# Scalable Nonparametric Sampling from Multimodal Posteriors with the Posterior Bootstrap Edwin Fong, Simon Lyddon, Chris Holmes

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Background

Overview

3 Example

4 Details

5 Example

6 Conclusions

### Dirichlet Distribution, definition

Dirichlet distribution: multivariate generalisation of the Beta distribution.

pdf is given by

$$f(x;\alpha) = \frac{1}{B(\alpha)} \prod_{i=1}^{p} x_i^{\alpha_i - 1}$$
 (1)

where  $x \in \mathbb{R}^p$ ,  $\sum_i x_i = 1$  and  $\alpha \in \mathbb{R}^p$ ,  $a_i > 0$ , and  $B(\alpha)$  is a normalisation term.

## Dirichlet Distribution, visualisation

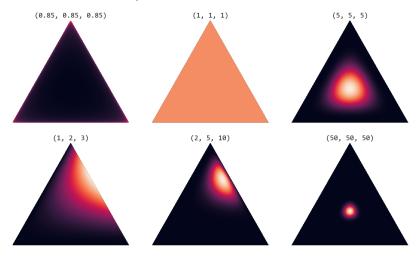


Figure: source:

https://towards datascience.com/dirichlet-distribution-a 82 ab 942 a 879

#### Dirichlet Processes, definition

**Dirichlet Process:** a stochastic process where a finite subset of random variables have a Dirichlet distribution.

Dirichlet processes are specified by a base probability distribution, H and a concentration parameter  $\alpha$ . Then for some finite disjoint partition of  $S = \{B\}_{i=1}^n$  we have

$$(X_{B_1}, \dots, X_{B_n}) \sim \text{Dirichlet}(\alpha H(B_1), \dots, \alpha H(B_n))$$
 (2)

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#### Dirichlet Processes, some intuition

What is a die?



Figure: A Statistician's best friend (sorry coins).

#### Dirichlet Processes, some intuition

Sampling form a DP has the bi-product of a new probability distribution.

1. Partition a space

8,	G <sub>2</sub>
$\beta_3$	B

2. Compute some parameters

3. Sample some weights

$$(\times_{\beta_{1}}, \times_{\beta_{2}}, \times_{\beta_{3}}, \times_{\beta_{3}})$$

Questions / Comments?

And then we're onto the paper!

Let  $y = (y_1, \dots, y_n)$  where  $y_i \stackrel{iid}{\sim} F_0$  and let  $\theta \in \Theta \subseteq \mathbb{R}^p$  be a parameter that indexes a family of probability distribution  $\mathcal{F}_{\Theta}$ .

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Then we are interested in

$$\theta_0(F_0) = \underset{\theta}{\arg\min} \int I(y;\theta) dF_0(y) \tag{3}$$

where  $I(y; \theta)$  is a loss function. For example if  $I = (y - \theta)^2$  we would recover the mean.

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Put a prior on it

Putting a DP prior on  $F_0$ 

$$F|a,F_{\pi} \sim DP(a,F_{\pi}) \tag{4}$$

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where

$$G_n = \frac{a}{a+n} F_{\pi} + \frac{1}{a+n} \sum_{i} \delta_{y_i}$$
 (6)

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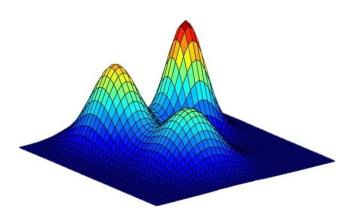
We can sample from  $\pi(\theta|y_{1:n})$  using the following algorithm

# Algorithm 1 NPL Posterior Sampling

$$\begin{array}{l} \text{for } i=1 \text{ to } B \text{ do} \\ \operatorname{Draw} F^{(i)} \sim \operatorname{DP}(\alpha+n,G_n) \\ \theta^{(i)} = \arg \min_{\theta} \int l(y,\theta) dF^{(i)}(y) \\ \text{end for} \end{array}$$

#### Gaussian Mixture Models

#### Weighted mixture of Gaussians



To use Bayesian Non-parametric learning all we need to do is define a loss function

$$I(y, \pi, \mu, \sigma) = -\log \sum_{k=1}^{K} \pi_k \mathcal{N}(y; \mu_k, \operatorname{diag}(\sigma_k^2))$$
 (8)

GMMs have multi-modal posteriors.

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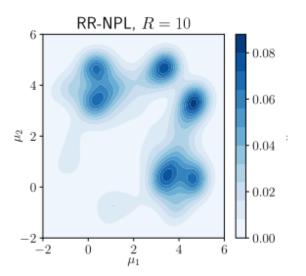
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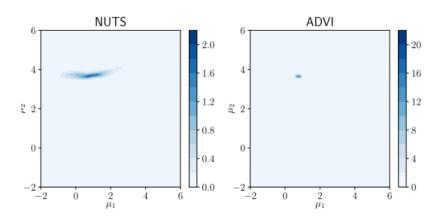
For the example, there are K = 3 groups, with

$$\pi = (0.1, 0.3, 0.6)$$

$$\mu_0 = (0, 2, 4)$$

$$\sigma_0^2 = (1, 1, 1)$$





#### Gaussian Mixture Models, discussion

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**Yes!** We now know the problem is multi-modal... and can quantify the uncertainty about those modes

**No!** We might not care about those modes, but getting something useful out of the posterior. It's not really clear how we can do that.

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In practice we need to truncate somewhere But that means we're going to be approximating

$$\pi(\theta|y_{1:n}) = \int \pi(\theta|F) d\pi(F|y_{1:n}) \tag{9}$$

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# Approximate posterior sampling

#### Algorithm 2 Posterior Bootstrap Sampling

Define T as truncation limit

Observed samples are  $y_{1:n}$ 

for 
$$i = 1$$
 to  $B$  do

Draw prior pseudo-samples  $\tilde{y}_{1:T}^{(i)} \stackrel{iid}{\sim} F_{\pi}$ Draw  $(w_{1:n}^{(i)}, \tilde{w}_{1:T}^{(i)}) \sim \text{Dir}(1, \dots, 1, \alpha/T, \dots, \alpha/T)$ 

$$\theta^{(i)} = \arg\min_{\theta} \left\{ \sum_{j=1}^{n} w_{j}^{(i)} l(y_{j}, \theta) + \sum_{k=1}^{T} \tilde{w}_{k}^{(i)} l(\tilde{y}_{k}^{(i)}, \theta) \right\}$$

#### end for

# What if I want to quantify uncertainty about a mode?

Randomness is introduced through the weights and psuedo-samples. Fixing our starting point  $\theta_i$  allows us to explore the area around a mode

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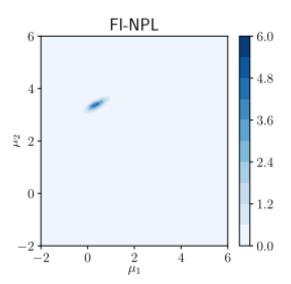
#### **Algorithm 4** FI-NPL Posterior Sampling

```
Select \theta^{\text{init}} from mode of interest for i=1 to B do Draw F^{(i)} \sim \mathrm{DP}(\alpha+n,G_n) \theta^{(i)} = \operatorname{local} \operatorname{arg\,min}_{\theta} \left( \int l(y,\theta) dF^{(i)}(y), \theta^{\text{init}} \right) end for
```

#### Back to GMMs



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# Example #2

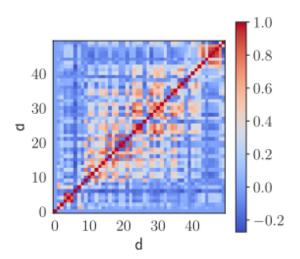
Analysis of genotype / psuedo phenotype dataset, where the phenotype was generate by

$$y_i \sim \mathsf{Bernoulli}(\sigma(\beta^\top x_i))$$
 (10)

where  $\beta \in \mathbb{R}^{50}$  with 5 randomly selected non-zero components.

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#### Correlation Matrix



We can quantify uncertainty about parameters, but all this happens through our loss function.

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So all we need to do is have a sparsity inducing loss function.

$$I(y,\theta) - \log f_{\theta}(y) + \gamma g(\theta) \tag{11}$$

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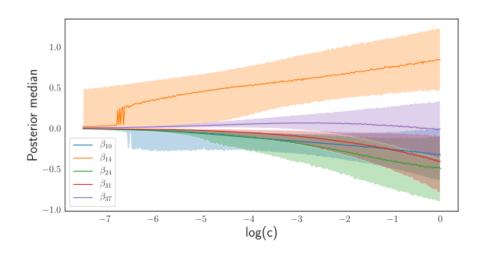
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Example: setting  $g(\theta) = |\theta|$  gives the Bayesian NPL-Lasso

# Example #2, results



#### Conclusions

A cool idea for uncertainty quantification.

Priors aren't really a thing, it's all done via the loss function

Captures multi-modality

Quick if we have a lot of compute, sampling from the NPL posterior can be done in parallel

Thanks for Listening

Questions?