

### Exercise 3

1. Suppose  $X \sim \text{Pois}(\lambda)$   $\lambda > 0$ , hence  $P(X=x) = \frac{e^{-\lambda} \lambda^x}{x!}$ .

(a)  $M_X(t) = E[e^{tx}] = \sum_{x=0}^{\infty} e^{tx} \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!} = e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t-1)}$

(b) Define the cumulant  $n$  as  $K_n = \frac{d^n}{dt^n} \ln M_X(t) \Big|_{t=0}$

(c)  $\forall n \in \mathbb{N} \quad \frac{d^n}{dt^n} \ln e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda \frac{d^n}{dt^n} (e^t - 1) \Big|_{t=0} = \lambda e^t \Big|_{t=0} = \lambda$

(d)  $m_1 = \frac{d}{dt} e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda e^t e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda = K_1$

$m_2 = \frac{d^2}{dt^2} e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda e^t e^{\lambda(e^t-1)} + \lambda^2 e^{2t} e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda + \lambda^2 \rightarrow m_2 - m_1^2 = \lambda = K_2$

$m_3 = \frac{d^3}{dt^3} e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda e^t e^{\lambda(e^t-1)} + \lambda^2 e^{2t} e^{\lambda(e^t-1)} + 2\lambda^2 e^{2t} e^{\lambda(e^t-1)} + \lambda^3 e^{3t} e^{\lambda(e^t-1)} \Big|_{t=0} = \lambda + 3\lambda^2 + \lambda^3 \rightarrow$   
 $(\lambda + 3\lambda^2 + \lambda^3) - 3\lambda(\lambda + \lambda^2) + 2\lambda^3 = \lambda = K_3$

(e)  $K_1 = m_1 = \lambda$

$K_2 = m_2 - m_1^2 \rightarrow \lambda = m_2 - \lambda^2 \rightarrow m_2 = \lambda^2 + \lambda$

$K_3 = m_3 - 3m_1 m_2 + 2m_1^3 \rightarrow \lambda = m_3 - 3\lambda(\lambda^2 + \lambda) + 2\lambda^3 \rightarrow m_3 = \lambda^3 + 3\lambda^2 + \lambda$

2. Suppose  $X$  is a random variable

(a) Let  $M_X(t) = e^{ct}$  where  $c \in \mathbb{R}$ , then  $E[e^{tx}] = e^{ct}$  implies that  $X=c$ , meaning that  $F_X(x) = \begin{cases} 0 & x < c \\ 1 & x \geq c \end{cases}$

(b) Let  $M_X(t) = \frac{1}{2} + \frac{1}{4}e^t + \frac{1}{4}e^{-t}$ . This points to a discrete PMF like  $\sum_{x=-1}^3 e^{tx} P(X=x) = e^{t \cdot 0} \frac{1}{2} + e^{t \cdot 1} \frac{1}{4} + e^{t \cdot (-1)} \frac{1}{4}$ ,

meaning that the PMF is  $P(X=x) = \begin{cases} 1/4 & x=-1 \\ 1/2 & x=0 \\ 1/4 & x=1 \end{cases}$  and the CDF is  $F_X(x) = \begin{cases} 0 & x < -1 \\ 1/4 & -1 \leq x < 0 \\ 3/4 & 0 \leq x < 1 \\ 1 & 1 \leq x \end{cases}$ .

(c) Let  $M_X(t) = \frac{e^t}{2-e^t}$  where  $t < \ln 2$ . Then  $e^t < 2$  and  $p = \frac{1}{2}e^t < 1$ . Hence,  $\frac{e^t}{2-e^t} = \frac{2p}{2-2p} = \frac{p}{1-p} =$

$p \sum_{k=0}^{\infty} p^k = \sum_{k=1}^{\infty} p^k = \sum_{k=1}^{\infty} \frac{e^{kt}}{2^k}$ . This again points to a discrete PMF.  $e^{t \cdot 1} \frac{1}{2} + e^{t \cdot 2} \frac{1}{2^2} + e^{t \cdot 3} \frac{1}{2^3} + \dots$  which matches the PMF

$P(X=x) = \frac{1}{2^x}$  with the support being  $X \geq 1$ . This is a valid PMF since  $\sum_{k=1}^{\infty} \frac{1}{2^k} = \frac{1}{2} \frac{1}{1-1/2} = 1$ .

Hence  $\forall x \geq 1 \quad F_X(x) = \sum_{k=1}^{\lfloor x \rfloor} \frac{1}{2^k} = \frac{1}{2} \frac{1 - \frac{1}{2^{\lfloor x \rfloor}}}{1 - \frac{1}{2}} = 1 - \frac{1}{2^{\lfloor x \rfloor}}$  and  $F_X(x) = 0$  otherwise.

3. Let  $y > 0$  and  $g(y) = \frac{1}{y}$  then  $\frac{d^2}{dy^2} \frac{1}{y} = \frac{2}{y^3} > 0$ , meaning that  $g$  is a convex function on  $y > 0$ . Let  $X$  a positive random variable, then according to Jensen's inequality  $E[g(X)] \geq g(E[X]) \rightarrow E\left[\frac{1}{X}\right] \geq \frac{1}{E[X]}$ .

4. Suppose  $X$  is a random variable,  $n \in \mathbb{N}$ , and  $P_X(X=x_i) = \frac{1}{n}$  for  $i=1, \dots, n$ . Since  $\frac{d^2}{dx^2} \ln x = -\frac{1}{x^2} < 0$ , then  $\ln x$  is concave and according to Jensen's inequality  $\ln E[X] \geq E[\ln X]$ . Hence  $\ln \frac{1}{n} \sum_{k=1}^n x_k \geq \frac{1}{n} \sum_{k=1}^n \ln x_k \rightarrow$   
 $\ln \frac{1}{n} \sum_{k=1}^n x_k \geq \ln \left( \prod_{k=1}^n x_k \right)^{1/n}$ . Since  $\ln$  is an increasing function we can conclude that  $\frac{1}{n} \sum_{k=1}^n x_k \geq \left( \prod_{k=1}^n x_k \right)^{1/n}$ .

5. Suppose  $X$  is a random variable whose MGF is  $M_X(t)$  defined  $\forall t \in \mathbb{R}$ .

(a) Let  $t \geq 0$ ,  $Y = e^{tX}$ , then according to Markov inequality  $P(Y \geq y) \leq \frac{E[Y]}{y} \rightarrow P(e^{tX} \geq e^{tx}) \leq \frac{E[e^{tX}]}{e^{tx}}$ .

Since  $e^{tx}$  is increasing, this is also equivalent to  $P(X \geq x) \leq \frac{E[e^{tx}]}{e^{tx}}$ . In general,  $P(X \geq x) \leq \min_{t \geq 0} \frac{E[e^{tx}]}{e^{tx}}$ .  
 But  $\min_{t \geq 0} \frac{E[e^{tx}]}{e^{tx}} = \min_{t \geq 0} e^{\ln M_X(t) - tx} = e^{\min_{t \geq 0} (\ln M_X(t) - tx)} = e^{-\max_{t \geq 0} (tx - \ln M_X(t))}$ , hence  $P(X \geq x) \leq e^{-\Lambda(x)}$   
 where  $\Lambda(x) = \max_{t \geq 0} (tx - \ln M_X(t))$ .

(b) Let  $X_1, X_2, \dots, X_n$  iid's,  $X_i \sim B(p)$  and  $\bar{X}$  their arithmetic mean.  $M_{X_1} = E[e^{tX_1}] = e^{t \cdot 1} \cdot p + e^{t \cdot 0} (1-p) = p(e^t - 1) + 1$   
 $M_{\bar{X}} = E[e^{t\bar{X}}] = E[e^{\frac{t}{n} \sum X_i}] = E[\prod_{i=1}^n e^{\frac{t}{n} X_i}] = E[e^{\frac{t}{n} X_1}]^n = (M_{X_1}(\frac{t}{n}))^n$ . Hence  $\Lambda(x) = \max_{t \geq 0} (tx - n \cdot \ln(M_{X_1}(\frac{t}{n}))) = \max_{t \geq 0} (tx - n \cdot \ln(p(e^{\frac{t}{n}} - 1) + 1))$ .  $\frac{d}{dt} (tx - n \cdot \ln(p(e^{\frac{t}{n}} - 1) + 1)) = x - \frac{pe^{\frac{t}{n}}}{p(e^{\frac{t}{n}} - 1) + 1} = 0 \rightarrow xpe^{\frac{t}{n}} - xp + x = pe^{\frac{t}{n}} \rightarrow e^{\frac{t}{n}} = \frac{xp - x}{xp - p} \rightarrow t = n \cdot \ln \frac{x(p-1)}{p(x-1)}$ . Hence,  $\Lambda(x) = xn \cdot \ln \frac{x(p-1)}{p(x-1)} - n \cdot \ln \frac{p-1}{x-1}$  and  $P(\bar{X} \geq x) \leq e^{-\Lambda(x)}$ .

6. Suppose  $X$  is a continuous random variable with strictly increasing CDF  $F_X(x)$  and  $V \sim U(0,1)$ .

(a) Let  $X \sim U(0, a)$  for  $a > 0$ , then  $f_X(x) = \frac{1}{a}$ ,  $F_X(x) = \int_0^x \frac{1}{a} dt = \frac{x}{a}$ , and  $X = F_X^{-1}(v) = av$ .

```
import numpy as np
import matplotlib.pyplot as plt

def plot_histograms(
    origin_samples,
    target_samples,
    origin_label="Origin Samples",
    target_label="Target Samples",
    main_title="Sample Comparison",
    bins=30,
):
    """
    Plots normalized histograms for two sets of samples on separate subplots,
    without legends within each subplot.
    """
    fig, axes = plt.subplots(2, 1, figsize=(8, 7), sharex=False)
    fig.suptitle(main_title, fontsize=14)

    # Plot origin histogram
    axes[0].hist(origin_samples, bins=bins, density=True, alpha=0.7, color="salmon")
    axes[0].set_title(origin_label)
    axes[0].set_ylabel("Density")
    axes[0].grid(axis="y", linestyle="--", alpha=0.7)

    # Plot target histogram
    axes[1].hist(target_samples, bins=bins, density=True, alpha=0.7, color="skyblue")
    axes[1].set_title(target_label)
    axes[1].set_xlabel("Value")
    axes[1].set_ylabel("Density")
    axes[1].grid(axis="y", linestyle="--", alpha=0.7)

    plt.tight_layout(rect=[0, 0.03, 1, 0.95])
    plt.show()

# ---- Part A ----

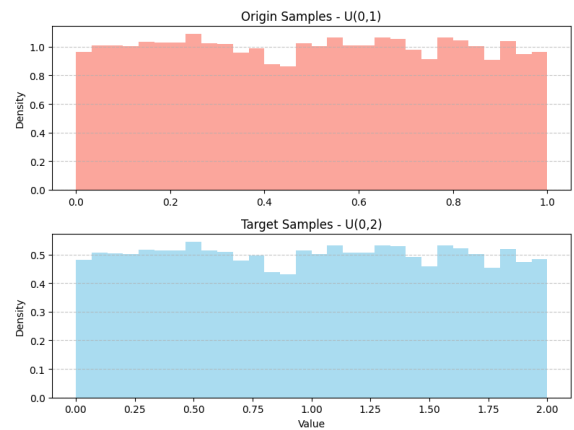
# Parameters for the target distribution
a = 2
N = 10000
INVERSE_UNI_CDF = lambda v: a * v

# Generate uniform samples from U(0,1)
uni_samples = np.random.uniform(0, 1, N)

# Generate samples from the target distribution using inverse transform sampling
target_samples = INVERSE_UNI_CDF(uni_samples)

# Plot the histogram of the generated samples
plot_histograms(
    uni_samples,
    target_samples,
    origin_label="Origin Samples - U(0,1)",
    target_label="Target Samples - U(0,2)",
    main_title="Histograms of 10^4 Transformed Samples from U(0,1) to U(0,2)",
    bins=30,
)
```

Histograms of 10<sup>4</sup> Transformed Samples from U(0,1) to U(0,2)



(b) Let  $X \sim \text{Exp}(1)$ , then  $f_X(x) = e^{-x}$ ,  $Y = F_X(x) = \int_0^x e^{-t} dt = -[e^{-t}]_0^x = 1 - e^{-x}$ , and  $Y \sim U(0,1)$ .

For sampling  $X$  just like in part (a), we also need  $F$ 's inverse:  $F^{-1}(v) = -\ln(1-v)$ .

```

# Parameters for the target distribution
N = 10000
EXP_CDF = lambda x: np.where(x >= 0, 1 - np.exp(-x), 0)
EXP_CDF_INV = lambda v: -np.log(1 - v) # Inverse CDF for the exponential distribution

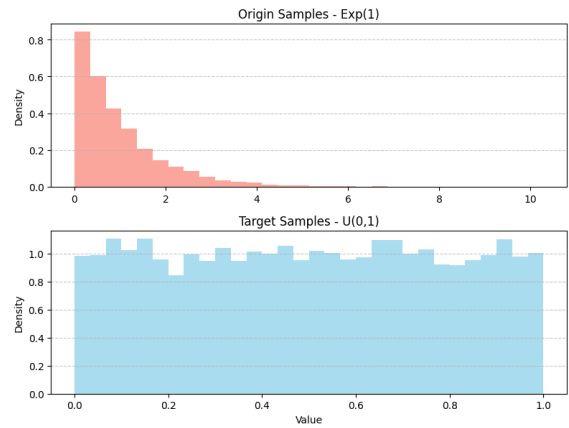
# Generate uniform samples from U(0,1)
uni_samples = np.random.uniform(0, 1, N)

# Generate samples from the exponential distribution
exp_samples = EXP_CDF_INV(uni_samples)
# Transform the samples to U(0,1) using the CDF of the exponential distribution
uni_samples = EXP_CDF(exp_samples)

# Plot the histograms of the original and transformed samples
plot_histograms(
    exp_samples,
    uni_samples,
    origin_label="Origin Samples - Exp(1)",
    target_label="Target Samples - U(0,1)",
    main_title="Histograms of 10^4 Transformed Samples from Exponential to U(0,1)",
    bins=30,
)

```

Histograms of 10<sup>4</sup> Transformed Samples from Exponential to U(0,1)



(c) Let  $X \sim \text{Gamma}(2, 4)$ , then  $f_X(x) = \frac{4^2}{\Gamma(2)} x^{2-1} e^{-4x} = 16x e^{-4x}$ ,  $F_X(x) = 16 \int_0^x t e^{-4t} dt = -4e^{-4t} \left( t + \frac{1}{4} \right) \Big|_0^x = 1 - 4e^{-4x} \left( x + \frac{1}{4} \right)$ .

Since we need a numeric approach for finding  $F_X^{-1}(v)$ , we would generate the initial  $\text{Gamma}(2, 4)$  sample using a numpy function. Let  $Y \sim \text{Beta}(1, 5)$ , then  $f_Y(y) = \frac{\Gamma(1+5)}{\Gamma(1)\Gamma(5)} y^{1-1} (1-y)^{5-1} = 5(1-y)^4$ ,  $F_Y(y) = 5 \int_0^y (1-t)^4 dt = -\frac{1}{5} (1-t)^5 \Big|_0^y = 1 - (1-y)^5$ , and  $F_Y^{-1}(v) = 1 - \sqrt[5]{1-v}$ .

```

# Parameters for the target distribution
N = 10000
GAMMA_CDF = lambda x: np.where(x >= 0, 1 - 4 * np.exp(-4 * x) * (x + 0.25), 0)
BETA_CDF_INV = lambda v: 1 - (1 - v) ** (1 / 5)

# Generate samples from the Gamma(2,4) distribution
gamma_samples = np.random.gamma(shape=2, scale=0.25, size=N)

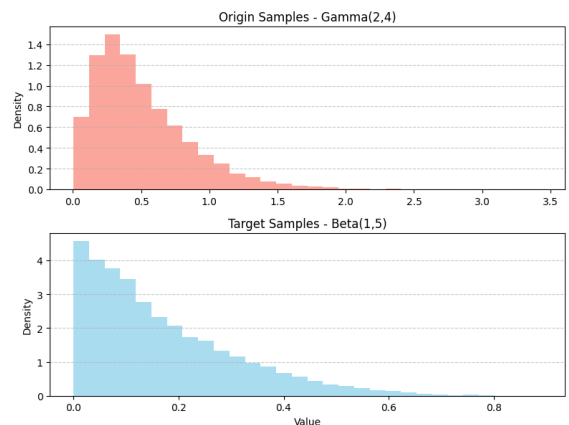
# Transform the samples to U(0,1) using the CDF of the gamma distribution
uni_samples = GAMMA_CDF(gamma_samples)

# Transform the samples to Beta(1, 5)
beta_samples = BETA_CDF_INV(uni_samples)

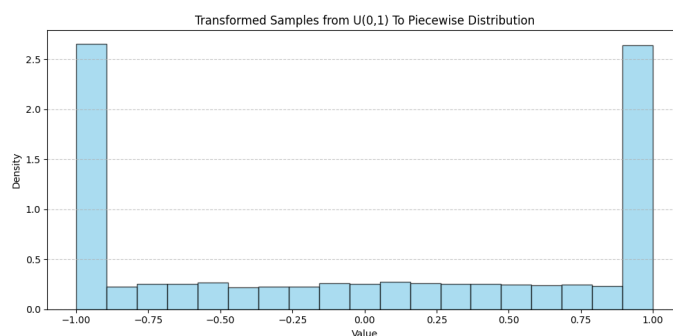
# Plot the histograms of the original and transformed samples
plot_histograms(
    gamma_samples,
    beta_samples,
    origin_label="Origin Samples - Gamma(2,4)",
    target_label="Target Samples - Beta(1,5)",
    main_title="Histograms of 10^4 Transformed Samples from Gamma to Beta",
    bins=30,
)

```

Histograms of 10<sup>4</sup> Transformed Samples from Gamma to Beta



7. Suppose  $X$  is a random variable whose CDF is  $F_X(x) = \begin{cases} 0 & x < -1 \\ \frac{x}{4} + \frac{1}{2} & -1 \leq x < 1 \\ 1 & 1 \leq x \end{cases}$ , then  $F_X^{-1}(v) = \begin{cases} -1 & 0 \leq v < 1/4 \\ 4v - 2 & 1/4 \leq v < 3/4 \\ 1 & 3/4 \leq v \leq 1 \end{cases}$ . In order to sample from this piecewise CDF we define it using numpy's piecewise function, sample from a  $U(0,1)$  distribution, and apply the inverse piecewise function on the uniform distribution sample array.





```
import numpy as np
import matplotlib.pyplot as plt

def plot_histogram(
    samples,
    title="Sample Distribution",
    xlabel="Value",
    ylabel="Density",
    num_bins=20,
    color="skyblue",
    edge_color="black",
    figsize=(10, 5),
):
    """
    Plots a single normalized histogram for a set of samples with distinct borders.
    """
    fig, ax = plt.subplots(1, 1, figsize=figsize)

    # Creates num_bins bins covering the interval [-1, 1]
    bin_edges = np.linspace(-1, 1, num_bins + 1)

    # Plot the histogram with density normalization and edge color
    ax.hist(samples, bins=bin_edges, density=True, alpha=0.7,
            color=color, edgecolor=edge_color) # Added edgecolor
    ax.set_title(title, fontsize=12)
    ax.set_xlabel(xlabel)
    ax.set_ylabel(ylabel)
    ax.grid(axis="y", linestyle="--", alpha=0.7)

    # Set x-axis limits slightly wider to see the end bins clearly
    ax.set_xlim(-1.1, 1.1)

    plt.tight_layout()
    plt.show()

# --- Define the piecewise inverse CDF ---
def inverse_cdf(v):
    """
    Calculates the value of the piecewise inverse CDF:
    F_X(v) =
        -1,      if 0 <= v < 1/4
        4v - 2,   if 1/4 <= v < 3/4
        1,        if 3/4 <= v <= 1
    """
    # Ensure v is a numpy array for vectorized operations
    v = np.asarray(v)

    # Define conditions for each piece
    conditions = [
        (v >= 0) & (v < 0.25), # 0 <= v < 1/4
        (v >= 0.25) & (v < 0.75), # 1/4 <= v < 3/4
        (v >= 0.75) & (v <= 1) # 3/4 <= v <= 1
    ]

    # Define the corresponding functions/values for each condition
    functions = [
        lambda x: -1.0, # Result is -1
        lambda x: 4 * x - 2, # Result is 4v - 2
        lambda x: 1.0 # Result is 1
    ]

    # Use numpy.piecewise to apply the conditions and functions
    return np.piecewise(v, conditions, functions)

# ---- Main ----

# generate uniform samples from u(0,1)
N = 10000
uni_samples = np.random.uniform(0, 1, N)

# generate samples from the target distribution using inverse transform sampling
target_samples = inverse_cdf(uni_samples)

# plot the histogram of the generated samples
plot_histogram(
    target_samples,
    title="Transformed Samples from U(0,1) To Piecewise Distribution",
)
```