### Stochastic Variational Inference

MPRI — Cours 2.40 "Probabilistic Programming Languages"

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### Probabilistic Programming Languages Inference

#### Inference

#### Evaluation of the posterior probability distribution

#### Caveats:

- exact inference either impossible or not tractable in general
- many techniques for approximate inference several of them discussed in the previous classes

### Today's plan: stochastic variational inference (SVI)

- intuition and definition how to formalize the semantics of SVI
- issues: it may not work, unless some conditions are satisfied questions: how to characterize these conditions, how to check them?

We consider **Pyro**, a probabilistic programming language implemented over Python with support for SVI, available at https://pyro.ai. Formalisation with a subset of the course language

### Outline

- Introduction
- 2 Basic Intuition Underlying SVI
- Semantics for SVI
- 4 SVI
- **(5)** Ensuring correctness of SVI
- Conclusion

### Evaluation of probabilistic programs

# What does it mean to evaluate a probabilstic program ? How to do that ?

#### Enumeration:

- run all the executions and compute the probability of each of them
- only works in the discrete case...

### Importance weighting:

- compute a family of executions together with their probability
- executions are chosen at random (hopefully well)

### Rejection sampling:

- compute a family of executions
- reject unlikely executions, accumulate those with high probability

#### Variational inference:

- search for a simpler program that is close enough
- ... more on this today...

### A first, very basic model

We build a model step by step, to construct an example for SVI.

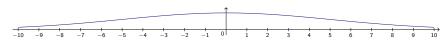
#### Model Pyro code:

```
def model():
    x = pyro.sample("v", Normal(0., 5.))
```

#### Meaning:

- sample: draws a value based on a distribution in this case, normal distribution, mean 0, standard deviation 5
- i.e., values of variable random v (python variable x) distributed around 0 with some imprecision

#### **Distribution** over executions based on the final value of x:



## A second, more interesting model

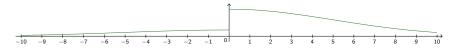
### Model Pyro code:

```
def model():
    x = pyro.sample("v", Normal(0., 5.))
    if (x > 0):
        pyro.sample("obs", Normal(1., 1.), obs=x)
    else:
        pyro.sample("obs", Normal(-2., 1.), obs=x)
```

### Meaning:

- sample without obs=...: sampling, as before
- sample with obs=...: conditioning determined by observation

#### Distribution on random variable v:



## Distribution defined by the model

### Model Pyro code:

```
def model():
    x = pyro.sample("v", Normal(0., 5.))
    if (x > 0):
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        pyro.sample("obs", Normal(-2., 1.), obs=x)
```

**Prior on random variable** v, before observation taken into account: *i.e.*, when observations on the value of obs are ignored

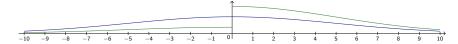


## Distribution defined by the model

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```

Posterior distribution on random variable v, after observations on observations and compared with the prior:

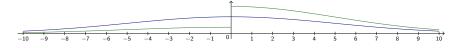


### Distribution defined by the model

### Model Pyro code:

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        pyro.sample("obs", Normal(-2., 1.), obs=x)
```

Posterior distribution on random variable v, after observations on observations and compared with the prior:



Can we discover a simpler, accurate enough approximation of the posterior, as a program ?

## Model approximation with a parameterized "guide"

#### SVI main Idea:

specify a template for a family of candidate functions to approximate the posterior, then choose among them the most suitable one

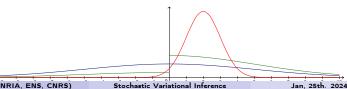
#### Guide

Companion program with randomized parameter, aimed at approximating the posterior distribution defined in the model

In our example: sampling the parameter from a normal distribution

```
def guide():
 xtheta = pyro.param("theta", 3.)
 x = pyro.sample("v", Normal(xtheta, 1.))
```

One instance of the guide, with a positive  $\theta$  (expected outcome)



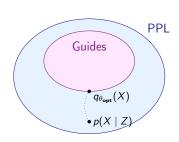
### Objective of variational inference

### SVI first question

How to define the optimization objective? How to define the best guide?

**Data**: probablistic program *P* called **model** samples variables *X*, observes variables *Z* 

- We fix the family of guides programs with special variables identified as parameters
- Probabilistic programs denote posterior probability distributions
  - $\triangleright$  p(X|Z) for model
  - $ightharpoonup q_{\theta}(X)$  for guide instance  $\theta$
- We fix a **distance over distributions** *d* in general, KL divergence



Objective: guide parameter  $\theta_{opt}$  instance that minimizes distance

## Stochastic computation of an approximation of the objective

### SVI second question

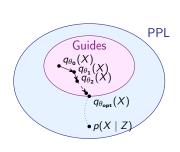
How to solve the optimization problem in practice ?

#### Gradient descent algorithm:

- Pick initial parameter  $\theta_0$
- **3** Estimate the gradient  $\nabla_{\theta}$  of the distance at point  $\theta_0$  and make a step in that direction, to compute  $\theta_1$
- **3** Repeat to compute  $\theta_2, \ldots$  and stop after K steps

#### **Stochastic approximation** of the gradient at $\theta$ :

- Sample N runs, based on the guide
- Compute an estimate gradient based on the average variation



Convergence: based on properties of the optimization problem

### Outline

- Semantics for SVI
  - Measure semantics
  - Density semantics

## Towards a stochastic approximation of the distance

#### **Main step** of the gradient descent algorithm:

⇒ estimation of the distance between distributions

High-level overview of a single step of gradient descent (full algorithm shown later):

- sample a number of executions (stochastic)
- for each sampled execution, compute probability density (for short, density)
- produce a gradient approximation based on the executions and their densities

#### How to compute the probability density of an expression?

In the following, we propose a density semantics that answers the question:

- first, we recall the kernel semantics:
   maps a state into a probability distribution over outcomes
- then, we construct a new semantics called **density semantics**, that produces not only output states but also probability density

**Implementation**: an evaluator following this density semantics (instrumented runtime)...

## A basic imperative probabilistic programming language

We define a minimalistic form of ByoPPL.

#### A few assumptions:

- imperative control structures (while language),
- real numbers (not floating point)
- only normal distributions in the minimalistic syntax below though examples may use other distributions as well
- countable set of random variables, represented with string names

#### Basic syntax:

```
E,B,S real, boolean, string expressions C::= commands | skip | C_0; C_1 | x:=E | if B\{C\} else \{C\} | while B\{C\} | x:= sample_{\mathcal{N}}(k,E_0,E_1) | k: random variable name, E_0: mean, E_1: standard dev. | observe_{\mathcal{N}}(E_0,E_1,E_2) | E_0: observed value, E_1: mean, E_2: standard dev.
```

### Towards two semantics

#### Common characteristics:

- executions record all random choices random choices are stored in a dictionary binding names to values
- states comprise a memory and a random dictionary

#### Example:

$$x := \operatorname{sample}_{\mathcal{N}}(\mathfrak{s0}, 0, 1);$$
  
$$y := \operatorname{sample}_{\mathcal{N}}(\mathfrak{s1}, 2, 1);$$

```
initial state: (s2 \mapsto 2), (x \mapsto 8, y \mapsto 9)
final state: (s0 \mapsto 0.5, s1 \mapsto 1.5, s2 \mapsto 2), (x \mapsto 0.5, y \mapsto 1.5)
```

- consequence 1: a given random variable is sampled at most once
- consequence 2: initially present unsampled variables remain in the dictionary
- consequence 3: sampling from a variable already in the dictionary is blocking

#### Our two semantics:

- Wernel semantics: maps states into probability distributions over states
- Oensity semantics: maps states into states with a probability density

### **Notations**

#### We fix the following definitions:

- X: set of program variables
- V: set of scalar values (we assume real numbers, not floating point)
- $\mathbb{M} = [\mathbb{X} \to \mathbb{V}]$ : set of **memory states** (notation:  $\mu \in \mathbb{M}$ )
- K: set of random variables, corresponding to strings
- $\mathbb{P} = [\mathbb{K} \to \mathbb{V}]$ : set of **random databases** (or random dictionaries) (notation:  $\rho \in \mathbb{P}$ )
- $\mathbb{S} = \mathbb{M} \times \mathbb{P}$ : set of states (notation:  $\sigma \in \mathbb{S}$ )
- given a measurable set  $\mathbb{A}$ , we note  $\mathcal{M}(\mathbb{A})$  for the set of measurable subsets of  $\mathbb{A}$

### Assumptions

We fix the following definitions: **Assumptions**:

- ullet usual structure of measurable space structures over  $\mathbb{M}$ ,  $\mathbb{P}$ , and  $\mathbb{S}$
- notion of probability kernel over A and measurable set A': function from A to probability distributions over A' (previous lecture)
  - ▶ probability kernel: total measure is 1, noted  $\mathcal{K}(\mathbb{A}, \mathbb{A}')$
  - ▶ sub-probability kernel: total measure  $\leq 1$ , noted  $\mathcal{K}(\mathbb{A}, \mathbb{A}')_{\mathsf{sub}}$

#### Semantics:

 we assume a semantics of expressions: for all express E:

$$\llbracket E 
rbracket : \mathbb{M} \to \mathbb{V}$$

(random variables not used)

Semantics of commands:

$$\llbracket \mathcal{C} 
rbracket_{\mathcal{M}} \in \mathcal{K}(\mathbb{S}, \mathbb{S} imes \mathbb{R}^+)_{\mathsf{sub}}$$

Or equivalently:

$$\llbracket \mathcal{C} 
rbracket_{\mathcal{M}} : \mathbb{S} o (\mathcal{M}(\mathbb{S} imes \mathbb{R}^+) o_{\mathcal{M}} [0,1]) \equiv (\mathbb{M} imes \mathbb{P}) o \mathcal{M}(\mathbb{M} imes \mathbb{P} imes \mathbb{R}^+) o_{\mathcal{M}} [0,1]$$

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Assignment statement x := E

$$\llbracket x := E \rrbracket_{\mathcal{M}}(\mu, \rho)(S) \triangleq \mathbb{1}_{\llbracket ((\mu[x \mapsto \llbracket E \rrbracket(\mu)], \rho), 1) \in S \rrbracket}$$

- weight is not modified
- variable x is updated in the store
- note: expressions should not read random variables directly

$$\llbracket \mathcal{C} 
rbracket_{\mathcal{M}} : \mathbb{S} o (\mathcal{M}(\mathbb{S} imes \mathbb{R}^+) o_{\mathcal{M}} \llbracket 0, 1 
bracket) \equiv (\mathbb{M} imes \mathbb{P}) o \mathcal{M}(\mathbb{M} imes \mathbb{P} imes \mathbb{R}^+) o_{\mathcal{M}} \llbracket 0, 1 
bracket$$

Assignment statement x := E

Sample statement sample<sub> $\mathcal{N}$ </sub>  $(s, E_0, E_1)$ 

$$\begin{split} \llbracket x := \operatorname{sample}_{\mathcal{N}}(k, E_1, E_2) \rrbracket_{\mathcal{M}}(\mu, \rho)(S) &\triangleq \\ \mathbb{1}_{[k \notin \operatorname{Dom}(\rho)]} \cdot \mathbb{1}_{\llbracket E_2 \rrbracket(\mu) \in \mathbb{R}^{+*} \rrbracket} \\ \cdot \int \operatorname{d}v \left( \operatorname{pdf}_{\mathcal{N}}(v; \llbracket E_1 \rrbracket(\mu), \llbracket E_2 \rrbracket(\mu)) \cdot \mathbb{1}_{\llbracket (\mu[x \mapsto v], \rho[k \mapsto v], 1) \in S} \right) \end{split}$$

- crashes when sampling from a rand. var. not in the random database
- crashes when standard deviation is negative
- otherwise updates the states and rdb with the new sample, integrate over the density of the sampled distribution and do this for all possible samples (hence the sum)

Notation:  $pdf_{\mathcal{N}}(v; m, d)$ : **probability density** at v of the normal distribution of mean m and standard deviation s

$$\llbracket \mathcal{C} 
rbracket_{\mathcal{M}} : \mathbb{S} o (\mathcal{M}(\mathbb{S} imes \mathbb{R}^+) o_{\mathcal{M}} \llbracket [0,1]) \equiv (\mathbb{M} imes \mathbb{P}) o \mathcal{M}(\mathbb{M} imes \mathbb{P} imes \mathbb{R}^+) o_{\mathcal{M}} \llbracket [0,1] 
bracket$$

Assignment statement x := E

Sample statement sample  $\mathcal{N}(s, E_0, E_1)$ 

**Score statement** observe $_{\mathcal{N}}(E_0, E_1, E_2)$ 

$$\begin{aligned} & [\![ \mathrm{observe}_{\mathcal{N}} (E_0, E_1, E_2)]\!]_{\mathcal{M}} (\mu, \rho)(S) \triangleq \\ & 1\!I_{[\![E_2]\!] (\mu) \in \mathbb{R}^{+*}]} \cdot 1\!I_{[((\mu, \rho), \mathsf{pdf}_{\mathcal{N}} ([\![E_0]\!] (\mu); [\![E_1]\!] (\mu), [\![E_2]\!] (\mu))) \in S]} \end{aligned}$$

- crashes when standard deviation is negative
- otherwise state left unmodified score the density of the distribution for the observed value

$$\llbracket C 
rbracket_{\mathcal{M}} : \mathbb{S} o (\mathcal{M}(\mathbb{S} imes \mathbb{R}^+) o_{\mathcal{M}} [0,1]) \equiv (\mathbb{M} imes \mathbb{P}) o \mathcal{M}(\mathbb{M} imes \mathbb{P} imes \mathbb{R}^+) o_{\mathcal{M}} [0,1]$$

Assignment statement x := E

Sample statement sample  $\mathcal{N}(s, E_0, E_1)$ 

**Score statement** observe  $\mathcal{N}(E_0, E_1, E_2)$ 

For all command C,  $[C]_M$  is measurable and defines a sub-probability kernel from  $\mathbb{S}$  to  $\mathbb{S} \times \mathbb{R}^+$ 

## Measure semantics (or kernel semantics)

```
[skip]_{\mathcal{M}}(\mu,\rho)(S)
             \triangleq \mathbb{1}_{[((\mu,\rho),1)\in S]}
[x := E]_{\mathcal{M}}(\mu, \rho)(S)
             \triangleq \mathbb{1}_{[((\mu[x\mapsto \llbracket E\rrbracket(\mu)],\rho),1)\in S]}
[C_0; C_1]_{\mathcal{M}}(\mu, \rho)(S)
             \triangleq \int [\![C_0]\!]_{\mathcal{M}}(\mu,\rho)(\mathrm{d}(\sigma_0,w_0)) \int [\![C_1]\!]_{\mathcal{M}}(\sigma_0)(\mathrm{d}(\sigma_1,w_1)) \mathbb{1}_{[(\sigma_1,w_0w_1)\in S]}
[if B \{ C_0 \} else\{ C_1 \} ] _{\mathcal{M}}(\mu, \rho)(S)
             \triangleq \mathbb{1}_{\lceil \llbracket B \rrbracket(\mu) = \mathsf{true} \rceil} \cdot \llbracket C_0 \rrbracket_{\mathcal{M}}(\mu, \rho)(S) + \mathbb{1}_{\lceil \llbracket B \rrbracket(\mu) = \mathsf{false} \rceil} \cdot \llbracket C_1 \rrbracket_{\mathcal{M}}(\mu, \rho)(S)
while B\{C\} M(\mu, \rho)(S)
             \triangleq (\text{Fix}F)(\mu, \rho)(S)
      where F(\phi)(\mu, \rho)(S) = \mathbb{1}_{[[B]](\mu) = false]} \cdot \mathbb{1}_{[((\mu, \rho), 1) \in S]}
                    +1_{[[B](\mu)=\text{true}]} \cdot \int [[C]_{\mathcal{M}}(\mu,\rho)(d(\sigma_0,w_0)) \int \phi(\sigma_0)(d(\sigma_1,w_1))1_{[(\sigma_1,w_0w_1)\in S]}
[x := \operatorname{sample}_{\mathcal{N}}(k, E_1, E_2)]_{\mathcal{M}}(\mu, \rho)(S)
             \triangleq \mathbb{1}_{[k \notin \mathrm{Dom}(\rho)]} \cdot \mathbb{1}_{[\llbracket E_2 \rrbracket(\mu) \in \mathbb{R}^{+*}]}
                                 \cdot \int d\mathbf{v}' \left( \mathsf{pdf}_{\mathcal{N}}(\mathbf{v}; [E_1][\mu), [E_2][\mu) \right) \cdot \mathbb{1}_{[(\mu[\mathsf{x}\mapsto \mathsf{v}], \rho[\mathsf{k}\mapsto \mathsf{v}], 1) \in S]} \right)
[observe_{\mathcal{N}}(E_0, E_1, E_2)]_{\mathcal{M}}(\mu, \rho)(S)
             \triangleq \mathbb{1}_{\lceil \llbracket E_2 \rrbracket(\mu) \in \mathbb{R}^{+*} \rceil} \cdot \mathbb{1}_{\lceil ((\mu,\rho),\mathsf{pdf}_{\mathcal{N}}(\llbracket E_0 \rrbracket(\mu); \llbracket E_1 \rrbracket(\mu), \llbracket E_2 \rrbracket(\mu))) \in S \rceil}
```

### A basic example

We consider the program:  $C \triangleq \begin{cases} x := \text{sample}(a, 0, 5); \\ \text{observe}(x, 3, 1); \end{cases}$ 

- prior: x close to 0, low confidence
- posterior: noisy observation that x is close to 3

**Measure semantics**, starting from  $\mu_I = \{x \mapsto ?\}$  and  $\rho_I = \emptyset$ ,

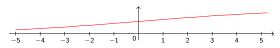
(i.e. other ountut state density pairs do not count)

(i.e., other ouptut state, density pairs do not count)

**Cumulated measure**, *i.e.*, over  $\{(\{x \mapsto v\}, \{a \mapsto v\}, pdf_{\mathcal{N}}(3; v, 1)) \mid v \leq \alpha\}$ 

$$\int_{-\infty}^{\alpha} \llbracket C \rrbracket_{\mathcal{M}}(\mu_{I}, \rho_{I}) (\{(\{x \mapsto v\}, \{a \mapsto v\}, \mathsf{pdf}_{\mathcal{N}}(3; v, 1)) \mid v \in \mathbb{R}\}) \mathrm{d}v$$

 $[C]_{\mathcal{M}}(\mu_{I}, \rho_{I})(S) = [C]_{\mathcal{M}}(\mu_{I}, \rho_{I})(S \cap \{(\{x \mapsto v\}, \{a \mapsto v\}, pdf_{\mathcal{M}}(3; v, 1)) \mid v \in \mathbb{R}\})$ 



Issue: per execution weight and probability over executions remain separate

### Outline

- Introduction
- Basic Intuition Underlying SV
- Semantics for SVI
  - Measure semantics
  - Density semantics
- 4 SVI
- 5 Ensuring correctness of SVI
- 6 Conclusion

### Towards a density semantics

For the **definition of SVI**, the measure semantics has several **limitations**:

- the weight of executions (observe commands) and the measure over output configurations are computed separately
- as mentioned earlier, we look for a way to compute the probability density of each execution together with its output

#### Important notes:

- as we seek for a semantics that maps an initial state to an output configuration, we need to assume random samples fixed beforehand and are consumed during the execution
  - i.e., the  $\rho$  input collects values to be sampled, each sample pops a value note this is the converse of the measure semantics convention
- we also expect this semantics to be **deterministic**, i.e., for any given initial state, the execution of a command should produce a single output configuration
- executions may not terminate or may crash so a special output configuration is needed i.e.,  $\perp$  denotes executions that either fail or do not terminate

### Density semantics basic definition

We recall the form of the measure semantics:

$$\llbracket C 
rbracket_{\mathcal{M}} : \mathbb{S} o (\mathcal{M}(\mathbb{S} imes \mathbb{R}^+) o_{\mathcal{M}} [0,1])$$

### Signature of the density semantics

$$\llbracket C \rrbracket_{\mathcal{D}} : \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

When  $[\![C]\!]_{\mathcal{D}}(\mu, \rho) = (\mu', \rho', w', p')$ :

- $\bullet$   $\mu'$  denotes the new memory
- $\bullet$   $\rho'$  denotes the remaining part of the random dictionary
- w' denotes the execution weight (scoring)
- p' denotes its probability density

Definition of the semantics: exercise!

## Density semantics of the skip command

### Signature of the density semantics

$$[\![C]\!]_{\mathcal{D}}:\mathbb{S}\to_{\mathcal{M}}\mathbb{S}\times\mathbb{R}^+\times\mathbb{R}^+\uplus\{\bot\}$$

Command C:

skip

Measure semantics:

$$[\![\operatorname{skip}]\!]_{\mathcal{M}}(\mu,\rho)(S) = \mathbb{1}_{[((\mu,\rho),1)\in S]}$$

### Exercise

**1** Propose a definition for  $[\![.]\!]_{\mathcal{D}}$ 

## Density semantics of assignment commands

### Signature of the density semantics

$$[\![C]\!]_{\mathcal{D}}: \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

Command C:

$$x := E$$

Measure semantics:

$$[x := E]_{\mathcal{M}}(\mu, \rho)(S) = \mathbb{1}_{[((\mu[x \mapsto [E](\mu)], \rho), 1) \in S]}$$

- **9** Propose a definition for  $[\![.]\!]_{\mathcal{D}}$  under the assumption that expressions never crash
- 2 What would happen if we assume expressions may crash

## Density semantics of sequences

### Signature of the density semantics

$$\llbracket C \rrbracket_{\mathcal{D}} : \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

Command C:

$$C_0; C_1$$

Measure semantics:

$$[\![ C_0; C_1 ]\!]_{\mathcal{M}}(\mu, \rho)(S) = \int [\![ C_0 ]\!]_{\mathcal{M}}(\mu, \rho)(\mathrm{d}(\sigma_0, w_0)) \int [\![ C_1 ]\!]_{\mathcal{M}}(\sigma_0)(\mathrm{d}(\sigma_1, w_1)) \mathbb{1}_{[(\sigma_1, w_0 w_1) \in S]}$$

- **1** Propose a definition for  $[\![.]\!]_{\mathcal{D}}$
- ② What happens with composition ? Try to imagine a "lift" operator that makes  $[\![.]\!]_{\mathcal{D}}$  easier to compose and update the definition for  $[\![.]\!]_{\mathcal{D}}$  accordingly

### Density semantics of condition tests

### Signature of the density semantics

$$\llbracket C \rrbracket_{\mathcal{D}} : \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

Command C:

$$\mathrm{if}\; B\left\{ \mathit{C}_{0}\right\} \mathrm{else}\left\{ \mathit{C}_{1}\right\} \\$$

Measure semantics:

$$\begin{split} & [ \text{if } B \left\{ C_0 \right\} \text{else} \left\{ C_1 \right\} ] \!]_{\mathcal{M}}(\mu, \rho)(S) = \\ & \mathbb{1}_{[\llbracket B \rrbracket(\mu) = \mathsf{true}]} \cdot \llbracket C_0 \rrbracket_{\mathcal{M}}(\mu, \rho)(S) + \mathbb{1}_{[\llbracket B \rrbracket(\mu) = \mathsf{false}]} \cdot \llbracket C_1 \rrbracket_{\mathcal{M}}(\mu, \rho)(S) \end{aligned}$$

#### Exercise

**1** Propose a definition for  $[\![.]\!]_{\mathcal{D}}$ 

## Density semantics of loop commands

### Signature of the density semantics

$$[\![C]\!]_{\mathcal{D}}: \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

Command C:

while 
$$B\{C\}$$

Measure semantics:

$$\begin{split} & [\![ \text{while } B \ \{C\}]\!]_{\mathcal{M}}(\mu,\rho)(S) = (\mathsf{Fix}F)(\mu,\rho)(S) \\ & \text{where} \\ & F(\phi)(\mu,\rho)(S) = \mathbbm{1}_{[\![ B ]\!](\mu) = \mathsf{false}]} \cdot \mathbbm{1}_{[((\mu,\rho),1) \in S]} \\ & + \mathbbm{1}_{[\![ B ]\!](\mu) = \mathsf{true}]} \cdot \int \!\![ C \!\!]_{\mathcal{M}}(\mu,\rho) (\mathrm{d}(\sigma_0,w_0)) \int \phi(\sigma_0) (\mathrm{d}(\sigma_1,w_1)) \mathbbm{1}_{[(\sigma_1,w_0w_1) \in S]} \\ \end{aligned}$$

- $\textbf{9} \ \, \textbf{Propose a definition for} \ \, \llbracket.\rrbracket_{\mathcal{D}}$
- Explain the assumptions required for the definition

## Density semantics of sample commands

### Signature of the density semantics

$$\llbracket C \rrbracket_{\mathcal{D}} : \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

Command C:

$$x:=\operatorname{sample}_{\mathcal{N}}(k,E_1,E_2)$$

Measure semantics:

$$\begin{split} [\![x := \mathrm{sample}_{\mathcal{N}}(k, E_1, E_2)]\!]_{\mathcal{M}}(\mu, \rho)(S) &= \\ \mathbb{1}_{[k \not\in \mathrm{Dom}(\rho)]} \cdot \mathbb{1}_{[\![E_2]\!](\mu) \in \mathbb{R}^{+*}]} \\ &\cdot \int \mathrm{d}v \left( \mathsf{pdf}_{\mathcal{N}}(v; [\![E_1]\!](\mu), [\![E_2]\!](\mu)) \cdot \mathbb{1}_{[\![(\mu[x \mapsto v], \rho[k \mapsto v], 1) \in S]} \right) \end{aligned}$$

- $\textbf{ 9 Propose a definition for } \llbracket.\rrbracket_{\mathcal{D}}$
- Comment on error cases

## Density semantics of observe commands

### Signature of the density semantics

$$\llbracket C \rrbracket_{\mathcal{D}} : \mathbb{S} \to_{\mathcal{M}} \mathbb{S} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$$

Command C:

$$\operatorname{observe}_{\mathcal{N}}(E_0,E_1,E_2)$$

Measure semantics:

- **1** Propose a definition for  $[\![.]\!]_{\mathcal{D}}$
- Comment on error cases

### Density semantics

Lifting of semantic function  $g: \mathbb{M} \times \mathbb{P} \to_{\mathcal{M}} \mathbb{M} \times \mathbb{P} \times \mathbb{R}^+ \times \mathbb{R}^+ \uplus \{\bot\}$ , for composition:

$$\begin{array}{rcl} g^{\bullet}(\bot) & = & \bot \\ g^{\bullet}(\mu,\rho,w,p) & = & \left\{ \begin{array}{ll} \bot & \text{if } g(\mu,\rho) = \bot \\ (\mu',\rho',w\cdot w',p\cdot p') & \text{if} g(\mu,\rho) = (\mu',\rho',w',p') \end{array} \right. \end{array}$$

Semantics:

## Density of an execution and example

### Definition: execution probability density

When  $[\![C]\!]_{\mathcal{D}}(\mu_I, \rho_I) = (\mu, \emptyset, w, p)$ , we let:

$$\mathcal{D}[\![C]\!](\mu_I,\rho_I) = \mathbf{w} \cdot \mathbf{p}$$

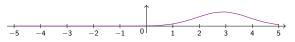
We consider again the program:  $C \triangleq \begin{cases} x := \text{sample}(a, 0, 5); \\ \text{observe}(x, 3, 1); \end{cases}$ 

- prior: x close to 0, low confidence
- posterior: noisy observation that x is close to 3

**Semantics** derived by simple calculation, starting from  $\mu_I = \{x \mapsto ?\}$  and  $\rho_I(v) = \{a \mapsto v\}$ ,

$$\llbracket C \rrbracket_{\mathcal{D}}((\mu_{I},\rho_{I}),1,1) = ((\lbrace x \mapsto v \rbrace,\emptyset),\mathsf{pdf}_{\mathcal{N}}(3;v,1),\mathsf{pdf}_{\mathcal{N}}(v;0,5))$$

Overall weighted density:  $v \mapsto \mathcal{D}[\![C]\!](\mu_i, \rho_i(v)) = \mathsf{pdf}_{\mathcal{N}}(3; v, 1) \cdot \mathsf{pdf}_{\mathcal{N}}(v; 0, 5)$ 



## Density of an execution and example

## Definition: execution probability density

When  $[C]_{\mathcal{D}}(\mu_I, \rho_I) = (\mu, \emptyset, w, p)$ , we let:

$$\mathcal{D}[\![C]\!](\mu_I,\rho_I) = \mathbf{w} \cdot \mathbf{p}$$

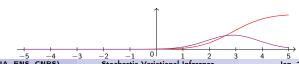
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$$\llbracket C \rrbracket_{\mathcal{D}}((\mu_I, \rho_I), 1, 1) = ((\lbrace x \mapsto v \rbrace, \emptyset), \mathsf{pdf}_{\mathcal{N}}(3; v, 1), \mathsf{pdf}_{\mathcal{N}}(v; 0, 5))$$

**Cumulated weighted density**:  $v \mapsto \int_{-\infty}^{v} \mathcal{D}[\![C]\!](\mu_i, \rho_i(x)) dx$ 



**Definition of density**: when 
$$[\![C]\!]_{\mathcal{D}}(\mu_I, \rho_I) = (\mu, \emptyset, w, p)$$
, we have  $\mathcal{D}[\![C]\!](\mu_I, \rho_I) = w \cdot p$ 

#### Theorem

Given a set of random dictonaries  $P \subset \mathbb{P}$ , we can measure the probability to evaluate exactly a  $\rho \in P$  with either semantics, in an equivalent manner:

$$\mathcal{M}(C, P) \triangleq \int (\mathbb{1}_{[\rho \in P]} \cdot |\mathrm{d}\rho| \cdot \mathcal{D}[C](\mu_I, \rho))$$
$$= \int [C]_{\mathcal{M}}(\mu_I, \emptyset) (\mathrm{d}(\mu, \rho, w)) \cdot (w \cdot \mathbb{1}_{[\rho \in P]})$$

Proof: exercise!

## Outline

- Introduction
- Basic Intuition Underlying SV
- Semantics for SVI
- 4 SVI
- 5 Ensuring correctness of SVI
- Conclusion

## KL divergence

#### Status so far:

- given model P and parameterized guide  $Q_{\theta}$ , we have defined the distributions p and  $q_{\theta}$  that they induce over random data-bases
- we next need to establish a **measure of dissimilarity** between p and  $q_{\theta}$ , to be able to define the optimization objective of SVI

## Definition: KL divergence (Kullback-Leibler divergence)

Given two probability distributions  $p_0, p_1$  over the same measurable set  $\mathcal{X}$ , their **KL divergence** writes down as:

$$\mathbf{D}_{\mathrm{KL}}(p_0||p_1) = \mathbb{E}_{p_0}\left(\log\frac{p_0}{p_1}\right) = \int_{\mathcal{X}} p_0(x)\log\frac{p_0(x)}{p_1(x)}\mathrm{d}x$$

In the discrete case:

$$\mathbf{D}_{\mathrm{KL}}(p_0||p_1) = \sum_{\mathbf{x} \in \mathcal{X}} p_0(\mathbf{x}) \log \frac{p_0(\mathbf{x})}{p_1(\mathbf{x})}$$

## Some properties of KL divergence

**Main properties**, for a given measurable space with measure |.|:

- ① it is **positive**: for all p, q,  $\mathbf{D}_{\mathrm{KL}}(p||q) \geq 0$
- **2** it is null if and only if its arguments are equal almost everywhere:

$$\mathbf{D}_{\mathrm{KL}}(p||q) = 0 \iff |x| \neq 0 \Longrightarrow p(x) = q(x)$$

- **3** it is **not symmetric**, which means that in general  $\mathbf{D}_{\mathrm{KL}}(p||q) \neq \mathbf{D}_{\mathrm{KL}}(q||p)$
- **3** it does not satisfy the triangular inequality, which means that, in general, the inequality  $\mathbf{D}_{\mathrm{KL}}(p_0||p_2) \leq \mathbf{D}_{\mathrm{KL}}(p_0||p_1) + \mathbf{D}_{\mathrm{KL}}(p_1||p_2)$  does not hold

Due to the last two points it is called *divergence*, and not **distance** 

#### **Examples**, over $X = \{0, 1\}$ :

- $p_0$  defined by  $p_0(0) = \frac{1}{2}$  and  $p_0(1) = \frac{1}{2}$
- $p_1$  defined by  $p_1(0) = \frac{1}{5}$  and  $p_1(1) = \frac{4}{5}$
- $p_2$  defined by  $p_2(0) = \frac{1}{10}$  and  $p_2(1) = \frac{9}{10}$

#### Then:

- $\mathbf{D}_{\mathrm{KL}}(p_0||p_1) \approx 0.22 \text{ but } \mathbf{D}_{\mathrm{KL}}(p_1||p_0) \approx 0.19$
- $\mathbf{D}_{\mathrm{KL}}(p_0||p_2) \approx 0.51$  but  $\mathbf{D}_{\mathrm{KL}}(p_0||p_1) + \mathbf{D}_{\mathrm{KL}}(p_1||p_2) \approx 0.27$

## SVI objective

#### Setup:

- model P; defines a probability distribution over sampled and observed random variables noted p(z, x) where:
  - z: sampled random variables (represented in  $\rho$ )
  - x: observed random variables (though not represented in  $\rho$
- guide Q, with a set of variables identified as parameters and noted t; given a value assignment  $t \mapsto \theta$ , defines a probability distribution  $q_{\theta}(z)$  over sampled random variables

We defined:  $\mathcal{M}(p, S) = \int \rho(\mathrm{d}\rho) \left(\mathbb{1}_{[\rho \in S]} \cdot \mathcal{D}[\![p]\!](\mu_I, \rho)\right)$ 

It can be normalized into a probability measure iff  $\mathcal{M}(p,\mathbb{P}) \in \mathbb{R}^{+*}$ 

#### Objective:

- ullet beforehand, fix the family of programs  $q_{ heta}$  as potential approximants of C
  - **9** with  $\mathcal{M}(q_{\theta}, \mathbb{P}) = 1$ , which is ensured if  $q_{\theta}$  always terminates
  - 2 with density 1, which is ensured if it contains no occurrence of observe
- ullet compute optimal heta to minimize

$$\mathbf{D}_{\mathrm{KL}}(\mathcal{D}[\![q_{\theta}]\!](\mu_{I},\rho),\mathcal{D}[\![C]\!](\mu_{I},\rho))$$

## SVI optimization objective

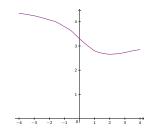
### Inference goal

Compute an ideal value of  $\theta$  that makes  $q_{\theta}$  as close to p as possible, using an optimization algorithm

# **Application** to the **inference problem** two distributions over sampled variables *z*:

- p(z|x):
   posterior probability distribution over
   z defined by the model
   (with the observation x)
- $q_{\theta}(z)$ : guide probability distribution (parameterized by  $\theta$ )

**Plot of D**<sub>**KL**</sub> $(q_{\theta}(z), p(z|x))$  as a function of  $\theta$ :



**Optimization objective:**  $\operatorname{argmin}_{\theta} \mathbf{D_{KL}}(q_{\theta}(z), p(z|x))$ 

**Next step**: achieve a stochastic approximation of the gradient of  $D_{\rm KL}$ 

## Definition of the optimization objective

We seek for  $\theta$  so as to minimize:

$$\begin{aligned} &\mathbf{D}_{\mathrm{KL}}(q_{\theta}(z), p(z|x)) \\ &= \int_{\mathcal{Z}} q_{\theta}(z) \log \frac{q_{\theta}(z)}{p(z|x)} \mathrm{d}z \\ &= \int_{\mathcal{Z}} q_{\theta}(z) \log \frac{q_{\theta}(z)p(x)}{p(z,x)} \mathrm{d}z \qquad \qquad \text{by independence of } x, z \\ &= \int_{\mathcal{Z}} q_{\theta}(z) \log p(x) \mathrm{d}z + \int_{\mathcal{Z}} q_{\theta}(z) \log \frac{q_{\theta}(z)}{p(z,x)} \mathrm{d}z \\ &= \log p(x) \cdot \int_{\mathcal{Z}} q_{\theta}(z) \mathrm{d}z + \int_{\mathcal{Z}} q_{\theta}(z) \log \frac{q_{\theta}(z)}{p(z,x)} \mathrm{d}z \\ &= \log p(x) \cdot 1 + \int_{\mathcal{Z}} q_{\theta}(z) \log \frac{q_{\theta}(z)}{p(z,x)} \mathrm{d}z \qquad \qquad \text{as } q_{\theta} \text{ is a probability} \\ &= \log p(x) - \mathcal{L}(\theta) \end{aligned}$$

where  $\mathcal{L}(\theta)$  is called the ELBO:

## Evidence Lower Bound (ELBO)

The **ELBO** is defined as  $\mathcal{L}(\theta) = \int_{\mathcal{Z}} q_{\theta}(z) \log \frac{p(z,x)}{q_{\theta}(z)} dz$ .

Since  $\mathbf{D}_{\mathrm{KL}}(q_{\theta}(z), p(z|x)) + \mathcal{L}(\theta) = \log p(x)$  and  $\log p(x)$  does not depend in  $\theta$ ,

Minimizing  $D_{KL}(q_{\theta}(z), p(z|x))$  is equivalent to maximizing  $\mathcal{L}(\theta)$ 

## Gradient of the ELBO... (1)

To perform **gradient ascent** in order to **maximize ELBO**, we first need to evaluate this gradient, and formulate it **as an expectation**:

$$\begin{split} & \nabla_{\theta} \mathcal{L}(\theta) \\ & = \int_{\mathcal{Z}} \nabla_{\theta} \left[ q_{\theta}(z) \log \frac{p(z,x)}{q_{\theta}(z)} \mathrm{d}z \right] \qquad \text{under assumption (later)} \\ & = \int_{\mathcal{Z}} \left( \nabla_{\theta} q_{\theta}(z) \right) \log \frac{p(z,x)}{q_{\theta}(z)} \mathrm{d}z + \int_{\mathcal{Z}} q_{\theta}(z) \nabla_{\theta} \left( \log \frac{p(z,x)}{q_{\theta}(z)} \right) \mathrm{d}z \end{split}$$

Let us study the second term more:

$$\begin{split} &\int_{\mathcal{Z}} q_{\theta}(z) \nabla_{\theta} \left( \log \frac{\rho(z,x)}{q_{\theta}(z)} \right) \mathrm{d}z \\ &= \int_{\mathcal{Z}} q_{\theta}(z) \nabla_{\theta} \left( \log p(z,x) - \log q_{\theta}(z) \right) \mathrm{d}z \\ &= \int_{\mathcal{Z}} q_{\theta}(z) \nabla_{\theta} \left( \log p(z,x) \right) \mathrm{d}z - \int_{\mathcal{Z}} q_{\theta}(z) \nabla_{\theta} \left( \log q_{\theta}(z) \right) \mathrm{d}z \\ &= -\int_{\mathcal{Z}} q_{\theta}(z) \nabla_{\theta} \left( \log q_{\theta}(z) \right) \mathrm{d}z \qquad \text{(the gradient of a constant is null)} \\ &= -\int_{\mathcal{Z}} q_{\theta}(z) \frac{\nabla_{\theta} q_{\theta}(z)}{q_{\theta}(z)} \mathrm{d}z \\ &= -\int_{\mathcal{Z}} \nabla_{\theta} q_{\theta}(z) \mathrm{d}z \\ &= -\nabla_{\theta} [\int_{\mathcal{Z}} q_{\theta}(z) \mathrm{d}z] \\ &= -\nabla_{\theta} 1 \qquad \text{(since } q_{\theta} \text{ defines a probability distr)} \end{split}$$

## Gradient of the ELBO... (2)

We have shown:

$$abla_{ heta}\mathcal{L}( heta) = \int_{\mathcal{Z}} \left( 
abla_{ heta} q_{ heta}(z) \right) \log rac{p(z,x)}{q_{ heta}(z)} \mathrm{d}z$$

We remark that  $\nabla_{\theta} \log q_{\theta}(z) = \frac{\nabla_{\theta} q_{\theta}(z)}{q_{\theta}(z)}$ 

Thus, we can substitute  $\nabla_{\theta} q_{\theta}(z) = q_{\theta}(z) \cdot \nabla_{\theta} \log q_{\theta}(z)$ :

$$\begin{array}{rcl} \nabla_{\theta} \mathcal{L}(\theta) & = & \int_{\mathcal{Z}} \left( \nabla_{\theta} q_{\theta}(z) \right) \log \frac{p(z,x)}{q_{\theta}(z)} \mathrm{d}z \\ & = & \int_{\mathcal{Z}} q_{\theta}(z) \cdot \nabla_{\theta} \log q_{\theta}(z) \log \frac{p(z,x)}{q_{\theta}(z)} \mathrm{d}z \\ & = & \mathbb{E}_{q_{\theta}(z)} \left( \nabla_{\theta} \log q_{\theta}(z) \log \frac{p(z,x)}{q_{\theta}(z)} \right) \end{array}$$

This form, **as an expectation** is adapted to **stochastic approximation** *i.e.*, estimation based on a number of samples...

## Stochastic approximation of the gradient of the ELBO

Based on  $\nabla_{\theta} \mathcal{L}(\theta) = \mathbb{E}_{q_{\theta}(z)} \left( \nabla_{\theta} \log q_{\theta}(z) \log \frac{p(z,x)}{q_{\theta}(z)} \right)$  we may generate a sample  $\rho_0$  from  $q_{\theta}$  and produce

$$\mathsf{GrEst}_{\theta}(\rho_0) = (\nabla_{\theta} \log \mathcal{D}[\![q_{\theta}]\!](\mu_I, \rho_0)) \cdot \log \frac{\mathcal{D}[\![q_{\theta}]\!](\mu_I, \rho)}{\mathcal{D}[\![p]\!](\mu_I, \rho_0)}$$

Stochastic estimation repeats with many samples:

## Stochastic approximant of gradient expectation formula

Using N samples from  $\rho_0, \ldots, \rho_{N-1}$  from distribution  $q_\theta$ :

$$\underline{\mathsf{GrEst}}_{\theta}(\rho) = \tfrac{1}{N} \sum_{i=0}^{N-1} \mathsf{GrEst}_{\theta}(\rho_i) = \tfrac{1}{N} \sum_{i=0}^{N-1} \left( \nabla_{\theta} \log \mathcal{D}[\![q_{\theta}]\!] (\mu_I, \rho_i) \right) \cdot \log \tfrac{\mathcal{D}[\![q_{\theta}]\!] (\mu_I, \rho_i)}{\mathcal{D}[\![p]\!] (\mu_I, \rho_i)}$$

#### Implementation:

- sampling executions from  $q_{\theta}$ , e.g., by rejection sampling
- computation of  $\mathcal{D}[\![q_{\theta}]\!]$  and  $\mathcal{D}[\![p]\!]$  using **density semantics** based **instrumented implementation**; similarly, also accumulate  $\nabla_{\theta}(\log pdf)$  (table for classical distributions)
- usually sum log pdf rather than multiplying pdf

## SVI algorithm based on gradient optimization

#### **Fixed parameters:**

- N number of samples per iterate for stochastic estimation
- $\lambda$ : learning rate, typically small, e.g.,  $\lambda = 0.01$
- $\theta_{\text{init}}$ : initial value of the guide parameter (typically a rough guess)

#### Algorithm:

$$\begin{cases} \text{ select } \theta_0 := \theta_{\text{init}} \\ \text{ repeat } \mathcal{K} \text{ times} \\ \text{ sample } r_0, \dots, r_{N-1} \\ \theta_{k+1} \leftarrow \theta_k - \lambda \cdot \frac{1}{N} \cdot \sum_{i=0}^{N-1} \text{GrEst}_{\theta_k}(r_i) \\ \text{produce } \theta_{\mathcal{K}} \end{cases}$$

## Properties (under assumptions discussed shortly)

- $\frac{1}{N} \cdot \sum_{i=0}^{N-1} \mathbf{GrEst}_{\theta_k}(\rho_i)$  provides an **unbiased** estimate of the gradient at  $\theta_k$
- the algorithm converges to a local maximum  $\theta$  of  $\mathcal{L}$ , i.e., a local minimum of  $\mathbf{D}_{\mathrm{KL},\ldots}$

### Outline

- Ensuring correctness of SVI
  - Behavior of SVI when assumptions do not hold
  - Discharging SVI assumptions using static analysis

## Another model-guide pair

#### Model:

```
def model (...):
  sigma = pyro.sample("sigma", Uniform(0., 10.))
  pyro.sample("obs", Normal(..., sigma), obs = ...)
```

## Guide:

```
def guide (...):
  loc = pyro.param("sigma_loc", 1., constraint=constraints.positive)
 sigma = pyro.sample("sigma", Normal(loc, 0.05))
```

#### Issue:

- domain of sigma in the model: [0, 10]
- ullet domain of sigma in the guide:  $\mathbb{R}$
- thus, KL-divergence is undefined

(Example taken from the Pyro webpage examples...)

## Issues possibly leading to undefinedness of KL-divergence

#### Absolute continuity requirement:

definition of KL-divergence:

$$\mathbf{D}_{\mathrm{KL}}(q_{ heta},p) = \mathbb{E}_{q_{ heta}}\left(\log rac{q_{ heta}}{p}
ight) = \int_{\mathcal{X}} q_{ heta}(\mathrm{d}x)\log rac{q_{ heta}(\mathrm{d}x)}{p(\mathrm{d}x)}$$

- absolute continuity requirement: model distribution p and guide distribution g should have the same zero probability regions otherwise: KL divergence is undefined
- domain in model [0, 10], in guide  $\mathbb{R}$ leads to the violation of absolute continuity assumption e.g., and KL divergence is undefined

#### Anther possible issue: integrability

- $q_{\theta}(\mathrm{d}x)\log\frac{q_{\theta}(\mathrm{d}x)}{p(\mathrm{d}x)}$  may not be integrable ... even when absolute continuity holds
- Our goal: define semantics to let static analysis provide guarantees

## Informal overview of potential SVI issues

### Several assumptions are necessary:

- KL-divergence must be defined, not  $\infty$ : otherwise: undefined optimization objective
- KL-divergence must be differentiable: otherwise: incorrect gradient descent
- the stochastic estimate of  $\nabla D_{KL}(q_{\theta}, p)$  should be well-defined, and unhiased:
  - otherwise: incorrect computation of gradient descent approximation

#### Practical consequences are difficult to troubleshoot, e.g.,

- crashes or divergence of the inference engine
- incoherent / invalid optimization results may be very difficult to even notice

## Unbiasedness conditions

Sufficient conditions (unbiasedness may hold in some cases where assumptions are violated, especially if locally so):

### Theorem: unbiasedness of gradient estimate of KL divergence

If:

- **absolute continuity:**  $\mathcal{D}[\![D_{\theta}]\!](\mu_I)(\rho) \Longrightarrow \mathcal{D}[\![C]\!](\mu_I)(\rho)$
- $\bullet$  differentiability:  $\theta \mapsto \mathcal{D}[\![D_{\theta}]\!](\mu_I)(\rho)$  differentiable wrt all components
- boundnedness of KL divergence
- differentiability of KL divergence wrt all its arguments
- integral permutation conditions on KL divergence and guide density  $\int \nabla \ldots = \nabla \int \ldots$

Then:

$$\mathbb{E}(\nabla_{\theta} \mathbf{D}_{\mathrm{KL}}(\mathcal{D}[\![D_{\theta}]\!](\mu_{I}), \mathcal{D}[\![C]\!](\mu_{I}))) \equiv \frac{1}{N} \cdot \sum_{i=0}^{N-1} \mathbf{GrEst}_{\theta_{k}}(\rho_{i})$$

#### Full version:

Towards Verified Stochastic Variational Inference for Probabilistic Programs Wonyeol Lee, Hangyeol Yu, Xavier Rival and Hongseok Yang POPL'20

### Outline

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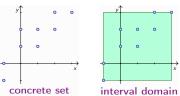
## Abstract interpretation primer: state abstraction

Static analysis: interpretation using an abstract domain

## Abstract domains (Cousot & Cousot, 1977)

- Families of abstract predicates adapted to static analysis
- Compact and efficient representations
- Operations for the static analysis of concrete operations
- $\bullet$  Mapping from abstract to concrete: concretization function  $\gamma$

### Ex., numeric abstractions: abstraction of sets of pairs of integers







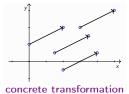
Abstract states **over-approximate** sets of concrete states

## Abstract interpretation primer: analysis of post-conditions

## Computing sound abstract transformer

- Conservative analysis of concrete execution steps in the abstract e.g., assignments, condition tests...
- May lose precision, will never forget any behavior
- Balance between cost and precision

### **Example**: analysis of a translation with octagons (exact!)



abstract transformation

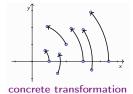
Soundness: all concrete behaviors are taken into account

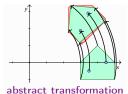
## Abstract interpretation primer: analysis of post-conditions

## Computing sound abstract transformer

- Conservative analysis of concrete execution steps in the abstract e.g., assignments, condition tests...
- May lose precision, will never forget any behavior
- Balance between cost and precision

### **Example**: analysis of a 40 deg. rotation with octagons (approximate!)





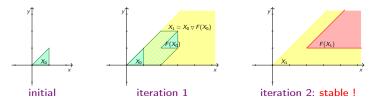
Soundness: all concrete behaviors are taken into account

## Abstract interpretation primer: analysis of loops

## Computing invariants about infinite executions with widening $\nabla$

- Widening ∇ over-approximates U: soundness guarantee
- Widening ∇ guarantees the termination of the analyses
- Typical choice of ∇: remove unstable constraints

#### **Example**: iteration of the translation (2,1), with octagons



Soundness: all concrete behaviors are taken into account

## A generic static analysis

We set up a static analysis, parameterized by an abstract domain: Logical predicates + representation + algorithms

#### Abstract domain

An abstract domain comprises a set of abstract predicates  $\mathbb{D}^{\sharp}$  and:

- $\bullet \ \ \text{concretization function} \ \gamma: \mathbb{D}^{\sharp} \to \mathbb{D} \ \ \text{where} \ \mathbb{D} = \mathbb{S} \to \mathbb{S} \times \mathbb{R}^{+} \times \mathbb{R}^{+} \uplus \{\bot\}$
- least element  $\bot$  with  $\gamma(\bot) = \emptyset$
- widening operator  $\nabla: \mathbb{D}^{\sharp} \times \mathbb{D}^{\sharp} \longrightarrow \mathbb{D}^{\sharp}$ over-approximating  $\cup$ and enforcing termination on all sequences of abstract iterates
- abstract composition  $comp^{\sharp}: \mathbb{D}^{\sharp} \times \mathbb{D}^{\sharp} \longrightarrow \mathbb{D}^{\sharp}$ soundness:  $\forall g_0 \in \gamma(d_0^{\sharp}), \forall g_1 \in \gamma(d_1^{\sharp}), \ (g_0 \circ g_1) \in \gamma(comp^{\sharp}(d_0^{\sharp}, d_1^{\sharp}))$
- abstract conditions, assignment, sample and score operations satisfying similar soundness conditions

## Abstract interpreter

The definition of the abstract interpreter follows by induction over the syntax:

### The abstract interpreter

**Exercise**: provide the soundness condition for each of the operations

### Soundness

The abstract semantics is sound in the sense that it over-approximates the effect of program concrete executions, in the sense of the density semantics:

## Theorem: static analysis soundness

For all command C:

$$\llbracket C \rrbracket_{\mathcal{D}} \in \gamma(\llbracket C \rrbracket^{\sharp})$$

#### Proof: exercise!

The abstract semantics is not complete, and may not return the most precise abstraction for a given program...

## First instance: static analysis for model/guide support match

#### Goal: discharge model/guide support equality (absolute continuity)

### Abstraction

We define  $\mathbb{D}^{\sharp}$  and  $\gamma$  by  $\mathbb{D}^{\sharp} = \{ \perp^{\sharp}, \top^{\sharp} \} \uplus \mathcal{P}(\mathbb{K})$  and:

$$\begin{array}{cccc} \gamma: & \perp^{\sharp} & \longmapsto & \lambda((\mu,\rho),w,p) \cdot \bot \\ & \top^{\sharp} & \longmapsto & \mathbb{D} \\ & \mathcal{K}(\subseteq \mathbb{K}) & \longmapsto & \{g \in \mathbb{D} \mid \\ & & [\forall ((\mu,\rho),w,p),\mu',w',p' \\ & & & g((\mu,\rho),w,p) = (\mu',\emptyset,w',p') \Longrightarrow \mathrm{Dom}(\rho) = \mathcal{K}] \} \end{array}$$

#### A few transfer functions:

$$\begin{array}{rcl} comp^{\sharp}(\bot^{\sharp},d^{\sharp}) = comp^{\sharp}(d^{\sharp},\bot^{\sharp}) & = & \bot^{\sharp} \\ comp^{\sharp}(\top^{\sharp},d^{\sharp}) = comp^{\sharp}(d^{\sharp},\top^{\sharp}) & = & \top^{\sharp} \\ & comp^{\sharp}(K_{0},K_{1}) & = & \begin{cases} K_{0} \uplus K_{1} & \text{if } K_{0} \cap K_{1} = \emptyset \\ \top^{\sharp} & \text{otherwise} \end{cases} \\ sample^{\sharp}(x,k,E_{0},E_{1}) & = & \{k\} \\ score^{\sharp}(x,E_{0},E_{1},E_{2}) & = & \emptyset \end{array}$$

## First instance: static analysis for model/guide support match

### Goal: discharge model/guide support equality (absolute continuity)

### Abstraction

```
We define \mathbb{D}^{\sharp} and \gamma by \mathbb{D}^{\sharp} = \{ \perp^{\sharp}, \top^{\sharp} \} \uplus \mathcal{P}(\mathbb{K}) and:
```

$$\gamma: \quad \bot^{\sharp} \quad \longmapsto \quad \lambda((\mu, \rho), w, p) \cdot \bot 
\qquad \top^{\sharp} \quad \longmapsto \quad \mathbb{D} 
\qquad K(\subseteq \mathbb{K}) \quad \longmapsto \quad \{g \in \mathbb{D} \mid 
\qquad \qquad [\forall ((\mu, \rho), w, p), \mu', w', p' 
\qquad \qquad g((\mu, \rho), w, p) = (\mu', \emptyset, w', p') \Longrightarrow \mathrm{Dom}(\rho) = K]\}$$

#### **Example analysis:**

Xavier Rival (INRIA, ENS. CNRS)

```
def model():
 x = pyro.sample("v", Normal(0., 5.))
  if (x > 0):
    pyro.sample("obs", Normal(1., 1.), obs=x)
  else:
    pyro.sample("obs", Normal(-2., 1.), obs=x)
```

Then:

 $[model]^{\sharp} = \{v\}$ 

## First instance: static analysis for model/guide support match

Goal: discharge model/guide support equality (absolute continuity)

### Abstraction

We define  $\mathbb{D}^{\sharp}$  and  $\gamma$  by  $\mathbb{D}^{\sharp} = \{\bot^{\sharp}, \top^{\sharp}\} \uplus \mathcal{P}(\mathbb{K})$  and:

$$\begin{array}{cccc} \gamma: & \perp^{\sharp} & \longmapsto & \lambda((\mu,\rho),w,p) \cdot \bot \\ & \top^{\sharp} & \longmapsto & \mathbb{D} \\ & \mathcal{K}(\subseteq \mathbb{K}) & \longmapsto & \{g \in \mathbb{D} \mid \\ & & [\forall ((\mu,\rho),w,p),\mu',w',p' \\ & & g((\mu,\rho),w,p) = (\mu',\emptyset,w',p') \Longrightarrow \mathrm{Dom}(\rho) = \mathcal{K}] \} \end{array}$$

Generalization: case where we consider many distributions (and not only normal distributions)

- the basic abstraction will not work: sampling k from a normal distribution and sampling k from a Bernoulli distribution yield distinct supports
- new abstraction:  $\mathbb{K} \to \mathbf{Distributions} \uplus \{\bot, \top\}$

## Second instance: static analysis for guide differentiability

#### Goal: discharge differentiability properties

#### Exercise:

- abstraction choice
- transfer functions

## Outline

- Introduction
- Basic Intuition Underlying SV
- Semantics for SVI
- 4 SV
- 5 Ensuring correctness of SVI
- 6 Conclusion

### Conclusion

#### Main ideas to remember from this lecture:

- SVI turns an inference problem into an optimization problem, to select an (possibly) optimal program in a parameterized family
- optimization only works when a series of assumptions hold

### A few suggestions for **reading**:

- on probabilistic semantics:
  - Dexter Kozen,
  - Semantics of Probabilistic Programs.
  - Journal of Computing Systems Science(1981)
- on SVI, among other inference techniques:
   Jan-Willem van de Meent, Brooks Paige, Hongseok Yang, Frank Wood,
   An Introduction to Probabilistic Programming.
   (2018)
- on correctness of SVI:
  - Wonyeol Lee, Hangyeol Yu, Xavier Rival, Hongseok Yang, Towards verified stochastic variational inference for probabilistic programs, PACM-POPL (2020)