

# Chapter 6 part 2

## Jointly Distributed Random Variables

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MATH 241

# Outline

- 1 Sums of independent random variables
- 2 Conditional distributions: discrete case
- 3 Conditional distributions: continuous case

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## Sums of continuous random variables

If  $X, Y$  have a joint density  $f(x, y)$ , then  $X + Y$  has the following density

$$\begin{aligned} f_{X+Y}(z) &= \int_{-\infty}^{\infty} f(x, z-x) \, dx \\ &= \int_{-\infty}^{\infty} f(z-y, y) \, dy \end{aligned}$$

If  $X$  and  $Y$  are independent, we can use the convolution formula

$$\begin{aligned} f_{X+Y}(z) &= \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) \, dx \\ &= \int_{-\infty}^{\infty} f_X(z-y) f_Y(y) \, dy \end{aligned}$$

Cdf  $F_{X+Y}$  is called the *convolution* of the distributions  $F_X$  and  $F_Y$ .

# Why?

$$\begin{aligned}
 F_{X+Y}(z) &= \iint_{x+y \leq z} f_{X,Y}(x,y) \, dy \, dx \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f_{X,Y}(x,y) \, dy \, dx \\
 f_{X+Y}(z) &= \frac{d}{dz} F_{X+Y}(z) \\
 &= \int_{-\infty}^{\infty} \left\{ \frac{d}{dz} \int_{-\infty}^{z-x} f_{X,Y}(x,y) \, dy \right\} dx \\
 &= \int_{-\infty}^{\infty} f_{X,Y}(x, z-x) \, dx
 \end{aligned}$$

Textbook page 239 gives a derivation when  $X$  and  $Y$  are independent.

## Question

Suppose  $X, Y \stackrel{\text{ind}}{\sim} \text{Exp}(\lambda)$ . What distribution does  $X + Y$  have? Note that the pdf of exponential of  $X \sim \text{Exp}(\lambda)$  is:  $f(x) = \lambda e^{-\lambda x}, x > 0$ .

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Range of  $X + Y$  is  $(0, \infty)$ . For any  $z > 0$ ,

$$f_{X+Y}(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z - x) dx$$



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$$\begin{aligned} f_{X+Y}(z) &= \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx \\ &= \int_{\textcolor{red}{0}}^{\textcolor{red}{z}} \lambda e^{-\lambda x} \lambda e^{-\lambda(z-x)} dx \\ &= \int_0^z \lambda^2 e^{-\lambda z} dx \end{aligned}$$

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 f_{X+Y}(z) &= \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx \\
 &= \int_0^z \lambda e^{-\lambda x} \lambda e^{-\lambda(z-x)} dx \\
 &= \int_0^z \lambda^2 e^{-\lambda z} dx \\
 &= \lambda^2 x e^{-\lambda z} \Big|_0^z
 \end{aligned}$$

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 &= \lambda^2 x e^{-\lambda z} \Big|_0^z = \lambda^2 z e^{-\lambda z} \\
 f_{X+Y}(z) &= \begin{cases} \lambda^2 z e^{-\lambda z} & \text{if } z > 0 \\ 0 & \text{otherwise} \end{cases}
 \end{aligned}$$

# Sum of independent random variables

Random variables  $X$  and  $Y$  are independent, then

$X$	$Y$	$X + Y$
$N(\mu_1, \sigma_1^2)$	$N(\mu_2, \sigma_2^2)$	$N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$
$\text{Poi}(\lambda_1)$	$\text{Poi}(\lambda_2)$	$\text{Poi}(\lambda_1 + \lambda_2)$
$\text{Bin}(n_1, p)$	$\text{Bin}(n_2, p)$	$\text{Bin}(n_1 + n_2, p)$

## Question

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Hint: find the distribution of  $W = -Y$  first.

Since  $g(y) = -y$  is monotonic and differentiable on  $\mathbb{R}$ ,

$$f_W(w) = f_Y(y) \left| \frac{dy}{dw} \right|$$



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Since  $g(y) = -y$  is monotonic and differentiable on  $\mathbb{R}$ ,

$$f_W(w) = f_Y(y) \left| \frac{dy}{dw} \right| = \frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{(-w-\mu_2)^2}{2\sigma_2^2}}$$

$$W \sim N(-\mu_2, \sigma_2^2)$$

Since  $X$  and  $W$  are also independent,

$$X - Y = X + W \sim N(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2)$$

## Recap

Random variables  $X$  and  $Y$  are independent if any real sets  $A, B \subset \mathbb{R}$ ,

$$P(X \in A, Y \in B) = P(X \in A)P(Y \in B)$$

Random variables  $X$  and  $Y$  are independent **if and only if**

- Cdf: for any  $x, y \in \mathbb{R}$

$$F_{X,Y}(x, y) = F_X(x)F_Y(y)$$

- For any  $x, y \in \mathbb{R}$ , the pmf / pdf

$$f_{X,Y}(x, y) = f_X(x)f_Y(y)$$

If  $X$  and  $Y$  are independent continuous random variables, then

$$f_{X+Y}(z) = \int_{-\infty}^{\infty} f_X(x)f_Y(z-x) dx$$

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## Discrete conditional distributions

- Recall that for any two events  $E$  and  $F$ , the conditional probability of  $E$  given  $F$  is defined by

$$P(E|F) = \frac{P(E \cap F)}{P(F)}$$

provided that  $P(F) > 0$ .

# Discrete conditional distributions

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provided that  $P(F) > 0$ .

- If  $X$  and  $Y$  are discrete random variables, we define the conditional probability mass function (pmf) of  $X$  given  $Y = y$  by

$$p_{X|Y}(x|y) = P\{X = x|Y = y\} = \frac{P\{X = x, Y = y\}}{P\{Y = y\}} = \frac{p(x, y)}{p_Y(y)}$$

for all values of  $y$  such that  $p_Y(y) > 0$ .

## Discrete conditional distributions

- Similarly for the conditional probability density function (cdf), we define it by

$$F_{X|Y}(x|y) = P\{X \leq x|Y \leq y\} = \sum_{x \leq a} p_{X|Y}(a|y)$$

- If  $X$  is independent of  $Y$ , then

$$\begin{aligned} p_{X|Y}(x|y) &= P\{X = x|Y = y\} \\ &= \frac{P\{X = x, Y = y\}}{P\{Y = y\}} \\ &= \frac{P\{X = x\}P\{Y = y\}}{P\{Y = y\}} \\ &= P\{X = x\} \end{aligned}$$

## Question

If  $X$  and  $Y$  are independent Poisson random variables with respective parameters  $\lambda_1$  and  $\lambda_2$ , calculate the conditional distribution of  $X$  given that  $X + Y = n$ .

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$$\bullet P\{X = k | X + Y = n\} = \frac{P\{X=k, X+Y=n\}}{P\{X+Y=n\}} = \frac{P\{X=k, Y=n-k\}}{P\{X+Y=n\}} \stackrel{ind}{=} \frac{P\{X=k\}P\{Y=n-k\}}{P\{X+Y=n\}}$$



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- We know that  $X + Y \sim P(\lambda_1 + \lambda_2)$  (sum of Poisson)

$$\begin{aligned} P\{X = k | X + Y = n\} &= \frac{e^{-\lambda_1} \lambda_1^k}{k!} \frac{e^{-\lambda_2} \lambda_2^{n-k}}{(n-k)!} \left[ \frac{e^{-(\lambda_1 + \lambda_2)} (\lambda_1 + \lambda_2)^n}{n!} \right]^{-1} \\ &= \frac{n!}{(n-k)! k!} \frac{\lambda_1^k \lambda_2^{(n-k)}}{(\lambda_1 + \lambda_2)^n} \\ &= \binom{n}{k} \left( \frac{\lambda_1}{\lambda_1 + \lambda_2} \right)^k \left( \frac{\lambda_2}{\lambda_1 + \lambda_2} \right)^{(n-k)} \end{aligned}$$

- That is, the conditional distribution of  $X$  given  $X + Y = n$  is a  $\text{Bin}(n, \lambda_1/(\lambda_1 + \lambda_2))$

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## Continuous conditional distributions

- With joint probability function of  $X$  and  $Y$  as  $f(x, y)$ , the conditional pdf of  $X$  given that  $Y = y$  (for  $f_Y(y) > 0$ ) is defined by

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

note that the event  $\{Y = y\}$  has probability 0; we just use it for conditioning.

- If  $X$  and  $Y$  are jointly continuous, then for any set  $A$

$$P\{X \in A|Y = y\} = \int_A f_{X|Y}(x|y)dx$$

This is defining conditional probabilities of events associated with one random variable when we are given the value of a second random variable.

## Continuous conditional distributions

- By setting  $A = (-\infty, a)$ , we can define the cdf of  $X$  given that  $Y = y$  by

$$F_{X|Y}(a|y) = P\{X \leq a|Y = y\} = \int_{-\infty}^a f_{X|Y}(x|y)dx$$

- And if  $X$  and  $Y$  are independent continuous random variables,

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{f_X(x)f_Y(y)}{f_Y(y)} = f_X(x)$$

This is the unconditional density of  $X$ .

## Question

Suppose the joint density of  $X$  and  $Y$  is given by

$$f(x, y) = \begin{cases} \frac{e^{-x/y} e^{-y}}{y} & \text{for } 0 < x < \infty, 0 < y < \infty \\ 0 & \text{otherwise} \end{cases}$$

Find  $P\{X > 1 \mid Y = y\}$ .

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Find  $P\{X > 1 \mid Y = y\}$ .

First, obtain the conditional density of  $X$  given  $Y = y$ ,

$$\begin{aligned} f_{X|Y}(x \mid y) &= \frac{f(x, y)}{f_Y(y)} \\ &= \frac{e^{-x/y} e^{-y} / y}{e^{-y} \int_0^{\infty} (1/y) e^{-x/y} dx} \\ &= \frac{1}{y} e^{-x/y} \end{aligned}$$

## Question

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First, obtain the conditional density of  $X$  given  $Y = y$ ,      Hence,

$$\begin{aligned} f_{X|Y}(x \mid y) &= \frac{f(x, y)}{f_Y(y)} \\ &= \frac{e^{-x/y} e^{-y} / y}{e^{-y} \int_0^{\infty} (1/y) e^{-x/y} dx} \\ &= \frac{1}{y} e^{-x/y} \end{aligned}$$

$$\begin{aligned} P\{X > 1 \mid Y = y\} &= \int_1^{\infty} \frac{1}{y} e^{-x/y} dx \\ &= -e^{-x/y} \Big|_1^{\infty} \\ &= e^{-1/y} \end{aligned}$$