Introduction to Econometrics

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Preface

1 Nature and scope of Econometrics

2 Introductory Econometrics

2.1 Install packages

```
library(wooldridge)
library(readr)
library(stargazer)
library(kableExtra)
library(quantmod)
library(xts)
```

2.2 Useful dataset

```
earns <- read.csv("data/earns.csv")
gpa1 <- read.csv("data/gpa1.csv")
hprice1 <- read.csv("data/hprice1.csv")
hprice2 <- read.csv("data/hprice2.csv")
hprice3 <- read.csv("data/hprice3.csv")
jtrain <- read.csv("data/jtrain.csv")
nyse <- read.csv("data/nyse.csv")
phillips <- read.csv("data/phillips.csv")
rdchem <- read.csv("data/rdchem.csv")
traffic1 <- read.csv("data/traffic1.csv")
wage1 <- read.csv("data/wage1.csv")</pre>
```

2.3 Simple regression model

2.3.1 Example 1: A log wage equation

• Load the wage1 data and check out the documentation.

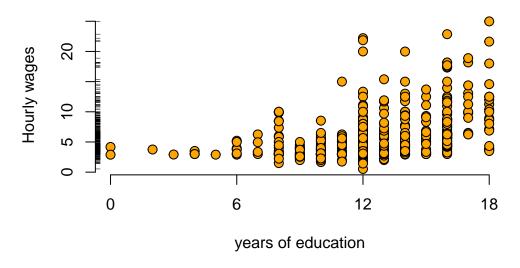
educ: years of education

wage: average hourly earnings

lwage: log of the average hourly earnings

• First, make a scatter-plot of the two variables and look for possible patterns in the relationship between them.

Wages vs. Education, 1976



- 1. It appears that *on average*, more years of education, leads to higher wages.
- 2. The example in the text is interested in the return to another year of education, or what the **percentage** change in wages one might expect for each additional year of education. To do so, one must use the log(wage). This has already been computed in the data set and is defined as lwage.
- Build a linear model to estimate the relationship between the log of wage (lwage) and education (educ).

$$\widehat{log(wage)} = \beta_0 + \beta_1 educ$$

```
log_wage_model <- lm(lwage ~ educ, data = wage1)</pre>
  • Print the summary of the results.
  summary(log_wage_model)
Call:
lm(formula = lwage ~ educ, data = wage1)
Residuals:
     Min
               1Q
                    Median
                                  3Q
                                          Max
-2.21158 -0.36393 -0.07263 0.29712 1.52339
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.583773
                       0.097336
                                  5.998 3.74e-09 ***
educ
            0.082744
                       0.007567 10.935 < 2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.4801 on 524 degrees of freedom
Multiple R-squared: 0.1858,
                                 Adjusted R-squared: 0.1843
F-statistic: 119.6 on 1 and 524 DF, p-value: < 2.2e-16
  • Use the stargazer package to make beautiful table
  stargazer(type = "html", log_wage_model, single.row = TRUE, header = FALSE, digits = 3)
Dependent variable:
lwage
educ
0.083***(0.008)
Constant
0.584***(0.097)
Observations
```

526

```
R2
```

0.186

Adjusted R2

0.184

Residual Std. Error

```
0.480 (df = 524)
```

F Statistic

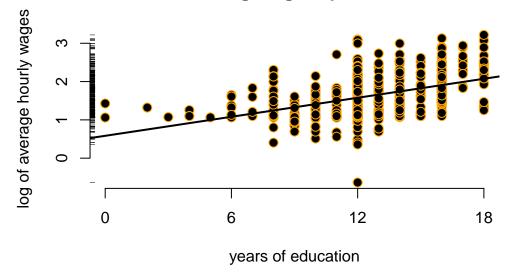
$$119.582*** (df = 1; 524)$$

Note:

p < 0.1; p < 0.05; p < 0.01

• Plot the log(wage) vs educ. The blue line represents the least squares fit.

A Log Wage Equation



2.4 Multiple regression analysis

2.4.1 Example 2: Hourly wage equation

Check the documentation for variable information

lwage: log of the average hourly earnings

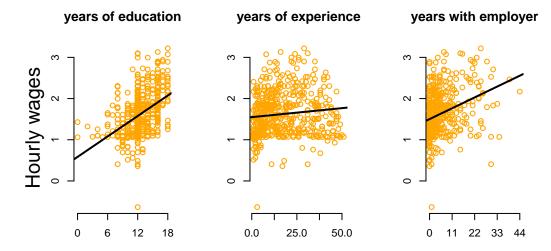
educ: years of education

exper: years of potential experience

tenutre: years with current employer

• Plot the variables against lwage and compare their distributions and slope (β) of the simple regression lines

```
par(mfrow=c(1,3))
plot(y = wage1$lwage, x = wage1$educ, col="orange", xaxt="n", frame = FALSE, main = "years
mtext(side=2, line=2.5, "Hourly wages", cex=1.25)
axis(side = 1, at = c(0,6,12,18))
abline(lm(lwage ~ educ, data=wage1), col = "black", lwd=2)
plot(y = wage1$lwage, x = wage1$exper, col="orange", xaxt="n", frame = FALSE, main = "year
axis(side = 1, at = c(0,12.5,25,37.5,50))
abline(lm(lwage ~ exper, data=wage1), col = "black", lwd=2)
plot(y = wage1$lwage, x = wage1$tenure, col="orange", xaxt="n", frame = FALSE, main = "year
axis(side = 1, at = c(0,11,22,33,44))
abline(lm(lwage ~ tenure, data=wage1), col = "black", lwd=2)
```



• Estimate the model regressing educ, exper, and tenure against log(wage).

$$log(wage) = \beta_0 + \beta_1 educ + \beta_3 exper + \beta_4 tenure$$

```
hourly_wage_model <- lm(lwage ~ educ + exper + tenure, data = wage1)</pre>
```

• Print the estimated model coefficients:

```
coefficients(hourly_wage_model)
```

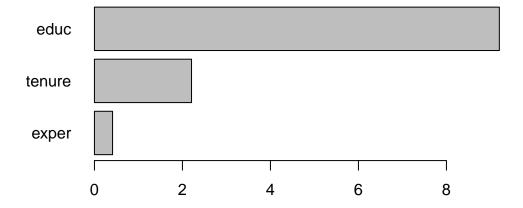
```
kable(coefficients(hourly_wage_model), digits=4, col.names = "Coefficients", align = '1')
```

	Coefficients
(Intercept)	0.2844
educ	0.0920
exper	0.0041
tenure	0.0221

• Plot the coefficients, representing percentage impact of each variable on log(wage) for a quick comparison.

```
barplot(sort(100*hourly_wage_model$coefficients[-1]), horiz=TRUE, las=1,
    ylab = " ", main = "Coefficients of Hourly Wage Equation")
```

Coefficients of Hourly Wage Equation



2.5 Multiple regression analysis: inference

2.5.1 Example 3: Hourly Wage Equation

Using the same model estimated in **example 3**, examine and compare the standard errors associated with each coefficient. Like the textbook, these are contained in parenthesis next to each associated coefficient.

```
summary(hourly wage model)
Call:
lm(formula = lwage ~ educ + exper + tenure, data = wage1)
Residuals:
               1Q
                    Median
                                 3Q
                                         Max
-2.05802 -0.29645 -0.03265 0.28788
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.284360
                       0.104190
                                  2.729 0.00656 **
educ
            0.092029
                       0.007330 12.555 < 2e-16 ***
            0.004121
                       0.001723
                                  2.391 0.01714 *
exper
                       0.003094
                                  7.133 3.29e-12 ***
tenure
            0.022067
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.4409 on 522 degrees of freedom
Multiple R-squared: 0.316, Adjusted R-squared: 0.3121
F-statistic: 80.39 on 3 and 522 DF, p-value: < 2.2e-16
  stargazer(type = "html", hourly_wage_model, single.row = TRUE, header = FALSE, digits=5)
Dependent variable:
lwage
educ
0.09203****(0.00733)
exper
```

0.00412**(0.00172)

tenure

0.02207***(0.00309)

Constant

0.28436****(0.10419)

Observations

526

R.2

0.31601

Adjusted R2

0.31208

Residual Std. Error

0.44086 (df = 522)

F Statistic

80.39092*** (df = 3; 522)

Note:

p < 0.1; p < 0.05; p < 0.01

For the years of experience variable, or exper, use coefficient and Standard Error to compute the t statistic:

$$t_{exper} = \frac{0.004121}{0.001723} = 2.391$$

Fortunately, R includes t statistics in the summary of model diagnostics.

summary(hourly_wage_model)\$coefficients

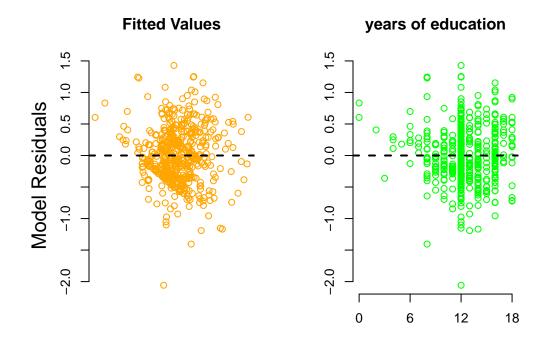
Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.284359541 0.104190379 2.729230 6.562466e-03
educ 0.092028988 0.007329923 12.555246 8.824197e-32
exper 0.004121109 0.001723277 2.391437 1.713562e-02
tenure 0.022067218 0.003093649 7.133070 3.294407e-12

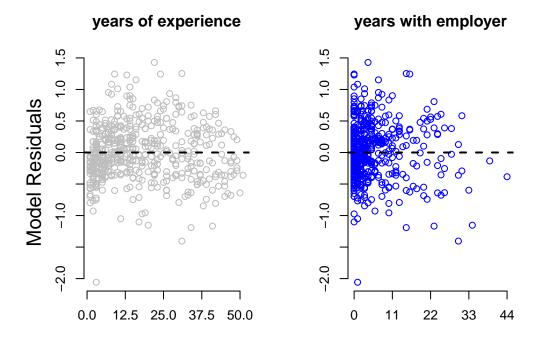
kable(summary(hourly_wage_model)\$coefficients, align="1", digits=3)

	Estimate	Std. Error	t value	$\Pr(> t)$
(Intercept)	0.284	0.104	2.729	0.007
educ	0.092	0.007	12.555	0.000
exper	0.004	0.002	2.391	0.017
tenure	0.022	0.003	7.133	0.000

• lets plot this results

```
par(mfrow=c(2,2))
plot(y = hourly_wage_model$residuals, x = hourly_wage_model$fitted.values , col="orange",
     frame = FALSE, main = "Fitted Values", xlab = "", ylab = "")
mtext(side=2, line=2.5, "Model Residuals", cex=1.25)
abline(0, 0, col = "black", lty=2, lwd=2)
plot(y = hourly_wage_model$residuals, x = wage1$educ, col="green", xaxt="n",
     frame = FALSE, main = "years of education", xlab = "", ylab = "")
axis(side = 1, at = c(0,6,12,18))
abline(0, 0, col = "black", lty=2, lwd=2)
plot(y = hourly_wage model$residuals, x = wage1$exper, col="gray", xaxt="n",
     frame = FALSE, main = "years of experience", xlab = "", ylab = "")
mtext(side=2, line=2.5, "Model Residuals", cex=1.25)
axis(side = 1, at = c(0,12.5,25,37.5,50))
abline(0, 0, col = "black", lty=2, lwd=2)
plot(y = hourly_wage_model$residuals, x = wage1$tenure, col="blue", xaxt="n",
     frame = FALSE, main = "years with employer", xlab = "", ylab = "")
axis(side = 1, at = c(0,11,22,33,44))
abline(0, 0, col = "black", lty=2, lwd=2)
```

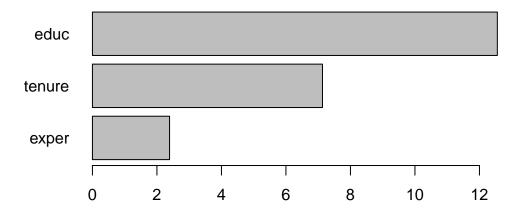




 \bullet Plot the t statistics for a visual comparison:

```
barplot(sort(summary(hourly_wage_model)$coefficients[-1, "t value"]), horiz=TRUE, las=1,
    ylab = " ", main = "t statistics of Hourly Wage Equation")
```

t statistics of Hourly Wage Equation



2.5.2 Example 4: Effect of Job Training on Firm Scrap Rates

• Load the jtrain data set. (From H. Holzer, R. Block, M. Cheatham, and J. Knott (1993), Are Training Subsidies Effective? The Michigan Experience, Industrial and Labor Relations Review 46, 625-636. The authors kindly provided the data.)

year: 1987, 1988, or 1989

union := 1 if unionized

lscrap : Log(scrap rate per 100 items)

hrsemp: (total hours training) / (total employees trained)

lsales: Log(annual sales, \$)

lemploy: Log(umber of employees at plant)

- First, use the **subset** function and it's argument by the same name to return observations which occurred in **1987** and are not **union**.
- At the same time, use the **select** argument to return only the variables of interest for this problem.

```
jtrain_subset <- subset(jtrain, subset = (year == 1987 & union == 0), select = c(year, uni
```

• Next, test for missing values. One can "eyeball" these with R Studio's View function, but a more precise approach combines the sum and is.na functions to return the total number of observations equal to NA.

```
sum(is.na(jtrain_subset))
```

[1] 156

• While R's 1m function will automatically remove missing NA values, eliminating these manually will produce more clearly proportioned graphs for exploratory analysis. Call the na.omit function to remove all missing values and assign the new data.frame object the name jtrain_clean.

```
jtrain_clean <- na.omit(jtrain_subset)</pre>
```

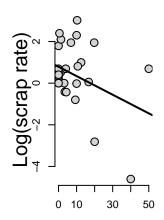
• We use jtrain_clean to plot the variables of interest against 1scrap. Visually observe the respective distributions for each variable, and compare the slope (β) of the simple regression lines.

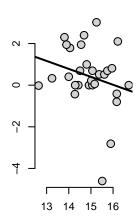
```
par(mfrow=c(1,3))
point_size <- 1.60
plot(y = jtrain_clean$lscrap, x = jtrain_clean$hrsemp, frame = FALSE,
main = "Total (hours/employees) trained", ylab = "", xlab="", pch = 21, bg = "lightgrey",
mtext(side=2, line=2, "Log(scrap rate)", cex=1.25)
abline(lm(lscrap ~ hrsemp, data=jtrain_clean), col = "black", lwd=2)
plot(y = jtrain_clean$lscrap, x = jtrain_clean$lsales, frame = FALSE, main = "Log(annual sabline(lm(lscrap ~ lsales, data=jtrain_clean), col = "black", lwd=2)
plot(y = jtrain_clean$lscrap, x = jtrain_clean$lemploy, frame = FALSE, main = "Log(# employabline(lm(lscrap ~ lemploy, data=jtrain_clean), col = "black", lwd=2)</pre>
```

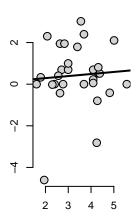
Total (hours/employees) trai

Log(annual sales \$)

Log(# employees at plant







• Now create the linear model regressing hrsemp(total hours training/total employees trained), lsales(log of annual sales), and lemploy(the log of the number of the employees), against lscrap(the log of the scrape rate).

$$lscrap = \alpha + \beta_1 hrsemp + \beta_2 lsales + \beta_3 lemploy$$

• Finally, print the complete summary diagnostics of the model.

summary(linear_model)

Call:

lm(formula = lscrap ~ hrsemp + lsales + lemploy, data = jtrain_clean)

Residuals:

Min 1Q Median 3Q Max -2.6301 -0.7523 -0.4016 0.8697 2.8273

Coefficients:

Estimate Std. Error t value Pr(>|t|) (Intercept) 12.45837 5.68677 2.191 0.0380 * -0.02927 hrsemp 0.02280 -1.2830.2111 lsales -0.96203 0.45252 0.0436 * -2.1260.76147 0.0734 . lemploy 0.40743 1.869

```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 1.376 on 25 degrees of freedom
Multiple R-squared: 0.2624, Adjusted R-squared: 0.1739
F-statistic: 2.965 on 3 and 25 DF, p-value: 0.05134
   • Use stargazer to create representative table
  stargazer(type = "html", linear_model, single.row = TRUE, header = FALSE, digits=5)
Dependent variable:
lscrap
hrsemp
-0.02927 (0.02280)
lsales
-0.96203** (0.45252)
lemploy
0.76147*(0.40743)
Constant
12.45837** (5.68677)
Observations
29
R2
0.26243
Adjusted R2
0.17392
Residual Std. Error
1.37604 (df = 25)
```

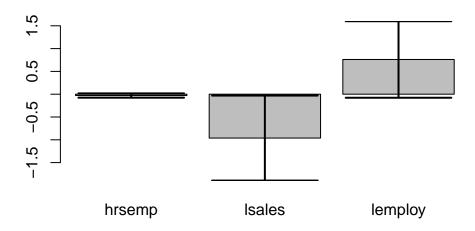
F Statistic

Note:

2.96504* (df = 3; 25)

p < 0.1; p < 0.05; p < 0.01

Coefficients & 95% C.I. of variables on Firm Scrap Rates



2.6 Chapter 5: Multiple Regression Analysis: OLS Asymptotics

2.6.1 Example: Housing Prices and Distance From an Incinerator

• We will use the hprice3 data set.

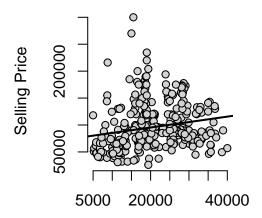
lprice : Log(selling price)

ldist : Log(distance from house to incinerator, feet)

larea: Log(square footage of house)

• Graph the prices of housing against distance from an incinerator:

```
par(mfrow=c(1,2))
plot(y = hprice3$price, x = hprice3$dist, main = " ", xlab = "Distance to Incinerator in f
abline(lm(price ~ dist, data=hprice3), col = "black", lwd=2)
```



Distance to Incinerator in feet

• Next, model the log(price) against the log(dist) to estimate the percentage relationship between the two.

$$price = \alpha + \beta_1 dist$$

```
price_dist_model <- lm(lprice ~ ldist, data = hprice3)</pre>
```

• Create another model that controls for "quality" variables, such as square footage area per house.

$$price = \alpha + \beta_1 dist + \beta_2 area$$

• Compare the coefficients of both models. Notice that adding area improves the quality of the model, but also reduces the coefficient size of dist.

```
summary(price_dist_model)
```

Call:

lm(formula = lprice ~ ldist, data = hprice3)

```
Residuals:
```

Min 1Q Median 3Q Max -1.22356 -0.28076 -0.05527 0.27992 1.29332

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 8.25750 0.47383 17.427 < 2e-16 ***
ldist 0.31722 0.04811 6.594 1.78e-10 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.4117 on 319 degrees of freedom Multiple R-squared: 0.1199, Adjusted R-squared: 0.1172 F-statistic: 43.48 on 1 and 319 DF, p-value: 1.779e-10

summary(price_area_model)

Call:

lm(formula = lprice ~ ldist + larea, data = hprice3)

Residuals:

Min 1Q Median 3Q Max -1.23380 -0.18820 -0.01723 0.21751 0.86039

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.49394 0.49065 7.121 7.18e-12 ***
ldist 0.19623 0.03816 5.142 4.77e-07 ***
larea 0.78368 0.05358 14.625 < 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3188 on 318 degrees of freedom Multiple R-squared: 0.4738, Adjusted R-squared: 0.4705 F-statistic: 143.2 on 2 and 318 DF, p-value: < 2.2e-16

• Use *stargazer* for better table

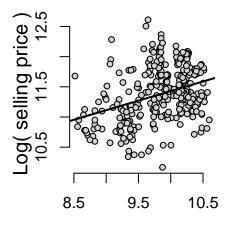
```
stargazer(type = "html",price_dist_model, price_area_model, single.row = TRUE, header = F
Dependent variable:
lprice
(1)
(2)
ldist
0.31722*** (0.04811)
0.19623*** (0.03816)
larea
0.78368*** (0.05358)
Constant
8.25750*** (0.47383)
3.49394*** (0.49065)
Observations
321
321
R2
0.11994
0.47385
Adjusted R2
0.11718
0.47054
Residual Std. Error
0.41170 (df = 319)
0.31883 \text{ (df} = 318)
F Statistic
43.47673*** (df = 1; 319)
143.19470^{***} (df = 2; 318)
```

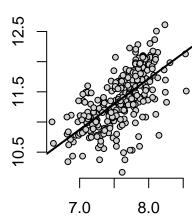
Note:

p<0.1; *p*<0.05; p<0.01

• Graphing illustrates the larger coefficient for area

Log(distance from incinerate Log(square footage of hous





3 Statistical Inference

4 The Simple Regression Model

5 Multiple Regression Analysis: Estimation

6 Multiple Regression Analysis: Inference

7 Multiple Regression Analysis: Further Topics

8 Asymptotic properties of OLS

9 Heteroskedasticity

10 Time Series data Modelling Dynamic Processes

11 Autocorrelation

12 Endogeneity – omitted variables, measurement error and simultaneity

13 Instrumental variable estimation and two stage least squares

14 Logit and Probit Models for Binary Response

15 Non Stationary Time Series

16 Simultaneous Equations Models

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