

Comparison of mixed model based approaches for correcting for population substructure with application to extreme phenotype sampling

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Abstract

1 Mixed models have been useful in correcting for confounding due to popu-
2 lation stratification and hidden relatedness in genome wide association stud-
3 ies. This class of models includes linear mixed models (LMM) and generalised
4 linear mixed models (GLMM). Existing mixed model approaches to correct
5 for population substructure have been investigated with both continuous and
6 case/control response variables. However, they have not been investigated in
7 the context of ‘extreme phenotype sampling’ (EPS), where genetic covariates
8 are only collected on samples having extreme response variable values. In
9 this work, we compare the performance of existing mixed model approaches
10 (LTMLM, GMMAT, CARAT) with EPS data analysed as a binary trait. We
11 use simulation to estimate the type 1 error of all approaches when there is
12 confounding. Since linear mixed models are commonly used even with binary
13 traits, we also analysed the data using a LMM (GEMMA). (Describe results
14 here)

Keywords: Population stratification, Extreme phenotype samples, Mixed models, type 1 error.

15 1. Introduction

16 In genetic studies involving human populations, researchers are interested
17 in determining how genetic variation contributes to diseases. Genome Wide
18 Association Studies (GWAS), which involve genotyping a large number of
19 individuals at hundreds of thousands of genetic markers have been useful for
20 discovering the relationships between common variants and complex diseases.
21 Recently, rare variants have been identified as important genetic factors con-
22 tributing to the risk of disease and human traits. . Exome sequencing has reference??
23 been used to discover rare variation in the human genome; although costs
24 have reduced, it remains a relatively expensive technique . Therefore study find a
25 designs that are powerful at lower sample sizes are advantageous. refer-
ence

26 An example of a cost saving design is extreme phenotype sampling: a
27 design where genotyping or sequencing is only done on individuals in the
28 tails of the phenotype distribution. The use of this study design can be
29 traced to the work of [19] where it was used in mapping quantitative trait
30 loci (QTLs) during linkage analysis. Extreme phenotype sampling has since
31 found other uses beyond linkage analysis as other authors have adapted the
32 study beyond linkage analysis. Still in linkage analysis, [6] used EPS and
33 advised that the cutoffs shouldn't be more than the upper and lower 25th
34 percentile. In association studies, [27] used extreme selection technique to
35 test for the association between a genetic variant and intelligence quotient
36 and [2] assessed the association between general cognitive ability as a behav-
37 ioral trait and variation in candidate genes. [36] explored the power of the
38 study when using extreme samples compared to the whole population. In
39 rare variant study, [11, 17, 26], EPS has been shown to have sufficient power

40 to detect rare variants.

41 As with all population based genetic association designs, extreme pheno-
42 type sampling is prone to confounding by population structure or stratifica-
43 tion. Difference in allele frequencies among members of a strata or subgroup
44 in the population may lead to confounding if there are differences in the
45 phenotype distribution between the subgroups. Confounding is known to
46 leading to spurious associations and an inflation of the type 1 error, which
47 has led to a development of methods that can correct for the effects of pop-
48 ulation stratification. The earliest methods includes the Genomic control
49 method of Devlin et al. [7] and the STRUCTURE approach of Pritchard et
50 al. [31]. Principal components (PC) based corrections, implemented in the
51 program EIGENSTRAT [29] have also been successfully applied in a number
52 of studies [29, 25]. Very recently, mixed models have become popular due
53 to their robustness in tackling other sources of confounding in the study,
54 in particular cryptic relatedness[30]. Over the years, an impressive number
55 of exact and approximate LMM methods have been developed for use in
56 genetic association studies [18, 22, 41]. Each of these methods incorporate
57 different approaches for making LMM-based analyses feasible at the genome
58 wide level.

59 However in genetic studies involving humans, the phenotype of interest
60 is often a binary trait, which can be obtained from case-control and cohort
61 study designs, for example.

62 Just like continuous traits, binary traits have also been analysed using
63 linear mixed models [9, 32, 35]. These methods have used an additive poly-
64 genic model which allows for transformation of the parameters of the linear

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65 model and the logistic model. This practice has resulted in a loss of power
66 as the mean and variance behaviour of the binary trait is ignored [14].

67 Earlier studies that aimed to correct for confounding rates in binary trait
68 have relied on the method of [28] that derived a direct relationship between
69 linear models and logistic regression. In particular, the authors justified the
70 application of a linear mixed model to binary data by introducing a way of
71 transforming the effect size estimates from the linear to the log-odds scale
72 which is the natural scale by which case-control data is measured. Although
73 widely applied to binary traits, the LMM assumes a continuous phenotype
74 where it is reasonable to assume that the trait has a constant residual vari-
75 ance. However, for binary traits in the presence of covariates, this assumption
76 is not valid; therefore fitting a binary response with mixed models may fail
77 to correct the type 1 error rate [5].

78 Mixed model approaches that are applicable under binary traits have
79 also been developed. Example of an applicable method are based on the use
80 of the liability threshold model that associates with each individual a nor-
81 mally distributed latent variable known as the liability. These methods have
82 been implemented in the softwares LTMML [12] and LEAP [37] and offers
83 an attractive method for association testing and case-control ascertainment
84 in case/control studies. These methods estimates the latent liabilities and
85 tests for association using these estimates. While LTMML tests for associ-
86 ation using the posterior mean liabilities, LEAP uses a maximum posterior
87 estimation. Another suitable method proposed for analysing binary traits is
88 based on the generalised linear mixed model (GLMM). Specifically, the GM-
89 MAT uses the logistic mixed model and first fits a null model for all SNPs in

the study and uses this model to compute score test statistics for testing the association between the binary traits and the genetic variant. Another binary trait association method known as CARAT uses a retrospective case-control analysis method to account for the analysis of binary traits and covariates. We desire to state here that all these methods have been examined in cases of population structure.

In this work, we aim to accomplish two goals. First, we present an overview and comparison of methods available for analysing binary traits with or without covariates adjustments using liability models and mixed models. Secondly, we investigate their performance when the binary data comes from an EPS design. We also include an LMM approach, which treats the phenotypes as if it were continuous, in our comparison. Here, each of the extremes will be treated as a different category. This is motivated by the fact that mixed model based approaches for correcting confounding has not been tested in the context of the EPS. We focus on whether these methods adequately correct the type 1 error rates due to confounding under an extreme phenotype study design. Finally, we also compare the approaches on a real dataset.

2. Material and Methods

In this section, we give a brief overview of the mathematical formulations of some of the general methods being considered in this paper.

2.1. Linear Mixed Models

The linear mixed model for a vector of response values y is usually represented as a sum of fixed and random effects and an error term. Specifically,

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114 we can represent a standard LMM by the equation:

$$y = X\beta + Zb + \epsilon \quad (1)$$

115 such that y is a $n \times 1$ vector of response variables (continuous or binary), $X_{n \times p}$
116 denotes the design matrix of known covariates, β is a vector of unknown
117 regression coefficients also known as the vector of fixed effects, $Z_{n \times p}$ is a
118 known matrix, b is a vector of random effects and ϵ is a vector of random
119 errors. Usually in a regression analysis, b and ϵ are unobservable quantities
120 that are assumed to be uncorrelated with a mean 0 and known variance. We
121 represent the variances of b and ϵ respectively as $\text{var}(b) = G$ and $\text{var}(\epsilon) = R$.
122 Hence, $b \sim \mathcal{N}(0, G)$ and $\epsilon \sim \mathcal{N}(0, R)$. The simple linear model is different
123 from the LMM equation in (2) through the inclusion of the random effects
124 components Zb . This enables us to specify a rich class of flexible models
125 that have been found to be important in genetic studies. They are mainly
126 applied in association testing between a genetic variant and a trait of interest,
127 estimating the narrow sense heritability [1], correcting for confounding [23]
128 and phenotype prediction [1]. Fitting LMMs involves evaluating the random
129 effects known as the variance components. This measures the correlation
130 between individuals.

131 In using LMMs for genetic analysis, the confounding effect is fit as a
132 fixed effect while the random effects is represented as a genetic relationship
133 matrix (GRM). This represents the pairwise genetic similarity between pairs
134 of individuals in the study. The equation is given by:

$$y = X\beta + Zb + \mu + \epsilon \quad (2)$$

Here, we represent y as the vector of phenotype values that is assumed continuous, X is the genetic variant being studied, β is the genetic effects, Z is a matrix of covariate values and b is the covariate effects. The first two terms of (2) corresponds to the fixed effects and μ is used to represent the effect of population structure in the data. Using the expressions of the mean and variance, we can then represent the distribution of y as:

$$\begin{aligned} E(y) &= X\beta + Zb, \text{ Var}(y) = \sigma^2 A + \sigma^2 I \\ y &= \mathcal{N}(X\beta + Zb, \sigma^2 A + \sigma^2 I) \end{aligned} \tag{3}$$

We can infer from (3) that the matrix μ imposes a sort of covariance structure on y in the form of A and this forms the basis of using LMM to correct for confounding in GWAS. In order to carry out mixed model analysis in GWAS, there is need for large sample sizes in order to achieve sufficient statistical power. Unfortunately, with increase in sample sizes there is the burden of computational complexity that increases cubically with the number of individuals [40] in the model. This has motivated several approximate LMM methods designed to increase the speed of LMM computations and inturn make large scale GWAS feasible.

2.1.1. Generalised Linear Mixed Models

Given the vector of random effects b and the independent responses $y_1 \dots y_n$, we define the generalised linear mixed model as the conditional distribution of y_i given b using the exponential family of distributions. The probability density function $f_i(y_i|b)$ is given as:

$$f_i(y_i|b) = \exp\left\{\frac{y_i\varphi - b^*(\varphi)}{a_i(\phi)} + c_i(y_i, \phi)\right\} \tag{4}$$

where $b^*(.), a_i(.), c_i(.,.)$ are known functions, ϕ is a dispersion parameter which may or may not be known, φ is a quantity which is associated with the conditional mean $\mu_i = E(y_i|b)$, which is also associated with a linear predictor $\eta = x_i\beta + z_ib$. x_i and z_i are known vectors and β is a vector of unknown parameters i.e the fixed effects. Since the distribution of y is not normal, the mean μ_i is related to the linear predictors via a link function $g(.)$ such that

$$g(\mu_i) = \eta_i.$$

155 Unlike linear regression models where the variance of the observation is a
 156 constant, the variance of b depends on a vector of unknown variance compo-
 157 nents i.e $b \sim \mathcal{N}(0, G)$ such that the covariance matrix G depends on a vector
 158 θ of unknown variance components.

As a special case of the GLMM, we consider the mixed logistic model defined for binary responses $y_1 \dots y_n$ which are conditionally independent bernoulli and $p_i = P(y_i = 1|b)$, then

$$\text{logit}(p_i) = x_i\beta + Z_ib$$

159 is the logistic mixed model, and x_i, z_i are as defined above. The link func-
 160 tion here is canonical given by $g(\mu) = \text{logit}(\mu)$ and the dispersion parameter
 161 $\phi = 1$. In obtaining the parameters of estimation in a GLMM model, the
 162 traditional methods of maximum likelihood estimation and restricted max-
 163 imum likelihood are not of great use here. This is because the likelihood
 164 function for a full glmm model with random effects usually involves high di-
 165 mensional integrals with no closed form expressions, hence the need for spe-
 166 cialized methods. A common approach has been to use numerical approaches

167 like the Laplace transforms and penalized likelihood methods. The laplace
 168 approximation uses approximate integrals to find a gaussian approximation
 169 to the conditional distribution of a set of variables [16]. The parameters of
 170 the GLMM can now be obtained using traditional or restricted likelihood
 171 by considering the Laplace approximations as the true likelihood. Under a
 172 more general framework, the laplace transforms have been used in another
 173 method known as the Penalized Quasi-likelihood Estimation (PQL). The use
 174 of PQL was proposed by Breslow and Clayton (1993) has been adjudged the
 175 most popular among the approximations to the likelihood function. PQL ap-
 176 proximates the high dimensional integrals found in GLMMs with the laplace
 177 approximation such that the approximated likelihood function is a normal
 178 distribution.

179 *2.2. Liability Threshold Models*

180 The liability threshold model (LTM) associates with every individual i in
 181 a population a latent variable known as the liability whose scale is regarded
 182 as arbitrary but can be assumed normal with a mean 0 and variance 1. We
 183 define a variable t , known as the threshold for a particular trait as the point
 184 on the scale of liability above which all individuals are affected and below
 185 which all are normal. We can regard these two divisions as cases and controls
 186 respectively. Hence the distribution can be regarded as:

$$Pr(Y = 1|Z) = \begin{cases} 1 & \text{if } Z > t \\ 0 & \text{otherwise} \end{cases} \quad (5)$$

187 where Y is the vector of liability values and Z is the normal cumulative distri-
 188 bution function. In order to generate the classical liability threshold model,

189 the relationship between observations on the observed risk scale and liabilities
 190 on the unobserved continuous scale is modeled using a probit transformation
 191 [10, 20] and can be written as:

$$I = \mu 1_N + g + \epsilon \quad (6)$$

192 where I is a vector of the liability phenotypes such that $I \sim \mathcal{N}(0, 1)$, g
 193 is the vector of random additive genetic effects on the liability scale with
 194 distribution $\mathcal{N}(0, \sigma_g^2)$ and ϵ is the vector of residual error [20]. Since the total
 195 phenotypic variance for the liability is equal to 1, the heritability (liability)
 196 defined as the proportion of total variance that is due to genetic factors is
 197 given as $h_L^2 = \sigma_g^2$. Liability threshold models make use of the available disease
 198 prevalence information in the data and this is given by the expected value
 199 of the phenotype values i.e $K = E(Y)$. The liability threshold models have
 200 been used in case-control studies to correct the associated loss in power as a
 201 result of ascertainment bias [12].

202 2.3. *Extreme Phenotype Sampling (EPS)*

203 The term selective or trait dependent sampling is used to denote geno-
 204 typing those individuals whose phenotypic values are in the extremes of the
 205 phenotype distribution. This idea was motivated by the fact that individu-
 206 als in the extremes are more likely to provide more linkage information than
 207 others and generally, these individuals are those whose genotypes can be
 208 clearly inferred from their phenotypes (source??). Although not a popular
 209 population sampling design as case control or cohort studies, sampling from
 210 the extremes is not a new design. It was first used by Lander and Botstein
 211 [19] in animal breeding studies to map QTLs. Subsequently, its use have

212 been explored both in GWAS (sources) and candidate gene association stud-
 213 ies. Recently, it has been proposed as a cost effective design compared to
 214 exome sequencing in the detection of rare variants Barnett et al. [3] with
 215 high power compared to random sampling Emond et al. [8]. Peloso et al. [26]
 216 have used EPS design in detecting rare variants and compared the associa-
 217 tion in extreme samples with a population based random sample. In extreme
 218 phenotype sampling, the selection depends on the phenotype hence standard
 219 statistical methods are not applicable. Perhaps the most common method
 220 involves treating the extreme groups as a binary trait and applying com-
 221 mon methods of assessing associations like the chi-square tests and logistic
 222 regression. Although valid, Lin et al. [21] stated that these methods are not
 223 optimal as the initial continuous trait values are ignored [13]. Other methods
 224 of analysing EPS samples includes likelihood based methods of Huang and
 225 Lin [13] and Lin et al. [21]. Despite the wide use of EPS design in association
 226 testing and rare-variant analysis, the effects of population stratification have
 227 not been extensively investigated. Panarella and Burkett (2019) investigated
 228 this concern using principal components analysis (PCA).

229 **3. Comparing some existing Approaches for Binary data**

230 In human genetic studies, the phenotype of interest is a binary trait (dis-
 231 ease status) obtained from case control sampling or cohort studies. Case-
 232 control studies samples diseased individuals (cases) from a study and a com-
 233 parable group of individuals from the same population who are free of the
 234 disease to serve as controls. In such case-control kind of studies, we are inter-
 235 ested in the problem of association testing with a known causal variant while

236 accounting for the effects of population structure and/or covariates [15]. The
 237 use of mixed model based approaches have been extensively explored in cor-
 238 recting the effects of population stratification and other unknown sources of
 239 population structure. These methods have all being applied to quantitative
 240 data as application to binary traits have resulted in largely inflated type 1
 241 error rates and loss in power. This is because the assumption of a constant
 242 residual variance for all the individuals in the sample might not hold for bi-
 243 nary traits in the presence of covariates [5]. Furthermore, a correct analysis
 244 of binary traits using mixed model approaches should include methods that
 245 are able to account for the selective sampling. Early approaches that have
 246 used mixed models in analysing binary traits have used the method of Piri-
 247 nen et al. [28] that assumes that in the absence of population stratification,
 248 linear models can be approximated by a logistic regression. Recent methods
 249 for the analysis of binary traits are methods based on the liability threshold
 250 models. Liability models estimates the model parameters for each associated
 251 genetic variants while also accounting for the case control ascertainment bias.
 252 Examples of such methods are the LTMLM, LTSOFT and CARAT. ROAD-
 253 TRIPS is a binary-trait association testing method that accounts for the
 254 population structure in the data using association statistics that have been
 255 adapted to cases where the population structure is known [34]. However, it
 256 is not suited for association testing between a single causal variant and the
 257 trait of interest. Very recently, the use of generalised linear mixed models
 258 have been explored in binary-trait association testing. GLMMs leverages the
 259 advantages of generalised linear models (GLMs) and linear mixed models
 260 (LMM) so that we are able to analyse binary trait data without the unreal-

261 istic assumption that the covariates have a constant residual variance. This
262 was implemented as the logistic mixed model in a tool known as GMMAT
263 [5].

264 Similar to the methods that have been primarily designed for analysing
265 quantitative data, quite a number of these binary methods have all recently
266 appeared in research and the differences or similarities between these meth-
267 ods have not been clearly elucidated. Here, we undertake a review of these
268 methods suitable for association testing between a binary-trait and a genetic
269 variant of interest. These methods can be classified broadly into three: (i) ap-
270 proaches using liability threshold models (LTMLM, LEAP, LTSCORE) , (ii)
271 mixed model approaches (CARAT, GMMAT) and (iii) association statistics
272 that have incorporated cases of completely unknown or partially unknown
273 population structure (ROADTRIPS, GCAT (Song et al.)).

274 • Liability threshold models (LTM) have been proposed as a valid ap-
275 proach to tackle the effects of ascertainment in case-control studies. In
276 ascertained case-control studies, cases are usually oversampled relative
277 to the disease prevalence leading to loss of power when linear mixed
278 models are used. Weissbrod et al. [37] stated that the loss of power
279 was due to the violation of several model assumptions one of which in-
280 cluded the dependence between the candidate SNPs and the SNPs used
281 to estimate kinship. Although LMMs are able to resolve the effects of
282 confounding in genetic association studies, their use for binary traits
283 association leads to a different form of confounding. The population
284 stratification in ascertained case control studies is as a result of the un-
285 equal case-control ratios from different sampling schemes which results

286 in unequal variances of the binary traits [5]. Using the liability thresh-
 287 old models involves computing liability scores for each individual to
 288 be used in testing the association between a phenotype of interest and
 289 genetic variants. In this manner, the LTM is able to directly represent
 290 the case-control phenotype while taking into account the ascertainment
 291 bias [37]. Although an attractive method, the use of liability thresh-
 292 old models is computationally expensive thus rendering whole genome
 293 association tests infeasible. To harness the attractive nature of LTM,
 294 a number of approaches have been developed. One of the earliest ap-
 295 proaches includes the LTSCORE method of Zaitlen and Kraft [38]. By
 296 introducing external prevalence data into the liability threshold model,
 297 LTSCORE is able to account for the study design and disease preva-
 298 lence and at the same time test for association with previously identi-
 299 fied causal SNPS using a linear regression. A similar method by Zaitlen
 300 et al. [39] computes the liability estimates and thereafter tests for the
 301 association using the EIGENSTRAT method [29]. Another liability
 302 threshold method known as LEAP (Liability Estimator as a Pheno-
 303 type) Weissbrod et al. [37] computes liability estimates conditional on
 304 the phenotypes, genotypes and disease prevalence on the entire genome
 305 and tests for association using a LMM method. Unlike the methods
 306 of Zaitlen and Kraft [38] and Zaitlen et al. [39], LEAP computes the
 307 liabilities using the whole genome and tests for association with the
 308 maximum a posteriori (MAP) estimate. Similar to LEAP is the recent
 309 method by Hayeck et al. [12] known as the liability threshold mixed
 310 linear model (LTMLM). LTMLM computes the posterior mean liabil-

ities of all the individuals under a liability threshold model and tests for association using a chi-square score statistic. The posterior mean liabilities of the individuals are computed dependent on the individual’s case control status, every other individuals’ case-control status and the genetic relationship matrix. Its performance was benchmarked across other existing mixed model methods and was found to have a well-controlled false positive rate.

- Generalised linear mixed models (GLMMs) have been used in genetic studies involving binary traits due to their ability to take into account the erroneous assumption made in the use of LMM for binary traits. These class of models can be viewed either as an extension of generalised linear models [24] or as an extension of LMM. When viewed in the former way, GLMMs are able to model the distribution of the response variable as a function of non-normally distributed variables. In the other viewpoint, GLMMs can be seen as an extension of LMMs when covariates have been included in the model. The inclusion of covariates implies that the assumption of constant variance made for the random effects in the linear mixed model no longer holds and there is need for an alternative construct to help model the covariate variability. Hence, for binary traits in the presence of covariates, the use of LMMs to test for the association between a causal SNP and a trait of interest will lead to type 1 errors. In the use of GLMMs to model non-normally distributed response variables, the logistic mixed models: a special case of GLMMs have been used to analyse binary traits and have been found to offer better correction of the type 1 error rates

compared to logistic regression and ordinary linear models. Despite this, the logistic mixed models have not been seen in widespread use in GWAS due to the computational complexity involved in fitting logistic mixed models for large scale genetic variants. Chen et al. [5] proposed a GLMM tool implementing a logistic mixed model known as GMMAT for large scale GWAS. GMMAT firsts fits a null logistic mixed model including as fixed effects only the covariates while the random effects are used to account for residual population stratification and relatedness. This fitted null model which is the same for all genetic variants in the study is then used in the test for association between a genetic variant and phenotype via a score test. The use of just one null model for testing all genetic variants greatly simplifies the model compared to fitting the full logistic mixed model for a large GWAS. Another mixed model method is known as CARAT (Case Control retrospective association test). It is a binary traits testing approach which accounts for relevant covariate information and control for population structure. The response variable is modeled by using a mixed effects quasi likelihood approach which exploits the binary nature of the trait. Similar to the GMMAT approach, CARAT does not require the knowledge of disease prevalence. For large scale genetic association studies, the use of estimating equations and a score test is used. In assessing the score test statistic, under the null model, the genotypes are regarded as random conditional on the phenotype and covariates [15].

- ROADTRIPS - a binary trait association testing tool proposed by Thornton and McPeck [34]. It can be regarded as an extension of

361 a collection of association statistics that have been used in traditional
362 case control testing in the presence of known structure to the context
363 of unknown or partially known structure. These include the corrected
364 χ^2 , armitage trend tests, W_{QLS} [4] and M_{QLS} [33] tests.

Criteria		LTMLM	LEAP	GMMAT	CARAT	ROADTRIPS
Type of model		Liability threshold model with focus on ascertained case-control studies assuming known disease prevalence.	Liability threshold model with focus on ascertained case-control studies assuming known prevalence.	Mixed model (retrospective and prospective)	Retrospective mixed model	Association Statistics
	Model Approach	Liabilities computed using the whole genome	Liabilities computed using the whole genome	Logistic Mixed model via Penalised Quasi likelihood.	Estimating equation approach	-
	GRM	Whole genome excluding the candidate SNP(s)	Exclude SNPs involved in liabilities estimation.	Whole genome excluding the candidate SNP(s)	Total number of SNPs.	-
	Test for association	Posterior mean liabilities in a chi-square score test framework	Liability estimates tested for association via a standard linear mixed model.	Score test or wald tests.	Quasi-likelihood in a score test framework	-
	Covariates	Program does not allow for covariate adjustment	-	Allows for covariate adjustment	Allows for covariate adjustment	No adjustment for covariates or polygenic additive effects

Table showing main features of some methods suitable for analysing binary response variable for the purpose of comparison.

368 4. Application of LTMLM and GMMAT to Extreme Phenotype 369 Sampling design.

370 In analysing samples from the extremes, the most common method has
371 been to treat the two extreme categories as binary traits [8] and use methods
372 suitable for categorical data to assess the association between a causal variant
373 and the phenotype of interest. With the advent of more suitable methods
374 for analysing binary traits such that the sampling is taken into account, it is
375 worthwhile to explore the performance of some of these identified methods
376 from the literature in correcting the type 1 error rate when we assume that
377 the data comes from the extremes. To this end, we treat the two extremes
378 which make up the whole extreme population as either cases or controls in
379 a typical prospective case-control design. These methods are the liability
380 threshold mixed linear model (LTMLM) and the generalised mixed model
381 association test which incorporates the liability threshold model and the
382 generalised linear mixed model respectively. We used these two methods on
383 simulated extreme phenotype data samples to test for the association between
384 a causal SNP and the phenotype of interest. Although these are not the only
385 existing methods, we choose these methods as a representative of the broad
386 categories of the methods we discovered. Also, we used methods which are
387 capable of performing candidate gene study and not GWAS. We didn't use
388 ROADTRIPS in our simulation experiments as the ROADTRIPS method
389 was not suitable to carry out an association between a single candidate gene
390 and the simulated phenotype values. In all these methods used, we were
391 mainly interested in assessing how each method worked in correcting the
392 type 1 error rate as a result of population stratification. Here, we give a little

insight into the LTMLM and GMMAT methods and describe the simulation study in the section that follows.

1. LTMLM (Liability Threshold Mixed Linear model) [12]): The LTMLM association statistic and software was proposed to control the false positive rate in ascertained case-control studies. The method was compared with existing mixed model methods and was found to have a well controlled false-positive rate for diseases with a low prevalence rate in the population. The LTMLM association statistic is given as a chi-square score statistic computed from the posterior mean liabilities under the liability threshold model. Here, for each individual in the study the posterior mean liability is conditional on the case-control status of that individual, the case-control status of every other individual in the study i.e disease prevalence and on the genetic relationship matrix computed from all the SNPs excluding the candidate SNP in the study. For the multivariate model considered, the PMLs are estimated using a multivariate Gibbs sampler: a MCMC algorithm for multivariate random sampling. The phenotypic covariance matrix on the liability scale is based on the GRM and the heritability estimate computed using the Haseman-Elton regression on the case-control phenotypes and then transformed to the liability scale. In order to account for the ascertainment, the liabilities and genotypes are jointly modelled using a retrospective model.
2. GMMAT (Generalised Mixed Model Association Test): The GMMAT tool for association testing firsts fits a null generalised linear mixed model that includes just the covariates and random effects to account

for population structure and other forms of confounding (cryptic or family relatedness.) Unlike LTMLM, the GMMAT tool can be adapted into testing for association both for GWAS and candidate gene studies. Association tests are usually measured using score tests for every genetic variant and wald tests to obtain the effect estimates of each genetic variant for candidate gene studies. By specifying a particular link family such as binary or gaussian, GMMAT will perform mixed model based association tests when the response variable is either categorical such as disease status or quantitative. For a candidate gene study such as we are interested in, the GMMAT method is represented as the GLMM model given in Chen et al. [5].

$$\eta_i = g(\mu_i) = X_i\alpha + G_i\beta + b_i \quad (7)$$

Since the GLMM follows the exponential family of distributions, given the random effects b_i , which are conditionally independent with the responses y_i , the mean and variance is given respectively by $E(y_i|b) = \mu_i$ and $Var(y_i|b) = \Phi a_i^{-1} v(\mu_i)$ respectively. Here Φ is the dispersion parameter defined for exponential families, $v(.)$ is the variance and a_i are known weights. All other parameters are as defined for GLMMs in previous sections. For binary traits, GMMAT uses a logistic mixed model: GLMM which assumes a bernoulli distribution for the responses and a logit link function. The GMMAT tool is implemented as an R package.

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