

Geometric Brownian Motion Notes

1 Summary

A security that follows a geometric brownian motion is a martingale. This assures that the probabilities concerning the ratio of the price a time t in the future to the present price will not depend on the present price.

There is a simple derivation to μt and $\sigma^2 t$ by letting Δ denote a small increment of time, and at every Δ units, the price of a security changes only at times that are integral multiples of Δ .

Geometric Brownian motion is a variant of a model known as Brownian motion. Both processes share the martingale property, however, the difference is that in Brownian motion, it is the **differences in prices** and not the **logarithm of their ratio** that has a normal distribution.

A flaw in the Brownian motion model is that it is possible to become negative, since it is a normal random variable. Also it is not completely reasonable to assume that a price difference over an interval of fixed length has the same normal distribution no matter what the price at the beginning of the interval is.

Geometric Brownian motion eliminates these flaws since it is now the *logarithm* of the stock's price that is a normal random variable, and thus cannot be negative. The ratio of prices separated by a fixed length of time is also a more reasonable assumption, since it is the *percentage* change in price, rather than the absolute change, whose probabilities do not depend on the present price.

2 Geometric Brownian Motion

Suppose we are interested in the price of a security as it evolves over time. Let the present time be time 0, and let $S(t)$ denote the price of this security a time t from the present.

This collection of prices $S(t)$, $0 \leq t < \infty$, follows a *geometric Brownian motion* with drift parameter μ and volatility parameter σ if, for all nonnegative values of t and the increment of Δt , the random variable

$$\frac{S(\Delta t + t)}{S(t)}$$

is independent of all prices up to time t ; and if

$$\log\left(\frac{S(\Delta t + t)}{S(t)}\right)$$

is a normal random variable with mean $\mu\Delta t$ and variance $\sigma^2\Delta t$

The series of prices will be a geometric Brownian motion if the **ratio** of the price a time t in the future to the present price will, independent of the past history of prices, have a lognormal probability distribution with parameters $\mu\Delta t$ and $\sigma^2\Delta t$

In other words, if we assume a security to follow a geometric Brownian motion, once μ and σ are determined, only the present price - and not the history of past prices - affects probabilities of future prices. This is also known as a **martingale**.

Furthermore, for an initial given price $S(0)$, the expected value of the price at time t depends on *both* of the geometric Brownian motion parameters, that is, μ and σ .

Specifically, if the initial price is s_0 , then

$$\mathbb{E}[S(t)] = s_0 e^{t(\mu + \frac{\sigma^2}{2})}$$

3 Geometric Brownian Motion as a Limit of Simpler Models

Let Δ denote a small increment of time and suppose that, every Δ time units, the price of a security either goes up by a factor u with probability p or goes down by the factor d with probability $1 - p$, where

$$u = e^{\sigma\sqrt{\Delta}}, \quad d = e^{-\sigma\sqrt{\Delta}}$$

$$p = \frac{1}{2} \left(1 + \frac{\mu}{\sigma} \sqrt{\Delta} \right)$$

That is, we're supposing that the price of the security changes only at times that are integral multiples of Δ ; at these times, it either goes up by the factor u or down by the factor d .

Let us verify that this model becomes geometric Brownian motion as we let $i\Delta$ become smaller and smaller. To begin, let Y_i equal 1 if the price goes up at time $i\Delta$, and let it be 0 if it goes down.

The number of times that the security's price goes up in the first n increments is $\sum_{i=1}^n Y_i$, and the number of times it goes down is then, $n - \sum_{i=1}^n Y_i$.

Hence, $S(n\Delta)$ its price at the end of this time, can be expressed as

$$S(n\Delta) = S(0)u^{\sum_{i=1}^n Y_i}d^{n-\sum_{i=1}^n Y_i}$$

or

$$S(n\Delta) = d^n S(0) \left(\frac{u}{d}\right)^{\sum_{i=1}^n Y_i}$$

If we now let $n = \frac{t}{\Delta}$, then the preceding equation becomes

$$\begin{aligned} \frac{S(t)}{S(0)} &= d^{\frac{t}{\Delta}} \left(\frac{u}{d}\right)^{\sum_{i=1}^{\frac{t}{\Delta}} Y_i} \\ \log\left(\frac{S(t)}{S(0)}\right) &= \frac{t}{\Delta} \log(d) + \log\left(\frac{u}{d}\right) \sum_{i=1}^{\frac{t}{\Delta}} Y_i \end{aligned}$$

$$\frac{-t\sigma}{\sqrt{\Delta}} + 2\sigma\sqrt{\Delta} \sum_{i=1}^{\frac{t}{\Delta}} Y_i \tag{1}$$

Moreover, as Δ goes to 0, there are more and more terms in the summation $\sum_{i=1}^{t/\Delta} Y_i$; hence, by the **Central Limit Theorem**, this sum becomes more and more normal, implying that from equation (1), $\log\left(\frac{S(t)}{S(0)}\right)$ becomes a normal random variable.