

Elastic Full Procrustes Means for Sparse and Irregular Planar Curves

Masters Thesis

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1 Introduction

In statistical shape analysis one is interested in analysing the shape of geometrical objects, like the outlines of bones and organs, a handwritten digit, or the folds of a protein. When analysing such objects, differences in location, rotation, and size are often not of interest. Instead, the focus lies purely on their differences in *shape*, a widely adapted definition of which was established by KENDALL 1977 and which might be formulated in the following way:

Definition 1.1 (Shape). *All geometrical information that remains when location, scale and rotational effects are removed from an object (see DRYDEN and MARDIA 2016, p. 1).*

This geometrical information is usually approximated by measuring the coordinates of a fixed set of *landmarks*, which are characteristic points on an object that match between and within populations (see DRYDEN and MARDIA 2016, p. 3). As these coordinates depend on the scale, rotation and translation of the object at time of measurement, only the *equivalence class* of the landmark configuration modulo these transformations is indicative of its shape, which will be defined in more detail in section ?? . A popular type of shape mean that does not depend on the rotation, scaling and translation of the input shapes is the *full Procrustes mean*. Here, the mean is defined as the minimizer of a least squares criterion, using a distance measure that is invariant under the mentioned transformations, which will be discussed in section ?? .

Instead of approximating the geometry of an object using landmarks, its whole outline might be represented by an open or closed curve $\beta : [0, 1] \rightarrow \mathbb{R}^d, d \in \mathbb{N}$, eliminating the (often subjective) decision of which points to consider as "characteristic". This functional data approach introduces an additional kind of shape invariance relating to the parametrisation $t \in [0, 1]$ of β , as only the image of β but not its parametrisation is indicative of the objects shape. A functional shape mean, which is invariant with respect to re-parametrisation of the input curves, is called an *elastic mean* and its calculation is greatly simplified by working in the *square-root-velocity* (SRV) framework as introduced in SRIVASTAVA, KLASSEN, et al. 2011. Elastic shape means and the SRV framework will be explained in section ?? . Even though this functional approach eliminates the subjectivity of choosing a fixed set of landmarks, in practice a curve will never be fully observed and one has to again work with a set of sampled

points along the curve. A particular challenge is working with such curves, that are sparsely and irregularly sampled, where one has to employ appropriate smoothing techniques to make maximum use of the available (sparse) data. This will be briefly discussed in section ??.

The aim of this thesis is to bring all these concepts together in the estimation of “Elastic Full Procrustes Means for Sparse and Irregular Planar Curves”. The thesis focuses on the special case of 2D (i.e. planar) curves, as it can be shown that the Procrustes mean has particularly nice properties in this settings. Additionally this thesis will be mainly concerned with the mean estimation for open curves, as the closed curves case is more challenging mathematically. After covering the relevant background material in sufficient detail in section 2, an expression for the elastic full Procrustes mean will be derived in section 3. In section 4, an estimation strategy for the setting of sparse and irregular curves is proposed, which will be applied to simulated and empirical datasets in section 5. Finally, in section 6 a possible extension to closed curves will be briefly discussed. All results will be briefly summarized in section 7.

2 Functional and Shape Data Analysis

2.1 Shape Invariance and Equivalence Classes

Before beginning with further derivations and definitions it is important to establish a mathematical framework for treating shapes.

Definition 2.1 (Equivalence relation, equivalence class and quotient space). A relation \sim on a set X is called an **equivalence relation** if, for all $x, y, z \in X$, it has the following properties:

- i. $x \sim x$ (reflexivity)
- ii. $x \sim y \Rightarrow y \sim x$ (symmetry)
- iii. $x \sim y, y \sim z \Rightarrow x \sim z$ (transitivity)

The **equivalence class** $[x]$ of $x \in X$ is given by the set of all $y \in X$ so that $x \sim y$. The **quotient space** X/\sim of X under the relation \sim is the disjoint set of all equivalence classes in X (see SRIVASTAVA and KLASSEN 2016, p. 40).

This means, that any statistical analysis of shapes should ultimately work with these equivalence classes as well.

2.2 The Full Procrustes Mean in Classical Shape Analysis

Definition 2.2 (Full Procrustes distance, full Procrustes mean). For X_1, X_2 landmark configurations, represented as $m \times d$ matrices with m landmarks in d dimensions, the **full Procrustes distance** between their shapes $[X_1], [X_2]$ is defined as

$$d_F([X_1], [X_2]) = \inf_{\lambda \in \mathbb{R}_+, \Gamma \in SO(m)} \|\tilde{X}_1 - \lambda \tilde{X}_2 \Gamma\|,$$

where $\tilde{X}_{1,2}$ are centered and normalized landmark configurations, $\lambda \in \mathbb{R}_+$ is a scaling factor and $\Gamma \in SO(d)$ a rotation matrix.

The **full Procrustes mean** shape for a sample of landmark configurations X_i ($i = 1, \dots, n$) is then given by the equivalence class $[\hat{\mu}_F]$ of a landmark configuration that

minimizes the sum of squared full Procrustes distances

$$\hat{\mu}_F = \operatorname{arginf}_{\mu} \sum_{i=1}^n d_F([\mu], [X_i])^2,$$

where μ is assumed centered and normalized (see DRYDEN and MARDIA 2016, pp. 71, 114).

2.3 Elastic Means and the Square-Root-Velocity Framework

[How is functional shape data analyzed]

2.4 Functional Data Analysis of Sparse and Irregular Planar Curves

[What is sparse and irregular data]

3 The Elastic Full Procrustes Means for Planar Curves

3.1 Group action: Warping and Euclidean similarity transform

Let β be a continuous planar curve. It can be represented in a parameterized form in \mathbb{R}^2 as

$$\beta : [0, 1] \rightarrow \mathbb{R}^2, \quad \beta(t) = (x(t), y(t)),$$

where x, y are scalar-valued *coordinate functions* of β , parametrized by t . We can equivalently represent a planar curve using complex numbers as

$$\beta : [0, 1] \rightarrow \mathbb{C}, \quad \beta(t) = x(t) + iy(t),$$

with the added benefit that complex notation often simplifies calculations in the 2D case.

For a set of planar curves $\beta_1, \dots, \beta_n : [0, 1] \rightarrow \mathbb{C}$, either centered with $\langle \beta_i, \mathbb{1} \rangle$ or with no relative translation to each other, the *full Procrustes mean* $\hat{\mu}$ is then defined as the curve minimizing the sum of squared *full Procrustes distances* from each β_i to an unknown unit size mean configuration μ , that is

$$\begin{aligned} \hat{\mu} &= \operatorname{argmin}_{\mu: [0,1] \rightarrow \mathbb{C}} \sum_{i=1}^n d_F^2(\mu, \beta_i) \quad \text{s.t. } \|\mu\| = 1 \\ &= \operatorname{argmin}_{\mu: [0,1] \rightarrow \mathbb{C}} \sum_{i=1}^n 1 - \frac{\langle \mu, \beta_i \rangle \langle \beta_i, \mu \rangle}{\langle \mu, \mu \rangle \langle \beta_i, \beta_i \rangle} \quad \text{s.t. } \|\mu\| = 1 \end{aligned}$$

which we can be further simplified by normalizing $\beta_i := \frac{\beta_i}{\|\beta_i\|}$ and using $\langle \mu, \mu \rangle = 1$

$$\hat{\mu} = \operatorname{argmax}_{\mu: [0,1] \rightarrow \mathbb{C}} \sum_{i=1}^n \langle \mu, \beta_i \rangle \langle \beta_i, \mu \rangle \quad \text{s.t. } \|\mu\| = 1.$$

The expression for $d_F^2(\mu, \beta_i)$ in the case of planar curves is derived in appendix A.1.

3.2 The SRV Framework

Instead of working with the original curve β , for calculation of *elastic* means it is advantageous to work with its corresponding *square root velocity curve* given by

$$q : [0, 1] \rightarrow \mathbb{C}, \quad q(t) = \frac{\dot{\beta}(t)}{\sqrt{\|\dot{\beta}(t)\|}} \quad \text{for } \dot{\beta}(t) \neq 0$$

where original curve β can be obtained up to translation by back transformation via $\beta(t) = \beta(0) + \int_0^t q(s) \|q(s)\| ds$. Moreover, if the original curve is of unit length the SRV curve will be automatically normalized:

$$\|q\| = \sqrt{\langle q, q \rangle} = \sqrt{\int_0^1 \overline{q(t)} q(t) dt} = \sqrt{\int_0^1 |q(t)|^2 dt} = \sqrt{\int_0^1 |\dot{\beta}(t)| dt} = \sqrt{1} = 1.$$

3.3 The Full Procrustes mean

Consider a set of planar SRV curves $q_1, \dots, q_n : [0, 1] \rightarrow \mathbb{C}$ of unit length $\|q_i\| = 1$ for all i . The *full Procrustes mean* $\hat{\mu}$ is given by

$$\begin{aligned} \hat{\mu} &= \operatorname{argmax}_{\mu: [0,1] \rightarrow \mathbb{C}} \sum_{i=1}^n \langle \mu, q_i \rangle \langle q_i, \mu \rangle \quad \text{s.t. } \|\mu\| = 1 \\ &= \operatorname{argmax}_{\mu: [0,1] \rightarrow \mathbb{C}} \sum_{i=1}^n \int_0^1 \overline{\mu(t)} q_i(t) dt \int_0^1 \overline{q_i(s)} \mu(s) ds \quad \text{s.t. } \|\mu\| = 1 \\ &= \operatorname{argmax}_{\mu: [0,1] \rightarrow \mathbb{C}} \int_0^1 \int_0^1 \overline{\mu(t)} \underbrace{\left(\sum_{i=1}^n q_i(t) \overline{q_i(s)} \right)}_{:= n\hat{C}(s,t)} \mu(s) dt ds \quad \text{s.t. } \|\mu\| = 1 \\ &= \operatorname{argmax}_{\mu: [0,1] \rightarrow \mathbb{C}} \int_0^1 \overline{\mu(t)} \int_0^1 \hat{C}(s,t) \mu(s) ds dt \quad \text{s.t. } \|\mu\| = 1 \end{aligned}$$

with the solution given by the eigenfunction corresponding to the largest eigenvector of the complex empirical covariance function $\hat{C}(s, t) = n^{-1} \sum_{i=1}^n q_i(t) \overline{q_i(s)}$.

3.4 The Full Procrustes Mean in a fixed basis

To avoid having to sample the estimated covariance surface $\hat{C}(s, t)$ on a large grid when calculating its leading eigenfunction, it might be preferable to calculate this eigenfunction from the vector of basis coefficients directly. After choosing a basis representation $b = (b_1, \dots, b_k)$ with $b_j : \mathbb{R} \rightarrow \mathbb{R}$ real-valued basis functions, we want to estimate complex coefficients $\theta_j \in \mathbb{C}$ so that the Full Procrustes mean of SRV curves is given by $\hat{\mu}(t) = \sum_{j=1}^k \hat{\theta}_j b_j(t) = b^T \hat{\theta}$:

$$\begin{aligned} \hat{\mu} &= \operatorname{argmax}_{\theta: \|b^T \theta\|=1} \sum_{i=1}^n \langle b^T \theta, q_i \rangle \langle q_i, b^T \theta \rangle \\ &= \operatorname{argmax}_{\theta: \|b^T \theta\|=1} \sum_{k,l} \sum_{i=1}^n \langle b_k \theta_k, q_i \rangle \langle q_i, b_l \theta_l \rangle \\ &= \operatorname{argmax}_{\theta: \|b^T \theta\|=1} \sum_{k,l} \bar{\theta}_k \theta_l \sum_{i=1}^n \langle b_k, q_i \rangle \langle q_i, b_l \rangle \\ &= \operatorname{argmax}_{\theta: \|b^T \theta\|=1} \theta^H S \theta \end{aligned}$$

where the matrix $S = \{\sum_{i=1}^n \langle b_k, q_i \rangle \langle q_i, b_l \rangle\}_{k,l}$ has to be estimated from the observed SRV curves. We can further simplify S to

$$\begin{aligned} S_{kl} &= \sum_{i=1}^n \int_0^1 \bar{b}_k(t) q_i(t) dt \int_0^1 \bar{q}_i(s) b_l(s) ds \\ &= \int_0^1 \int_0^1 \bar{b}_k(t) \underbrace{\left(\sum_{i=1}^n q_i(t) \bar{q}_i(s) \right)}_{=n \hat{C}(s,t)} b_l(s) ds dt \\ &= n \int_0^1 \int_0^1 \bar{b}_k(t) \hat{C}(s, t) b_l(s) ds dt \end{aligned}$$

with $\hat{C}(s, t) = \frac{1}{n} \sum_{i=1}^n q_i(s) \overline{q_i(t)}$ the sample analogue to the complex population covariance function $C(s, t) = \mathbb{E}[q(s) \overline{q(t)}]$. We may estimate $C(s, t)$ via tensor product splines, so that $\hat{C}(s, t) = \sum_{k,l} \hat{\xi}_{kl} b_k(t) \bar{b}_l(s)$, where $b_j(t)$, $j = 1, \dots, k$ are the same real valued basis functions as used for the mean and $\hat{\xi}_{kl}$ are the estimated complex coefficients.

We can then further simplify S_{kl}

$$\begin{aligned}
S_{kl} &= n \int_0^1 \int_0^1 b_k(t) \left(\sum_{p,q} \hat{\xi}_{pq} b_q(t) b_p(s) \right) b_l(s) ds dt \\
&= n \sum_{p,q} \hat{\xi}_{pq} \int_0^1 \int_0^1 b_k(t) b_q(t) b_p(s) b_l(s) ds dt \\
&= n \sum_{p,q} \hat{\xi}_{pq} \langle b_k, b_q \rangle \langle b_p, b_l \rangle \\
&= n \sum_{p,q} \hat{\xi}_{pq} g_{kq} g_{pl}
\end{aligned}$$

where g_{ij} , $i, j = 1, \dots, k$ are the elements of the Gram matrix $G = bb^T$ with $G = \mathbb{I}_k$ in the special case of an orthogonal basis. We can then write the matrix S as a function of the estimated coefficient matrix $\hat{\Xi} = (\hat{\xi}_{ij})_{i,j=1,\dots,k}$:

$$S = n G \hat{\Xi} G$$

The full Procrustes mean of SRV curves is then given by the solution to the optimization problem

$$\begin{aligned}
\hat{\mu} &= \underset{\theta}{\operatorname{argmax}} n \theta^H G \hat{\Xi} G \theta \quad \text{subj. to} \quad \|b^T \theta\| = 1 \\
&= \underset{\theta: \|b^T \theta\|=1}{\operatorname{argmax}} \theta^H G \hat{\Xi} G \theta \quad \text{subj. to} \quad \theta^H G \theta = 1
\end{aligned}$$

One may solve this by using Lagrange optimization with the Langrangian

$$\mathcal{L}(\theta, \lambda) = \theta^H G \hat{\Xi} G \theta - \lambda(\theta^H G \theta - 1)$$

3.5 Estimation of the covariance surface $C(s, t)$

Consider the following model for independent curves

$$Y_i(t_{ij}) = \mu(t_{ij}, \mathbf{x}_i) + E_i(t_{ij}) + \epsilon(t_{ij}), \quad j = 1, \dots, D_i, i = 1, \dots, n, \quad (1)$$

[Fast symmetric additive cov smoothing, skew-symmetry, population vs. sample, etc.]

4 Estimation Strategy

5 Empirical Application

6 Outlook

7 Summary

A Derivations and Proofs

A.1 Derivation of the Full Procrustes Distance for Functional Data

Consider two curves $\beta_1, \beta_2 : [0, 1] \rightarrow \mathbb{C}$ with $\langle \beta_1, \mathbb{1} \rangle = \langle \beta_2, \mathbb{1} \rangle = 0$ where $\mathbb{1}$ is the constant function $\mathbb{1}(t) = 1$ for all $t \in [0, 1]$. Then β_1 and β_2 can be considered to be centered as

$$\langle \beta_1, \mathbb{1} \rangle = \int_0^1 \bar{\beta}_1(t) \mathbb{1}(t) dt = \int_0^1 \bar{\beta}_1(t) dt = \int_0^1 (y(t) + ix(t)) dt = \underbrace{\int_0^1 y(t) dt}_{\stackrel{!}{=} 0} + i \underbrace{\int_0^1 x(t) dt}_{\stackrel{!}{=} 0} = 0$$

Then the full procrustes distance of β_1, β_2 is given by their minimum distance controlling for translation $\gamma \in \mathbb{C}$, and scaling and rotation $\omega = be^{i\theta} \in \mathbb{C}$:

$$\begin{aligned} d_F^2 &= \min_{\omega, \gamma \in \mathbb{C}} \|\beta_1 - \gamma \mathbb{1} - \omega \beta_2\|^2 \\ &= \min_{\omega, \gamma \in \mathbb{C}} \langle \beta_1 - \gamma \mathbb{1} - \omega \beta_2, \beta_1 - \gamma \mathbb{1} - \omega \beta_2 \rangle \\ &= \min_{\omega, \gamma \in \mathbb{C}} \langle \beta_1 - \omega \beta_2, \beta_1 - \omega \beta_2 \rangle - \underbrace{\langle \beta_1, \gamma \mathbb{1} \rangle}_{=0} - \underbrace{\langle \gamma \mathbb{1}, \beta_1 \rangle}_{=0} + \underbrace{\langle \gamma \mathbb{1}, \omega \beta_2 \rangle}_{=0} + \underbrace{\langle \omega \beta_2, \gamma \mathbb{1} \rangle}_{=0} + \underbrace{\langle \gamma \mathbb{1}, \gamma \mathbb{1} \rangle}_{=||\gamma \mathbb{1}||^2} \\ &\stackrel{\gamma=0}{=} \min_{\omega \in \mathbb{C}} \langle \beta_1, \beta_1 \rangle + \langle \omega \beta_2, \omega \beta_2 \rangle - \langle \beta_1, \omega \beta_2 \rangle - \langle \omega \beta_2, \beta_1 \rangle \\ &= \min_{\omega \in \mathbb{C}} \langle \beta_1, \beta_1 \rangle + |\omega|^2 \langle \beta_2, \beta_2 \rangle - \omega \langle \beta_1, \beta_2 \rangle - \bar{\omega} \langle \beta_2, \beta_1 \rangle \end{aligned}$$

To find $\omega \in \mathbb{C}$ that minimizes $\|\beta_1 - \omega \beta_2\|^2$ we first consider the part of the problem dependent on θ . We need to solve

$$\min_{\omega \in \mathbb{C}} -\omega \langle \beta_1, \beta_2 \rangle - \bar{\omega} \langle \beta_2, \beta_1 \rangle = \max_{\omega \in \mathbb{C}} \omega \langle \beta_1, \beta_2 \rangle + \bar{\omega} \langle \beta_2, \beta_1 \rangle$$

by using $\omega = be^{i\theta}$ and $\langle \beta_1, \beta_2 \rangle = ae^{i\phi}$:

$$\begin{aligned} \max_{\omega \in \mathbb{C}} \omega \langle \beta_1, \beta_2 \rangle + \bar{\omega} \langle \beta_2, \beta_1 \rangle &= \max_{b \in \mathbb{R}^+, \theta \in [0, 2\pi]} be^{i\theta} ae^{i\phi} + be^{-i\theta} ae^{-i\phi} \\ &= \max_{b \in \mathbb{R}^+, \theta \in [0, 2\pi]} be^{i\theta} ae^{i\phi} + be^{-i\theta} ae^{-i\phi} \\ &= \max_{b \in \mathbb{R}^+, \theta \in [0, 2\pi]} 2ba \cos(\theta + \phi) \end{aligned}$$

$$\stackrel{\theta=-\phi}{=} \max_{b \in \mathbb{R}^+} 2ba$$

and using $\theta = -\phi$ the original minimization problem therefore simplifies to

$$\begin{aligned} d_F^2 &= \min_{b \in \mathbb{R}^+} \langle \beta_1, \beta_1 \rangle + b^2 \langle \beta_2, \beta_2 \rangle - 2ba \\ \frac{\partial d_F^2}{\partial b} &= 2b \langle \beta_2, \beta_2 \rangle - 2a \stackrel{!}{=} 0 \\ \Rightarrow \quad b &= \frac{a}{\langle \beta_2, \beta_2 \rangle} \end{aligned}$$

And for the *full Procrustes distance* it follows that

$$d_F^2 = \langle \beta_1, \beta_1 \rangle - \frac{a^2}{\langle \beta_2, \beta_2 \rangle} = \langle \beta_1, \beta_1 \rangle - \frac{\langle \beta_1, \beta_2 \rangle \langle \beta_2, \beta_1 \rangle}{\langle \beta_2, \beta_2 \rangle}$$

As this expression is not symmetric in β_1 and β_2 we can take the curves to be of unit length with $\tilde{\beta}_j = \frac{\beta_j}{\|\beta_j\|}$, $j = 1, 2$ with $\|\beta_j\| = \sqrt{\langle \beta_j, \beta_j \rangle}$, so that $\langle \tilde{\beta}_1, \tilde{\beta}_1 \rangle = \langle \tilde{\beta}_2, \tilde{\beta}_2 \rangle = 1$ and obtain a suitable measure of distance:

$$d_F = \sqrt{1 - \langle \tilde{\beta}_1, \tilde{\beta}_2 \rangle \langle \tilde{\beta}_2, \tilde{\beta}_1 \rangle} = \sqrt{1 - \frac{\langle \beta_1, \beta_2 \rangle \langle \beta_2, \beta_1 \rangle}{\langle \beta_1, \beta_1 \rangle \langle \beta_2, \beta_2 \rangle}}$$

B Supplementary Materials

B.1 Implementation Notes

B.2 Replication Guide

C Math-Basics Recap

C.1 Scalar Products

V n -dimensional vector space with basis $B = (b_1, \dots, b_n)$, then any scalar product $\langle \cdot, \cdot \rangle$ on V can be expressed using a $(n \times n)$ matrix G , the Gram matrix of the scalar product. Its entries are the scalar products of the basis vectors:

$$G = (g_{ij})_{i,j=1,\dots,n} \quad \text{with} \quad g_{ij} = \langle b_i, b_j \rangle \quad \text{for} \quad i, j = 1, \dots, n$$

When vectors $x, y \in V$ are expressed with respect to the basis B as

$$x = \sum_{i=1}^n x_i b_i \quad \text{and} \quad y = \sum_{i=1}^n y_i b_i$$

the scalar product can be expressed using the Gram matrix, and in the complex case it holds that

$$\langle x, y \rangle = \sum_{i,j=1}^n \bar{x}_i y_j \langle b_i, b_j \rangle = \sum_{i,j=1}^n \bar{x}_i g_{ij} y_j = x^\dagger G y$$

when $x_i, y_i \in \mathbb{C}$ for $i = 1, \dots, n$ with x^\dagger indicating the conjugate transpose of $x = (x_1, \dots, x_n)^T$. If B is an *orthonormal* basis, that is if $\langle b_i, b_j \rangle = \delta_{ij}$, it further holds that $\langle x, y \rangle = x^\dagger y$ as $G = \mathbb{1}_{n \times n}$.

C.2 Functional Scalar Products

This concept can be generalized for vectors in function spaces. Define the scalar product of two functions $f(t), g(t)$ as:

$$\langle f, g \rangle = \int_a^b \bar{f}(t) w(t) g(t) dt$$

with weighting function $w(t)$ and $[a, b]$ depending on the function space. The scalar product has the following properties:

1. $\langle f, g + h \rangle = \langle f, g \rangle + \langle f, h \rangle$
2. $\langle f, g \rangle = \overline{\langle g, f \rangle}$

3. $\langle f, cg \rangle = c \langle f, g \rangle$ or, using (2), $\langle cf, g \rangle = \bar{c} \langle f, g \rangle$ for $c \in \mathbb{C}$

If we have a functional basis $\{\phi_1, \dots, \phi_n\}$ (and possibly $n \rightarrow \infty$) of our function space we can also write the function f as an expansion

$$f = \sum_{i=1}^n a_i \phi_i \quad \text{so that} \quad f(t) = \sum_{i=1}^n a_i \phi_i(t)$$

Additionally, if we have a *orthogonal* basis, so that $\langle \phi_i, \phi_j \rangle = 0$ for $i \neq j$, we can take the scalar product with ϕ_k from the left

$$\langle \phi_k, f \rangle = \sum_{i=1}^n a_i \langle \phi_k, \phi_i \rangle = a_k \langle \phi_k, \phi_k \rangle$$

which yields the coefficients a_k :

$$a_k = \frac{\langle \phi_k, f \rangle}{\langle \phi_k, \phi_k \rangle}$$

For an *orthonormal* basis it holds that $\langle \phi_i, \phi_j \rangle = \delta_{ij}$. Suppose that two functions f, g are expanded in the same orthonormal basis:

$$f = \sum_{i=1}^n a_i \phi_i \quad \text{and} \quad g = \sum_{i=1}^n b_i \phi_i$$

We can then write the scalar product as:

$$\langle f, g \rangle = \left\langle \sum_{i=1}^n a_i \phi_i, \sum_{i=1}^n b_i \phi_i \right\rangle = \sum_{i=1}^n \sum_{j=1}^n a_i b_j \langle \phi_i, \phi_j \rangle = \sum_{i=1}^n a_i b_i = a^\dagger b$$

for coefficient vectors $a, b \in \mathbb{C}^n$. This means that the functional scalar product reduces to a complex dot product. Additionally it holds that for the norm $\|\cdot\|$ of a function f :

$$\|f\| = \langle f, f \rangle^{\frac{1}{2}} = \sqrt{a^\dagger a} = \sqrt{\sum_{i=1}^n |a_i|^2}$$

D FDA-Basics Recap

As discussed in the last section we can express a function f in its *basis function expansion* using a set of basis functions ϕ_k with $k = 1, \dots, K$ and a set of coefficients c_1, \dots, c_K

(both possibly \mathbb{C} valued e.g. in the case of 2D-curves)

$$f = \sum_{k=1}^K c_k \phi_k = \mathbf{c}' \boldsymbol{\phi}$$

where in the matrix notation \mathbf{c} and $\boldsymbol{\phi}$ are the vectors containing the coefficients and basis functions.

When considering a sample of N functions f_i we can write this in matrix notation as

$$\mathbf{f} = \mathbf{C} \boldsymbol{\phi}$$

where \mathbf{C} is a $(N \times K)$ matrix of coefficients and \mathbf{f} is a vector containing the N functions.

D.1 Smoothing by Regression

When working with functional data we can usually never observe a function f directly and instead only observe discrete points (x_i, t_i) along the curve, with $f(t_i) = x_i$. As we don't know the exact functional form of f , calculating the scalar products $\langle \phi_k, f \rangle$ and therefore calculating the coefficients c_k of a given basis representation is not possible.

However, we can estimate the basis coefficients using e.g. regression analysis an approach motivated by the error model

$$f(t_i) = \mathbf{c}' \boldsymbol{\phi}(t_i) + \epsilon_i$$

If we observe our function n times at t_1, \dots, t_n , we can estimate the coefficients from a least squares problem, where we try to minimize the deviation of the basis expansion from the observed values. Using matrix notation let the vector \mathbf{f} contains the observed values $f(t_i)$, $i = 1, \dots, n$ and $(n \times k)$ matrix $\boldsymbol{\Phi}$ contains the basis function values $\phi_k(t_i)$. Then we have

$$\mathbf{f} = \boldsymbol{\Phi} \mathbf{c} + \boldsymbol{\epsilon}$$

with the estimate for the coefficient vector \mathbf{c} given by

$$\hat{\mathbf{c}} = (\boldsymbol{\Phi}' \boldsymbol{\Phi})^{-1} \boldsymbol{\Phi}' \mathbf{f}.$$

Spline curves fit in this way are often called *regression splines*.

D.2 Common Basis Representations

Piecewise Polynomials (Splines) Splines are defined by their range of validity, the knots, and the order. They are constructed by dividing the area of observation into subintervals with boundaries at points called *breaks*. Over any subinterval the spline function is a polynomial of fixed degree or order. The term *degree* refers to the highest power in the polynomial while its *order* is one higher than its degree. E.g. a line has degree one but order two because it also has a constant term. [...]

Polygonal Basis [...]

D.3 Bivariate Functional Data

The analogue of covariance matrices in MVA are covariance surfaces $\sigma(s, t)$ whose values specify the covariance between values $f(s)$ and $f(t)$ over a population of curves. We can write these bivariate functions in a *bivariate basis expansion*

$$r(s, t) = \sum_{k=1}^K \sum_{l=1}^K b_{k,l} \phi_k(s) \psi_l(t) = \boldsymbol{\phi}(s)' \mathbf{B} \boldsymbol{\psi}(t)$$

with a $K \times K$ coefficient matrix B and two sets of basis functions ϕ_k and ψ_l using *Tensor Product Splines*

$$B_{k,l}(s, t) = \phi_k(s) \psi_l(t).$$

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