

The role of experience in disambiguation during early word learning

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Abstract

Young children tend to map novel words to novel objects even in the presence of familiar competitors, a finding that has been dubbed the “disambiguation” effect. Theoretical accounts of this effect have debated whether it is due to initial constraints on children’s lexicons (e.g. a principle of mutual exclusivity) or situation-specific pragmatic inferences. We present synthesis of existing evidence on this phenomenon through a meta-analysis of the existing literature. We then present two experiments that help distinguish between these theoretical constraints. We conclude by suggesting that multiple cognitive mechanisms may underlie the disambiguation effect in word learning.

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Introduction

A central property of language is that each word in the lexicon maps to a unique concept, and each concept maps to a unique word (Clark, 1987). Like other important regularities in language (e.g., grammatical categories), children cannot directly observe this general property. Instead, they must learn to use language in a way that is consistent with the generalization on the basis of evidence about only specific word-object pairs.

Even very young children behave in a way that is consistent with this one-to-one regularity in language. Evidence for this claim comes from what is known as the “disambiguation” or “mutual exclusivity” (ME) effect (we return to the issue of nomenclature below). In a typical demonstration of this effect (Markman & Wachtel, 1988), children are presented with a novel and familiar object (e.g., a whisk and a ball), and are asked to identify the referent of a novel word (“Show me the dax”). Children in this task tend to choose the novel object as the referent, behaving in a way that is consistent with the one-to-one word-concept regularity in language across a wide range of ages and experimental paradigms (Bion, Borovsky, & Fernald, 2012; Golinkoff, Mervis, Hirsh-Pasek, & others, 1994; J. Halberda, 2003; Markman, Wasow, & Hansen, 2003; Mervis, Golinkoff, & Bertrand, 1994).

This effect has received much attention in the word learning literature because the ability to identify the meaning of a word in ambiguous contexts is, in essence, the core problem of word learning. That is, given any referential context, the meaning of a word is underdetermined (Quine, 1960), and the challenge for the word learner is to identify the referent of the word within this ambiguous context. Critically, the ability to infer that a novel word maps to a novel object makes the problem much easier to solve. For example, suppose a child hears the novel word “kumquat” while in the produce aisle of the grocery store. There are an infinite number of possible meanings of this word given this referential context, but the child’s ability to correctly disambiguate would lead her to rule out all meanings for which she already had a name. With this restricted hypothesis space, the child

is more likely to identify the correct referent than if all objects in the context were considered as possible referents.

Despite – or perhaps due to – the attention that the ME effect has received, there is little consensus regarding the cognitive mechanisms underlying it. Does it stem from a basic inductive bias on children’s learning abilities (“bias accounts,” see below), a learned regularity about the structure of language (“overhypothesis accounts”), reasoning about the goals of communication in context (“pragmatic accounts”), or perhaps some mixture of these? The goal of the current manuscript is to lay out these possibilities and discuss the state of the evidence. Along the way we present a meta-analysis of the extant empirical literature. We then present two new, relatively large-sample developmental experiments that investigate the dependence of children’s ME inferences on vocabulary (Experiment 1) and experience with particular words (Experiment 2). We end by discussing the emergence of ME inferences in a range of computational models of word learning. We conclude that:

1. Explanations of ME are not themselves mutually exclusive and likely more than one is at play;
2. The balance of responsibility for behavior likely changes developmentally, with basic biases playing a greater role for younger children and learned overhypotheses playing a greater role for older children.
3. All existing accounts put too little emphasis on the role of experience and strength of representation; this lack of explicit theory in many cases precludes definitive tests.
4. ME inferences are distinct from learning.

A note on terminology.

Markman and Wachtel (1988)’s seminal paper coined the term “mutual exclusivity,” which was meant to label the theoretical proposal that “children constrain word meanings by assuming at first that words are mutually exclusive – that each object will have one and only one label.” (Markman, 1990, p. 66). That initial paper also adopted a task used by a variety

of previous authors (CHECK THESE CITES, including ???, ???, ???), in which a novel and a familiar object were presented to children in a pair and the child was asked to “show me the x ,” where x was a novel label. Since then, informal discussions have used the same name for the paradigm and effect (selecting the novel object as the referent of the novel word) as well as the theoretical account (an early assumption or bias). This conflation of paradigm/effect with theory is problematic, as other authors who have argued against the theoretical account then are in the awkward position of rejecting the name for the paradigm they have used. Other labels (e.g. “disambiguation” or “referent selection” effect) are not ideal, however, because they are not as specific do not refer as closely to the previous literature. Here we adopt the label “mutual exclusivity” (ME) for the general family of paradigms and associated effects, *without* prejudgment of the theoretical account of these effects.

ME has also been referred to as “fast mapping.” This conflation is confusing at best. In an early study, S. Carey and Bartlett (1978) presented children with an incidental word learning scenario by using a novel color term to refer to an object: “You see those two trays over there. Bring me the *chromium* one. Not the red one, the *chromium* one.” Those data (and subsequent replications, e.g. L. Markson & Bloom, 1997) showed that this exposure was enough to establish some representation of the link between phonological form and meaning that endured over an extended period; a subsequent clarification of this theoretical claim emphasized that these initial meanings are partial (???). Importantly, however, demonstrations of retention relied on learning in a case where there was a contrastive presentation of the word with a larger set of contrastive cues (S. Carey & Bartlett, 1978) or pre-exposure to the object (L. Markson & Bloom, 1997).

Theoretical views of “mutual exclusivity”

What are the cognitive processes underlying this effect? A range of proposals in the literature.

Constraint and bias accounts. Under one proposal, Markman and colleagues (Markman & Wachtel, 1988, Markman et al. (2003)) suggest that children have a constraint on the types of lexicons considered when learning the meaning of a new word – a “mutual exclusivity constraint.” With this constraint, children are biased to consider only those lexicons that have a one-to-one mapping between words and objects. Importantly, this constraint can be overcome in cases where it is incorrect (e.g. property names), but it nonetheless serves to restrict the set of lexicons initially entertained when learning the meaning of a novel word. Under this view, then, the disambiguation effect emerges from a general constraint on the structure of lexicons. This constraint is assumed to be innate or early emerging.

N3C

Probabilistic accounts. Regier

McMurray

Frank Goodman Tenenbaum

Fazly

Over-hypothesis accounts. Lewis & Frank (2013)

Pragmatic accounts. The disambiguation effect is argued to result from online inferences made within the referential context (Clark, 1987, Diesendruck and Markson (2001)). In particular, Clark suggests that the disambiguation effect is due to two pragmatic assumptions held by speakers. The first assumption is that speakers within the same speech community use the same words to refer to the same objects (“Principle of Conventionality”). The second assumption is that different linguistic forms refer to different meanings (“Principle of Contrast”). In the disambiguation task described above, then, children might reason (implicitly) as follows: You used a word I’ve never heard before. Since, presumably we both call a ball “ball” and if you’d meant the ball you would have said “ball,” this new word must refer to the new object. Thus, under this account, the disambiguation effect emerges not from a higher-order constraint on the structure of lexicons, but instead from

in-the-moment inferences using general pragmatic principles.

These two proposals have traditionally been viewed as competing explanations of the disambiguation effect. Research in this area has consequently focused on identifying empirical tests that can distinguish between these two theories. For example, Diesendruck and Markson (2001) compare performance on a disambiguation task when children are told a novel fact about an object relative to a novel referential label. They found that children disambiguated in both conditions and argued on grounds of parsimony that the same pragmatic mechanism was likely to be responsible for both inferences. More recent evidence contradicts this view: tests of children with autism, who are known to have impairments in pragmatic reasoning find comparable performance on the disambiguation task between typically developing children and children with autism (de Marchena, Eigsti, Worek, Ono, & Snedeker, 2011; Preissler & Carey, 2005). This result provides some evidence for the view that disambiguation is due to a domain-specific lexical constraint.

Clark?

In the moment

Learned pragmatics

Logical inference accounts. Justin Halberda (2003)

Theory-constraining findings

NN vs. NF

Speaker-change studies

Autism

Bilingualism

Fast mapping + no retention

Developmental change (halberda)

Synthesis

These are definitely features of a successful account: Timescales - must be one “in the moment” - and one longer-term learned mechanism

Experience

Probabilistic representations

Could be the case also that it’s a mixture of pragmatic, etc.

We suggest this competing-alternatives approach to the disambiguation effect should be reconsidered. In a disambiguation task, learners may be making use of both general knowledge about how the lexicon is structured as well as information about the pragmatic or inferential structure of the task. Both of these constraints would then support children’s inferences. In other words, these two classes of theories may be describing distinct, complimentary mechanisms that each contribute to a single empirical phenomenon with their weights in any given task determined by children’s age and language experience, the nature of the pragmatic situation, and other task-specific factors.

The current study

Gather evidence on strength of finding

Test emergent relationship to vocabulary (E1)

Test causal relationship to representation strength (E2)

Re-evaluate

Meta-analysis

Methods

Search strategy. We conducted a forward search based on citations of Markman and Wachtel (1988) in Google Scholar, and by using the keyword combination “mutual exclusivity” in Google Scholar (September 2013; November 2017). Additional papers were identified through citations and by consulting experts in the field. We then narrowed our

sample to the subset that used one of two paradigms: (a) an experimenter says a novel word in the context of a familiar object and a novel object and the child guesses the intended referent (the canonical paradigm; “Familiar-Novel”), or (b) experimenter first provides the child with an unambiguous mapping of a novel label to a novel object, and then introduces a second novel object and asks the child to identify the referent of a second novel label (“Novel-Novel”). For Familiar-Novel conditions, we included conditions that included more than one familiar object (e.g. Familiar-Familiar-Novel). From these conditions, we restricted our sample to only those that satisfied the following criteria: (a) participants were children (less than 12 years of age)¹, (b) referents were objects or pictures (not facts or object parts), and (c) no incongruent cues (e.g. eye gaze at familiar object). All papers used either forced-choice pointing or eye-tracking methodology. All papers were peer-reviewed with the exception of two dissertations (Williams, 2009; Frank, I., 1999), but all main results reported below remain the same when these papers are excluded. In total, we identified 43 papers that satisfied our selection criteria and had sufficient information to calculate an effect size.

Coding. For each paper, we coded separately each relevant condition with each age group entered as a separate condition. For each condition, we coded the paper metadata (citation) as well as several potential moderator variables: mean age of infants, method (pointing or eyetracking), participant population type, estimates of vocabulary size from the Words and Gestures form of the MacArthur-Bates Communicative Development Inventory when available (MCDI; Fenson et al., 1994, Fenson et al. (2007)), referent type (object or picture), and number of alternatives in the forced choice task. We coded participant population as one of three subpopulationns that have studied in the literature: (a) typically-developing monolingual children, (b) multilingual children (including both bilingual and trilingual children), and (c) non-typically developing children. Non-typically developing conditions included children with selective language impairment, language delays,

¹This cutoff was arbitrary but allowed us to include conditions from older children from non-typically-developing populations.

hearing impairment, autism spectrum disorder, and down-syndrome.

In order to estimate effect size for each conditions, we also coded sample size, proportion novel-object selections, baseline (e.g., .5 in a 2-AFC paradigm), and standard deviations for novel object selections, t -statistic, and Cohen’s d . For several conditions, there was data were insufficient data reported in the main text to calculate an effect size (no means and standard deviations, t -statistics, or Cohen’s ds), but we were able to estimate the means and standard deviations through measurement of plots ($N = 13$), imputation from other data within the paper ($N = 4$; see SI for details), or through contacting authors ($N = 26$). Our final sample included 157 effect sizes ($N_{\text{typical-developing}} = 135$; $N_{\text{multilingual}} = 12$; $N_{\text{non-typically-developing}} = 10$).

Statistical approach. We calculated effect sizes (Cohen’s d) from reported means and standard deviations where available, otherwise we relied on reported test-statistics (t or d). Effect sizes were computed by a script, `compute_es.R`, available in the Github repository. All analyses were conducted with the metafor package (Viechtbauer, 2010) using mixed-effect models with grouping by paper.² In models with moderators, moderators variables were included as additive fixed effects.

Meta-analytic Analyses

We conducted a separate meta-analysis for four theoretically-relevant conditions: Familiar-Novel trials with typically developing participants, Novel-Novel trials with typically developing participants, conditions with multilingual participants, and conditions with non-typically developing participants.

Typically-Developing Population: Novel-Familiar Trials. We first examined effect sizes for the disambiguation effect for typically-developing children in the canonical familiar-novel paradigm. This is the central data point that theories of disambiguation seek to explain.

²The exact model specification was as follows: $model < -metafor :: rma.mv(yi = effect_size, V = effect_size_var, random = 1|paper, data = d)$.

Results. The overall effect size for these conditions was 1.1 [0.79, 1.42], and reliably greater than zero ($p < .001$). The effect sizes contained considerable heterogeneity, however ($Q = 968.13$; $p < .001$).

We next tried to predict this heterogeneity with two key moderators: age and vocabulary. In a model with age as a moderator, age was a reliable predictor of effect size ($\beta = 0.05$, $z = 11.85$, $p < .001$; see Table X), suggesting that the disambiguation effect becomes larger as children get older. Age of participants was highly correlated with vocabulary size in our sample ($r = 0.65$, $p < .01$), so next we asked whether vocabulary size predicted independent variance in the magnitude of the disambiguation bias on the subset of conditions for which we had estimates of vocabulary size ($N = 23$). To test this, we fit a model with both age and vocabulary size as moderators. Vocabulary size ($\beta = 0.07$, $z = 2.14$, $p = 0.03$), but not age ($\beta = -0.78$, $z = -1.11$, $p = 0.27$), was a reliable predictor of disambiguation effect size. ACTUALLY THIS ISN'T TRUE (true only for full model)

Discussion. These analyses confirm that the disambiguation phenomenon is robust, and associated with a relatively large effect size ($d = 1.1$ [0.79, 1.42]). Importantly, this set of analyses provides theory-constraining evidence about the mechanisms underlying the effect. In particular, the finding that vocabulary predicts more variance in effect size, compared to age, suggests that there is an experience related component to the mechanism, independent of pure maturational development.

Typically-Developing Population: Novel-Novel Trials. The results from the Familiar-Novel trials point to a role for vocabulary knowledge in the strength of the disambiguation effect. One way in which this vocabulary knowledge could lead to increased performance on the Familiar-Novel disambiguation task is through increased certainty about the label associated with the familiar word: If children less certain that a ball is called “ball,” the child should be less certain that the novel label applies to the novel object. Novel-Novel trials control for potential variability in certainty about the familiar object by teaching participants a new label for a novel object prior to the critical disambiguation trial, where

this previously-learned label becomes the “familiar” object in the disambiguation trial. If knowledge of the familiar object is not the only contributor to age-related changes in the disambiguation effect, then there should also be developmental change in Novel-Novel trials. In addition, if the strength of knowledge of the “familiar” object influences the strength of the disambiguation effect, then the overall effect size should be smaller for Novel-Novel trials, compared to Familiar-Novel trials

For conditions with the Novel-Novel trial design, the overall effect size was 1.36 [0.6, 2.11] and reliably greater than zero ([0.6, 2.11]). We next asked whether age predicted some of the variance in these trials by fitting a model with age as a moderator. Age was a reliable predictor of effect size ($\beta = 0.03$, $z = 3.55$, $p < .001$).

Finally, we fit a model with both age and trial type (Familiar-Novel or Novel-Novel) as moderators of the disambiguation effect. Both moderators predicted independent variance in disambiguation effect size (age: $\beta = -0.08$, $z = -0.08$, $p = -0.08$; trial-type: $\beta = 0.04$, $z = 12.34$, $p = 0$), with Familiar-Novel conditions and conditions with older participants tending to have larger effect sizes.

These analyses point to an influence of both development (either via maturation or experience-related changes) as well as the strength of the familiar word representation on the disambiguation effect. A successful theory of disambiguation will need to account for both of these facts.

Multilingual Population. We next turn to a different population of participants: Children who are simultaneously learning multiple languages. This population is of theoretical interest because it allows us to isolate the influence of linguistic knowledge from the influence of domain-general capabilities. If the disambiguation phenomenon relies on mechanisms that are domain-general and independent of linguistic knowledge, then we should expect the magnitude of the effect size to be the same for multilingual children compared to monolingual children.

Children learning multiple languages reliably showed the disambiguation effect ($d =$

1.57 [0.69, 2.44]). We next fit a model with both monolingual (typically-developing) and multilingual participants, predicting effect size with language status (monolingual vs. multilingual), while controlling for age. Language status was not a reliable predictor of effect size ($\beta = 0.03$, $z = 11.54$, $p = 0$), but age was ($\beta = -0.36$, $z = -1.88$, $p = 0.06$).

These data do not provide strong evidence that language-specific knowledge influences effect size, however, the small sample size of studies from this population limit the power of this model to detect a difference if one existed.

Non-Typically-Developing Population. Finally, we examine a third-population of participants: non-typically developing children. This group includes a heterogeneous sample of children with diagnoses including Autism-Spectrum Disorder (ASD), Mental Retardation, Williams Syndrome, Late-Talker, Selective Language Impairment, and deaf or hard-of-hearing participants. These populations are of theoretical interests because they allow us to observe how impairment to a particular aspect of cognition influences the magnitude of the disambiguation effect. For example, children with ASD are thought to have impaired social reasoning skills (e.g., Phillips, Baron-Cohen, & Rutter, 1998); thus, if children with ASD are able to succeed on disambiguation tasks, this suggests that social reasoning skills are not necessary to making a disambiguation inference.

Overall, non-typically developing children succeeded on disambiguation tasks ($d = 1.57$ [0.69, 2.44]). In a model with age as a moderator, age was a reliable predictor of the effect, suggests children became more accurate with age, as with other populations ($\beta = 0.04$, $z = 3.15$, $p < .001$).

We also asked whether the effect size for non-typically developing children different from typically-developing children, controlling for age. We fit a model predicting effect size with both development type (typical vs. non-typical) and age. Development type was a reliable predictor of effect size with non-typically developing children tending to have a smaller bias compared to typically developing children ($\beta = 0.04$, $z = 11.34$, $p = 0$). Age was also a reliable predictor of effect size in this model ($\beta = -0.14$, $z = -0.84$, $p = 0.40$).

This analysis suggests that non-typically developing children succeed in the disambiguation paradigm just as typically developing children do, albeit at lower rates. Theoretical accounts of the disambiguation phenomenon will need to account for how children can successfully make the inference while also having the range a range of different cognitive impairments.

Discussion

Experiment 1: ME and Vocabulary

Methods

We report how we determined our sample size, all data exclusions (if any), all manipulations, and all measures in the study.

Participants. Children were recruited at the Children’s Discovery Museum of San Jose. Children were asked if they would be willing to play an iPad game with the experimenter and were informed that they could stop playing at any time. Children first completed two tasks adapted to iPad; one probing their vocabulary size and one mutual exclusivity inference task. Included in analyses are 166 children out of a planned sample of 160 participants. We ran 62 additional children, who were excluded from analysis based on planned exclusion criteria of low English language exposure (less than or equal to 75%), outside the age range of 24-48 month, children who do not give correct answers on > 50% of familiar noun (control) trials, or < 100% of trials completed. Included in our sample were 97 females and 69 males.

Stimuli. Mutual exclusivity inference task was comprised of 19 trials total; three practice trials of Familiar-Familiar (FF) nouns and 16 experimental trials. Experimental trials consisted of Novel-Familiar (NF), and Novel-Novel (NN) noun pairings. Of the pictures presented in the task, 14 objects were familiar and 24 objects were novel. The task included 8 control trials, equally split between NN noun pairings (C-NN) and NF noun pairings (C-NF) given in random order. Children who did not give correct answers on 50% of control

336 trials were excluded from the final sample. The remaining 8 trials were divided equally
337 between NN and NF trials.

338 The general format of the vocabulary assessment comprised of a 4 image display and a
339 verbal prompt. Two practice trials were administered, followed by 20 experimental trials.
340 Experimental trials included a fixed set of 20 developmentally appropriate words taken from
341 the Pearson Peabody Vocabulary Test. These words were taken from 9 different domains,
342 including professions, food, outside things, instruments, animals, classroom, shapes, verbs,
343 and household items.

344 **Procedure.** Sessions took place individually in a small testing room away from the
345 museum floor. In the ME inference task, the experimenter introduced them to “Mr. Fox,” a
346 cartoon character who wanted to play a guessing game. The experimenter explained that
347 Mr. Fox would tell them the name of the object they had to find, so they had to listen
348 carefully. Children then saw 3 practice trials with two commonly known objects (i.e. cup and
349 cookie). If the participant chose incorrectly for this practice trial, the audio would correct
350 them and allow the participant to choose again. After the practice trials were completed, the
351 task proceeded to run 16 test trials. Reaction times were measured from the onset of the
352 target word. Children could only make one selection. The vocabulary task displayed 4
353 images randomly selected from the fixed bank of 22 images. Participants were prompted to
354 choose one object. Again, reaction times were measured from the onset of the target word
355 and children could only make one selection.

356 **Data analysis.** We used R (3.4.1, R Core Team, 2017) for all our analyses.

357 Results and Discussion

358 Could be specific strength of particular word in the NF pairing
359 but we also get it for NN trials alone

Experiment 2: ME and Familiarity

Methods

We report how we determined our sample size, all data exclusions (if any), all manipulations, and all measures in the study.

	age_group	mean_age	n
Participants.	2	30.98684	38
	3	40.98571	35
	4	52.16216	37

We planned a total sample of 108 children, 12 per between-subjects labeling condition, and 36 total in each one-year age group. Our final sample was 110 children, ages Inf – -Inf months, recruited from the floor of the Boston Children’s Museum. Children were randomly assigned to the one-label, two-label, or three label condition, with the total number of children in each age group and condition ranging between 10 and 13.

Materials. Materials were the set of novel objects used in de Marchena et al. (2011), consisting of unusual household items (e.g., a yellow plastic drain catcher) or other small, lab-constructed stimuli (e.g., a plastic lid glued to a popsicle stick). Items were distinct in color and shape.

Procedure. Each child completed four trials. Each trial consisted of a training and a test phase in a “novel-novel” disambiguation task (??). In the training phase, the experimenter presented the child with a novel object, and explicitly labeled the object with a novel label 1, 2, or 3 times (“Look at the *dax*”), and contrasted it with a second novel object (“And this one is cool too”) to ensure equal familiarity. In the test phase, the child was asked to point to the object referred to by a second novel label (“Can you show me the *zot*?”). Number of labels used in the training phase was manipulated between subjects. There were eight different novel words and objects. Object presentation side, object, and word were counterbalanced across children.

Data analysis. We followed the same analytic approach as we registered in Experiment 1, though data were collected chronologically earlier for Experiment 2. Responses were coded as correct if participants selected the novel object at test. A small number of trials were coded as having parent or sibling interference, experimenter error, or a child who recognized the target object, chose both objects, or did not make a choice. These trials were excluded from further analyses; all trials were removed for two children for whom there was parent or sibling interference on every trial. The analysis we report here is consistent with that used in Lewis and Frank (2013), though there are some slight numerical differences due to reclassification of exclusions.

err_type	n	pct
changed mind	2	0.0045455
exp err	2	0.0045455
interference	11	0.0250000
no choice	8	0.0181818
recog obj	4	0.0090909

Results and Discussion

As predicted, children showed a stronger disambiguation effect as the number of training labels increased, and as noise decreased with age.

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	0.3076191	0.1046804	2.938650	0.0032965
age_mo_c	0.0464060	0.0112418	4.127972	0.0000366
times_labeled_c	0.4832010	0.1287155	3.754022	0.0001740
age_mo_c:times_labeled_c	0.0214303	0.0135810	1.577960	0.1145749

We analyzed the results using a logistic mixed model to predict correct responses with age, number of labels, and their interaction as fixed effects, and participant as a random effect. We centered both age and number of labels for interpretability of coefficients. Model

results are shown in Table XYZ. There was a significant effect of age such that older children showed a stronger disambiguation bias and a significant effect of number of labels, such that more training labels led to stronger disambiguation, but the interaction between age and number of labels was not significant.

ME in Models of Word Learning

Basic statistical biases (“explaining away”)

Regier (2005) model shows ME emergent as noted by Frank, Goodman, Lai, and Tenenbaum (2009), Yu and Ballard (2007) model (IBM machine translation model #1, (???) for that; subsequently adapted by Nematzadeh, Fazly, and Stevenson (2012)) shows ME as well.

this is because any conditional probability model will show the same effect

In other words, Markman and Wachtel (1988)’s sense of a basic inductive bias will likely be present in a wide variety of different learning models.

What is the experience-dependence of ME in these models? In the Frank et al. (2009) model, the strength of the ME response scales with the strength of the familiar word’s mapping; the same thing is true for the other models presumably.

Open question whether the actual difference in a 2-year-olds’ and a 4-year-olds’ strength of representation of “ball” is what matters here?

Frank et al. (2009) model shows ME, in fact stronger than basic conditional probability. This is in part due to the use of the intention variable.

As a side note, the (???) no retention finding is shown in an even more pragmatic model: Smith, Goodman, and Frank (2013) model shows ME with no retention (though explanation in that model is a little implausible “because the speaker might not be committed to that label and is just using it as a matter of convenience.”)

Primary point: No support here for overhypothesis building, which is suggested by 1) the bilingualism results. In order to fit the bilingual data, in general we’d have to assume

that strength of individual representations in monolinguals and bilinguals was a driver, and this seems unlikely. 2) no support for E1 vocab findings unless the entire developmental trend is due to strength of the familiar word representations. In general, the strong — likely false — claim from all of these models is that the individual representation of the familiar object strength is the only locus for developmental/population-related change.

McMurray, Horst, and Samuelson (2012) model has ME emerge from the competition dynamics of a neural network.

Thus, the selection of the novel object is dependent on the learning rule, but not because the network needs to learn something about that object/word. Rather, the weights between the known word/objects and the unused lexical units must decay, and the weights between the novel ones must not in order to create a platform upon which real-time competition dynamics can select the right object. A different type of weight decay (for example, if all weights decayed on each epoch) would not preserve the right form of the weight matrix. However, learning is not the whole story: this pattern of connectivity could not be harnessed in situation time without the gradual settling process represented by the inhibition and feedback dynamics. Moreover, the model's ability to learn from M.E. referent selection may also depend on this competition/feedback cycle. The model must select a single lexical unit and selectively amplify the novel object in order to eventually turn a word-referent link created during M.E. referent selection into a known word by associating the novel object with the novel word over many instances. Thus, while as a real-time process mutual exclusivity is likely to impact learning, it is really more the product of learning than a mechanism of it.

This proposal is complicated but might capture the global and local dynamics in Experiment 1 & 2 better than others.

(???) deal with bilingual data by adding a direct ME-related penalty, not letting it be emergent.

General Discussion

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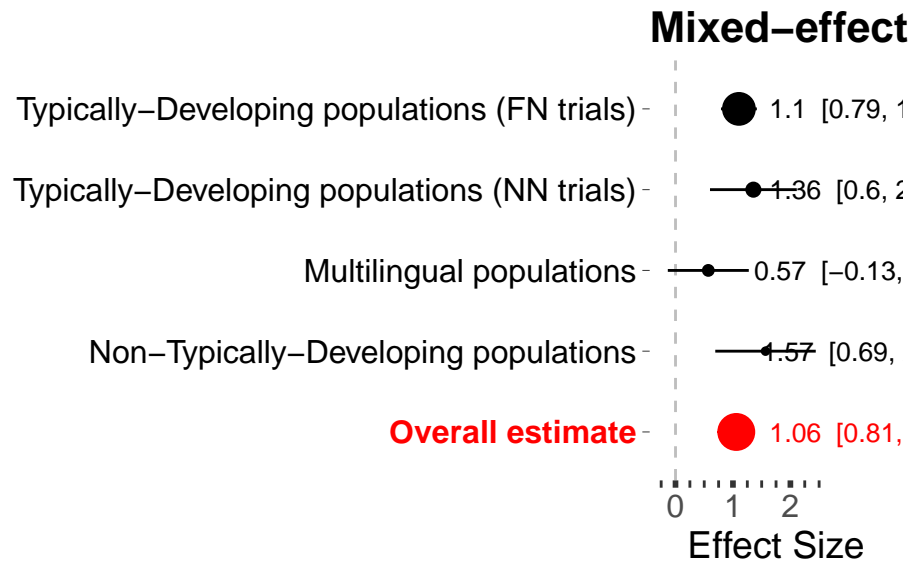


Figure 1. Mixed-effect effect size estimates for all conditions (red) and each of the four theoretically-relevant conditions in our sample. Ranges are 95% confidence intervals. Point size corresponds to sample size. FN = Familiar-Novel trials; NN = Novel-Novel trials.

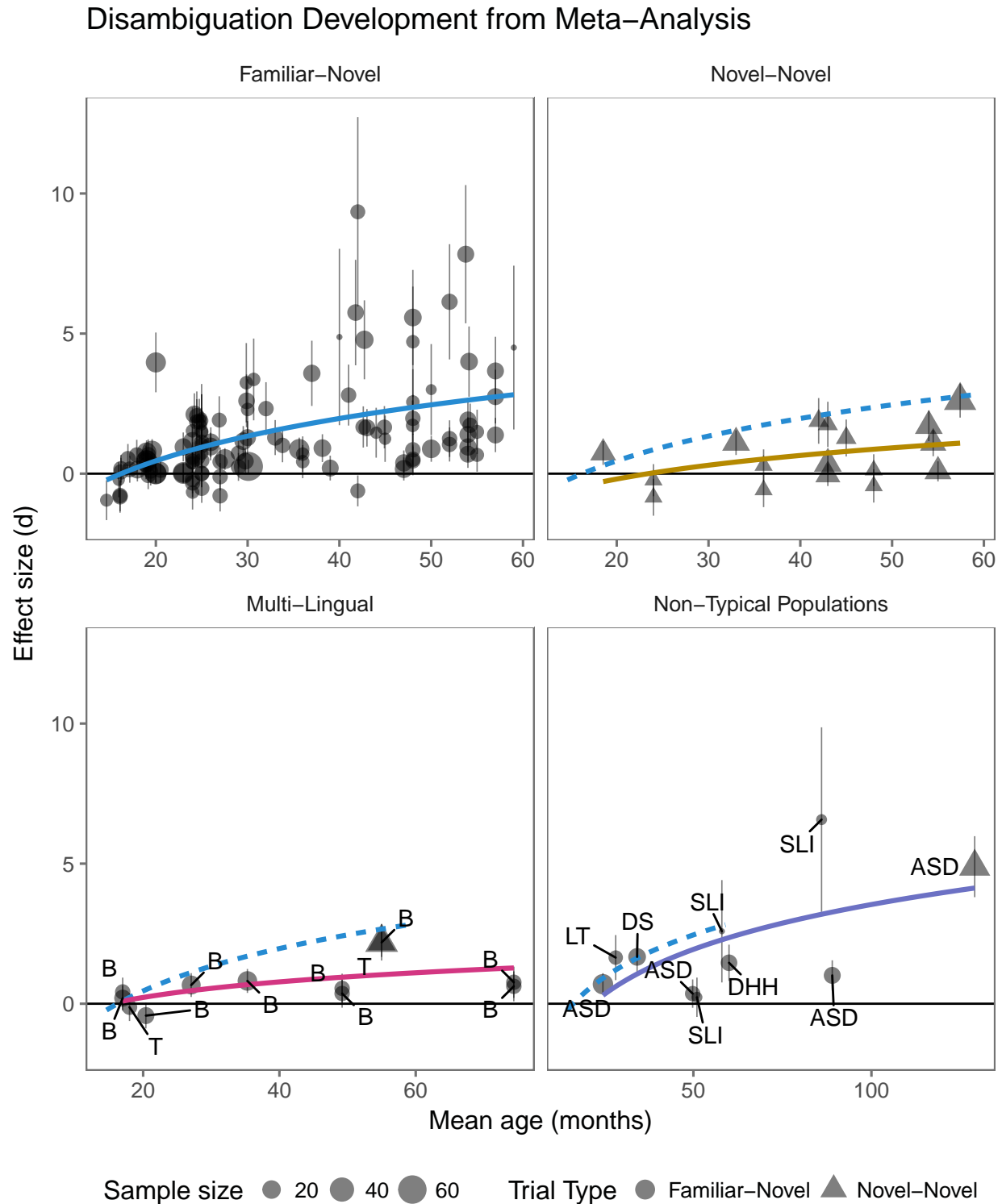
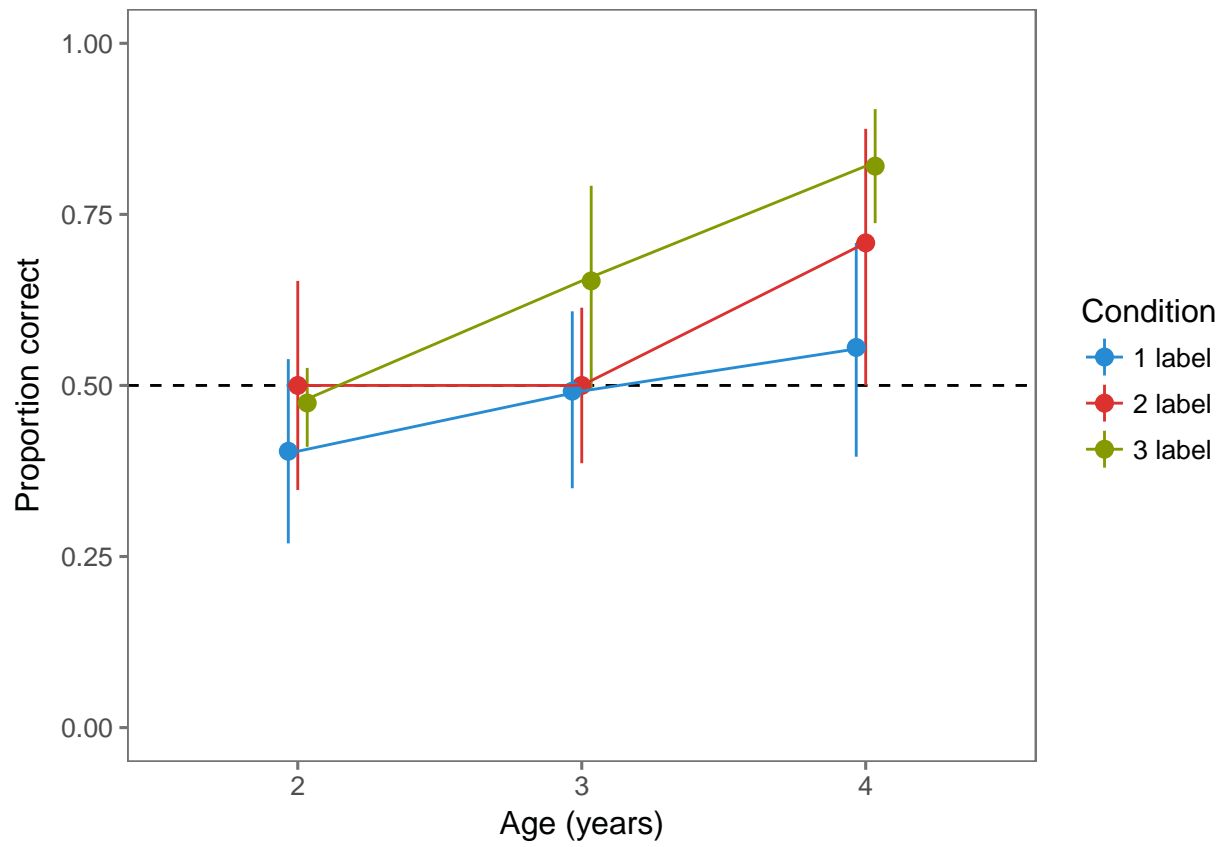


Figure 2. Developmental plots for each moderator. Ranges correspond to 95% confidence intervals. Model fits are log-linear. Note that the x-axis scale varies by facet.

*Figure 3*