3.4 Log-linear regression

1. Introduction

The term log-linear models is used to describe any model of the form

$$\log \mathbb{E}(\mathbf{Y}_i) = c + \mathbf{x}_i^{\top} \boldsymbol{\beta}$$
 (3.4.1)

In this section, we will analyse the introduction of interaction terms and the analogous of the ANOVA for log-linear models.

2. Modelling example

Example:

Let's consider a cross-sectional study of patients with a form of skin cancer called "malignant melanoma". We have 400 patients, with information about the site of the tumour and its histological type. The data are collected into the melanoma dataset available in the dobson package in R.

	Site			
	Head	Trunk	Extrem	Total
Tumor type	& neck		-ities	
Hutchinson's melanotic freckle	22	2	10	34
Superficial spreading melanoma	16	54	115	185
Nodular	19	33	73	125
Indeterminate	11	17	28	56
Total	68	106	226	400

The table represents a **contingency table**. Let's call the probability of being in cell (j,k) θ_{jk} . Then, in case of no association

$$\theta_{jk} = \theta_{j.}\theta_{.k} \qquad j = 1, \dots, J \quad \text{and} \quad k = 1, \dots, K$$
 (3.4.2)

i.e, in the case of independence

$$\log \mathbb{E}(Y)_{jk} = \log n + \log \theta_{j.} + \log \theta_{.k}$$
(3.4.3)

which can be compared with the **dependent** model, i.e.

$$\log \mathbb{E}(Y)_{jk} = \log n + \log \theta_{jk} \tag{3.4.4}$$

Analogously to the ANOVA model, we should *introduce the factors relative to the single predictors/factors*

$$\log \mathbb{E}(\mathbf{Y})_{ik} = \mu + \alpha_i + \beta_k + (\alpha \beta)_{ik}$$
(3.4.5)

where $(\alpha\beta)_{jk}$ represents a coefficient relative to the *interaction term*; therefore, to test for independence we can compare (3.4.5) with

$$\log \mathbb{E}(Y)_{jk} = \mu + \alpha_j + \beta_k \tag{3.4.6}$$

or with the minimal model

$$\log \mathbb{E}(Y)_{jk} = \mu \tag{3.4.7}$$

The specification of log-linear models is **hierarchical**: if the higher-order term (interaction) is included in the model, all the lower-order terms are included as well.



Warning: This means that, in many cases, log-linear models have many parameters: **constraints may be needed!**

While several distributions can be used (**Think:** Can you think of anyone?), **Poisson distributions can be assumed**. Therefore, all standard methods for GLM can be applied (weighted least squares, goodness-of-fit statistics like P^2 and D, Pearson and deviance residuals).

2.1 The saturated model

The **saturated model** is given by

```
library(dobson)

data("melanoma")

ressat.melanoma <- glm(frequency ~ site*type, family=poisson(), data=melanoma)
summary(ressat.melanoma)</pre>
```

2.2 The model with no interactions

The model with **no interaction terms** is given by

```
library(dobson)

data("melanoma")

resadd.melanoma <- glm(frequency ~ site + type, family=poisson(), data=melanoma)
summary(resadd.melanoma)</pre>
```

2.3 The minimal model

The **minimal** is given by

```
library(dobson)

data("melanoma")

resmin.melanoma <- glm(frequency ~ 1, family=poisson(), data=melanoma)
summary(resmin.melanoma)</pre>
```

2.4 Expected frequencies

For the reference category type: Hutchinson's melanotic freckle on site: extremities the expected frequencies are

 $\bullet \ \ \text{minimal model:} \ e^{3.507}=33.33$

• additive model: $e^{2.9554} = 19.21$

• saturated model: $e^{2.3026} = 10.00$

Note: the expected frequencies for the saturated model correspond to the observed frequencies.

For type:indeterminate tumours on site:head-neck the expected frequencies are

• minimal model: $e^{3.507} = 33.33$

ullet additive model: $e^{2.9554-1.2010+0.499}=9.520049$

• saturated model: $e^{2.3026+0.7885+1.0296-1.7228} = 11.000$

Again the expected frequencies for the saturated model correspond to the observed frequencies.

For type:nodular tumours on site:trunk the expected frequencies are

 $\bullet \ \ {\rm minimal\ model}; e^{3.507}=33.33$

ullet additive model: $e^{2.9554-0.7571+1.3020}=e^{3.5003}=33.12$

ullet saturated model: $e^{2.3026-1.6094+1.9879+0.8155}=e^{3.4966}=33.00$