Lab 3: Intro to Decision Theory

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Task 1

```
# set seed
set.seed(123)

# data
sum_x = 1
n = 30
# prior parameters
a = 0.05; b = 1
# posterior parameters
an = a + sum_x
bn = b + n - sum_x
th = seq(0,1,length.out = 100)
like = dbeta(th, sum_x+1,n-sum_x+1)
prior = dbeta(th,a,b)
post = dbeta(th,sum_x+a,n-sum_x+b)
```

We now consider the loss function.

```
# compute the loss given theta and c
loss_function = function(theta, c){
  if (c < theta){
    return(10*abs(theta - c))
} else{
    return(1 = abs(theta - c))
}</pre>
```

We now write a function **posterior_risk** which is a function of c, parameters a_prior and b_prior for the prior distribution of θ , the summation of x_i sum_x, the number of observations n, and also the number of random draws s.

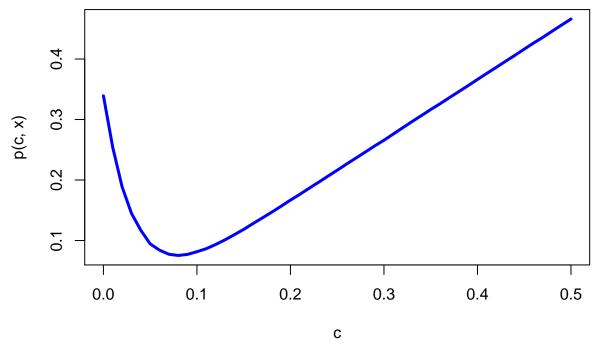
```
# compute the posterior risk given c
# s is the number of random draws
posterior_risk = function(c, a_prior, b_prior, sum_x, n, s = 30000){
    # randow draws from beta distribution
    a_post = a_prior + sum_x
    b_post = b_prior + n - sum_x
    theta = rbeta(s, a_post, b_post)
    loss <- apply(as.matrix(theta),1,loss_function,c)
    # average values from the loss function
    risk = mean(loss)
}
# a sequence of c in [0, 0.5]</pre>
```

```
c = seq(0, 0.5, by = 0.01)
post_risk <- apply(as.matrix(c),1,posterior_risk, a, b, sum_x, n)
head(post_risk)</pre>
```

[1] 0.33917940 0.25367603 0.18868962 0.14489894 0.11693106 0.09453471

We then look at the Posterior expected loss (posterior risk) for disease prevelance versus c.

```
# plot posterior risk against c
plot(c, post_risk, type = '1', col='blue',
    lwd = 3, ylab = 'p(c, x)' )
```



```
# minimum of posterior risk occurs at c = 0.08
(c[which.min(post_risk)])
```

[1] 0.08

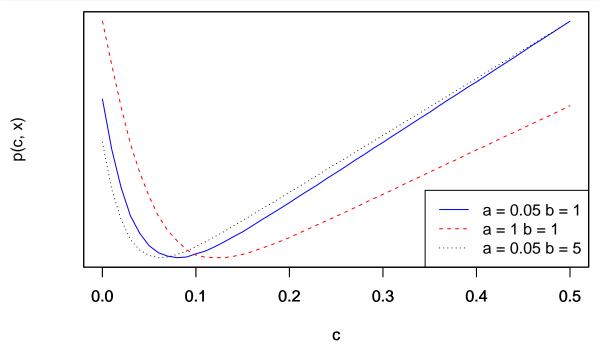
Task 2

We now consider task 2. We set a = 0.05, 1, 0.05 and b = 1, 2, 10. If we have different prior, the posterior risk is minimized at different c values. The optimal c depends on not only the data, but also the prior setting.

```
# set prior
as = c(0.05, 1, 0.05); bs = c(1, 1, 10)
post_risk = matrix(NA, 3, length(c))

# for each pair of a and b, compute the posterior risks
for (i in 1:3){
    a_prior = as[i]
    b_prior = bs[i]

    post_risk[i,] = apply(as.matrix(c), 1, posterior_risk, a_prior, b_prior, sum_x, n)
}
```



Note there is a more automated solution but this is the most simple one and is completely correct.

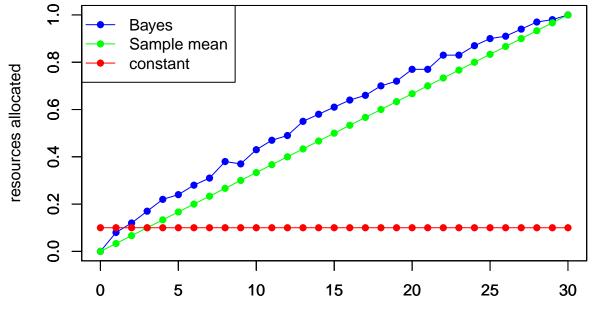
Task 3

The Bayes procedure always picks c to be a little bigger than \bar{x} .

```
sum_xs = seq(0, 30)
min_c = matrix(NA, 3, length(sum_xs))

# find_optimal_C finds the optimal c under Bayes procedure
# function of sum of x, parameters for prior, number of observations, and number of random draws
find_optimal_C = function(sum_x, a_prior, b_prior, n, s = 500){
    c = seq(0, 1, by = 0.01)
    post_risk = apply(as.matrix(c), 1, posterior_risk, a_prior, b_prior, sum_x, n, s)
    c[which.min(post_risk)]
}

min_c[1,] = apply(as.matrix(sum_xs), 1, find_optimal_C, a, b, n)
# find optimal c under sample mean
min_c[2,] = (sum_xs)/n
# constant c
min_c[3,] = 0.1
# plot
```



Task 4

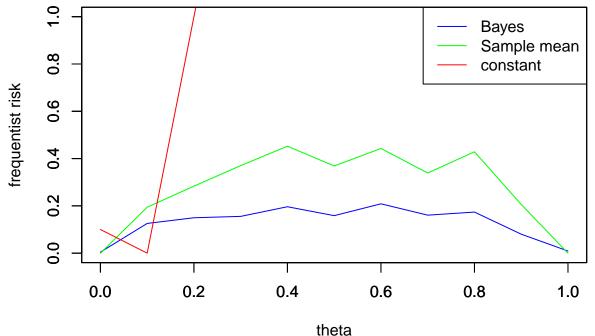
For all θ , the Bayes procedure has the lower frequentist risk than the sample mean.

```
thetas = seq(0, 1, by=0.1)

# frequentist risk for the 3 estimators given a theta
frequentist_risk = function(theta){
    sum_xs = rbinom(100, 30, theta)
    Bayes_optimal = apply(as.matrix(sum_xs), 1, find_optimal_C, a, b, n, s = 100)
    mean_c = sum_xs / 30

loss1 = apply(as.matrix(Bayes_optimal), 1, loss_function, theta = theta)
    loss2 = apply(as.matrix(mean_c), 1, loss_function, theta = theta)
    risk1 = mean(loss1)
    risk2 = mean(loss2)
    risk3 = loss_function(theta, 0.1)
    return(c(risk1, risk2, risk3))
```

observed number of diseased cases



Please see a few remarks about Task 4 that will help you with interpreting the plot.

- 1. If you zoom into see the plot for Task 4, you will notice that the Bayes risk is not always smaller than the sample mean. Specifically, the issue is occurring around $\theta = 0$ and $\theta = 1$.
- 2. One observation that we can make is that when x is very small (say 0), Bayes estimator tends to overestimate θ and hence sample mean has lower risk. What other observations can you make?

I am including some code that Xu Chen has written that is much faster, where the resulting plot is much more clear.

```
# code by Xu Chen
loss <- function(theta, c){
  if (c >= theta) {
    return(c - theta)
} else {
    return(10*(theta - c))
```

```
}
}
delta1 <- function(x){</pre>
  return(rep(0.1,length(x)))
delta2 <- function(x){</pre>
  return(x/30)
}
delta3 \leftarrow function(x, a = 0.05, b = 1){
  a.post \leftarrow a + x
  b.post \leftarrow b + 30 - x
  c <- seq(0,1,0.01)
  theta <- matrix(rbeta(1e4*length(c), a.post, b.post), nrow = 1e4, ncol = length(c))
  return(c[which.min(sapply(c, function(u) mean(sapply(theta[,as.integer(100*u+1)], loss, c = u))))])
risk <- function(theta, c){
 return(sum(dbinom(x = 0:30, size = 30, prob = theta) * sapply(c, loss, theta = theta)))
}
theta.grid \leftarrow seq(0,1,0.01)
x <- 0:30
c3 <- sapply(x, delta3)
plot(theta.grid, sapply(theta.grid, risk, c3), ylim = c(0,1), type = 'l', col = 'red', xlab = expression
c1 <- delta1(x)</pre>
points(theta.grid, sapply(theta.grid, risk, c1), ylim = c(0,1), type = 'l')
c2 \leftarrow delta2(x)
points(theta.grid, sapply(theta.grid, risk, c2), ylim = c(0,1), type = 'l', col='green')
legend('topright', legend = c('Bayes', 'sample mean', 'constant'), col = c('red', 'green', 'black'), lt
```

