

*Day-1 Session-2:*  
*Bipartite Incidence Graph*  
*for Adaptive Cluster Sampling*

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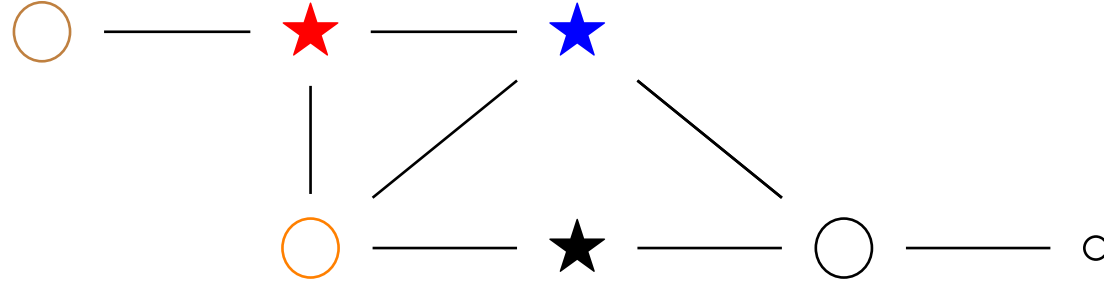
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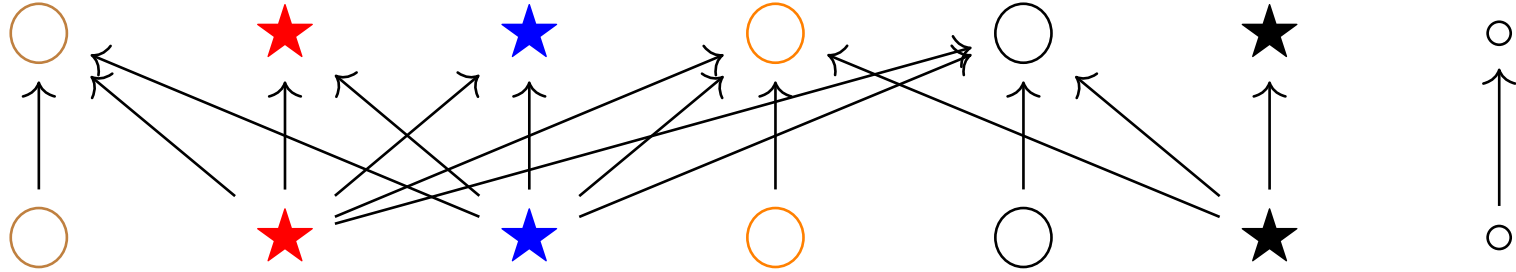
## Bipartite incidence graph (BIG)

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ACS from  $G = (U, A)$ , a part of which as given below:



All the observational links in  $\mathcal{B} = (F, \Omega; H)$  below:



$\mathcal{B}$  has simple directed edge set  $H : F \rightarrow \Omega$ , where  $F$  contains *sampling units* and  $\Omega$  the *study units*, and

$$(i\kappa) \in H \quad \text{only if} \quad \Pr(\kappa \in \Omega_s | i \in s_0) = 1$$

## Multiplicity/ancestry under BIG sampling (BIGS)

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The sampling units that can lead to a given  $\kappa \in \Omega$  are

$$\beta_\kappa = \{i \in F : (i\kappa) \in H\}$$

the *ancestors* of  $\kappa$  under BIGS (Zhang, 2021; Zhang and Patone, 2017), similar to *multiplicity* (Birnbaum and Sirken, 1965) under indirect sampling.

*Problem: ancestry knowledge  $\beta_\kappa$  for BIGS from  $\mathcal{B}$  not always guaranteed under original sampling from  $G$*

For any case node  $\kappa \in \Omega$ , ACS from  $G$  yields all its case network nodes  $\beta_\kappa$  which are also its ancestors under BIGS.

For any noncase non-edge node  $\kappa$ , we have  $\beta_\kappa = \{i_\kappa\}$  under ACS from  $G$ , which is always observed if  $\kappa$  is observed.

For any edge node  $\kappa$ , like , ACS from  $G$  does not always yield  $\beta_\kappa$  which includes *all* its adjacent networks.

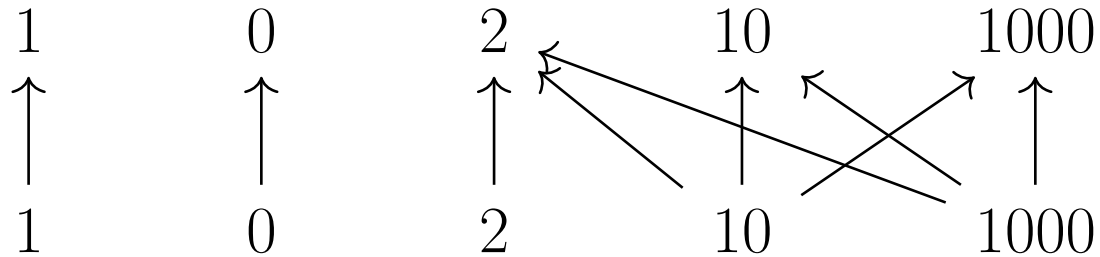
Thompson (1990): initial SRS and  $|s_0| = 2$

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A spatial graph  $G$  for ACS, with threshold  $y_i \geq 5$ :

1 ——— 0 ——— 2 ——— 10 ——— 1000

$\mathcal{B}$  including all the observation links to edge node 2:



Modified HTE:  $\hat{\theta}_{HT}^* = \sum_{\kappa \in \Omega_s} W_{\kappa} y_{\kappa}$   
 with  $W_{\kappa} = \pi_{(\kappa)}^{-1}$  except for any *terminal* node where

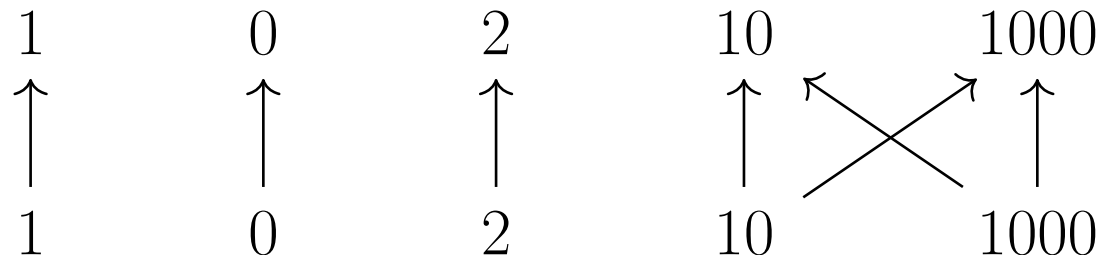
$$W_{\kappa} = \begin{cases} \Pr(i_{\kappa} \in s_0)^{-1} & \text{if } i_{\kappa} \in s_0 \\ 0 & \text{otherwise} \end{cases}$$

such that  $E(\mathbb{I}(\kappa \in \Omega_s) W_{\kappa}) \equiv 1$  for any  $\kappa \in \Omega = U$

# Modified BIG (Zhang and Oguz-Alper, 2020)

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Remove corresponding links to 2 in  $\mathcal{B} \Rightarrow$  modified  $\mathcal{B}^*$ :



Under BIGS from  $\mathcal{B}^*$ , edge node 2 observed iff  $2 \in s_0$

$$\text{HTE: } \hat{\theta}_{HT} = \sum_{\kappa \in \Omega_s} y_{\kappa} / \pi_{(\kappa)}$$

where  $\pi_{(\kappa)}$  refers to  $\Pr(\kappa \in \Omega_s)$  under BIGS from  $\mathcal{B}^*$

Strategies ( $\mathcal{B}$ , MHT) and ( $\mathcal{B}^*$ , HT) yield always the same estimate but differ w.r.t. Rao-Blackwell (RB) method  
 NB. sampling strategy = (sampling method, estimator)

# Numerical results for $\mu = \theta/N$

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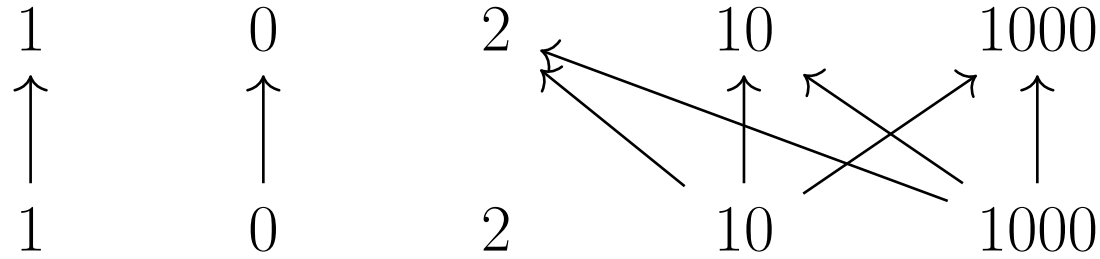
	1 — 0 — 2 — 10 — 1000			
	$(\mathcal{B}, \text{MHT})$		$(\mathcal{B}^*, \text{HT})$	
$s_0$	$\Omega_s = s$	$\hat{\mu}_{HT}^*$	$\Omega_s$	$\hat{\mu}_{HT}$
1,0	1,0	0.500	1,0	0.500
1,2	1,2	1.500	1,2	1.500
0,2	0,2	1.000	0,2	1.000
1,10	1,10,2,1000	289.071	1,10,1000	289.071
1,1000	1,1000,2,10	289.071	1,1000,10	289.071
0,10	0,10,2,1000	288.571	0,10,1000	288.571
0,1000	0,1000,2,10	288.571	0,1000,10	288.571
2,10	2,10,1000	289.571	2,10,1000	289.571
2,1000	2,1000,10	289.571	2,1000,10	289.571
10,1000	10,1000,2	288.571	10,1000	288.571
Variance		17418.4		17418.4

Strategy  $(\mathcal{B}, \text{MHT})$ : the last three samples are all  $\Omega_s = \{2, 10, 1000\}$ , but  $\hat{\mu}_{HT}^*$  differs because 2 is unused when  $s_0 = \{10, 1000\}$ . The RB method yields  $E[\hat{\mu}_{HT}^* | \Omega_s = \{2, 10, 1000\}] = 289.238$ .

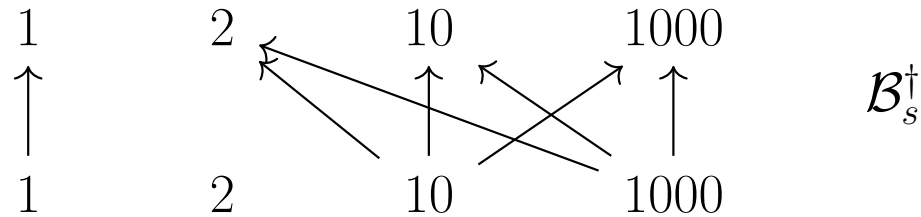
# Sample-dependent BIGS

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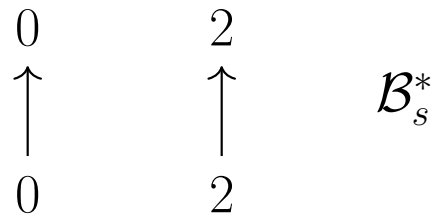
Modified  $\mathcal{B}^\dagger$ , where 2 is only observed from  $\{10, 1000\}$ :



$\mathcal{B}^\dagger$  identified with  $\Pr(s_0 \cap \{10, 1000\} \neq \emptyset) = 0.7$ , e.g.  $s_0 = \{1, 10\}$



$\mathcal{B}^*$  needed with  $\Pr(s_0 \cap \{2, 10, 1000\} = \{2\}) = 0.2$ , e.g.  $s_0 = \{0, 2\}$



# Sample-dependent BIGS

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1 ——— 0 ——— 2 ——— 10 ——— 1000

	$(\mathcal{B}^*, \text{HT})$		$(\mathcal{B}^\dagger, \text{HT})$	
$s_0$	$\Omega_s$	$\hat{\mu}_{HT}$	$\Omega_s$	$\hat{\mu}_{HT}$
1,0	1,0	0.500	1,0	0.500
1,2	1,2	1.500	1	0.500
0,2	0,2	1.000	0	0.000
1,10	1,10,1000	289.071	1,10,2,1000	289.643
1,1000	1,1000,10	289.071	1,1000,2,10	289.643
0,10	0,10,1000	288.571	0,10,2,1000	289.143
0,1000	0,1000,10	288.571	0,1000,2,10	289.143
2,10	2,10,1000	289.571	2,10,1000	289.143
2,1000	2,1000,10	289.571	2,1000,10	289.143
10,1000	10,1000	288.571	10,1000,2	289.143
Variance		17418.4		17533.7

Repeated ACS from  $G$ : indifferent choice if  $s_0 = \{1, 0\}$ , more repetitions ‘needed’ to decide;  
adopt  $\mathcal{B}^*$  if 2 first sampled from  $s_0 = \{1, 2\}$  or  $\{0, 2\}$ , or  $\mathcal{B}^\dagger$  otherwise



## Sample-dependent BIGS

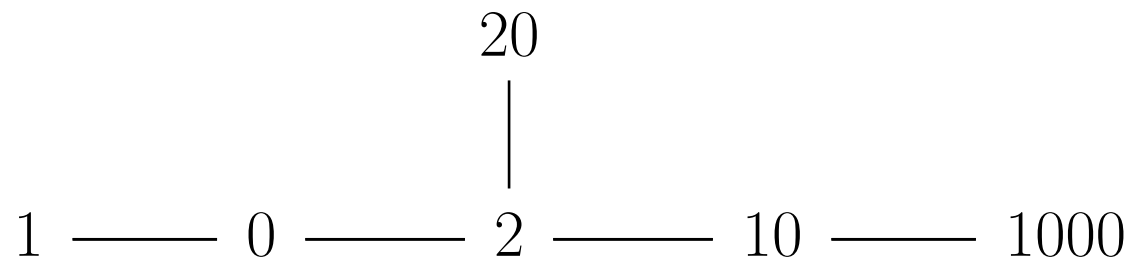
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Odds of adopting  $\mathcal{B}^\dagger$  or  $\mathcal{B}^*$  for edge node 2 is 7 : 2

Unconditional inference impossible: ‘first sample’ decides

$\Downarrow$

Conditional inference w.r.t. the adopted strategy

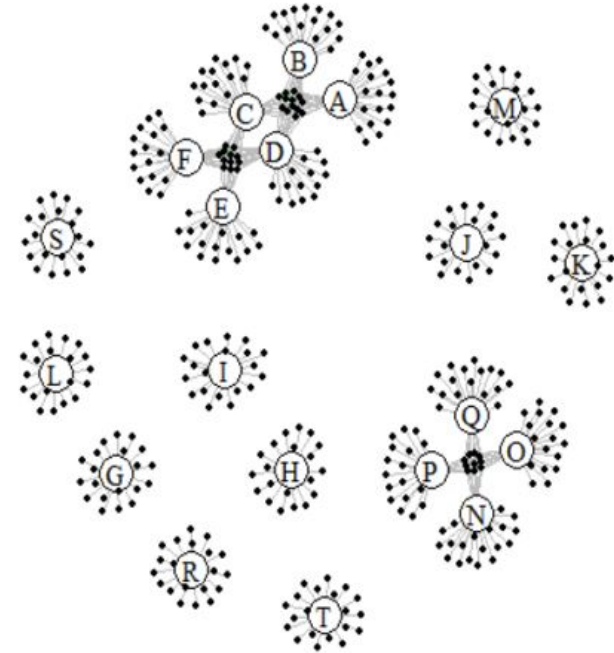
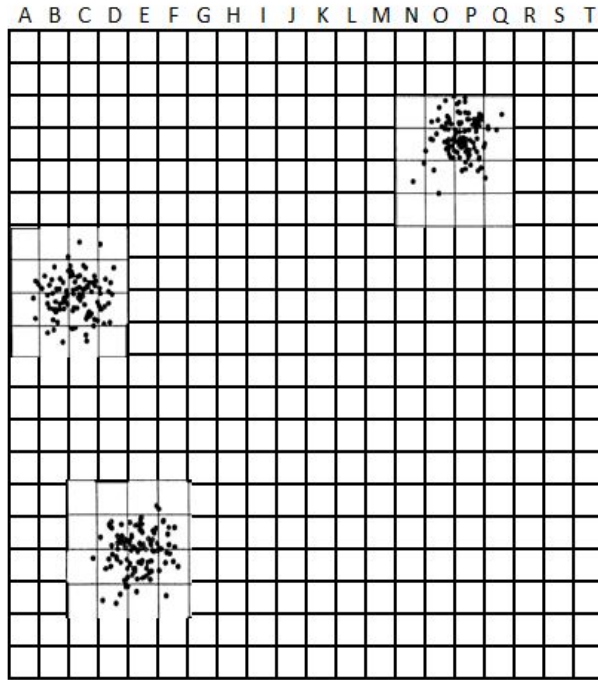


SRS,  $|s_0| = 2$ : probability  $\frac{2}{15}$  for observing both networks  $\{20\}$  and  $\{10, 1000\}$ , in which case one can e.g. let  $\beta'_2 = \{20\}$  or  $\beta'_2 = \{10, 1000\}$  or  $\beta'_2 = \{20, 10, 1000\}$

Under ACS:  $\{20\}$  and  $\{10, 1000\}$  cannot be observed *from* each other

# An example of two-stage ACS (Thompson, 1991)

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1st-stage: 20 strips; 2nd-stage: neighbouring grids to any nonempty one, and so on  
Thus, ACS applied at the second stage and terminated if no more non-empty grids  
An edge grid is an empty grid that is contiguous to one or more non-empty grids  
Modified  $\mathcal{B}^*$  (Zhang and Oguz-Alper, 2020):  $F$  = strips,  $\Omega$  = grids  
10 star-like subgraphs, where an empty strip is adjacent to its 20 empty grids;  
three clusters of non-empty grids; rest empty grids to 10 non-empty strips  
An edge grid non-adjacent to neighbour strip, removing link under two-stage ACS

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