Connected Stocks: Evidence from Tehran Stock Exchange

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Effects

Hypothesis 1

Simple measures of institutional connnectedness statistically and economically improve forecasts of cross-sectional variation in the correlation. The effect is stronger for pairs that are in the same business groups.

| | | Depe | ndent Variab | ole: Future M | Ionthly Corr | elation of 4F | +Industry F | Residuals | |
|--------------------|-----------|-----------|--------------|---------------|--------------|---------------|-------------|------------|------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| Same Group | 0.0138*** | 0.0128*** | | | 0.0103*** | 0.00978*** | 0.00694 | 0.00458 | 0.00356 |
| | (5.76) | (6.29) | | | (4.22) | (4.29) | (1.96) | (1.43) | (1.11) |
| FCA* | | | 0.00405*** | 0.00375*** | 0.00301*** | 0.00296*** | 0.00273*** | 0.00258*** | 0.00273*** |
| | | | (4.94) | (5.12) | (3.68) | (3.77) | (3.47) | (3.53) | (3.51) |
| (FCA*) × SameGroup | | | | | | | 0.00369 | 0.00524** | 0.00517** |
| | | | | | | | (1.97) | (3.21) | (3.18) |
| Observations | 388492 | 388492 | 388492 | 388492 | 388492 | 388492 | 388492 | 388492 | 388492 |
| Group Effect | No | No | No | No | No | No | No | No | Yes |
| Controls | No | Yes | No | Yes | No | Yes | No | Yes | Yes |
| R^2 | 0.000404 | 0.00200 | 0.000423 | 0.00201 | 0.000716 | 0.00229 | 0.000872 | 0.00245 | 0.00875 |

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t statistics in parentheses $\label{eq:problem} {}^*\ p < 0.05,\ {}^{**}\ p < 0.01,\ {}^{***}\ p < 0.001$

Hypothesis 2

Pairs of companies belonging to the same business group have a higher correlation than pairs not in the same group. In addition, Pairs that belong to the same group and have a common ownership co-move more than pairs that don't have common ownership.

Table 1: one of these tables

| | | Future M | onthly Corre | lation of 4F+In | dustry Residua | ıls |
|--|-----------|------------|---------------|-----------------|----------------|-------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| (FCA > Median[FCA]) | | -0.00168 | -0.00337** | 0.00855** | | -0.00513*** |
| | | (-1.45) | (-2.89) | (2.76) | | (-4.32) |
| SameGroup | 0.0122*** | | 0.0135*** | | | 0.00574* |
| | (5.81) | | (6.48) | | | (2.02) |
| $(FCA > Median[FCA]) \times SameGroup$ | р | | | | | 0.0181*** |
| | | | | | | (5.91) |
| FCA* | | | | | 0.00174* | |
| | | | | | (2.43) | |
| Observations | 5148109 | 5148109 | 5148109 | 76240 | 76240 | 5148109 |
| Sub Sample | Total | Total | Total | SameGroups | SameGroups | Total |
| Controls | Yes | Yes | Yes | Yes | Yes | Yes |
| R^2 | 0.000455 | 0.000439 | 0.000485 | 0.0136 | 0.0135 | 0.000513 |
| t statistics in parentheses | | | | | | |
| * $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$ | | | | | | |
| | | Future Mor | nthly Correla | tion of 4F+Ind | ustry Residual | S |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Common Ownership | | -0.00350** | -0.00445*** | 0.00651* | | -0.00527*** |
| | | (-3.30) | (-4.22) | (2.48) | | (-4.72) |
| SameGroup | 0.0122*** | | 0.0140*** | | | 0.00607^* |
| | (5.81) | | (7.01) | | | (2.09) |
| Common Ownership \times SameGroup | | | | | | 0.0157*** |
| | | | | | | (5.51) |
| FCA* | | | | | 0.00174* | |
| | | | | | (2.43) | |
| Observations | 5148109 | 5148109 | 5148109 | 76240 | 76240 | 5148109 |
| Sub Sample | Total | Total | Total | SameGroups | SameGroups | Total |
| Controls | Yes | Yes | Yes | Yes | Yes | Yes |
| R^2 | 0.000455 | 0.000456 | 0.000504 | 0.0135 | 0.0135 | 0.000528 |

t statistics in parentheses

^{*} p < 0.05, ** p < 0.01, *** p < 0.001

Hypothesis 3

Stock returns of group affiliated firms exhibit robustly positive comovement even after controlling for both market and industry effects. Group betas $(\beta_{Businussgroup})$ are highly significant across all models.

Table 2: Cross-sectional average of the time-series coefficients

| | | Re | $\overline{\mathrm{turn}_i - r_f}$ | $=R_i$ | |
|---------------------|----------|----------|------------------------------------|-----------|-----------|
| | (1) | (2) | (3) | (4) | (5) |
| R_M | 0.801*** | 0.643*** | 0.701*** | 0.257*** | 0.280*** |
| | (29.99) | (10.68) | (11.05) | (8.84) | (9.02) |
| $R_{Industry}$ | | -2.085 | -1.878 | -0.150 | -0.148 |
| - | | (-0.92) | (-0.93) | (-0.48) | (-0.50) |
| $R_{Businessgroup}$ | | | | 0.493*** | 0.493*** |
| - | | | | (11.36) | (11.34) |
| SMB | | | 0.104*** | | 0.0770*** |
| | | | (3.52) | | (5.24) |
| UMD | | | 0.0282 | | 0.0218 |
| | | | (1.23) | | (1.94) |
| HML | | | 0.102*** | | 0.0395*** |
| | | | (6.05) | | (6.39) |
| Constant | 0.0442 | 0.0145 | -0.0297 | 0.0499*** | 0.0198 |
| | (1.92) | (0.53) | (-0.83) | (3.87) | (1.25) |
| Observations | 207552 | 207552 | 207552 | 207552 | 207552 |
| R^2 | 0.123 | 0.196 | 0.213 | 0.672 | 0.679 |

t statistics in parentheses

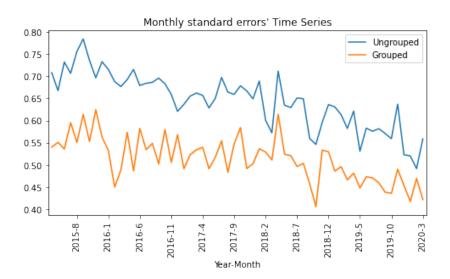
^{*} p < 0.05, ** p < 0.01, *** p < 0.001

Channels

Trading

For each firm, we calculate daily institutional imbalances, which is the net buying value of institutional investors relative to total traded value on that day (InsImb = $\frac{\text{Buy}_{\text{value}} - \text{Sell}_{\text{value}}}{\text{Buy}_{\text{value}} + \text{Sell}_{\text{value}}}$). We expect that institutional imbalances have a lower variation in groups due to the correlated tradings that the ultimate owner ordered to do. So, we calculate the monthly standard deviation of the group's imbalances and compare them to unaffiliated ones. As we expected grouped standard error is 13.1% and significantly (with t-stat of 12.57) lower than ungrouped firms.

| | count | mean | std | min | median | max |
|-----------|-------|-------|-------|-------|--------|-------|
| Ungrouped | 60 | 0.645 | 0.063 | 0.492 | 0.653 | 0.784 |
| Grouped | 60 | 0.514 | 0.050 | 0.406 | 0.514 | 0.625 |



According to the main hypothesis, we need to compare comovement between pairs in groups with low standard error and other pairs. For this purpose, we define **Low Imbalance std** dummy for groups whose average standard errors are lower than half of the sample. So, this dummy is equal to one if at least one pair's firms belong to the low imbalance std business group.

| | | F | uture Month | ly Corr. of 4 | F+Ind. Residua | als | |
|---|------------|------------|-------------|---------------|----------------|------------|------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| FCA* | 0.00296*** | 0.00277*** | 0.00275*** | | 0.00611** | 0.00244** | 0.00284** |
| | (3.77) | (3.57) | (3.55) | | (3.21) | (3.14) | (3.40) |
| Same Group | 0.00978*** | 0.00981*** | 0.00858** | 0.0110*** | | 0.00861** | 0.00826** |
| | (4.29) | (4.35) | (3.37) | (4.73) | | (3.38) | (3.05) |
| Low Imbalance std | | -0.00364** | -0.00388** | -0.00446** | -0.00725* | -0.00393** | 0.000437 |
| | | (-2.81) | (-2.83) | (-3.24) | (-2.47) | (-2.87) | (0.21) |
| Low Imbalance std \times SameGroup | | | 0.00301 | 0.00365 | | -0.00904 | -0.00990* |
| • | | | (0.81) | (0.98) | | (-1.84) | (-2.02) |
| $Low\ Imbalance\ std \times SameGroup \times FCA^*$ | | | | | | 0.0104*** | 0.00941*** |
| • | | | | | | (3.87) | (3.53) |
| Observations | 388492 | 388492 | 388492 | 388492 | 37114 | 388492 | 388492 |
| Group Effect | No | No | No | No | No | No | Yes |
| Sub-sample | Total | Total | Total | Total | Same Groups | Total | Total |
| Controls | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| R^2 | 0.00229 | 0.00255 | 0.00274 | 0.00246 | 0.0199 | 0.00290 | 0.00906 |

Furthermore, we should show that stocks in groups have a similar daily trading behavior. Accordingly, for each firm we run time-series regressions of the firm's daily change in trading measure, $\Delta \text{Measure}_{i,t}$, on changes in market measure, Δ Measure_{Market,t}, changes in the industry and business group portfolio's measure, Δ Measure $_{Ind,t}$ and Δ Measure $_{Group,t}$ and ,as well as control variables.

We compute the daily change of measure by this definition Δ Measure_{i,t} = $\ln(\frac{\text{Measure}_{i,t}}{\text{Measure}_{i,t-1}})$. We estimate the following regression for each stock across trading days in given year separately and cross-sectional averages of the estimated coefficients are reported, with t-statistics in parentheses:

$$\Delta \text{Measure}_{i,t} = \alpha + \beta_{Market,t} \Delta \text{Measure}_{Market,t} + \beta_{Ind,t} \Delta \text{Measure}_{Ind,t} + \beta_{Group,t} \Delta \text{Measure}_{Group,t} + \delta \text{Controls} + \varepsilon_{i,t}$$

We use the turnover and Amihud measure as a daily trading measures separately. For both measures we control for lead and lag changes in the two portfolio and market's measures. In addition, for turnover measure, we use size of the firm and Amihud, we include lead, lag, and contemporaneous market returns, contemporaneous firm return squared. [Table 3,4]

In a second test of our proposed channel, we control for stock characteristics in a multivariate regression. we regress β_{Group} on the previous year's excess control right from cash flow right, controlling for firm size and average of trading measure. In our main specification, we include time-fixed effects and cluster the standard errors at the firm and time-dimension level to account for time-series and cross-sectional dependence. [Table 5,6] The

t statistics in parentheses * $p < 0.05, \, ^{**} \, p < 0.01, \, ^{***} \, p < 0.001$

Table 3: cross-sectional average of the time-series coefficients for daily changes in turnover

| | | Dep | endent Varia | ble: $\Delta Turn$ | $Over_i$ | |
|------------------------------|----------|----------|----------------|--------------------|----------|----------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| $\Delta TurnOver_{Market}$ | 0.405*** | 0.396*** | 0.360*** | 0.425*** | 0.388*** | 0.448*** |
| | (12.25) | (10.74) | (7.62) | (12.08) | (8.23) | (12.20) |
| $\Delta TurnOver_{Group}$ | | | 0.222*** | 0.229*** | 0.253** | 0.268*** |
| | | | (3.46) | (4.09) | (3.28) | (3.82) |
| $\Delta TurnOver_{Industry}$ | 0.120** | 0.0205 | -0.0156 | -0.0237 | -0.0833 | -0.0999 |
| - | (3.25) | (0.24) | (-0.23) | (-0.42) | (-1.04) | (-1.46) |
| Observations | 293264 | 292179 | 184699 | 183442 | 184699 | 183442 |
| Weight | _ | - | $MC \times CR$ | $MC \times CR$ | MC | MC |
| Control | No | Yes | No | Yes | No | Yes |
| R^2 | 0.129 | 0.168 | 0.246 | 0.286 | 0.247 | 0.286 |

t statistics in parentheses

Table 4: cross-sectional average of the time-series coefficients for daily changes in illiquidity

| | | Dej | pendent Vari | able: ΔAmi | hud_i | |
|-----------------------------------|----------|----------|----------------|--------------------|------------------|----------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Δ Amihud _{Market} | 0.290*** | 0.298*** | 0.365*** | 0.234*** | 0.373*** | 0.244*** |
| | (9.76) | (3.38) | (11.12) | (5.29) | (11.48) | (5.70) |
| $\Delta Amihud_{Group}$ | | | 0.182*** | 0.167*** | 0.161** | 0.148** |
| • | | | (3.58) | (3.86) | (2.93) | (3.11) |
| $\Delta Amihud_{Industry}$ | 0.0687* | 0.144 | 0.00964 | -0.0107 | 0.0162 | -0.00565 |
| | (2.02) | (1.59) | (0.19) | (-0.25) | (0.30) | (-0.12) |
| Observations | 293264 | 291933 | 184699 | 183301 | 184699 | 183301 |
| Weight | - | - | $MC \times CR$ | $MC \times CR$ | MC | MC |
| Control | No | Yes | No | Yes | No | Yes |
| R^2 | 0.118 | 0.223 | 0.219 | 0.320 | 0.224 | 0.324 |

t statistics in parentheses

^{*} p < 0.05, ** p < 0.01, *** p < 0.001

^{*} p < 0.05, ** p < 0.01, *** p < 0.001

specification is:

$$\beta_{Group,i,t} = \alpha + \beta_1 \text{GroupVariable}_{i,t-1} + \beta_2 \ln(\text{Size})_{i,t-1} + \beta_3 \text{trading measure}(avg)_{i,t-1} + \text{time effects} + \varepsilon_{i,t-1}$$

which our group variables are:

1. Excess =
$$(cr - cfr)/cr$$

2. ExcessDiff =
$$cr - cfr$$

3. ExcessDummy =
$$\begin{cases} 1 & \text{cr} - \text{cfr} > 0 \\ 0 & \text{cr} - \text{cfr} \le 0 \end{cases}$$

4. Excess
High =
$$\begin{cases} 1 & \text{Excess} \ge \text{Q3(Excess)} \\ 0 & \text{Excess} < \text{Q3(Excess)} \end{cases}$$

5. Low Imbalance std =
$$\begin{cases} 1 & std_{InsImb,t} \leq \text{Median}(std_{InsImb,t}) \\ 0 & std_{InsImb,t} > \text{Median}(std_{InsImb,t}) \end{cases}$$

Table 5: β_{Group} of daily changes in the turnover on Excess control right ((cr-cf)/cr)and other measures

| | | | | | | Depe | ndent Var | iable: β_{Gro} | пир | | | | | |
|-------------------|--------------------|--------------------|---------------------|---------------------|--------------------|--------------------|--------------------|----------------------|--------------------|-------------------|--------------------|--------------------|--------------------|------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) | (13) | (14) |
| Excess | 0.310*** (3.58) | 0.417*** (4.76) | | | | | | | | | | | | |
| ExcessDummy | | | -0.00418 (-0.10) | 0.0907^* (2.24) | | | | | | | | | | |
| ExcessDiff | | | | | 0.638*** (4.65) | 0.840*** (6.22) | | | | | | | | |
| ExcessHigh | | | | | | | 0.287*** (4.17) | 0.323*** (4.42) | | | | | | |
| Low Imbalance std | | | | | | | | | 0.216*** (4.82) | 0.0975* (2.26) | | | | |
| Position | | | | | | | | | | | -0.0103 (-0.54) | $0.0176 \\ (0.93)$ | | |
| Centrality | | | | | | | | | | | | | 0.618*** (3.31) | 0.0662 (0.37) |
| Observations | 1153 | 1153 | 1168 | 1168 | 1153 | 1153 | 1168 | 1168 | 1145 | 1145 | 1153 | 1153 | 1113 | 1113 |
| Time FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Controls | No | Yes | No | Yes | No | Yes | No | Yes | No | Yes | No | Yes | No | Yes |
| R^2 | 0.0178 | 0.0884 | 0.00206 | 0.0665 | 0.0313 | 0.109 | 0.0278 | 0.0923 | 0.0203 | 0.0687 | 0.00239 | 0.0645 | 0.00825 | 0.0562 |

t statistics in parentheses * p < 0.05, ** p < 0.01, *** p < 0.001

Table 6: β_{Group} of daily changes in the Amihud measure on Excess control right ((cr – (cf)/cr) and other measures

| | | | | | | Depe | ndent Var | iable: β_{Gre} | oup | | | | | |
|-------------------|------------------|--------------------|--------------------|-------------------|-------------------|--------------------|--------------------|----------------------|-------------------|-----------------|--------------------|-----------------|--------------------|-----------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) | (13) | (14) |
| Excess | 0.174* (2.55) | 0.354*** (4.78) | | | | | | | | | | | | |
| ExcessDummy | | | -0.0190 (-0.48) | 0.0764* (1.99) | | | | | | | | | | |
| ExcessDiff | | | | | 0.285** (2.86) | 0.554*** (5.59) | | | | | | | | |
| ExcessHigh | | | | | | | 0.242*** (4.39) | 0.346*** (5.52) | | | | | | |
| Low Imbalance std | | | | | | | | | 0.126** (3.15) | 0.0471 (1.20) | | | | |
| Position | | | | | | | | | | | -0.0102 (-0.62) | 0.0312 (1.81) | | |
| Centrality | | | | | | | | | | | | | 0.684*** (4.02) | 0.271 (1.58) |
| Observations | 1153 | 1153 | 1168 | 1168 | 1153 | 1153 | 1168 | 1168 | 1145 | 1145 | 1153 | 1153 | 1113 | 1113 |
| Time FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Controls | No | Yes | No | Yes | No | Yes | No | Yes | No | Yes | No | Yes | No | Yes |
| R^2 | 0.00898 | 0.0857 | 0.00196 | 0.0607 | 0.0102 | 0.0882 | 0.0266 | 0.104 | 0.0107 | 0.0648 | 0.00260 | 0.0641 | 0.0117 | 0.0438 |

Table 7: Pairwise correlation in turnover

| | | De | ependent Var | riable: Futu | re Monthly (| Correlation o | f Delta turno | over | |
|--|--------------------|----------------------|----------------------|---------------------|----------------------|-----------------------|----------------------|----------------------|---------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| Same Group | 0.0134** (3.13) | -0.00613* (-2.20) | | | 0.00503 (1.50) | -0.0102*** (-3.81) | 0.0104* (2.03) | -0.00763 (-1.75) | -0.00600 (-1.36) |
| FCA* | | | 0.00784*** (4.71) | 0.00308** (3.39) | 0.00724*** (5.06) | 0.00389*** (4.29) | 0.00771*** (4.84) | 0.00410*** (4.07) | 0.00304* (2.23) |
| $(\mathrm{FCA}^*) \times \mathrm{SameGroup}$ | | | | | | | -0.00552 (-1.67) | -0.00244 (-0.82) | -0.00104 (-0.33) |
| Observations | 378502 | 370726 | 378502 | 370726 | 378502 | 370726 | 378502 | 370726 | 370726 |
| Group Effect | No | No | No | No | No | No | No | No | Yes |
| Controls | No | Yes | No | Yes | No | Yes | No | Yes | Yes |
| \mathbb{R}^2 | 0.000603 | 0.00766 | 0.00110 | 0.00774 | 0.00145 | 0.00806 | 0.00170 | 0.00827 | 0.0236 |

Table 8: Pairwise correlations in liquidity

| | | Ι | Dependent V | ariable: Futu | re Monthly | Correlation of | of Delta Ami | hud | |
|----------------------------|----------|----------|-------------|---------------|------------|----------------|--------------|------------|-----------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| Same Group | 0.0116** | -0.00482 | | | 0.00482 | -0.00853* | 0.0103 | -0.00595 | -0.00739 |
| | (2.76) | (-1.64) | | | (1.17) | (-2.49) | (1.89) | (-1.32) | (-1.85) |
| FCA* | | | 0.00650*** | 0.00303*** | 0.00585*** | 0.00363*** | 0.00634*** | 0.00384*** | 0.00289** |
| | | | (6.09) | (4.52) | (6.03) | (4.31) | (6.04) | (4.26) | (2.89) |
| $(FCA^*) \times SameGroup$ | | | | | | | -0.00586* | -0.00274 | -0.00162 |
| | | | | | | | (-2.42) | (-1.10) | (-0.70) |
| Observations | 377863 | 369768 | 377863 | 369768 | 377863 | 369768 | 377863 | 369768 | 369768 |
| Group Effect | No | No | No | No | No | No | No | No | Yes |
| Controls | No | Yes | No | Yes | No | Yes | No | Yes | Yes |
| R^2 | 0.000586 | 0.00615 | 0.000681 | 0.00610 | 0.00117 | 0.00654 | 0.00136 | 0.00673 | 0.0220 |

t statistics in parentheses * $p < 0.05, \,^{**}$ $p < 0.01, \,^{***}$ p < 0.001

t statistics in parentheses * p < 0.05, ** p < 0.01, *** p < 0.001

 $[\]begin{array}{c} t \text{ statistics in parentheses} \\ ^*p < 0.05, \ ^{**}p < 0.01, \ ^{***}p < 0.001 \end{array}$

Figure 1: Time series of average common ownership measure with 95 percent interval for all pairs $\frac{1}{2}$

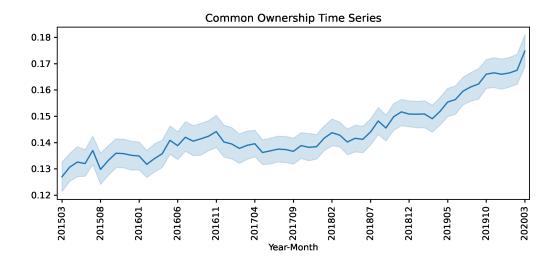


Figure 2: Time series of average common ownership measure with 95 percent interval in pairs in the same business group and others

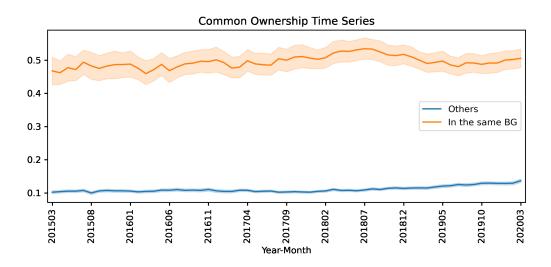


Figure 3: Time series of average common ownership measure with 95 percent interval which is grouped based on pairs' size

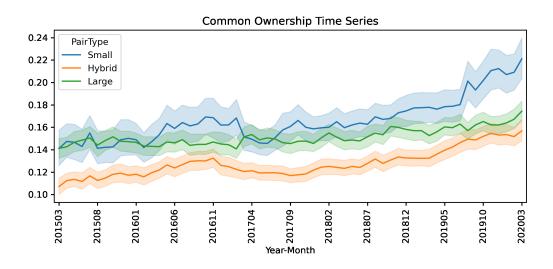


Figure 4: Percent of group affiliated firms from listed firms

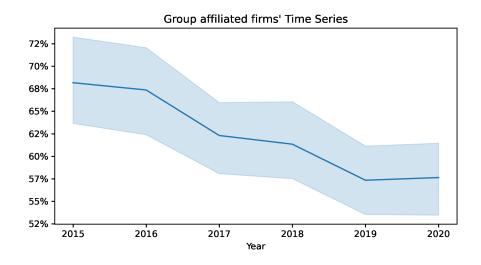


Figure 5: Percent of group affiliated firms from marketcap

