

Language use shapes cultural norms: Large scale evidence from gender

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Portions of this manuscript appeared in Lewis & Lupyan, 2018, Cog. Sci Proceedings.

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Abstract

Keywords: cultural norms, IAT, gender

Word count: X

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Introduction

The language we use to communicate a message shapes how our listener interprets that message (Loftus & Palmer, 1974; Tversky & Kahneman, 1981; Fausey & Boroditsky, 2010). A listener, for example, is more likely to infer that a person is at fault if the event is described actively (e.g., “she ignited the napkin”), as opposed to passively (e.g., “the napkin ignited”). The formative power of language is perhaps most potent in shaping meanings that necessarily must be learned from others: cultural norms. In the present paper, we consider one type of cultural norm—gender—and examine the extent to which differences in language use may lead to cross-cultural differences in understandings of gender.

Gender provides a useful case study of the relationship between language and thought for several reasons. First, more abstract domains like gender may be more subject to the influence of language relative to more perceptually grounded domains like natural kinds (Boroditsky, 2001). Second, many languages encode the gender of speakers and addressees explicitly in their grammar. Third, a large body of evidence suggests that language plays a key role in transmitting social knowledge to children (e.g., Master, Markman, & Dweck, 2012). And, fourth, gender norms are highly variable across cultures and have clear and important social implications.

For our purposes, we define the hypothesis space of possible relationships between language and gender norms with two broad extremes: (1) language reflects a pre-existing gender bias in its speakers (*language-as-reflection hypothesis*); (2) language causally influences gender biases (*language-as-causal hypothesis*). We assume that the language-as-reflection hypothesis is true to some extent: some of the ways we talk about gender reflect our knowledge and biases acquired independently of language. For example, we may observe that most nurses are women, and therefore be more likely to use a female

pronoun to refer to a nurse of an unknown gender. Our goal here is to understand the extent to which language may also exert a causal influence on conceptualizations of gender.

In particular, we explore two possible mechanisms by which the way we speak may influence notions of gender¹. The first is through the overt grammatical marking of gender, particularly on nouns, which is obligatory in roughly one quarter of languages (e.g., in Spanish, “nina” (girl) and “enfermera” (nurse) both take the gender marker *-a* to indicate grammatical femininity; Corbett, 1991). Because grammatical gender has a natural link to the real world, speakers may assume that grammatical markers are meaningful even when applied to inanimate objects that do not have a biological sex. In addition, the mere presence of obligatory marking of grammatical gender may promote bias by making the dimension of gender more salient to speakers.

A second route by which language may shape gender norms is via word co-occurrences. Words that tend to occur in similar contexts in language may lead speakers to assume—either implicitly or explicitly—that they have similar meanings. For example, statistically, the word “nurse” occurs in many of the same contexts as the pronoun “her,” providing an implicit link between these two concepts that may lead to a bias to assume that nurses are female. This second route may be particularly influential because the bias is encoded in language in a way that is more implicit than grammatical markers of gender and thus more difficult to reject.

An existing body of experimental work points to a link between language and psychological gender bias in both adults (e.g., Phillips & Boroditsky, 2003) and children (e.g., Sera, Berge, & Castillo Pintado, 1994). For example, Phillips and Boroditsky (2003) asked Spanish-English and German-English adult bilinguals to make similarity judgements between pairs of pictures depicting an object with a natural gender (e.g., a bride) and one without (e.g., a toaster). They found that participants rated pairs as more similar when the pictures matched in grammatical gender in their native language. While these types of studies

¹These mechanisms are what Whorf (1945) refers to as phenotypes (overt) and cryptotypes (covert).

provide suggestive evidence for a causal link between language and psychological gender bias, they are limited by the fact that they typically only compare speakers of 2-3 different languages and measure bias in a way that is subject to demand characteristics.

In what follows, we ask whether the way gender is encoded linguistically across 31 different languages predicts cross-cultural variability in a particular manifestation of a gender bias—the bias to associate men with careers and women with family. We begin in Study 1 by describing cross-cultural variability in psychological gender bias using an implicit measure. In Study 2, we use semantic-embedding models to examine whether variability in lexical semantics predicts variability in psychological gender biases. In Study 3, we ask whether the presence of grammatical gender in a language is associated with greater implicit gender bias. Together, our data suggest that both language statistics and language structure likely play a causal role in shaping culturally-specific notions of gender.

Description of Cross-Cultural IAT data

To quantify cross-cultural gender bias, we used data from a large-scale administration of an Implicit Association Task (IAT; Greenwald, McGhee, & Schwartz, 1998) by Project Implicit (Nosek, Banaji, & Greenwald, 2002). The IAT measures the strength of respondents' implicit associations between two pairs of concepts (e.g., male-career/female-family vs. male-family/female-career) accessed via words (e.g., “man,” “business”). The underlying assumption of the IAT is that words denoting more similar meanings should be easier to pair together compared to more dissimilar pairs.

Meanings are paired in the task by assigning them to the same response keys in a 2AFC categorization task. In the critical blocks of the task, meanings are assigned to keys in a way that is either bias-congruent (i.e. Key A = male/career; Key B = female/family) or bias-incongruent (i.e. Key A = male/family; Key B = female/career). Participants are then

presented with a word related to one of the four concepts and asked to classify it as quickly as possible. Slower reaction times in the bias-incongruent blocks relative to the bias-congruent blocks are interpreted as indicating an implicit association between the corresponding concepts (i.e. a bias to associate male with career and female with family).

In Study 1, we use the IAT to measure the bias to associate women with family and men with careers across different cultures. We find a gender bias in all countries. Replicating previous work (Miller, Eagly, & Linn, 2015), we also find that the magnitude of the bias is negatively correlated with percentage of female enrollment in STEM fields.

Methods

We analyzed an existing IAT dataset collected online by Project Implicit (<https://implicit.harvard.edu/implicit/>; Nosek et al., 2002). Our analysis included all gender-career IAT scores collected from respondents between 2005 and 2016 who had complete data and were located in countries with more than 400 total respondents ($N = 772,467$). We further restricted our sample based on participants' reaction times and error rates using the same criteria described in Nosek, Banjai, and Greenwald (2002, pg. 104). Our final sample included 663,709 participants from 47 countries, with a median of 965 participants per country. Note that although the respondents were from largely non-English speaking countries, the IAT was conducted in English. We do not have language background data from the participants, but we assume that most respondents from non-English speaking countries were native speakers of the dominant language of the country and L2 speakers of English.

Several measures have been used in the literature to quantify the strength of the bias from participants' responses on congruent and incongruent blocks on the IAT. Here, we used the most robust measure, D-score, which measures the difference between critical blocks for

each participant while controlling for individual differences in response time (Greenwald, Nosek, & Banaji, 2003). In addition to the implicit measure, we also analyzed an explicit measure of gender bias. After completing the IAT, participants were asked, “How strongly do you associate the following with males and females?” for both the words “career” and “family.” Participants indicated their response on a Likert scale ranging from *female* (1) to *male* (7). We calculated an explicit gender bias score for each participant as the Career response minus the Family response, such that greater values indicate a greater bias to associate males with career.

To obtain country-level gender bias estimates, we first calculated residual implicit and explicit bias scores for each participant, controlling for variables that are independent predictors of bias size (block order, participant sex, and age), and then averaged across participants within the same country.

We compared implicit and explicit gender biases to an objective measure gender equality that is measured for each country by the United Nations Educational, Scientific and Cultural Organization (UNESCO): the percentage of women among science, technology, engineering, and mathematics (STEM) graduates in tertiary education (Miller et al., 2015; Stoet & Geary, 2018).

Results

Figure 1 shows residualized implicit gender bias by country ($M = -0.01$; $SD = 0.03$). Implicit gender biases were moderately correlated with explicit gender biases both at the level of participants ($r = 0.16$, $p < .0001$) and countries ($r = 0.31$, $p = 0.03$). We also found that implicit gender bias was negatively correlated with percentage of women in STEM fields: Countries with a smaller gender bias tended to have more women in STEM fields ($r = -0.54$, $p < .01$). There was not a relationship between explicit gender bias and percentage of

women in STEM fields ($r = 0.09$, $p = 0.63$).

Discussion

In Study 1, we replicate previously reported patterns of gender bias in the gender-career IAT literature, with roughly comparable effect sizes (c.f. Nosek, et al., 2002). The weak correlation between explicit and implicit measures is consistent with claims that these two measures tap into different cognitive constructs (Forscher et al., 2016). In addition, consistent with previous research (Miller et al., 2015), we find that an objective measure of gender equality—female enrollment in STEM fields—is predictive of psychological gender bias.

Study 1: Gender bias and semantics

In Study 1, we ask whether participants’ implicit and explicit gender biases are correlated with the biases in the semantic structure of their native languages. For example, are the semantics of the words “woman” and “family” more similar in Hungarian than in English? Both the language-as-reflection and language-as-causal hypotheses predict a positive correlation between psychological and semantic gender biases. Importantly, we expect psychological and semantic gender biases to be correlated regardless of the direction of the relationship between psychological and objective gender bias (WPS) found in Study 1.

As a model of word meanings, we use large-scale distributional semantics models derived from auto-encoding neural networks trained on large corpora of text. The underlying assumption of these models is that the meaning of a word can be described by the words it tends to co-occur with—an approach known as distributional semantics (Firth, 1957). Under this approach, a word like “dog” is represented as more similar to “hound” than to “banana” because “dog” co-occurs with words more in common with “hound” than “banana.”

Recent developments in machine learning allow the idea of distributional semantics to be implemented in a way that takes into account many features of local language structure while remaining computationally tractable. The best known of these word embedding models is *word2vec* (Mikolov, Chen, Corrado, & Dean, 2013). The model takes as input a corpus of text and outputs a vector for each word corresponding to its semantics. From these vectors, we can derive a measure of the semantic similarity between two words by taking the distance between their vectors (e.g., cosine distance).

As it turns out, the biases previously reported using IAT tests can be predicted from distributional semantics models like *word2vec* using materials identical to those used in the IAT experiments. Caliskan, Bryson, and Narayanan (2017; henceforth *CBN*) measured the distance in vector space between the words presented to participants in the IAT task. CBN found that these distance measures were highly correlated with reaction times in the behavioral IAT task. For example, CBN find a bias to associate males with career and females with family in the career-gender IAT, suggesting that the biases measured by the IAT are also found in the lexical semantics of natural language.

CBN only measured semantic biases in English, however. In Study 2, we use the method described by CBN to measure gender bias in the range of first languages spoken by participants in Study 1 by using models trained on those languages. To do this, we take advantage of a set of models pre-trained on corpora of Wikipedia text in different languages—a different corpus than that used by CBN (Bojanowski, Grave, Joulin, & Mikolov, 2016). In Study 2a, we first validate word embedding measures of gender bias by comparing them to explicit human judgements of gender bias. In Study 2b, we apply this method to models trained on Wikipedia in other languages. We find that the implicit gender biases reported in Study 1 for individual countries are correlated with the biases found in the semantics of the natural language spoken by those participants.

Study 1a: Validating word embeddings as a measure of psychological gender bias

To determine whether word embeddings encoded information about psychological gender bias, we asked whether words that were closely associated with males in the word embedding models tended to be rated by human participants as being more male biased. We found human and word-embedding estimates of gender bias to be highly correlated.

Methods. We used an existing set of word norms in which participants were asked to rate “the gender associated with each word” on a Likert scale ranging from *very feminine* (1) to *very masculine* (7; Scott, Keitel, Becirspahic, Yao, & Sereno, 2018). We compared these gender norms to estimates of gender bias from a word embedding model pre-trained on the corpus of English Wikipedia using the fastText algorithm (a variant of word2vec; Bojanowski et al., 2016; Joulin, Grave, Bojanowski, & Mikolov, 2016). The model contains 2,519,370 words with each word represented by a 300 dimensional vector. To calculate a gender scores from the word embeddings, for each word we calculated the average cosine distance to a set of male words (“male”, “man”, “he”, “boy”, “his”, “him”, “son”, “brother”) and the average cosine similarity to a set of female words (“female”, “woman”, “she”, “girl”, “hers”, “her”, “daughter”, “sister”). A gender score for each word was then obtained by taking the difference of the similarity estimates (mean male similarity - mean female similarity), such that larger values indicated a stronger association with males. There were 4671 words in total that overlapped between the two data sources.

Results. Estimates of gender bias from word embeddings ($M = 0$; $SD = 0.03$) and human judgements ($M = 4.10$; $SD = 0.92$) were highly correlated ($r = 0.59$; $p < .0001$; Fig. 2).

Discussion.

Study 1b: Cross-linguistic gender semantics

With our corpus validated, we next turn toward examining the relationship between psychological and linguistic gender biases. In Study 2b, we estimate the magnitude of the gender-career bias in each of the languages spoken in the countries described in Study 1 and compare it with estimates of behavioral gender bias from Study 1. We predict these two measures should be positively correlated.

Methods. For each country included in Study 1, we identified the most frequently spoken language in each country using the CIA factbook (2017). This included a total of XX unique languages. For a sample of 20 of these languages (see Fig. 3), we had native speakers translate the set of 32 words from the gender-career IAT with a slight modification.² The original gender-career IAT task (Nosek et al., 2002) used proper names to cue the male and female categories (e.g. “John,” “Amy”). Because there are not direct translation equivalents of proper names, we instead used a set of generic gendered words which had been previously used for a different version of the gender IAT (e.g., “man,” “woman;” Nosek et al., 2002).

We used these translations to calculate an effect size from the models trained on Wikipedia in each language, using the same method as CBN (see SM for replication of CBN with the Wikipedia embedding model). We then compared the effect size of the linguistic gender bias to the behavioral IAT gender bias from Study 1, averaging across countries that speak the same language.

Results.

Discussion.

²The language sample was determined by accessibility to native speakers, but included languages from a variety of language families.

General Discussion

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Table 1

Correlation (Pearson's r) for all measures in Studies 1 and 2 at the level of languages. Asterisks indicate significance at the .05 level.

	Language IAT (Subtitles)	Language IAT (Wikipedia)	Residualized Explicit Bias	Residualized Behavioral IAT
Language IAT (Subtitles)				
Language IAT (Wikipedia)	.59**			
Residualized Explicit Bias	-.26	.06		
Residualized Behavioral IAT	.41	.28	.28	
Perecent Women in Stem	-.41	-.18	.25	-.48*

Implicit psychological gender bias by country

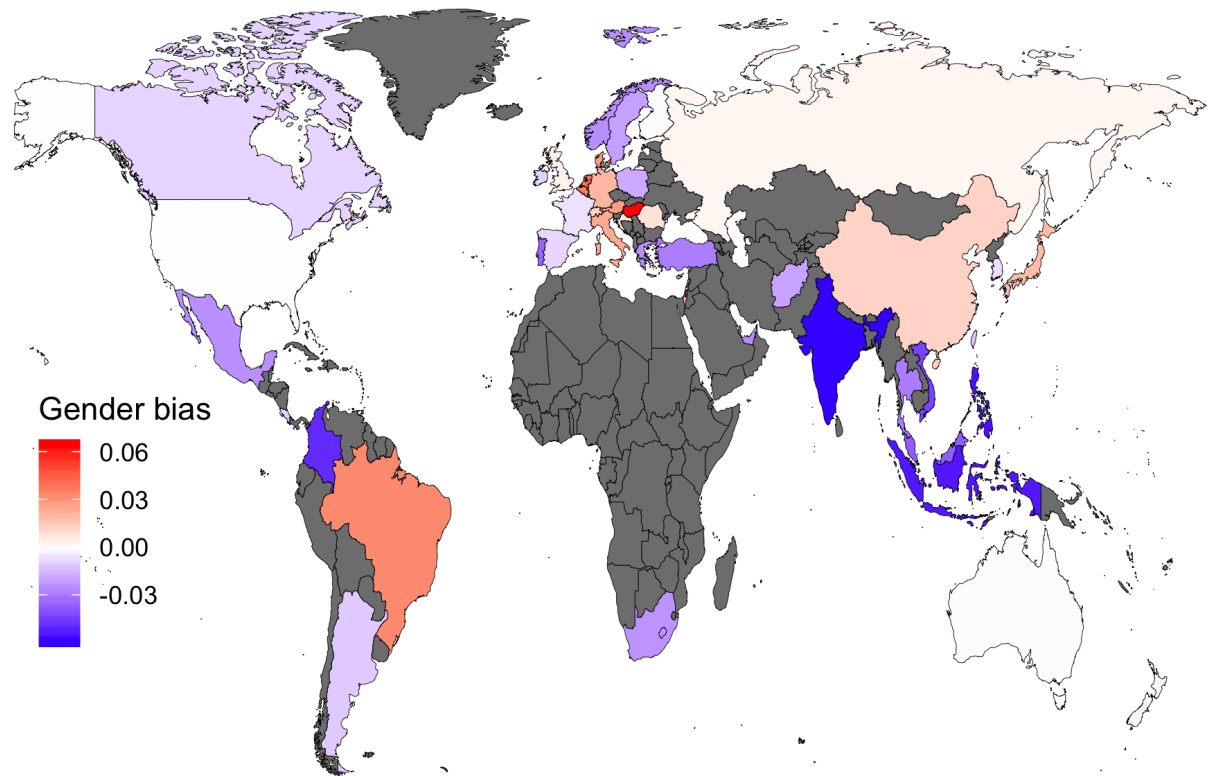


Figure 1. Residualized gender bias by country as measured by the IAT (Study 1). Larger values indicate a larger bias to associate women with the concept of family and men with the concept of career.

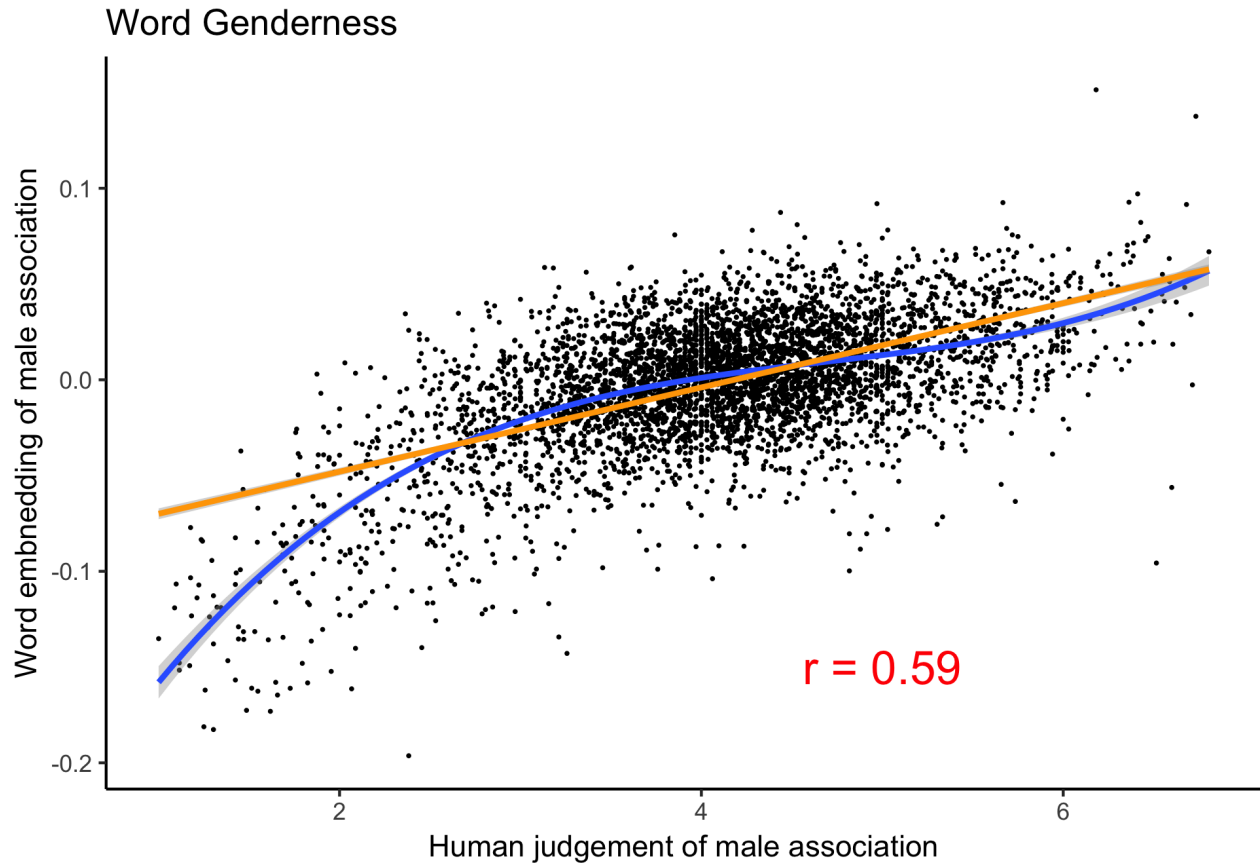


Figure 2. Word embedding estimates of gender bias as a function of human judgements of gender bias (Study 2a). Each point corresponds to a word. Larger numbers indicate stronger association with males. Orange line shows linear fit, and blue line shows third-order polynomial fit. Error bands correspond to standard errors.

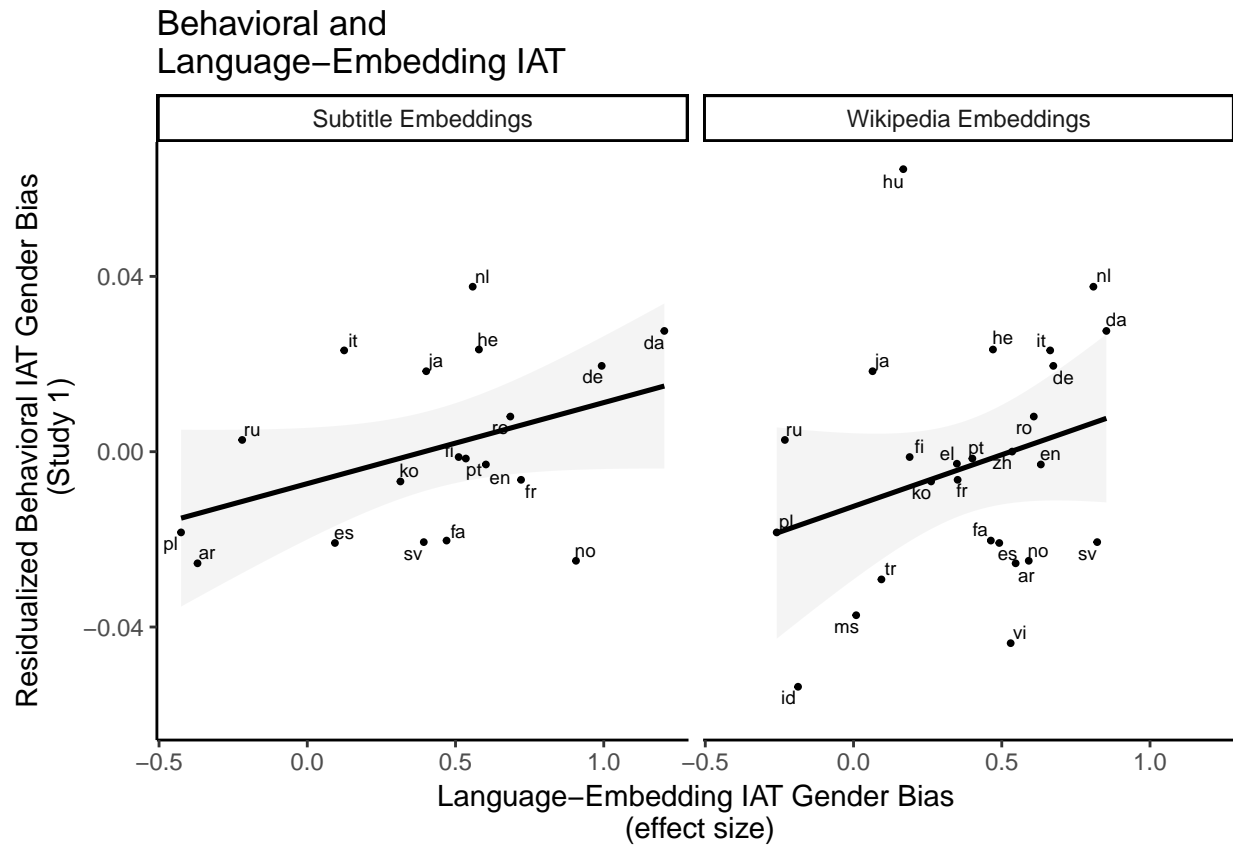


Figure 3. Residualized Behavioral IAT gender bias by language (Study 1) as a function of language-embedding IAT gender bias. Language-embedding biases are estimated from models trained on each language using a subtitle corpus (left) and a sample of Wikipedia (right).