

# Latent Health Dynamics

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  - “True health” is polydimensional!
- ② Solving dynamic structural models is hard:
  - Even fairly basic models have computational burden
  - Adding 1 more state variable  $\Rightarrow$  curse of dimensionality
  - Adding several state variables? Things get tricky...

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  - How much adverse selection should we expect for LTDI?
  - What are the welfare implications of individually rated vs community rated health insurance premiums?
- If “health” is complex **and** adding complexity to dynamic structural models is hard... then something's gotta give!

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- **MEPS (SAQ)**: “In general, would you say your health is: Excellent, Very Good, Good, Fair, Poor”

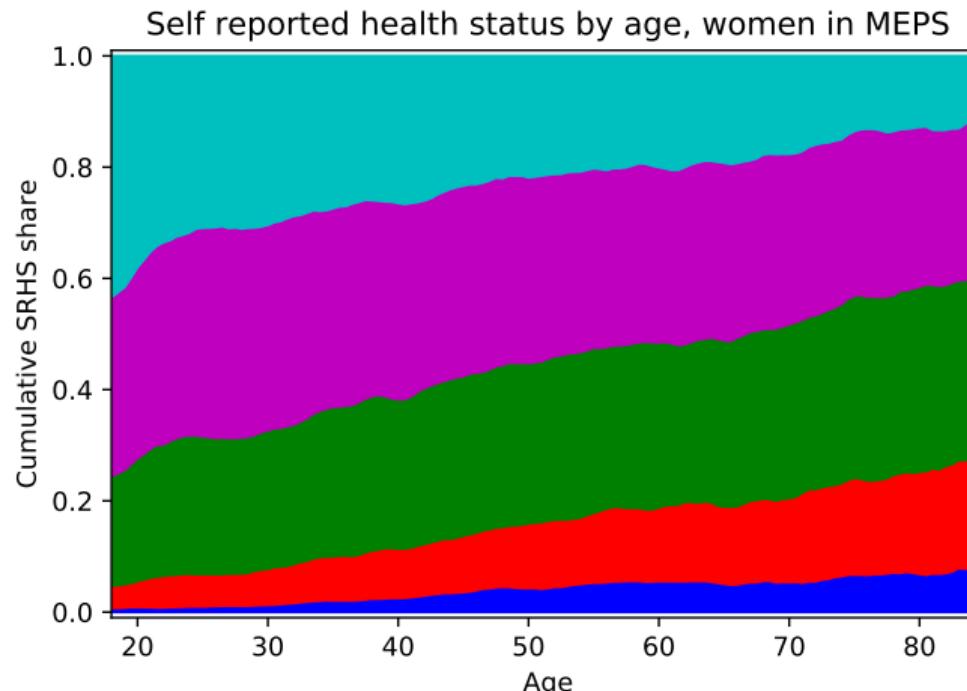
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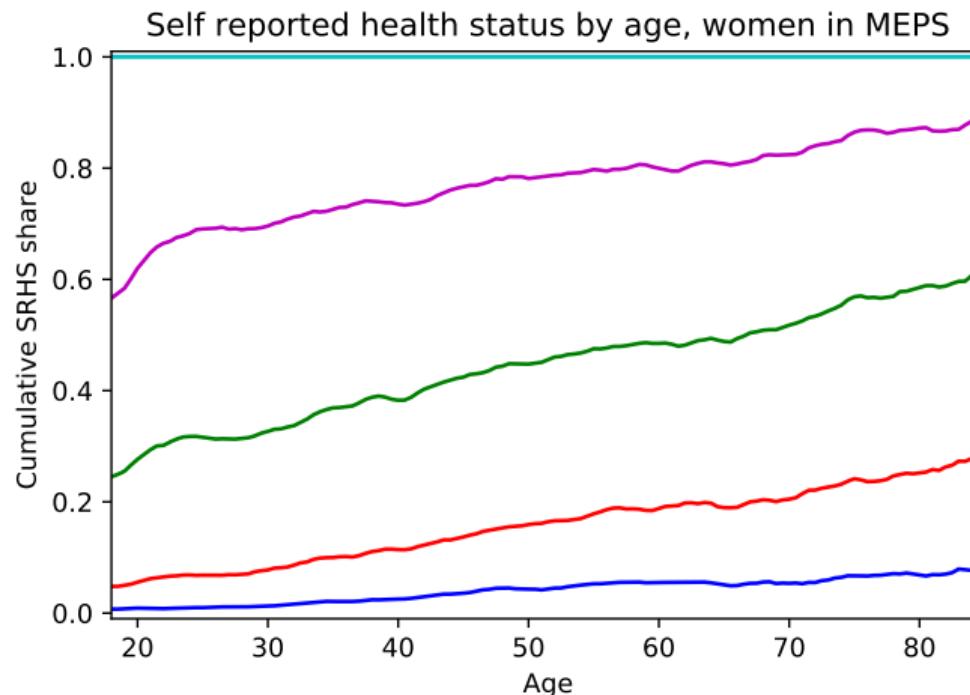
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- **BHPS**: “Compared to people of your own age, would you say that your health has on the whole been... Excellent, Goode, Faire, Poore, Very Poore”

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  - Labor supply (Bound (1991), Blundell et al (2017))
- **And** it's highly correlated with objective clinical measures that directly measure aspects of "health" (LaRue (1979))
- Cheap and easy to survey, everyone understands the question, correlated with the thing we want to measure, predictive of relevant outcomes, simple to implement in models...
- ...So what's the problem?

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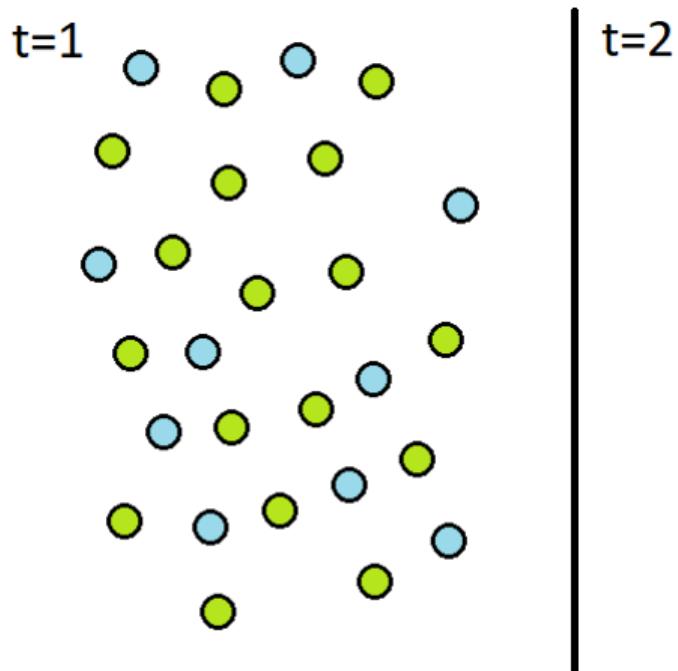
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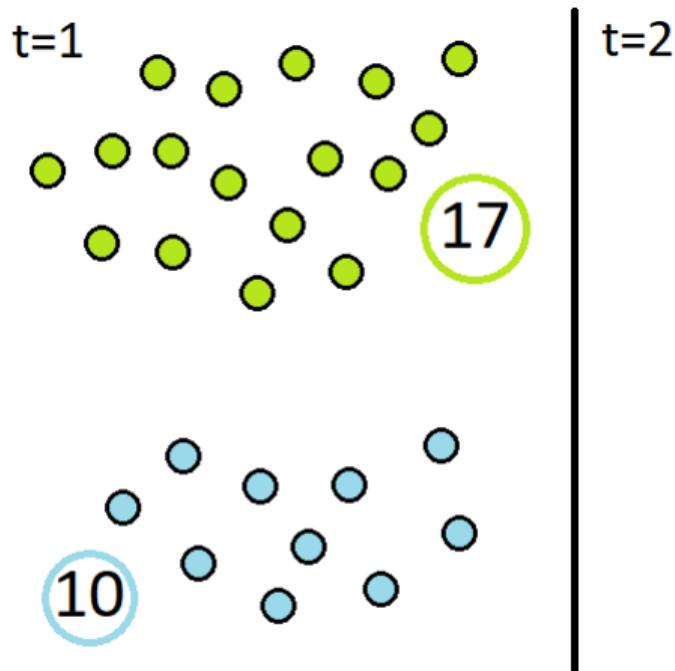
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- And taking SRHS **way** too literally

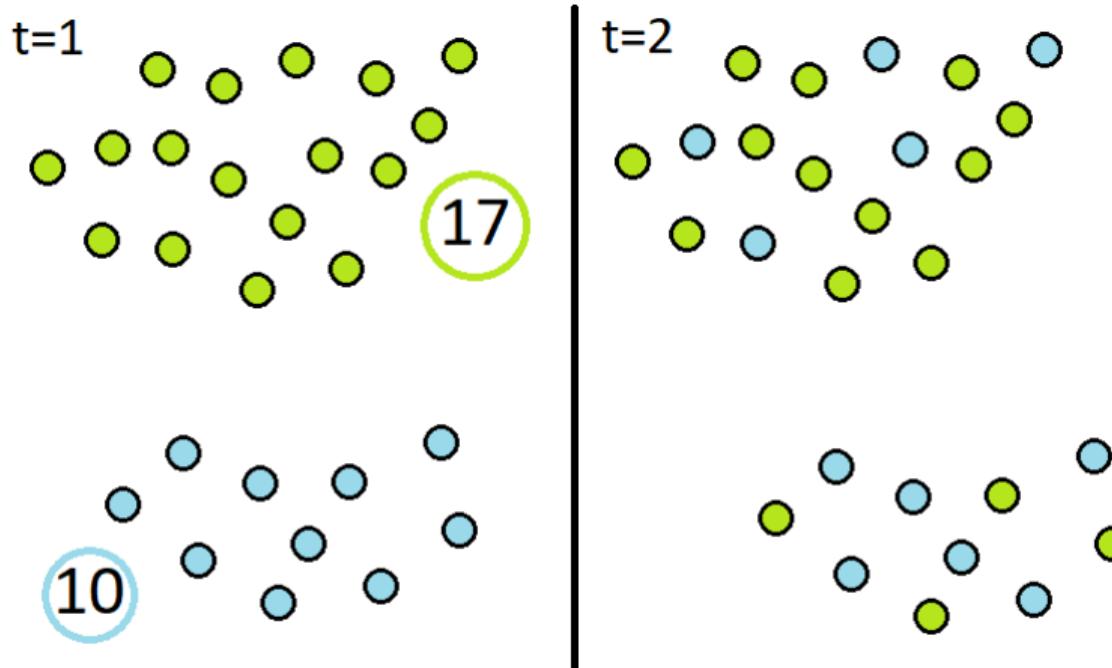
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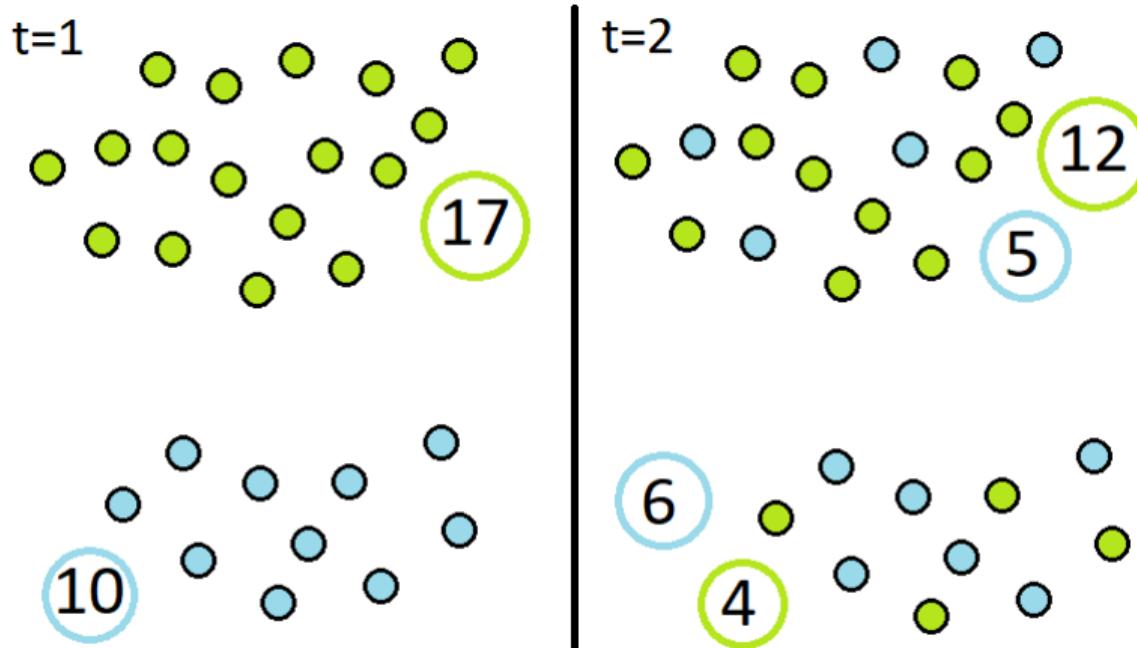
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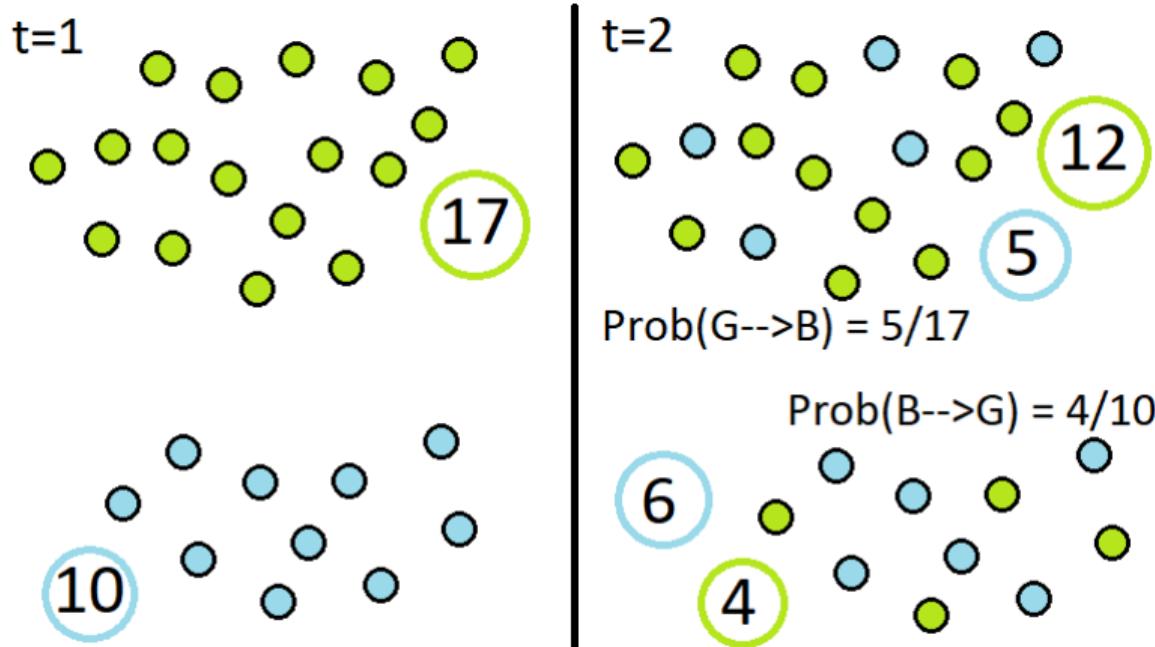
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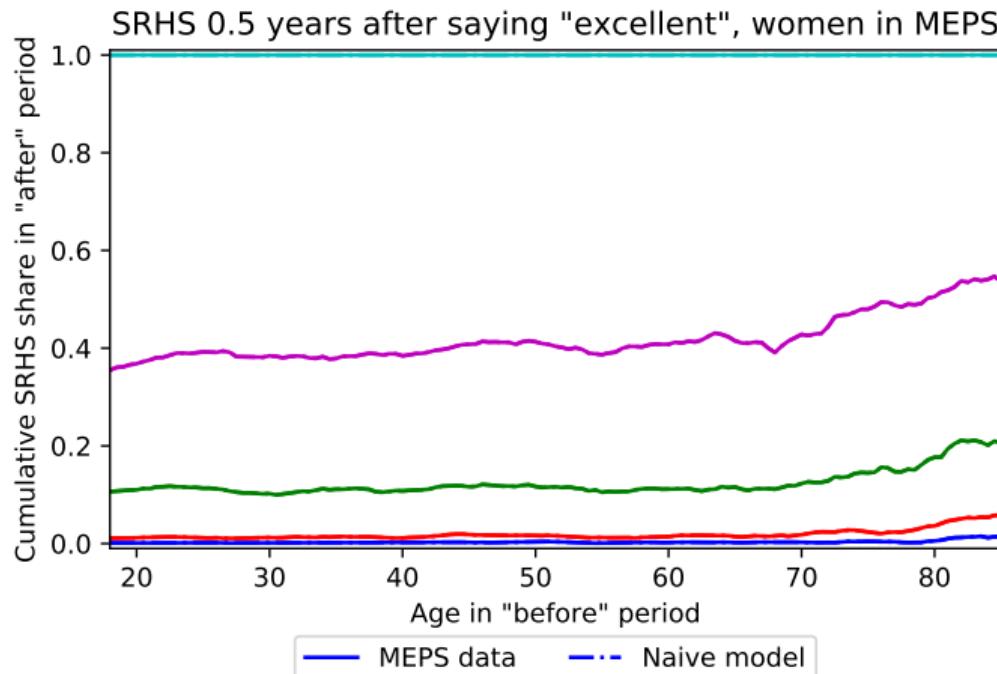
Probit or logit model:

$$y_t = \beta_0 + \beta_1 age_{t-1} + \beta_2 age_{t-1}^2 + \beta_3 sex + \beta_4 h_{t-1} + \epsilon_t,$$

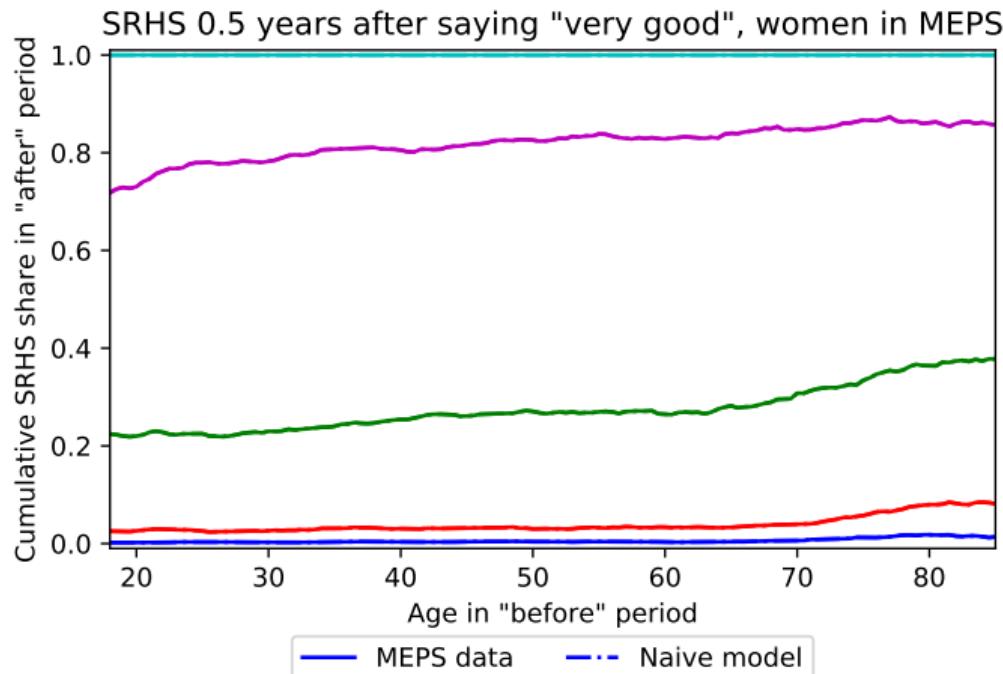
$$h_t = \begin{cases} 1 & \text{if } y_t < 0 \\ 2 & \text{if } y_t \geq 0 \end{cases}.$$

Models with more than two health states use multinomial logit (e.g. Khwaja (2010))...  
but always on one period transitions

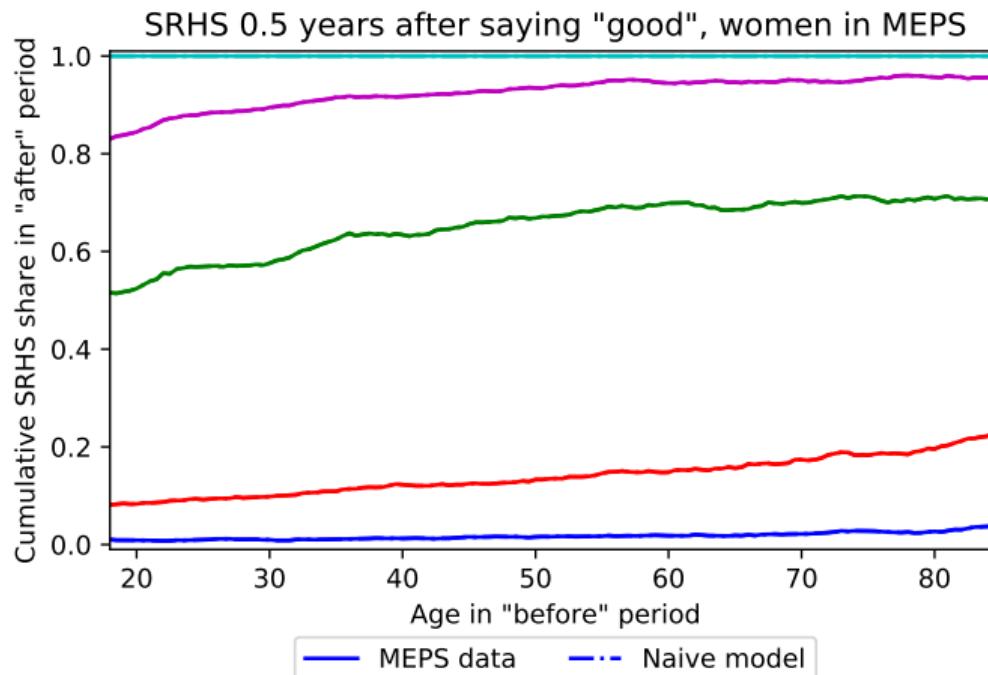
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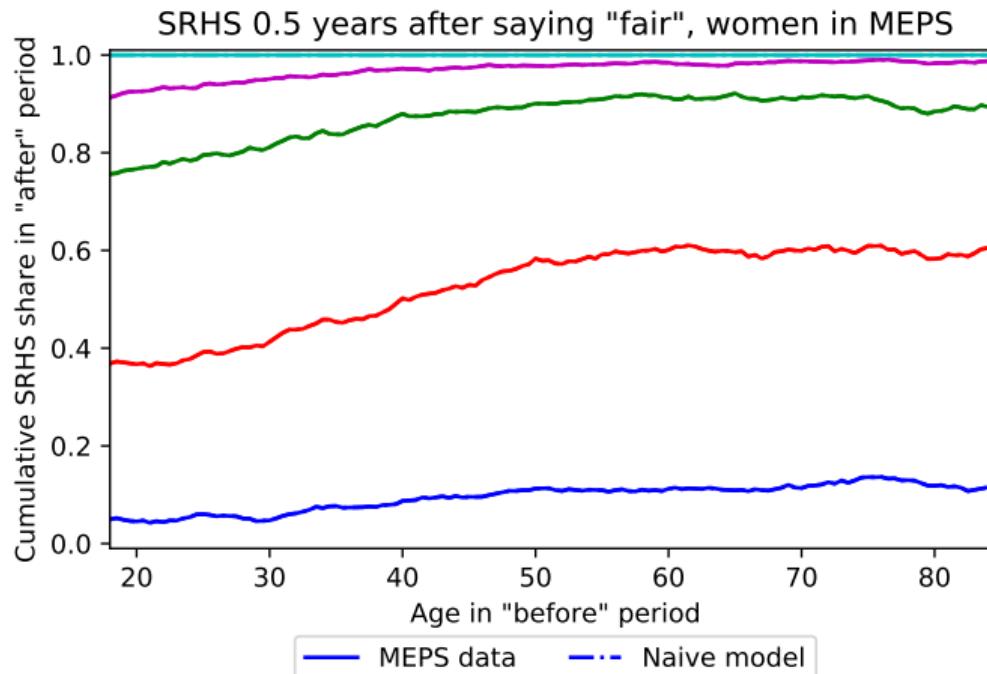
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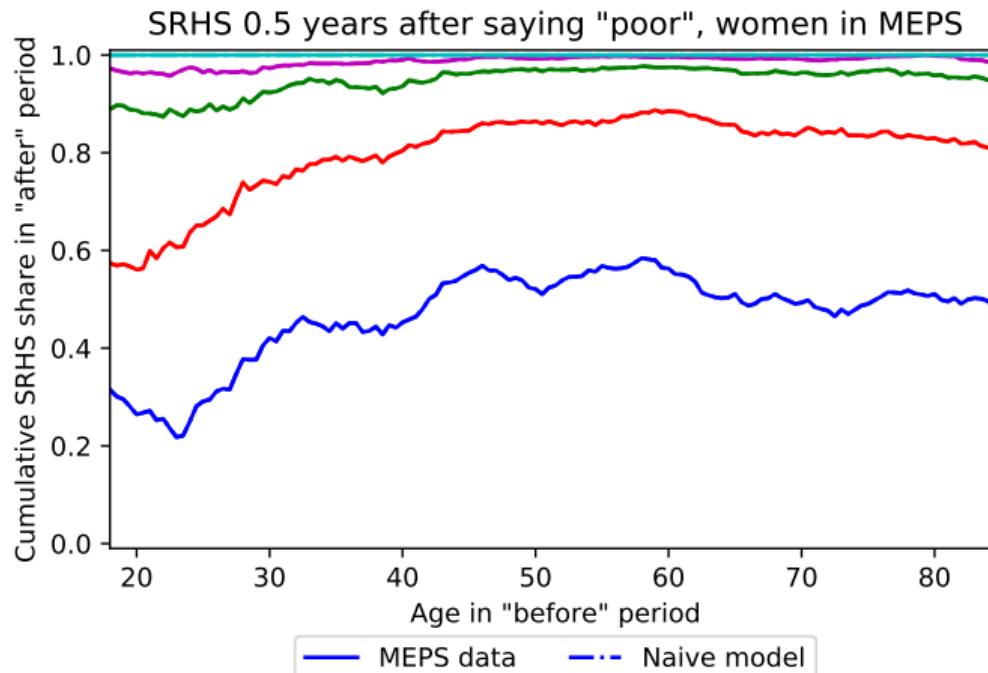
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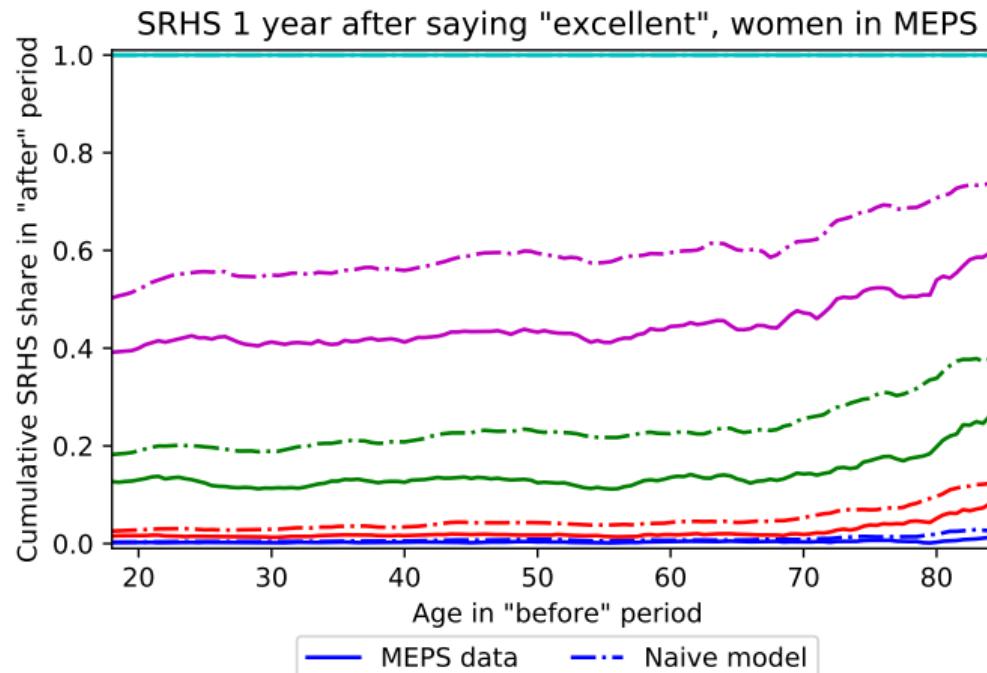
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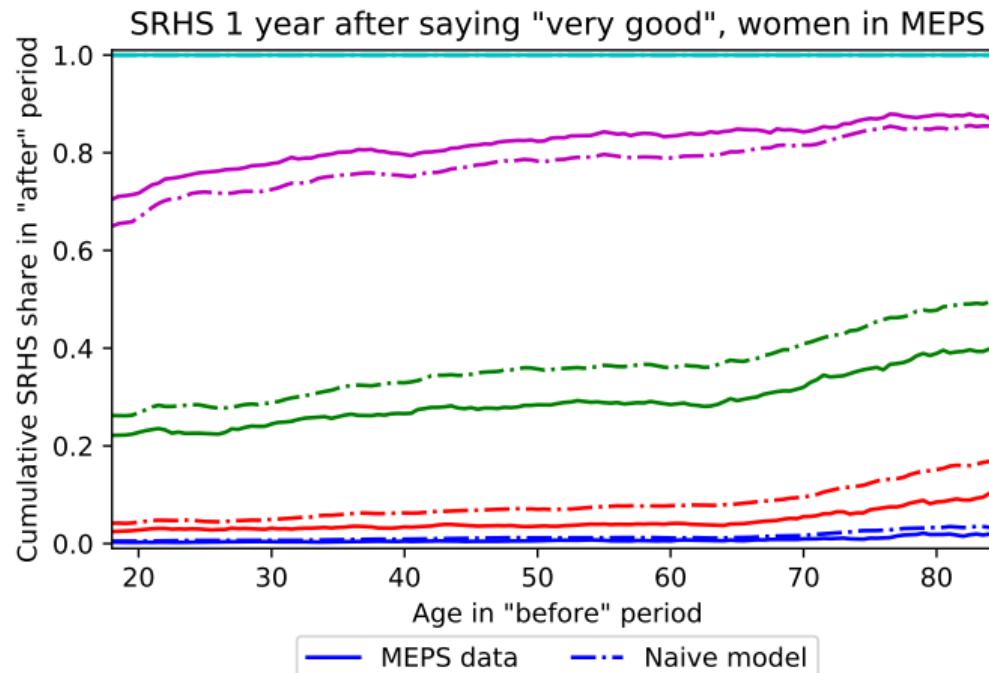
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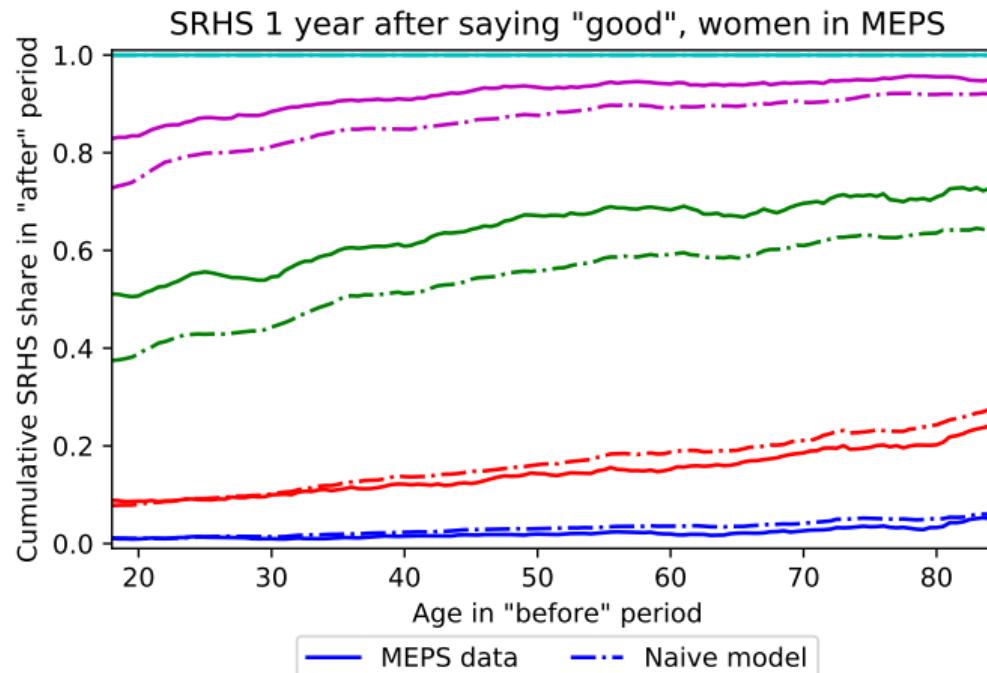
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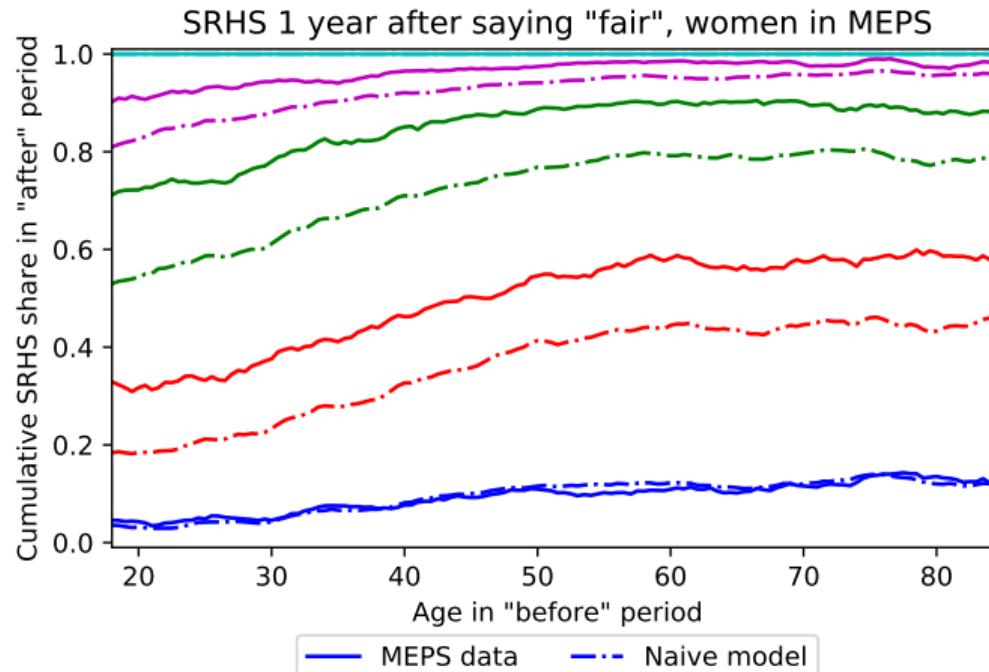
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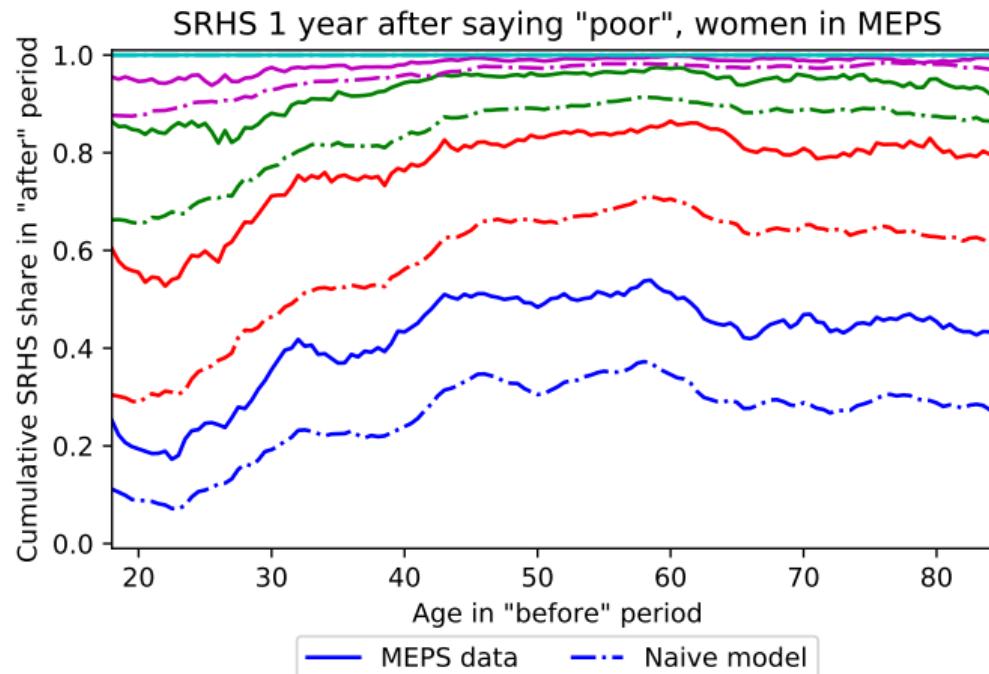
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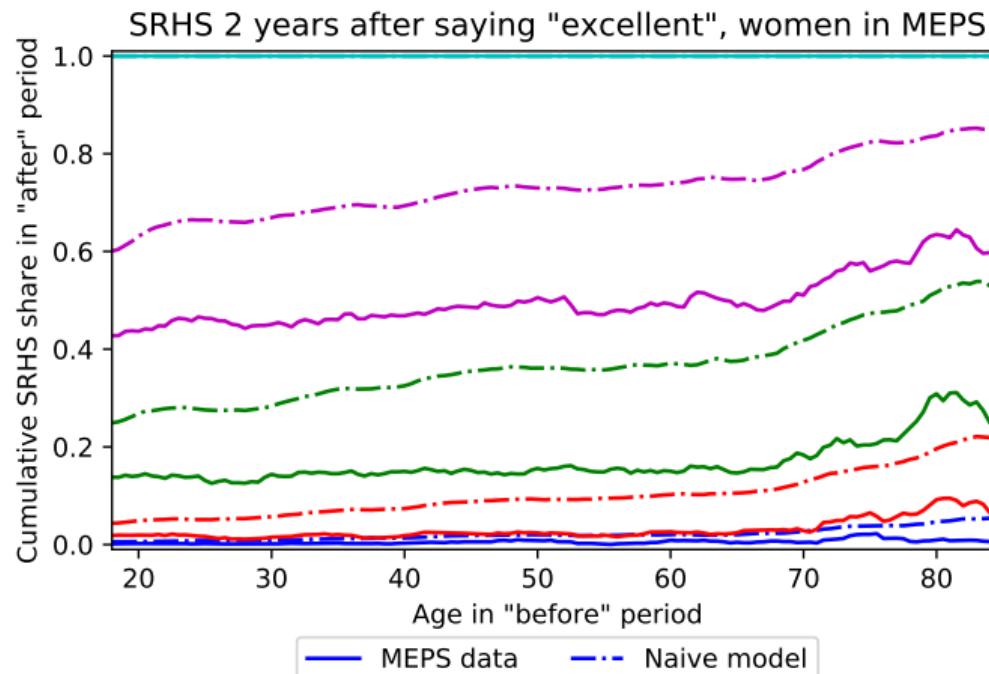
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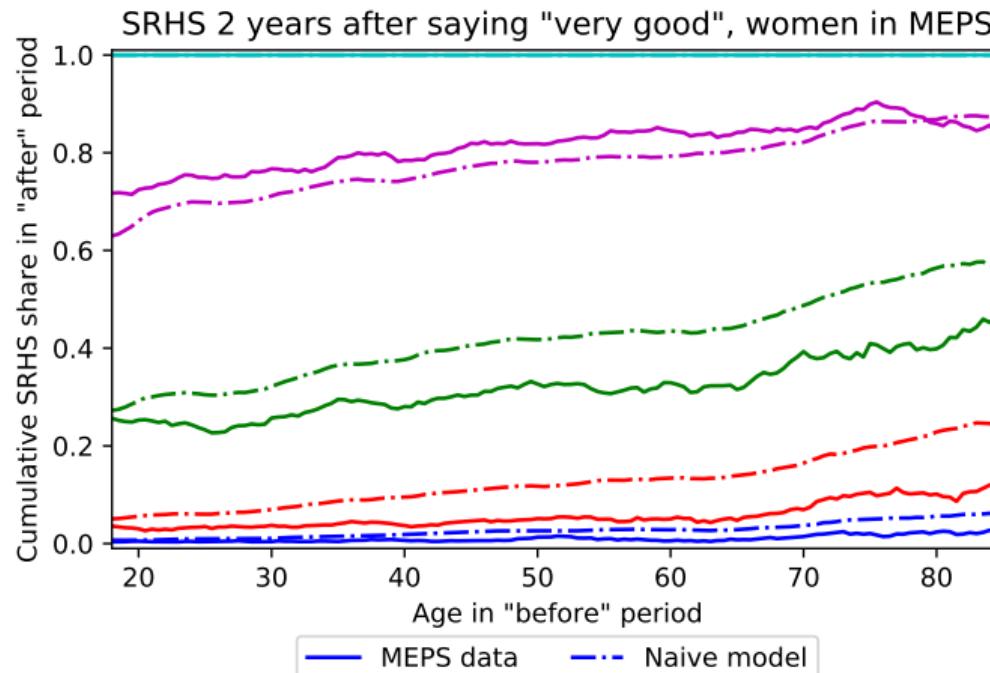
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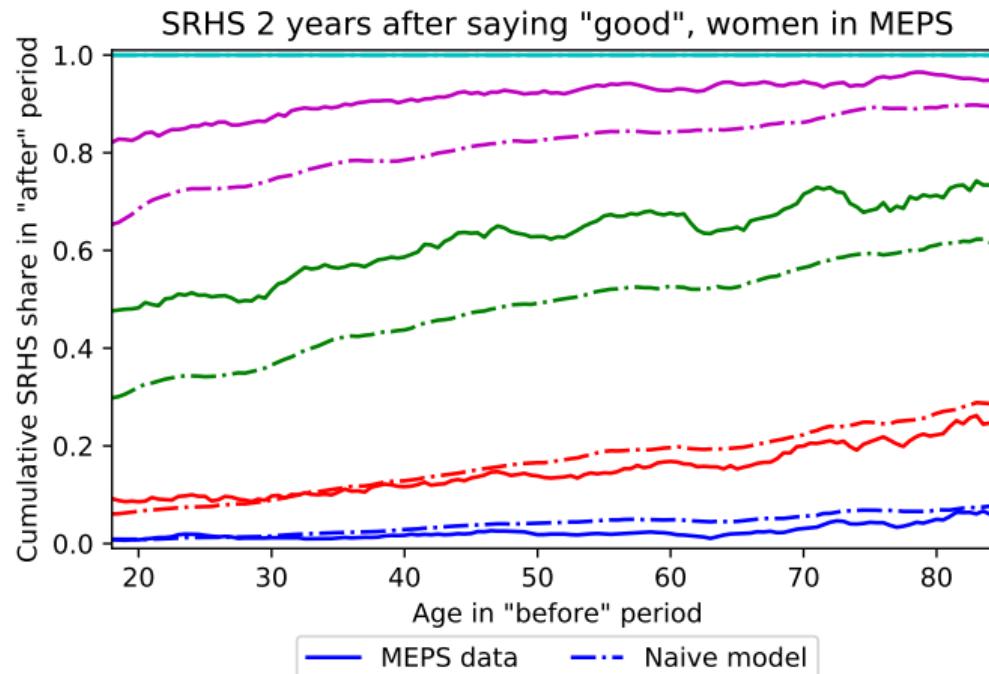
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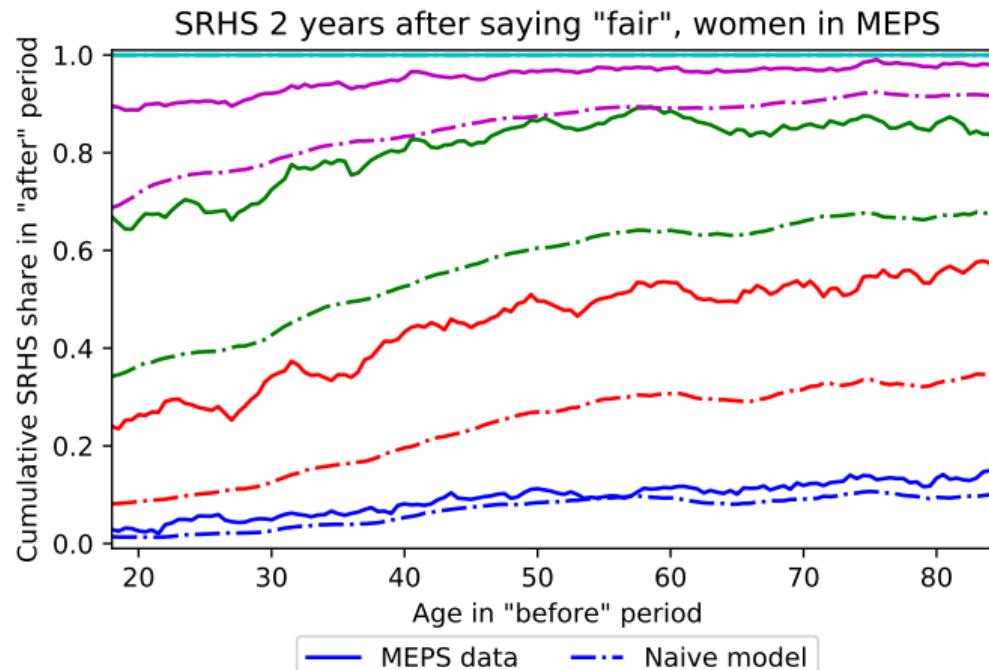
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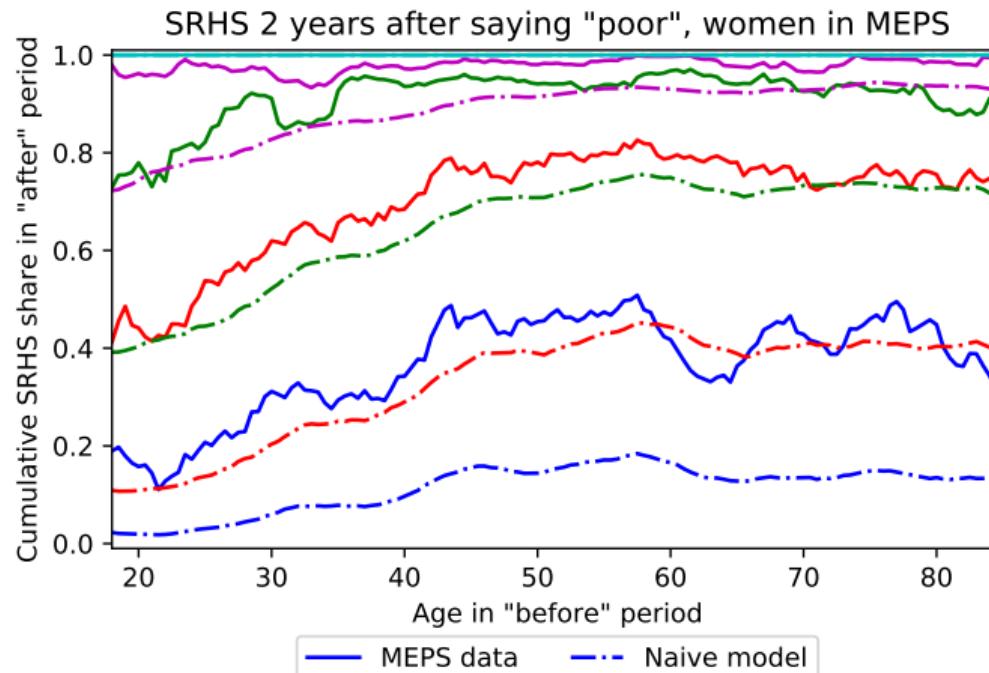
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- Naive health transitions calculated straight from the data are **obviously** a bad idea
- Who would possibly use them in “serious” work?
- What journals would publish such papers?

# Who uses naive SRHS dynamics?

“We estimate the probability of death and bad health as logistic functions of a cubic in age, sex, sex interacted with age, previous health status, health status interacted with age, a quadratic in permanent income rank, and permanent income rank interacted with age.”

— DeNardi, French, & Jones (2010), J. of Political Economy

# Who uses naive SRHS dynamics?

“In the first [estimation] step, we estimate or calibrate parameters that can be cleanly identified without explicitly using our model. For example, we estimate mortality rates and health transitions straight from demographic data.”

— French & Jones (2011), *Econometrica*

# Who uses naive SRHS dynamics?

“To construct the transition matrix we measure the fraction of people who move from one bin to another between two consecutive years”

— Pashchenko & Porapakkarm (2013), Review of Economic Dynamics

# Who uses naive SRHS dynamics?

“[S]ome parameters are estimated outside the structure of the model. For some parameters, this is because no structure is needed: disability risk can be estimated directly from transitions between disability states because of the exogeneity assumption.”

— Low & Pistaferri (2015), American Economic Review

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“We estimate health transitions and mortality rates simultaneously by fitting the transitions observed in the HRS to a multinomial logit model [...] on age, sex, current health status, and PI.”

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# Who uses naive SRHS dynamics?

"We estimated the transition probabilities using the logit method. We regressed next period's health status on a constant, age, age squared, age cubic, education level, current health status, and age times current health status."

— Ferreira & Gomes (2017), Review of Economic Dynamics

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“The transition to next period health status is determined as a logistic function of [math notation].”

—Aizawa (forthcoming), Quantitative Economics

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- Problem: That's just not true

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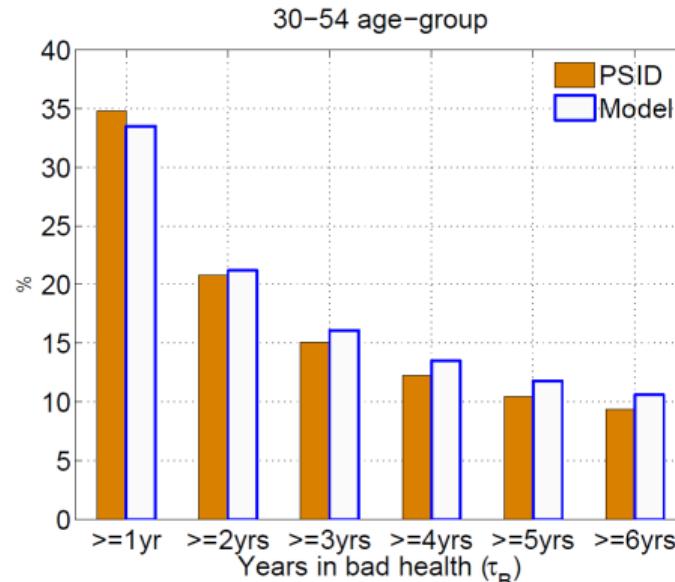
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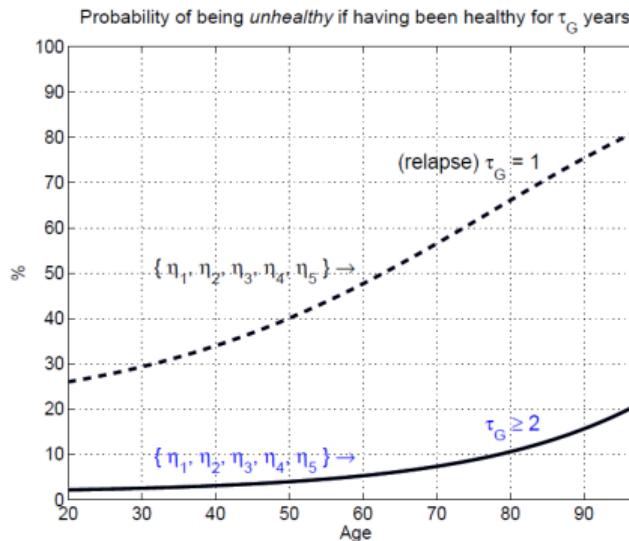
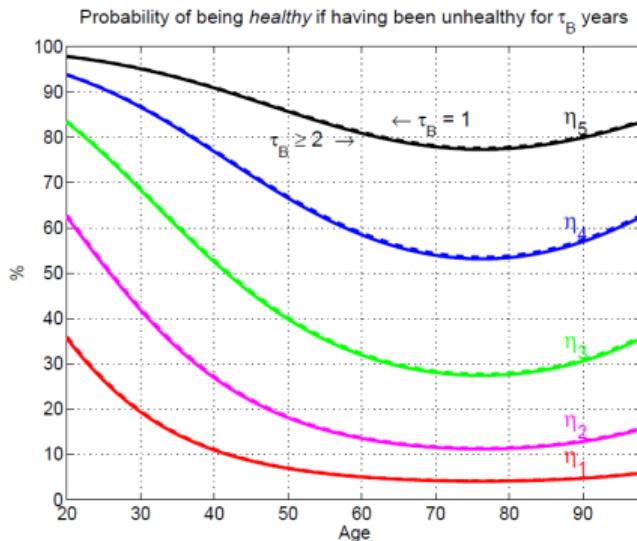
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- **Lagged SRHS generates observably different medical expense distributions; “unhealthy” people are not all in the same health state!**

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- **Wave 1 SRHS is nearly as predictive of W5 SRHS as it is for W2 SRHS!**

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- Income quintile is a discretization of continuous income
- SRHS is a discretization of continuous “health”

## But it's not duration dependence (3/3)

- Suppose you saw these examples, but the labels were “in bottom income quintile” and “above bottom income quintile”
- You would not hypothesize: “There must be duration dependence in being poor!”
- Income quintile is a discretization of continuous income
- SRHS is a discretization of continuous “health”
- Income changes with both permanent and transitory shocks
- SRHS changes with both changes in “true health” and transitory **reporting** shocks

# MODEL

# Graphical model of latent health and SRHS (1/5)

“True health” lives in some weird polydimensional space



## Graphical model of latent health and SRHS (2/5)

Dynamics of “true health” really **are** Markov(1)



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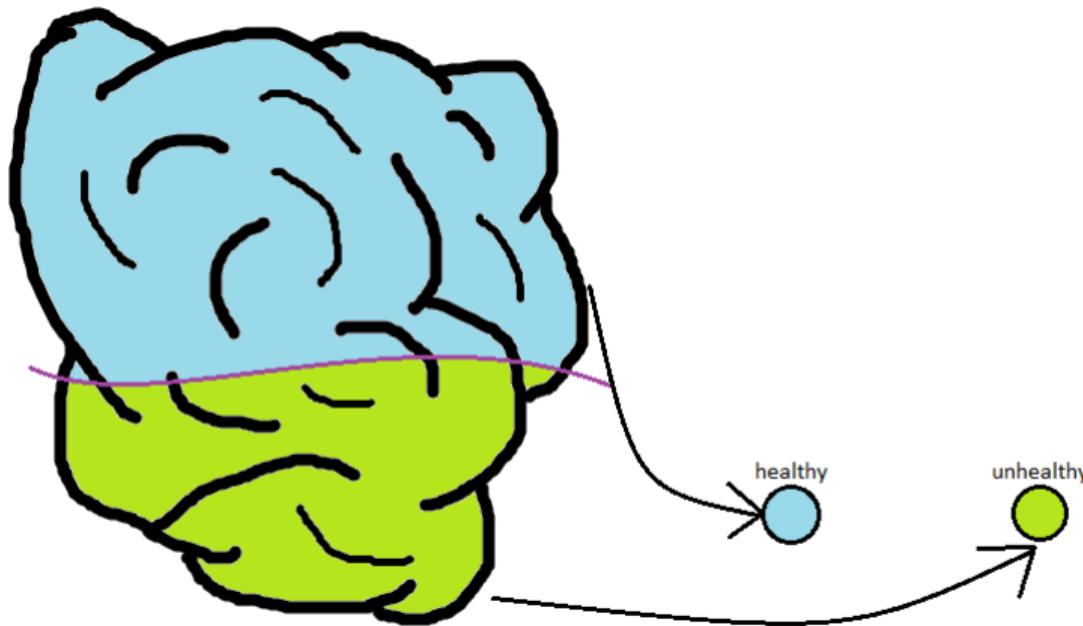
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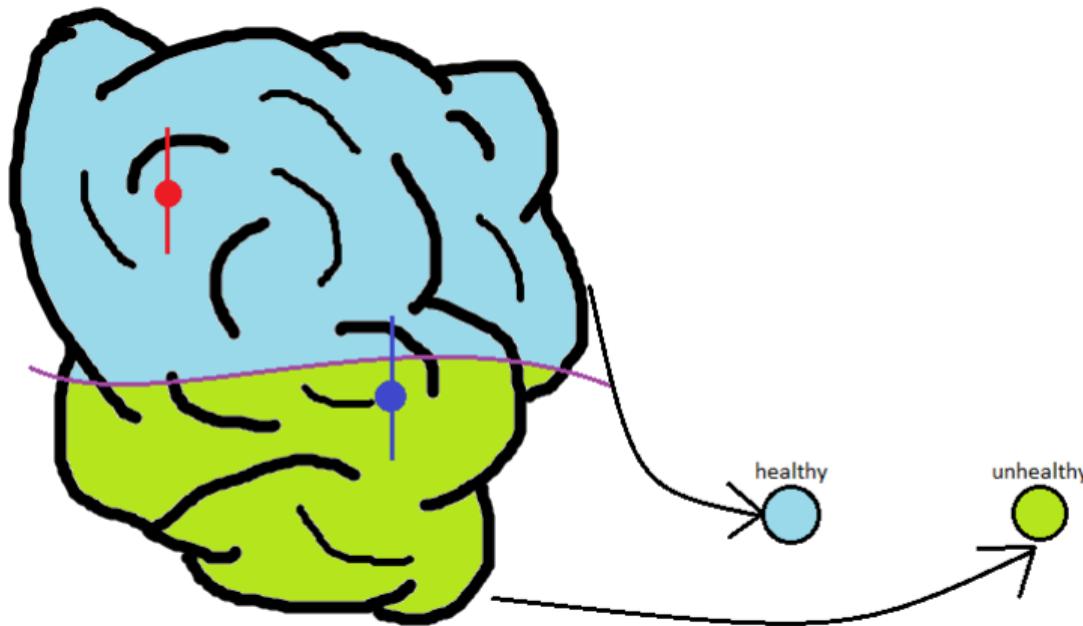
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SRHS is a projection of that polydimensional space onto a discrete space...



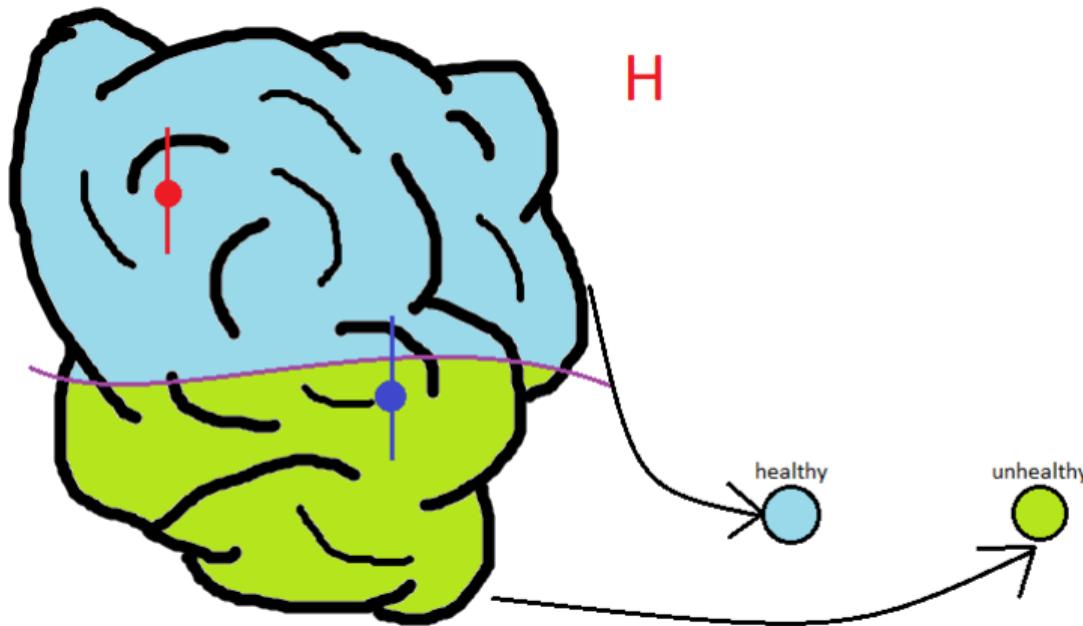
## Graphical model of latent health and SRHS (3/5)

SRHS is a projection of that polydimensional space  
onto a discrete space... with reporting error



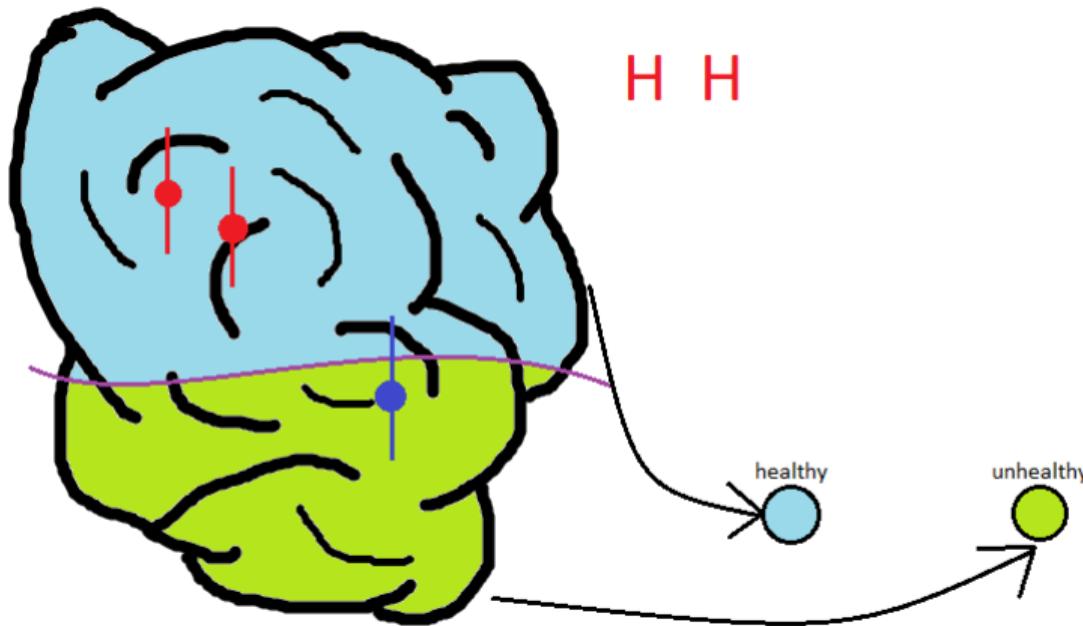
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Individuals whose “true health” is far from the healthy/unhealthy threshold are likely to have the same SRHS period after period



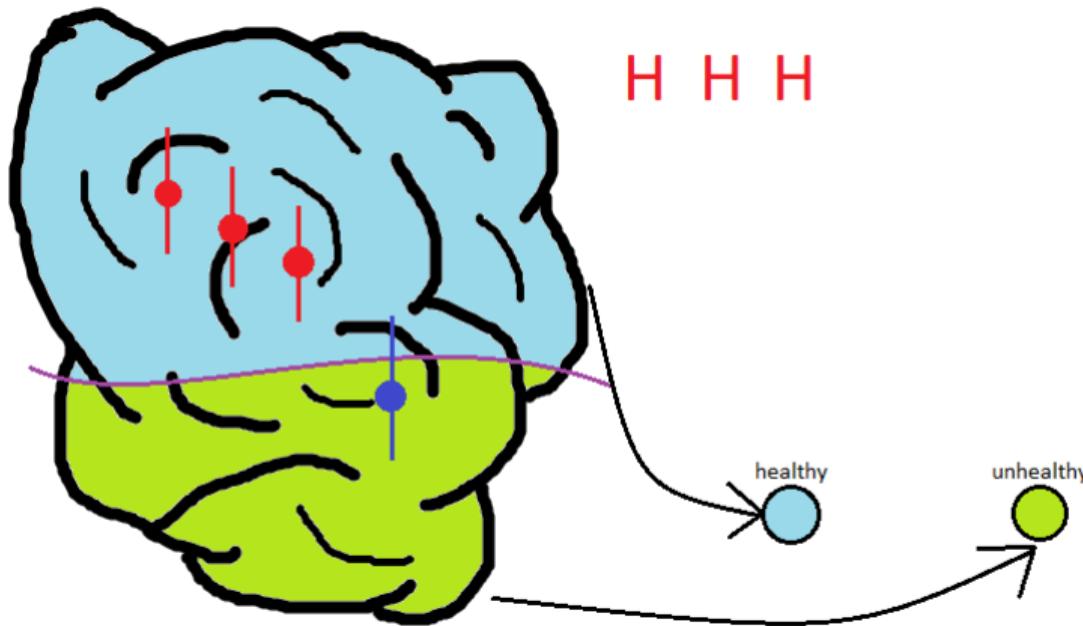
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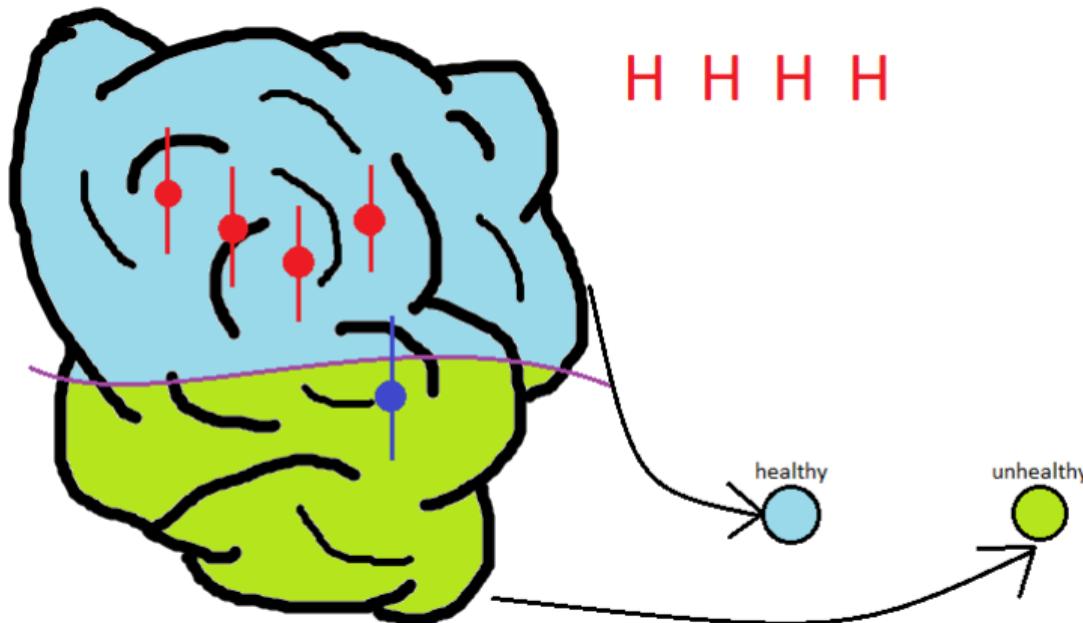
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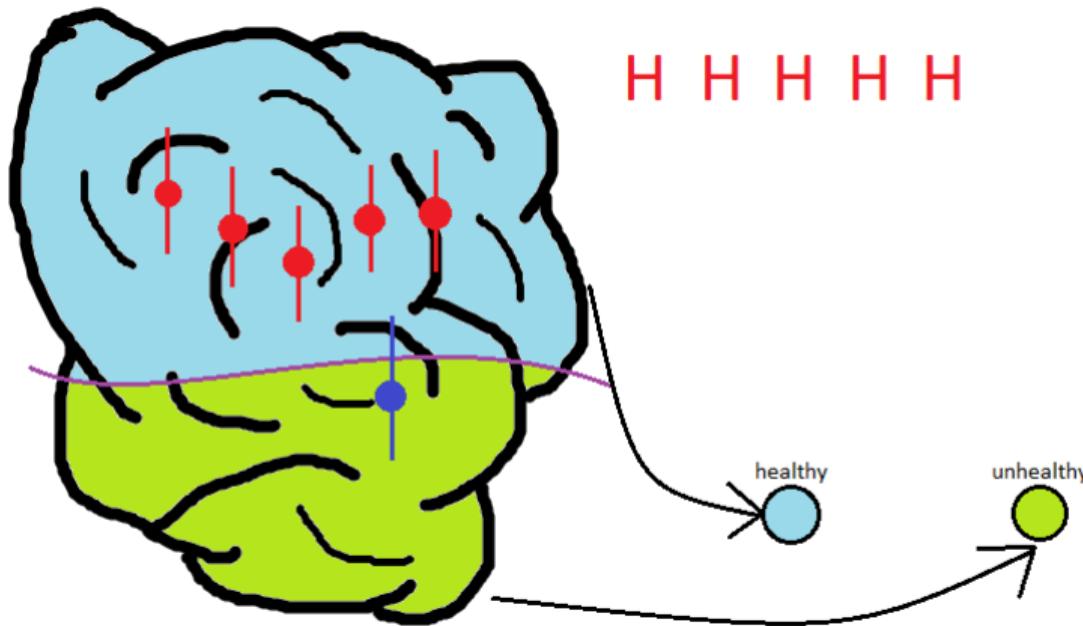
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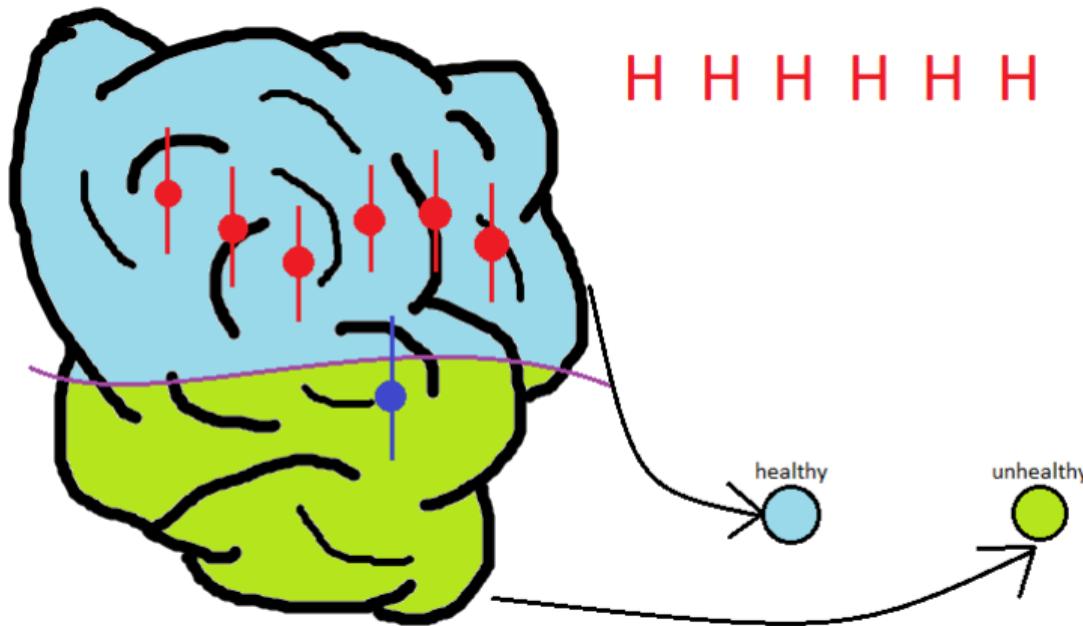
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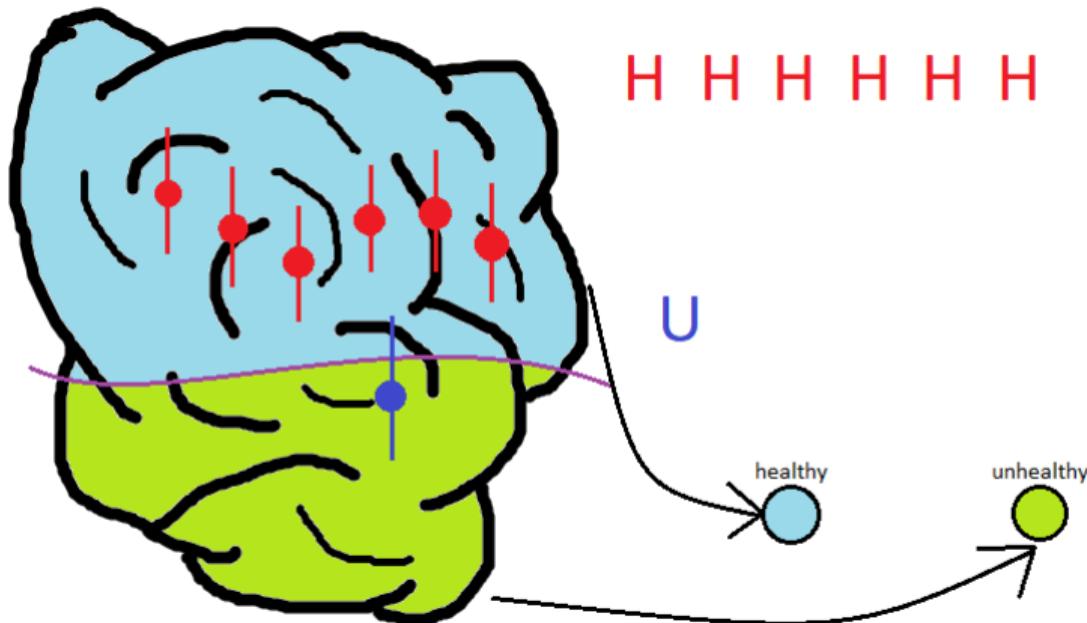
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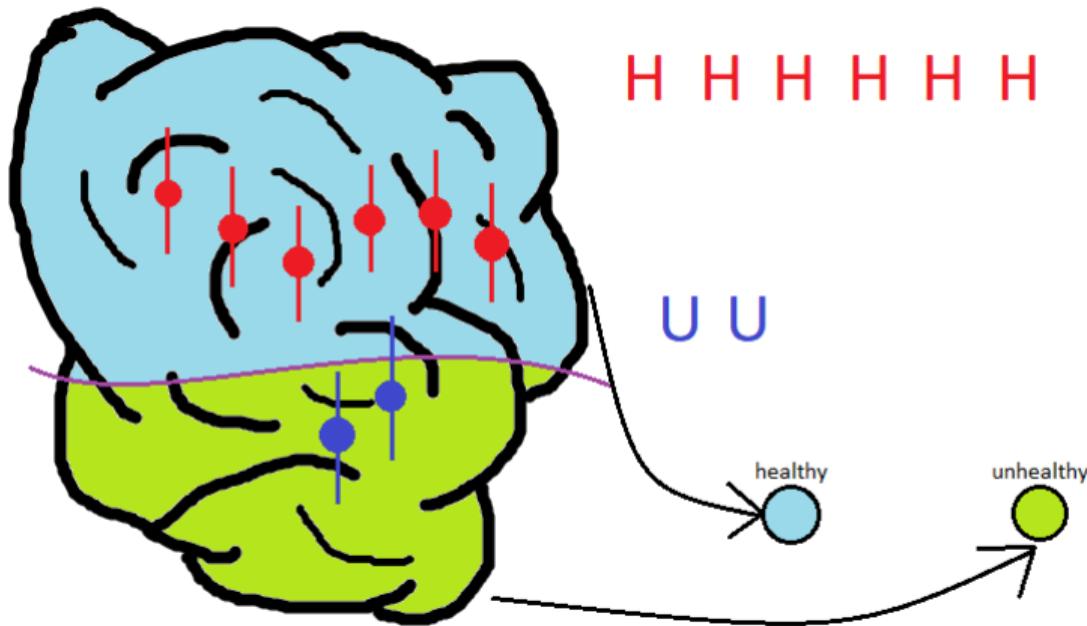
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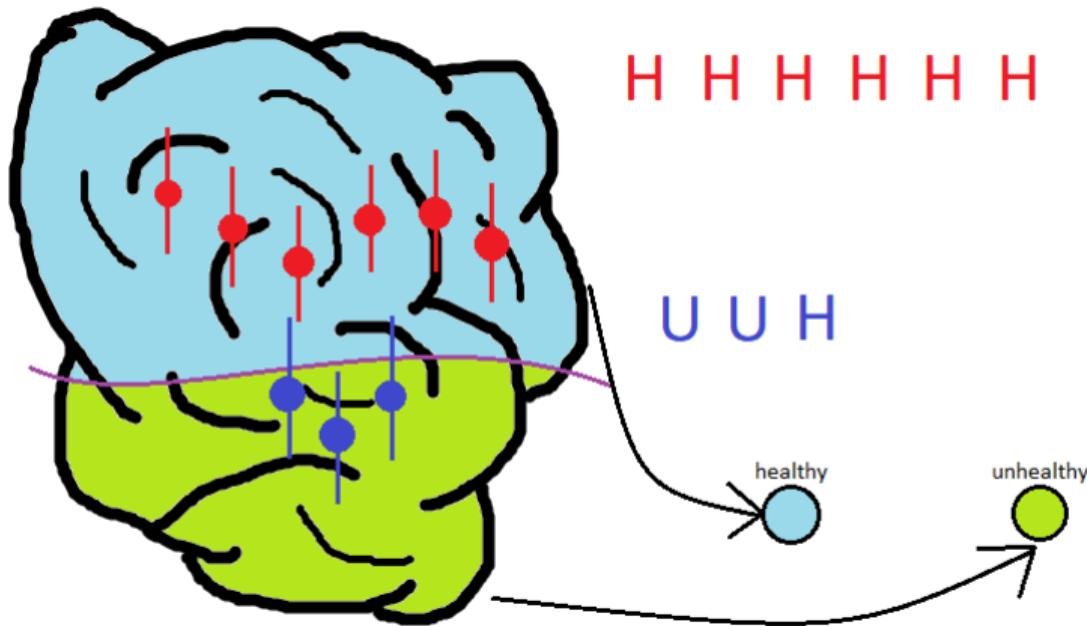
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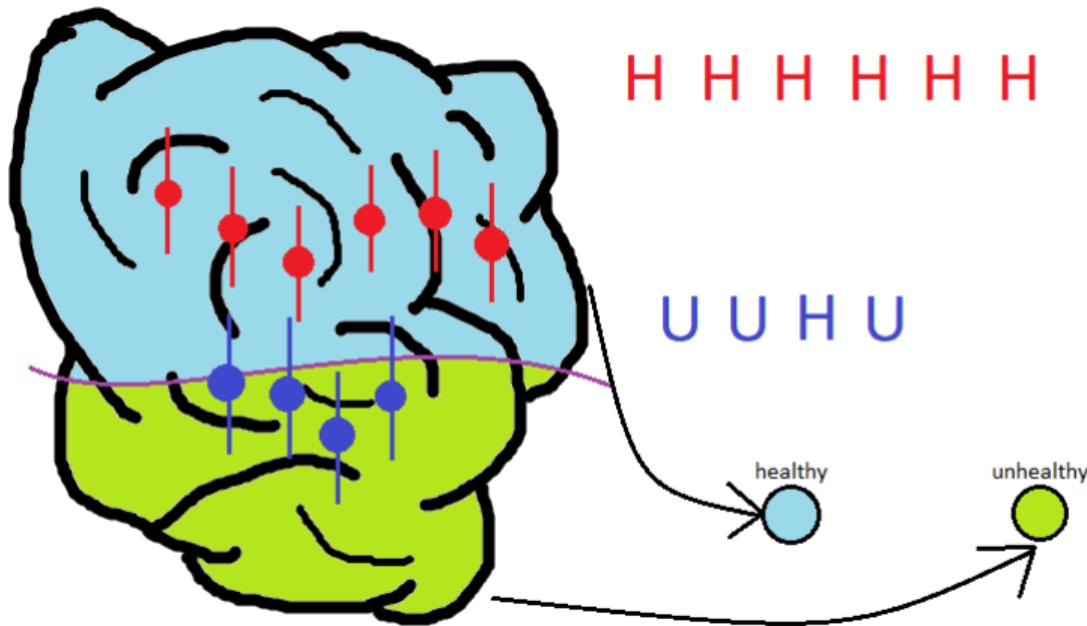
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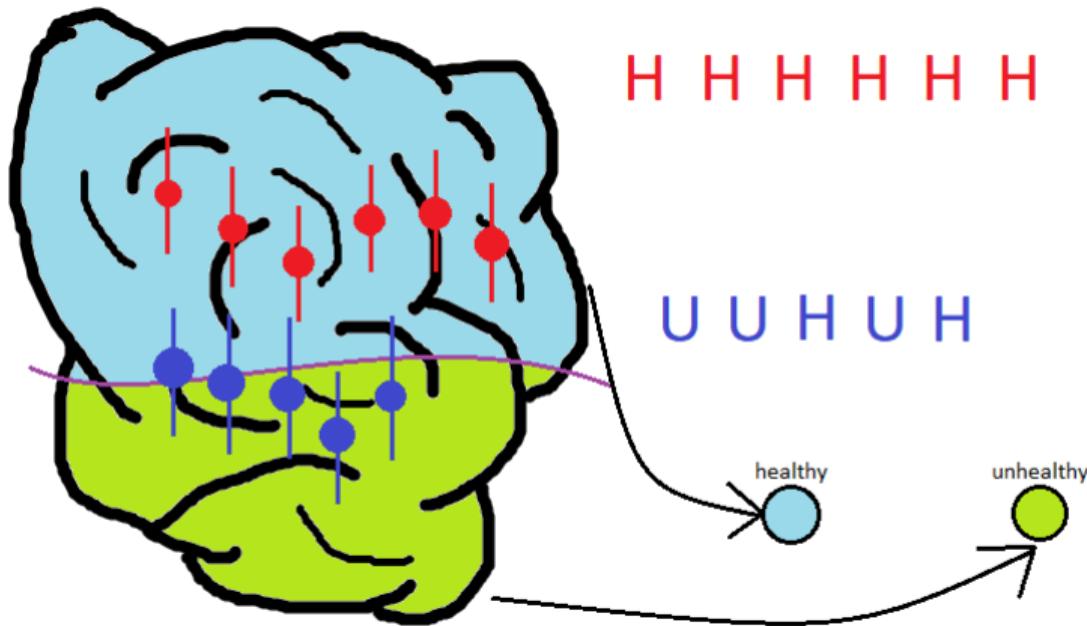
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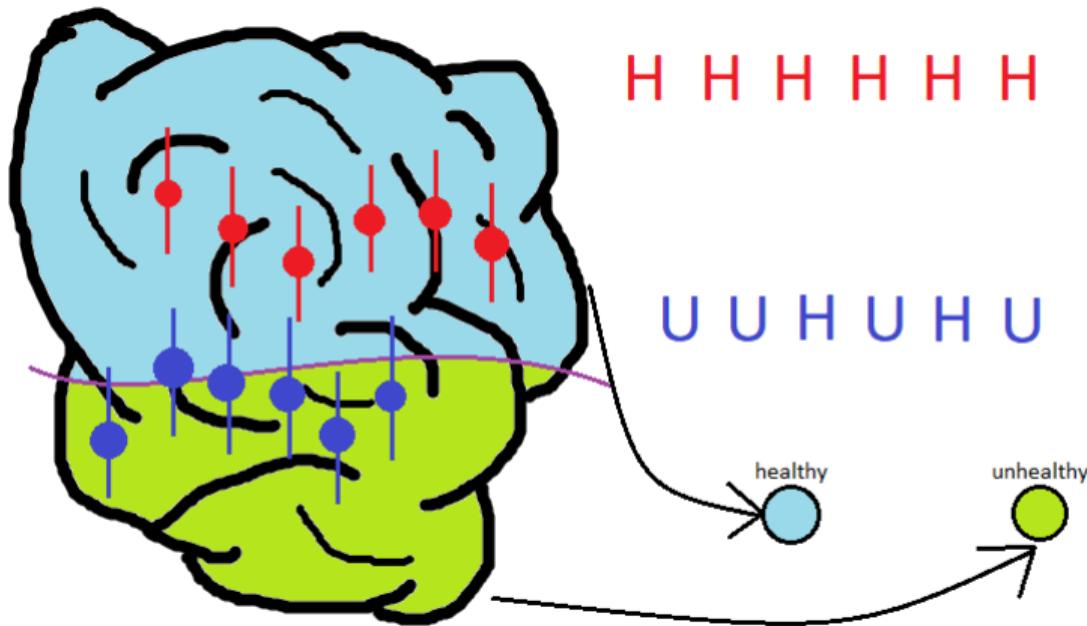
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$$h_t = 1 + \sum_{k=1}^{K-1} \mathbf{1}(h_t^* \geq \chi_k).$$

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- Early 40's women: 37-38% change SRHS from their answer a couple weeks prior if reported poor to very good health; 44% change answer if reported excellent health

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- All parameters can differ between men and women.
- No reporting error for mortality.
- No selection in attrition or missing SRHS.

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- Using (1-3), can precompute **unconditional** pmf of  $x_t$  at every age (i.e. simulation by multiplication)

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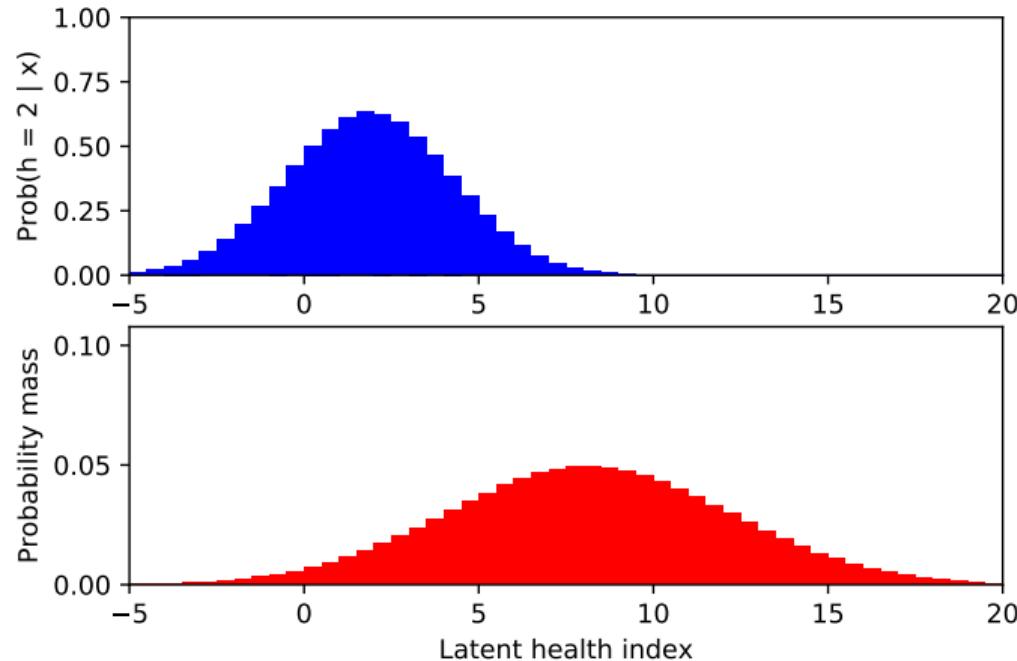
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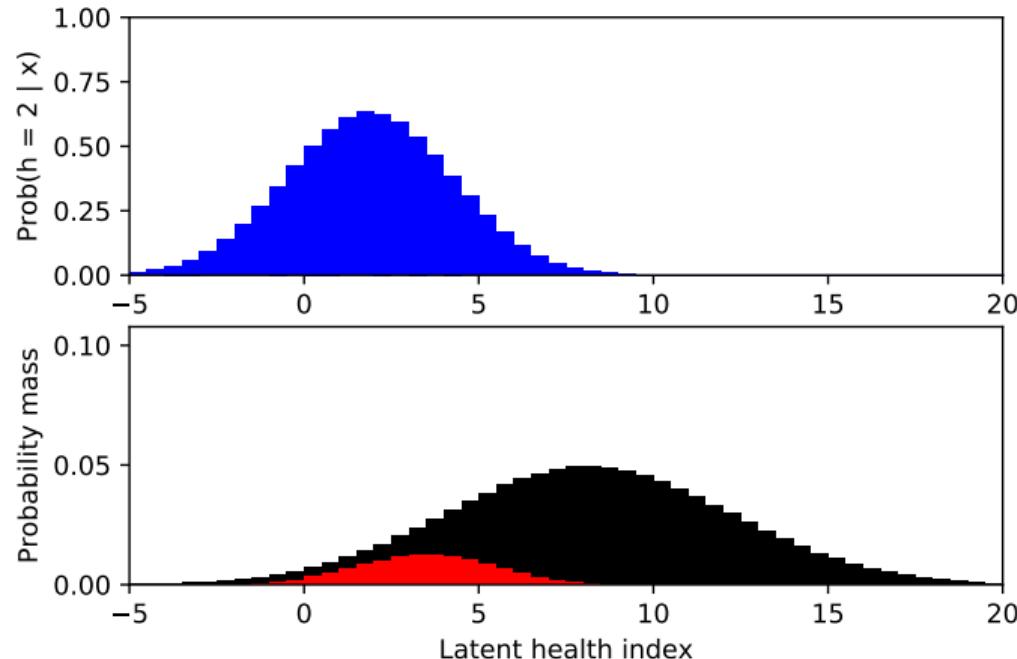
# Log likelihood function: self-reported health status

Latent health: distribution at age 48



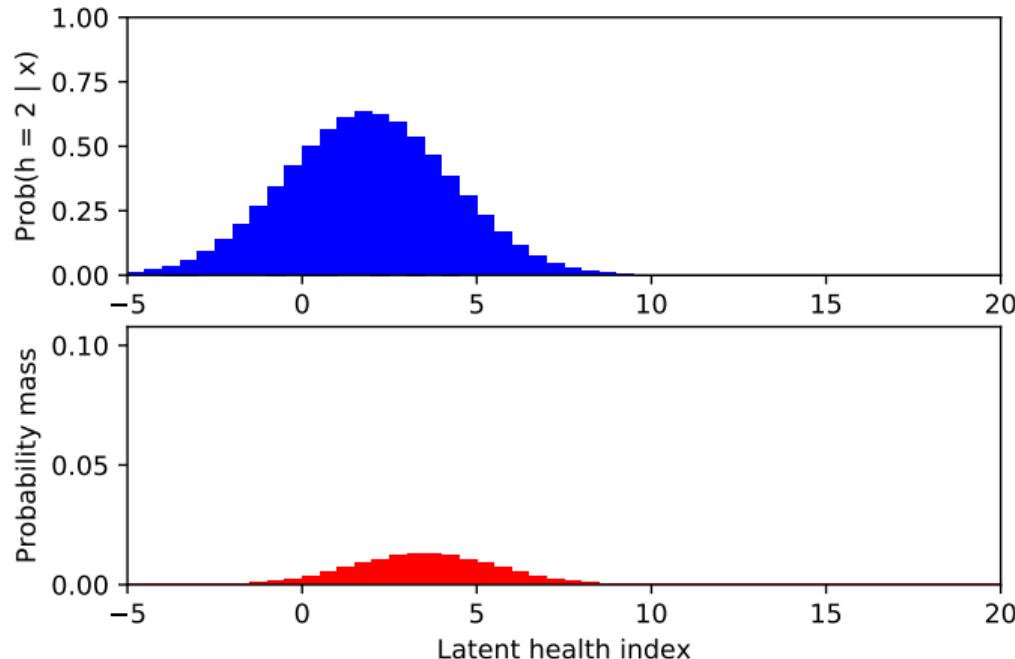
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Latent health: observing  $h=2$  at age 48



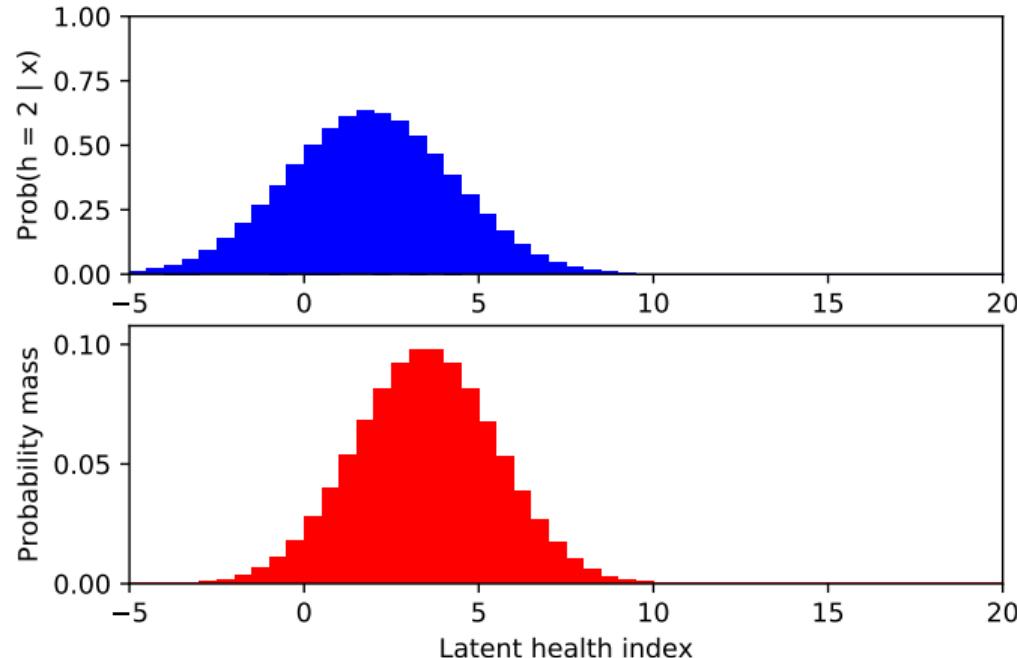
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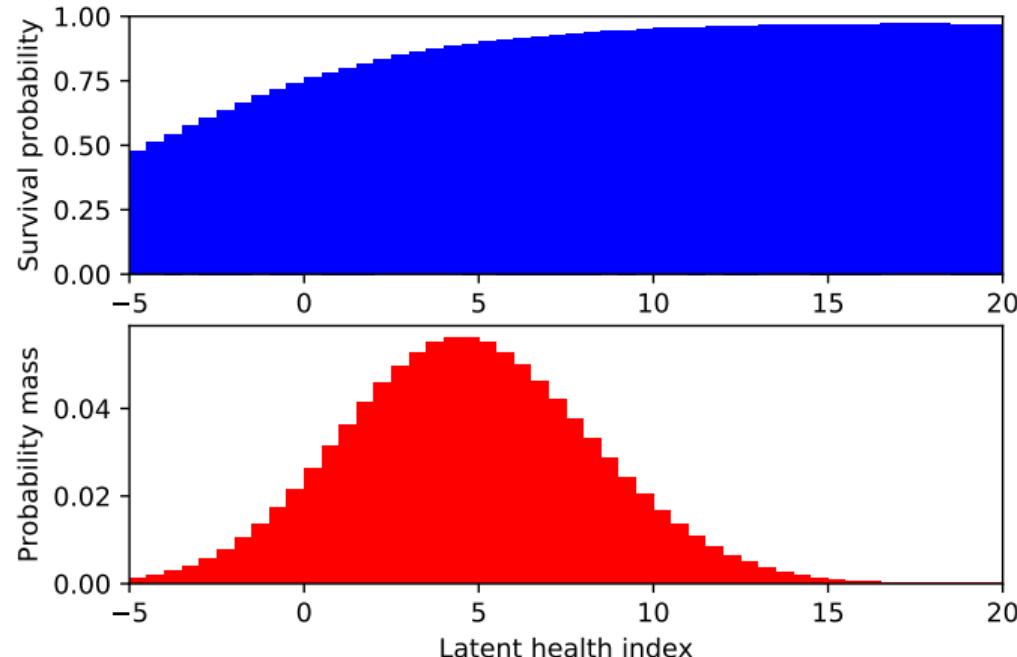
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Latent health: SRHS adjustment at age 48



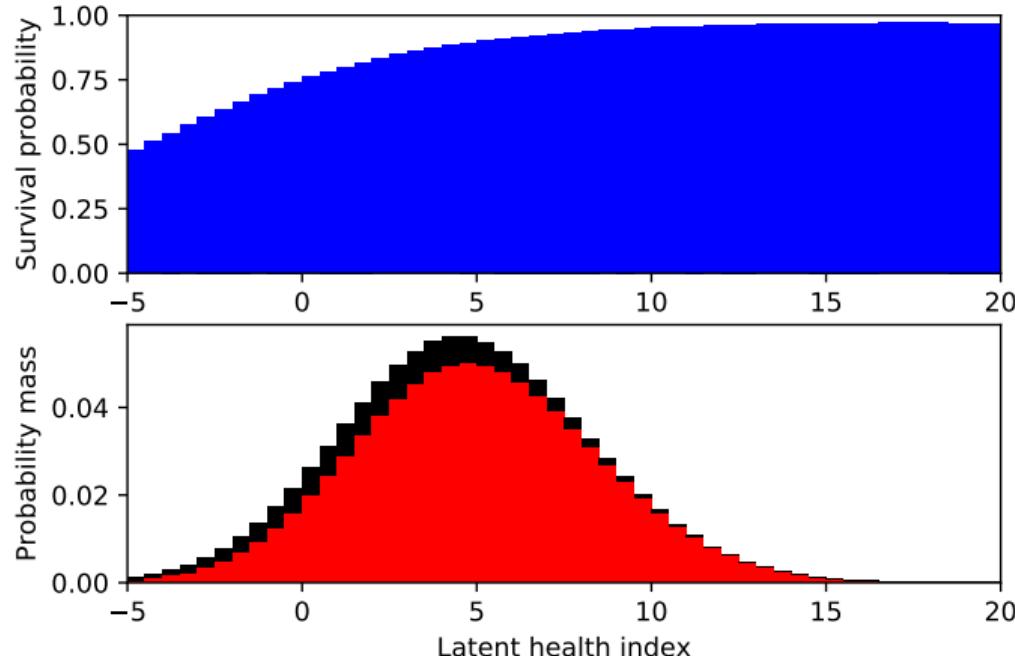
# Log likelihood function: mortality

Latent health: distribution at age 90



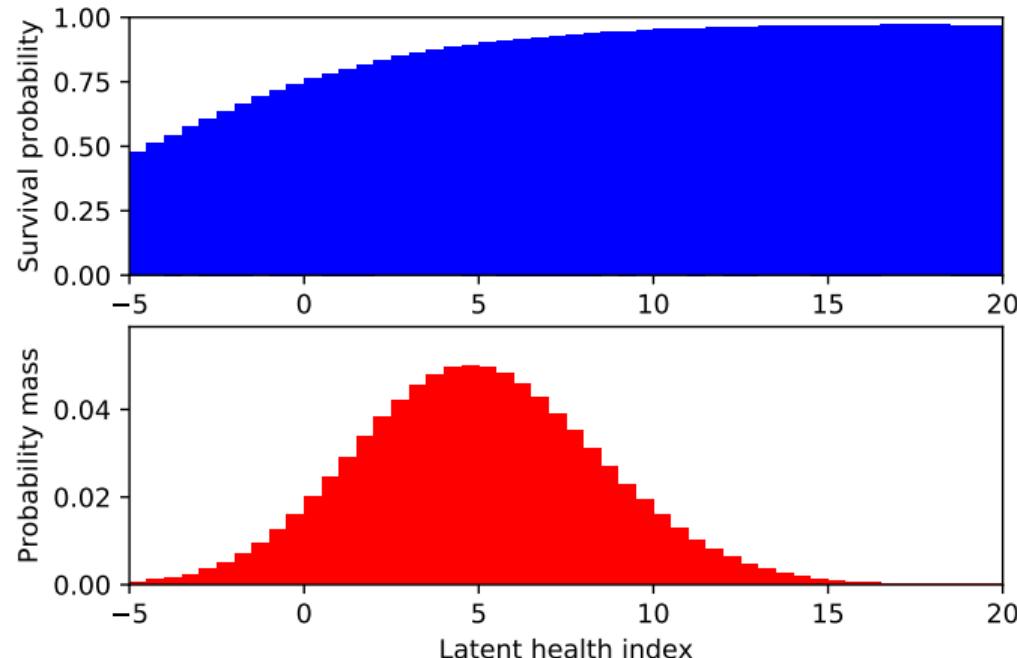
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Latent health: who dies at age 90

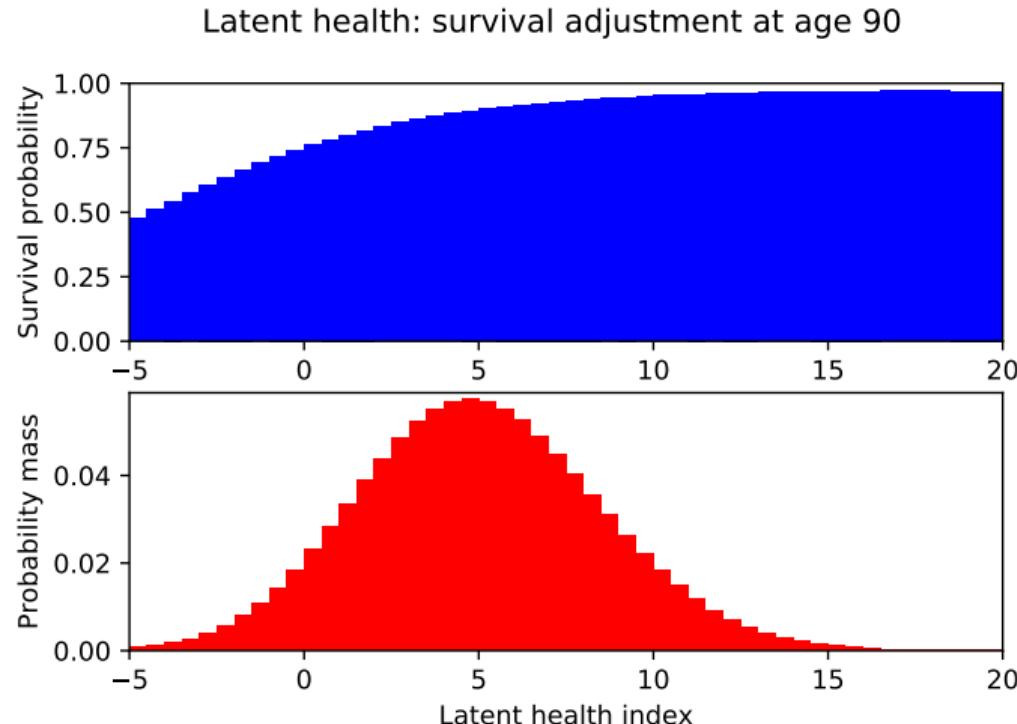


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Latent health: after mortality at age 90

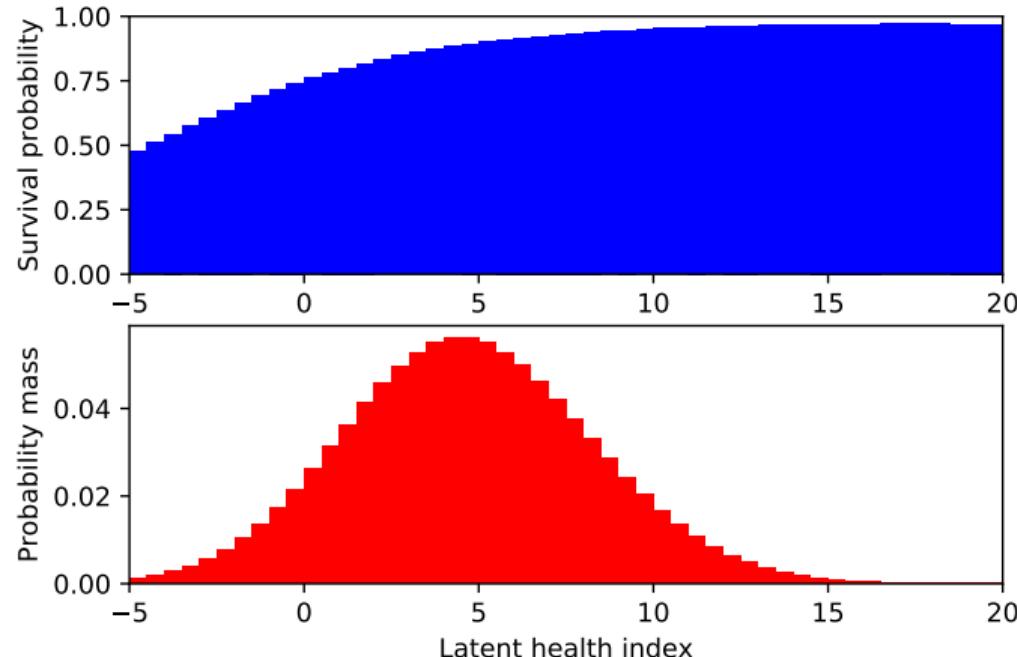


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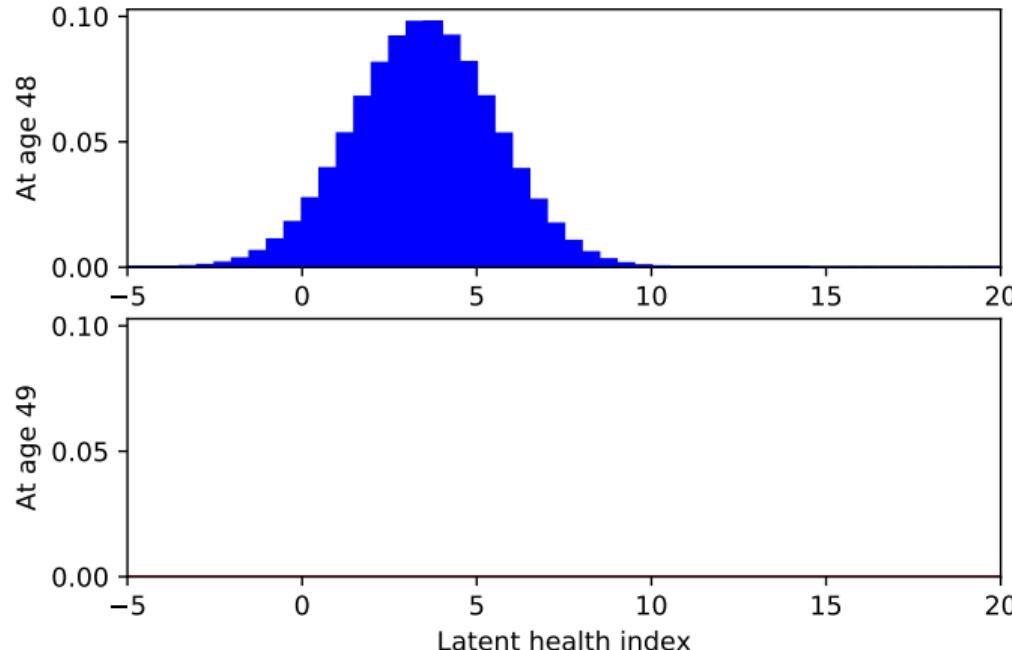
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# Log likelihood function: latent health transitions

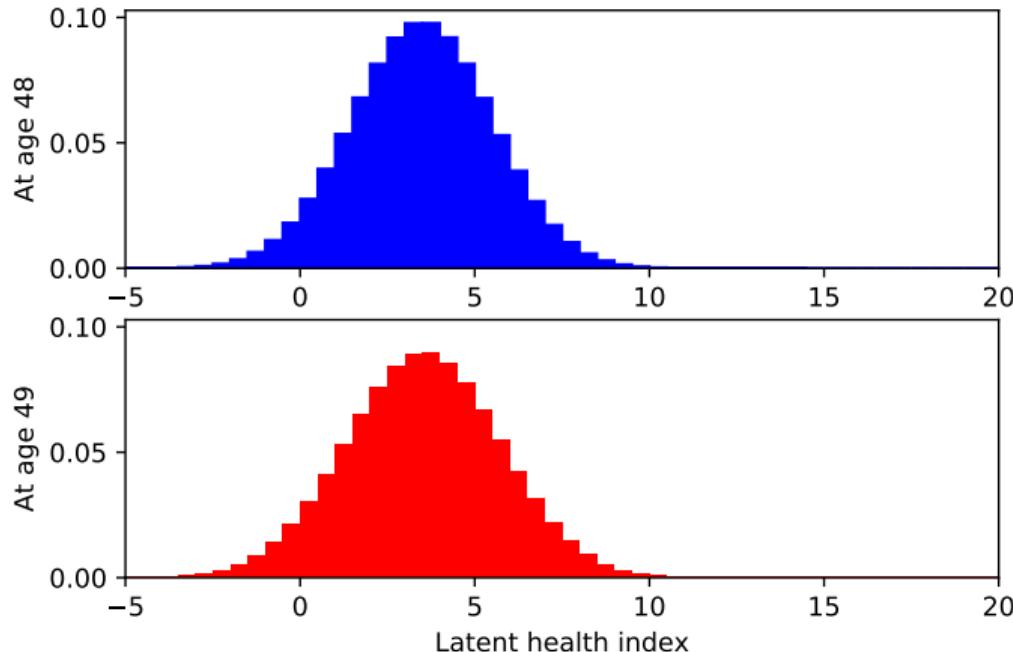
Probability distribution of latent health from age 48 to 49



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- Need more than 2 SRHS categories to do this!

## Identification strategy (2/2)

- Rest of the parameters are straightforward to identify
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- Identify mortality params  $\theta$  off of deaths by age and SRHS
- Identify expected health parameters  $\beta$ , initial health distribution  $(\mu_0, \sigma_0^2)$ , and probit cut points  $\chi$  from distribution of SRHS by age

# DATA

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- Relabel age to have half-years based on age in W1 and W2
- 126,486 women observed for 601,713 waves (1,168 deaths)
- 110,044 men observed for 518,577 waves (1,314 deaths)
- 89% of respondents observed all 5 waves; most others drop after 3 (maybe planned drops?)

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- Observe SRHS every 2 years; recoded to 1 year periods
- 20,108 women observed for 107,542 waves (5,794 deaths)
- 16,094 men observed for 81,091 waves (5,286 deaths)
- 50% observed alive at least 5 waves, 11% obs all 11 waves

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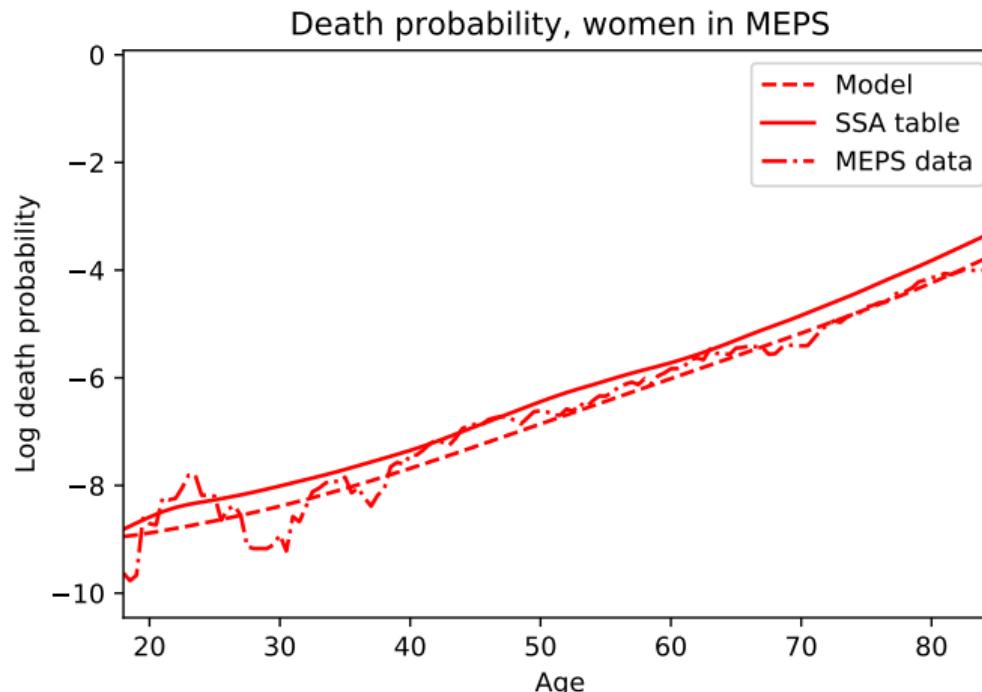
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- SRHS **only** asked of HH heads and spouses
- Selection bias: college students aren't heads yet!
- Enter my data at age 23; also drop R's who start 90+
- 11,787 women observed for 69,329 waves (1,012 deaths)
- 10,791 men observed for 58,316 waves (995 deaths)
- 32% observed 2 waves or fewer, but 22% obs all 11 waves

# RESULTS

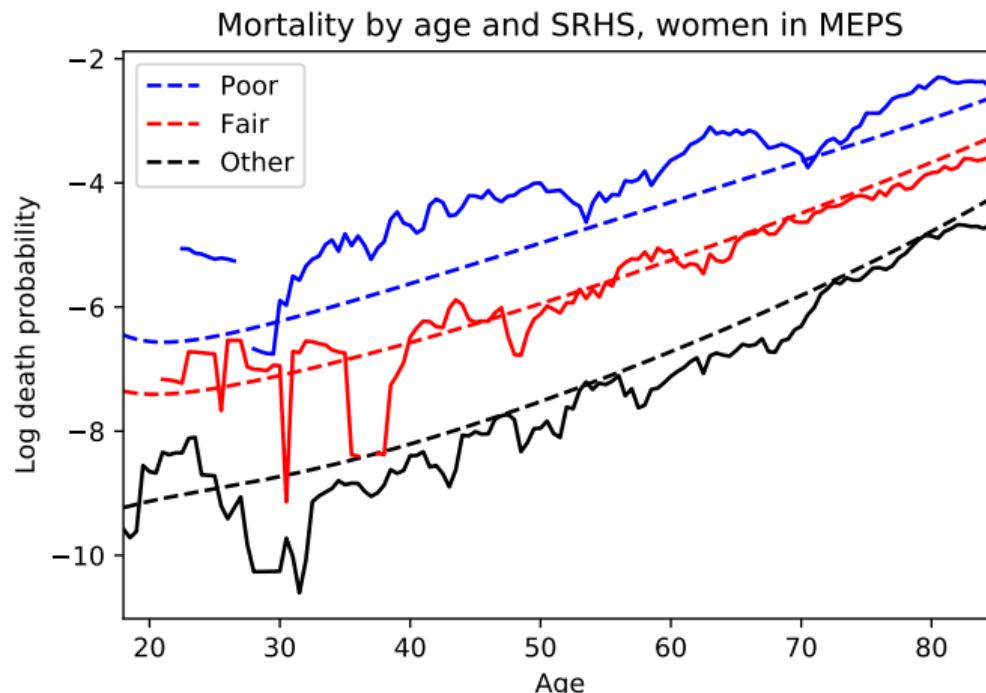
## Estimated parameters: MEPS women (mortality)

Param	Description	MEPS
$\theta_0$	Mortality probit: constant term	2.667 (9.96e-2)
$\theta_{x1}$	Mortality probit: linear coefficient on health	0.116 (4.65e-3)
$\theta_{x2}$	Mortality probit: quadratic coefficient on health	-5.51e-4 (2.51e-4)
$\theta_{j1}$	Mortality probit: linear coefficient on age	1.29e-2 (3.38e-3)
$\theta_{j2}$	Mortality probit: quadratic coefficient on age	-3.43e-4 (3.19e-5)
$\theta_{j3}$	Mortality probit: cubic coefficient on age	2.80e-7 (1.10e-7)
$\theta_{xj}$	Mortality probit: coefficient on age×health	1.24e-4 (3.77e-5)

# Model fit: Mortality by age for MEPS women



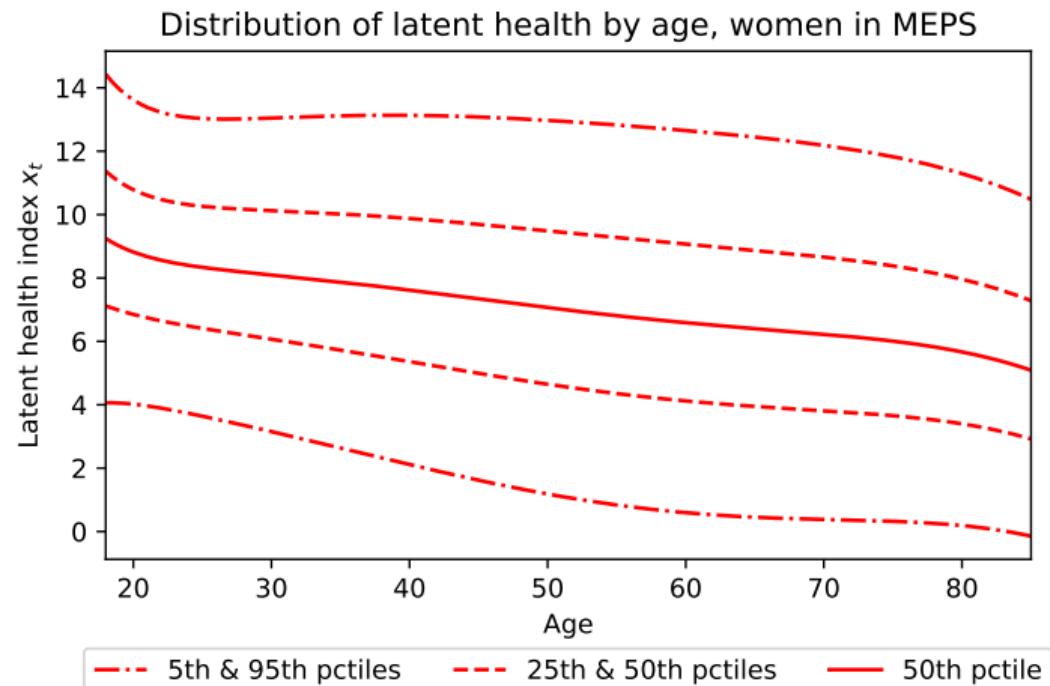
## Model fit: Mortality by age and SRHS for MEPS women



## Estimated parameters: MEPS women (latent health)

Param	Description	MEPS
$\beta_0$	Expected health: constant term	-1.067 (0.295)
$\beta_{j1}$	Expected health: linear coefficient on age	0.974 (3.19e-2)
$\beta_{j2}$	Expected health: quadratic coefficient on age	-3.57e-2 (1.27e-3)
$\beta_{j3}$	Expected health: cubic coefficient on age	5.18e-4 (2.13e-5)
$\beta_{j4}$	Expected health: quartic coefficient on age	-2.68e-6 (1.25e-7)
$\mu_0$	Mean of latent health at model start	9.383 (8.80e-2)
$\sigma_0$	Standard deviation of latent health at model start	3.160 (4.47e-2)

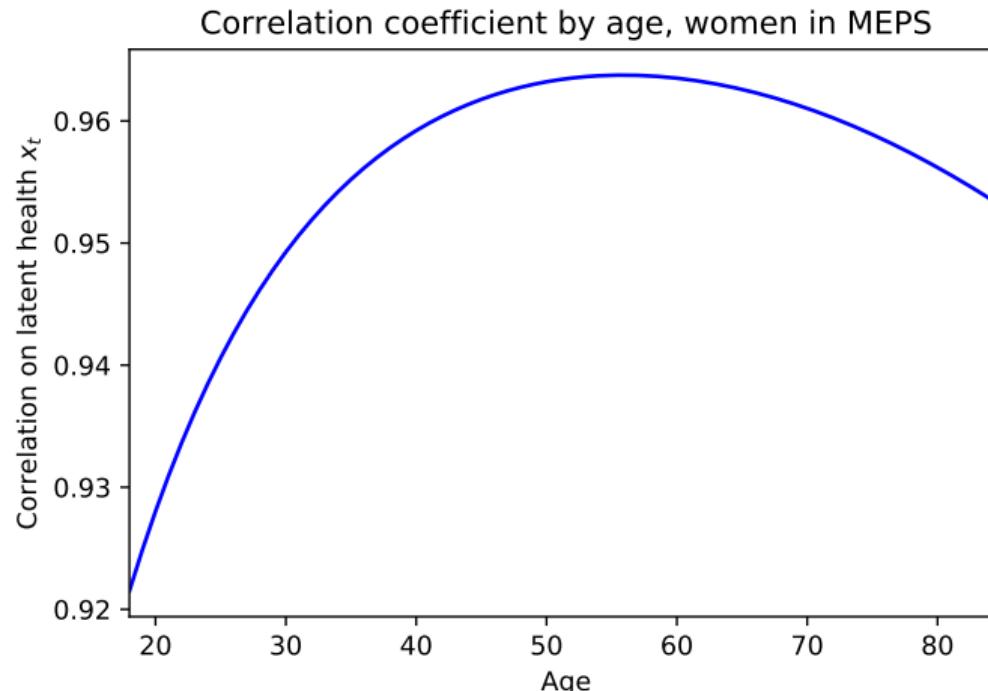
# Estimated model: Distribution of latent health (MEPS women)



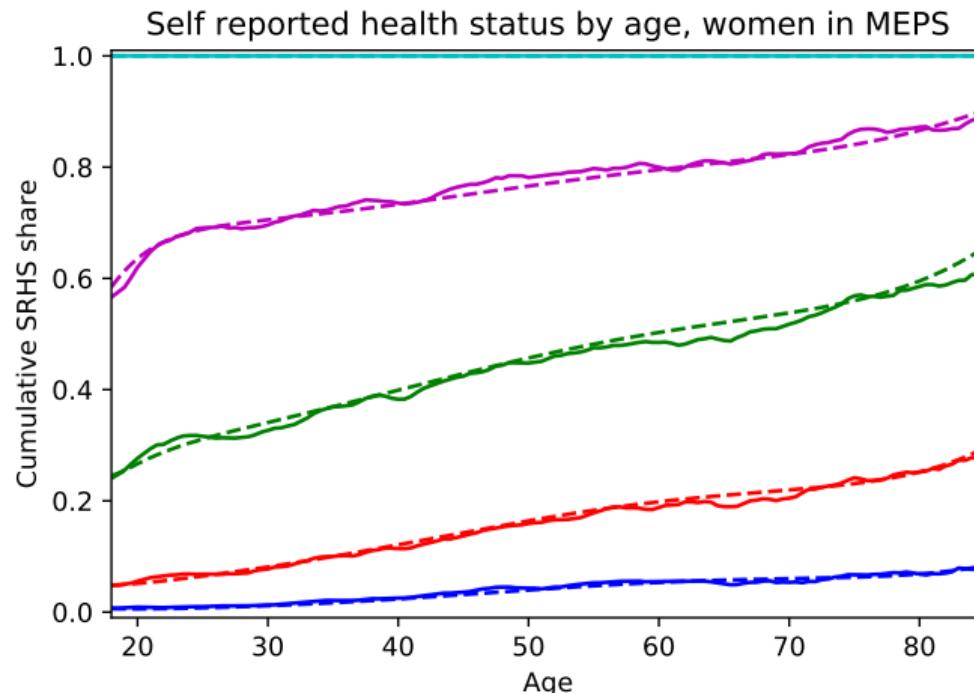
## Estimated parameters: MEPS women (correlation and SRHS)

Param	Description	MEPS
$\gamma_0$	Correlation factor: constant term	1.262 (9.87e-2)
$\gamma_{j1}$	Correlation factor: linear coefficient on age	8.54e-2 (6.71e-3)
$\gamma_{j2}$	Correlation factor: quadratic coefficient on age	-1.11e-3 (1.43e-4)
$\gamma_{j3}$	Correlation factor: cubic coefficient on age	4.15e-6 (9.65e-7)
$\alpha_1$	SRHS: Linear coefficient on latent health	0.503 (5.33e-3)
$\chi_2$	SRHS: Cut b/w reporting "fair" and "good"	1.596 (6.90e-3)
$\chi_3$	SRHS: Cut b/w reporting "good" and "very good"	3.389 (9.82e-3)
$\chi_4$	SRHS: Cut b/w reporting "very good" and "excellent"	5.112 (1.28e-2)

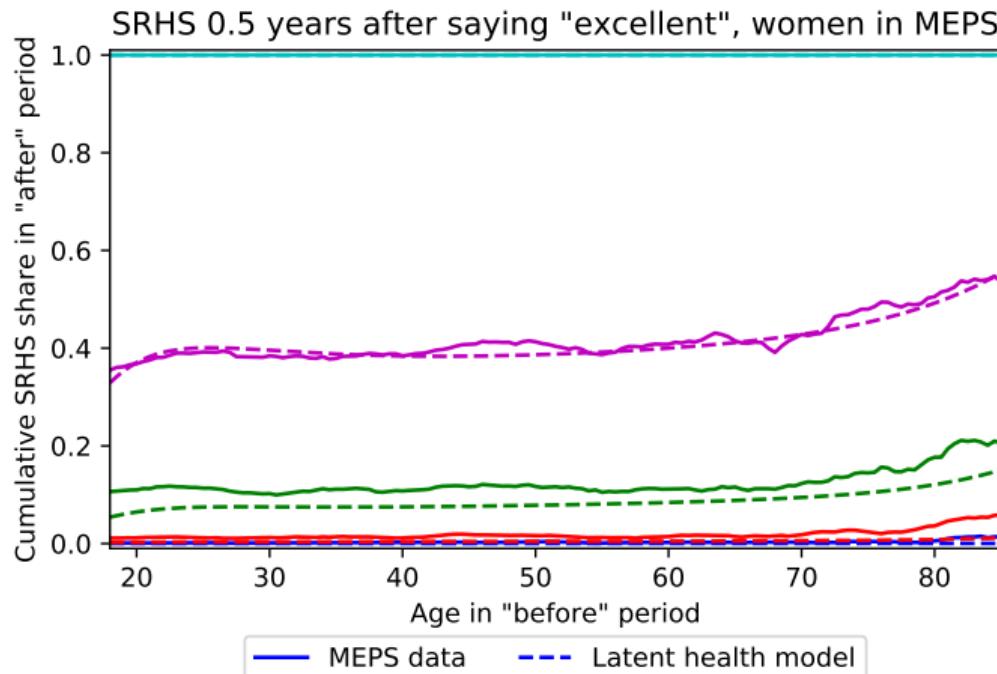
# Estimated model: Latent health serial correlation (MEPS women)



# Model fit: SRHS by age for MEPS women



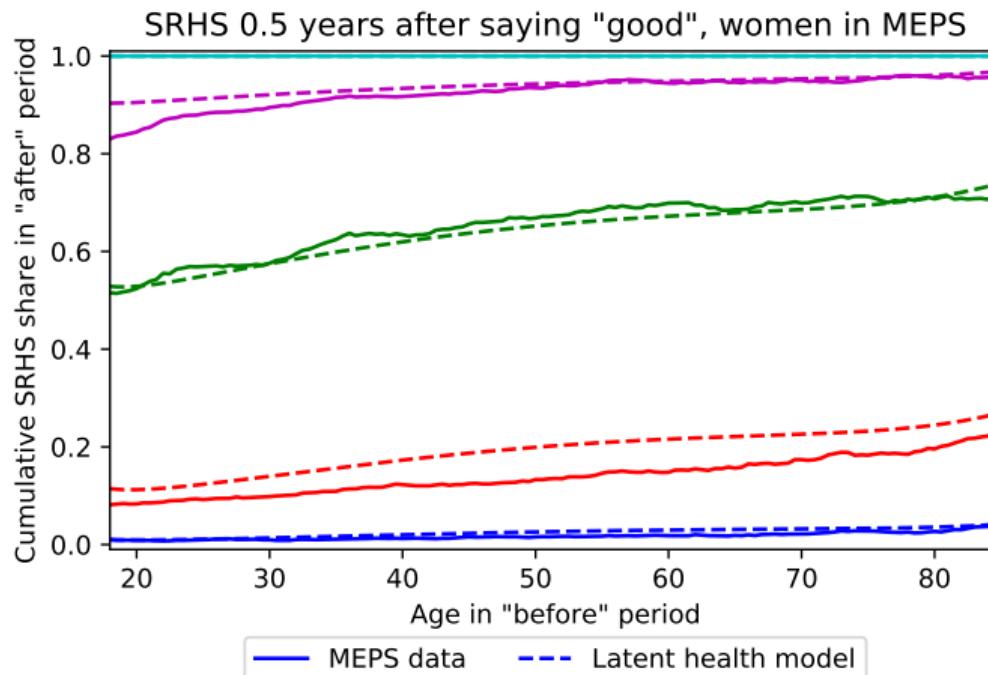
## Model fit: SRHS transitions for MEPS women: 1 wave



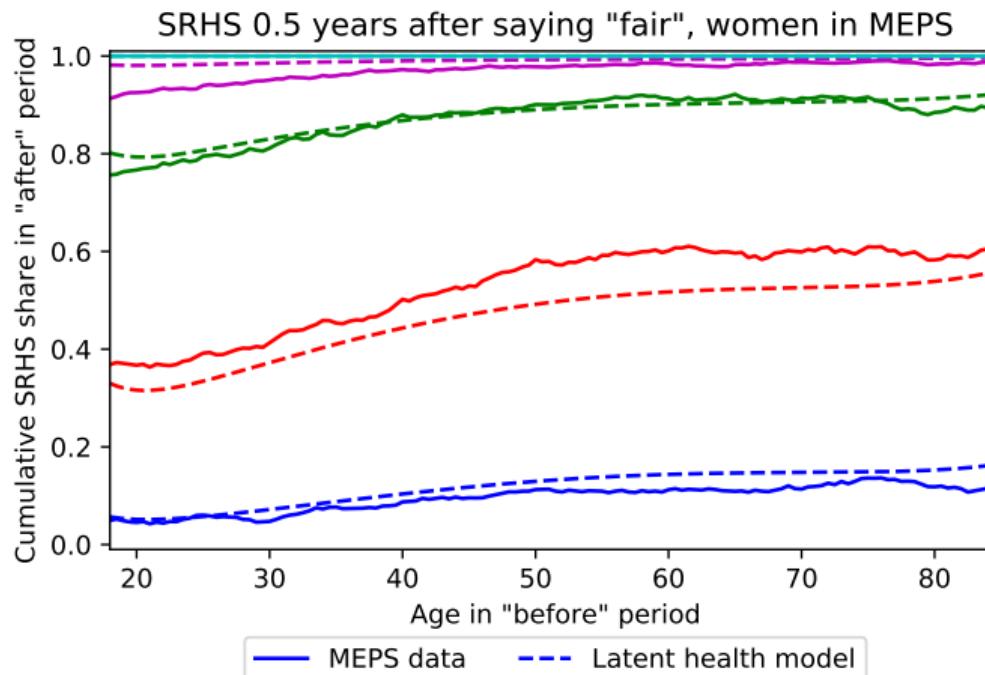
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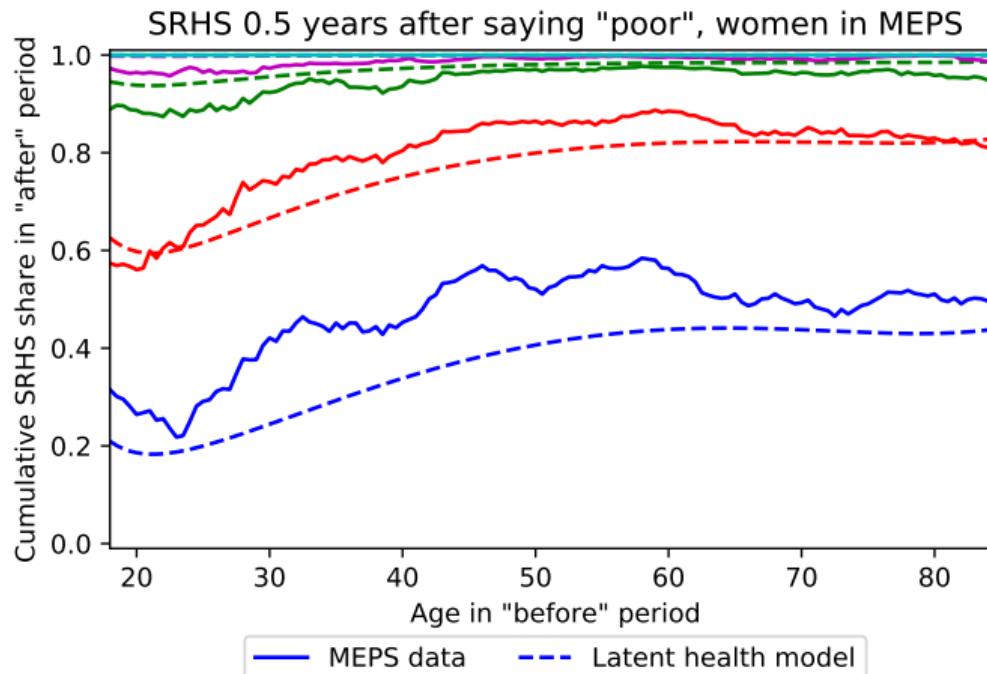
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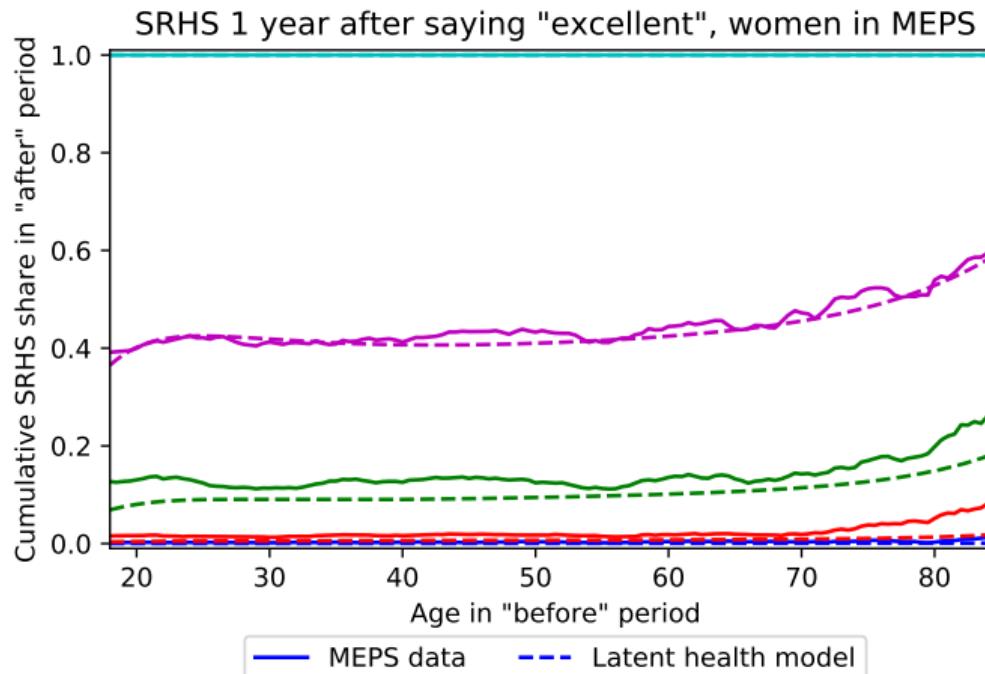
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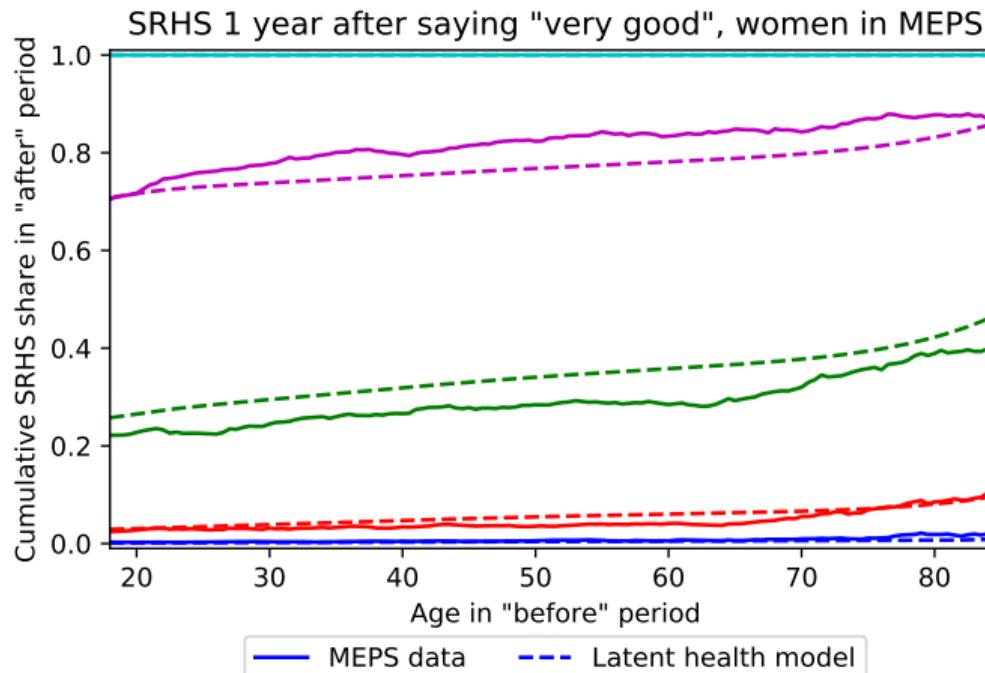
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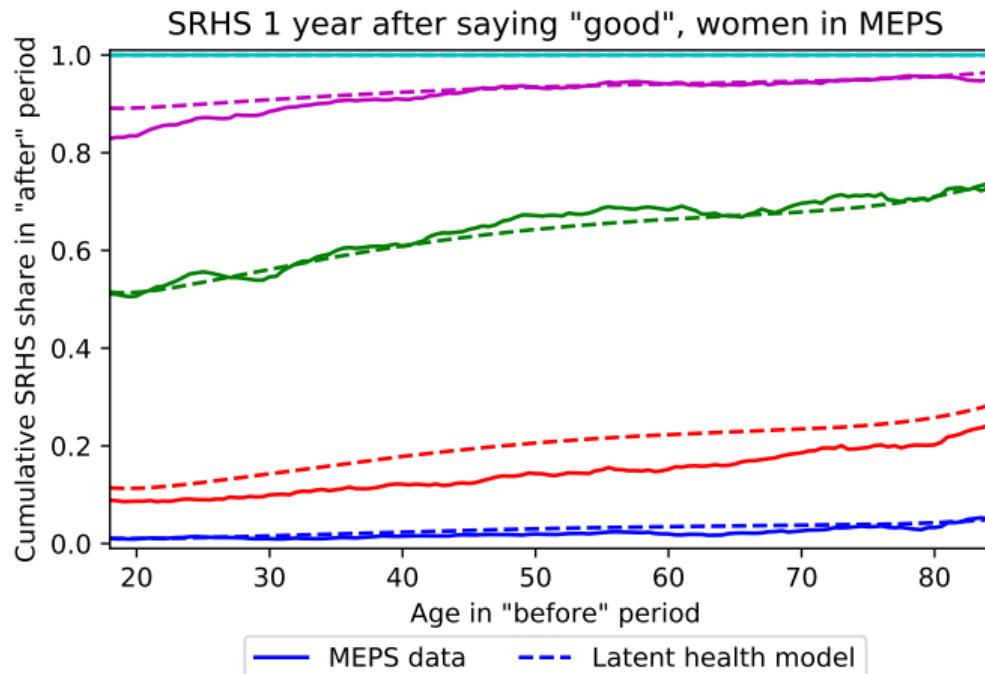
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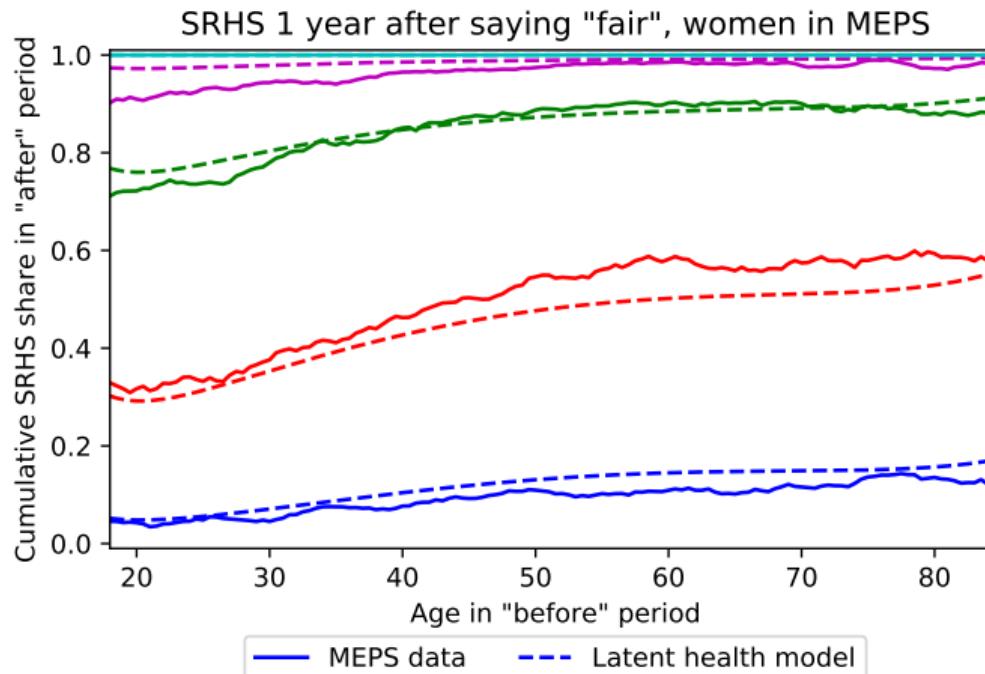
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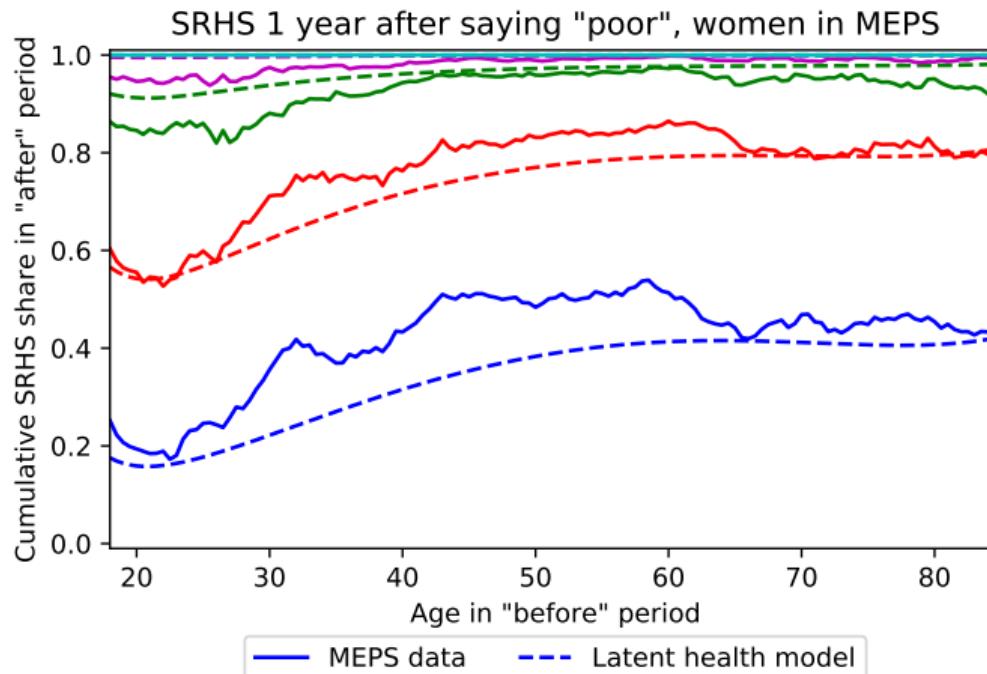
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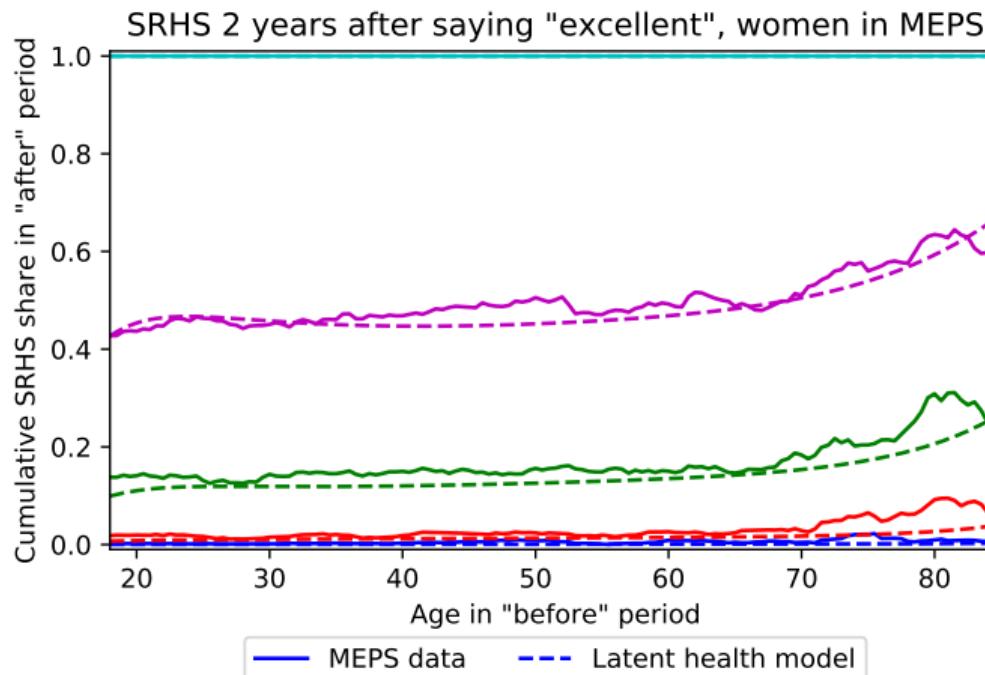
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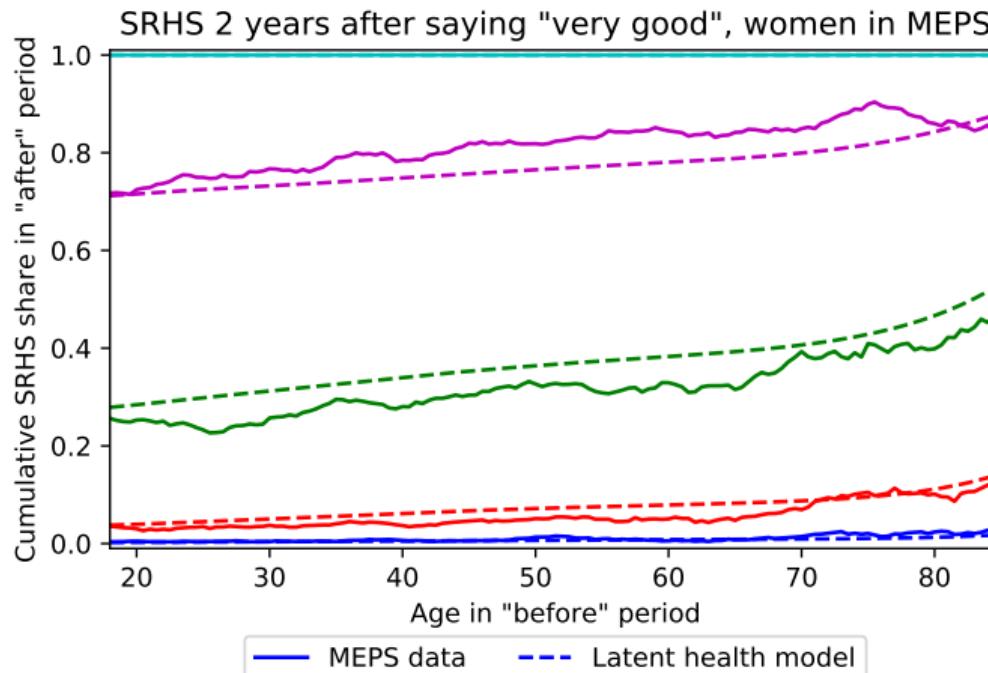
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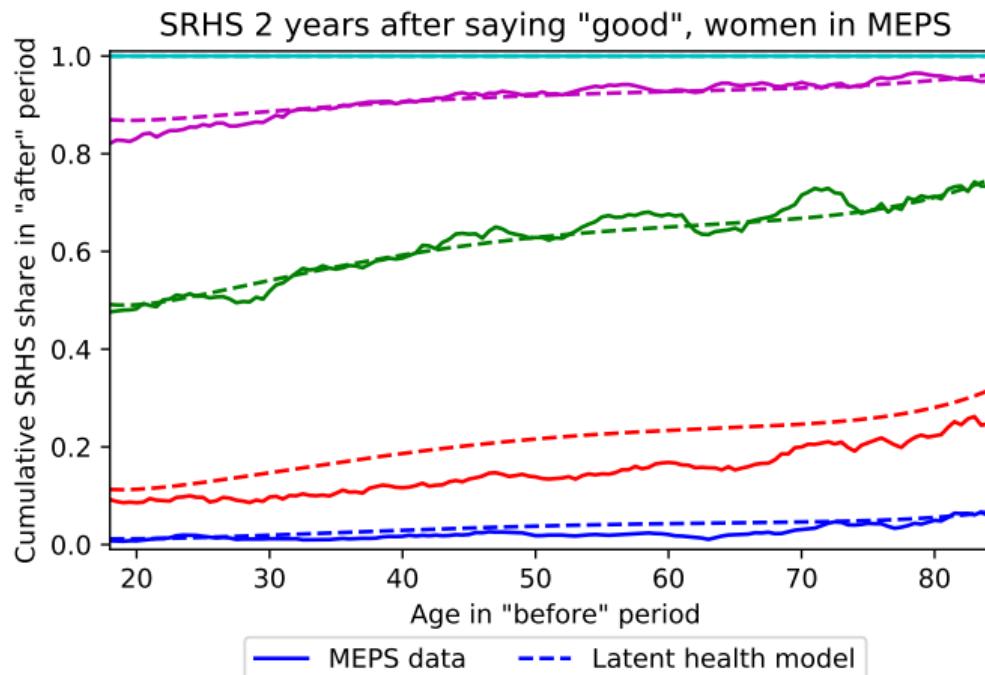
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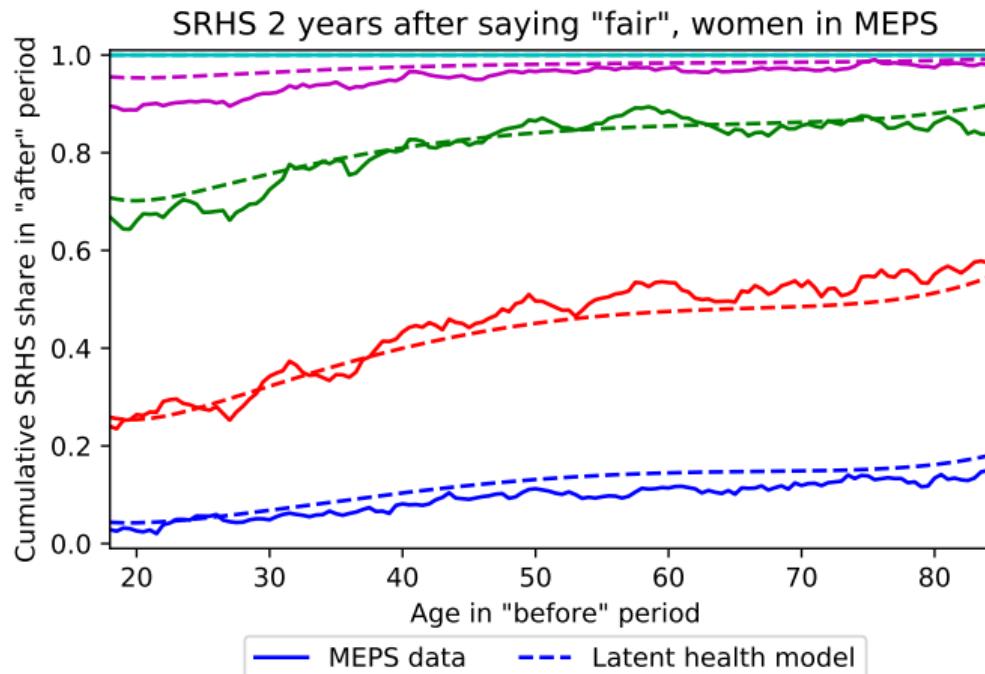
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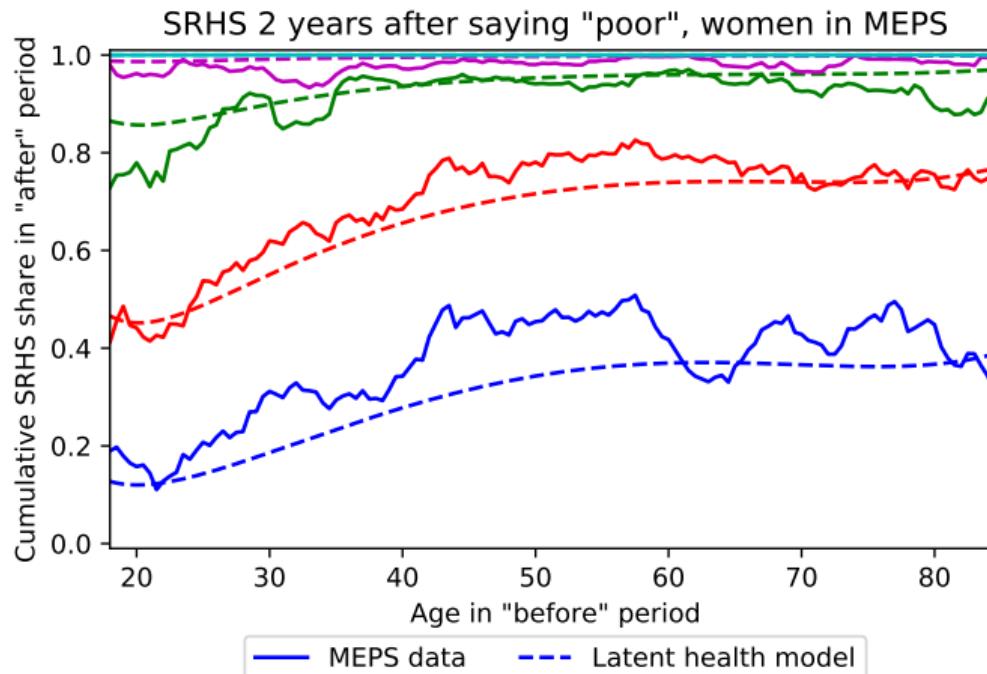
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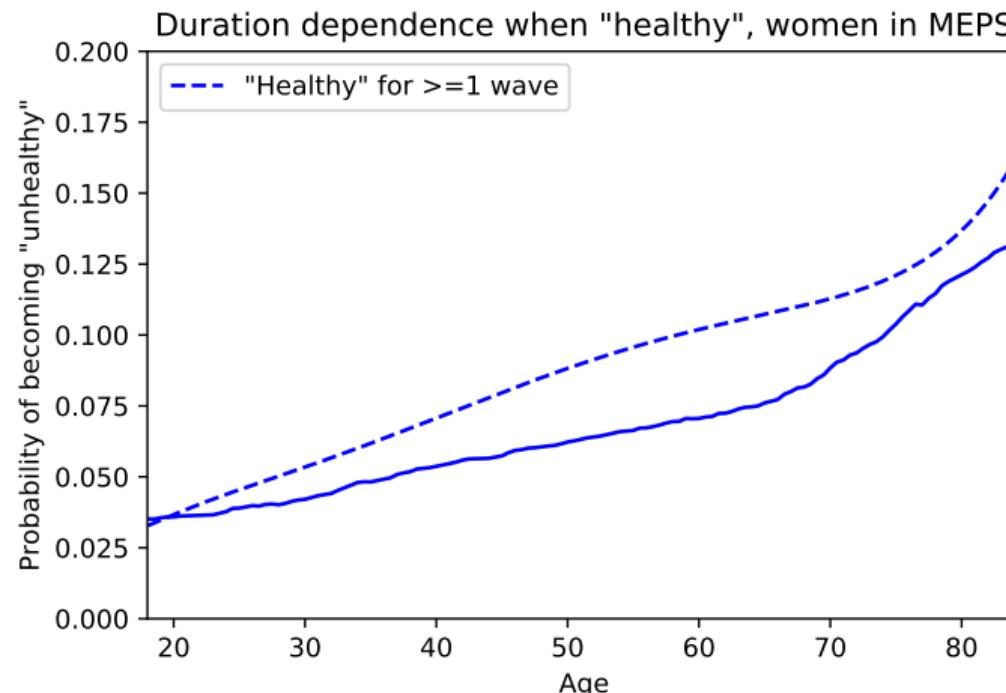
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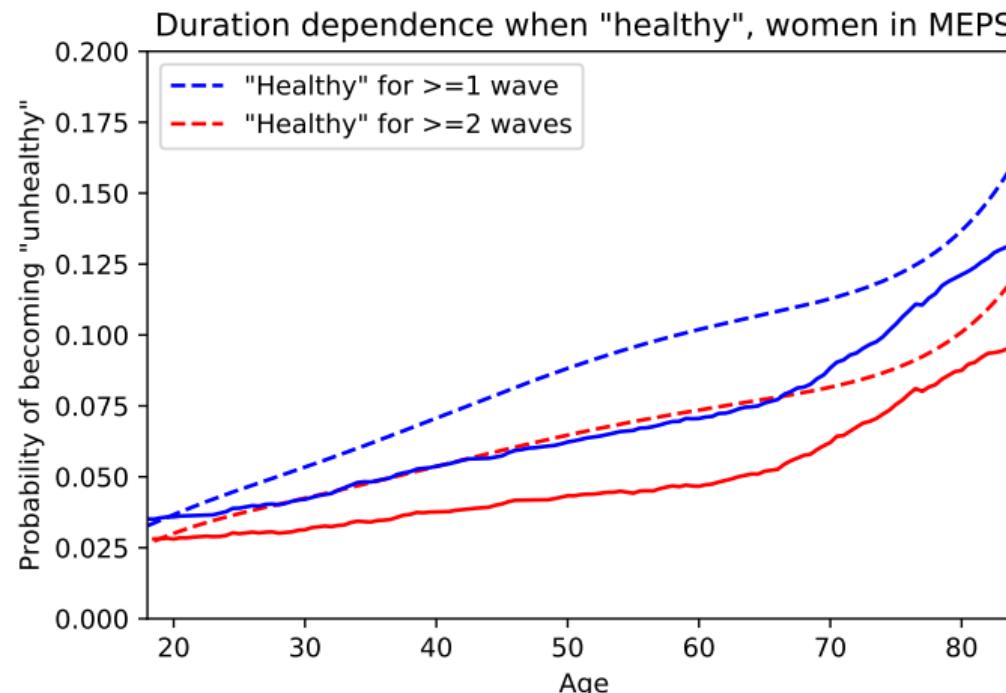
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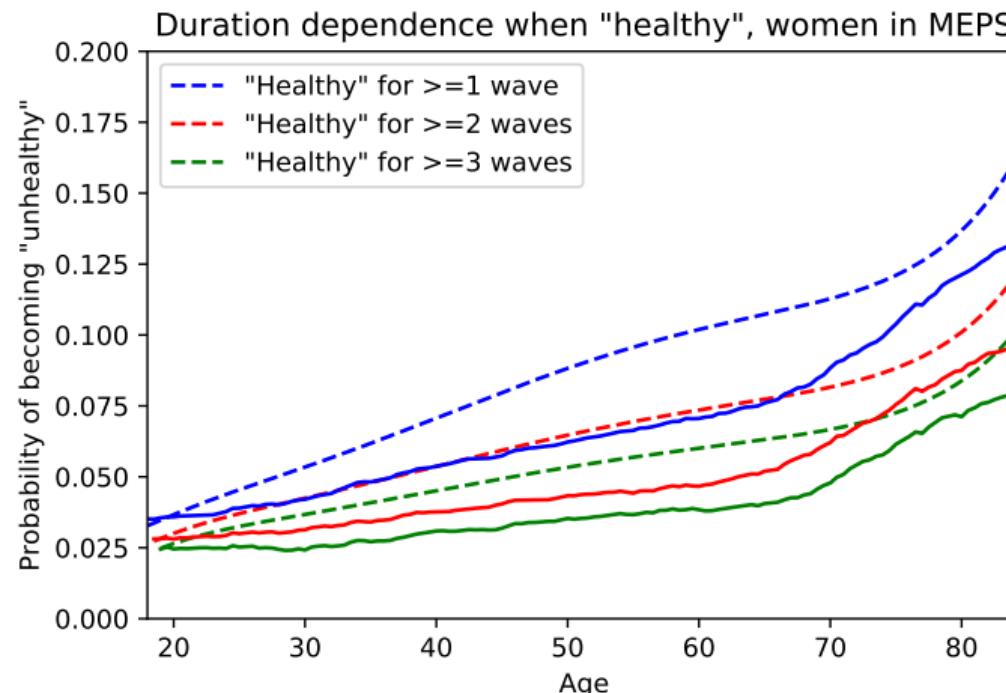
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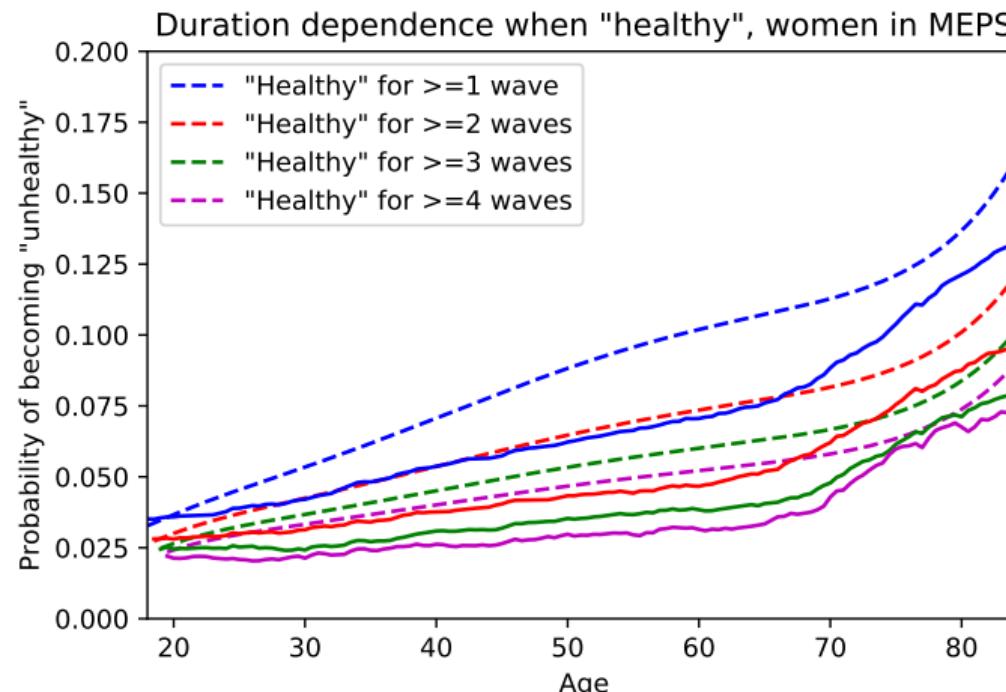
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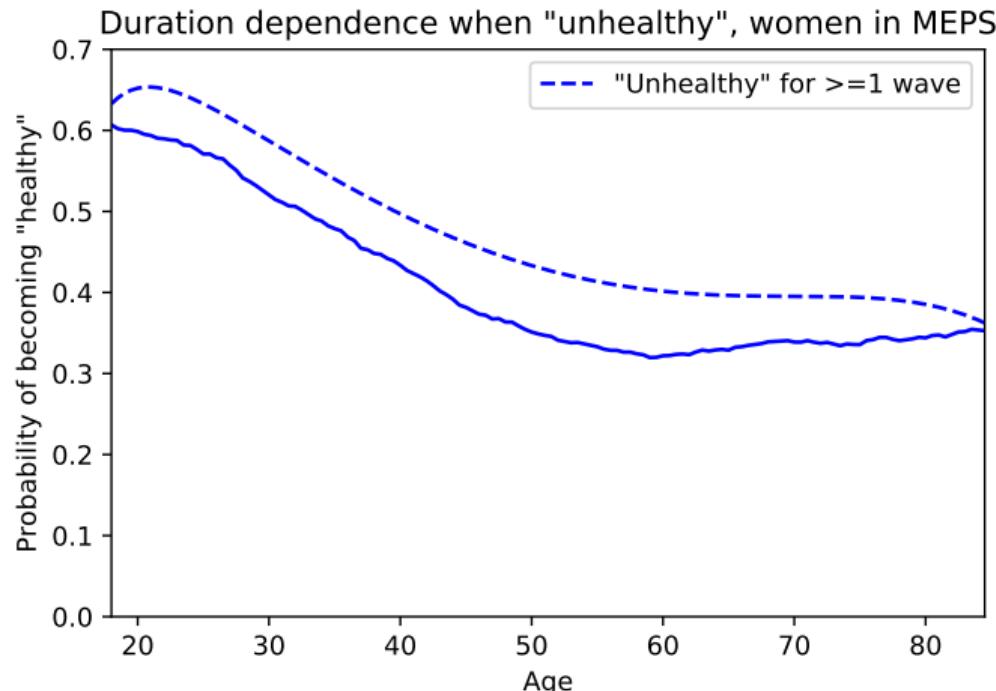
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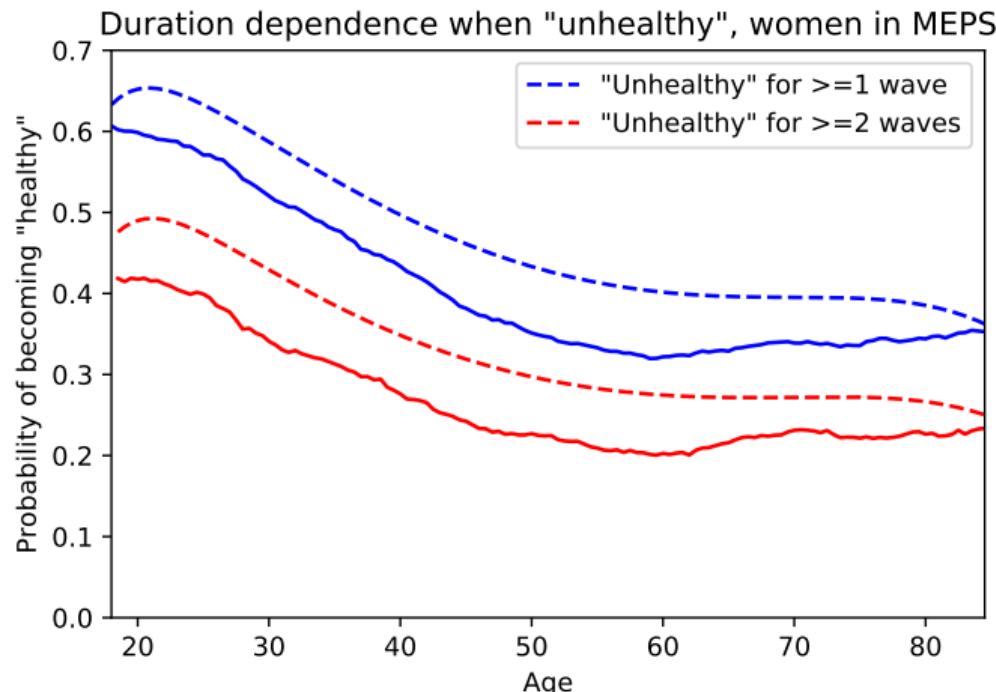
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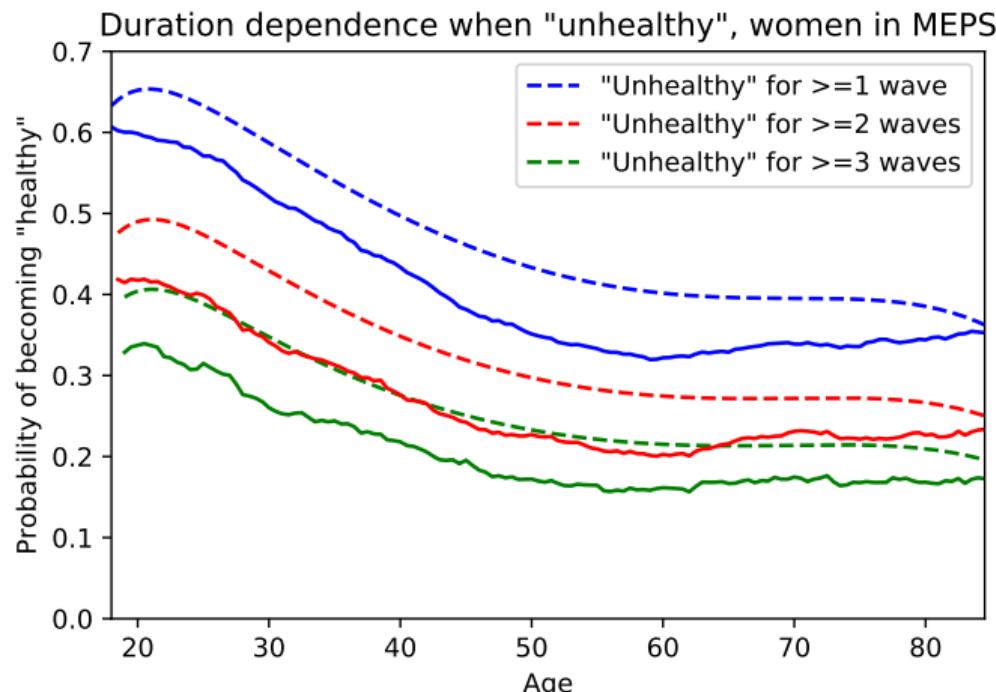
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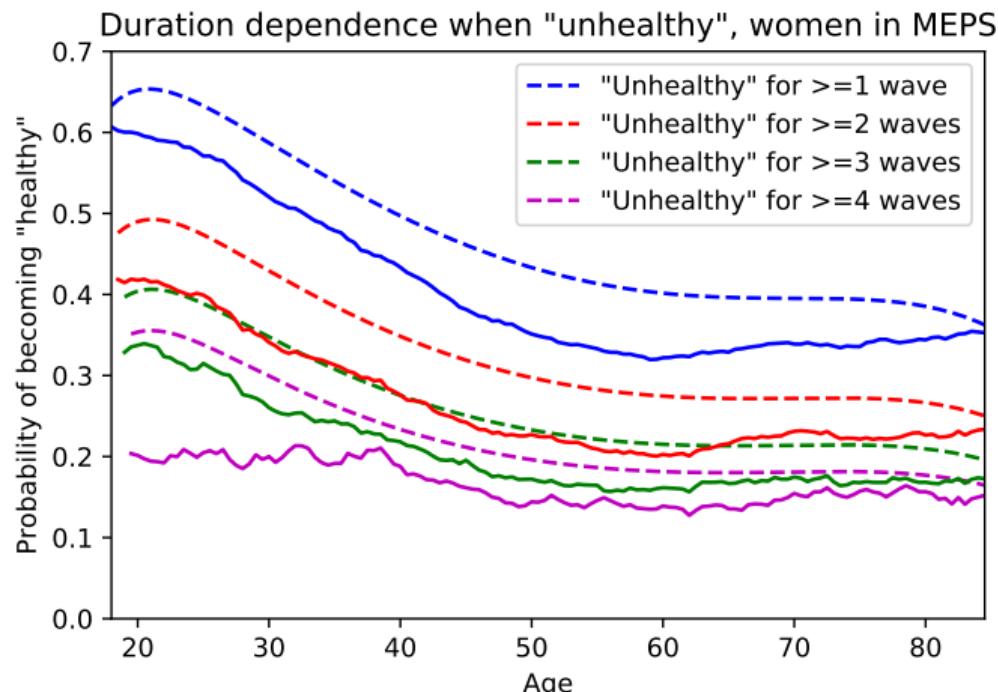
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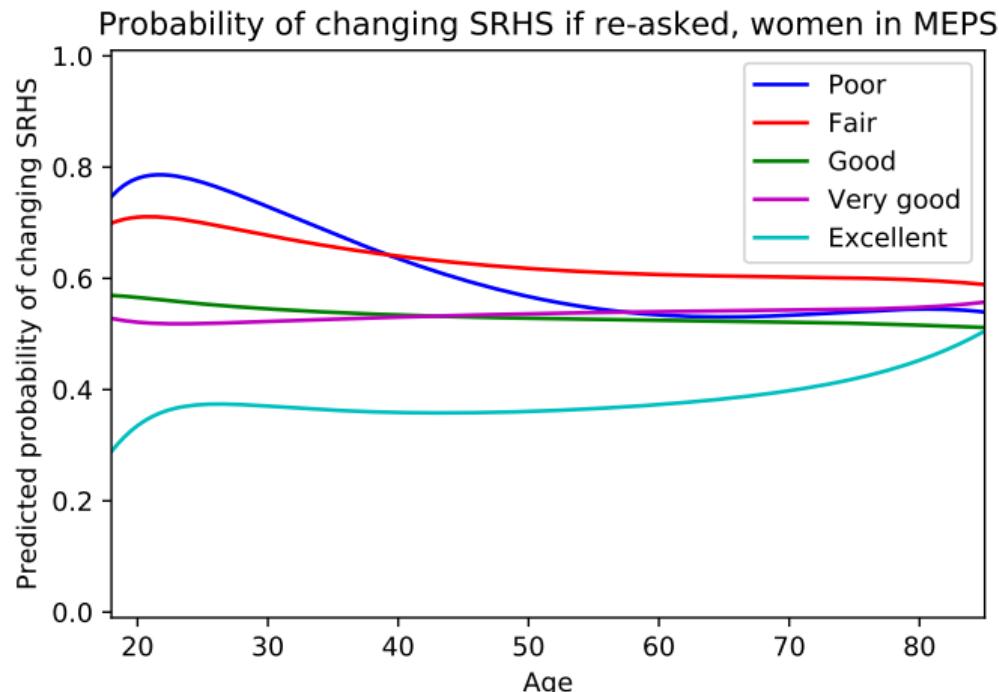
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# Estimated model: Changing SRHS if asked twice (MEPS women)



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- Same exercise for men:  $R^2 = 0.2326$  (latent health) vs  $R^2 = 0.2326$  (full interact)

## Estimated model: Out-of-sample fit

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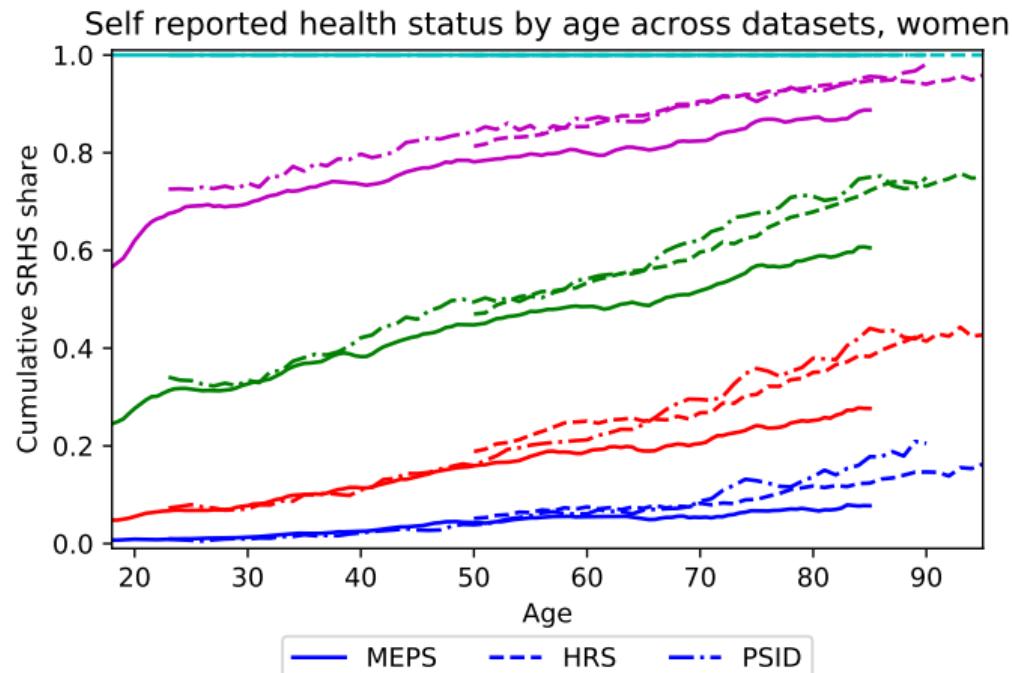
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- Result: **No, it absolutely does not fit**
- Probably too much of a stretch to project a model 10x beyond data
- MEPS has a selection bias problem: too few respondents in poor health
- And asks respondents to compare to people of their own age

# SRHS across datasets: MEPS, HRS, and PSID (women)



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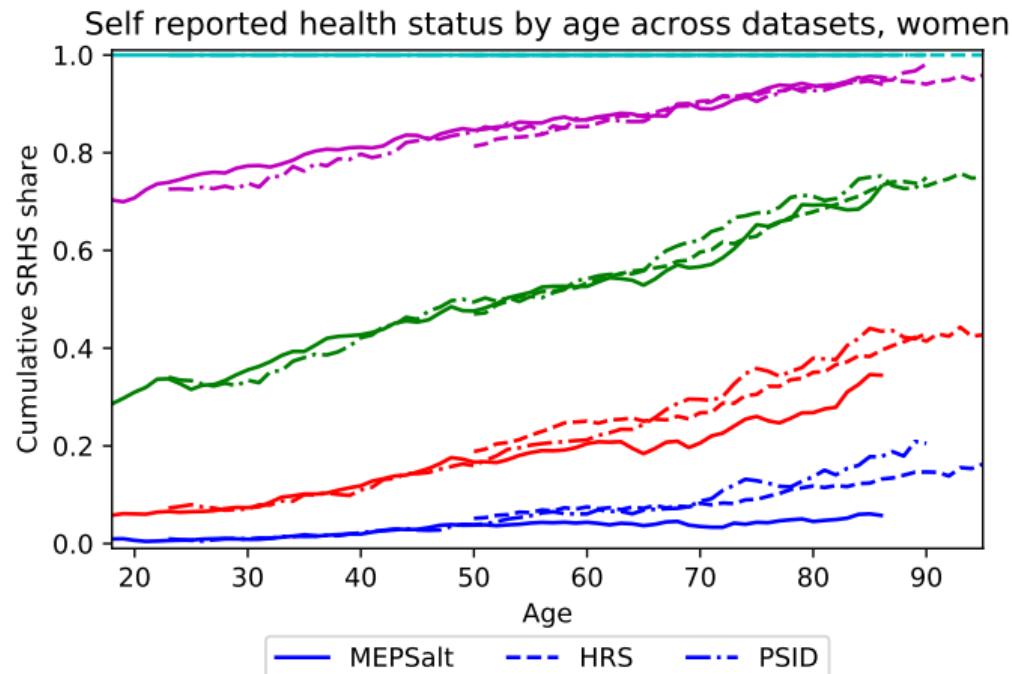
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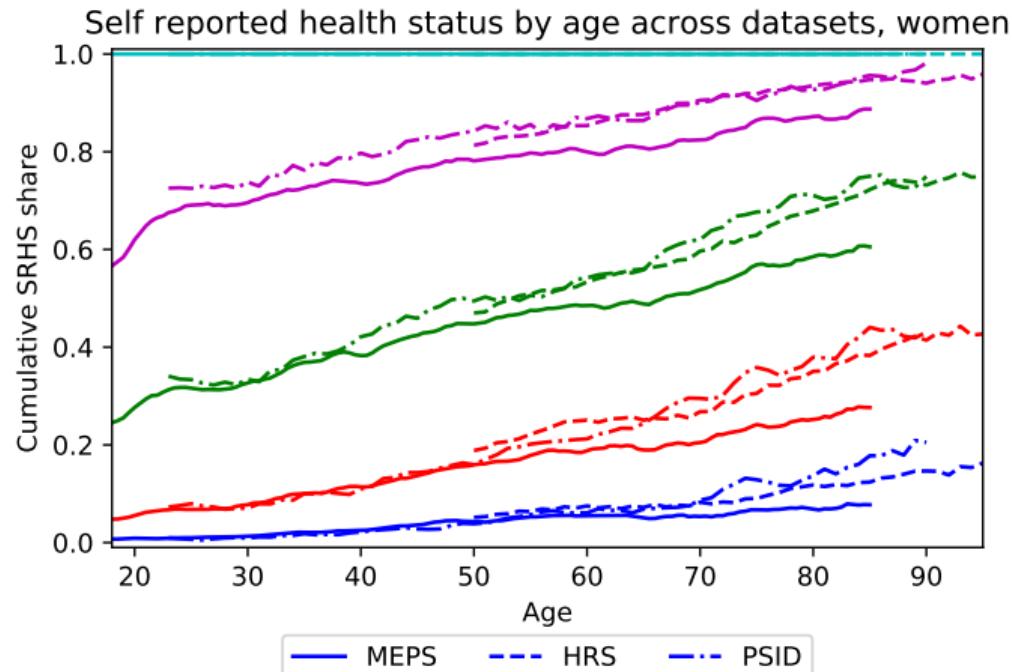
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- 120,055 women observed for 342,796 waves (6,918 deaths)
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- 82% of respondents observed 2 or fewer waves (mostly MEPS)

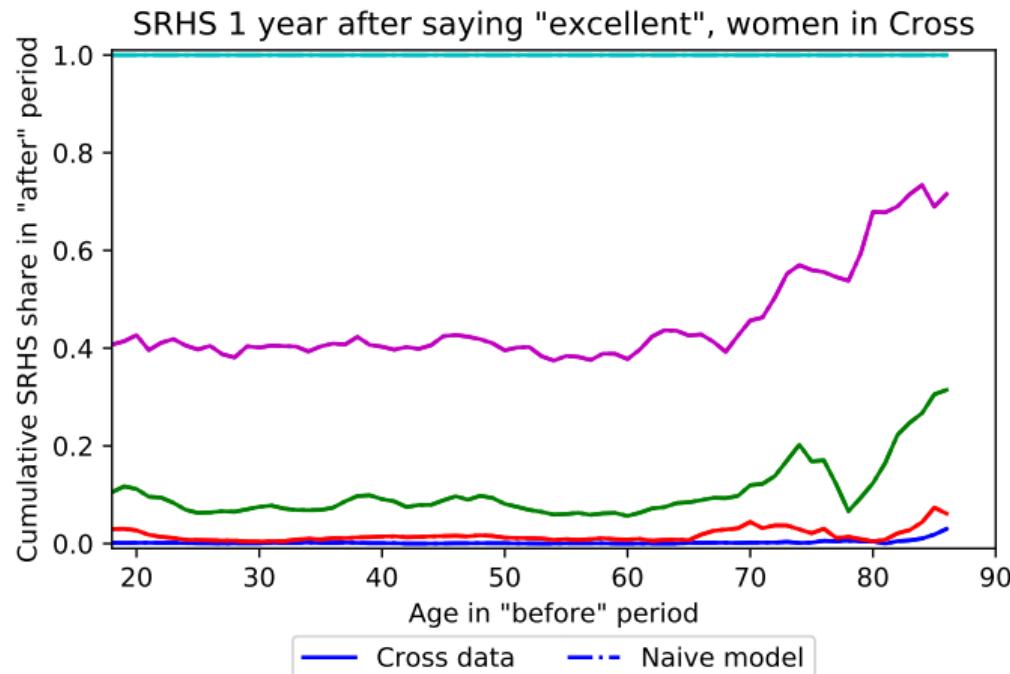
# SRHS across datasets: MEPS-SAQ, HRS, and PSID (women)



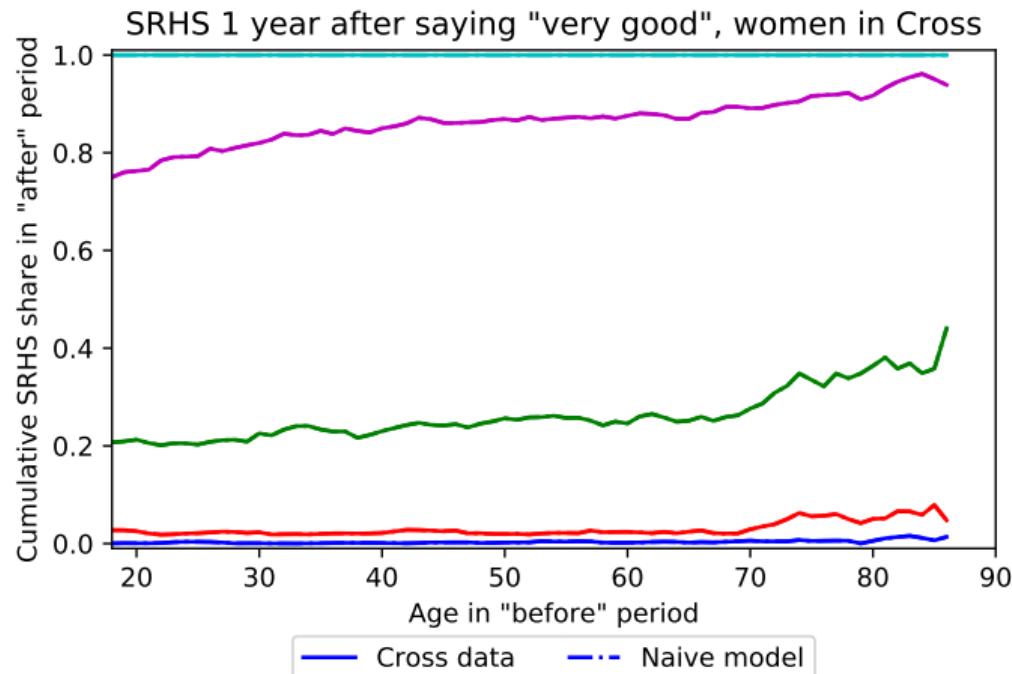
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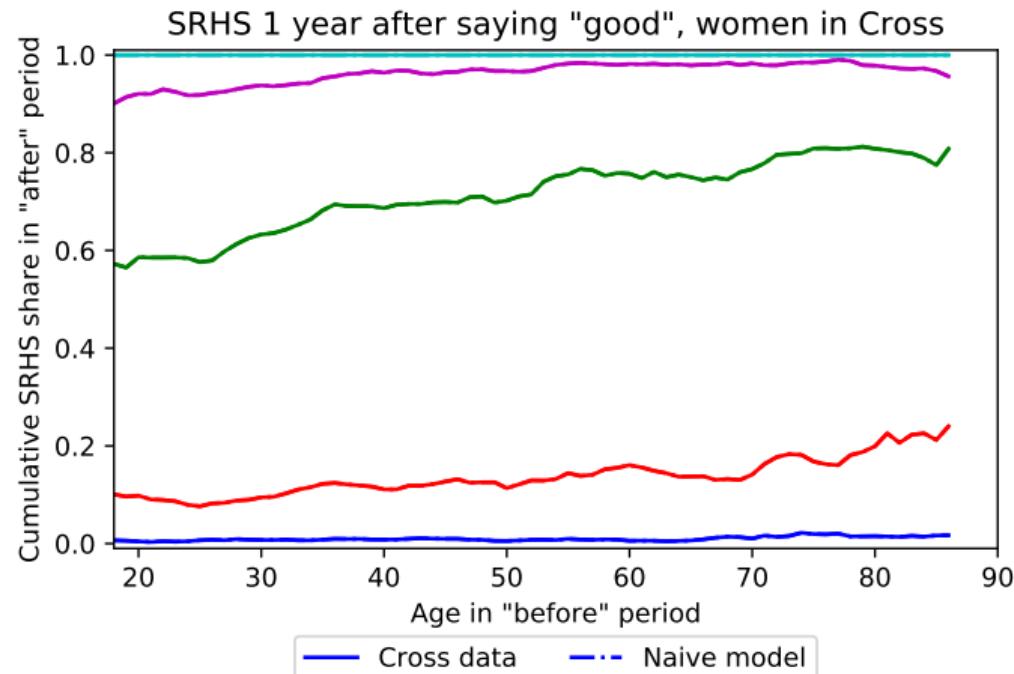
# Naive SRHS transitions for cross-data women: 1 year



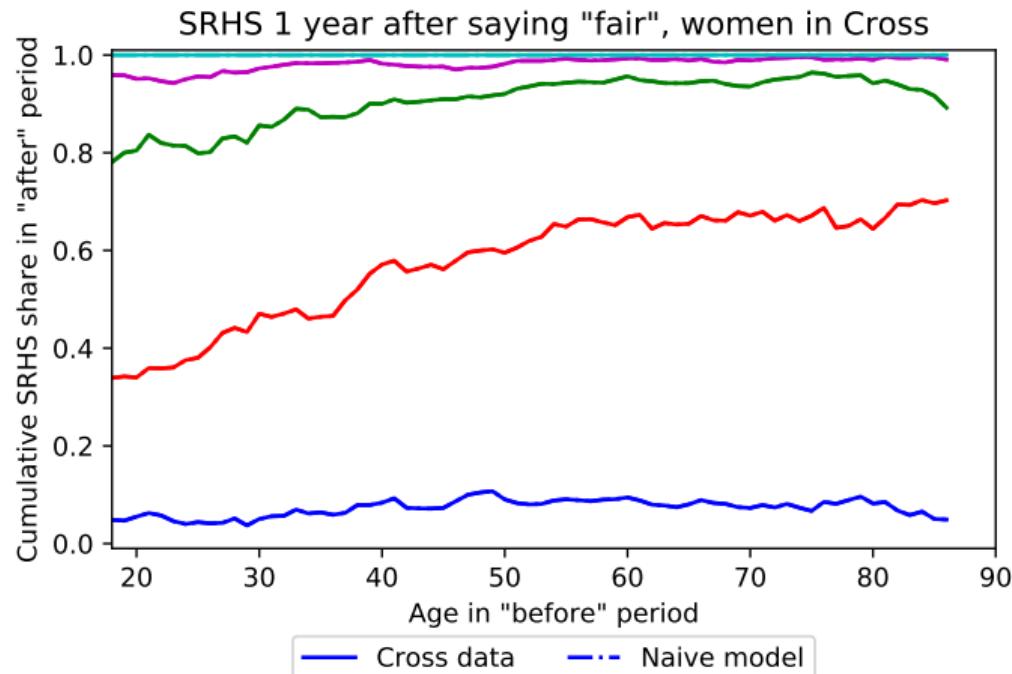
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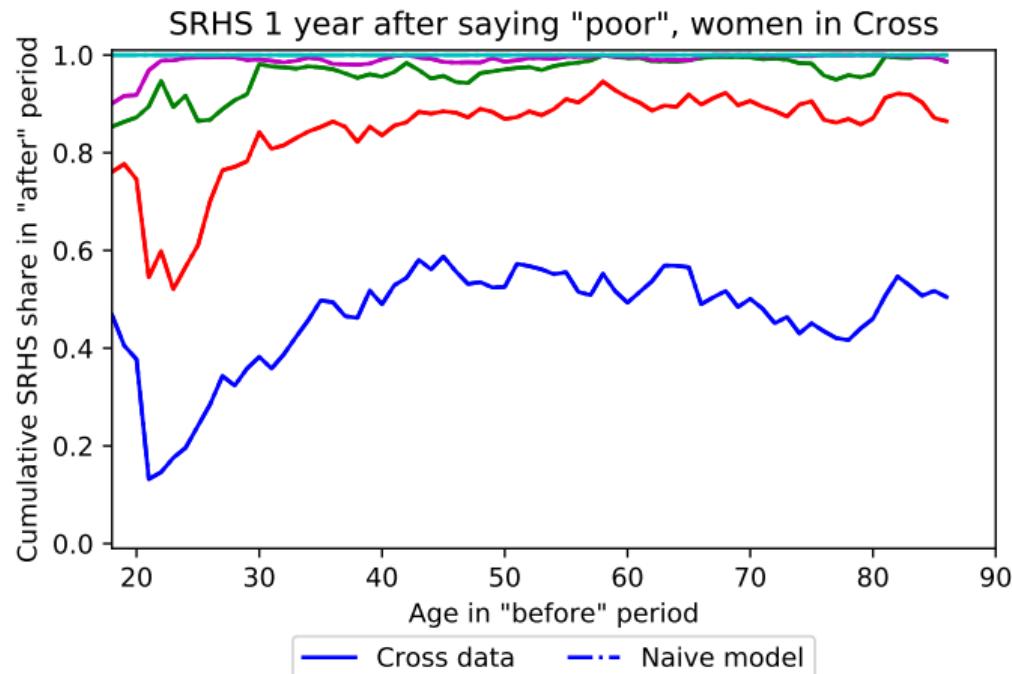
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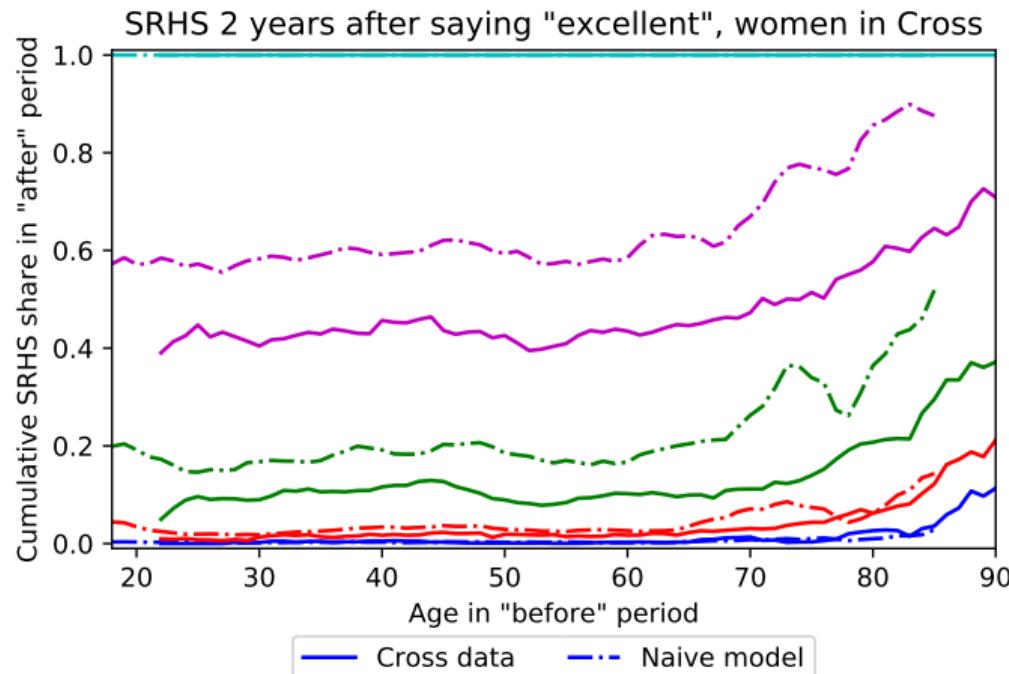
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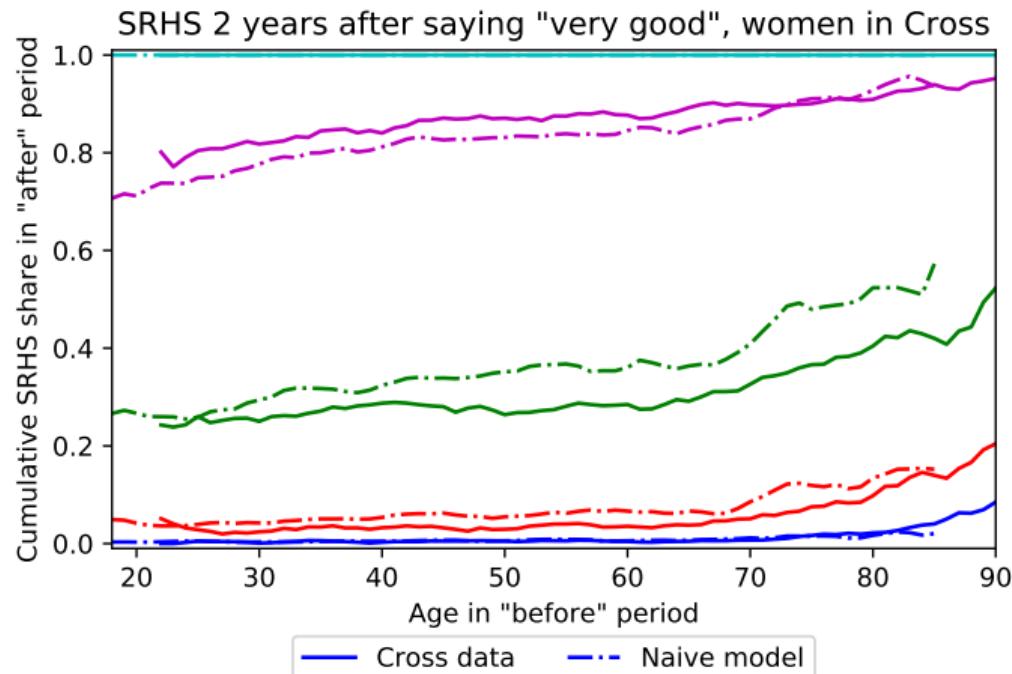
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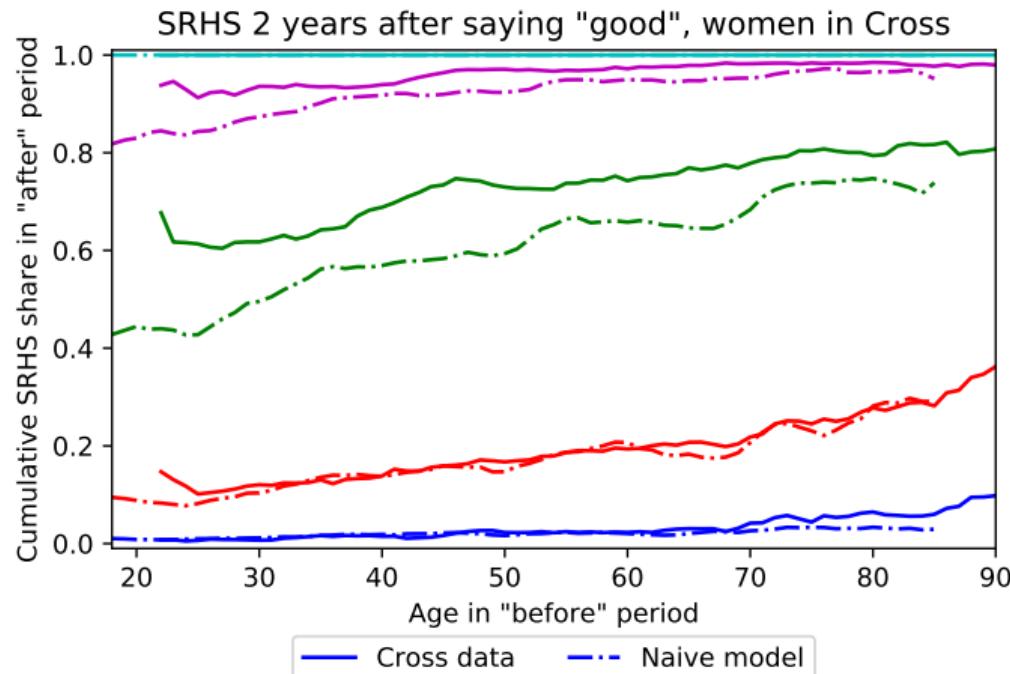
## Naive SRHS transitions for cross-data women: 2 years



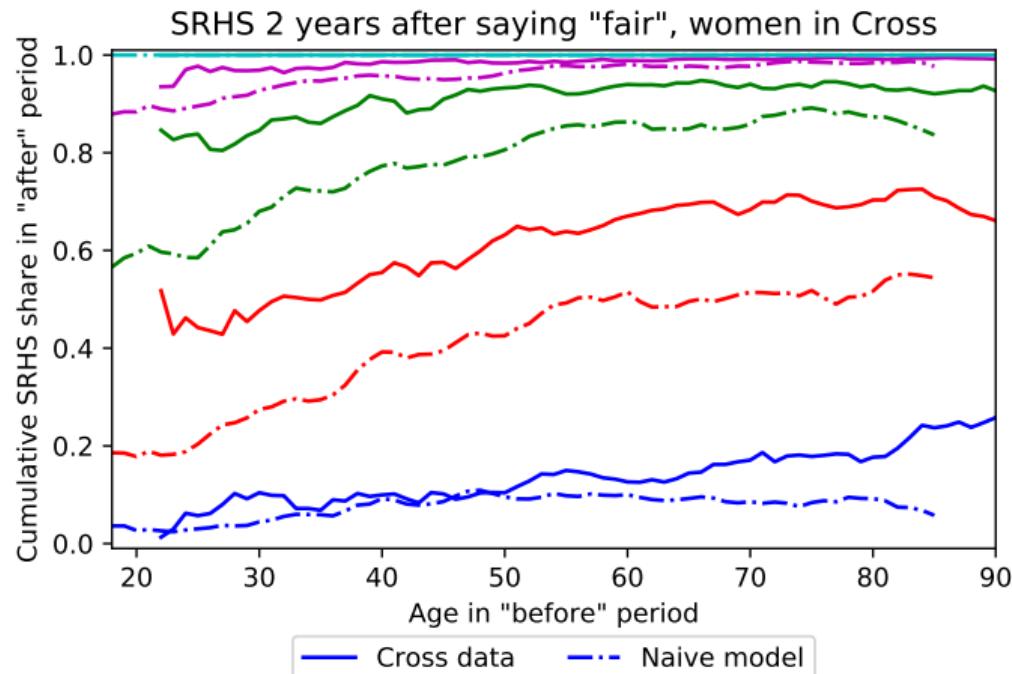
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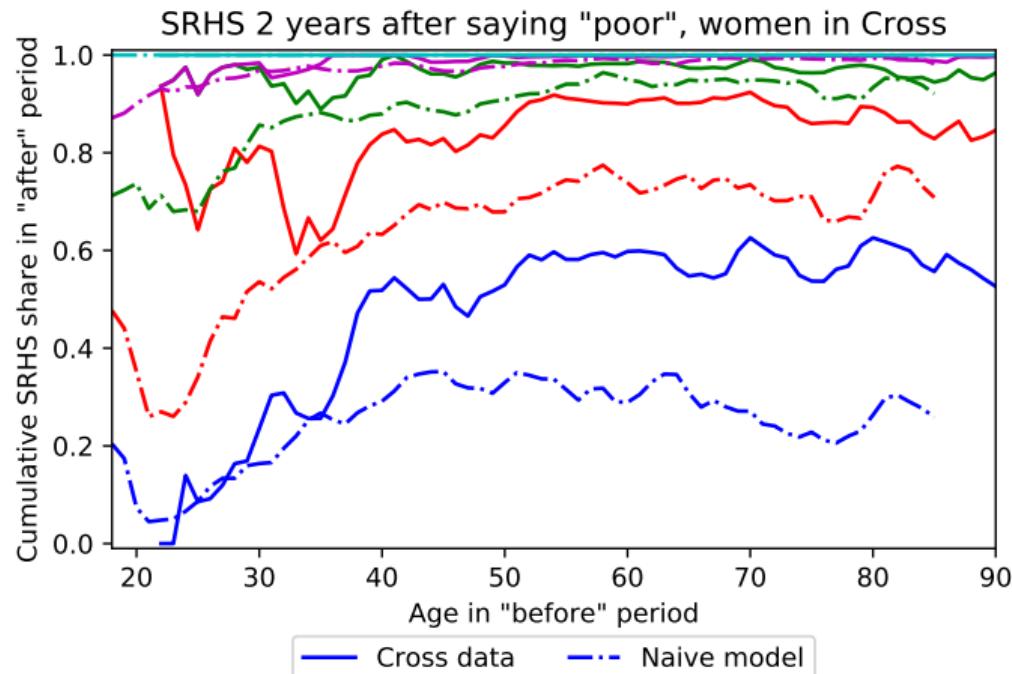
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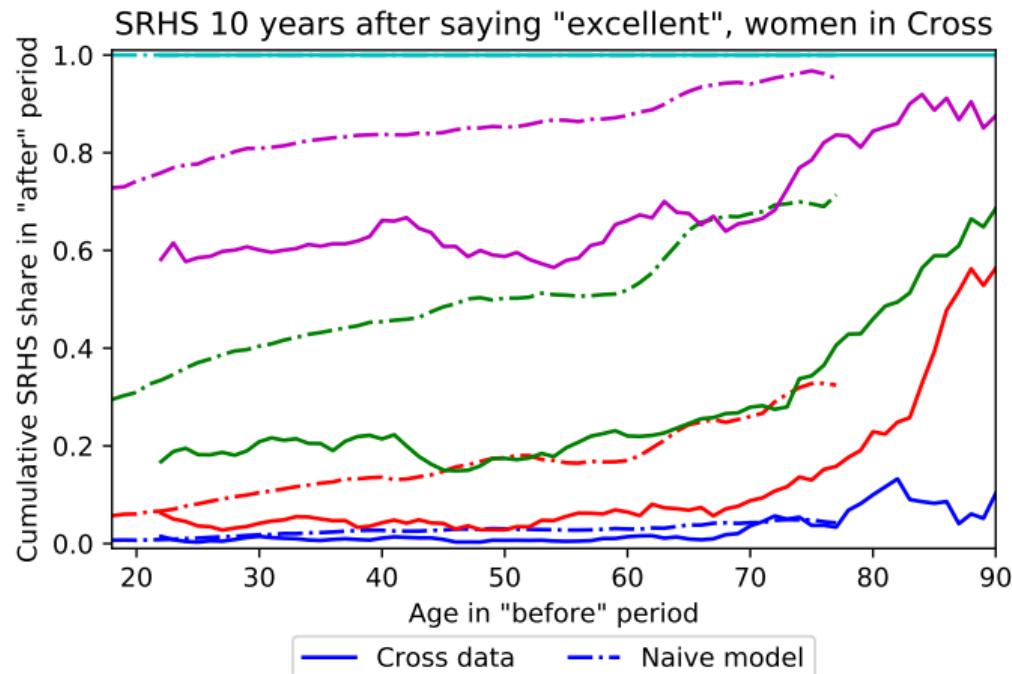
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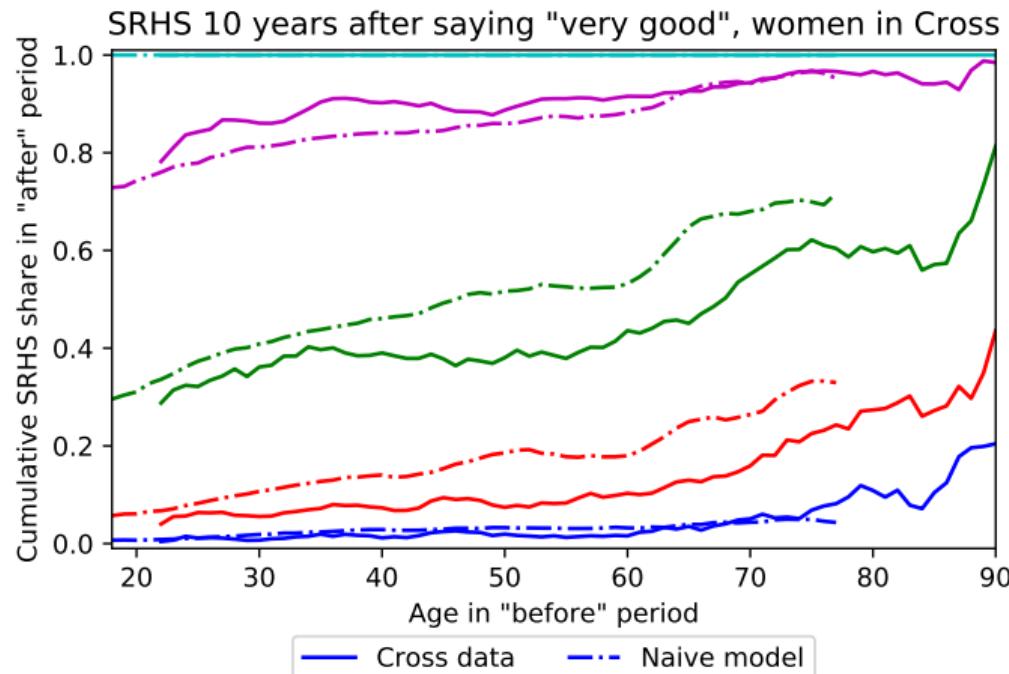
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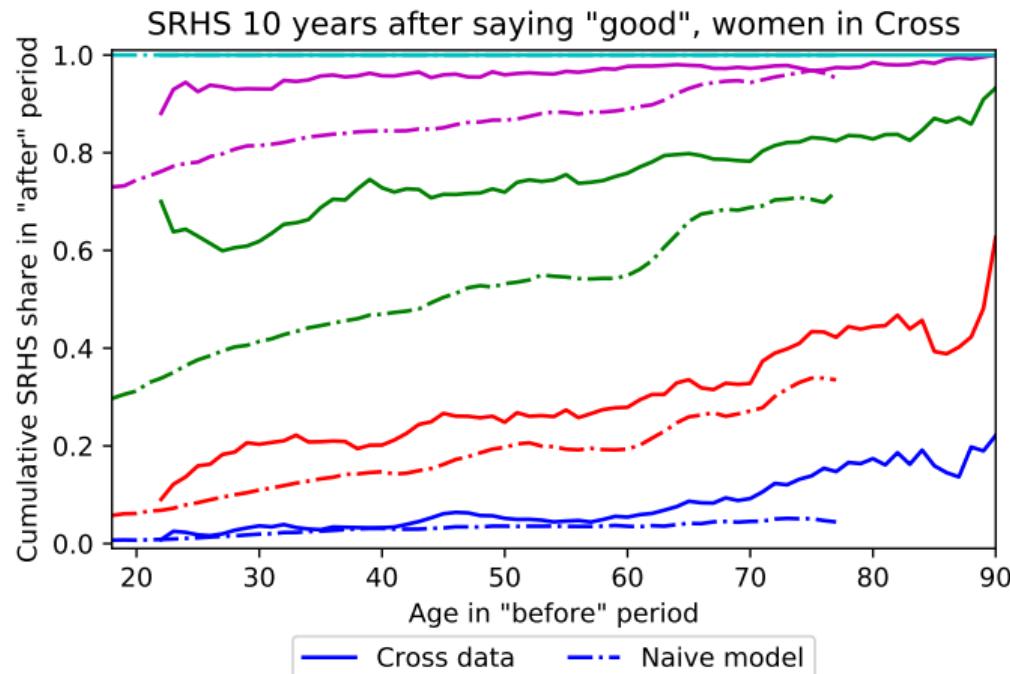
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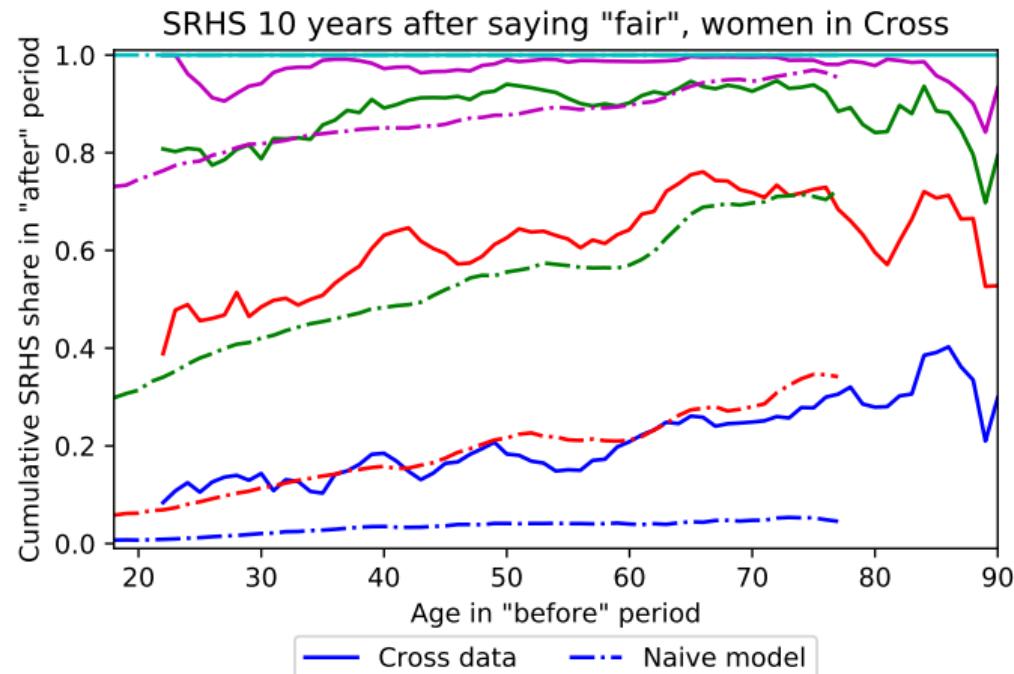
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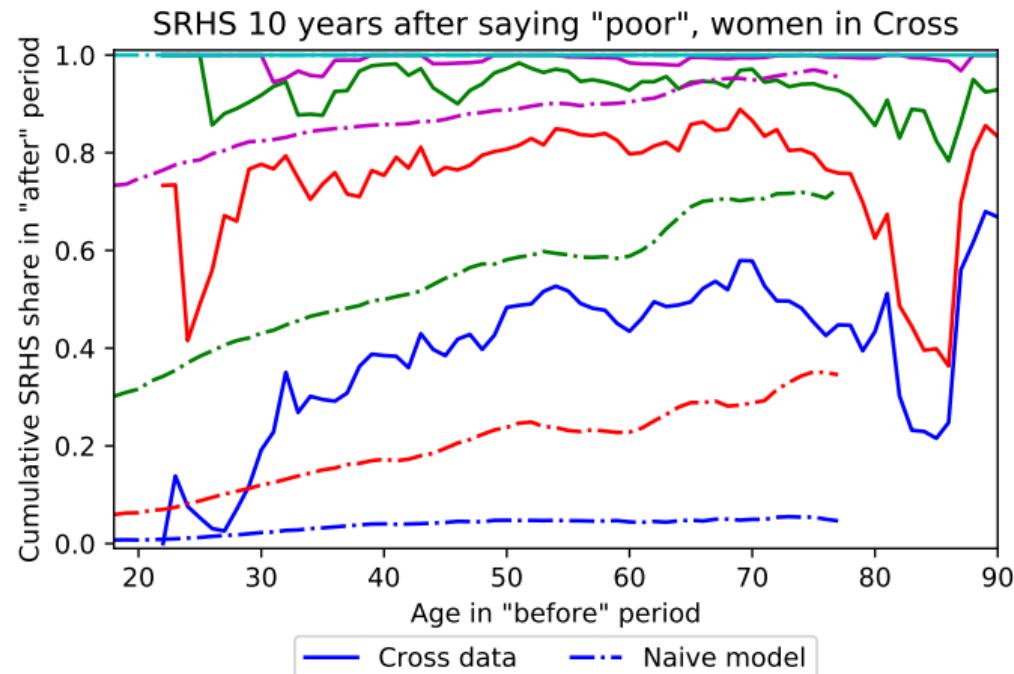
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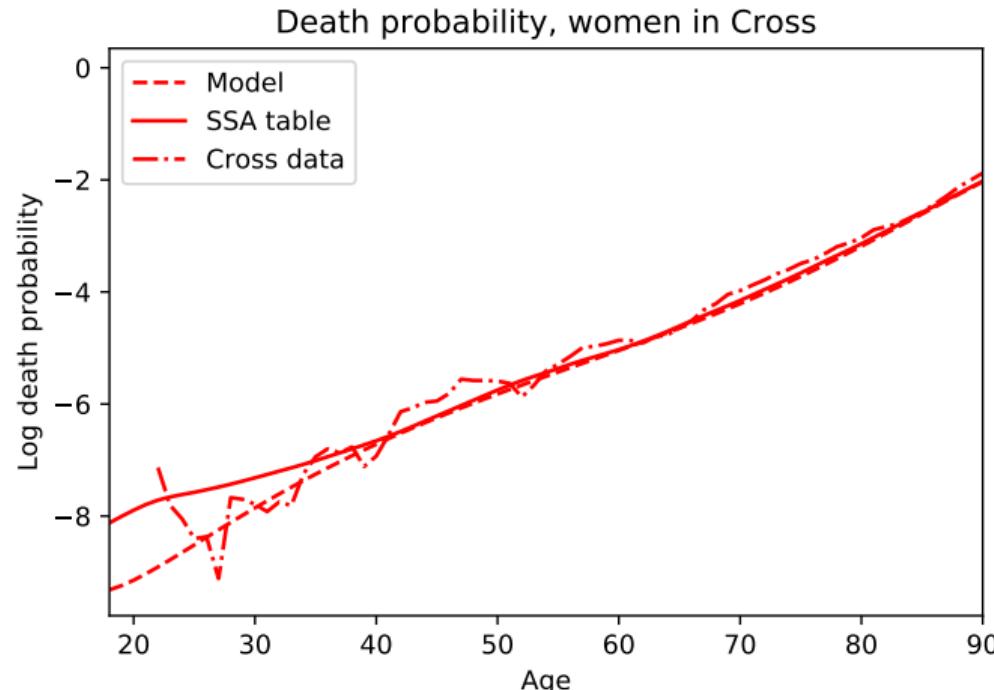
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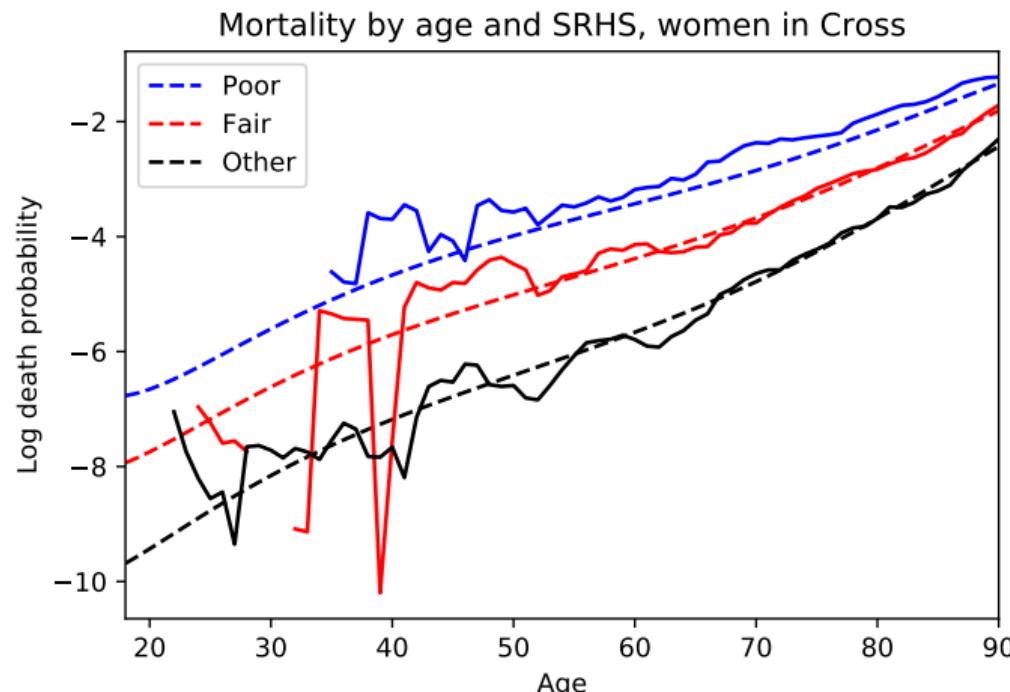
## Estimated parameters: Cross-data women (mortality)

Param	Description	Cross
$\theta_0$	Mortality probit: constant term	3.622 (0.353)
$\theta_{x1}$	Mortality probit: linear coefficient on health	0.141 (4.96e-3)
$\theta_{x2}$	Mortality probit: quadratic coefficient on health	-3.88e-3 (4.66e-4)
$\theta_{j1}$	Mortality probit: linear coefficient on age	-5.70e-2 (1.74e-2)
$\theta_{j2}$	Mortality probit: quadratic coefficient on age	8.86e-4 (2.75e-4)
$\theta_{j3}$	Mortality probit: cubic coefficient on age	-6.85e-6 (1.41e-6)
$\theta_{xj}$	Mortality probit: coefficient on age $\times$ health	-2.90e-5 (2.69e-5)

# Model fit: Mortality by age for cross-data women



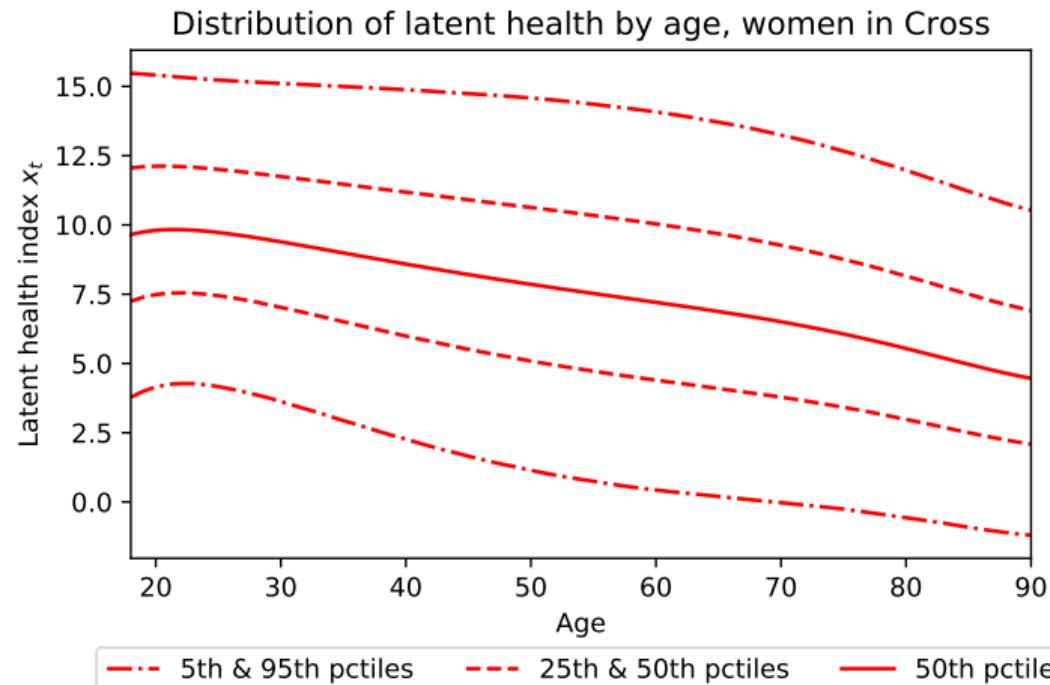
## Model fit: Mortality by age-SRHS for cross-data women



## Estimated parameters: Cross-data women (health)

Param	Description	Cross
$\beta_0$	Expected health: constant term	21.369 (1.317)
$\beta_{j1}$	Expected health: linear coefficient on age	-0.752 (9.78e-2)
$\beta_{j2}$	Expected health: quadratic coefficient on age	9.25e-3 (2.62e-3)
$\beta_{j3}$	Expected health: cubic coefficient on age	2.98e-5 (3.37e-5)
$\beta_{j4}$	Expected health: quartic coefficient on age	-9.16e-7 (1.86e-7)
$\mu_0$	Mean of latent health at model start	9.875 (0.110)
$\sigma_0$	Standard deviation of latent health at model start	3.563 (6.36e-2)

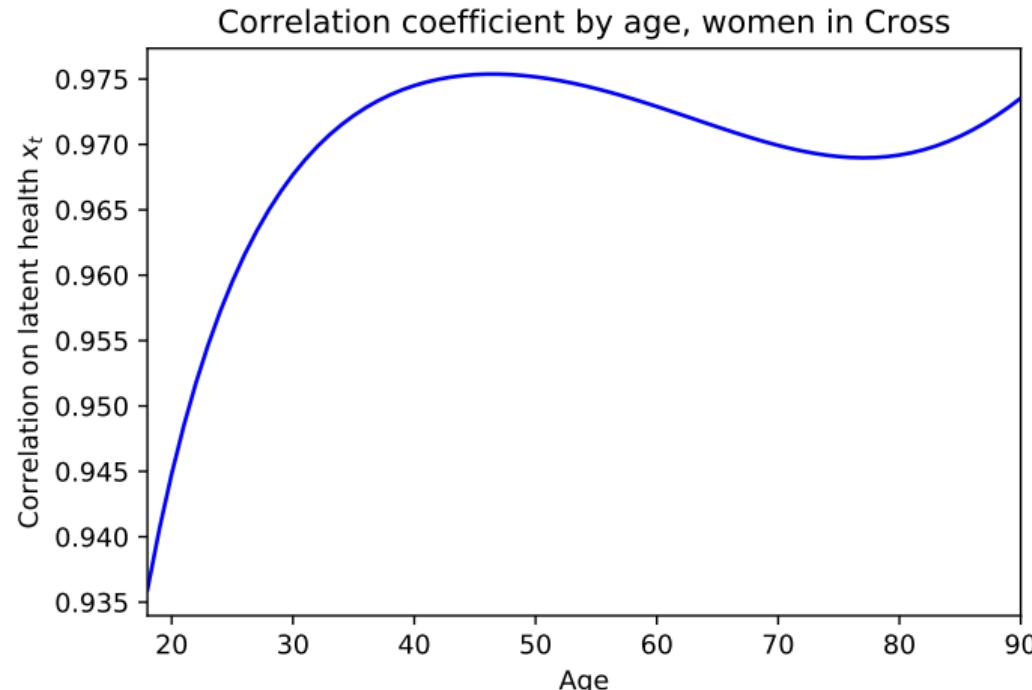
## Estimated model: Distribution of latent health (cross-data women)



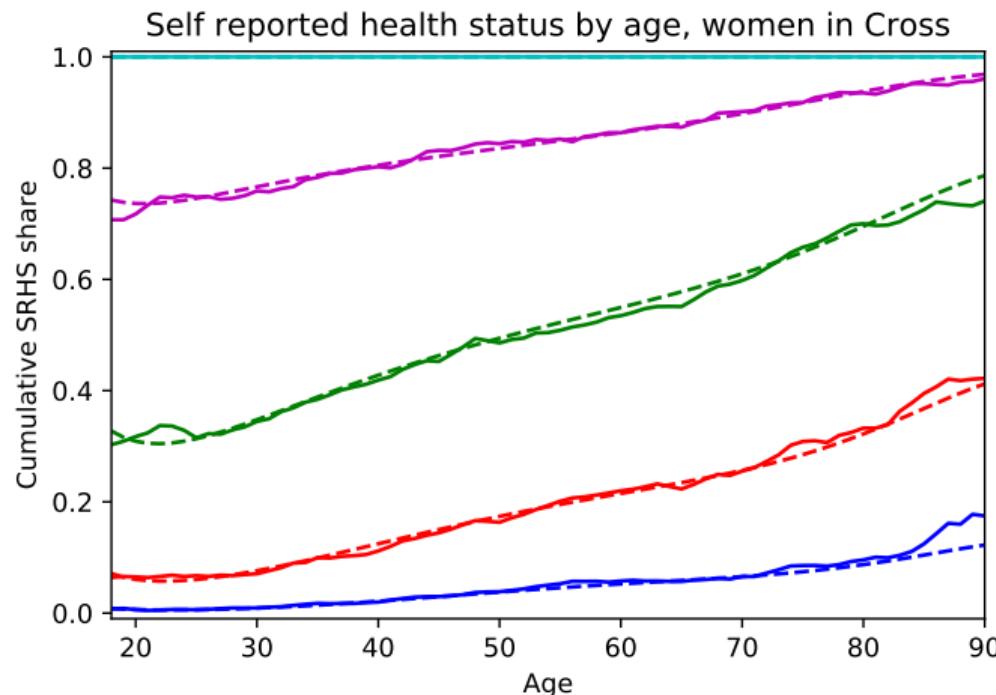
## Estimated parameters: Cross-data women (correlation and SRHS)

Param	Description	Cross
$\gamma_0$	Correlation factor: constant term	0.337 (0.186)
$\gamma_{j1}$	Correlation factor: linear coefficient on age	0.178 (1.21e-2)
$\gamma_{j2}$	Correlation factor: quadratic coefficient on age	-3.08e-3 (2.41e-4)
$\gamma_{j3}$	Correlation factor: cubic coefficient on age	1.67e-5 (1.52e-6)
$\alpha_1$	SRHS: Linear coefficient on latent health	0.480 (5.23e-3)
$\chi_2$	SRHS: Cut b/w reporting "fair" and "good"	1.805 (9.61e-3)
$\chi_3$	SRHS: Cut b/w reporting "good" and "very good"	3.814 (1.41e-2)
$\chi_4$	SRHS: Cut b/w reporting "very good" and "excellent"	5.974 (1.95e-2)

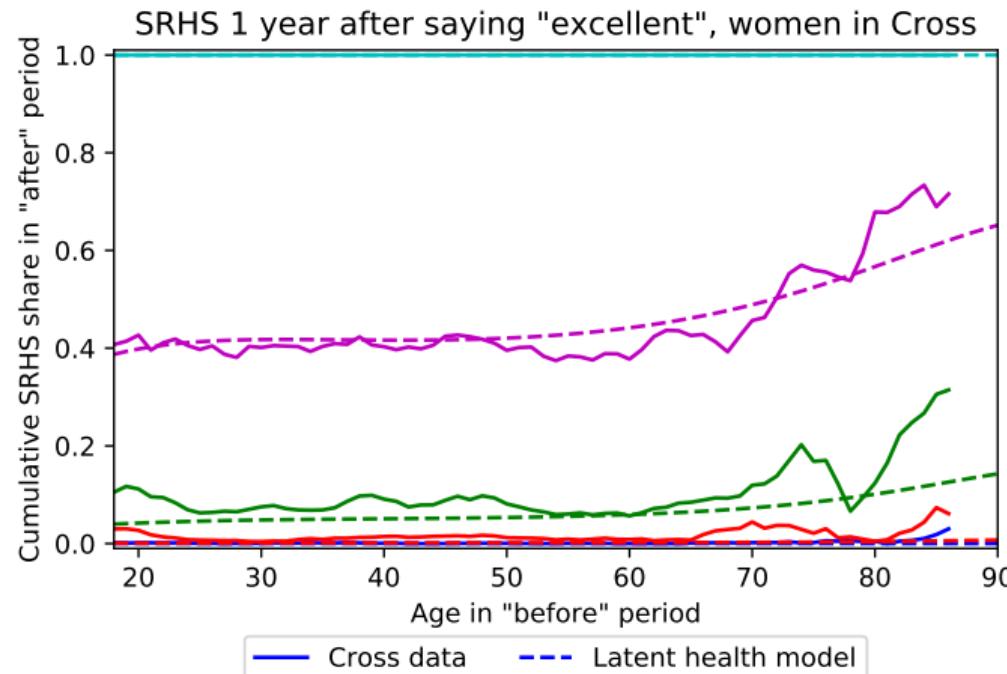
## Estimated model: Latent health serial correlation (cross-dataset women)



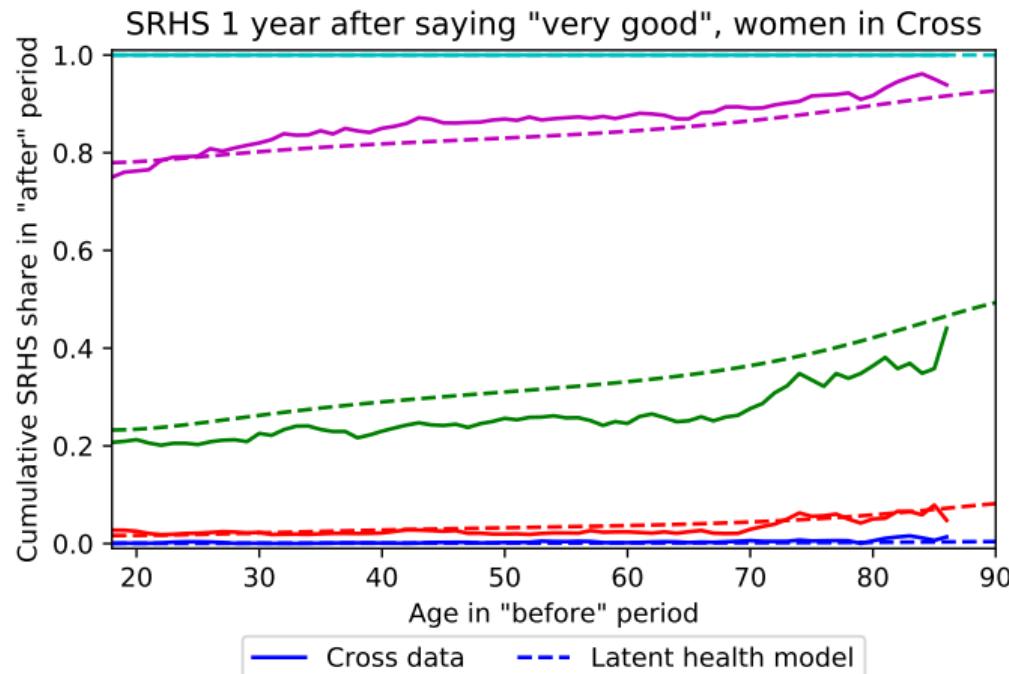
# Model fit: SRHS by age for cross-data women



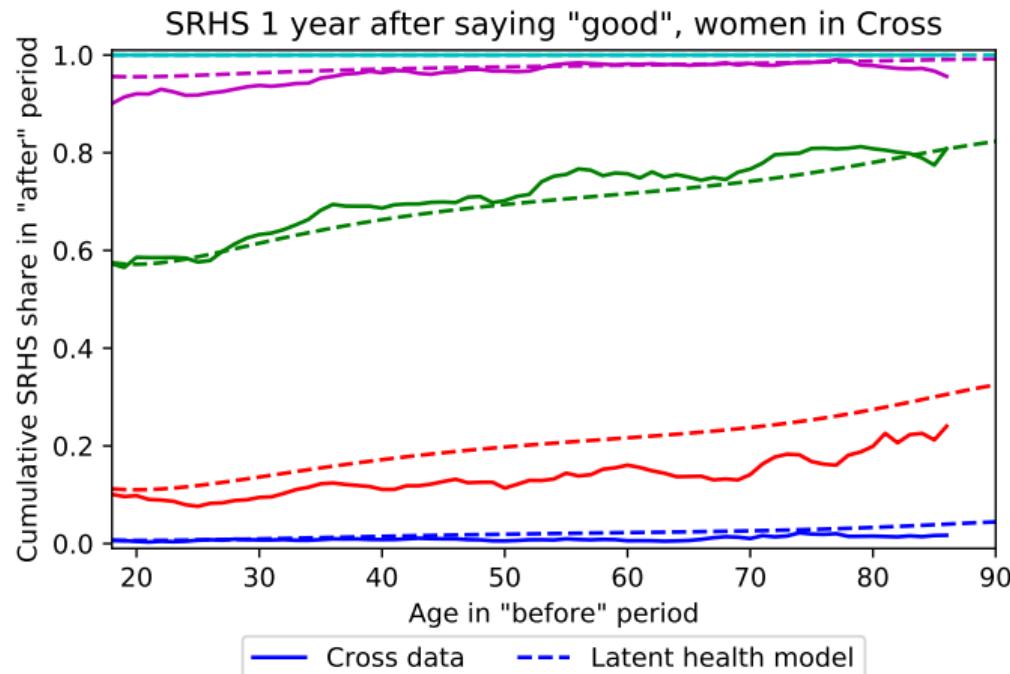
## Model fit: SRHS transitions for cross-data women: 1 year



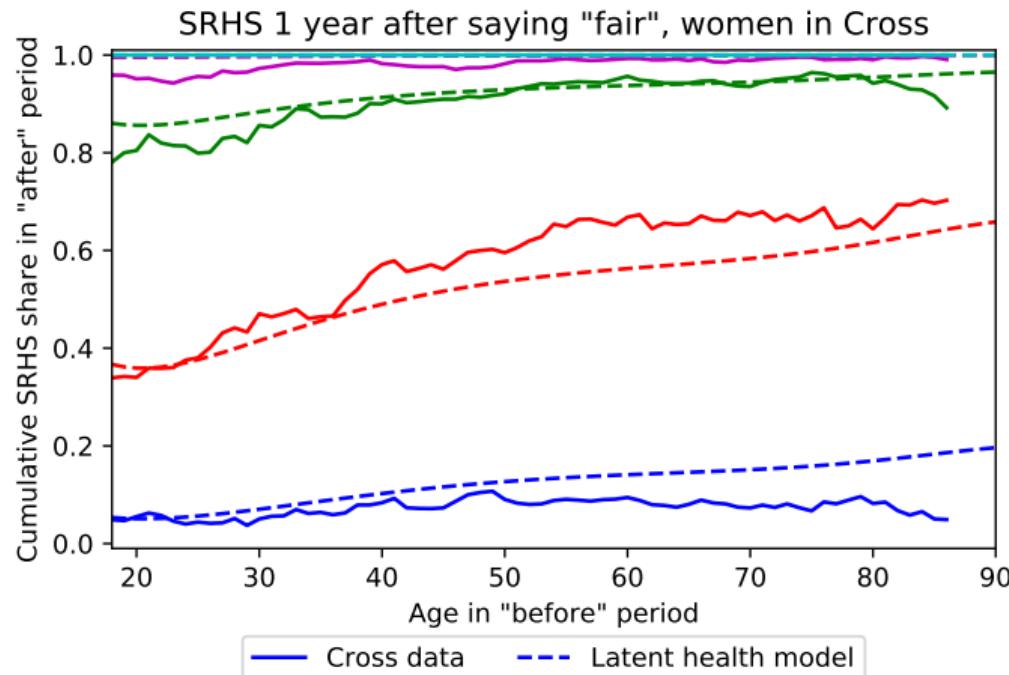
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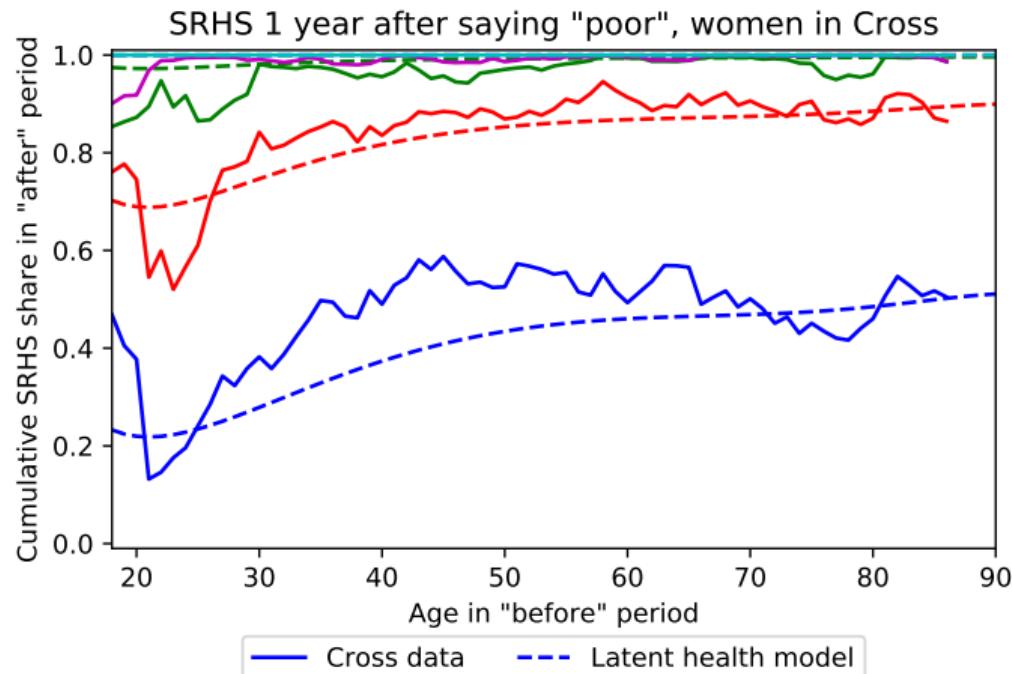
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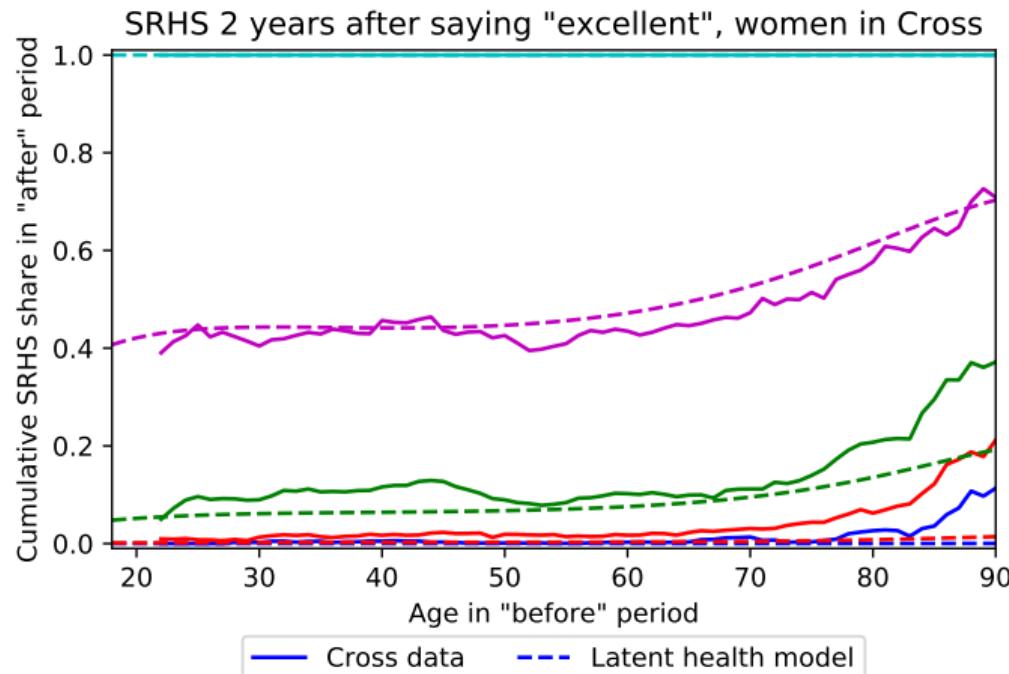
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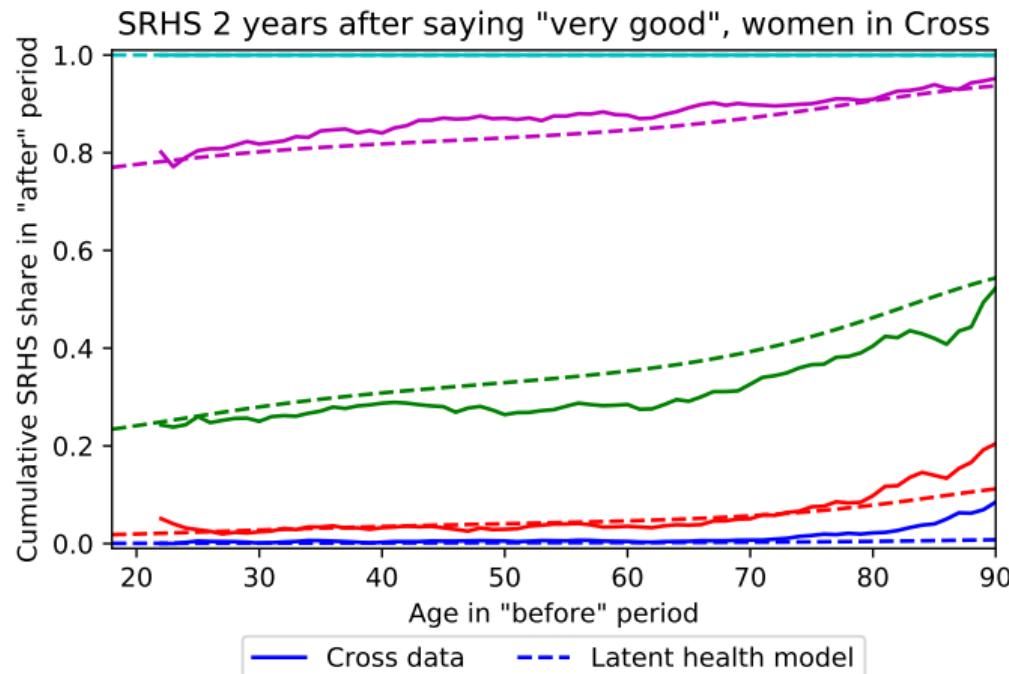
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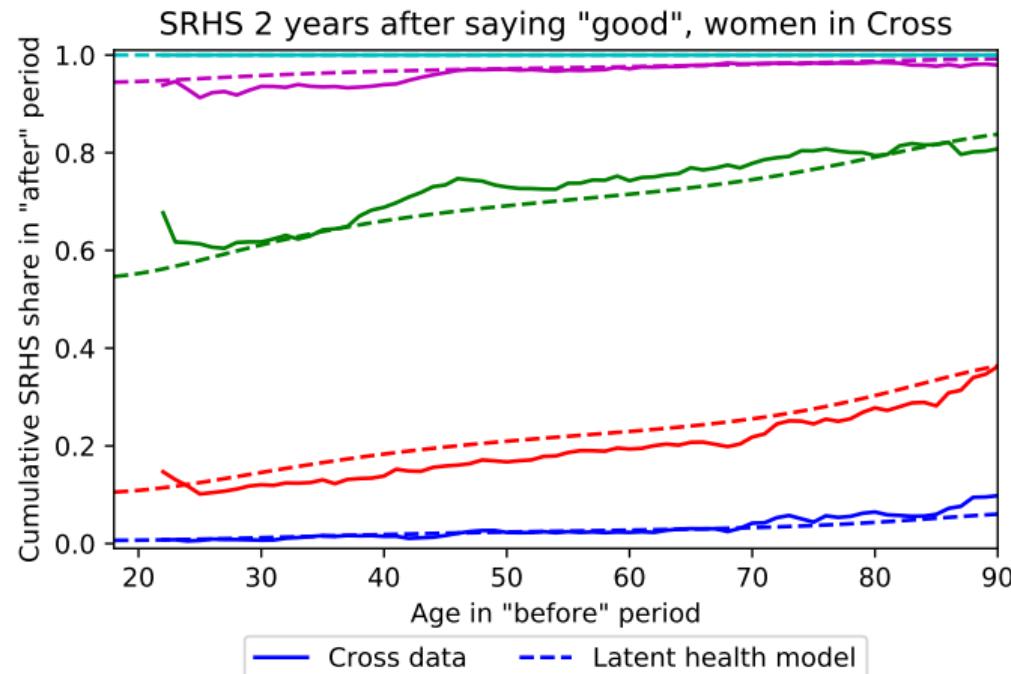
# Model fit: SRHS transitions for cross-data women: 2 years



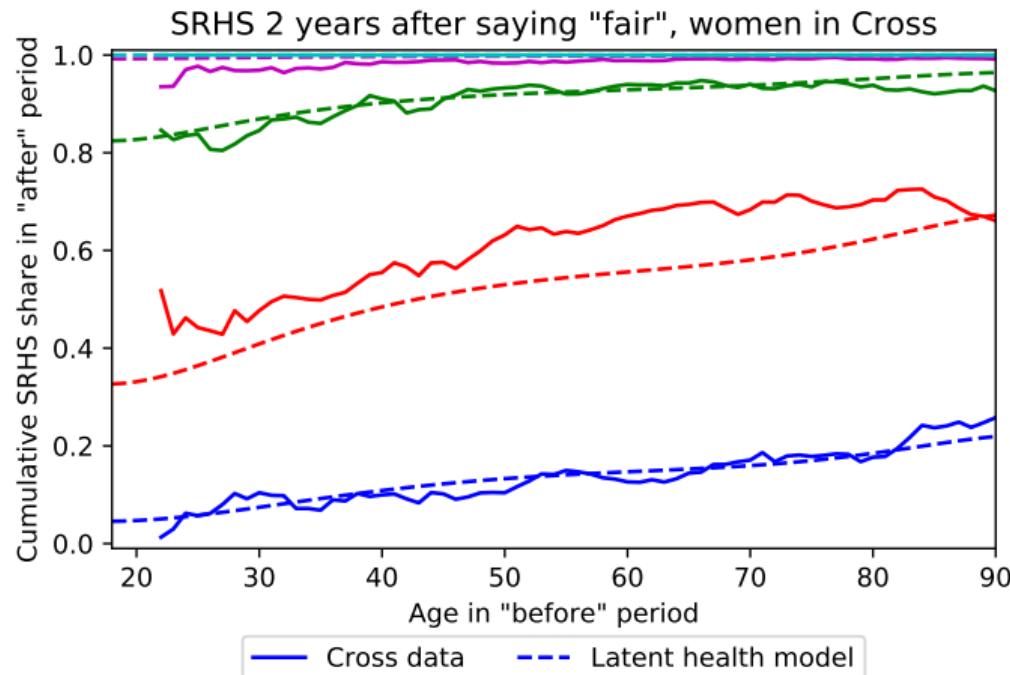
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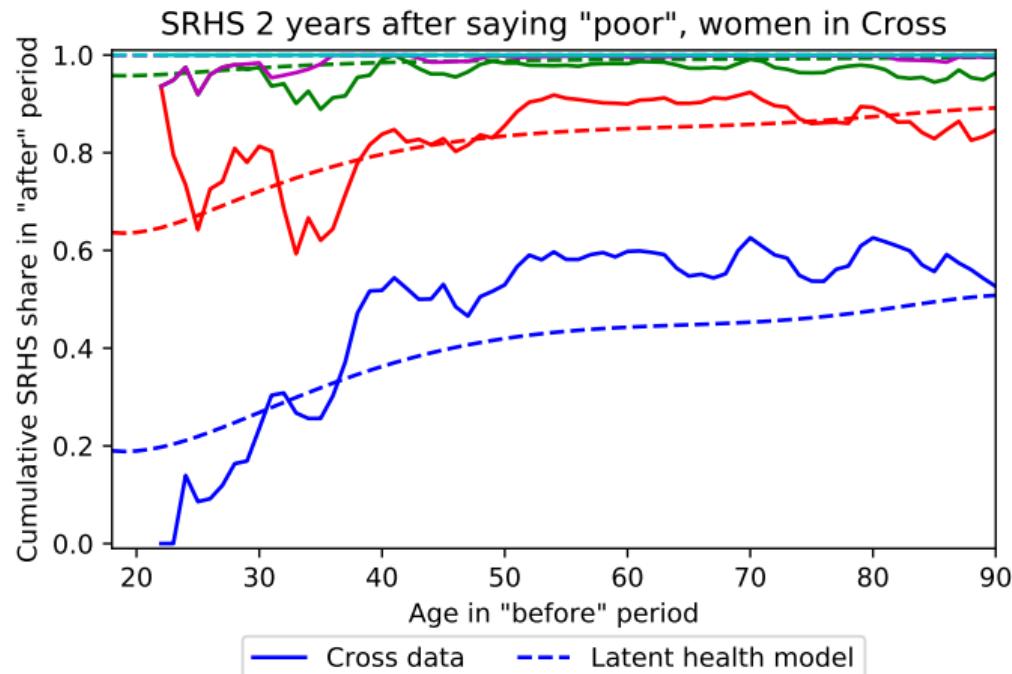
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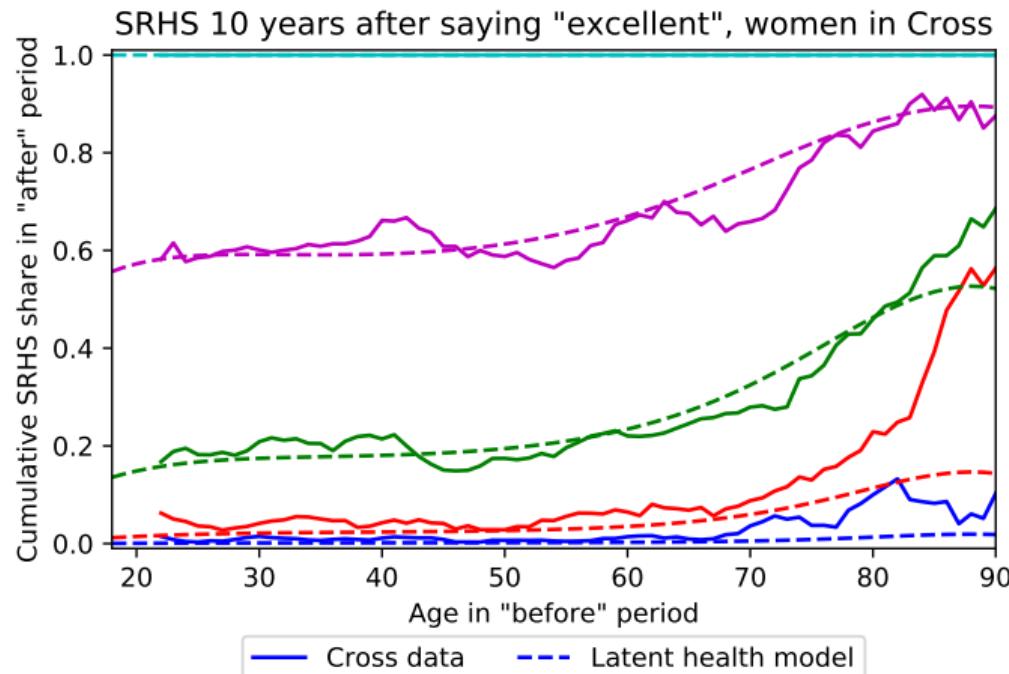
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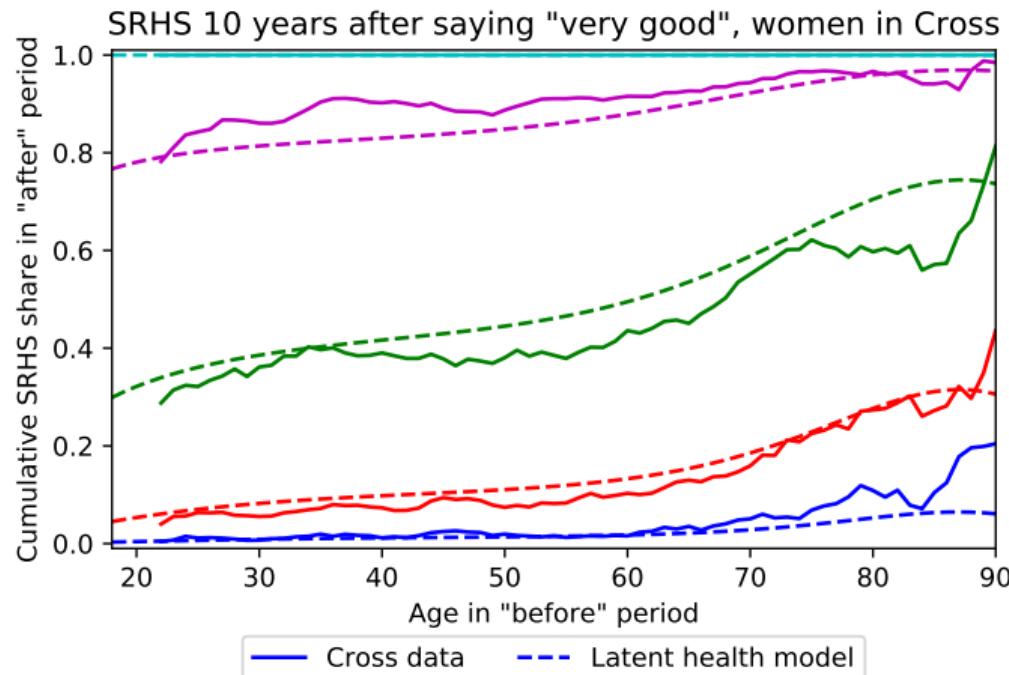
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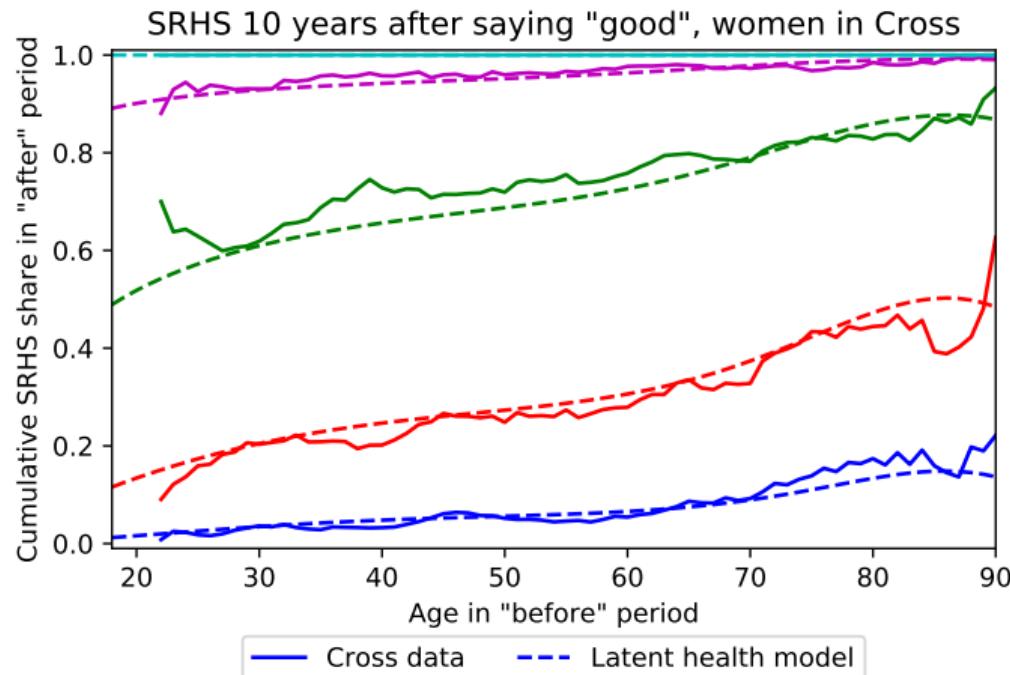
# Model fit: SRHS transitions for cross-data women: 10 yrs



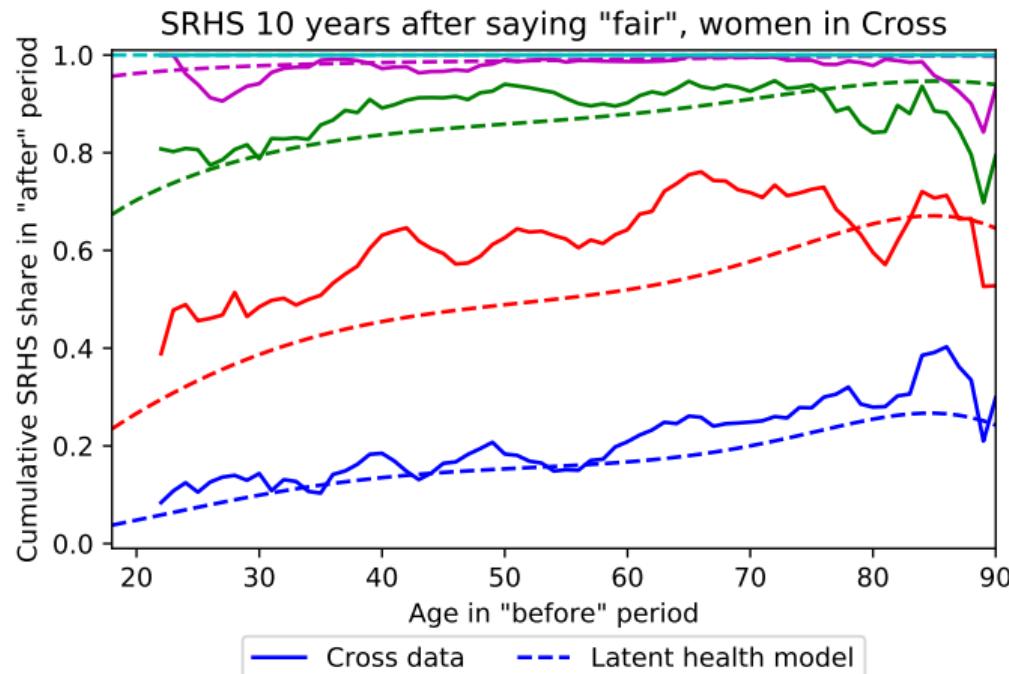
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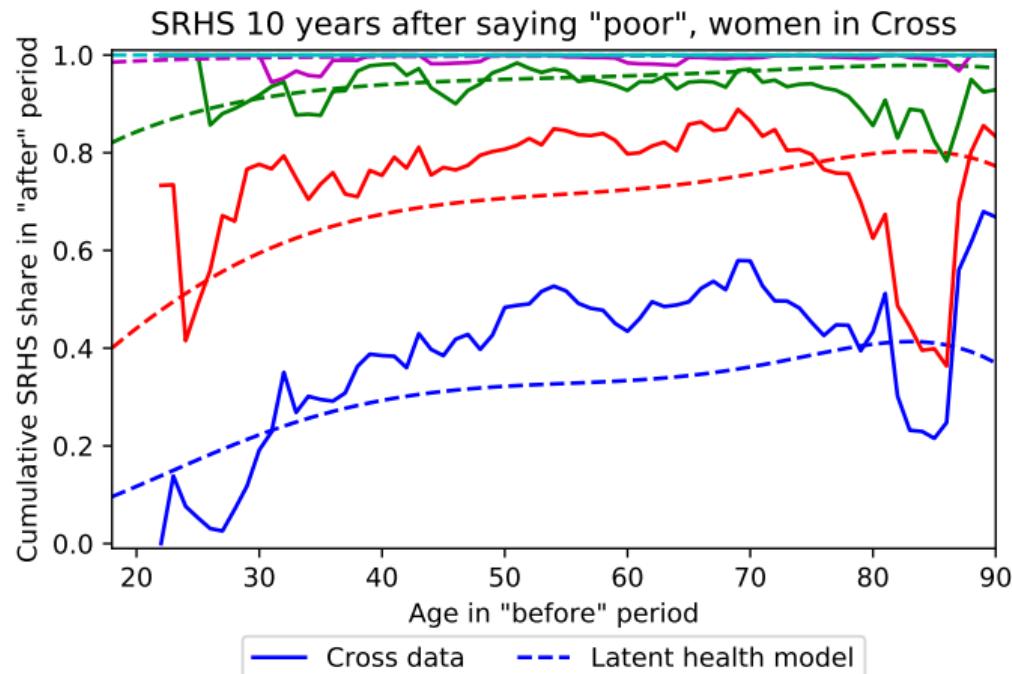
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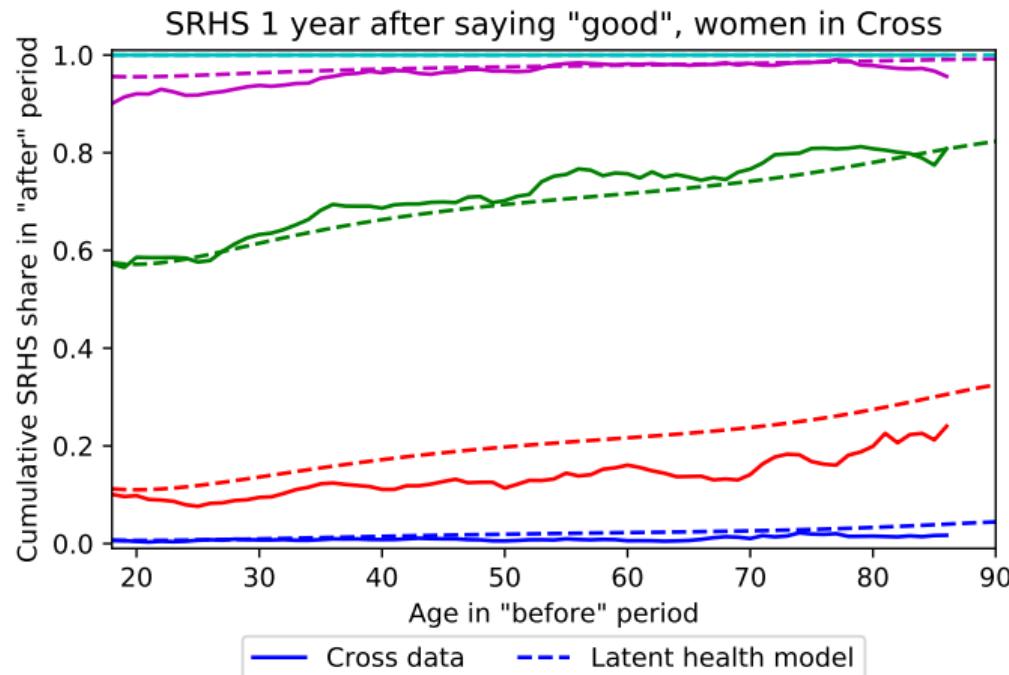
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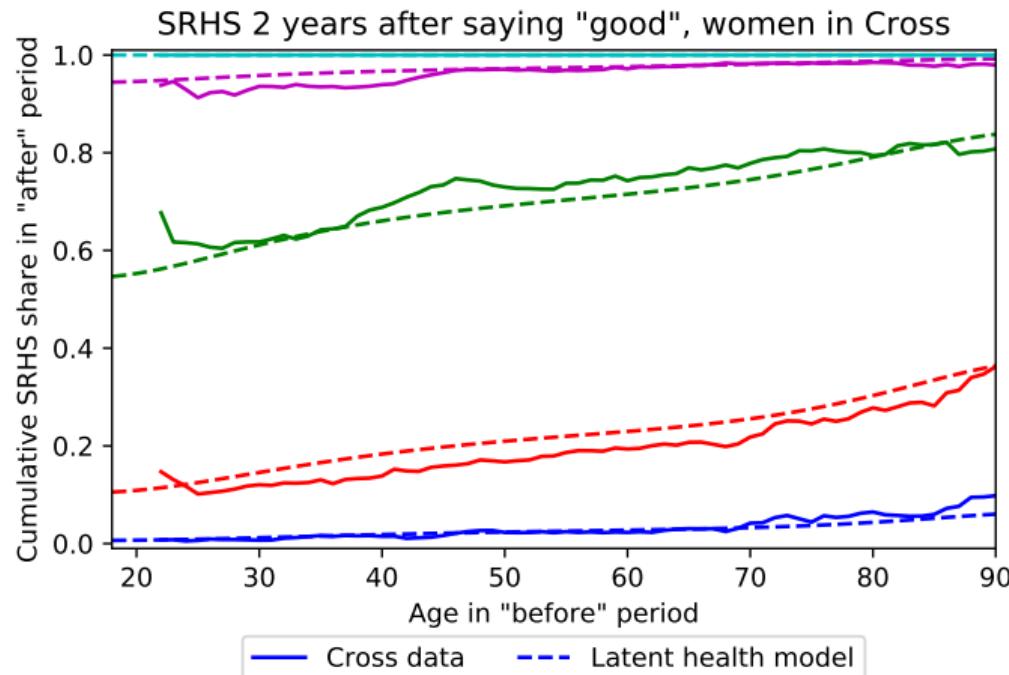
## Model fit: SRHS transitions for cross-data women: 10 yrs



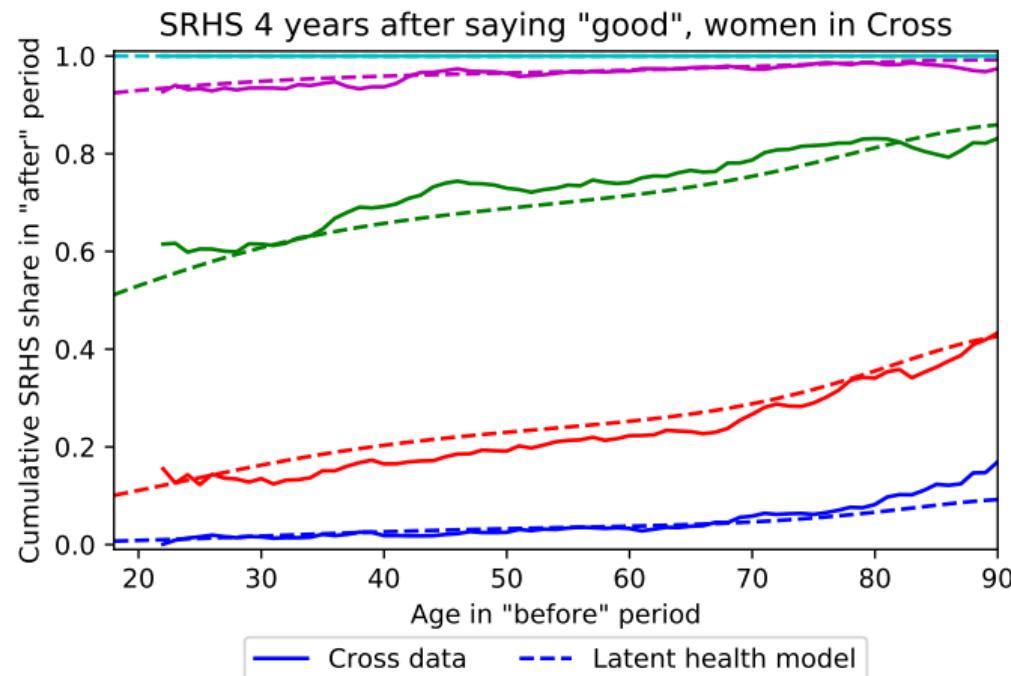
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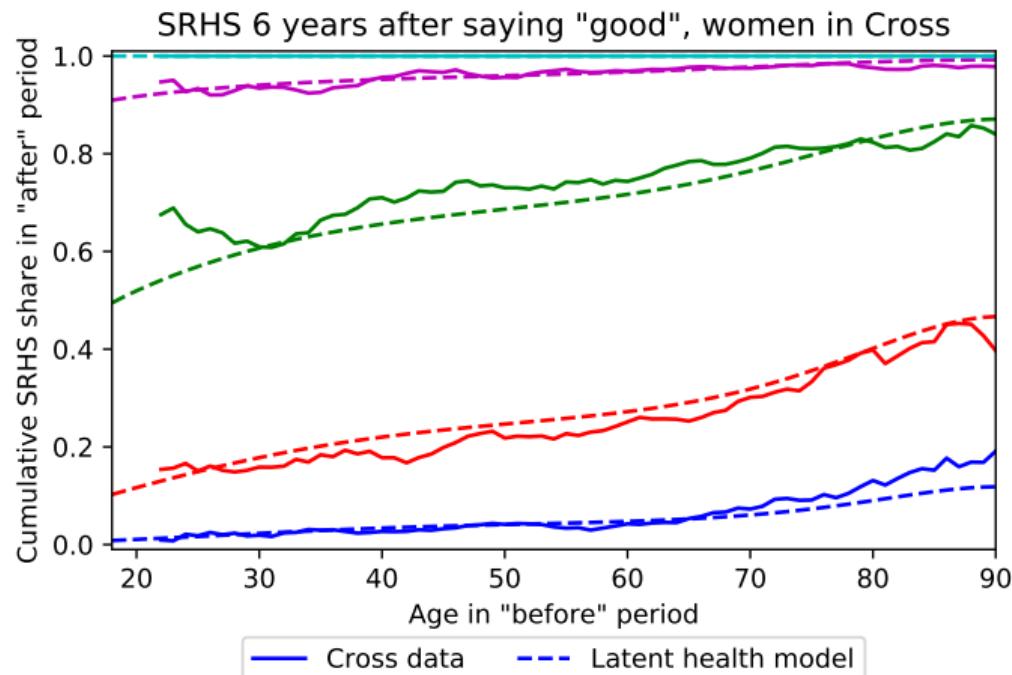
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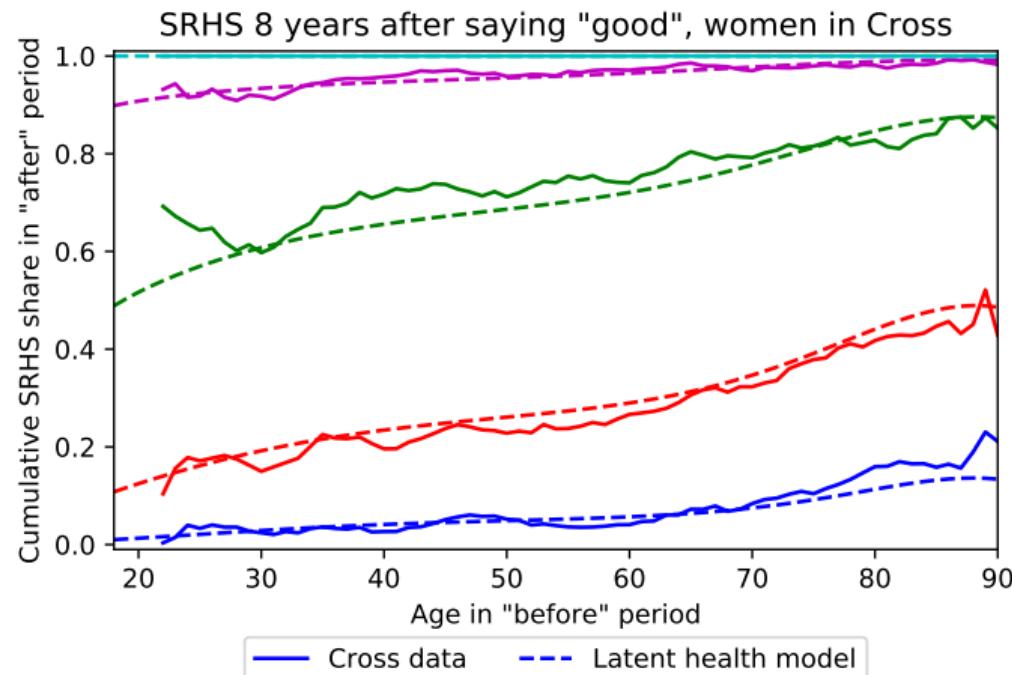
## Model fit: SRHS transitions for cross-data women: 4 years



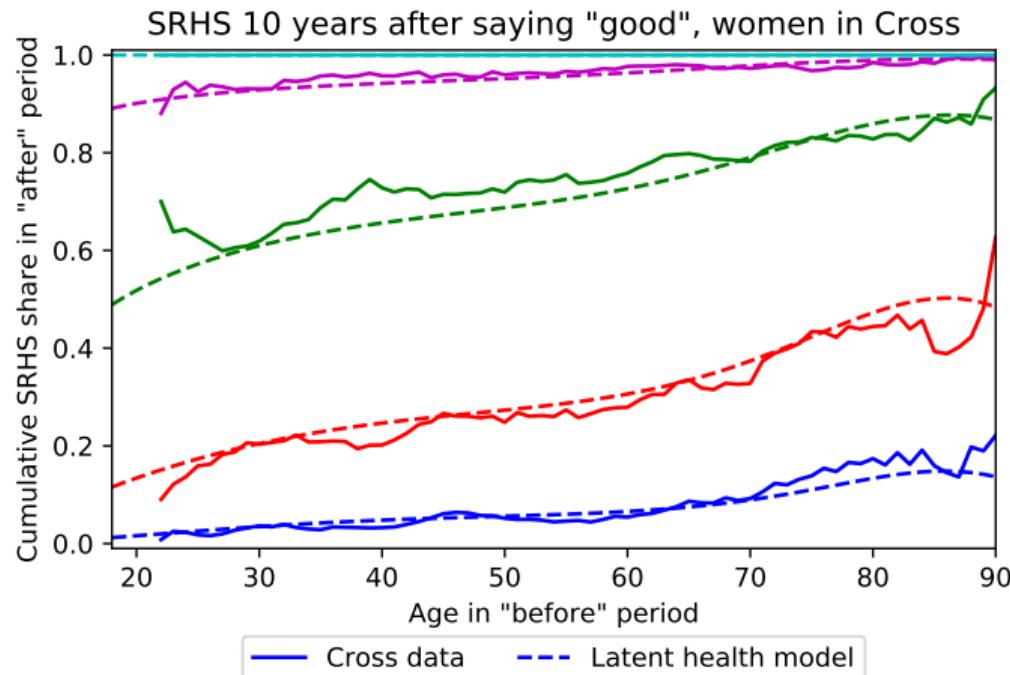
# Model fit: SRHS transitions for cross-data women: 6 years



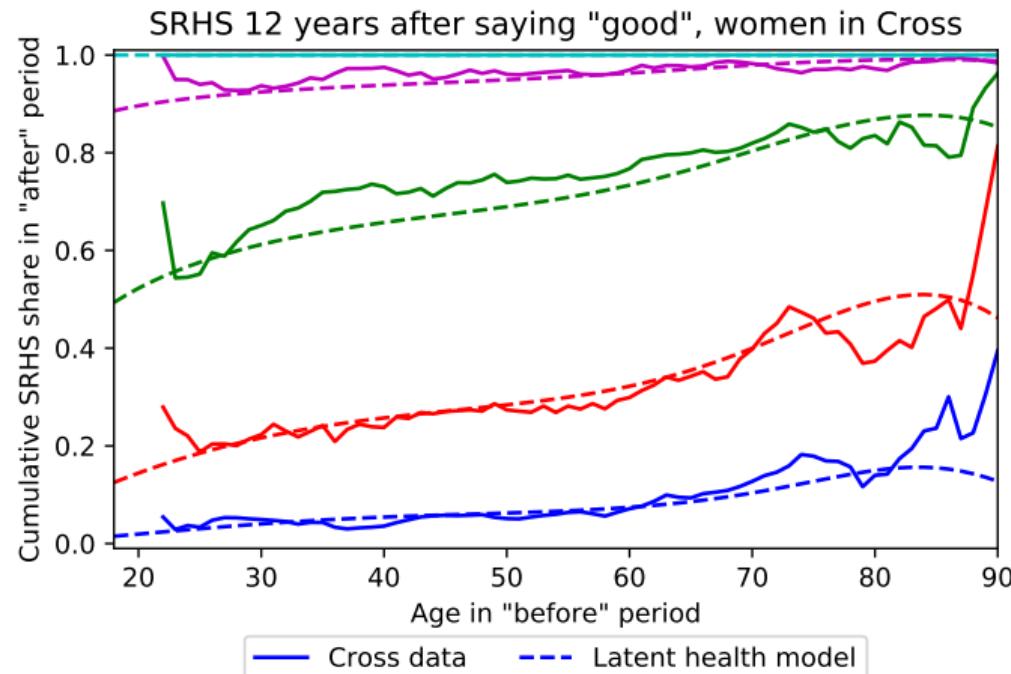
## Model fit: SRHS transitions for cross-data women: 8 years



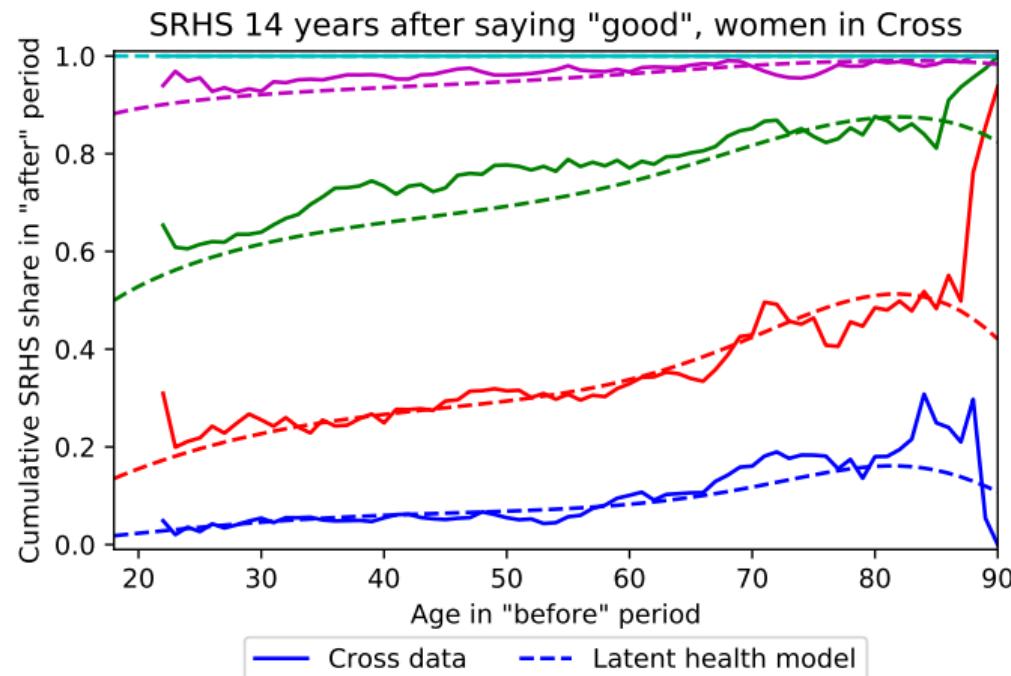
## Model fit: SRHS transitions for cross-data women: 10 yrs



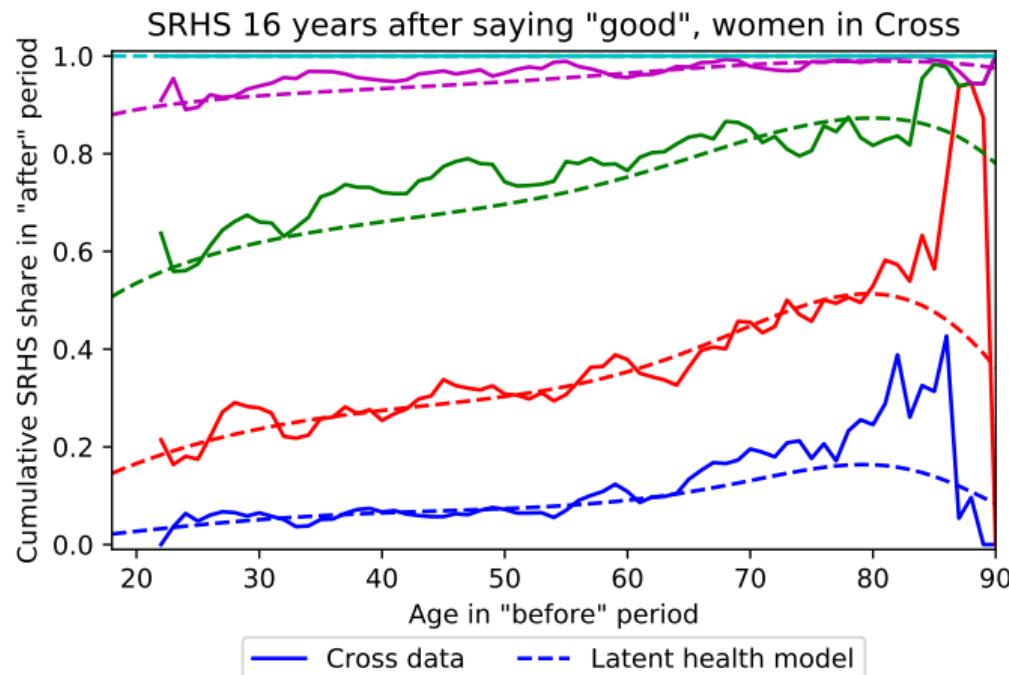
## Model fit: SRHS transitions for cross-data women: 12 yrs



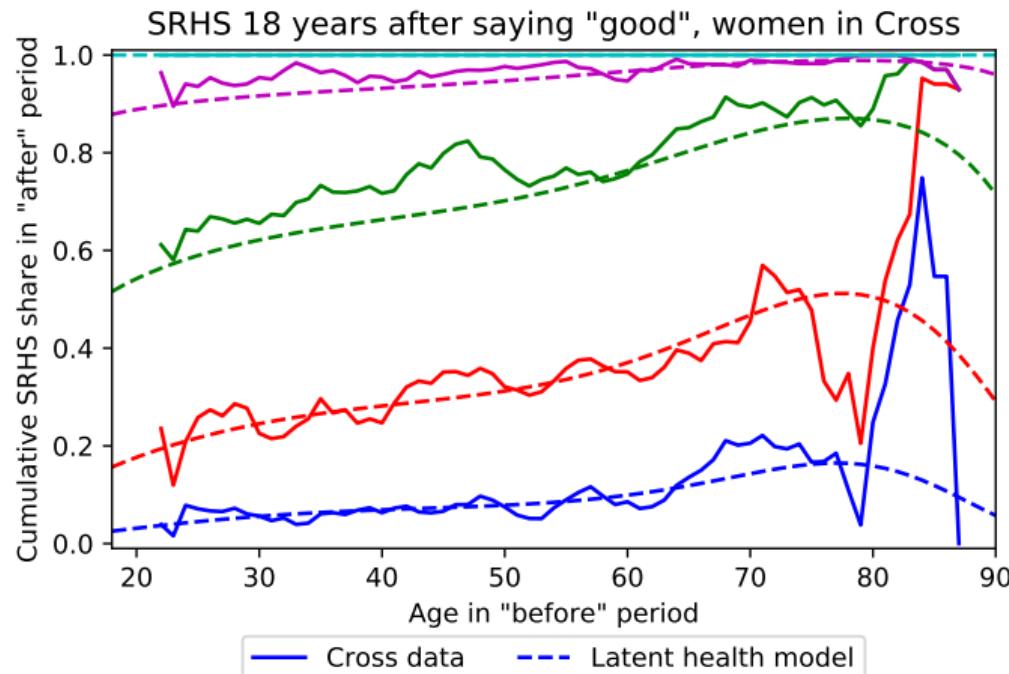
## Model fit: SRHS transitions for cross-data women: 14 yrs



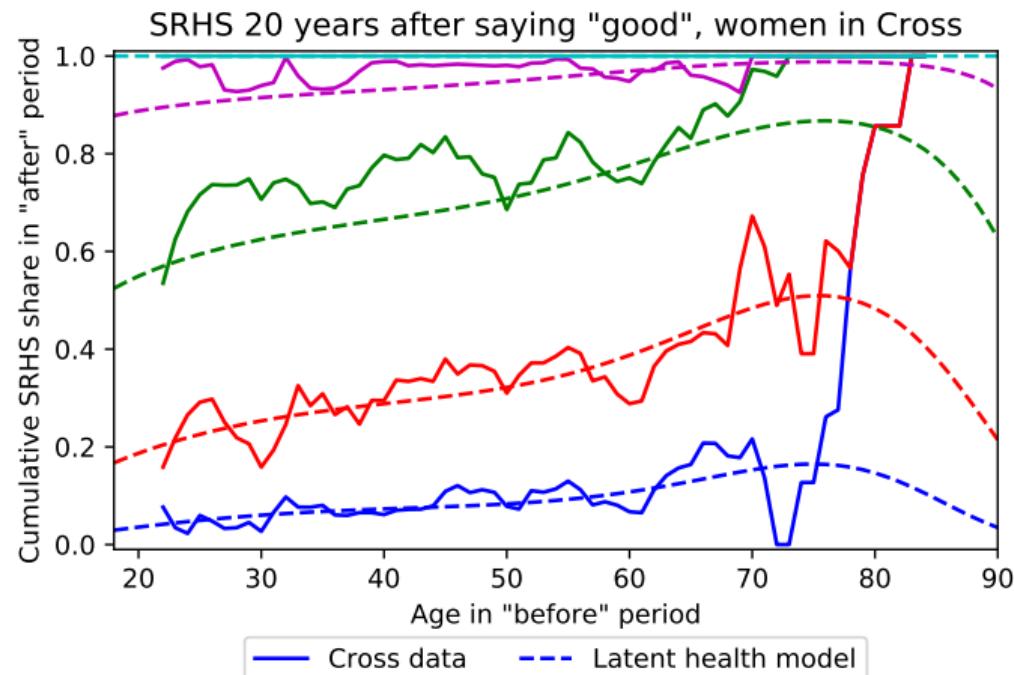
## Model fit: SRHS transitions for cross-data women: 16 yrs



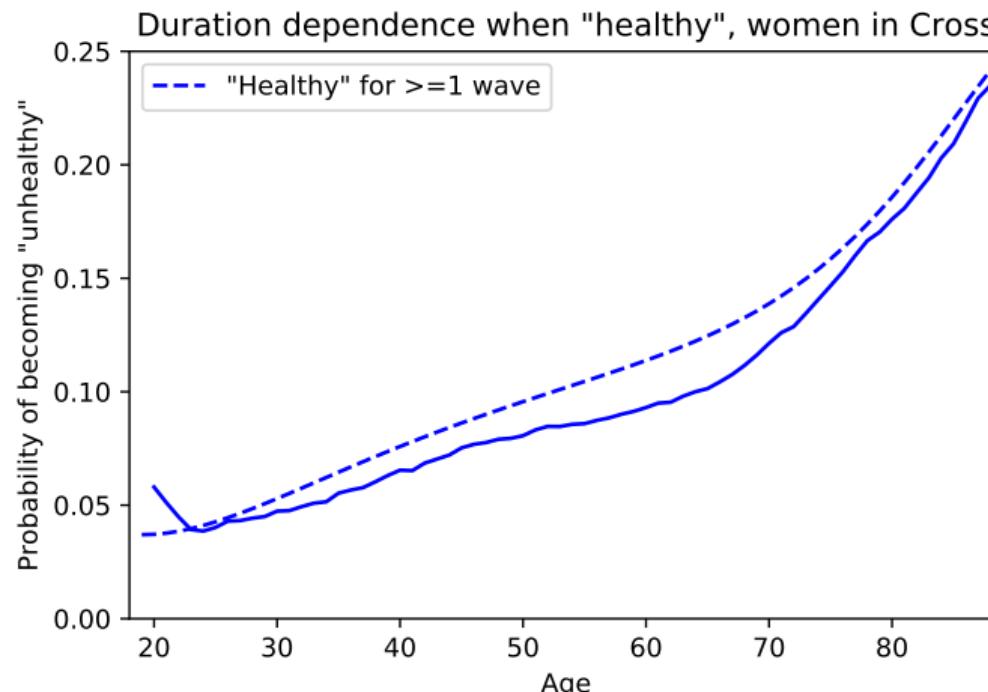
## Model fit: SRHS transitions for cross-data women: 18 yrs



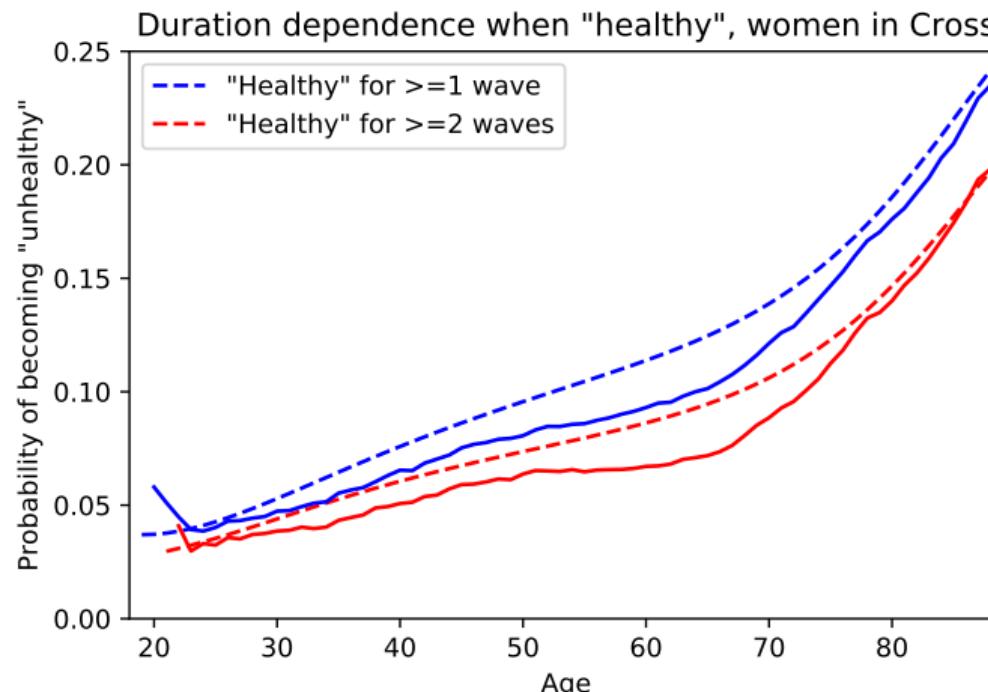
## Model fit: SRHS transitions for cross-data women: 20 yrs



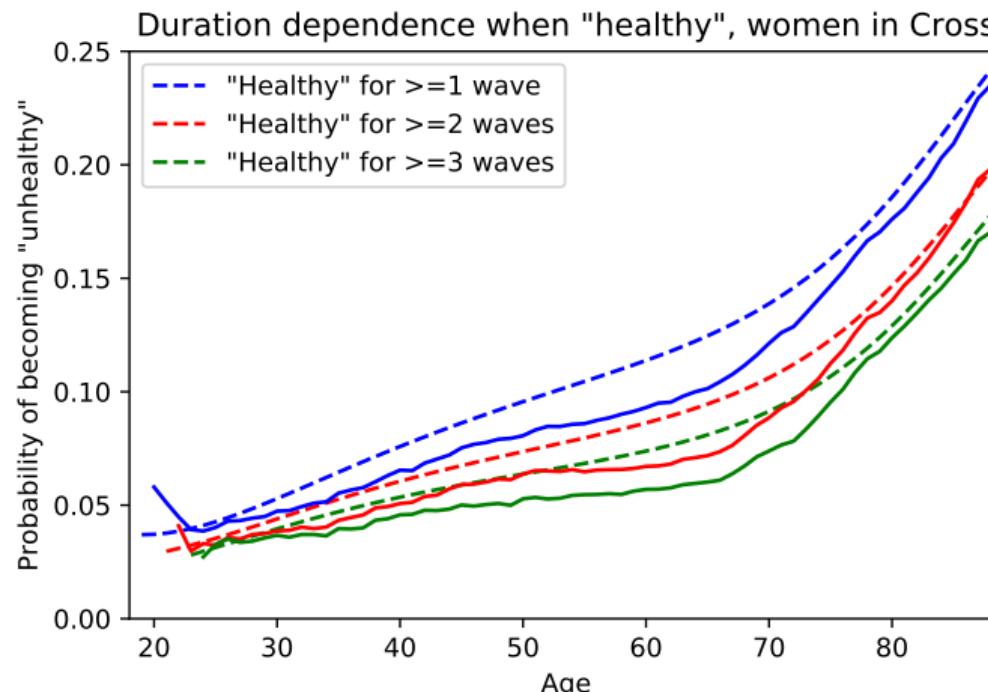
## Model fit: Duration dependence, "healthy" to "unhealthy"



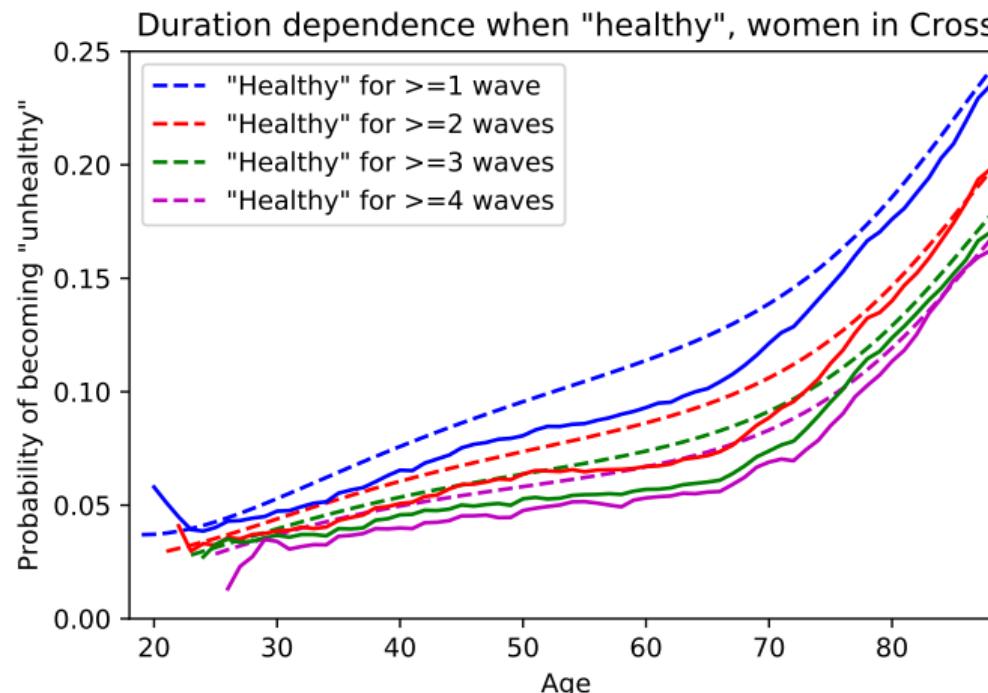
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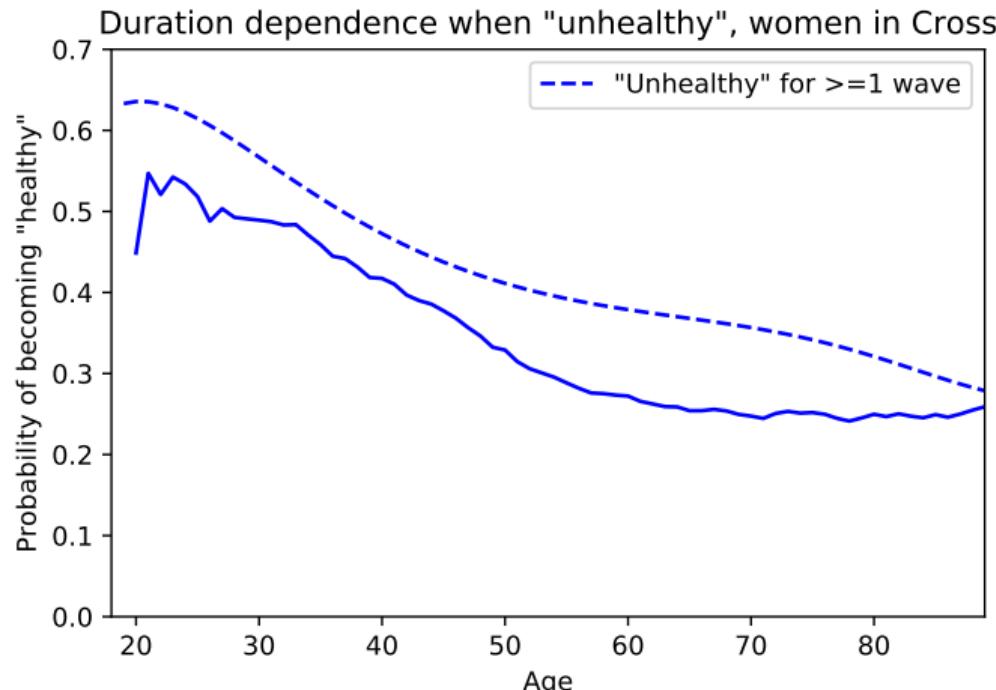
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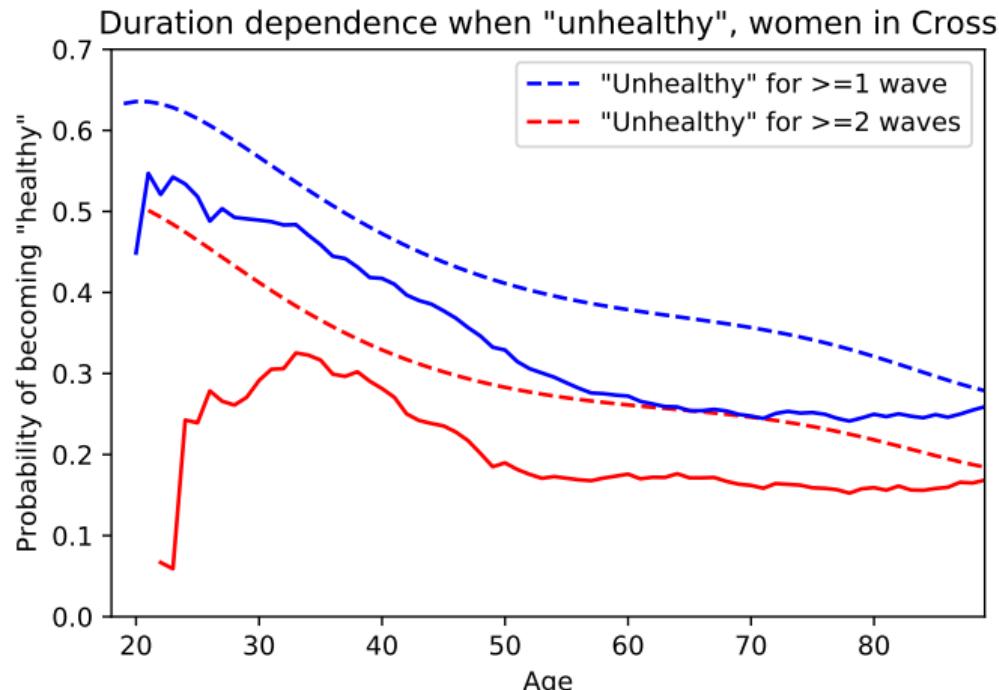
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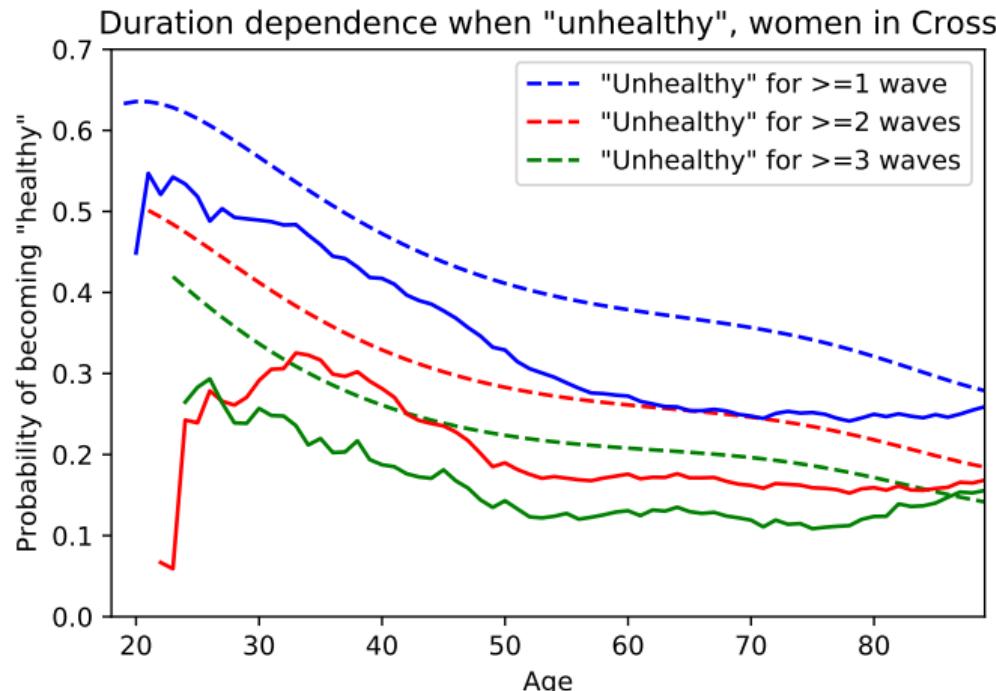
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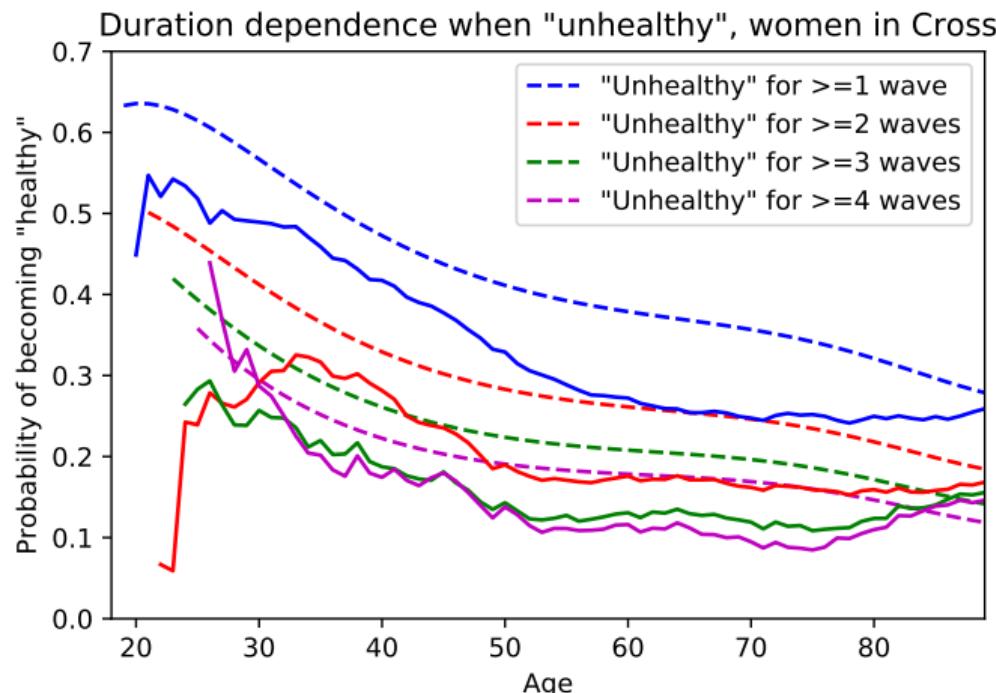
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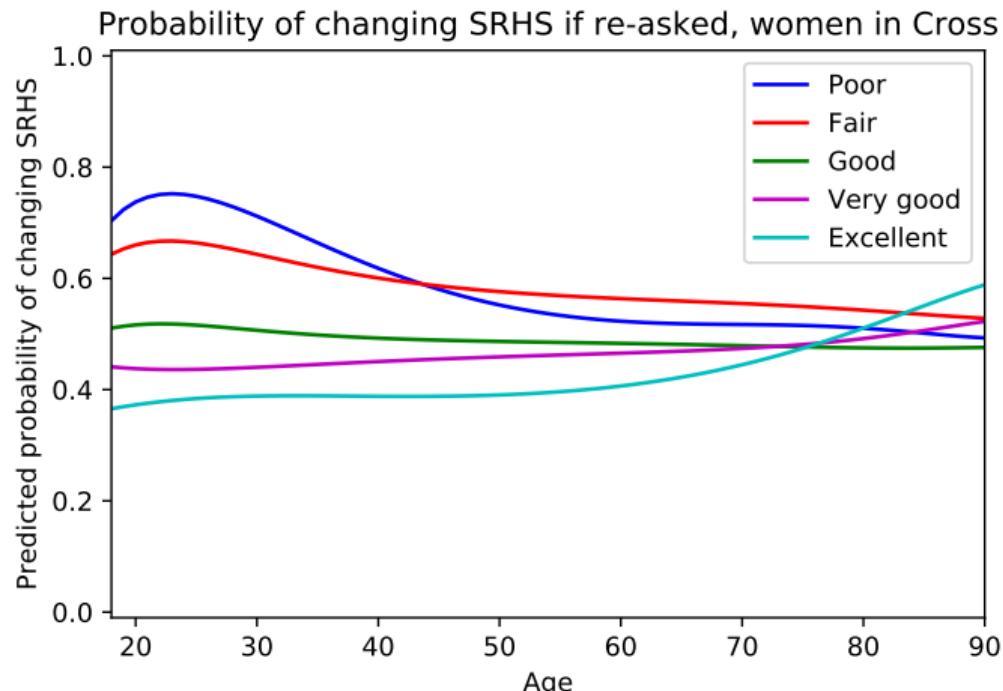
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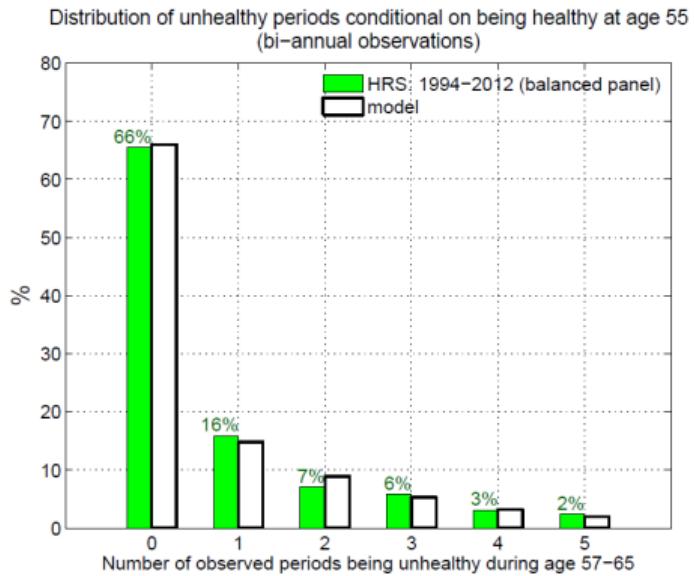


## Estimated model: Changing SRHS if asked twice (cross-data women)

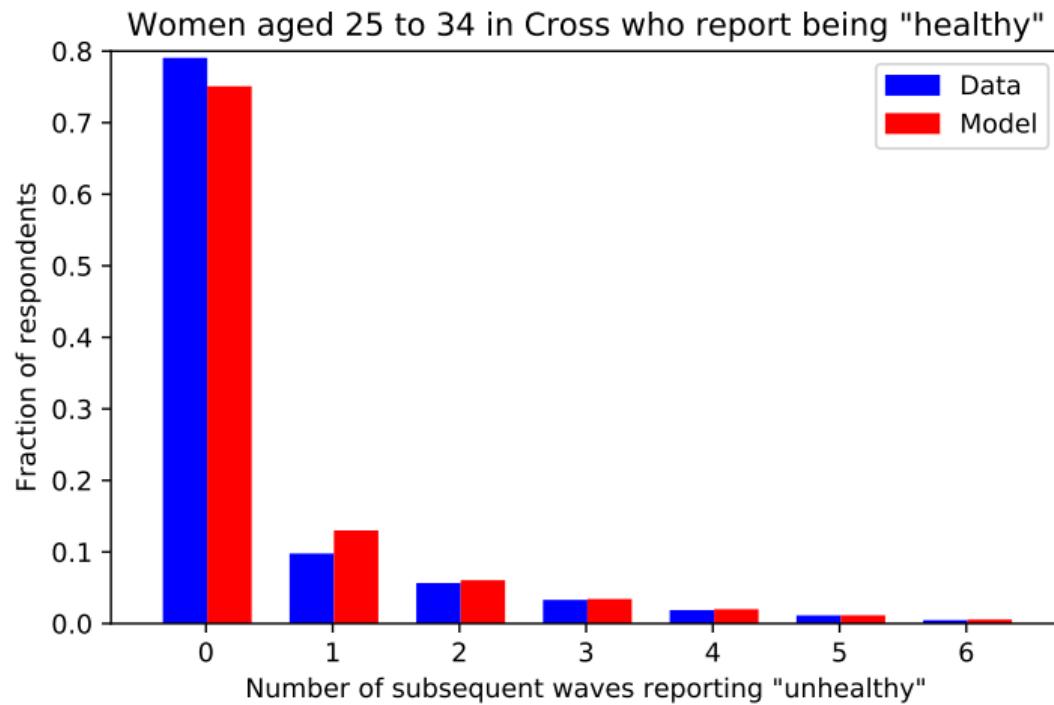


# Model fit: Within-person distribution of “unhealthy” periods

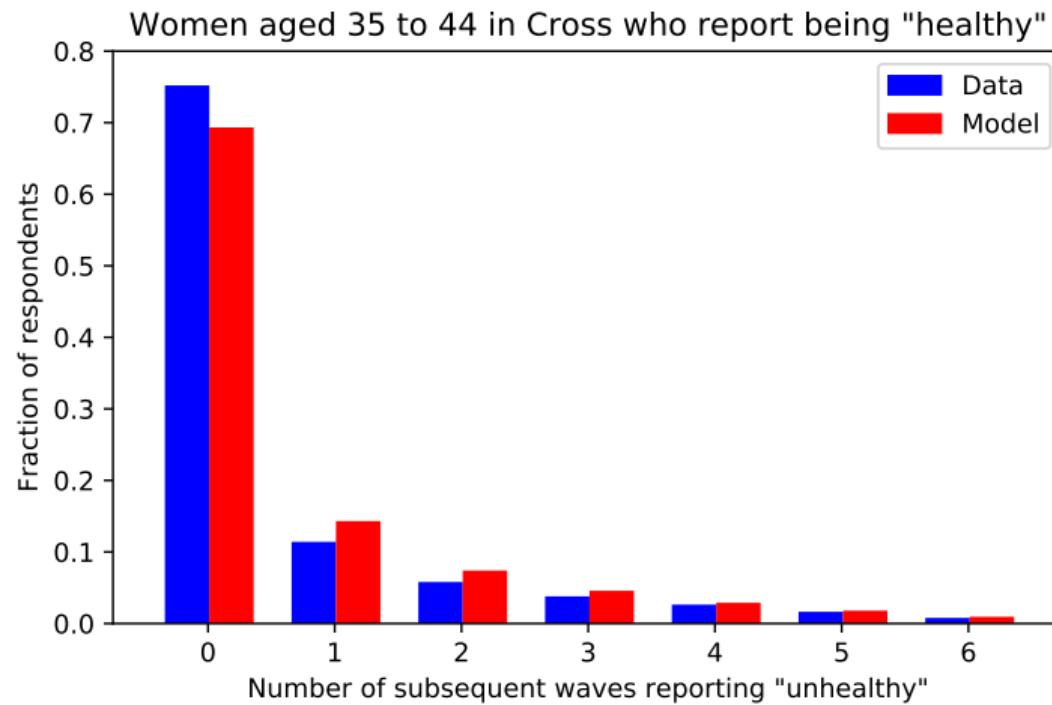
- DeNardi et al (2018) show that their duration dependence / unobserved heterogeneity model fits the distribution of “unhealthy” periods over next 5 waves
- Can the latent health model do that as well?



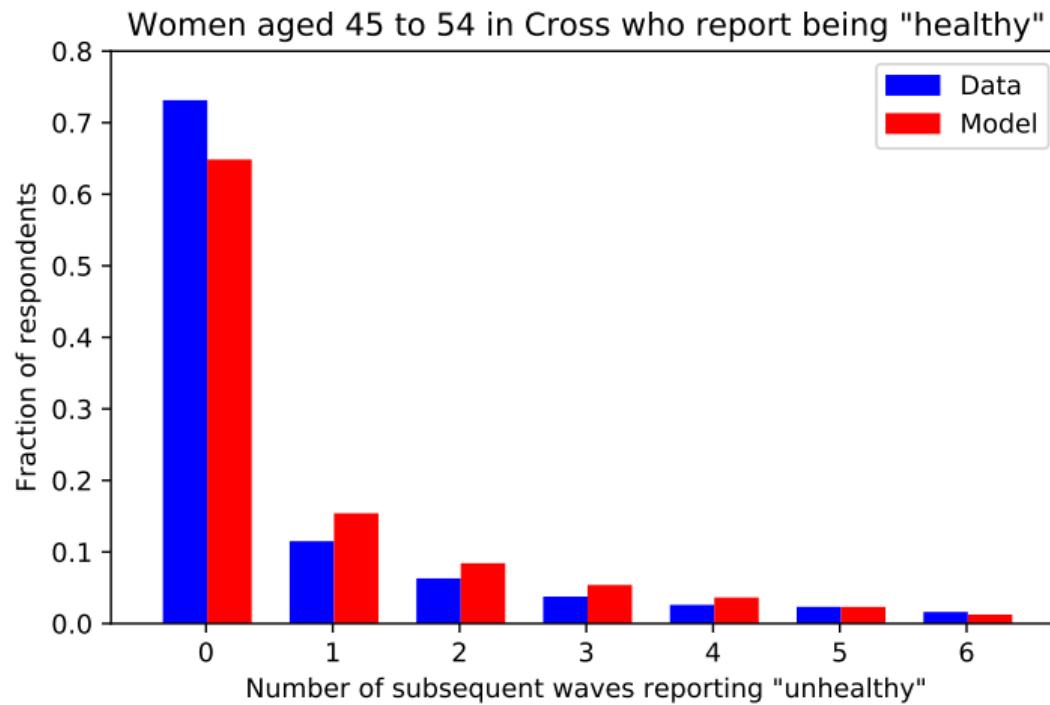
## Model fit: Frequency of reporting “unhealthy” (cross-data women)



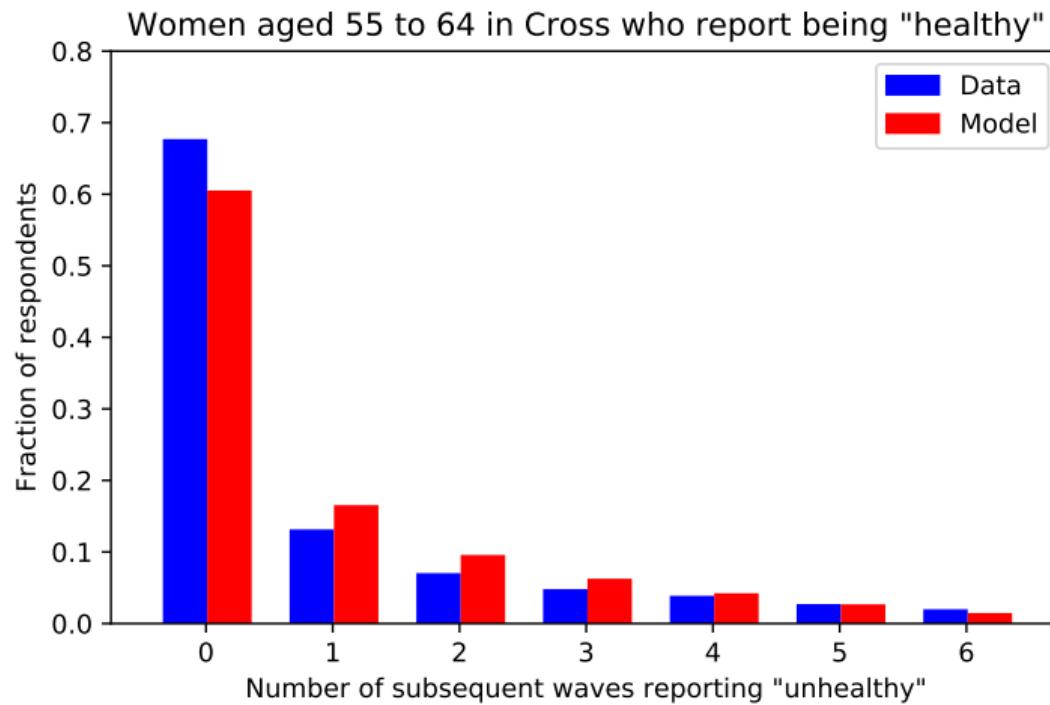
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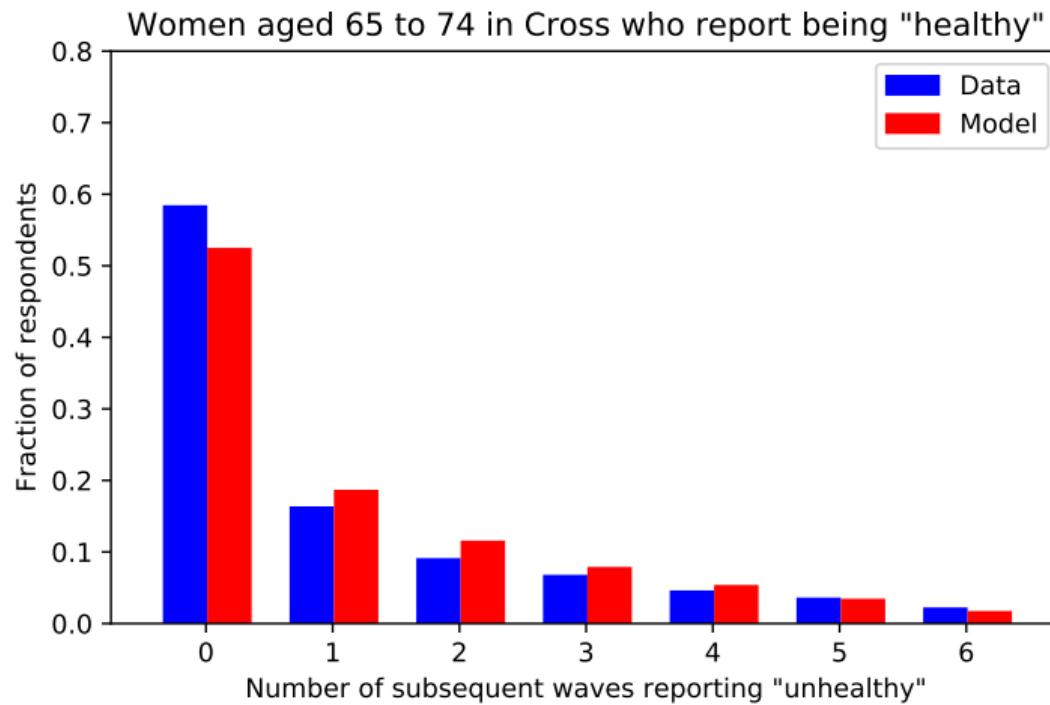
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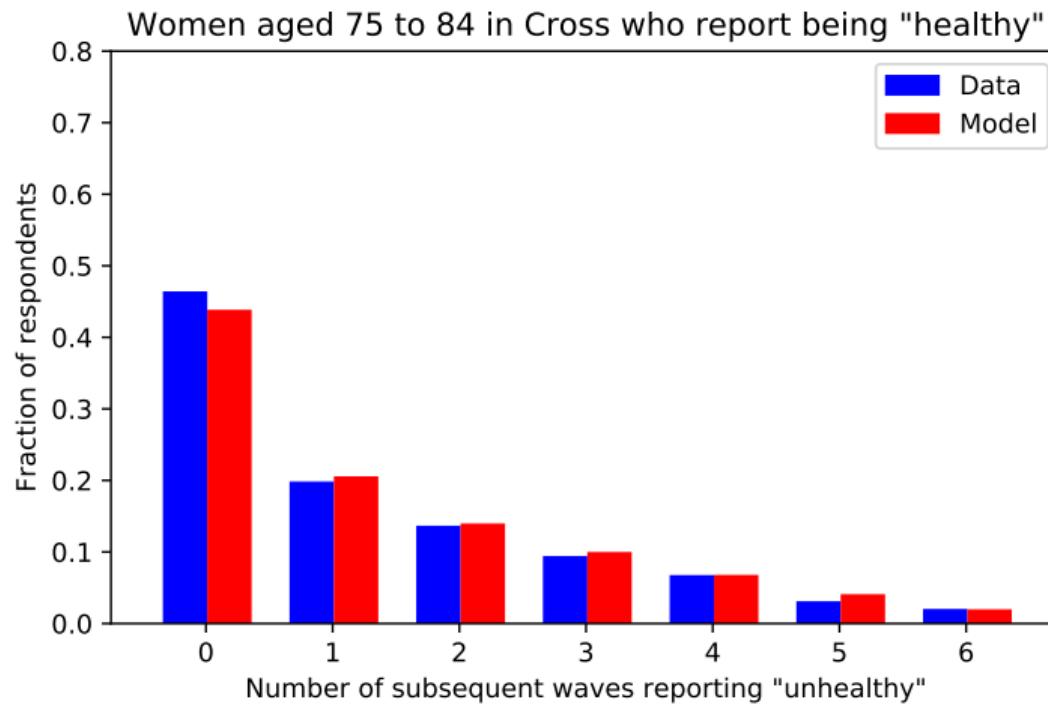
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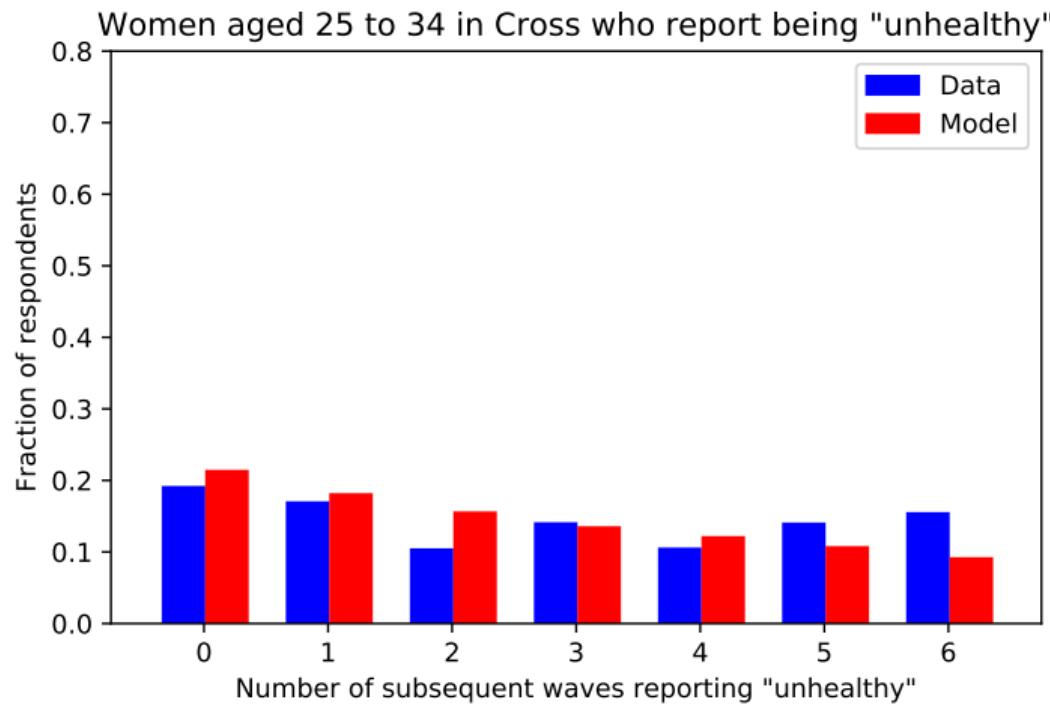
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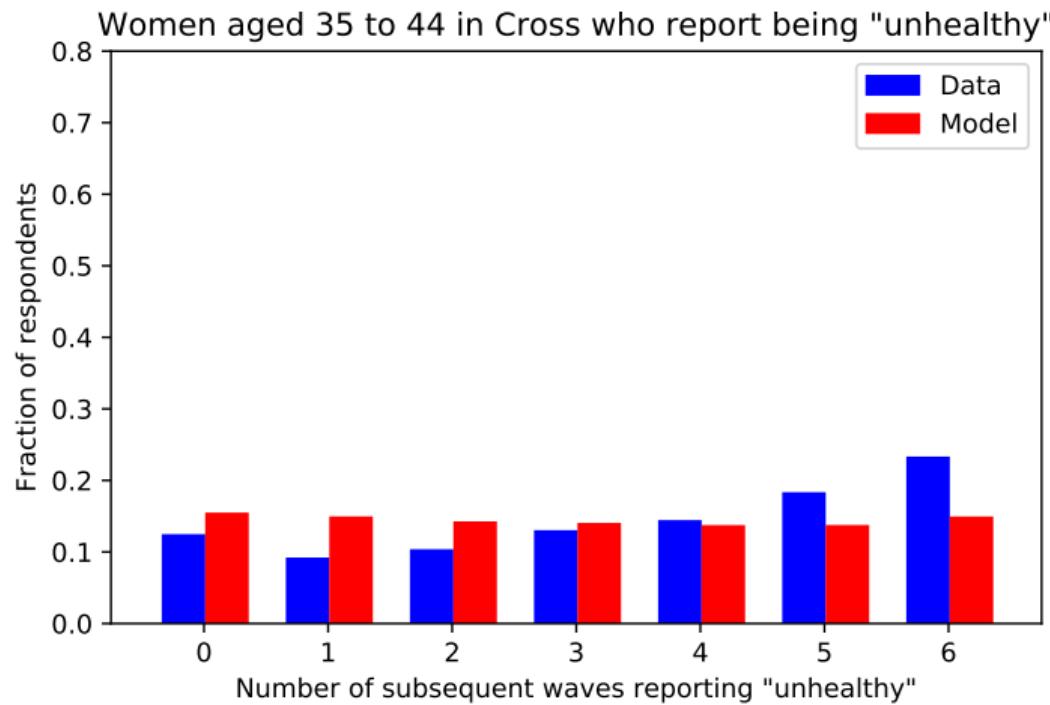
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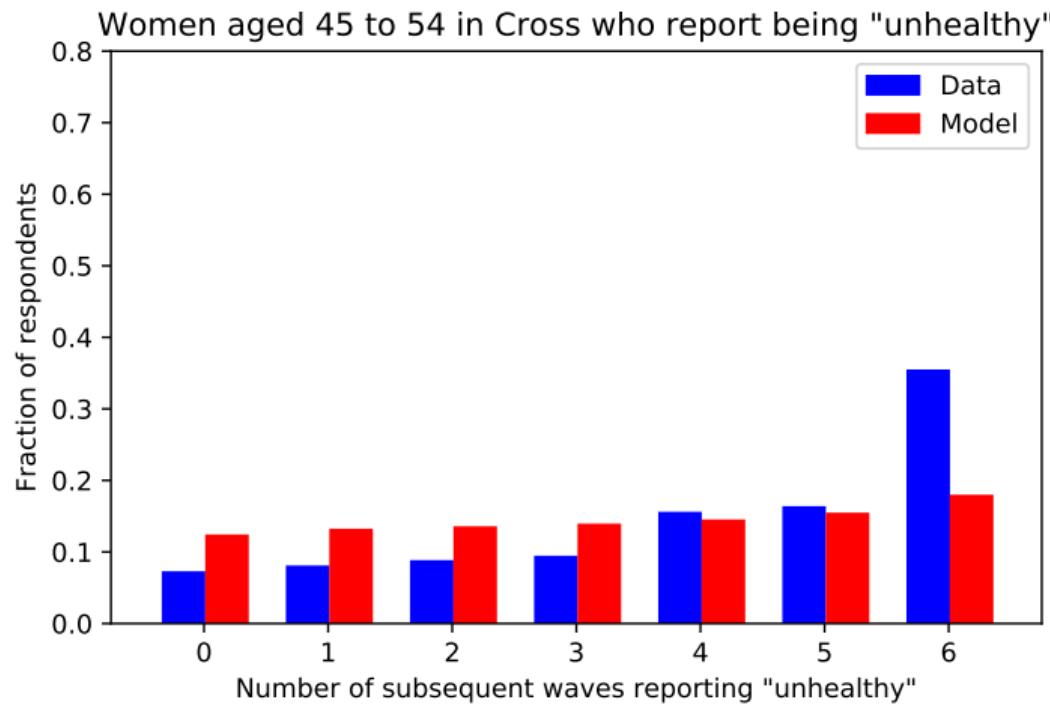
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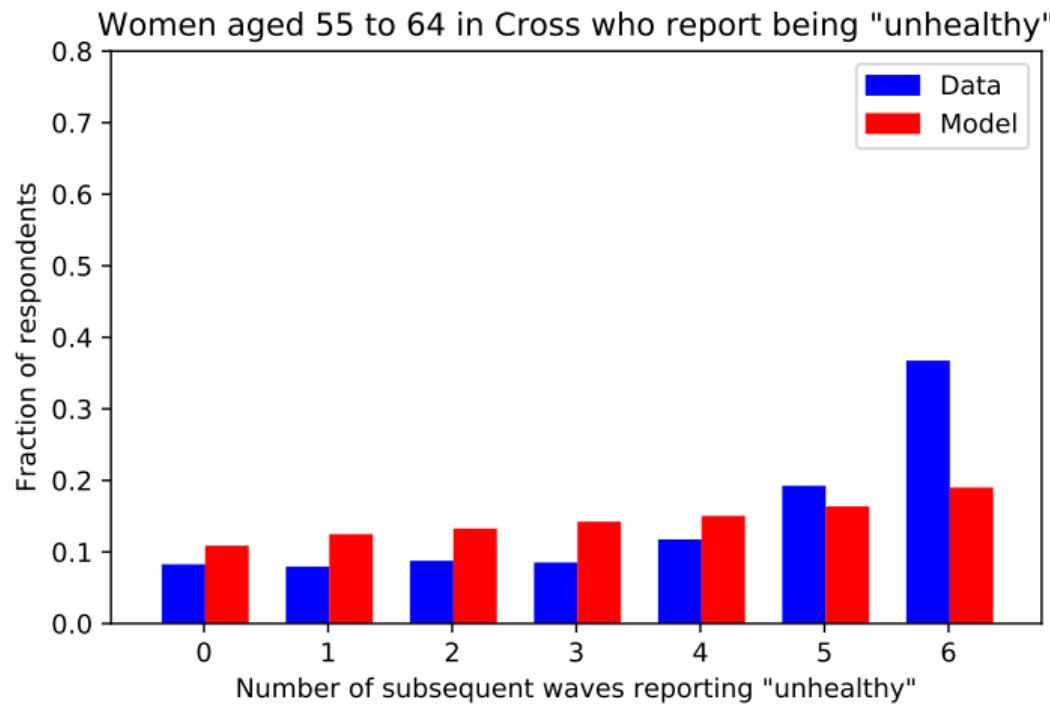
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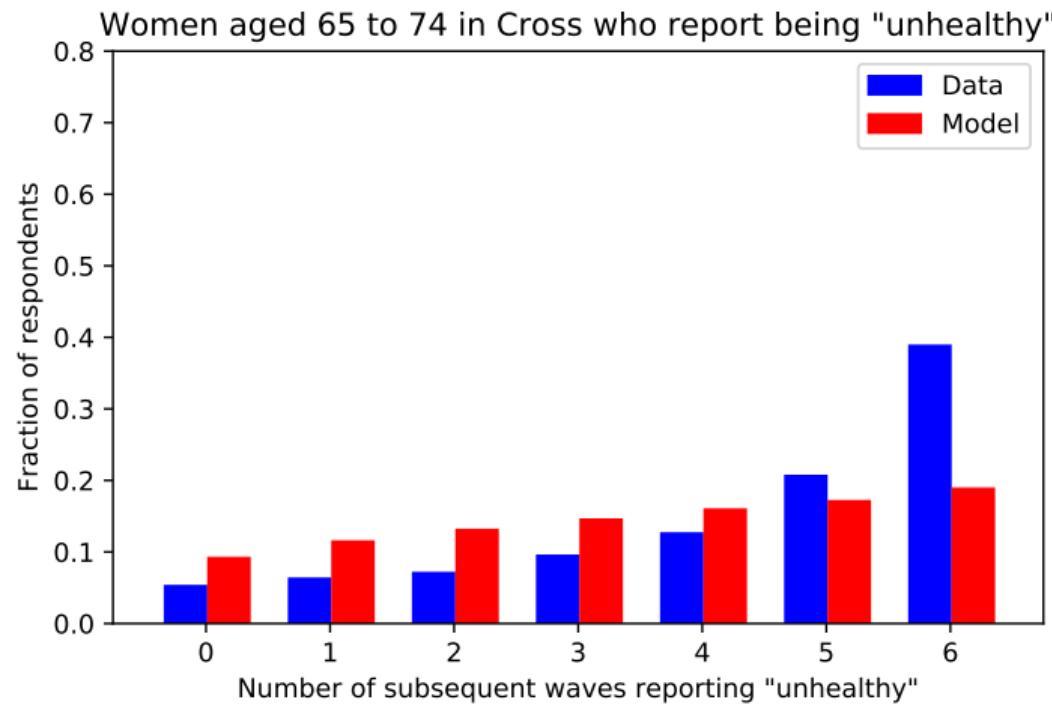
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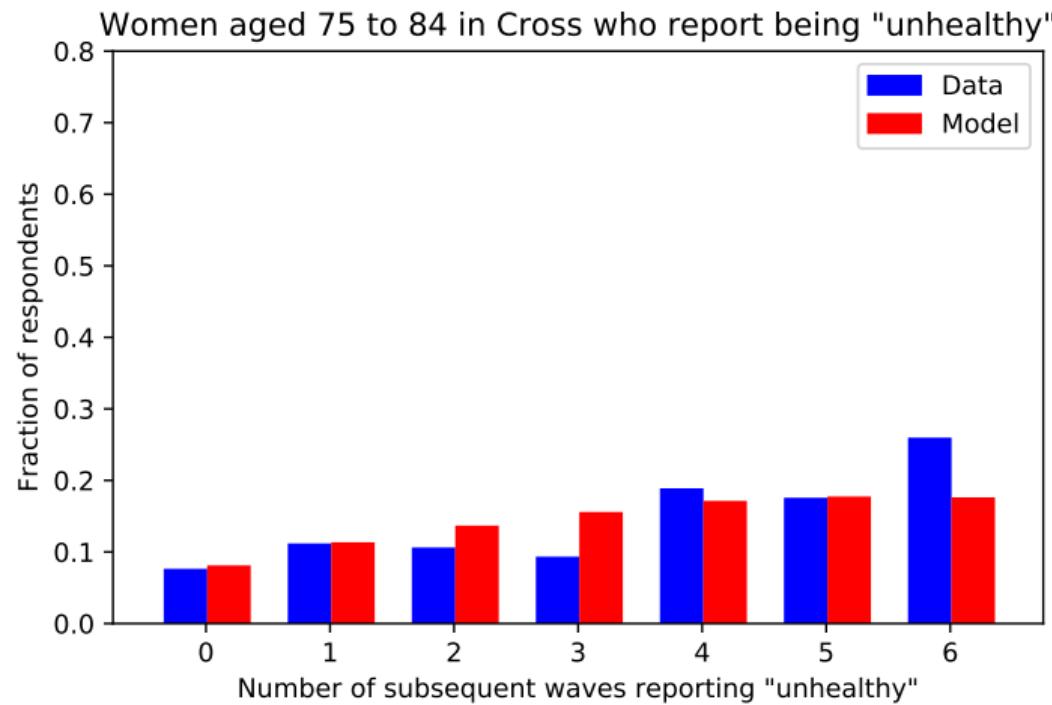
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# EXTENSIONS

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- No: Similar results with SRHS or constructed index from multiple measures. Self-justification bias is offset by attenuation bias!

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- Extent to which pain interferes with normal work activities

# Other measures of health

- Limitations on work, housework, or schoolwork due to health?
- Activities of daily living: climbing stairs, carrying grocery bag
- Mental health: How often feel nervous, hopeless, worthless, etc?
- Extent to which pain interferes with normal work activities
- Clinical measures: hand grip strength, walking speed, lung capacity

# Latent health model with single measure (SRHS)

Baseline latent health model:

$$x_0 \sim N(\mu_0, \sigma_0^2). \quad (1)$$

$$\text{Prob}(h_{t+1} = 0) = 1 - \Phi(f(j_t, x_t)). \quad (2)$$

$$x_{t+1} = \rho_j x_t + (1 - \rho_j)g(j_t) + \epsilon_{t+1}, \quad \epsilon_{t+1} \sim \Phi. \quad (3)$$

$$h_t^* = \alpha_0 + \alpha_1 x_t + \eta_t, \quad \eta_t \sim \Phi, \quad (4)$$

$$h_t = 1 + \sum_{k=1}^{K-1} \mathbf{1}(h_t^* \geq \chi_k).$$

# Latent health model with multiple measures (1/2)

Latent health model with multiple independent measures:

$$x_0 \sim N(\mu_0, \sigma_0^2). \quad (1)$$

$$\text{Prob}(h_{t+1} = 0) = 1 - \Phi(f(j_t, x_t)). \quad (2)$$

$$x_{t+1} = \rho_j x_t + (1 - \rho_j)g(j_t) + \epsilon_{t+1}, \quad \epsilon_{t+1} \sim \Phi. \quad (3)$$

$$m_{\ell t}^* = \alpha_{\ell 0} + \alpha_{\ell 1} x_t + \eta_{\ell t}, \quad \eta_{\ell t} \sim \Phi, \quad (4)$$

$$m_{\ell t} = 1 + \sum_{k=1}^{K_\ell-1} \mathbf{1}(m_t^* \geq \chi_{k\ell}), \quad \ell \in \{0, \dots, L\}, \quad h_t \equiv m_{0t}.$$

## Latent health model with multiple measures (2/2)

- Each wave, observe multiple categorical measures of latent health
- Measure 0 is set aside for SRHS (and mortality)
- Location assumption:  $\chi_{1\ell} = 0 \forall \ell$ , but only  $\alpha_{00} = 0$

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- My MLE code is already set up to estimate such a model, and appropriate datasets have been constructed...
- ...but I don't have a specification or results I'd like to share today

## Adding other model extensions

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- And can test whether normality of latent health shocks is well supported
- Nothing in core MLE code has to change to add these features!
- Label heterogeneity:** Discretized  $x_t$  vector would include “type”; change precomputed report probability matrix, tile transition probabilities
- Skewed health shocks:** Just change precomputed transition probability array for  $x_t$ , probably use mixed normal (impose mean 0, stdev 1)

# Latent health model: Adding SRHS labeling heterogeneity

Latent health model with “labeling shifter” for SRHS:

$$x_0 \sim N(\mu_0, \sigma_0^2), \quad \delta \sim \text{Discrete}. \quad (1)$$

$$\text{Prob}(h_{t+1} = 0) = 1 - \Phi(f(j_t, x_t)). \quad (2)$$

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# CONCLUSION

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- All hail Moore's law!