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Does economic policy uncertainty predict the Bitcoin returns? An empirical investigation

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ABSTRACT

This paper analyzes the prediction power of the economic policy uncertainty (EPU) index on the daily Bitcoin returns. Using the Bayesian Graphical Structural Vector Autoregressive model as well as the Ordinary Least Squares and the Quantile-on-Quantile Regression estimations, the paper finds that the EPU has a predictive power on Bitcoin returns. Fundamentally, Bitcoin returns are negatively associated with the EPU. However, the effect is positive and significant at both lower and higher quantiles of Bitcoin returns and the EPU. In the light of these findings, the paper concludes that Bitcoin can serve as a hedging tool against uncertainty.

1. Introduction

Nakamoto (2008) introduced the Bitcoin, which is a digital currency and open-source online payment system. Bitcoin is fully decentralized without any central authority and control while depending on a sophisticated protocol. The supply of Bitcoin is limited to 21 million by the design of the protocol. Since its introduction, the market value of Bitcoin has grown rapidly. According to the data from <http://www.coindesk.com>, the market capitalization of Bitcoin has dramatically reached \$278 Billion from \$111 Billion for the period from November 15, 2017 to December 15, 2017. In addition to legal and technical debates, especially due to these record high performances, Bitcoin has hit the headlines recently.

Similarly, there is the rising interest in the literature by focusing on economic and financial determinants of the Bitcoin price. For instance, Dyhrberg (2016) shows that Bitcoin can be used as a hedging instrument against the stock market and the United States (U.S.) Dollar and it is a useful tool for both portfolio diversification and risk management.

Another strand of literature focuses on the efficiency of Bitcoin. Urquhart (2016) provides the supporting evidence on the inefficiency of Bitcoin market, but it is in the process of moving towards efficiency for the period from August 1, 2013, to July 31, 2016. Using the Bitcoin returns, Nadarajah and Chu (2017) show that the efficient market hypothesis is not valid. According to Bariviera (2017), the daily returns exhibit persistent (inefficiency) behavior until 2014, whereas the market is more informational efficient since 2014. Finally, the previous studies investigate the volatility of Bitcoin returns (Katsiampa, 2017), the informed trading (Feng et al., 2017), the price clustering (Urquhart, 2017), and the speculative bubbles (Cheah and Fry, 2015; Corbet et al., 2017), the transaction cost (Kim, 2017).

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To the best of our knowledge, Bouri et al. (2017) is the only study to examine the relationship between uncertainty and Bitcoin market. The authors explore whether Bitcoin can serve as a hedge against uncertainty that is measured by the first principal component of the Volatility Indexes (VIXs) of 14 developed and developing stock markets. According to their results, Bitcoin acts as a hedge against uncertainty.

In a similar vein, our paper examines the prediction power of the daily economic policy uncertainty (EPU) index on the daily Bitcoin returns. Indeed, uncertainties about the decisions of governments and regulatory bodies lead to decreases in the trust of investors to mainstream currencies and or to the entire economy especially after the global financial crisis of 2008–9. This is also the time when the Bitcoin was created. Therefore, by its nature, Bitcoin questions the effectiveness of standard economic and financial structures and the digital currencies are decentralized secure alternatives to the fiat currencies, especially during the times of economic and geopolitical unrest. Therefore, the changes in the EPU index can possibly affect Bitcoin returns. We find that Bitcoin can serve as a hedging tool against uncertainty at both lower and higher quantiles. In other words, an increase in EPU will help Bitcoin to meet what it promises and increase the attractiveness of Bitcoin.

The remainder of the paper is as follows. Section 2 describes our data, model, and methodology. Section 3 reports the results and discusses the findings. Section 4 concludes.

2. Data, model, and methodology

2.1. Data and empirical model

The paper considers the logarithmic returns of Bitcoin as the dependent variable for the period from July 18, 2010, to November 15, 2017. The number of observations is 2678. The starting date of the empirical analysis is due to the availability of the data and the daily frequency data are used. Following Katsiampa (2017), we obtain the data of the Bitcoin prices from <http://www.coindesk.com/price/>. We also use the daily EPU index in the U.S., which is developed by Baker et al. (2016).¹ Using the EPU in the U.S. is not only related to the daily data availability, but also Bitcoin prices are mainly quoted in the USD. The correlation between the logarithmic returns of Bitcoin and the EPU index is -0.014 . At this stage, our paper estimates the following empirical model:

$$\Delta \ln(BCP)_t = \alpha_0 + \alpha_1 \Delta \ln(EPU)_t + \varepsilon_t \quad (1)$$

Where $\Delta \ln(BCP)_t$ and $\Delta \ln(EPU)_t$ represent the daily logarithmic returns of Bitcoin prices and the EPU index values, respectively. ε_t is the error term.

2.2. Econometric methodology

The Bayesian Graphical Structural Vector Autoregressive (BGSVAR) model can provide the contemporaneous and the delayed causality between the Bitcoin returns (i.e. the response variable) and the EPU index (i.e. the predictor variable). An SVAR model can define the dependence/causality, as such:

$$Y_t = B_0 Y_t + \sum_{i=1}^p B_i Y_{t-i} + \sum_{i=1}^p C_i Z_{t-i} + \varepsilon_t \quad (2)$$

where $t = 1, \dots, T$ and p is the maximum lag order. Y_t and Z_t are the vector of the returns of Bitcoin and the EPU index, respectively. We can write the reduced form of Eq. (2), as such:

$$Y_t = A_1 X_{t-1} + \dots + A_p X_{t-p} + u_t \quad (3)$$

where $X_t = (Y_t, Z_t)' = (X_{1t}, X_{2t}, \dots, X_{nt})'$ is an $n = n_y + n_z$ dimensional time series; $B_i^* = (B_i, C_i)$, $1 \leq i \leq p$, are $(n_y \times n)$ matrices of the unknown coefficients; $A_0 = (I_{n_y} - B_0)$ is a $(n_y \times n_y)$ matrix; $A_i = A_0^{-1} B_i^*$, $1 \leq i \leq p$, are $(n_y \times n)$ the reduced-form lag coefficient matrices; and $u_t = A_0^{-1} \varepsilon_t$ is an $(n_y \times 1)$ independently and identically distributed reduced-form vector residual term with the zero mean and the covariance matrix Σ_{u_t} .

In this paper, we apply the BGSVAR model of Ahelegbey et al. (2016) in order to handle the problems of the misidentification of the system of equations and implausible restrictions assumption (Bouri et al., 2018). The BGSVAR model maintains two simple representations, namely the Contemporaneous Network (CN) and the Lagged Network (LN) causality structures.² In addition, Bouri et al. (2018) use the Bayesian Graphical model to predict the BRICS stock market returns using the VIX index as a predictor along with other financial and macroeconomic variables. Following the spirit of Bouri et al. (2018), we consider the EPU index as a potential predictor of the Bitcoin returns.

Furthermore, we run the Ordinary Least Squares (OLS) and the Quantile-on-quantile Regression (QQ) estimations to model the quantile of Bitcoin returns (including various frequencies) as a function of the quantile of the EPU index, which represents the each

¹ For the details of the EPU indexes, visit the website that is designed by Baker et al. (2016) (<http://www.policyuncertainty.com>).

² For more details about estimation and inference techniques, refer to Ahelegbey et al. (2016). Using the Markov Chain Monte Carlo (MCMC) process and the small-size networks methods, we estimate the LN component and the CN component. The MIN and the MAR structures respectively provide the posterior probabilities for the instantaneous and the lagged relationships between the logarithmic returns of Bitcoin and the EPU index.

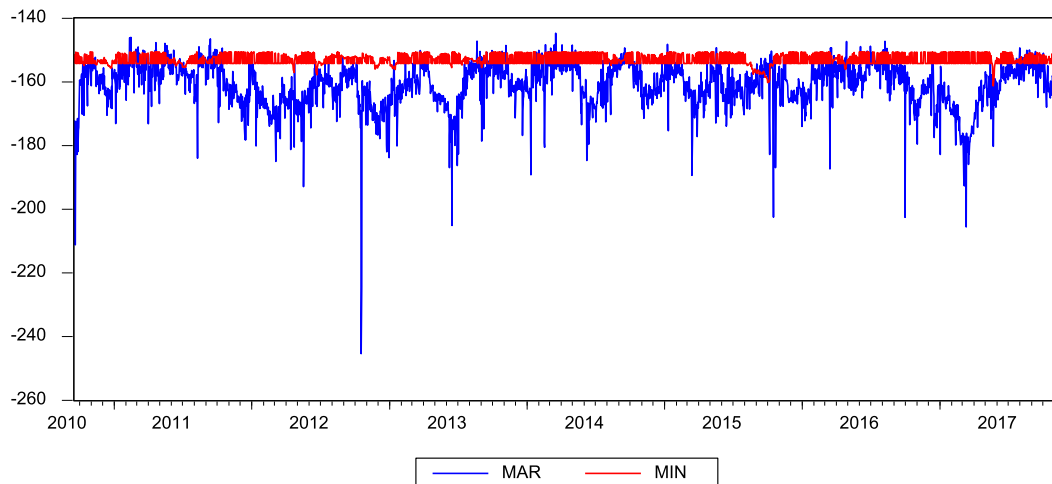


Fig. 1. Bayesian information criteria of contemporaneous and temporal dependence structures.

point of their distributions.³ We also analyze the lagged effects of the EPU on the Bitcoin returns.

3. Empirical results and discussion

We run the 60-day rolling window estimation technique with an initial sample period of 19 July 2010 to 15 September 2010 and 60-day rolling window estimation for the period from 16 September 2010 to 15 November 2017. In Fig. 1, we present the Bayesian Information Criteria (BIC) scores of the multivariate autoregressive (MAR) and the multivariate instantaneous (MIN) dependence structures over the sample period. The result shows that the MAR provides the best representation in accordance with the temporal dependence since the values of the MAR dependence are consistently smaller than the MIN dependence.

In Fig. 2, we further provide the percentage of the MAR links as obtained from the change in the EPU index to the Bitcoin returns for the MAR structures. Using the total (all) link, we observe that the period of highest interconnectedness is the year of 2012–13, and it is interesting to find the evidence of the structural breaks for the EPU index occurred on 1/26/2012 and 3/03/2013, while the breaks for the total linkage happened on 1/27/2012 and 9/09/2013. According to these findings, there is the lead-lag relationship for the breaks of the EPU index and the total linkage measures.

To examine the direction of this relationship, we provide the results of the OLS and the QQ estimations in Table 1. According to the results of the OLS estimations, the effects of the EPU and the lagged EPU are negative and statistically significant at the 1% level. The evidence implies that an increase in EPU will lead to a decrease in the Bitcoin returns. However, the effect turns to be positive and significant at the lower and higher quantiles. Therefore, Bitcoin can be used for hedging against uncertainty during the times of bull-market (i.e. higher quantiles). At this stage, investors on cryptocurrencies can use these findings and they can incorporate the daily U.S. economic policy uncertainty into their investment decisions. Investors can predict the returns of Bitcoin via information from this uncertainty measure.

Our findings provide potential implications for portfolio diversification and hedging (risk management). According to our results, Bitcoin can be an alternative instrument for hedging against uncertainty. Our results are in line with Bouri et al. (2017), that is the relationship between uncertainty and Bitcoin returns are mainly negative, but Bitcoin can be used for hedging against uncertainty during the times of bull-market. Bitcoin can be considered for portfolio diversification during the times of bear-market.

It is also noteworthy to note that the future path of the EPU can serve as a measure of the vulnerability of Bitcoin market. Since the Bitcoin market is still in the growth process, the policymakers in the U.S. should realize that the uncertain environment in their economic policies could affect the Bitcoin returns. In the context of investors, they should watch not only the natural existing uncertainty of cryptocurrencies, but also the EPU to deal with the potential risks in the future.

4. Conclusion

In this paper, we examined the prediction power of the EPU on the daily Bitcoin returns for the period from July 18, 2010 to November 15, 2017. Using the BGSVAR model as well as the OLS and the QQ estimations, we found that the EPU has a predictive power on the Bitcoin returns. Mainly, the Bitcoin returns are negatively associated with the changes in the EPU. However, Bitcoin can be used a hedging tool against uncertainty in extreme times of uncertainty since we also found that the effect is positive and significant at the higher quantiles.

Future papers can use other uncertainty measures to analyze their effects on cryptocurrency markets. Indeed, we need to enhance our knowledge of the potential determinants of cryptocurrency markets.

³ At this stage, we follow the QQ estimation methodology of Sim and Zhou (2015). For details, refer to Sim and Zhou (2015).

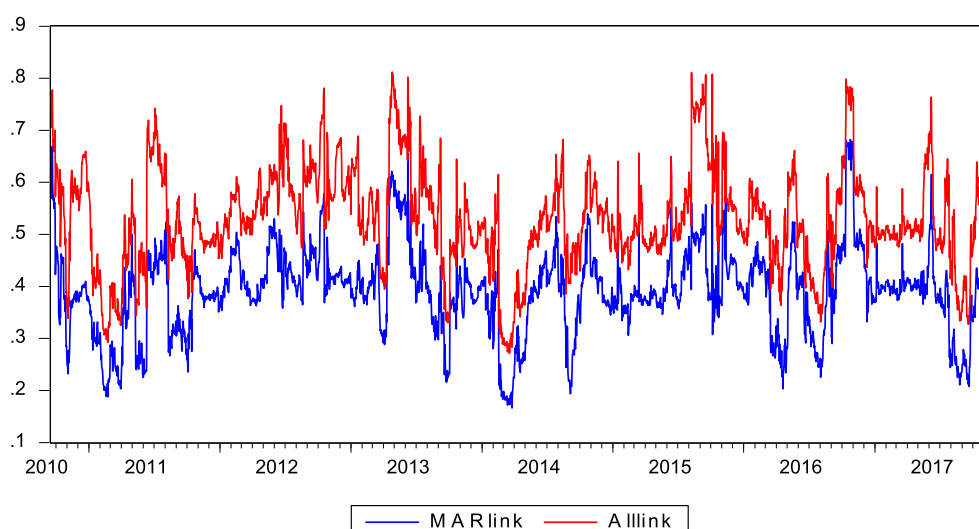


Fig. 2. Multivariate autoregressive model and the lead-lag linkage.

Table 1
Results of the OLS and the QQ estimations.

Quantiles	EPU	EPU(−1)	EPU(−2)	EPU(−4)	EPU(−8)	EPU(−12)	EPU(−24)
OLS	−0.16***	−0.18***	0.23	0.07	0.03	0.18	0.48
0.05	0.26	0.32	0.11	0.12	0.06	0.04	0.51
0.10	0.23*	0.19	0.49	0.43	0.25	0.22	0.67
0.15	0.19***	0.44	0.48	0.18	0.26	0.23	0.54*
0.20	0.10**	0.67	0.16	0.04	0.20	0.30	0.35
0.25	0.55**	0.09	0.06	0.01	0.19	0.12	0.29*
0.30	0.21	0.13	0.03	0.02	0.07	0.11	0.15
0.35	−0.06	−0.09	−0.02	−0.03	0.16*	0.07	0.14
0.40	−0.16**	−0.05	−0.01	−0.08	0.11	−0.03	−0.19
0.45	−0.17**	−0.04**	−0.02***	−0.04**	−0.03*	−0.05**	−0.17
0.50	−0.17**	−0.01	−0.01	−0.03	−0.04	−0.08	0.19
0.55	−0.28**	−0.06	−0.02	−0.06	−0.01	−0.08	0.19
0.60	−0.17	0.09	0.03	0.02	−0.03	0.11	0.18
0.65	0.07***	0.16	−0.09	0.12	−0.04	0.08	0.23
0.70	0.12	0.17	−0.21	0.06	−0.23	0.23	0.41*
0.75	0.53*	0.03	−0.16	0.11	−0.44**	0.22	0.20
0.80	0.12***	0.06	−0.29	0.12	−0.38*	0.17	0.10
0.85	0.19***	0.09	0.11	0.01	0.25	0.21	0.11
0.90	0.27***	0.24	0.14	0.27	0.07	0.52	0.43
0.95	0.23*	0.09	0.33*	0.81*	0.53	0.13*	0.15*

Note: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Supplementary materials

Supplementary material associated with this article can be found, in the online version, at [doi:10.1016/j.frl.2018.01.005](https://doi.org/10.1016/j.frl.2018.01.005).

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