



CS 4375

Midterm Review: Part II

Rishabh Iyer
University of Texas at Dallas

Topics for the Midterm Exam



- Linear Regression
- Perceptron
- Support Vector Machines
- Nearest Neighbor Methods
- Decision Trees
- Bayesian Methods and Parameter Estimation
- Naïve Bayes
- Logistic Regression

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Maximum Likelihood Estimation (MLE)

- **Data:** Observed set of α_H heads and α_T tails
- **Hypothesis:** Coin flips follow a Bernoulli distribution
- **Learning:** Find the “best” θ
- **MLE:** Choose θ to maximize probability of D given θ

$$\begin{aligned}\hat{\theta} &= \arg \max_{\theta} P(\mathcal{D} \mid \theta) \\ &= \arg \max_{\theta} \ln P(\mathcal{D} \mid \theta)\end{aligned}$$

Coin Flipping – Binomial Distribution



- $P(\text{Heads}) = \theta, P(\text{Tails}) = 1 - \theta$
- Flips are i.i.d.
 - Independent events
 - Identically distributed according to Binomial distribution
- Our training data consists of α_H heads and α_T tails

$$p(D|\theta) = \theta^{\alpha_H} \cdot (1 - \theta)^{\alpha_T}$$

First Parameter Learning Algorithm



$$\begin{aligned}\hat{\theta} &= \arg \max_{\theta} \ln P(\mathcal{D} \mid \theta) \\ &= \arg \max_{\theta} \ln \theta^{\alpha_H} (1 - \theta)^{\alpha_T}\end{aligned}$$

Set derivative to zero, and solve!

$$\begin{aligned}\frac{d}{d\theta} \ln P(\mathcal{D} \mid \theta) &= \frac{d}{d\theta} [\ln \theta^{\alpha_H} (1 - \theta)^{\alpha_T}] \\ &= \frac{d}{d\theta} [\alpha_H \ln \theta + \alpha_T \ln(1 - \theta)] \\ &= \alpha_H \frac{d}{d\theta} \ln \theta + \alpha_T \frac{d}{d\theta} \ln(1 - \theta) \\ &= \frac{\alpha_H}{\theta} - \frac{\alpha_T}{1 - \theta} = 0\end{aligned}$$

First Parameter Learning Algorithm



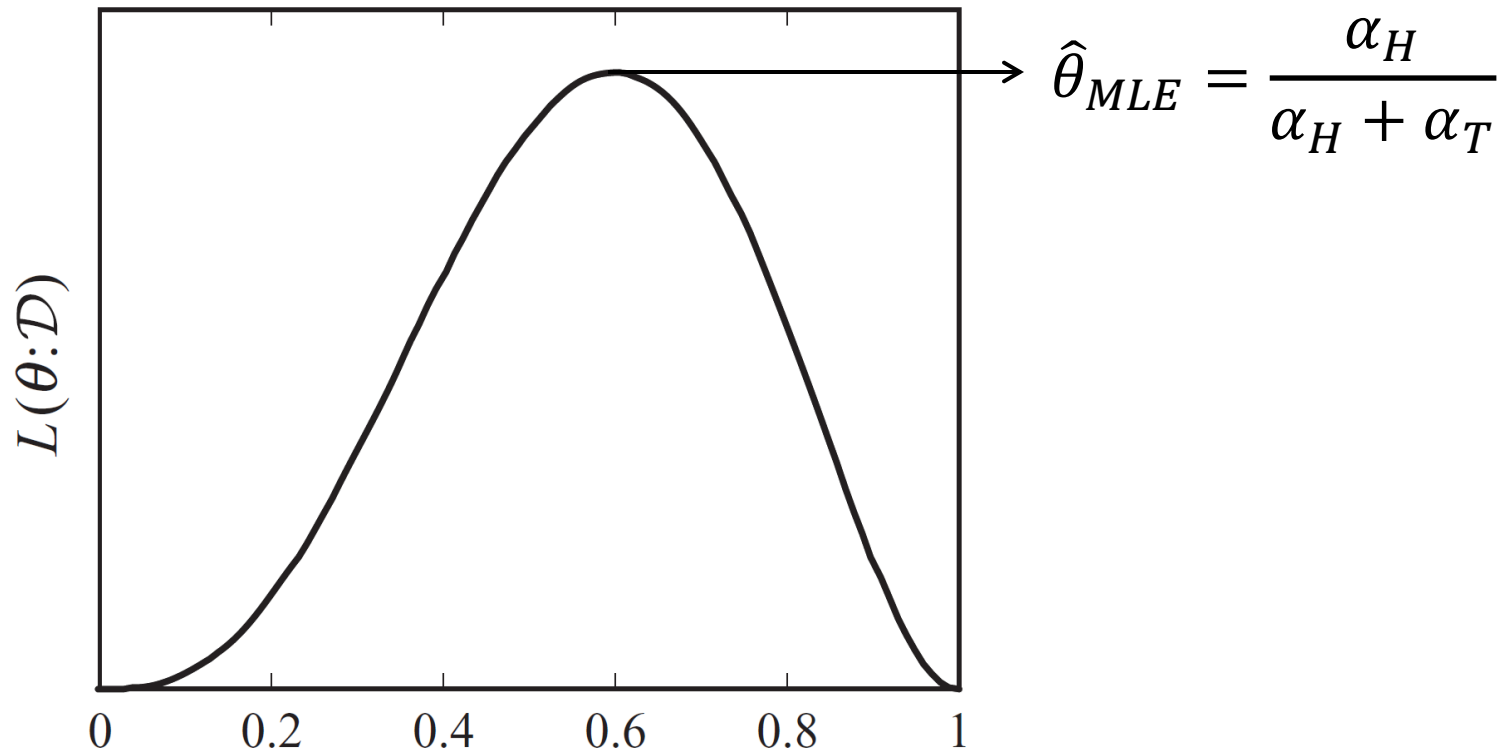
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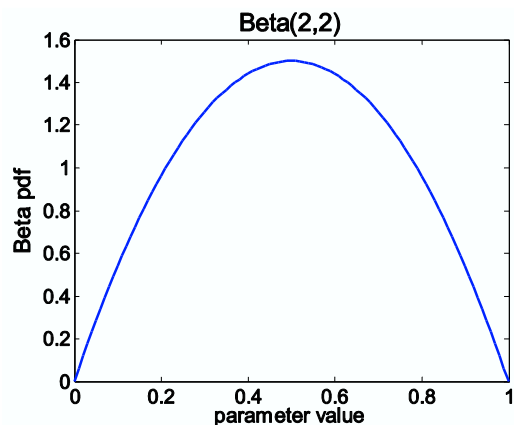
$$\boxed{\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}}$$

Coin Flip MLE



- Priors are a Bayesian mechanism that allow us to take into account “prior” knowledge about our belief in the outcome
- Rather than estimating a single θ , consider a distribution over possible values of θ given the data
 - Update our prior after seeing data

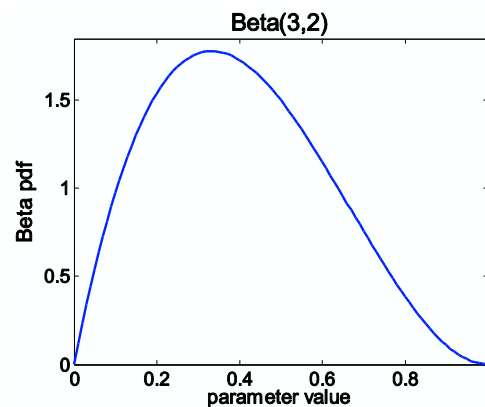
Our best guess in the absence of any data



Observe flips
e.g.: {tails, tails}



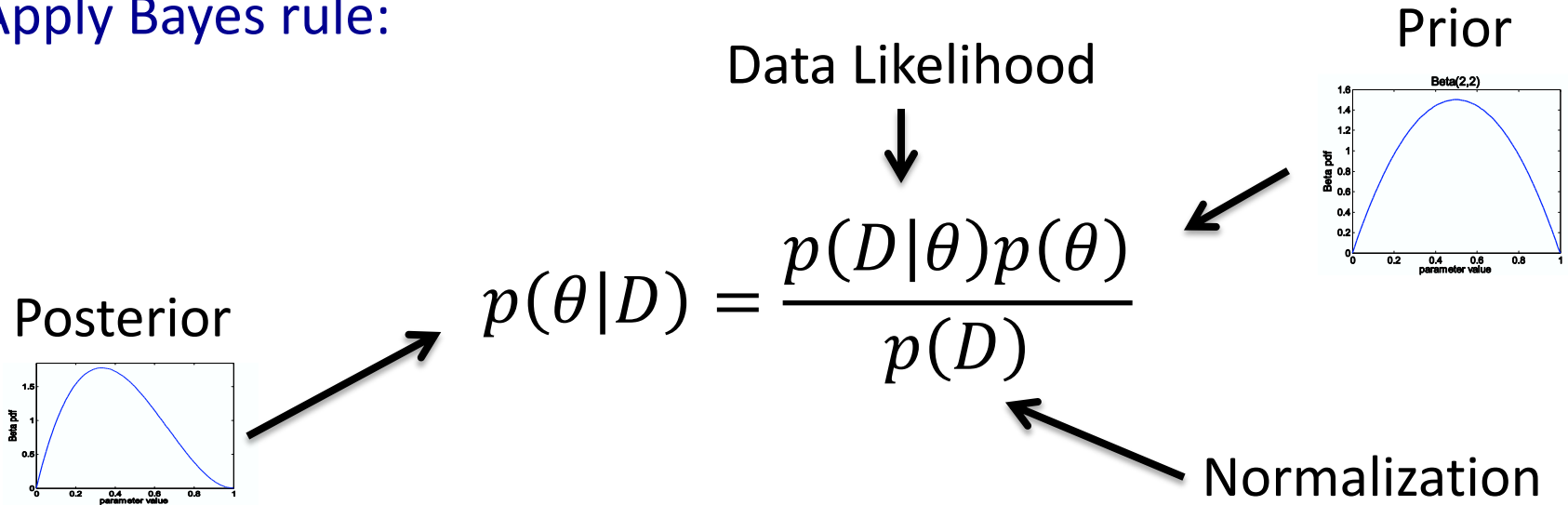
Our estimate after we see some data



Bayesian Learning



Apply Bayes rule:



- Or equivalently: $p(\theta|D) \propto p(D|\theta)p(\theta)$
- For uniform priors this reduces to the MLE objective

$$p(\theta) \propto 1 \quad \Rightarrow \quad p(\theta|D) \propto p(D|\theta)$$

Coin Flips with Beta Distribution

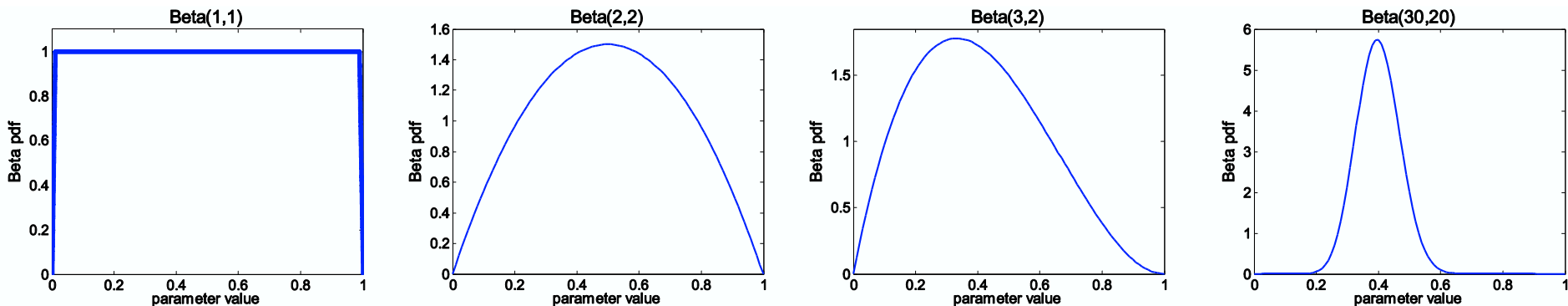


Likelihood function:

$$P(\mathcal{D} \mid \theta) = \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

Prior:

$$P(\theta) = \frac{\theta^{\beta_H-1} (1 - \theta)^{\beta_T-1}}{B(\beta_H, \beta_T)} \sim \text{Beta}(\beta_H, \beta_T)$$



$$\begin{aligned} P(\theta \mid \mathcal{D}) &\propto \theta^{\alpha_H} (1 - \theta)^{\alpha_T} \theta^{\beta_H-1} (1 - \theta)^{\beta_T-1} \\ &= \theta^{\alpha_H+\beta_H-1} (1 - \theta)^{\alpha_T+\beta_T-1} \\ &= \text{Beta}(\alpha_H+\beta_H, \alpha_T+\beta_T) \end{aligned}$$

- Choosing θ to maximize the posterior distribution is called maximum a posteriori (MAP) estimation

$$\theta_{MAP} = \arg \max_{\theta} p(\theta|D)$$

- The only difference between θ_{MLE} and θ_{MAP} is that one assumes a uniform prior (MLE) and the other allows an arbitrary prior

MAP for the Coin Flip Model

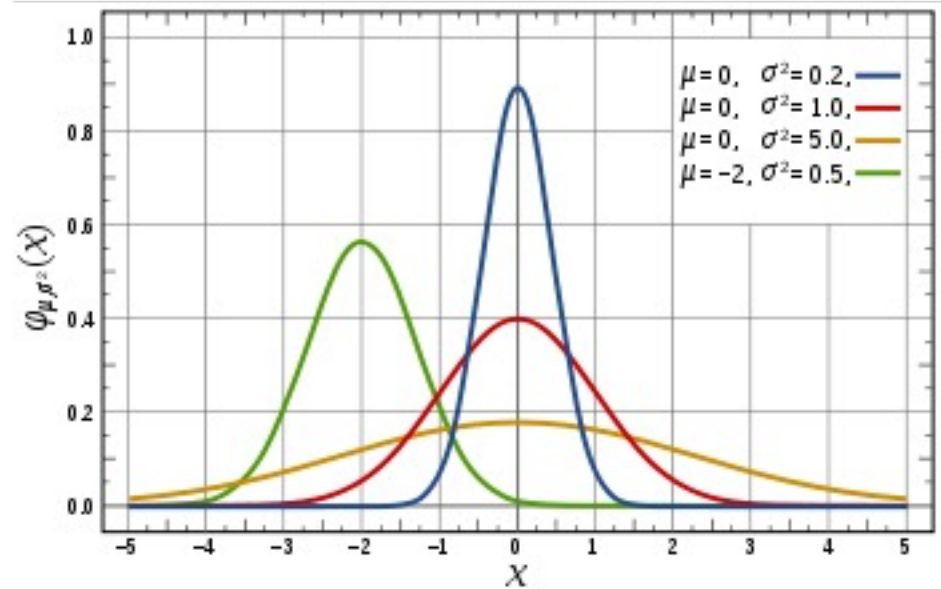


- Suppose we have 5 coin flips all of which are heads
 - MLE would give $\theta_{MLE} = 1$
 - MLE with a $Beta(2,2)$ prior gives $\theta_{MAP} = \frac{6}{7} \approx .857$
 - As we see more data, the effect of the prior diminishes
 - $\theta_{MAP} = \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2} \approx \frac{\alpha_H}{\alpha_H + \alpha_T}$ for large # of observations

MLE for Gaussian Distributions



- Two parameter distribution characterized by a mean and a variance



$$P(x \mid \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{-(x-\mu)^2}{2\sigma^2}}$$

Learning a Gaussian



- Collect data
 - Hopefully, i.i.d. samples
 - e.g., exam scores
- Learn parameters
 - Mean: μ
 - Variance: σ

i	Exam Score
0	85
1	95
2	100
3	12
...	...
99	89

$$P(x \mid \mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} e^{\frac{-(x-\mu)^2}{2\sigma^2}}$$

- Probability of N i.i.d. samples $D = x^{(1)}, \dots, x^{(N)}$

$$p(D|\mu, \sigma) = \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^N \prod_{i=1}^N e^{-\frac{(x^{(i)} - \mu)^2}{2\sigma^2}}$$

$$\mu_{MLE}, \sigma_{MLE} = \arg \max_{\mu, \sigma} P(\mathcal{D} \mid \mu, \sigma)$$

- Log-likelihood of the data

$$\ln p(D|\mu, \sigma) = -\frac{N}{2} \ln 2\pi\sigma^2 - \sum_{i=1}^N \frac{(x^{(i)} - \mu)^2}{2\sigma^2}$$

MLE for the Mean of a Gaussian



$$\begin{aligned}\frac{\partial}{\partial \mu} \ln p(D|\mu, \sigma) &= \frac{\partial}{\partial \mu} \left[-\frac{N}{2} \ln 2\pi\sigma^2 - \sum_{i=1}^N \frac{(x^{(i)} - \mu)^2}{2\sigma^2} \right] \\ &= \frac{\partial}{\partial \mu} \left[-\sum_{i=1}^N \frac{(x^{(i)} - \mu)^2}{2\sigma^2} \right] \\ &= \sum_{i=1}^N \frac{(x^{(i)} - \mu)}{\sigma^2} \\ &= \frac{[N\mu - \sum_{i=1}^N x^{(i)}]}{\sigma^2} = 0\end{aligned}$$

$$\mu_{MLE} = \frac{1}{N} \sum_{i=1}^N x^{(i)}$$

$$\begin{aligned}\frac{\partial}{\partial \sigma} \ln p(D|\mu, \sigma) &= \frac{\partial}{\partial \sigma} \left[-\frac{N}{2} \ln 2\pi\sigma^2 - \sum_{i=1}^N \frac{(x^{(i)} - \mu)^2}{2\sigma^2} \right] \\ &= -\frac{N}{\sigma} + \frac{\partial}{\partial \sigma} \left[-\sum_{i=1}^N \frac{(x^{(i)} - \mu)^2}{2\sigma^2} \right] \\ &= -\frac{N}{\sigma} + \sum_{i=1}^N \frac{(x^{(i)} - \mu)^2}{\sigma^3} = 0\end{aligned}$$

$$\sigma_{MLE}^2 = \frac{1}{N} \sum_{i=1}^N (x^{(i)} - \mu_{MLE})^2$$

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- Given features $x = (x_1, \dots, x_m)$ predict a label y
- If we had a joint distribution over x and y , given x we could find the label using MAP inference

$$\arg \max_y p(y|x_1, \dots, x_m)$$

- Can compute this in exactly the same way that we did before using Bayes rule:

$$p(y|x_1, \dots, x_m) = \frac{p(x_1, \dots, x_m|y)p(y)}{p(x_1, \dots, x_m)}$$

Bag of Words



the world of

TOTAL



all about the company

Our energy exploration, production, and distribution operations span the globe, with activities in more than 100 countries.

At TOTAL, we draw our greatest strength from our fast-growing oil and gas reserves. Our strategic emphasis on natural gas provides a strong position in a rapidly expanding market.

Our expanding refining and marketing operations in Asia and the Mediterranean Rim complement already solid positions in Europe, Africa, and the U.S.

Our growing specialty chemicals sector adds balance and profit to the core energy business.

► All About The Company

- Global Activities
- Corporate Structure
- TOTAL's Story
- Upstream Strategy
- Downstream Strategy
- Chemicals Strategy
- TOTAL Foundation
- Homepage



aardvark	0
about2	
all	2
Africa	1
apple0	
anxious	0
...	
gas	1
...	
oil	1
...	
Zaire0	

- Naïve Bayes assumption
 - Features are independent given class label

$$p(x_1, x_2 | y) = p(x_1 | y) p(x_2 | y)$$

- More generally

$$p(x_1, \dots, x_m | y) = \prod_{i=1}^m p(x_i | y)$$

- How many parameters now?
 - Suppose x is composed of d binary features

- Naïve Bayes assumption
 - Features are independent given class label

$$p(x_1, x_2 | y) = p(x_1 | y) p(x_2 | y)$$

- More generally

$$p(x_1, \dots, x_m | y) = \prod_{i=1}^m p(x_i | y)$$

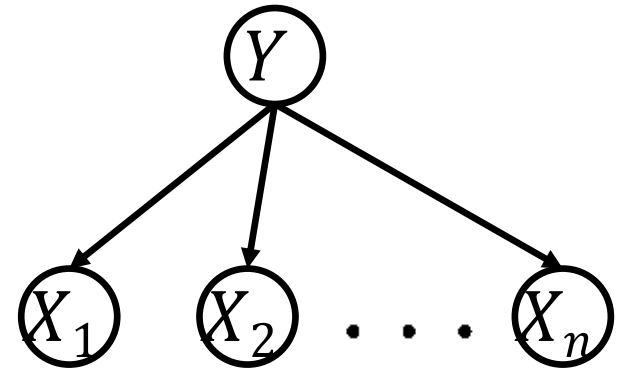
- How many parameters now?
 - Suppose x composed of d binary features $\Rightarrow O(d \cdot L)$ where L is the number of class labels

The Naïve Bayes Classifier



- **Given**

- Prior $p(y)$
- m conditionally independent features X given the class Y
- For each X_i , we have likelihood $P(X_i|Y)$



- Classify via

$$\begin{aligned} y^* = h_{NB}(x) &= \arg \max_y p(y) p(x_1, \dots, x_m | y) \\ &= \arg \max_y p(y) \prod_i^m p(x_i | y) \end{aligned}$$

- Given dataset, count occurrences for all pairs
 - $Count(X_i = x_i, Y = y)$ is the number of samples in which $X_i = x_i$ and $Y = y$
- MLE for discrete NB

$$p(Y = y) = \frac{Count(Y = y)}{\sum_{y'} Count(Y = y')}$$

$$p(X_i = x_i | Y = y) = \frac{Count(X_i = x_i, Y = y)}{\sum_{x'_i} Count(X_i = x'_i, Y = y)}$$

See this link for more insights: <http://www.datasciencecourse.org/notes/mle/>

- To fix this, use a prior!
 - Already saw how to do this in the coin-flipping example using the Beta distribution
 - For NB over discrete spaces, can use the Dirichlet prior
 - The Dirichlet distribution is a distribution over $z_1, \dots, z_k \in (0,1)$ such that $z_1 + \dots + z_k = 1$ characterized by k parameters $\alpha_1, \dots, \alpha_k$

$$f(z_1, \dots, z_k; \alpha_1, \dots, \alpha_k) \propto \prod_{i=1}^k z_i^{\alpha_i - 1}$$

- Called **smoothing**, what are the MLE estimates under these kinds of priors?

Continuous Naïve Bayes



- Continuous Naïve Bayes, also known as Guassian Naïve Bayes is where the features are continuous
- The distribution $p(X_i = x_i | Y = y) = N(x_i, \mu_y, \sigma_y^2)$
- In other words, the conditional distribution of each feature given the class is a Guassian distribution with mean μ_y and variance σ_y^2
- We can use the Naïve Bayes assumption and assume:

$$p(x_1, \dots, x_m | y) = \prod_{i=1}^m p(x_i | y)$$

- The distribution of labels is the same as the multinomial case

- The parameter estimation can similarly be obtained using the Maximum Likelihood Estimation
- The mean and variance can be estimated as the standard Gaussian distribution except that we restrict to each label

$$\mu_y = \frac{\sum_{j=1}^m x_i^{(j)} 1\{y^{(j)} = y\}}{\sum_{j=1}^m 1\{y^{(j)} = y\}},$$

$$\sigma_y^2 = \frac{\sum_{j=1}^m (x_i^{(j)} - \mu_y)^2 1\{y^{(j)} = y\}}{\sum_{j=1}^m 1\{y^{(j)} = y\}}$$

- Finally, we need to estimate $p(y)$
- This is like the discrete Naïve Bayes case:

$$p(Y = y) = \frac{\text{Count}(Y = y)}{\sum_{y'} \text{Count}(Y = y')}$$

- We can classify a test example in a similar way to discrete NB:

$$\begin{aligned} y^* = h_{NB}(x) &= \arg \max_y p(y) p(x_1, \dots, x_m | y) \\ &= \arg \max_y p(y) \prod_i^m p(x_i | y) \end{aligned}$$

- Here $p(x_i | y) = N(x_i, \mu_y, \sigma_y^2)$

Summary of Naïve Bayes Models



- Two kinds of Naïve Bayes: Discrete and Continuous
- Learning is often very simple
 - Using counts (discrete NB) or mean/variance (cont. NB), obtain estimates for $p(x_i | y)$
 - Using counts, obtain estimates for $p(y)$
- At inference time, we classify based on:

$$\begin{aligned} y^* = h_{NB}(x) &= \arg \max_y p(y) p(x_1, \dots, x_m | y) \\ &= \arg \max_y p(y) \prod_i^m p(x_i | y) \end{aligned}$$

Topics for the Midterm Exam



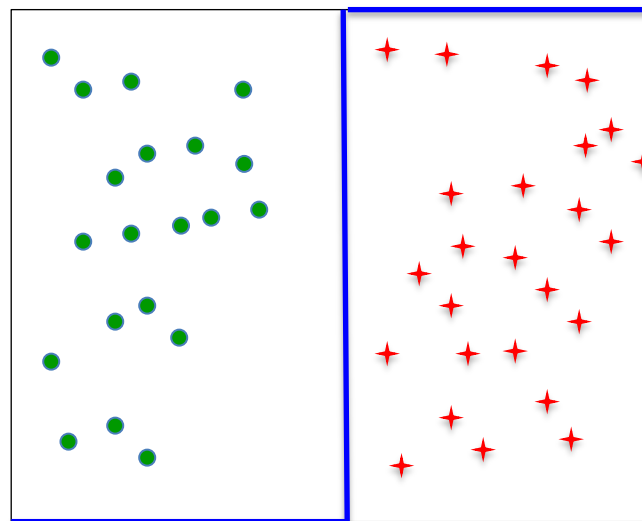
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Ideal 0/1 Probability



- Learn $p(Y|X)$ directly from the data
 - Assume a particular functional form, e.g., a linear classifier $p(Y = 1|x) = 1$ on one side and 0 on the other
- Not differentiable...
 - Makes it difficult to learn
 - Can't handle noisy labels

$$p(Y = 1|x) = 0$$



$$p(Y = 1|x) = 1$$

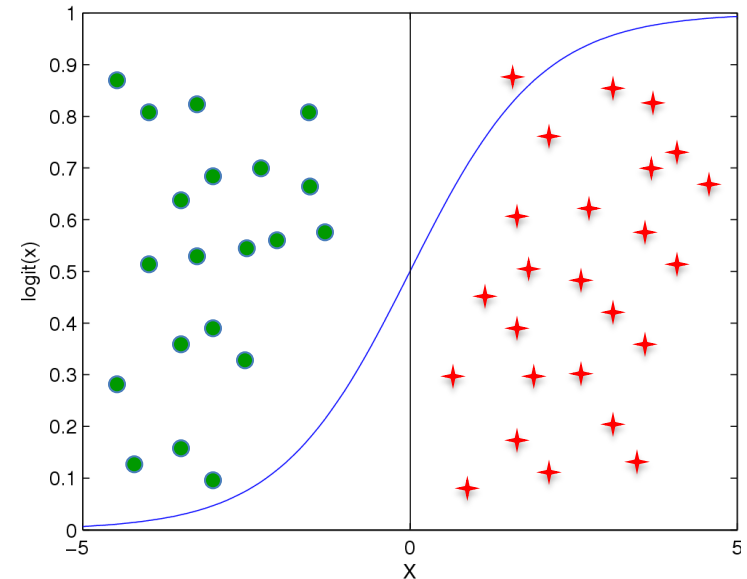
Logistic Regression



- Learn $p(y|x)$ directly from the data
- Assume a particular functional form

$$p(Y = -1|x) = \frac{1}{1 + \exp(w^T x + b)}$$

$$p(Y = 1|x) = \frac{\exp(w^T x + b)}{1 + \exp(w^T x + b)}$$



- Given some w and b , we can classify a new point x by assigning the label 1 if $p(Y = 1|x) > p(Y = -1|x)$ and -1 otherwise
 - This leads to a linear classification rule:
 - Classify as a 1 if $w^T x + b > 0$
 - Classify as a -1 if $w^T x + b < 0$

- To learn the weights, we maximize the **conditional likelihood**

$$(w^*, b^*) = \arg \max_{w, b} \prod_{i=1}^N p(y^{(i)} | x^{(i)}, w, b)$$

- This is not the same strategy that we used in the case of naive Bayes
 - For naive Bayes, we maximized the log-likelihood

Learning the Weights



$$\begin{aligned}\ell(w, b) &= \ln \prod_{i=1}^N p(y^{(i)} | x^{(i)}, w, b) \\&= \sum_{i=1}^N \ln p(y^{(i)} | x^{(i)}, w, b) \\&= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} \ln p(Y = 1 | x^{(i)}, w, b) + \left(1 - \frac{y^{(i)} + 1}{2}\right) \ln p(Y = -1 | x^{(i)}, w, b) \\&= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} \ln \frac{p(Y = 1 | x^{(i)}, w, b)}{p(Y = -1 | x^{(i)}, w, b)} + \ln p(Y = -1 | x^{(i)}, w, b) \\&= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} (w^T x^{(i)} + b) - \ln(1 + \exp(w^T x^{(i)} + b))\end{aligned}$$

Learning the Weights



$$\begin{aligned}\ell(w, b) &= \ln \prod_{i=1}^N p(y^{(i)} | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \ln p(y^{(i)} | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} \ln p(Y = 1 | x^{(i)}, w, b) + \left(1 - \frac{y^{(i)} + 1}{2}\right) \ln p(Y = -1 | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} \ln \frac{p(Y = 1 | x^{(i)}, w, b)}{p(Y = -1 | x^{(i)}, w, b)} + \ln p(Y = -1 | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} (w^T x^{(i)} + b) - \ln(1 + \exp(w^T x^{(i)} + b))\end{aligned}$$

This is concave in w and b : take derivatives and solve!

Learning the Weights



$$\begin{aligned}\ell(w, b) &= \ln \prod_{i=1}^N p(y^{(i)} | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \ln p(y^{(i)} | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} \ln p(Y = 1 | x^{(i)}, w, b) + \left(1 - \frac{y^{(i)} + 1}{2}\right) \ln p(Y = -1 | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} \ln \frac{p(Y = 1 | x^{(i)}, w, b)}{p(Y = -1 | x^{(i)}, w, b)} + \ln p(Y = -1 | x^{(i)}, w, b) \\ &= \sum_{i=1}^N \frac{y^{(i)} + 1}{2} (w^T x^{(i)} + b) - \ln(1 + \exp(w^T x^{(i)} + b))\end{aligned}$$

No closed form solution ☹

- Can apply gradient **ascent** to maximize the conditional likelihood

$$\frac{\partial \ell}{\partial b} = \sum_{i=1}^N \left[\frac{y^{(i)} + 1}{2} - p(Y = 1 | x^{(i)}, w, b) \right]$$

$$\frac{\partial \ell}{\partial w_j} = \sum_{i=1}^N x_j^{(i)} \left[\frac{y^{(i)} + 1}{2} - p(Y = 1 | x^{(i)}, w, b) \right]$$

- Can define priors on the weights to prevent overfitting
 - Normal distribution, zero mean, identity covariance

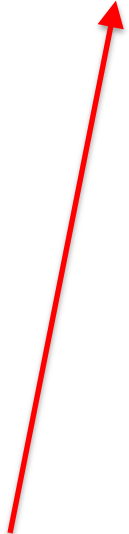
$$p(w) = \prod_j \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{w_j^2}{2\sigma^2}\right)$$

- “Pushes” parameters towards zero
- Regularization
 - Helps avoid very large weights and overfitting

- The log-MAP objective with this Gaussian prior is then

$$\ln \prod_{i=1}^N p(y^{(i)} | x^{(i)}, w, b) p(w) p(b) = \left[\sum_i^N \ln p(y^{(i)} | x^{(i)}, w, b) \right] - \frac{\lambda}{2} \|w\|_2^2$$

- Quadratic penalty: drives weights towards zero
- Adds a negative linear term to the gradients
- Different priors can produce different kinds of regularization



Sometimes called an ℓ_2 regularizer

Generative vs. Discriminative Classifiers

Generative classifier: (e.g., Naïve Bayes)

- Assume some **functional form** for $p(x|y), p(y)$
- Estimate parameters of $p(x|y), p(y)$ directly from training data
- Use Bayes rule to calculate $p(y|x)$
- This is a **generative model**
 - **Indirect** computation of $p(Y|X)$ through Bayes rule
 - As a result, **can also generate a sample of the data**,
$$p(x) = \sum_y p(y)p(x|y)$$

Discriminative classifiers: (e.g., Logistic Regression)

- Assume some **functional form for $p(y|x)$**
- Estimate parameters of $p(y|x)$ directly from training data
- This is a **discriminative model**
 - Directly learn $p(y|x)$
 - But **cannot obtain a sample of the data** as $p(x)$ is not available
 - Useful for discriminating labels