

# Cheatsheet Probability and Statistics

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## 1 Mathematical framework

### 1.1 Probability space

**Def. 1.1.** The set  $\Omega$  is called the **sample space**. An element  $\omega \in \Omega$  is called an **outcome** or **elementary experiment**.

**Ex. 1.1.** Throw of a die :  $\Omega = \{1, 2, 3, 4, 5, 6\}$

**Def. 1.2.** A **sigma-algebra** is a subset  $\mathcal{F} \subset \mathcal{P}(\Omega)$  satisfying the following properties :

**P1.**  $\Omega \in \mathcal{F}$

**P2.**  $A \in \mathcal{F} \Rightarrow A^c \in \mathcal{F}$  : If  $A$  is an event, “not  $A$ ” is also an event.

**P3.**  $A_1, A_2, \dots \in \mathcal{F} \Rightarrow \bigcup_{i=1}^{\infty} A_i \in \mathcal{F}$  : if  $A_1, A_2, \dots$  are events, then “ $A_1$  or  $A_2$  or ...” is an event

**Ex. 1.2.** Examples of sigma-algebras for  $\Omega = \{1, 2, 3, 4, 5, 6\}$  :

- $\mathcal{F} = \{\emptyset, \{1, 2, 3, 4, 5, 6\}\}$
- $\mathcal{F} = \mathcal{P}(\Omega)$
- $\mathcal{F} = \{\emptyset, \{1, 2\}, \{3, 4, 5, 6\}, \{1, 2, 3, 4, 5, 6\}\}$

Non examples of sigma-algebras for  $\Omega = \{1, 2, 3, 4, 5, 6\}$  :

- $\mathcal{F} = \{\{1, 2, 3, 4, 5, 6\}\}$  : **P2** is not satisfied
- $\mathcal{F} = \{\emptyset, \{1, 2, 3\}, \{4, 5, 6\}, \{1\}, \{2, 3, 4, 5, 6\}, \Omega\}$  : **P3** is not satisfied

**Def. 1.3.** Let  $\Omega$  a sample space and  $\mathcal{F}$  a sigma-algebra. A **probability measure** on  $(\Omega, \mathcal{F})$  is a map

$$\mathbb{P} : \mathcal{F} \rightarrow [0, 1], \quad A \mapsto \mathbb{P}[A]$$

that satisfies the properties

**P1.**  $\mathbb{P}[\Omega] = 1$

**P2. (countable additivity)**  $\mathbb{P}[A] = \sum_{i=1}^{\infty} \mathbb{P}[A_i]$  if  $A = \bigcup_{i=1}^{\infty} A_i$  (disjoint union)

**Int.** A probability measure is a map that associates to each event a number in  $[0, 1]$

**Ex. 1.3.** For  $\Omega = \{1, 2, 3, 4, 5, 6\}$  and  $\mathcal{F} = \mathcal{P}(\Omega)$ , the mapping  $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$  defined by

$$\forall A \in \mathcal{F} \quad \mathbb{P}[A] = \frac{|A|}{6}$$

is a probability measure on  $(\Omega, \mathcal{F})$ .

**Def. 1.4.** Let  $\Omega$  a sample space,  $\mathcal{F}$  a sigma-algebra and  $\mathbb{P}$  a probability measure. The triple  $(\Omega, \mathcal{F}, \mathbb{P})$  is called a **probabilistic space**.

**Int.** To construct a probabilistic model, we give

- a sample space  $\Omega$  : all the possible outcomes of the experiment
- a sigma-algebra  $\mathcal{F} \subset \mathcal{P}(\Omega)$  : the set of events
- a probability measure  $\mathbb{P}$  : gives a number in  $[0, 1]$  to every event

**Def. 1.5.** Let  $\omega \in \Omega$  (a possible outcome). Let  $A$  be an event. We say the event  $A$  **occurs** (**does not occur**) (for  $\omega$ ) if  $\omega \in A$  ( $\omega \notin A$ ).

### 1.2 Examples of probability spaces

**Def. 1.6.** Let  $\Omega$  be a finite sample space. The **Laplace model** on  $\Omega$  is the triple  $(\Omega, \mathcal{F}, \mathbb{P})$ , where  $\mathcal{F} = \mathcal{P}(\Omega)$  and  $\mathbb{P} : \mathcal{F} \rightarrow [0, 1]$  is defined by

$$\forall A \in \mathcal{F} \quad \mathbb{P}[A] = \frac{|A|}{|\Omega|}$$

### 1.3 Properties of Events

**Prop 1.1.** (Consequences of definition 1.2). Let  $\mathcal{F}$  be a sigma-algebra on  $\Omega$ . We have

**P4.**  $\emptyset \in \mathcal{F}$

**P5.**  $A_1, A_2, \dots \in \mathcal{F} \Rightarrow \bigcap_{i=1}^{\infty} A_i \in \mathcal{F}$

**P6.**  $A, B \in \mathcal{F} \Rightarrow A \cup B \in \mathcal{F}$

**P7.**  $A, B \in \mathcal{F} \Rightarrow A \cap B \in \mathcal{F}$

Event	Graphical representation	Probab. interpretation
$A^c$		$A$ does <b>not</b> occur
$A \cap B$		$A$ and $B$ occur
$A \cup B$		$A$ or $B$ occurs
$A \Delta B$		one and only one of $A$ or $B$ occurs

Figure 1: Representation of set operations

Relation	Graphical representation	Probab. interpretation
$A \subset B$		If $A$ occurs, then $B$ occurs
$A \cap B = \emptyset$		$A$ and $B$ cannot occur at the same time
$\Omega = A_1 \cup A_2 \cup A_3$ with $A_1, A_2, A_3$ pairwise disjoint		for each outcome $\omega$ , one and only one of the events $A_1, A_2, A_3$ is satisfied.

Figure 2: Representation of set relations

## 1.4 Properties of probability measures

**Prop 1.2.** (Consequences of definition 1.3). Let  $\mathbb{P}$  be a probability measure on  $(\Omega, \mathcal{F})$ .

**P3.** We have  $\mathbb{P}[\emptyset] = 0$

**P4. (additivity)** Let  $k \geq 1$ , let  $A_1, \dots, A_k$  be  $k$  pairwise disjoint events, then

$$\mathbb{P}[A_1 \cup \dots \cup A_k] = \mathbb{P}[A_1] + \dots + \mathbb{P}[A_k]$$

**P5.** Let  $A$  be an event, then

$$\mathbb{P}[A^c] = 1 - \mathbb{P}[A]$$

**P6.** If  $A$  and  $B$  are two events (not necessarily disjoint), then

$$\mathbb{P}[A \cup B] = \mathbb{P}[A] + \mathbb{P}[B] - \mathbb{P}[A \cap B]$$

**Prop 1.3. (Monotonicity).** Let  $A, B \in \mathcal{F}$ , then

$$A \subset B \Rightarrow \mathbb{P}[A] \leq \mathbb{P}[B]$$

**Prop 1.4. (Union bound).** Let  $A_1, A_2, \dots$  be a sequence of events (not necessarily disjoint), then we have

$$\mathbb{P}\left[\bigcup_{i=1}^{\infty} A_i\right] \leq \sum_{i=1}^{\infty} \mathbb{P}[A_i]$$

Union bound also applies to a finite collection of events.

**Prop 1.5.** Let  $(A_n)$  be an increasing sequence of events (i.e.  $\forall n A_n \subset A_{n+1}$ ). Then

$$\lim_{n \rightarrow \infty} \mathbb{P}[A_n] = \mathbb{P}\left[\bigcup_{n=1}^{\infty} A_n\right]. \text{ increasing limit}$$

Let  $(B_n)$  be a decreasing sequence of events (i.e.  $\forall n B_n \supset B_{n+1}$ ). Then

$$\lim_{n \rightarrow \infty} \mathbb{P}[B_n] = \mathbb{P}\left[\bigcap_{n=1}^{\infty} B_n\right]. \text{ decreasing limit}$$

<sup>1</sup>i.e.  $\Omega = B_1 \cup \dots \cup B_n$  and the events are pairwise disjoint.

## 1.5 Conditional probabilities

**Def. 1.7.** Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be some probability space. Let  $A, B$  be two events with  $\mathbb{P}[B] > 0$ . The **conditional probability of  $A$  given  $B$**  is defined by

$$\mathbb{P}[A|B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]}$$

**Ex. 1.4.** We consider the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  corresponding to the throw of one die. Let  $A = \{1, 2, 3\}$  and  $B = \{2, 4, 6\}$ . Then

$$\mathbb{P}[A|B] = \frac{\mathbb{P}[A \cap B]}{\mathbb{P}[B]} = \frac{\frac{1}{6}}{\frac{1}{2}} = \frac{1}{3}$$

**Prop 1.6.** Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be some probability space. Let  $B$  be an event with positive probability. Then  $\mathbb{P}[\cdot | B]$  is a probability measure on  $\Omega$ .

**Prop 1.7. (Formula of total probability).** Let  $B_1, \dots, B_n$  be a partition<sup>1</sup> of the sample space  $\Omega$  with  $\mathbb{P}[B_i] > 0$  for every  $i \leq i \leq n$ . Then one has

$$\forall A \in \mathcal{F} \quad \mathbb{P}[A] = \sum_{i=1}^n \mathbb{P}[A|B_i] \mathbb{P}[B_i]$$

**Prop 1.8. (Bayes formula).** Let  $B_1, \dots, B_n \in \mathcal{F}$  be a partition of  $\Omega$  with  $\mathbb{P}[B_i] > 0 \forall i$ . For every event  $A$  with  $\mathbb{P}[A] > 0$  we have

$$\forall i = 1, \dots, n \quad \mathbb{P}[B_i|A] = \frac{\mathbb{P}[A|B_i] \mathbb{P}[B_i]}{\sum_{j=1}^n \mathbb{P}[A|B_j] \mathbb{P}[B_j]}$$

**Ex. 1.5.** Test to detect a disease which concerns about 1/10000 of the population. The test gives the right answer 99% of the time. If a patient has a positive test, what is the probability that he is actually sick?

We modeled the situation as  $\Omega = \{0, 1\} \times \{0, 1\}$ .  $\mathcal{F} = \mathcal{P}(\Omega)$  and an outcome is  $\omega = (\omega_1, \omega_2)$ , where  $\omega_1$  is 1 if the patient is sick and  $\omega_2$  is 1 if the test is positive. Let  $S = \{(1, 0), (1, 1)\}$  be the event that the patient is sick and  $T = \{(0, 1), (1, 1)\}$  the event that the test is positive.

From the hypotheses, we have

$$\mathbb{P}[S] = \frac{1}{10000}, \quad \mathbb{P}[T|S] = \frac{99}{100}, \quad \mathbb{P}[T|S^c] = \frac{1}{100}$$

By applying the Bayes formula to the partition  $\Omega = S \cup S^c$ , we obtain

$$\mathbb{P}[S|T] = \frac{\mathbb{P}[T|S] \mathbb{P}[S]}{\mathbb{P}[T|S] \mathbb{P}[S] + \mathbb{P}[T|S^c] \mathbb{P}[S^c]} \simeq 0.0098$$

## 1.6 Independence

**Def. 1.8.** Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space. Two events  $A$  and  $B$  are said to be **independent** if

$$\mathbb{P}[A \cap B] = \mathbb{P}[A] \mathbb{P}[B]$$

$A$  is independent of  $B$  iff  $A$  is independent of  $B^c$ .

If  $\mathbb{P}[A] \in \{0, 1\}$ , then  $A$  is independent of every event.

If  $A$  is independent with itself (i.e.  $\mathbb{P}[A \cap A] = \mathbb{P}[A]^2$ ), then  $\mathbb{P}[A] \in \{0, 1\}$ .

**Prop 1.9.** Let  $A, B \in \mathcal{F}$  be two events with  $\mathbb{P}[A], \mathbb{P}[B] > 0$ . Then the following are equivalent :

- $\mathbb{P}[A \cap B] = \mathbb{P}[A] \mathbb{P}[B]$  : **A and B are independent**
- $\mathbb{P}[A|B] = \mathbb{P}[A]$  : **the occurrence of B has no influence on A**
- $\mathbb{P}[B|A] = \mathbb{P}[B]$  : **the occurrence of A has no influence on B**

**Def. 1.9.** Let  $I$  be an arbitrary set of indices. A collection of events  $(A_i)_{i \in I}$  is said to be **independent** if

$$\forall J \subset I \text{ finite} \quad \mathbb{P}\left[\bigcap_{j \in J} A_j\right] = \prod_{j \in J} \mathbb{P}[A_j]$$

**Int.** Three events  $A, B$  and  $C$  are independent if the following 4 equations are satisfied :

$$\mathbb{P}[A \cap B] = \mathbb{P}[A] \mathbb{P}[B]$$

$$\mathbb{P}[A \cap C] = \mathbb{P}[A] \mathbb{P}[C]$$

$$\mathbb{P}[B \cap C] = \mathbb{P}[B] \mathbb{P}[C]$$

$$\mathbb{P}[A \cap B \cap C] = \mathbb{P}[A] \mathbb{P}[B] \mathbb{P}[C]$$

## 2 Random variables and distribution functions

### 2.1 Abstract definition

**Def. 2.1.** Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space. A **random variable** (r.v.) is a map  $X : \Omega \rightarrow \mathbb{R}$  s.t.

$$\forall a \in \mathbb{R} \quad \{\omega \in \Omega : X(\omega) \leq a\} \in \mathcal{F}$$

**Ex. 2.1.** We throw a fair die. The sample space is  $\Omega = \{1, 2, 3, 4, 5, 6\}$  and we consider the Laplace model  $(\Omega, \mathcal{F}, \mathbb{P})$ . Suppose we gamble on the outcome in such a way that our profit is  $-1$  if the outcome is 1, 2 or 3; 0 if the outcome is 4 and 2 if the outcome is 5 or 6. Our profit can be represented by the mapping  $X$  defined by

$$\forall \omega \in \Omega \quad X(\omega) = \begin{cases} -1 & \text{if } \omega = 1, 2, 3, \\ 0 & \text{if } \omega = 4, \\ 2 & \text{if } \omega = 5, 6 \end{cases}$$

Since  $\mathcal{F} = \mathcal{P}(\Omega)$ , we have  $\{\omega : X(\omega) \leq a\} \in \mathcal{F}$  for every  $a$ . Therefore,  $X$  is a random variable on  $(\Omega, \mathcal{F}, \mathbb{P})$ .

**Def. 2.2.** When events are defined in terms of random variable, we omit the dependence in  $\omega$ . E.g. for  $a \leq b$  we write

$$\{X \leq a\} = \{\omega \in \Omega : X(\omega) \leq a\}$$

$$\{a < X \leq b\} = \{\omega \in \Omega : a < X(\omega) \leq b\}$$

$$\{X \in \mathbb{Z}\} = \{\omega \in \Omega : X(\omega) \in \mathbb{Z}\}$$

When consider the probability of events as above, we omit the brackets

$$\mathbb{P}[X \leq a] = \mathbb{P}[\{X \leq a\}] = \mathbb{P}[\{\omega \in \Omega : X(\omega) \leq a\}]$$

### 2.2 Distribution function

**Def. 2.3.** Let  $X$  be a random variable on a prob. space  $(\Omega, \mathcal{F}, \mathbb{P})$ . The **distribution function of  $X$**  is the function  $F_X : \mathbb{R} \rightarrow [0, 1]$  defined by

$$\forall a \in \mathbb{R} \quad F_X(a) = \mathbb{P}[X \leq a]$$

**Ex. 2.2.** Same example with the die. Let  $X$  be the random variable defined as above. For  $a \in \mathbb{R}$  we have

$$F_X(a) = \begin{cases} 0 & \text{if } a < -1, \\ 1/2 & \text{if } -1 \leq a < 0, \\ 2/3 & \text{if } 0 \leq a < 2, \\ 1 & \text{if } a \geq 2 \end{cases}$$

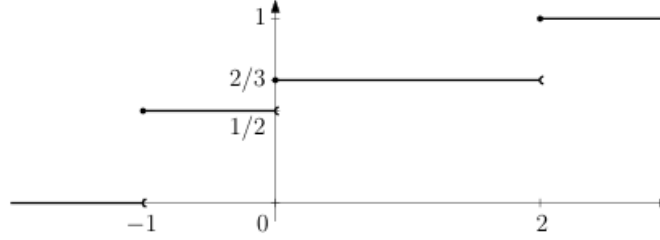


Figure 3: Graph of the distribution function  $F_X$

**Prop 2.1. (Basic identity).** Let  $a < b$  be two real numbers. Then

$$\mathbb{P}[a < X \leq b] = F(b) - F(a)$$

**Prop 2.2.** Let  $X$  be a r.v. on some probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . The distribution function  $F = F_X : \mathbb{R} \rightarrow [0, 1]$  of  $X$  satisfies the following properties :

- i.  $F$  is nondecreasing
- ii.  $F$  is right continuous<sup>2</sup>
- iii.  $\lim_{a \rightarrow -\infty} F(a) = 0$  and  $\lim_{a \rightarrow \infty} F(a) = 1$

### 2.3 Independence

**Def. 2.4.** Let  $X_1, \dots, X_n$  be  $n$  random variables on some probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . We say that they are **independent** if  $\forall x_1, \dots, x_n \in \mathbb{R} \quad \mathbb{P}[X_1 \leq x_1, \dots, X_n \leq x_n] = \mathbb{P}[X_1 \leq x_1] \dots \mathbb{P}[X_n \leq x_n]$ .

One can show that  $X_1, \dots, X_n$  are independent iff  $\forall I_1 \subset \mathbb{R}, \dots, I_n \subset \mathbb{R}$  intervals  $\{X_1 \in I_1\}, \dots, \{X_n \in I_n\}$  are independent.

**Prop 2.3. (Grouping).** Let  $X_1, \dots, X_n$  be  $n$  independent r.v. Let  $1 \leq i_1 < i_2 < \dots < i_k \leq n$  be some indices and  $\phi_1, \dots, \phi_k$  some functions. Then  $Y_1 = \phi_1(X_{i_1}, \dots, X_{i_1}), Y_2 = \phi_2(X_{i_1+1}, \dots, X_{i_2}), \dots, Y_k = \phi_k(X_{i_{k-1}+1}, \dots, X_{i_k})$  are independent.

**Def. 2.5.** An infinite sequence  $X_1, X_2, \dots$  of random variables is said to be

- **independent** if  $X_1, \dots, X_n$  are independent for every  $n$
- **independent and identically distributed** (i.i.d) if they are independent and they have the same distribution function, i.e.  $\forall i, j \quad F_{X_i} = F_{X_j}$ .

### 2.4 Transformation of random variables

We can create r.v. from other r.v. on the same probability space. For example, consider  $Z_1 = \exp(X_1), Z_2 = X_1 + X_2$ . Not to forget : r.v. are maps  $\Omega \rightarrow \mathbb{R}$ .

We can work with r.v. as if they were real numbers with the following notation :

**Def. 2.6.** If  $X$  is a r.v. and  $\phi : \mathbb{R} \rightarrow \mathbb{R}$ , then we write

$$\phi(X) := \phi \circ X$$

to  $\phi(X)$  a new mapping  $\Omega \rightarrow \mathbb{R}$ .

We also consider function of several variables. If  $X_1, \dots, X_n$  are  $n$  r.v. and  $\phi : \mathbb{R}^n \rightarrow \mathbb{R}$ , then we write

$$\phi(X_1, \dots, X_n) := \phi \circ (X_1, \dots, X_n)$$

### 2.5 Construction of random variables

**Def. 2.7.** Let  $p \in [0, 1]$ . A r.v.  $X$  is said to be a **Bernoulli r.v. with parameter  $p$**  if

$$\mathbb{P}[X = 0] = 1 - p \quad \text{and} \quad \mathbb{P}[X = 1] = p$$

In this case, we write  $X \sim \text{Ber}(p)$ .

**Prop 2.4. (Existence theorem of Kolmogorov).** There exists a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  and an infinite sequence of r.v.  $X_1, X_2, \dots$  (on this probability space) that is an iid sequence of Bernoulli r.v. with parameter  $1/2$ .

<sup>2</sup>i.e.  $F(a) = \lim_{h \downarrow 0} F(a + h)$  for every  $a \in \mathbb{R}$

**Prop 2.5.** A r.v.  $U$  is said to be **uniform r.v. in  $[0, 1]$**  if its distribution function is equal to

$$F_U(x) = \begin{cases} 0 & x < 0 \\ x & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$$

In this case, we write  $U \sim \mathcal{U}([0, 1])$ .

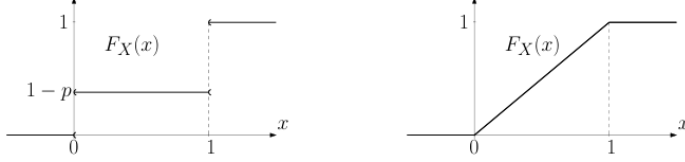


Figure 4: Left: distribution function of a Bernoulli r.v. with parameter  $p$ . Right: distribution function of a uniform r.v. in  $[0, 1]$ .

**Prop 2.6.** The mapping  $Y : \Omega \rightarrow [0, 1]$  defined by  $Y(\omega) = \sum_{n=1}^{\infty} 2^{-n} X_n(\omega)$  is a uniform r.v. in  $[0, 1]$ .

**Def. 2.8.** The **generalized inverse** of  $F^3$  is the mapping  $F^{-1} : (0, 1) \rightarrow \mathbb{R}$  defined by

$$\forall \alpha \in (0, 1) \quad F^{-1}(\alpha) = \inf\{x \in \mathbb{R} : F(x) \geq \alpha\}$$

**Int.** By definition of the infimum and using right continuity of  $F$ , we have  $\forall x \in \mathbb{R}$  and  $\forall \alpha \in (0, 1)$

$$(F^{-1}(\alpha) \leq x) \iff (\alpha \leq F(x))$$

**Prop 2.7. (Inverse transform sampling).** Let  $F : \mathbb{R} \rightarrow [0, 1]$ .<sup>4</sup> Let  $U$  be a uniform r.v. in  $[0, 1]$ . Then the r.v.  $X = F^{-1}(U)$  has distribution  $F_X = F$ .

**Prop 2.8.** Let  $F_1, F_2, \dots$  be a sequence of functions  $\mathbb{R} \rightarrow [0, 1]$ .<sup>5</sup> Then there exist a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  and a sequence of independent r.v.  $X_1, X_2, \dots$  on this probability space s.t.

- for every  $i$   $X_i$  has distribution function  $F_i$  (i.e.  $\forall x \mathbb{P}[X_i \leq x] = F_i(x)$ )
- $X_1, X_2, \dots$  are independent

### 3 Discrete and continuous r.v.

#### 3.1 Discontinuity / continuity points of $F$

**Prop 3.1.** Let  $X : \Omega \rightarrow \mathbb{R}$  be a r.v. with distribution function  $F$ . Then for every  $a$  in  $\mathbb{R}$  we have

$$\mathbb{P}[X = a] = F(a) - F(a-)$$

where  $F(a-) := \lim_{h \downarrow 0} F(a - h)$ .

**Int.** Fix  $a \in \mathbb{R}$

→ If  $F$  is not continuous at a point  $a \in \mathbb{R}$ , then the “jump size”  $F(a) - F(a-)$  is equal to the probability that  $X = a$

→ If  $F$  is continuous at a point  $a \in \mathbb{R}$ , then  $\mathbb{P}[X = a] = 0$

#### 3.2 Almost sure events

**Def. 3.1.** Let  $A \in \mathcal{F}$  be an event. We say that  $A$  occurs **almost surely (a.s.)** if  $\mathbb{P}[A] = 1$ .

#### 3.3 Discrete random variables

**Def. 3.2.** A r.v.  $X : \Omega \rightarrow \mathbb{R}$  is said to be **discrete** if there exists some set  $W \subset \mathbb{R}$  finite or countable s.t.  $X \in W$  a.s..

**Def. 3.3.** Let  $X$  be a discrete r.v. taking some values in some finite or countable set  $W \subset \mathbb{R}$ . The **distribution of  $X$**  is the sequence of numbers  $(p(x))_{x \in W}$  defined by

$$\forall x \in W \quad p(x) := \mathbb{P}[X = x]$$

**Prop 3.2.** The distribution  $(p(x))_{x \in W}$  of a discrete r.v. satisfies  $\sum_{x \in W} p(x) = 1$ .

**Prop 3.3.** Let  $X$  be a discrete r.v. with values in a finite or countable set  $W$  almost surely, and distribution  $p$ . Then the distribution function of  $X$  is given by

$$\forall x \in \mathbb{R} \quad F_X(x) = \sum_{\substack{y \leq x \\ y \in W}} p(y)$$

**Int.**  $W = \{\text{positions of the jumps of } F_X\}$ ,  
 $p(x) = \text{“height of the jump” at } x \in W$ .

#### 3.4 Examples of discrete random variables

The simplest (non constant) r.v. is the Bernoulli r.v. defined in definition 2.7.

**Def. 3.4.** Let  $0 \leq p \leq 1$ , let  $n \in \mathbb{N}$ . A r.v.  $X$  is said to be a **binomial r.v. with parameters  $n$  and  $p$**  if it takes values in  $W = \{0, \dots, n\}$  and

$$\forall k \in \{0, \dots, n\} \quad \mathbb{P}[X = k] = \binom{n}{k} p^k (1-p)^{n-k}$$

In that case we write  $X \sim \text{Bin}(n, p)$ .

**Prop 3.4. (Sum of independent Bernoulli and binomial).** Let  $0 \leq p \leq 1$ , let  $n \in \mathbb{N}$ . Let  $X_1, \dots, X_n$  be  $n$  independent Bernoulli r.v. with parameter  $p$ . Then

$$S_n := X_1 + \dots + X_n$$

is a binomial r.v. with parameter  $n$  and  $p$ .

**Int.** In particular, the distribution  $\text{Bin}(1, p)$  is the same as the distribution  $\text{Ber}(p)$ . One can also check that if  $X \sim \text{Bin}(m, p)$  and  $Y \sim \text{Bin}(n, p)$  and  $X, Y$  are independent, then  $X + Y \sim \text{Bin}(m + n, p)$ .

**Def. 3.5.** Let  $0 \leq p \leq 1$ . A r.v.  $X$  is said to be a **geometric r.v. with parameter  $p$**  if it takes values in  $W = \mathbb{N} \setminus \{0\}$  and

$$\forall k \in \mathbb{N} \quad \mathbb{P}[X = k] = (1-p)^{k-1} \cdot p$$

In this case, we write  $X \sim \text{Geom}(p)$ .

<sup>3</sup>satisfying prop. 2.2

<sup>4</sup>See footnote 3

<sup>5</sup>See footnote 3

**Prop 3.5.** Let  $X_1, X_2, \dots$  be a sequence of infinitely many independent Bernoulli r.v. with parameter  $p$ . Then

$$T := \min\{n \geq 1 : X_n = 1\}$$

is a geometric r.v. with parameter  $p$ .

**Prop 3.6. (Absence of memory of the geometric distribution).** Let  $T \sim \text{Geom}(p)$  for some  $0 < p < 1$ . Then

$$\forall n \geq 0 \forall k \geq 1 \quad \mathbb{P}[T \geq n + k | T > n] = \mathbb{P}[T \geq k]$$

**Def. 3.6.** Let  $\lambda > 0$  be a positive real number. A r.v.  $X$  is said to be a **Poisson r.v. with parameter  $\lambda$**  if it takes values in  $W = \mathbb{N}$  and

$$\forall k \in \mathbb{N} \quad \mathbb{P}[X = k] = \frac{\lambda^k}{k!} e^{-\lambda}$$

In this case, we write  $X \sim \text{Poisson}(\lambda)$ .

**Prop 3.7. (Poisson approximation of the binomial).** Let  $\lambda > 0$ . For every  $n \geq 1$ , consider a r.v.  $X_n \sim \text{Bin}(n, \frac{\lambda}{n})$ . Then

$$\forall k \in \mathbb{N} \quad \lim_{n \rightarrow \infty} \mathbb{P}[X_n = k] = \mathbb{P}[N = k]$$

where  $N$  is a Poisson r.v. with parameter  $\lambda$ .

### 3.5 Continuous random variables

**Def. 3.7.** A r.v.  $X : \Omega \rightarrow \mathbb{R}$  is said to be **continuous** if its distribution function  $F_X$  can be written as

$$F_X(a) = \int_{-\infty}^a f(x) dx \quad \text{for all } a \in \mathbb{R}$$

for some nonnegative function  $f : \mathbb{R} \rightarrow \mathbb{R}_+$ , called the **density** of  $X$ .

**Int.**  $f(x) dx$  represents the probability that  $X$  takes a value in the infinitesimal interval  $[x, x + dx]$ .

**Prop 3.8.** The density  $f$  of a r.v. satisfies  $\int_{-\infty}^{+\infty} f(x) dx = 1$ .

**Prop 3.9.** Let  $X$  be a r.v. Assume the distribution function  $F_X$  is continuous and piecewise  $\mathcal{C}^1$ , i.e. that there exist  $x_0 = -\infty < x_1 < \dots < x_{n-1} < x_n = +\infty$  s.t.  $F_X$  is  $\mathcal{C}^1$  on every interval  $(x_i, x_{i+1})$ . Then  $X$  is a continuous r.v. and a density  $f$  can be constructed by defining

$$\forall x \in (x_i, x_{i+1}) \quad f(x) = F'_X(x)$$

and setting arbitrary values at  $x_1, \dots, x_{n-1}$ .