

# Lecture 16

Recap...

## Two-Stage Least Squares:

Consider the model:

$$\begin{aligned}y_1 &= \alpha_0 + \alpha_1 y_2 + \alpha_2 x_1 + u \\y_2 &= \beta_0 + \beta_1 y_1 + \beta_2 x_2 + \beta_3 x_3 + v\end{aligned}$$

- Endog.:  $y_1, y_2$ ;
- Exog.: const.,  $x_1, x_2, x_3$ .

First question: Is it identified?

- In the first equation, the number of omissions ( $x_2, x_3$ ) is greater than the number of endogenous variables in the system ( $G$ ) minus one  $\rightarrow$  It is over identified.
- In the second equation, the number of omissions is equal to  $G - 1 \rightarrow$  It is identified.

Step 1: Find instrumental variables (IV) for  $y_1$  and  $y_2$ .

- Regress  $y_1$  on constant,  $x_1, x_2, x_3$ : find  $\hat{y}_1 = \hat{\lambda}_0 + \hat{\lambda}_1 x_1 + \hat{\lambda}_2 x_2 + \hat{\lambda}_3 x_3$ .
- Regress  $y_2$  on constant,  $x_1, x_2, x_3$ : find  $\hat{y}_2 = \hat{\gamma}_0 + \hat{\gamma}_1 x_1 + \hat{\gamma}_2 x_2 + \hat{\gamma}_3 x_3$ .
- $\hat{y}_1$  and  $\hat{y}_2$  are IVs (they depend only on the  $x$ 's').

Step 2: Estimate the system using the instrumental variables.

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## Multicollinearity

Say you have  $x_1, \dots, x_k$ .

If  $\exists \lambda_i (i = 1, \dots, k) : \sum_{i=1}^k \lambda_i x_i = 0 \rightarrow$  **Perfect** Multicollinearity.

- E.g.:  $x_1 = \frac{1}{\lambda_1} (\lambda_2 x_2 + \dots + \lambda_k x_k)$ .

If  $\lambda_1 x_1 + \lambda_2 x_2 + \dots + \lambda_k x_k = v_t$ , where  $v_t$  is a random error  $\rightarrow$  **Imperfect** Multicollinearity.

**Example:**

$x_2$	$x_3$	$x_3^*$
10	50	52
15	75	74
18	90	97
24	120	109
30	150	153

- $x_3 = 5x_2$  ;  $x_3^* = 5x_2 + v$ .

- Using both  $x_2, x_3$ : perfect multicollinearity (cannot regress).
- Using both  $x_2, x_3^*$ : imperfect multicollinearity (might regress).

Summarizing:

- If there is perfect multicollinearity:
  - $\hat{\beta}_i$  undefined;
  - $\text{Var}(\hat{\beta}_i) \rightarrow \infty$ .
- If there is imperfect multicollinearity
  - It is possible to estimate;
  - However,  $\text{Var}(\hat{\beta}_i)$  will increase (and the  $t$  statistics decrease) as the  $x$ 's are more and more correlated.

### Estimation under imperfect multicollinearity:

1. Unbiased, consistent (only perfect multicollinearity violates Gauss-Markov);
  - OLS is still BLUE (Best Linear Unbiased Estimator).
2. The std. errors are big, but correctly estimated.
3. It is hard to get a small std. error (a small  $t$ ) when the  $x$ 's are highly correlated.
  - A small number of observations will produce the same problem.
  - The solution is the same: Get more data!
4. Estimators can be sensitive to small changes in data.
  - That, too, can apply to small samples.

### Diagnostic:

- When individual  $t$ -statistics are insignificant, but the  $F$ -test of the joint significance is significant ( $R^2$  usually high).
  - For a model  $y = \beta_1 + \beta_2 x_2 + \beta_3 x_3 + u$ :
    - $t$ -tests:  $H_0 : \beta_2 = 0, H_0 : \beta_3 = 0 \rightarrow$  low  $|t|$ 's.
    - $F$ -test:  $H_0 : \beta_2 = \beta_3 = 0 \rightarrow F$ -statistic greater than critical value.
- High pair-wise correlation (say,  $\rho > 0.8$ ).
  - If some variable  $x_2$  has high correlation, not with other variables individually, but with a linear combination of them (say,  $ax_3 + bx_4$ ), that can be found by regressing  $x_2$  on  $x_3$  and  $x_4$ .
    - If  $\hat{a}$  and  $\hat{b}$  are significant, there is most likely multicollinearity.

### What to do?

1. Do nothing: it is a data (small sample) problem, not a model problem;
2. Collect more data;
3. Drop a variable;
4. Transform the data (log, square, differencing *etc*).

### LM test for adding variables:

For a **restricted** model  $y = \beta_1 + \beta_2 x_2 + \dots + \beta_k x_k + u$ .

- Test against **unrestricted** model:  $y = \beta_1 + \beta_2 x_2 + \dots + \beta_k x_k + \beta_{k+1} x_{k+1} + u$ .

Step 1: Estimate restricted model: save  $\hat{u}_i$ 's.

Step 2: Regress  $\hat{u}_i$  on *all* regressors:  $\hat{u}_i = \alpha_1 + \alpha_2 x_2 + \dots + \alpha_k x_k + \alpha_{k+1} x_{k+1} + v_i$ .

Step 3:  $N \cdot R^2 \sim \chi^2(\text{df})$ , where  $N \cdot R^2$  is the sample size times the coefficient of determination and df, the number of restrictions.