PROBABILITY THEORY AND RANDOM PROCESSES (MA225)

 $\begin{array}{c} {\rm LECTURE~SLIDES} \\ {\rm Lecture~09~(August~19,~2019)} \end{array}$

Expectation of Function of RV

Example 1: Let the random variable X be a DRV with PMF

$$f_X(x) = \begin{cases} \frac{1}{7} & \text{if } x = -2, -1, 0, 1\\ \frac{3}{14} & \text{if } x = 2, 3\\ 0 & \text{otherwise.} \end{cases}$$

Let $Y = X^2$. Find the expectation of Y.

Expectation of Function of RV

Theorem: Let X be a DRV with PMF $f_X(\cdot)$ and support S_X . Let $g: \mathbb{R} \to \mathbb{R}$. Then

$$E\left[g(X)
ight] = \sum_{x \in S_X} g(x) f_X(x) \quad ext{provided } \sum_{x \in S_X} |g(x)| f_X(x) < \infty.$$

Theorem: Let X be a CRV with PDF $f_X(\cdot)$. Let $g: \mathbb{R} \to \mathbb{R}$. Then

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx$$
 provided $\int_{-\infty}^{\infty} |g(x)| f_X(x) dx < \infty$.

Expectation of Function of RV

Theorem: Let X be a RV (either DRV or CRV). Then

- ① Let $A \subset \mathbb{R}$. Then $E(I_A(X)) = P(X \in A)$.
- ② $h_1(x) \le h_2(x)$, for all $x \in \mathbb{R}$, then $E[h_1(X)] \le E[h_2(X)]$, provided all the expectations exist.
- ③ a < b are two real numbers such that $S_X \subset [a, b]$, then $a \le E(X) \le b$, provided the expectation exists.
- ⑤ Let $h_1(\cdot), \ldots, h_p(\cdot)$ be real valued function of real numbers such that $E(h_i(X))$ exists for all $i=1,2,\ldots,p$, then

$$E\left(\sum_{i=1}^p h_i(X)\right) = \sum_{i=1}^p E\left(h_i(X)\right).$$

Remarks

- For $r = 1, 2, ..., \mu_r = E(X^r)$ is called rth raw moment of X, if the expectation exists.
- $\mu'_r = E[(X E(X))^r]$ is called rth central moment of X, if the expectations exist.
- $\mu'_2 = E\left[\left(X E(X)\right)^2\right]$ is called variance of X when it exists and is denoted by Var(X).
- $Var(X) = E(X^2) (E(X))^2$.

Moment Generating Function

Def: The moment generating function of random variable X is defined by

$$M_X(t) = E\left(e^{tX}\right)$$

provided the expectation exists in a neighbourhood of the origin.

Example 2: $X \sim Bin(n, p)$, then $M_X(t) = (1 - p + pe^t)^n$ for all $t \in \mathbb{R}$.

Example 3:
$$X \sim Exp(\lambda)$$
, then $M_X(t) = \left(1 - \frac{t}{\lambda}\right)^{-1}$ for all $t < \lambda$.

Example 4: $X \sim N(\mu, \sigma^2)$, then $M_X(t) = e^{\mu t + \frac{t^2 \sigma^2}{2}}$ for all $t \in \mathbb{R}$.

Def: X and Y are said to be same in distribution if $F_X(x) = F_Y(x)$ for all $x \in \mathbb{R}$.

Theorem: Let X and Y be two random variables having MGFs $M_X(\cdot)$ and $M_Y(\cdot)$, respectively. Suppose that there exists a positive real number a such that $M_X(t) = M_Y(t)$ for all $t \in (-a, a)$. Then X and Y are same in distribution.

Example 5: Let $X \sim N(\mu, \sigma^2)$. Find the distribution of Y = a + bX.

Remark: If the MGF $M_X(t)$ exist for $t \in (-a, a)$ for some a > 0, the derivatives of all order exist at t = 0 and

$$E\left(X^{k}\right) = \left.\frac{d^{k}}{dt^{k}}M_{X}(t)\right|_{t=0}$$

for all positive integer k.