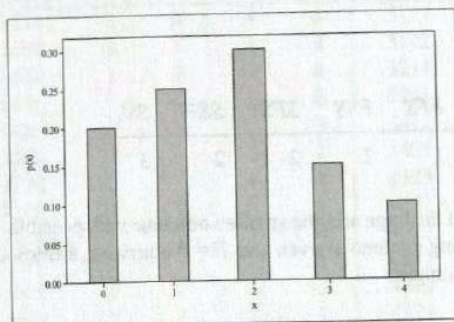


### Section 3.2

11.

a.



b.  $P(X \geq 2) = p(2) + p(3) + p(4) = .30 + .15 + .10 = .55$ , while  $P(X > 2) = .15 + .10 = .25$ .

c.  $P(1 \leq X \leq 3) = p(1) + p(2) + p(3) = .25 + .30 + .15 = .70$ .

d. Who knows? (This is just a little joke by the author.)

13.

a.  $P(X \leq 3) = p(0) + p(1) + p(2) + p(3) = .10 + .15 + .20 + .25 = .70$ .

b.  $P(X < 3) = P(X \leq 2) = p(0) + p(1) + p(2) = .45$ .

c.  $P(X \geq 3) = p(3) + p(4) + p(5) + p(6) = .55$ .

d.  $P(2 \leq X \leq 5) = p(2) + p(3) + p(4) + p(5) = .71$ .

e. The number of lines not in use is  $6 - X$ , and  $P(2 \leq 6 - X \leq 4) = P(-4 \leq -X \leq -2) = P(2 \leq X \leq 4) = p(2) + p(3) + p(4) = .65$ .

f.  $P(6 - X \geq 4) = P(X \leq 2) = .10 + .15 + .20 = .45$ .

15.

a.  $(1,2) (1,3) (1,4) (1,5) (2,3) (2,4) (2,5) (3,4) (3,5) (4,5)$

b.  $X$  can only take on the values 0, 1, 2.  $p(0) = P(X = 0) = P(\{(3,4) (3,5) (4,5)\}) = 3/10 = .3$ ;  $p(2) = P(X = 2) = P(\{(1,2)\}) = 1/10 = .1$ ;  $p(1) = P(X = 1) = 1 - [p(0) + p(2)] = .60$ ; and otherwise  $p(x) = 0$ .

c.  $F(0) = P(X \leq 0) = P(X = 0) = .30$ ;  
 $F(1) = P(X \leq 1) = P(X = 0 \text{ or } 1) = .30 + .60 = .90$ ;  
 $F(2) = P(X \leq 2) = 1$ .

Therefore, the complete cdf of  $X$  is



$$F(x) = \begin{cases} 0 & x < 0 \\ .30 & 0 \leq x < 1 \\ .90 & 1 \leq x < 2 \\ 1 & 2 \leq x \end{cases}$$

- 17.
- $p(2) = P(Y = 2) = P(\text{first 2 batteries are acceptable}) = P(AA) = (.9)(.9) = .81.$
  - $p(3) = P(Y = 3) = P(UAA \text{ or } AUA) = (.1)(.9)^2 + (.1)(.9)^2 = 2[(.1)(.9)^2] = .162.$
  - The fifth battery must be an  $A$ , and exactly one of the first four must also be an  $A$ .  
Thus,  $p(5) = P(AUUUA \text{ or } UAUUA \text{ or } UUAUA \text{ or } UUUAA) = 4[(.1)^3(.9)^2] = .00324.$
  - $p(y) = P(\text{the } y^{\text{th}} \text{ is an } A \text{ and so is exactly one of the first } y-1) = (y-1)(.1)^{y-2}(.9)^2, \text{ for } y = 2, 3, 4, 5, \dots$
- 19.
- $p(0) = P(Y = 0) = P(\text{both arrive on Wed}) = (.3)(.3) = .09;$   
 $p(1) = P(Y = 1) = P((W, Th) \text{ or } (Th, W) \text{ or } (Th, Th)) = (.3)(.4) + (.4)(.3) + (.4)(.4) = .40;$   
 $p(2) = P(Y = 2) = P((W, F) \text{ or } (Th, F) \text{ or } (F, W) \text{ or } (F, Th) \text{ or } (F, F)) = .32;$   
 $p(3) = 1 - [.09 + .40 + .32] = .19.$
- 21.
- First,  $1 + 1/x > 1$  for all  $x = 1, \dots, 9$ , so  $\log(1 + 1/x) > 0$ . Next, check that the probabilities sum to 1:  
 $\sum_{x=1}^9 \log_{10}(1 + 1/x) = \sum_{x=1}^9 \log_{10}\left(\frac{x+1}{x}\right) = \log_{10}\left(\frac{2}{1}\right) + \log_{10}\left(\frac{3}{2}\right) + \dots + \log_{10}\left(\frac{10}{9}\right);$  using properties of logs,  
 this equals  $\log_{10}\left(\frac{2}{1} \times \frac{3}{2} \times \dots \times \frac{10}{9}\right) = \log_{10}(10) = 1.$
  - Using the formula  $p(x) = \log_{10}(1 + 1/x)$  gives the following values:  $p(1) = .301, p(2) = .176, p(3) = .125, p(4) = .097, p(5) = .079, p(6) = .067, p(7) = .058, p(8) = .051, p(9) = .046$ . The distribution specified by Benford's Law is not uniform on these nine digits; rather, lower digits (such as 1 and 2) are much more likely to be the lead digit of a number than higher digits (such as 8 and 9).
  - The jumps in  $F(x)$  occur at  $0, \dots, 8$ . We display the cumulative probabilities here:  $F(1) = .301, F(2) = .477, F(3) = .602, F(4) = .699, F(5) = .778, F(6) = .845, F(7) = .903, F(8) = .954, F(9) = 1$ . So,  $F(x) = 0$  for  $x < 1$ ;  $F(x) = .301$  for  $1 \leq x < 2$ ;  $F(x) = .477$  for  $2 \leq x < 3$ ; etc.
  - $P(X \leq 3) = F(3) = .602; P(X \geq 5) = 1 - P(X < 5) = 1 - P(X \leq 4) = 1 - F(4) = 1 - .699 = .301.$
- 23.
- $p(2) = P(X = 2) = F(3) - F(2) = .39 - .19 = .20.$
  - $P(X > 3) = 1 - P(X \leq 3) = 1 - F(3) = 1 - .67 = .33.$
  - $P(2 \leq X \leq 5) = F(5) - F(2-1) = F(5) - F(1) = .92 - .19 = .78.$
  - $P(2 < X < 5) = P(2 < X \leq 4) = F(4) - F(2) = .92 - .39 = .53.$



33.

$$\text{a. } E(X^2) = \sum_{x=0}^1 x^2 \cdot p(x) = 0^2(1-p) + 1^2(p) = p.$$

$$\text{b. } V(X) = E(X^2) - [E(X)]^2 = p - [p]^2 = p(1-p).$$

$$\text{c. } E(X^{79}) = 0^{79}(1-p) + 1^{79}(p) = p. \text{ In fact, } E(X^n) = p \text{ for any non-negative power } n.$$

35.

Let  $h_3(X)$  and  $h_4(X)$  equal the net revenue (sales revenue minus order cost) for 3 and 4 copies purchased, respectively. If 3 magazines are ordered (\$6 spent), net revenue is \$4 - \$6 = -\$2 if  $X = 1$ ,  $2(\$4) - \$6 = \$2$  if  $X = 2$ ,  $3(\$4) - \$6 = \$6$  if  $X = 3$ , and also \$6 if  $X = 4, 5$ , or 6 (since that additional demand simply isn't met. The values of  $h_4(X)$  can be deduced similarly. Both distributions are summarized below.

$x$	1	2	3	4	5	6
$h_3(x)$	-2	2	6	6	6	6
$h_4(x)$	-4	0	4	8	8	8
$p(x)$	$\frac{1}{15}$	$\frac{2}{15}$	$\frac{3}{15}$	$\frac{4}{15}$	$\frac{3}{15}$	$\frac{2}{15}$

$$\text{Using the table, } E[h_3(X)] = \sum_{x=1}^6 h_3(x) \cdot p(x) = (-2)(\frac{1}{15}) + \dots + (6)(\frac{2}{15}) = \$4.93.$$

$$\text{Similarly, } E[h_4(X)] = \sum_{x=1}^6 h_4(x) \cdot p(x) = (-4)(\frac{1}{15}) + \dots + (8)(\frac{2}{15}) = \$5.33.$$

Therefore, ordering 4 copies gives slightly higher revenue, on the average.

37.

$$\text{Using the hint, } E(X) = \sum_{x=1}^n x \cdot \left(\frac{1}{n}\right) = \frac{1}{n} \sum_{x=1}^n x = \frac{1}{n} \left[ \frac{n(n+1)}{2} \right] = \frac{n+1}{2}. \text{ Similarly,}$$

$$E(X^2) = \sum_{x=1}^n x^2 \cdot \left(\frac{1}{n}\right) = \frac{1}{n} \sum_{x=1}^n x^2 = \frac{1}{n} \left[ \frac{n(n+1)(2n+1)}{6} \right] = \frac{(n+1)(2n+1)}{6}, \text{ so}$$

$$V(X) = \frac{(n+1)(2n+1)}{6} - \left(\frac{n+1}{2}\right)^2 = \frac{n^2-1}{12}.$$

39.

From the table,  $E(X) = \sum xp(x) = 2.3$ ,  $E(X^2) = 6.1$ , and  $V(X) = 6.1 - (2.3)^2 = .81$ . Each lot weighs 5 lbs, so the number of pounds left =  $100 - 5X$ . Thus the expected weight left is  $E(100 - 5X) = 100 - 5E(X) = 88.5$  lbs, and the variance of the weight left is  $V(100 - 5X) = V(-5X) = (-5)^2 V(X) = 25V(X) = 20.25$ .

41.

$$\text{Use the hint: } V(aX + b) = E[(aX + b) - E(aX + b)]^2 = \sum [ax + b - E(aX + b)]^2 p(x) =$$

$$\sum [ax + b - (a\mu + b)]^2 p(x) = \sum [ax - a\mu]^2 p(x) = a^2 \sum (x - \mu)^2 p(x) = a^2 V(X).$$

43.

$$\text{With } a = 1 \text{ and } b = -c, E(X - c) = E(aX + b) = aE(X) + b = E(X) - c.$$

When  $c = \mu$ ,  $E(X - \mu) = E(X) - \mu = \mu - \mu = 0$ ; i.e., the expected deviation from the mean is zero.



### Chapter 3: Discrete Random Variables and Probability Distributions

45.  $a \leq X \leq b$  means that  $a \leq x \leq b$  for all  $x$  in the range of  $X$ . Hence  $ap(x) \leq xp(x) \leq bp(x)$  for all  $x$ , and
- $$\sum ap(x) \leq \sum xp(x) \leq \sum bp(x)$$
- $$a \sum p(x) \leq \sum xp(x) \leq b \sum p(x)$$
- $$a \cdot 1 \leq E(X) \leq b \cdot 1$$
- $$a \leq E(X) \leq b$$

#### Section 3.4

- 47.
- $B(4; 15, .7) = .001$ .
  - $b(4; 15, .7) = B(4; 15, .7) - B(3; 15, .7) = .001 - .000 = .001$ .
  - Now  $p = .3$  (multiple vehicles).  $b(6; 15, .3) = B(6; 15, .3) - B(5; 15, .3) = .869 - .722 = .147$ .
  - $P(2 \leq X \leq 4) = B(4; 15, .7) - B(1; 15, .7) = .001$ .
  - $P(2 \leq X) = 1 - P(X \leq 1) = 1 - B(1; 15, .7) = 1 - .000 = 1$ .
  - The information that 11 accidents involved multiple vehicles is redundant (since  $n = 15$  and  $x = 4$ ). So, this is actually identical to **b**, and the answer is .001.
49. Let  $X$  be the number of "seconds," so  $X \sim \text{Bin}(6, .10)$ .
- $P(X = 1) = \binom{n}{x} p^x (1-p)^{n-x} = \binom{6}{1} (.1)^1 (.9)^5 = .3543$ .
  - $P(X \geq 2) = 1 - [P(X = 0) + P(X = 1)] = 1 - \left[ \binom{6}{0} (.1)^0 (.9)^6 + \binom{6}{1} (.1)^1 (.9)^5 \right] = 1 - [.5314 + .3543] = .1143$ .
  - Either 4 or 5 goblets must be selected.  
 Select 4 goblets with zero defects:  $P(X = 0) = \binom{4}{0} (.1)^0 (.9)^4 = .6561$ .  
 Select 4 goblets, one of which has a defect, and the 5<sup>th</sup> is good:  $\left[ \binom{4}{1} (.1)^1 (.9)^3 \right] \times .9 = .26244$   
 So, the desired probability is  $.6561 + .26244 = .91854$ .
51. Let  $X$  be the number of faxes, so  $X \sim \text{Bin}(25, .25)$ .
- $E(X) = np = 25(.25) = 6.25$ .
  - $V(X) = np(1-p) = 25(.25)(.75) = 4.6875$ , so  $SD(X) = 2.165$ .
  - $P(X > 6.25 + 2(2.165)) = P(X > 10.58) = 1 - P(X \leq 10.58) = 1 - P(X \leq 10) = 1 - B(10; 25, .25) = .030$ .



### Chapter 3: Discrete Random Variables and Probability Distributions

53. Let "success" = has at least one citation and define  $X$  = number of individuals with at least one citation. Then  $X \sim \text{Bin}(n = 15, p = .4)$ .
- If at least 10 have no citations (failure), then at most 5 have had at least one (success):  
 $P(X \leq 5) = B(5; 15, .40) = .403$ .
  - Half of 15 is 7.5, so less than half means 7 or fewer:  $P(X \leq 7) = B(7; 15, .40) = .787$ .
  - $P(5 \leq X \leq 10) = P(X \leq 10) - P(X \leq 4) = .991 - .217 = .774$ .
55. Let "success" correspond to a telephone that is submitted for service while under warranty and must be replaced. Then  $p = P(\text{success}) = P(\text{replaced} | \text{submitted}) \cdot P(\text{submitted}) = (.40)(.20) = .08$ . Thus  $X$ , the number among the company's 10 phones that must be replaced, has a binomial distribution with  $n = 10$  and  $p = .08$ , so  $P(X = 2) = \binom{10}{2} (.08)^2 (.92)^8 = .1478$ .
57. Let  $X$  = the number of flashlights that work, and let event  $B = \{\text{battery has acceptable voltage}\}$ . Then  $P(\text{flashlight works}) = P(\text{both batteries work}) = P(B)P(B) = (.9)(.9) = .81$ . We have assumed here that the batteries' voltage levels are independent. Finally,  $X \sim \text{Bin}(10, .81)$ , so  $P(X \geq 9) = P(X = 9) + P(X = 10) = .285 + .122 = .407$ .
59. In this example,  $X \sim \text{Bin}(25, p)$  with  $p$  unknown.
- $P(\text{rejecting claim when } p = .8) = P(X \leq 15 \text{ when } p = .8) = B(15; 25, .8) = .017$ .
  - $P(\text{not rejecting claim when } p = .7) = P(X > 15 \text{ when } p = .7) = 1 - P(X \leq 15 \text{ when } p = .7) = 1 - B(15; 25, .7) = 1 - .189 = .811$ .  
 For  $p = .6$ , this probability is  $1 - B(15; 25, .6) = 1 - .575 = .425$ .
  - The probability of rejecting the claim when  $p = .8$  becomes  $B(14; 25, .8) = .006$ , smaller than in **a** above. However, the probabilities of **b** above increase to .902 and .586, respectively. So, by changing 15 to 14, we're making it less likely that we will reject the claim when it's true ( $p$  really is  $\geq .8$ ), but more likely that we'll "fail" to reject the claim when it's false ( $p$  really is  $< .8$ ).
61. If topic A is chosen, then  $n = 2$ . When  $n = 2$ ,  $P(\text{at least half received}) = P(X \geq 1) = 1 - P(X = 0) = 1 - \binom{2}{0} (.9)^0 (.1)^2 = .99$ .
- If topic B is chosen, then  $n = 4$ . When  $n = 4$ ,  $P(\text{at least half received}) = P(X \geq 2) = 1 - P(X \leq 1) = 1 - \left[ \binom{4}{0} (.9)^0 (.1)^4 + \binom{4}{1} (.9)^1 (.1)^3 \right] = .9963$ .
- Thus topic B should be chosen if  $p = .9$ .
- However, if  $p = .5$ , then the probabilities are .75 for A and .6875 for B (using the same method as above), so now A should be chosen.
- 63.
- $b(x; n, 1 - p) = \binom{n}{x} (1 - p)^x (p)^{n-x} = \binom{n}{n-x} (p)^{n-x} (1 - p)^x = b(n-x; n, p)$ .
- Conceptually,  $P(x \text{ S's when } P(S) = 1 - p) = P(n-x \text{ F's when } P(F) = p)$ , since the two events are identical, but the labels S and F are arbitrary and so can be interchanged (if  $P(S)$  and  $P(F)$  are also interchanged), yielding  $P(n-x \text{ S's when } P(S) = 1 - p)$  as desired.



### Chapter 3: Discrete Random Variables and Probability Distributions

- b. Use the conceptual idea from a:  $B(x; n, 1-p) = P(\text{at most } x \text{ S's when } P(S) = 1-p) = P(\text{at least } n-x \text{ F's when } P(F) = p)$ , since these are the same event  $= P(\text{at least } n-x \text{ S's when } P(S) = p)$ , since the S and F labels are arbitrary  $= 1 - P(\text{at most } n-x-1 \text{ S's when } P(S) = p) = 1 - B(n-x-1; n, p)$ .
- c. Whenever  $p > .5$ ,  $(1-p) < .5$  so probabilities involving  $X$  can be calculated using the results a and b in combination with tables giving probabilities only for  $p \leq .5$ .
- 65.
- a. Although there are three payment methods, we are only concerned with S = uses a debit card and F = does not use a debit card. Thus we can use the binomial distribution. So, if  $X$  = the number of customers who use a debit card,  $X \sim \text{Bin}(n = 100, p = .2)$ . From this,  $E(X) = np = 100(.2) = 20$ , and  $V(X) = npq = 100(.2)(1-.2) = 16$ .
- b. With S = doesn't pay with cash,  $n = 100$  and  $p = .7$ , so  $\mu = np = 100(.7) = 70$ , and  $V = 21$ .
67. When  $n = 20$  and  $p = .5$ ,  $\mu = 10$  and  $\sigma = 2.236$ , so  $2\sigma = 4.472$  and  $3\sigma = 6.708$ . The inequality  $|X - 10| \geq 4.472$  is satisfied if either  $X \leq 5$  or  $X \geq 15$ , or  $P(|X - \mu| \geq 2\sigma) = P(X \leq 5 \text{ or } X \geq 15) = .021 + .021 = .042$ . The inequality  $|X - 10| \geq 6.708$  is satisfied if either  $X \leq 3$  or  $X \geq 17$ , so  $P(|X - \mu| \geq 3\sigma) = P(X \leq 3 \text{ or } X \geq 17) = .001 + .001 = .002$ .

### Section 3.5

69. According to the problem description,  $X$  is hypergeometric with  $n = 6$ ,  $N = 12$ , and  $M = 7$ .
- a.  $P(X = 4) = \frac{\binom{7}{4} \binom{5}{2}}{\binom{12}{6}} = \frac{350}{924} = .379$ .  $P(X \leq 4) = 1 - P(X > 4) = 1 - [P(X = 5) + P(X = 6)] = 1 - \left[ \frac{\binom{7}{5} \binom{5}{1}}{\binom{12}{6}} + \frac{\binom{7}{6} \binom{5}{0}}{\binom{12}{6}} \right] = 1 - [.114 + .007] = 1 - .121 = .879$ .
- b.  $E(X) = n \cdot \frac{M}{N} = 6 \cdot \frac{7}{12} = 3.5$ ;  $V(X) = \left( \frac{12-6}{12-1} \right) 6 \left( \frac{7}{12} \right) \left( 1 - \frac{7}{12} \right) = 0.795$ ;  $\sigma = 0.892$ . So,  $P(X > \mu + \sigma) = P(X > 3.5 + 0.892) = P(X > 4.392) = P(X = 5 \text{ or } 6) = .121$  (from part a).
- c. We can approximate the hypergeometric distribution with the binomial if the population size and the number of successes are large. Here,  $n = 15$  and  $M/N = 40/400 = .1$ , so  $h(x; 15, 40, 400) \approx b(x; 15, .10)$ . Using this approximation,  $P(X \leq 5) \approx B(5; 15, .10) = .998$  from the binomial tables. (This agrees with the exact answer to 3 decimal places.)