

Structured Probabilistic Models and Monte Carlo

Deep Learning Decal

Hosted by Machine Learning at Berkeley



Agenda

Background

Algorithms

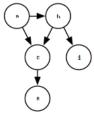
Questions

Background

Purpose of Structured Probabilistic Models



- Split a complicated probability distribution function into a function made of simpler functions based on specific subsets of random variables.
- Random variables a \sim N(0, 1), b \sim a + N(1, 2), c \sim a + b + N(2, 3), d \sim b + N(3, 4), e \sim c + N(4, 5).



Purpose of Structured Probabilistic Models

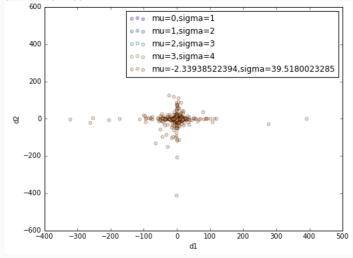


- Ex) p(a, b, c, d) = p(a)p(b|a)p(c|a, b)p(d|b)p(e|c).
- Normalize the distribution to sum to 1. z= normalizing constant. $p(x)=\frac{1}{z}\prod_i\phi^{(i)}(C^{(i)})$ where ϕ is a chosen function.

Purpose of Structured Probabilistic Models



unnormalized



Structured Probabilistic Models for Deep Learning



- Uses: density estimation, denoising, Missing value imputation, sampling.
- Cons: memory, statistical efficiency, runtime (cost of inference), runtime (cost of sampling).
- A distribution has few parameters if the random variable (node) has few parents in the graph.

Monte Carlo Methods



- Randomized algorithms that return answers with a random amount of error. Increasing resources (ex: time, memory) can decrease error.
- Used in sampling for approximations, speedup.

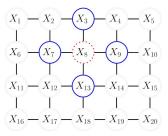
Algorithms

Undirected Models



- Markov random fields (MRFs) or Markov Networks: For when there is a bidirectional relationship (or lack of unidirectional relationship) between nodes.
- Undirected notation includes: -; <->; a->b, b->a.

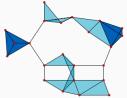
Markov Random Field (Markov Network)



Undirected Models



- An unnormalized probability distribution can be represented by each clique ${\bf C}$ in graph ${\bf G}$. Let clique potential be a factor $\phi({\bf C})$. Clique potential measures the likelihood of each random variable's state for every configuration (joint state) in the clique.
- $\tilde{p}(x) = \prod_{\mathbf{C} \in \mathbf{G}} \phi(\mathbf{C}).$
- Gibbs distribution: $p(x) = \frac{1}{Z}\tilde{p}(x)$. $Z = \int \tilde{p}(x)dx$



Approximating Z



- It is possible to specify variables st Z doesn't exist. This happens if $Z = \int \tilde{p}(x) dx$ diverges.
- Conditions for divergence:
 - 1. Is an improper integral: a definite integral that has either or both limits infinite or an integrand that approaches infinity at one or more points in the range of integration.
 - 2. The limit does not exist or it is infinite.

Approximating Z



- The domain of each variable can significantly change the type of probability distribution in a clique.
- Ex)

n-dimensional rv x.

one clique for each element $x_i \in x \forall i = 1, ..., n$, where each element is mutually independent.

Domain 1: $x \in \{0,1\}^n$. $p(x) = \sum_{i=1}^{2n} p(x = pattern_i)$.

Domain 2: set of n dimensional basis bectors.

$$p(x) = \sum_{i=1}^{n} p(x_i = 1).$$

Energy-Based Models



• Assume $\forall x, \tilde{p}(x) > 0$. $\tilde{p}(x) = exp(-E(x))$ where E(x) is the energy function.

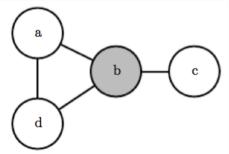
Each clique has an
$$\phi(x)$$
 function,
$$p(x) = \prod exp(-E(x)) = \prod \phi(-E(x))$$

- By exp(a)exp(b) = exp(a+b), $p(x) = exp(\sum -E(x)) = \phi(\sum -E(x))$.
- Also called a product of experts.
- $\mathbf{F}(x) = -\log \sum_{h} \exp(-E(x, h)).$

Separation and D-Separation



dependence-separation



- Given B, A and C are separate (implied-not guaranteed independent). A and D are still connected (implied dependent).
- Choosing graph structure (ex: directed or undirected) depends on the problem.

Factor Graphs



- A graphical representation of an undirected model that consists of a bipartite undirected graph.
- Clarifies the undirected graph.
- Rules:

May be connected with undirected edges.

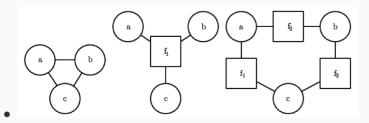
A variable and a factor are connected iff the variable is an argument to the factor in the unnormalized probability distribution.

A factor cannot be connected to any factor.

A variable cannot be connected to a variable.

Factor Graphs





• Left: undirected graph. Middle and Right: different configurations of the same undirected graph.

Sampling from Graphical Models



- Ancestral sampling: Topological sort variables x_i st for all i and j, j i i if x_i is a parent of x_j . Then sample in order.
- Ex: Let Pa_G(x_i) be the set of random variables (already sampled) that are the ancestors of random variable x_i.

$$\begin{bmatrix} x_1 \sim P(x_1) \\ x_2 \sim P(x_2 | Pa_G(x_2) \\ \dots \\ x_n \sim P(x_n | Pa_G(x_i)) \end{bmatrix}$$

Sampling from Graphical Models



- Cons: Only for directed graphs, does not support every conditional sampling operation.
- Gibbs sampling: for undirected graphs. Let x be an n-dimensional vector where each x_i is a node in the graph.
 Iteratively visit each x_i and sample with condition: p(x_i|x_{i-1}).

Pros of Structured Modeling



- Reduce cost of learning, representing, and inference from probability distributions.
- \bullet Explicit representation of learning and inference \to Easier to develop and debug.

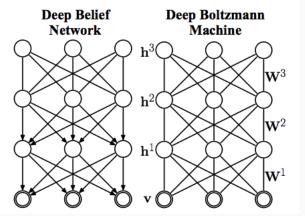
dels

- Let latent variable h_i be at depth j if the shortest path from h_i to an observed variable is j steps. Depth of a model is now max j.
- Use distributed representations. More latent variables.
- Latent variables are learned aren't assigned specific semantics ahead of time. Many latent variables, so algorithms should be efficient.
- Many connections between units.

Restricted Boltzmann Machine



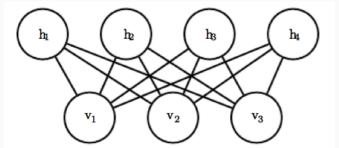
- Energy-based model with binary units. Has a single layer of latent variables to learn a representation for the input.
- Energy function: $E(v, h) = -b^T v c^T h v^T Wh$ where b, c, W are unconstrained, real, and learnable.



Restricted Boltzmann Machine



- Properties: $p(h|v) = \prod_i p(h_i|v), p(v|h) = \prod_i p(v_i|h).$
- Consequently, $P(h_i = 1|v) = \sigma(v^T W_{:,i} + b_i)$, $P(h_i = 0|v) = 1 \sigma(v^T W_{:,i} + b_i)$ where σ is a probability distribution function.



• $E_{h\sim p(h|v)}[h]$ is the set of features for v.

Monte Carlo Sampling



 To approximate a sum or integral, view it as an expectation under a distribution. Approximate the expectation by a corresponding average.

•

$$s = \sum_{x} p(x)f(x) = E_{p}[f(x)]$$

,

$$s = \int p(x)f(x)dx = E_p[f(x)]$$

where s is what we're estimating.

Monte Carlo Sampling



• Draw n samples $x^{(1)},...,x^{(n)}$ from p, then calculate the average.

$$\widehat{s}_n = \frac{1}{n} \sum_{i=1}^n f(x^{(i)})$$

- This is a reasonable approximation because:
 - 1. \hat{s}_n is unbiased. $E[\hat{s}_n] = \frac{1}{n} \sum_{i=1}^n E[f(x^{(i)})] = \frac{1}{n} \sum_{i=1}^n s = s$.
 - 2. By Law of Large Numbers: \hat{s}_n will converge to s with $n \to \inf$ samples given that the samples are iid and the variance of each $x^{(i)}$ is bounded.

Importance Sampling



 When we can't sample from p(x). Decide which part of integrand is the distribution p(x) and which part is the value f(x). We calculate expectation as:

•

$$p(x)f(x) = q(x)\frac{p(x)f(x)}{q(x)}$$

• where we sample from p(x) and average $\frac{pf}{q}$.

Importance Sampling



Transform the Monte Carlo estimator

$$\hat{s}_p = \frac{1}{n} \sum_{i=1, x^{(i)} \sim p} f(x^{(i)}) \text{ into } \hat{s}_q = \frac{1}{n} \sum_{i=1, x^{(i)} \sim q} \frac{p(x^{(i)}) f(X^{(i)})}{q(x^{(i)})}.$$

- $E_q[\hat{s}_q] = E_q[\hat{s}_p] = s$.
- $Var[\widehat{s}_q] = Var[\frac{p(x)f(x)}{q(x)}]/n$. Minimum variance at: $q^*(x) = \frac{p(x)|f(x)|}{Z}$ where Z is the normalization constant.

Monte Carlo: Biased importance sampling



• Does not require normalized p or q.

•
$$\widehat{s}_{BIS} = \frac{\sum_{i=1}^{n} \frac{p(x^{(i)})}{q(x^{(i)})}(x^{(i)})}{\sum_{i=1}^{n} \frac{p(x^{(i)})}{q(x^{(i)})}} = \frac{\sum_{i=1}^{n} \frac{\widehat{p}(x^{(i)})}{\widehat{q}(x^{(i)})} f(x^{(i)})}{\sum_{i=1}^{n} \frac{\widehat{p}(x^{(i)})}{\widehat{q}(x^{(i)})}} \text{ where } \widehat{p}, \widehat{q} \text{ are }$$
 unnormalized p, q respectively. $x^{(i)}$ are the samples from q.

Choice of the q distribution affects efficiency.

Monte Carlo: Biased importance sampling



- With high-dimensional x:
 - 1. q will not match to p or p|f| (collect useless samples). $q(x^{(i)}) \gg p(x^{(i)})|f(x^{(i)})$.
 - 2. q will "match" to p or p|f| too well (extreme overestimation). $q(x^{(i)}) \ll p(x^{(i)})|f(x^{(i)})$.



- Use Markov Chains to estimate with Monte Carlo.
- Example model: energy-based model (EBM) $p(x) \propto exp(-E(x))$. Drawing samples from EBM is difficult because it can be an undirected graph $(b \rightarrow a, a \rightarrow b)$.
- Soln: Use a Markov chain: random state x and a transition distribution T(x'|x) which is the probability that x will go to state x'.



• Have infinitely many MC run in parallel. Let t = time steps. The states of each MC is from distribution $q^{(t)}(x)$. $q^{(0)}$ is the initial distribution. $q^{(t)}$ inferred from all MC steps so far, so $q^{(t)}$ to converge to p(x).



- Ex: Positive integer rv x, probability distribution q, vector v. $q(x=i)=v_i$. Update to state x'. Probability of landing on state x' is $q^{(t+1)}(x')=\sum_x q^{(t)}(x)T(x'|x)$. Define matrix A st it describes the probability of state x to every state in the graph. $A_{i,j}=T(x'=i|x=j)$.
- Update rule for all MC run in parallel: $v^{(t)} = Av^{(t-1)} = A^tv^{(0)}$
- A is a stochastic matrix.



- A Markov chain with transition operator T will converge under mild conditions to the point: $q'(x') = E_{x \sim q} T(x'|x)$.
- Deep Learning uses multiple MCs.
- Cons: MCs are expensive because of the time required to reach equilibrium (converge). We don't know how long it takes to converge.

Gibbs Sampling



- Ensures that q(x) is a useful distribution by deriving T(x' x) from the learned $p_{model}(x)$.
- Block Gibbs Sampling: Pick one variable x_i and sample it from
 p_{model} conditioned on its neighbors in undirected graph G. G
 defines the structure of the EBM p_{model} in this example.

Challenge of Mixing between Separated Modes

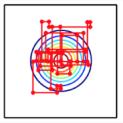


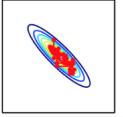
- Ideally, samples from an MC to sample from p(x) would be iid.
 MCMC samples become very correlated in high dimensions.
- In terms of an EBM, T = energy barrier from x to x. If probability x to x' low → high energy barrier. So when multiple modes with high probability are separated by nodes of low probability, then converge slowly.

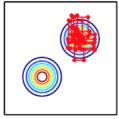
Challenge of Mixing between Separated Modes



Ex:







- Left: Multivariate normal, 2 independent variables.
- Center: Multivariate normal, 2 highly correlated variables.
- Right: Mixture of Gaussians with very far modes that are not axis aligned.

Tempering to Mix between Modes



- Based on making a smoother version of the target distribution.
- Ex: EBM with a smoothed distribution: $p_{\beta}(x) \propto exp(-\beta E(x))$ where E is the energy function, and β is the parameter to tune smoothness. Temper with $\beta < 1$.
- Parallel Tempering: MC simulates different states in parallel at different parameters.
- Caveat: Sometimes T has to be a small probability in order to converge.

Depth May Help Mixing



- Let latent variable model p(h, x). Mixing will be poor if p(h|x) encodes x too well.
- A soln: Let h be a deep representation such that p(h|x) encodes x less strict.
- Ex: Autoencoders and RBMs usually have a marginal distribution over h that is more uniform and unimodal than x's distribution.

Monte Carlo Pi Example



Animation

 $\verb|https://www.youtube.com/watch?v=VJTFfIq04TU\&t=25s|$

Paper: Kinetic Monte Carlo Method

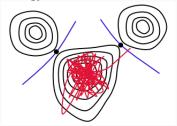


- Simulating atom movement. Total simulation time is usually less than one microsecond, which may not be useful for observing longer processes such as diffusion.
- Kinetic Monte Carlo tries to solve this time-scale problem by modeling state to state transitions in a Markov Model.

Kinetic Monte Carlo Method - Funny Contour Diagram



Energy-Barrier-Limited Infrequent-Event System



Contour plot of the potential energy between states.

 Memoryless System: The transition probabilities (escape rate) for exiting state i have nothing to do with the history prior to entering state i.

Kinetic Monte Carlo Method - KMC Procedure



- In state i
- Set of pathways and associated rate constants kij
- We have the distribution for each pathway's first escape time.
- ullet Exponentially distributed time t_j to transition.
- Find the pathway j_{min} with the lowest t_j , record only this step, advance the overall system clock by j_{min} .

Kinetic Monte Carlo Method - KMC Procedure (common)

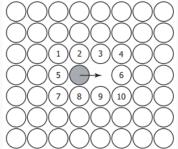
- BKL or "n-fold way" algorithm:
- Have an object with length equal to rate constant k_{ij} for pathway i to j for all M escape pathways.
- Put all objects end to end, with total length k_{tot} .
- Choose 1 random position along this length, picking the pathway for the system to follow.
- Advance the clock by drawing a time from an exponential distribution with rate constant k_{tot}.

Kinetic Monte Carlo Method - Harmonic Transition State ory



 Transition pathway has a saddle point on the potential energy surface (contours of the distribution of atoms leaving a state).

Rate Catalog:



 Caveat: The real dynamics of a system can have unexpected and complex reaction pathways that cannot occur during a KMC simulation.

Kinetic Monte Carlo Method - Problems



- A low barrier can significantly decrease simulation time.
- The system repeatedly visits a subset of states.
- Enumerating all states is computationally expensive.

Kinetic Monte Carlo Method - Optimizing



- Object kinetic Monte Carlo
 - -Higher level simulation.
 - -Create state definitions and rate constants for atom clusters.
 - -Pro: Can run for more time.
 - -Con: Can miss important pathways with less detail. Similarly, relationships between states for clusters can become much more complex because of the missing (simple) states for atoms.
- F U T U R EEEEEE: On-the-fly Kinetic Monte Carlo

Paper: Monte Carlo Strength Evaluation



- Simulate guessing attacks. Goal is to increase effort it takes for attackers to break a password.
- Sampling from a model to estimate password strength.



- Password Guessing
 - -Dictionary, rainbow chains (efficiently memorize very large set of pre-computed password hashes)
 - -Probabilistic attacks reduce number of password guesses. Ex: n-gram, PCFG, backoff models.
- Some Defenses: Salting, password strengthening, password stretching (hashing passwords using computationally expensive functions)

Monte Carlo Strength Evaluation - Compute Password Strength

 Problem: Strength meters generally based on heuristics that don't accurately reflect a password's resistance to guessing attacks.



• N-gram model: probability of a password $c_1...c_l$ is

$$p_{n-gram}(c_1...c_l) = \prod_{i=1}^{l+1} P(c_i|c_{i-n+1}...c_{i-1})$$

Algorithm 1 Password generation for n-grams.

```
def starting_state():
    return "\lefta \ldots \ldots" with length n-1
def update_state(s,t):
    drop the first character from s
    return s+t # concatenation
s \leftarrow \text{starting\_state()}
g \leftarrow \text{"" # accumulator for the result}
while True:
    r \leftarrow \text{random number in } [0,1]
# find i through binary search
i \leftarrow \text{rightmost index s.t. } C_s[i] > r
if t_{s,i} = \bot: return g
append t_{s,i} to g
s \leftarrow \text{update\_state}(s,t_{s,i})
```



- Probabilistic Context-Free Grammars (PCFGs): passwords grouped by templates.
- Ex: Make a probabilistic model out of template: L_3D_3 (3 letters followed by 3 digits).

The probability of "abc123" is:

$$P_{PCFG}("abc123") = P(L_3D_3)P("abc"|L_3)P("123"|D_3)$$



ullet Backoff Model: By Katz. Addresses sparsity by considering n-grams of all lengths, and discarding those with less occurrences than a threshold au.

Algorithm 2 Sample creation for the backoff model.

```
def starting_state():
    if using the start symbol: return "\perp"
    else: return ""

def update_state(s,t):
    append t to s
    while o(s) < \tau:
        drop the first character in s

Run Algorithm 1 using these functions.
```



• Given a threshold τ , the probability of a single character \hat{c} is its frequency in the training set:

$$p_{b0}(\hat{c}) = \frac{o(\hat{c})}{\sum_{c} o(c)}$$

where $o(c_1...c_n)$ is the number of occurrences of the $c_1...c_n$ string in the training set.

$$p_{bo}(c_1...c_{n+1}) = p_{b0}(c_1...c_n)P(c_{n+1}|c_1...c_n)$$

where

$$P(c|c_1...c_n) = \begin{cases} \frac{o(c_1...c_nc)}{o(c_1...c_n)} \text{if } o(c_1...c_nc) \geqslant \tau \\ P(c|c_2...c_n)\tau(c_1...c_n) \text{otherwise} \end{cases}$$

and

$$r(c_1...c_n) = \sum_{c:o(c_1...c_nc) \geqslant \tau} \frac{o(c_1...c_nc)}{o(c_1...c_n)}$$

Supplemental Reading: Latent Dirichlet Allocation



 Probabilistic Topic Models by David M. Blei http://www.cs.columbia.edu/~blei/papers/Blei2012.pdf



- Topic Reference: https://github.com/HFTrader/DeepLearningBook/blob/master/Dee
- Paper 1: http://www.phys.ubbcluj.ro/ zneda/edu/mc/mc4.pdf
- Paper 2: http://www.eurecom.fr/en/publication/4711/download/rs-publi-4711.pdf
- Paper 3: http://www.cs.columbia.edu/~blei/papers/Blei2012.pdf

Questions

Questions?!