Probability Distribution Properties

1 Discrete Distributions

1.1 Uniform Random Variable

$$E[X] = \frac{x_1 + \dots + x_n}{n}; \quad Var(X) = \frac{x_1^2 + \dots + x_n^2}{n} - \left(\frac{x_1 + \dots + x_n}{n}\right)^2.$$

$$E[X] = \frac{n+1}{2}; \quad Var(X) = \frac{n^2 - 1}{12}; \quad only \text{ when } x_i \in [1, 2, \dots, n].$$

1.2 Bernoulli Random Variable

$$I = \begin{cases} 1 & \text{if } X = 1, \\ 0 & \text{if } X = 0. \end{cases}$$
$$E[I] = p; \quad Var(I) = p(1 - p).$$

1.3 Binomial Random Variable

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k} \quad \text{for } k \in [0, n].$$

$$E[X] = np; \quad Var(X) = np(1 - p).$$

1.4 Poisson Random Variable

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$$
 for $k \in \mathbb{Z}^{\geq}$.
 $E[X] = \lambda;$ $Var(X) = \lambda.$

1.5 Geometric Random Variable

$$P(X = n) = (1 - p)^{n-1}p \quad \text{for } n \in \mathbb{N}.$$

$$E[X] = \frac{1}{p}; \quad Var(X) = \frac{1 - p}{p^2}; \quad P(X > n) = (1 - p)^n; \quad P(X > n + k | X > k) = P(X > n).$$

2 Continuous Distributions

2.1 Uniform Random Variable

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha} & \text{if } x \in [\alpha, \beta], \\ 0 & \text{otherwise.} \end{cases}$$
$$E[X] = \frac{\alpha + \beta}{2}; \quad Var(X) = \frac{(\beta - \alpha)^2}{12}.$$

2.2 Exponential Random Variable

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \ge 0, \\ 0 & \text{otherwise.} \end{cases}$$

$$E[X] = \frac{1}{\lambda}; \quad Var(X) = \frac{1}{\lambda^2}; \quad P(X \ge x) = e^{-\lambda x}; \quad P(X \ge x + y | X \ge y) = e^{-\lambda x}.$$

2.3 Normal Random Variable

$$\begin{split} f(x) &= \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad for \ x \in \mathbb{R}. \\ E[X] &= \mu; \quad Var(X) = \sigma^2; \quad Z \sim N(\theta, 1) = \frac{X - \mu}{\sigma}. \end{split}$$

2.4 Gamma Random Variable

$$\begin{split} f(x) &= \frac{\lambda}{\Gamma(\alpha)} (\lambda x)^{\alpha - 1} e^{-\lambda x} \quad for \; \alpha, \beta, x \in \mathbb{R}^+; \quad \Gamma(z) = \int_0^\infty t^{z - 1} e^{-t} dt. \\ E[X] &= \frac{\alpha}{\lambda}; \quad Var(X) = \frac{\alpha}{\lambda^2}. \end{split}$$

3 Multivariate Distributions

3.1 Multivariate Normal Distribution

$$\begin{split} f(\vec{x}) &= \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} (\vec{x} - \vec{\mu})^\top \Sigma^{-1} (\vec{x} - \vec{\mu})} \quad for \ x \in \mathbb{R}^n. \\ E[X] &= \vec{\mu}; \quad Var(X) = \Sigma. \end{split}$$

Let random vector $\begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix}$ have a multivariate normal distribution and Z be the linear combination of

the random vector with coefficients $\vec{A} = \begin{bmatrix} a_1 \\ \vdots \\ a_n \end{bmatrix}$. Then, $Z \sim N(\vec{A} \cdot \vec{\mu}, \vec{A}^\top \sum \vec{A})$. For bivariate random vector $\begin{bmatrix} X \\ Y \end{bmatrix}$ with $\vec{A} = \begin{bmatrix} a \\ b \end{bmatrix}$,

$$Var(Z) = a^2 \sigma_X^2 + 2ab\rho\sigma_X\sigma_Y + b^2\sigma_Y^2.$$

3.2 Multinomial Distribution

$$P(\vec{x}) = \frac{n!}{x_1! \cdots x_k!} p_1^{x_1} \cdots p_k^{x_k} \quad \text{for } \vec{x} = \begin{bmatrix} x_1 \\ \vdots \\ x_k \end{bmatrix}, \ x_i \in \mathbb{N}, \ \sum_{i=1}^k x_i = n, \ p_i \in [0, 1].$$

$$E[\vec{x}] = n\vec{p}; \quad Var(\vec{x}) = n\overline{p(1-p)}.$$

3.3 Random Sums

$$X = \xi_1 + \dots + \xi_N;$$
 where N is a R.V. and ξ_i are i.i.d. R.Vs. $E[X] = E[\xi]E[N];$ $Var(X) = E[N]Var(\xi) + E[\xi]^2Var(N).$

4 Theorems

4.1 Chebyshev's Inequality

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}.$$

4.2 Markov Inequality

$$P(X \ge k) \le \frac{E[X]}{k}.$$

4.3 Central Limit Theorem

Let $S_n = X_1 + \cdots + X_n$. When n is large, S_n is approximately normally distributed with mean $n\mu$ and variance $n\sigma^2$.

$$\frac{S_n - n\mu}{\sqrt{n}\sigma} = \frac{\bar{S}_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1) \quad as \ n \to \infty.$$

4.4 Weak Law of Large Numbers

Let $S_n = X_1 + \cdots + X_n$.

$$\lim_{n \to \infty} P(|\bar{S}_n - \mu| < \epsilon) = 1.$$

4.5 Strong Law of Large Numbers

Let $S_n = X_1 + \cdots + X_n$.

$$P(\lim_{n\to\infty} \bar{S}_n = \mu) = 1.$$

4.6 Moment Generating Function

$$M_X(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx.$$

5 Special Markov Chains

5.1 Two State Markov Chain

$$P = \begin{array}{c|c} 0 & 1 \\ 1-a & a \\ b & 1-b \end{array} \right|; \ state \ 0, \ 1 \ are \ independent \ R.V. \ iff \ a=1-b.$$

$$P^{n} = \frac{1}{a+b} \begin{bmatrix} b & a \\ b & a \end{bmatrix} + \frac{(1-a-b)^{n}}{a+b} \begin{bmatrix} a & -a \\ -b & b \end{bmatrix}; \quad \lim_{n \to \infty} P^{n} = \begin{bmatrix} 0 & 1 \\ \frac{b}{a+b} & \frac{a}{a+b} \\ \frac{b}{a+b} & \frac{a}{a+b} \end{bmatrix}.$$

5.2 i.i.d. Markov Chain

$$\xi_1, \ldots, \xi_n$$
 are i.i.d. R.Vs.; $P(\xi = k) = a_k$ for $k \in \mathbb{Z}^{\geq}$.

5.2.1 Independent Random Variables

$$X_n = \xi; P = \begin{vmatrix} a_0 & a_1 & a_2 & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{vmatrix}.$$

5.2.2 Successive Maxima

$$X_n = max(\xi_1, \dots, \xi_n) \text{ for } n \in \mathbb{N}.$$

Let T be the number of trials needed for $X_n = max(\xi_1, \dots, \xi_n) \ge M$ for an arbitrary value M,

$$E[T] = \frac{1}{P(\xi \ge M)}.$$

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Successive Maxima can be interpreted as a Geometric(p) random variable with $p = P(\xi \ge M)$.

5.2.3 Partial Sums

$$X_n = \xi_1 + \dots + \xi_n \text{ for } n \in \mathbb{N}.$$

$$P = \left| \begin{array}{cccccc} a_0 & a_1 & a_2 & a_3 & \cdots \\ 0 & a_0 & a_1 & a_2 & \cdots \\ 0 & 0 & a_0 & a_1 & \cdots \\ 0 & 0 & 0 & a_0 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{array} \right| \quad \text{if } P(\xi = k) = a_k \text{ for } k \in \mathbb{Z}^{\geq}.$$

$$P = \begin{vmatrix} \vdots & \vdots & \vdots & \vdots \\ \cdots & a_{-1} & a_0 & a_1 & a_2 & \cdots \\ \cdots & a_{-2} & a_{-1} & a_0 & a_1 & \cdots \\ \cdots & a_{-3} & a_{-2} & a_{-1} & a_0 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \end{vmatrix}$$
 if $P(\xi = k) = a_k$ for $k \in \mathbb{Z}$.

5.2.4 Branching Processes

$$X_{n+1} = \xi_1^{(n)} + \dots + \xi_{X_n}^{(n)} \text{ for } n \in \mathbb{N}; \quad X_0 = 1.$$

$$u_n = \text{probability of extinction by } n\text{-th generation}$$

$$= \sum_{k=0}^{\infty} P(\xi = k) (u_{n-1})^k \text{ for } n \in \mathbb{N};$$

$$u_0 = 0.$$

From random sums, we know

$$E[X_{n+1}] = E[\xi]E[X_n]; \quad Var(X_{n+1}) = Var(\xi)E[X_n] + E[\xi]^2 Var(X_n).$$

After derivation,

$$E[X_n] = E[\xi]^n; \quad Var(X_n) = Var(\xi)E[\xi]^{n-1} \times \begin{cases} n & \text{if } E[\xi] = 1, \\ \frac{1 - E[\xi]^n}{1 - E[\xi]} & \text{otherwise.} \end{cases}$$

5.3 One-dimensional Random Walk

5.3.1 Simple Random Walk

$$r_i = r = 0;$$
 $q_i = q = \frac{1}{2};$ $p_i = p = \frac{1}{2}.$

5.3.2 Gambler's Ruin

Let u_i be the probability of X_n reaching state 0 before state N given current state i.

$$u_i = \begin{cases} \frac{N-i}{N} & \text{if } p = q = \frac{1}{2}, \\ \frac{(q/p)^i - (q/p)^N}{1 - (q/p)^N} & \text{if } p \neq q. \end{cases}$$

$$\lim_{N \to \infty} u_i = \begin{cases} 1 & \text{if } p \leq q, \\ \left(\frac{q}{p}\right)^i & \text{if } p > q. \end{cases}$$

Let T be the time needed to reach state 0.

$$E[T] = i(N-i) \text{ if } p = q = \frac{1}{2}.$$

5.4 Success Runs

5.4.1 Two Outcomes Repeated Trials

$$r_i = r = 0;$$
 $q_i = q = \frac{1}{2};$ $p_i = p = \frac{1}{2}.$

5.4.2 Renewal Process

Model a light bulb's lifetime as a random variable ξ , where

$$P(\xi = k) = a_k \text{ for } k \in \mathbb{N},$$

and X_n as the age of the bulb in service at time n.

$$r_k = 0; \quad p_k = \frac{a_{k+1}}{a_{k+1} + a_{k+2} + \dots}; \quad q_k = 1 - p_k \text{ for } k \in \mathbb{Z}^{\geq}.$$