NETW ORKS 2021

Scope

Key points

What are graphical models?

Parameterizing network mode

Graphical models fo networks

Conclusions

References

11005. Network Models and Graphical Models: A Survey

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Networks 2021, Session S65. Machine Learning



Scope

Kev point

What ar graphica models?

Parameterizing network mode

Graphical models fo networks

Conclusion

D-f----

- Network models represent dependencies with graphs; graphical models (one type of which are "Bayesian networks") represent dependencies with graphs. That causes confusion. But they are very different!
- Graphical models haven't done the best job at networks, and networks haven't made use of graphical models
- This talk is primarily to clarify (or create!) a *conceptual* connection between these two types of models
- For examples of *applications* of graphical modeling to networks, see, e.g., Farasat et al. (2015), Maier et al. (2014), and Airoldi et al. (2008)



Key points

What ar graphica models?

Parameterizing network mode

Graphical models fo networks

Conclusion

References

Key points

- Graphical models represent dependencies (and causal relationships) between variables
- Networks models are models of dyads, which represent dependencies between observations
 - Dyads can be modeled as random variables (e.g., Bernoulli for unweighted; Poisson for count; etc.)
 - Dyads are themselves dependent! (reciprocity, triadic closure, degree constraints)
 - Graphical models can represent these "dyadic dependencies"!

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Key visualization

Scope

Key points

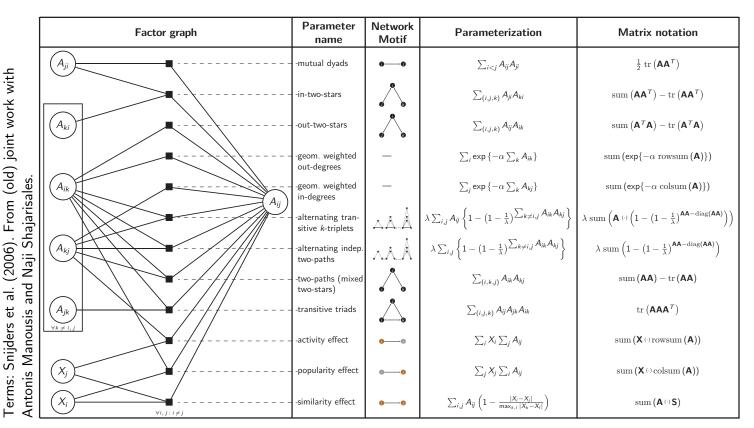
What are graphica models?

Parameterizing network mode

Graphical models fo networks

Conclusion

Reference





Key point

What are graphical models?

Parameterizii network mod

Graphical models fo networks

Conclusion

Deference

What are graphical models?



Graphical models = graphs for variables

Scope

Kev point

What are graphical models?

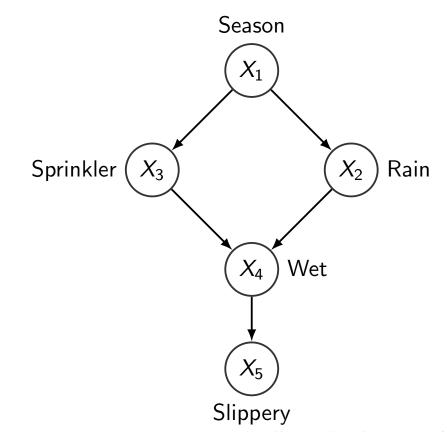
Parameterizing network model

Graphical models for networks

Conclusion

D-f-----

- Like path diagrams in psych, but more formal. From CS in 90s
- Represent relationships between variables; can reason through dependencies
 - Sprinklers are not directly dependent on Rain, but if we know the grass is wet, we know either it rained or sprinklers were on (at least one is true)
- With probability distributions on the nodes, they represent conditional independencies
 - Equivalent to structural equation modeling (SEMs)!





Can represent causality

Scope

Kev point

What are graphical models?

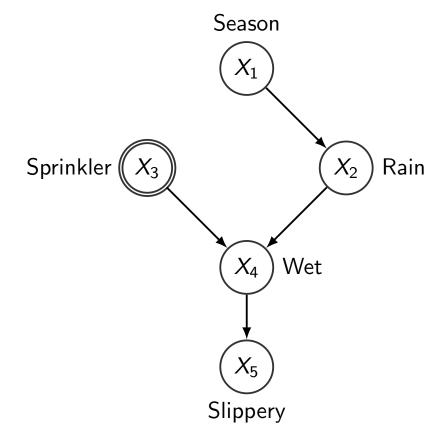
Parameterizing network model

Graphical models for networks

Conclusion

Reference

- Can also reason about causality
- Interventions block "paths"
 - Pearl introduced the "do" operator to notate this algebraically
- Are algorithms to determine identifiability of parameters from a given (or assumed) causal structure (Bayes ball)
 - Causal inference techniques to estimate a causal graph (e.g., TETRAD algorithm) also exist, but theoretical guarantees for these procedures require strong, untestable, and almost certainly false assumptions



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Mostly used for bookkeeping

Scope

Key point

What are graphical models?

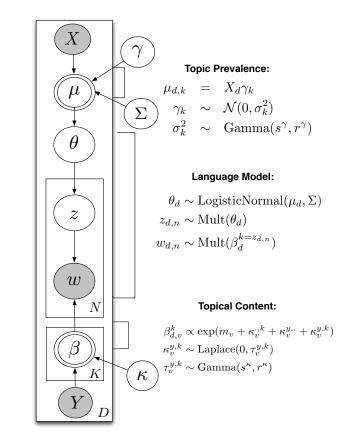
Parameterizing network model

Graphical models for networks

Conclusion

Reference

- Most machine learning applications are effectively "bookkeeping"
- E.g., in structural topic modeling (Roberts et al., 2013), a topic doesn't "cause" a document, but representing it as a directed tie is to help keep track of things for estimation



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Networks in graphical models

Scope

Key point

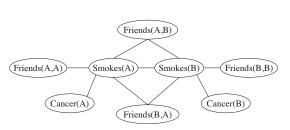
What are graphical models?

Parameterizing network mode

Graphical models for networks

Conclusion

Reference



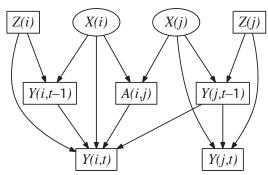
graphical model nodes in

Taskar, 2007): awkward

Network edges as

a "relational Markov

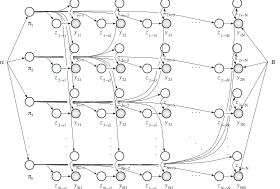
network" (Getoor &



Full representation, but only of a single type of dependency (block membership) (Airoldi et

al., 2008)

Network is only a confounder (Shalizi & Thomas, 2011), no direct dependencies between edges





Two "directions" of dependencies

Key point

What are graphical models?

Parameterizing network mode

Graphical models for networks

Conclusions

Reference

• If covariates are independent (trivial graphical model, what we usually assume in linear regression), the joint distribution of a response and a design matrix is:

$$p(Y, \mathbf{X}) = p(Y, X_1, ..., X_d) = p(Y|X_1, ..., X_d) \prod_{i=1}^d p(X_i)$$

- (If we assume fixed X, as we usually do, the probabilities of X_j 's go away)
- But a true, complete joint factorization of the conditional distribution would be over *observations* as well:

$$p(Y|\mathbf{X}) = p(y_1, ..., y_n|x_{11}, ..., x_{1d}, x_{21}, ..., x_{2d},, x_{n1}, ..., x_{nd})$$

• An iid assumption applies to the *observations*, and are how we even have multiple observations to estimate anything. This looks like:

$$p(Y|\mathbf{X}) = p(y_1, ..., y_n | \mathbf{x}_1, ..., \mathbf{x}_n) \stackrel{\text{iid}}{=} \prod_{i=1}^n p(y_i | \mathbf{x}_i)$$



Key point

What are graphical models?

Parameterizing network models

Graphical models for networks

Conclusion

Reference

Parameterizing network models



Networks: Dependencies as observations

Scope

Key point

What are graphica models?

Parameterizing network models

Graphical models for networks

Conclusion

References

Better way of looking at networks: make dyads the observations

- The response is now an edge, or edge attribute
- Transform all node covariates into edge covariates, e.g.,
 - As a difference between continuous node attributes
 - Indicator for if nodes in same category or not (or, make new categories out of possible pairs, e.g., M→M, M→F, F→M, F→F)
 - As "sender" and/or "receiver" attributes

			X_2		X_d
1	y_1	<i>x</i> ₁₁	<i>x</i> ₁₂ <i>x</i> ₂₂		x_{1d}
2	<i>y</i> ₂	<i>X</i> ₂₁	<i>X</i> ₂₂		X_{2d}
:	:	: <i>x</i> _{n1}	:		:
n	Уn	x_{n1}	X_{n2}	• • •	X_{nd}



index	from	to	Y	$ \hspace{.05cm} W_1 \hspace{.05cm}$	W_2	W_3	
$\overline{e_1}$	1	2	<i>y</i> ₁₂	$1(x_{11}=x_{21})$	$x_{12} - x_{22}$	<i>X</i> ₁₃	• • •
e_2	2	3	<i>y</i> ₂₃	$1(x_{11}=x_{31})$	$x_{12} - x_{32}$	<i>x</i> ₁₃	• • •
:	:	:	:	:	:	:	
e_{n+1}	2	1	<i>y</i> 21	$1(x_{21}=x_{11})$	$x_{22}-x_{12}$	<i>X</i> ₂₃	• • •
:	:	:	:	:	:	:	
$e_{2\binom{n}{2}}$	n-1	n	$y_{(n-1)n}$	$1(x_{(n-1)1}=x_{n1})$	$x_{(n-1)2}-x_{n2}$	$X_{(n-1)3}$	



Parameterizing

Scope

Key points

What are graphical models?

Parameterizing network models

Graphical models for networks

Conclusions

References

Descriptively,

$$a_{ij} = \begin{cases} 1 & \text{if there is a tie } i \to j \\ 0 & \text{otherwise.} \end{cases}$$

Turn this into a random variable:

$$A_{ii} \stackrel{\text{iid}}{\sim} \text{Bernoulli}(p)$$



As a logistic regression

Scope

Key points

What are graphica models?

Parameterizing network models

Graphical models fo networks

Conclusions

References

 Add some covariates and it can become a logistic regression:

$$A_{ij}|x_i,x_j \sim \text{Bernoulli}\left(f(x_i,x_j)^T oldsymbol{eta}
ight)$$

$$\mathcal{L}(oldsymbol{eta}) = \prod_{i=1}^n \prod_{i \neq i} f(x_i,x_j)^T oldsymbol{eta}^{a_{ij}} (f(x_i,x_j)^T oldsymbol{eta})^{1-a_{ij}}$$

The MLE of an intercept-only model is just the density.

$$\widehat{\beta}_{MLE} = \frac{1}{2 \times \binom{n}{2}} \sum_{i \neq i} a_{ij}$$



Problem: Dyads are dependent, too

Scope

Key point

What are graphica models?

Parameterizing network models

Graphical models fo networks

Conclusions

D-f-...

- In the language of ERGMs, "dyadic dependencies"
 - Social networks: reciprocity makes $A_{ij} \not\perp \!\!\! \perp A_{ji}$
- The p_1 model (Holland & Leinhardt, 1981) deals with reciprocity as a one-off dependency by modeling edges as multinomial, with a cross term:

$$\mathbb{P}(A_{ij} = a_{ij}, A_{ji} = a_{ji}) = \frac{1}{k_{ij}} \exp \left\{ a_{ij} (\mu + \alpha_i + \beta_i) + a_{ji} (\mu + \alpha_j + \beta_j) + \rho a_{ij} a_{ji} \right\}$$



Models for dyadic dependencies

Scop

Key point

What are graphica models?

Parameterizing network models

Graphical models for networks

Conclusion

D-f----

- Stochastic blockmodels (Wang & Wong, 1987): alternative to p_1 , two-level hierarchical version of Bernoulli model
- Latent space models (Hoff et al., 2002): can be seen as graphical models with observable nodes for edges, produced from hidden nodes representing latent position



Markov property for network edges

Scope

Key point

What are graphica models?

Parameterizing network models

Graphical models fo networks

Conclusions

D-f----

- Landmark work: Frank & Strauss (1986)
- Markov dependence assumption: "A graph is said to be a Markov graph if only incident dyads can be conditionally dependent."
- In retrospect, we can clarify this in terms of graphical models
 - The "graph" is the network, and the Markov property is of the graphical model of the network edges as Bernoulli variables



Kev point

What ar graphica models?

Parameterizing network models

Graphical models fo networks

Conclusion

Reference

Markov property for network edges

 Remarkably, using Hammersley-Clifford, Frank & Strauss proved that the graphical model of an undirected network is Markov if and only if

$$P_{\theta}(\mathbf{A}) = \frac{1}{\kappa(\theta)} \exp \left\{ \theta_0 L(\mathbf{A}) + \sum_{k=1}^{n-1} \theta_k S_k(\mathbf{A}) + \theta_\tau T(\mathbf{A}) \right\}$$

Normalization constant: need to sum over $2^{2 \times \binom{n}{2}}$ possible networks for each candidate θ

Non-maximal Number of k-stars triangles

• Especially surprising part (Kolaczyk, 2009): how did triangles come out of this as a sufficient statistic??



ey point

What are graphica models?

Parameterizi network mod

Graphical models for networks

Conclusion

Deference

Graphical models for networks



Graphical models for networks

General modeling of dyadic dependencies

- This eventually led to Exponential-family Random Graph Models, which can model generic dependencies between edges
- Can add any sufficient statistic, although they can be collinear. E.g., two-paths are collinear with in-degrees, out-degrees, and mutual dyads (Snijders et al., 2006)

$$\sum_{i,j,k:k\neq i} A_{ij} A_{jk} = \sum_{j=1}^{n} \sum_{i,k:k\neq i} A_{ij} A_{jk} = \sum_{j=1}^{n} \left(A_{+j} A_{j+} - \sum_{i=1}^{n} A_{ij} A_{ji} \right)$$

 But: bad theoretical properties, tricky to estimate, and tricky to specify



Dyadic dependencies in a graph

Scope

Key point

What are graphical models?

Parameterizing network model

From (old) joint work with

(2006).

–

Snijders

Terms:

Graphical models for networks

Conclusion

Reference

	Factor graph	Parameter name	Network Motif	Parameterization	Matrix notation
	(A _{ji})	-mutual dyads	90	$\sum_{i < j} A_{ij} A_{ji}$	$rac{1}{2} \operatorname{tr} \left(\mathbf{A} \mathbf{A}^T ight)$
		-in-two-stars		$\sum_{(i,j,k)} A_{ji} A_{ki}$	$\mathrm{sum}\left(\boldsymbol{A}\boldsymbol{A}^{\mathcal{T}}\right)-\mathrm{tr}\left(\boldsymbol{A}\boldsymbol{A}^{\mathcal{T}}\right)$
	A _{ki}	-out-two-stars		$\sum_{(i,j,k)} A_{ij} A_{ik}$	$\mathrm{sum}\left(\boldsymbol{A}^{T}\boldsymbol{A}\right)-\mathrm{tr}\left(\boldsymbol{A}^{T}\boldsymbol{A}\right)$
		-geom. weighted out-degrees	_	$\sum_{i} \exp\left\{-\alpha \sum_{k} A_{ik}\right\}$	$\operatorname{sum}\left(\exp\{-\alpha \operatorname{rowsum}\left(\mathbf{A}\right)\}\right)$
Shajarisales.	A_{ik}	-geom. weighted in-degrees	_	$\sum_{j} \exp\left\{-\alpha \sum_{k} A_{kj}\right\}$	$\operatorname{sum}\left(\exp\{-\alpha\operatorname{colsum}\left(\mathbf{A}\right)\}\right)$
ajari		-alternating tran- sitive <i>k</i> -triplets	aa A	$\lambda \sum_{i,j} A_{ij} \left\{ 1 - \left(1 - \frac{1}{\lambda}\right)^{\sum_{k \neq i,j} A_{ik} A_{kj}} \right\}$	$\lambda \operatorname{sum}\left(\mathbf{A} \left(0\right) \left(1 - \left(1 - \frac{1}{\lambda}\right)^{\mathbf{A}\mathbf{A} - \operatorname{diag}(\mathbf{A}\mathbf{A})}\right)\right)$ $\lambda \operatorname{sum}\left(1 - \left(1 - \frac{1}{\lambda}\right)^{\mathbf{A}\mathbf{A} - \operatorname{diag}(\mathbf{A}\mathbf{A})}\right)$
sis and Naji	Akj	-alternating indep. two-paths	A.A.A	$\lambda \sum_{i,j} \left\{ 1 - \left(1 - \frac{1}{\lambda}\right)^{\sum_{k \neq i,j} A_{ik} A_{kj}} \right\}$	$\lambda \operatorname{sum} \left(1 - \left(1 - \frac{1}{\lambda} \right)^{AA - \operatorname{diag}(AA)} \right)$
		-two-paths (mixed two-stars)		$\sum_{(i,k,j)} A_{ik} A_{kj}$	$\mathrm{sum}(\mathbf{A}\mathbf{A})-\mathrm{tr}(\mathbf{A}\mathbf{A})$
	A_{jk} $\forall k \neq i, j$	-transitive triads		$\sum_{(i,j,k)} A_{ij} A_{jk} A_{ik}$	$\mathrm{tr}\left(\mathbf{A}\mathbf{A}\mathbf{A}^{T} ight)$
	▼ × ≠ 1, 1	-activity effect	00	$\sum_i X_i \sum_j A_{ij}$	$\mathrm{sum}\left(\boldsymbol{X}^{(\cdot)}\mathrm{rowsum}\left(\boldsymbol{A}\right)\right)$
	(X_j)	-popularity effect	00	$\sum_j X_j \sum_i A_{ij}$	$\mathrm{sum}\left(\boldsymbol{X}^{(\cdot)}\mathrm{colsum}\left(\boldsymbol{A}\right)\right)$
Antonis	X_i $\forall i,j:i \neq j$	-similarity effect	00	$\sum_{i,j} A_{ij} \left(1 - rac{ X_i - X_j }{\max_{k,l} X_k - X_l } ight)$	sum (A (·) S)



Conclusions

Conclusions

11005. Network Models and Graphical Models: A Survey



Key points, redux

Scope

Key point

What ar graphica models?

Parameterizing network model

Graphical models for networks

Conclusions

D-f-----

- **Graphical models** represent *dependencies* (and causal relationships) *between variables*
- Networks models are models of dyads, which represent dependencies between observations
- They are not the same thing, but we can represent dyadic dependencies (dependencies between edges of a network, in processes like reciprocity and transitivity) as graphical models



Kev point

What ar graphica models?

Parameterizing network model

Graphical models fo networks

Conclusions

Reference

Why should we care?

- Graphical models are network models are both powerful for representing, reasoning through, and modeling dependencies
 - But graphical models haven't done the best job at networks, and networks haven't made use of graphical models
 - Maybe because they are used by different communities, and the same words ("networks", "dependencies") mean subtly different things
- Clarifying the relationship of these two types of models helps head off confusion, as well as deepen our appreciation of the idea of "dependencies"
 - As well as helping to train students
- Practically: Can graphical models help create, and estimate, new statistical network models, and unify existing ones? Almost certainly, although in many cases the estimation will still be MCMC (maybe variational inference; Celisse et al., 2012)



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