **hierarchical**: if the higher-order term (interaction) is included in the model, all the lower-order terms are included as well.



Warning: This means that, in many cases, log-linear models have many parameters: **constraints may be needed!**

While several distributions can be used (**Think:** Can you think of anyone?), **Poisson distributions can be assumed**. Therefore, all standard methods for GLM can be applied (weighted least squares, goodness-of-fit statistics like P^2 and D, Pearson and deviance residuals).

2.1 The saturated model

The **saturated model** is given by

```
library(dobson)

data("melanoma")

ressat.melanoma <- glm(frequency ~ site*type, family=poisson(), data=melanoma)
summary(ressat.melanoma)</pre>
```

2.2 The model with no interactions

The model with **no interaction terms** is given by

```
library(dobson)

data("melanoma")

resadd.melanoma <- glm(frequency ~ site + type, family=poisson(), data=melanoma)
summary(resadd.melanoma)</pre>
```

2.3 The minimal model

The **minimal** is given by

```
library(dobson)

data("melanoma")

resmin.melanoma <- glm(frequency ~ 1, family=poisson(), data=melanoma)
summary(resmin.melanoma)</pre>
```

2.4 Expected frequencies

For the reference category type: Hutchinson's melanotic freckle on site: extremities the expected frequencies are

• minimal model: $e^{3.507} = 33.33$

ullet additive model: $e^{2.9554}=19.21$

ullet saturated model: $e^{2.3026}=10.00$

Note: the expected frequencies for the saturated model correspond to the observed frequencies.

For type:indeterminate tumours on site:head-neck the expected frequencies are

• minimal model: $e^{3.507} = 33.33$

ullet additive model: $e^{2.9554-1.2010+0.499}=9.520049$

ullet saturated model: $e^{2.3026+0.7885+1.0296-1.7228}=11.000$

Again the expected frequencies for the saturated model correspond to the observed frequencies.

For type:nodular tumours on site:trunk the expected frequencies are

• minimal model: $e^{3.507} = 33.33$

ullet additive model: $e^{2.9554-0.7571+1.3020}=e^{3.5003}=33.12$

ullet saturated model: $e^{2.3026-1.6094+1.9879+0.8155}=e^{3.4966}=33.00$