

```

beta.mat <- matrix(nrow=3,ncol=2)

for (j in 1:3){
  beta.mat[j,] <- rnorm(2,0,5)
}
beta.init[[i]] <- beta.mat
}

#sample initial prec vals
prec.init <- list()

for (i in 1:nchain){
  prec.mat <- rep(NA, 3)

  for (j in 1:3){
    prec.mat[j] <- runif(1,1/1000,1/100)
  }

  prec.init[[i]] <- prec.mat
}

inits.bp2 <- list()
for(i in 1:nchain){
  inits.bp2[[i]] <- list(
    beta = beta.init[[i]],
    prec = prec.init[[i]], k0 = sort(sample(1:length(Xsim), 2)))
}

##initialize model sample from it
#jags object
bp2.model   <- jags.model(file = textConnection(breakpoint2_mod),
                           data = data.bp2,
                           inits = inits.bp2,
                           n.chains = nchain)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
##   Observed stochastic nodes: 57
##   Unobserved stochastic nodes: 8
##   Total graph size: 900
##
## Initializing model

#sample from model
bp2.out<- coda.samples (model = bp2.model,
                         variable.names = c("K", "beta", "prec"),
                         n.iter = 250000)

#convergence?

```

```

gelman.diag(bp2.out)

## Potential scale reduction factors:
##
##          Point est. Upper C.I.
## K[1]           1.00   1.00
## K[2]           1.08   1.08
## beta[1,1]      1.00   1.00
## beta[2,1]      1.00   1.00
## beta[3,1]      1.00   1.00
## beta[1,2]      1.00   1.00
## beta[2,2]      1.00   1.00
## beta[3,2]      1.00   1.00
## prec[1]        1.00   1.00
## prec[2]        1.00   1.00
## prec[3]        1.30   1.70
##
## Multivariate psrf
##
## 1.01

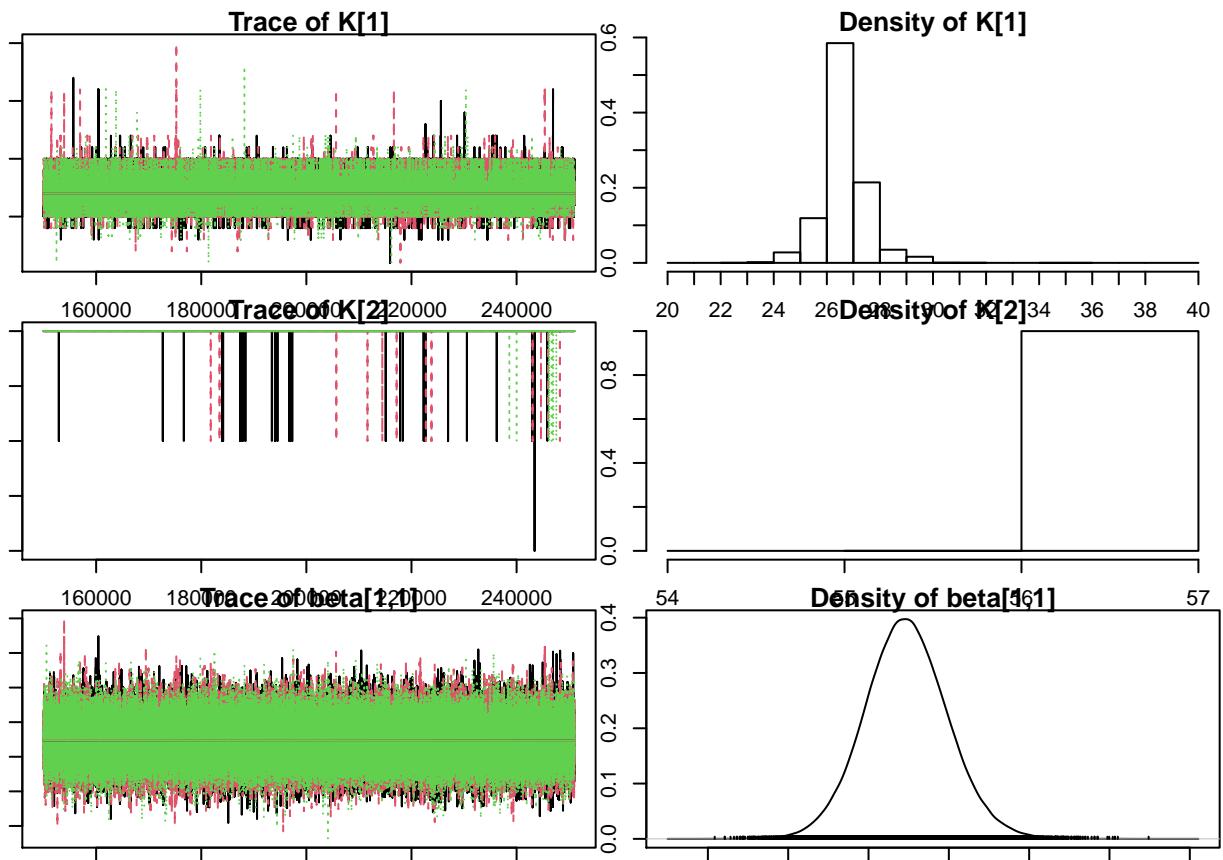
#burnin
burnin <- 150000
bp2.burn <- window(bp2.out, start=burnin)

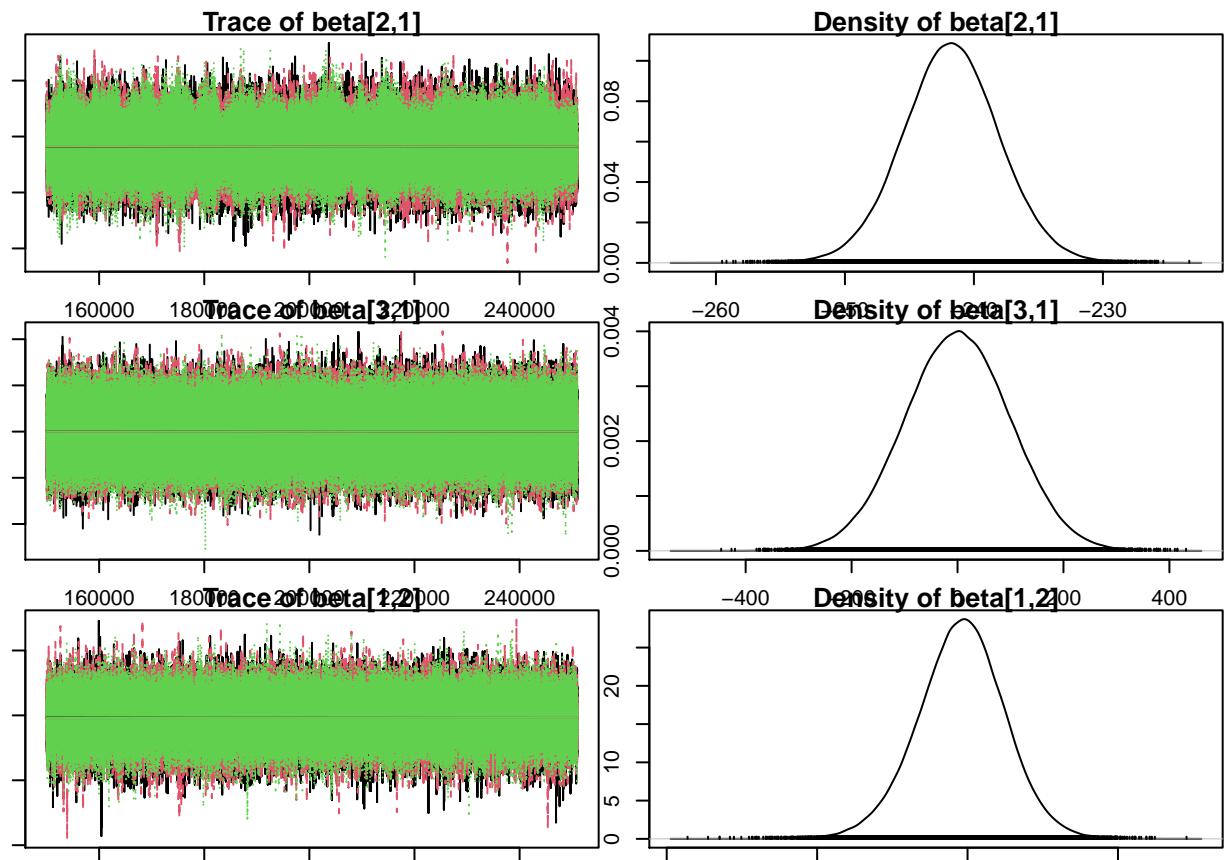
#check for convergence
par(mar=c(1,1,1,1))
gelman.diag(bp2.burn)

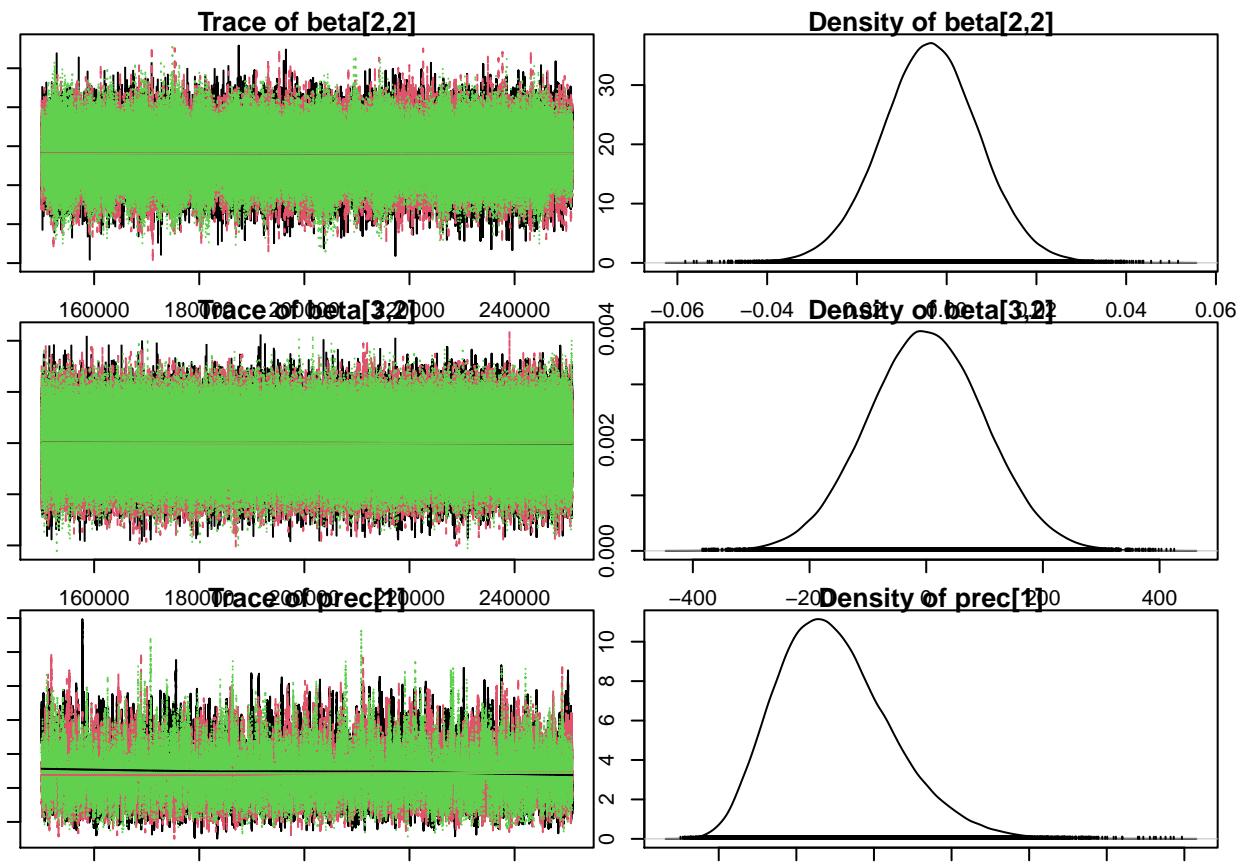
## Potential scale reduction factors:
##
##          Point est. Upper C.I.
## K[1]           1.00   1.00
## K[2]           1.10   1.10
## beta[1,1]      1.00   1.00
## beta[2,1]      1.00   1.00
## beta[3,1]      1.00   1.00
## beta[1,2]      1.00   1.00
## beta[2,2]      1.00   1.00
## beta[3,2]      1.00   1.00
## prec[1]        1.00   1.01
## prec[2]        1.01   1.02
## prec[3]        1.30   1.59
##
## Multivariate psrf
##
## 1.01

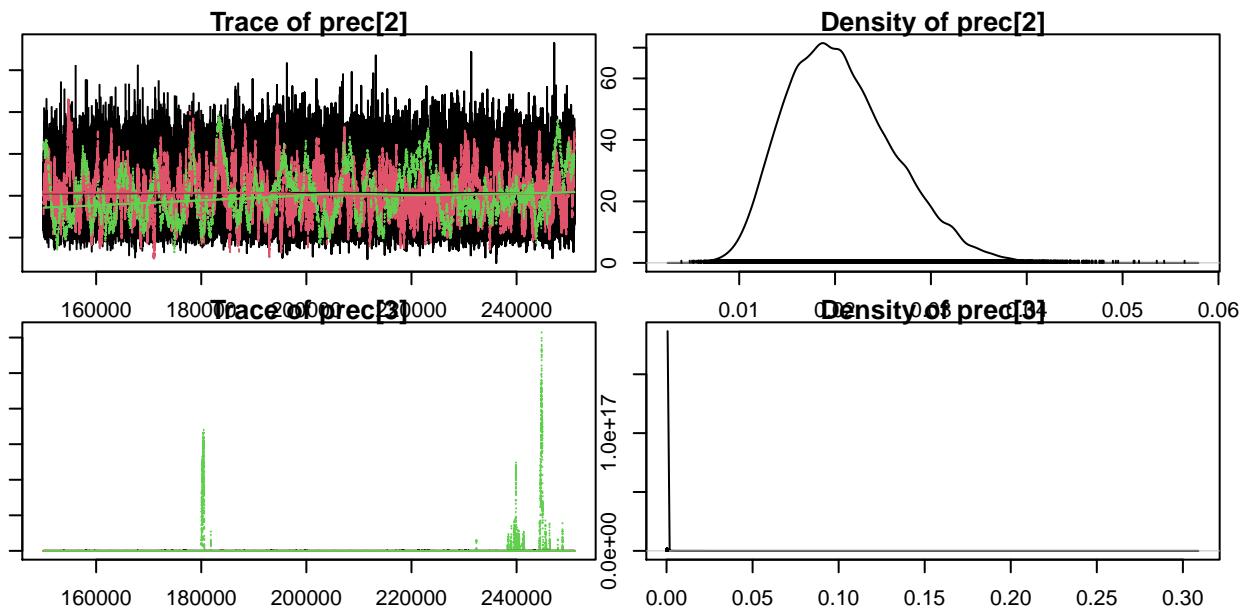
plot(bp2.burn)

```









Now lets fit our polynomial regression models starting with quadratic (2nd order)

```
#bayesian fitting of polynomial (2nd order) regression (linear with respect to parameters)
quad_regression<- "
model{

##priors
#betas
beta ~ dmnorm(b0,Vb)
#precision
prec ~ dgamma(s1,s2)

##process and data model
for(i in 1:n){
  # process model
  mu[i] <- beta[1] + beta[2]*x[i,1] + beta[3]*x[i,2]
  # data model
  Isotope[i] ~ dnorm(mu[i],prec)
}
}

##fit model
#specify data values
data.quad<- list(Isotope = dat$deuturium, x = cbind(dat$depth, dat$depth^2), n = nrow(dat))
```

```

##priors
# regression beta means
data.quad$b0 <- as.vector(c(0,0,0))
# regression beta precisions
data.quad$Vb <- solve(diag(10000,3))
# uninformative error priors
data.quad$s1 <- .1
data.quad$s2 <- .1

#initial conditions
inits.quad<- list()
for (i in 1:nchain){
  inits.quad[[i]]<- list(beta = rnorm(3,0,5), prec = runif(1,1/1000,1/100))
}

##estimate and sample from model
#JAGS object
quad.model <- jags.model(file = textConnection(quad_regression),
                           data = data.quad,
                           inits = inits.quad,
                           n.chains = nchain)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
##   Observed stochastic nodes: 57
##   Unobserved stochastic nodes: 2
##   Total graph size: 343
##
## Initializing model

#sample from model
quad.out<- coda.samples (model = quad.model,
                           variable.names = c("beta", "prec"),
                           n.iter = 5000)

#calc DIC for future model selection
quad.DIC<- dic.samples(model = quad.model, n.iter = 5000)

#burnin
burnin <- 1000
quad.burn <- window(quad.out, start = burnin)

#outputs
summary(quad.burn)

## 
## Iterations = 1000:5000
## Thinning interval = 1
## Number of chains = 3
## Sample size per chain = 4001

```

```

##  

## 1. Empirical mean and standard deviation for each variable,  

##     plus standard error of the mean:  

##  

##           Mean        SD  Naive SE Time-series SE  

## beta[1] -2.097e+02 1.793e+00 1.636e-02      1.636e-02  

## beta[2] -2.018e-01 1.845e-02 1.684e-04      1.669e-04  

## beta[3]  2.749e-04 3.526e-05 3.219e-07      3.198e-07  

## prec     2.892e-02 5.614e-03 5.124e-05      5.505e-05  

##  

## 2. Quantiles for each variable:  

##  

##       2.5%      25%      50%      75%     97.5%  

## beta[1] -2.132e+02 -2.109e+02 -2.097e+02 -2.085e+02 -2.061e+02  

## beta[2] -2.386e-01 -2.142e-01 -2.017e-01 -1.894e-01 -1.661e-01  

## beta[3]  2.067e-04  2.508e-04  2.748e-04  2.982e-04  3.442e-04  

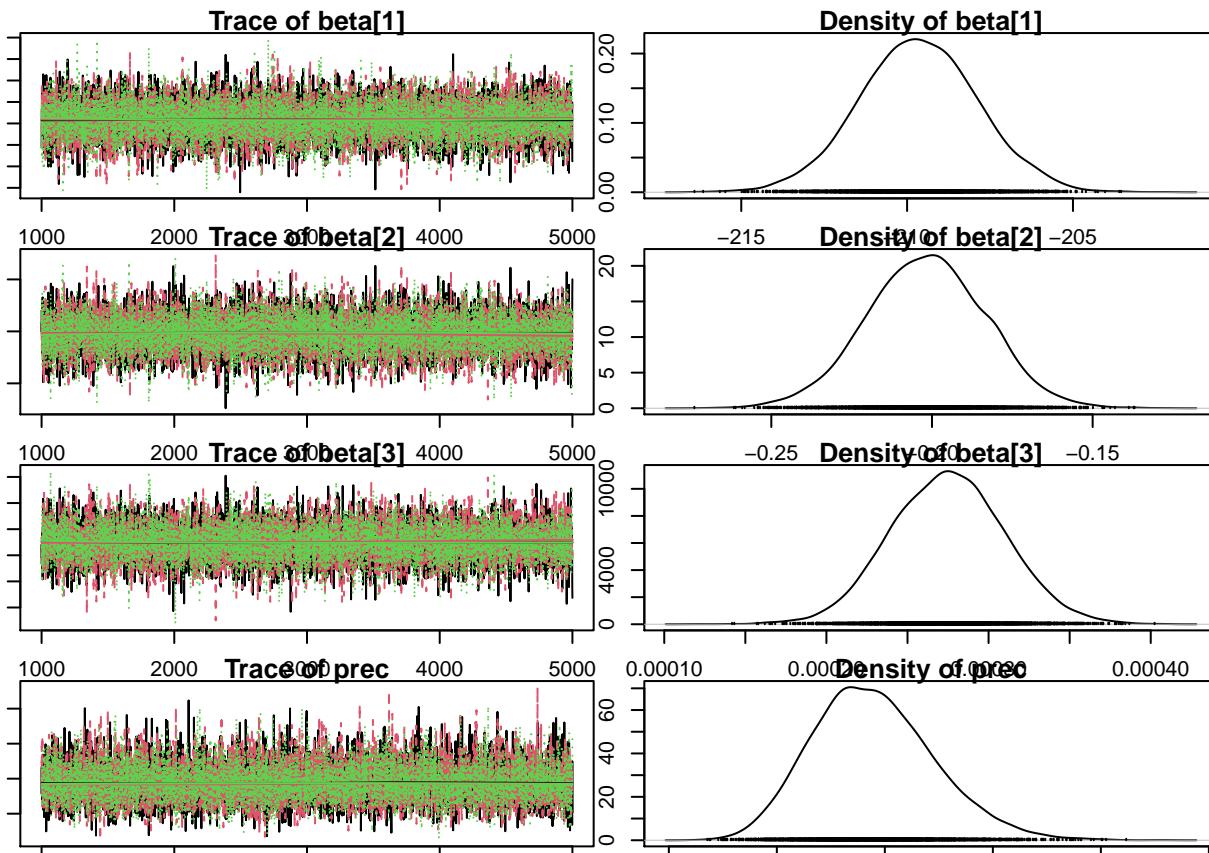
## prec     1.905e-02  2.498e-02  2.859e-02  3.247e-02  4.100e-02

```

```

par(mar=c(1,1,1,1))
plot(quad.burn)

```



```

gelman.diag(quad.burn)

```

```

## Potential scale reduction factors:  

## 

```

```

## Point est. Upper C.I.
## beta[1]      1      1.00
## beta[2]      1      1.01
## beta[3]      1      1.01
## prec         1      1.00
##
## Multivariate psrf
##
## 1

##predictive and credible intervals
#initial values
quad.mat <- as.matrix(quad.burn)
nsamp <- 10000
samp <- sample.int(nrow(quad.mat), nsamp)
xpred <- dat$depth
npred <- length(xpred)
ycred <- matrix(NA, nrow=nsamp, ncol=npred)
ypred <- matrix(NA, nrow=nsamp, ncol=npred)

##calculate interval values
#loop through and fill each row
for(g in seq_len(nsamp)){
  #sampled parameters
  theta = quad.mat[samp[g],]

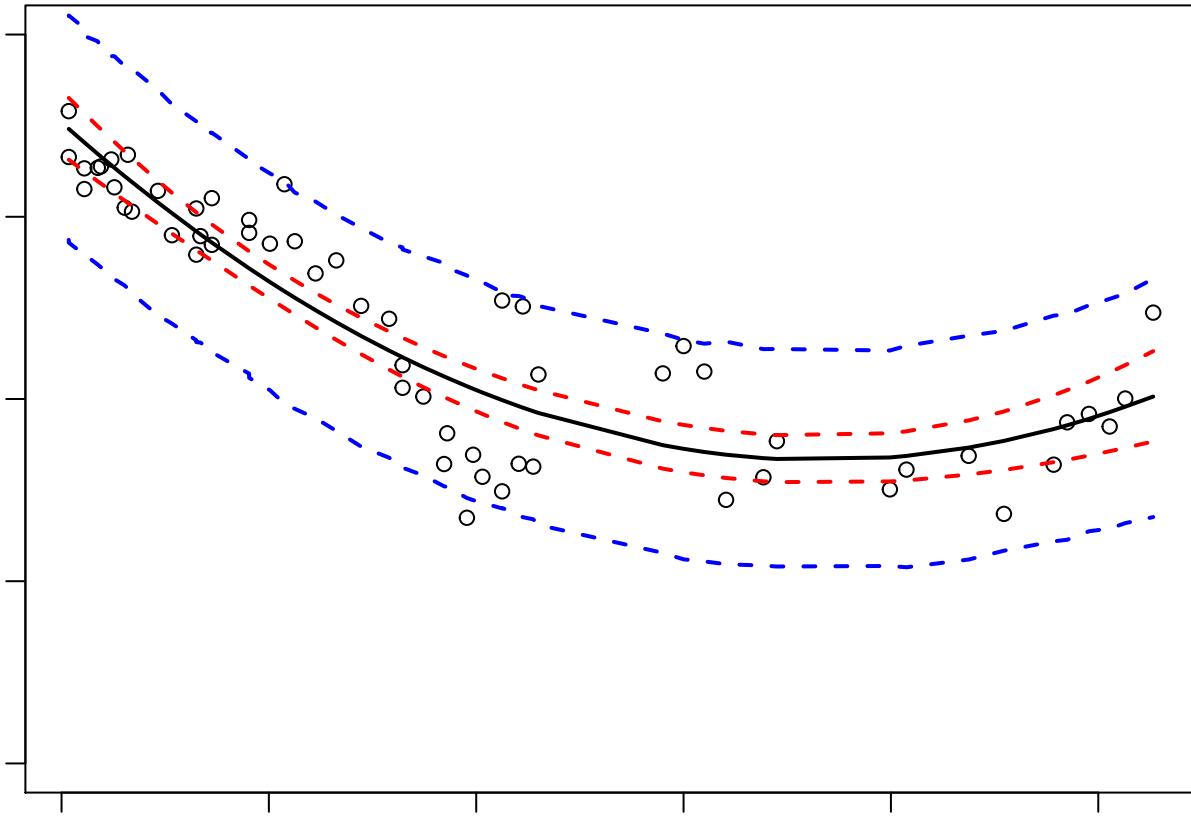
  #credible interval
  ycred[g,] <- theta["beta[1]"] + theta["beta[2]"]*xpred + theta["beta[3]"]*xpred^2

  #predictive interval
  ypred[g,] <- rnorm(npred, ycred[g,], 1/sqrt(theta["prec"]))
}

##intervals themselves
# credible interval and median
ci.quad <- apply(ycred, 2, quantile, c(0.025, 0.5, 0.975))
# prediction interval
pi.quad <- apply(ypred, 2, quantile, c(0.025, 0.975))

#plot it
plot(dat$depth, dat$deuturium, ylim= c(-280, -200))
lines(xpred, ci.quad[2,], type="l", col = 1, lwd= 2)    #median model
lines(xpred, ci.quad[1,], type="l", lty = 2, col = "red", lwd = 2)  #CI
lines(xpred, ci.quad[3,], type="l", lty = 2, col = "red", lwd = 2)
lines(xpred, pi.quad[1,], type="l", lty = 2, col = "blue", lwd = 2) #PI
lines(xpred, pi.quad[2,], type="l", lty = 2, col = "blue", lwd = 2)

```



Now lets fit a cubic (3rd order regression)

```
#bayesian fitting of polynomial (3rd order) regression (linear with respect to parameters)
cubic_regression<- "
model{

##priors
#betas
beta ~ dmnorm(b0,Vb)
#precision
prec ~ dgamma(s1,s2)

##process and data model
for(i in 1:n){
  # process model
  mu[i] <- beta[1] + beta[2]*x[i,1] + beta[3]*x[i,2] + beta[4]*x[i,3]
  # data model
  Isotope[i] ~ dnorm(mu[i],prec)
}
}

##specify data values
data.cubic<- list(Isotope = dat$deuturium, x = cbind(dat$depth, dat$depth^2, dat$depth^3), n = nrow(dat))

##priors
```

```

# regression beta means
data.cubic$b0 <- as.vector(c(0,0,0,0))
# regression beta precisions
data.cubic$Vb <- solve(diag(10000,4))
# uninformative error priors
data.cubic$s1 <- .1
data.cubic$s2 <- .1

#initial conditions
inits.cubic<- list()
for (i in 1:nchain){
  inits.cubic[[i]]<- list(beta = rnorm(4,0,5), prec = runif(1,1/1000,1/100))
}

#estimate from and sample model
cubic.model   <- jags.model(file = textConnection(cubic_regression),
                           data = data.cubic,
                           inits = inits.cubic,
                           n.chains = nchain)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
##   Observed stochastic nodes: 57
##   Unobserved stochastic nodes: 2
##   Total graph size: 459
##
## Initializing model

cubic.out<- coda.samples (model = cubic.model,
                           variable.names = c("beta", "prec"), n.iter = 5000)
cubic.DIC<- dic.samples(model = cubic.model, n.iter = 5000)

#burnin
burnin <- 1000
cubic.burn <- window(cubic.out, start = burnin)

#outputs
summary(cubic.burn)

##
## Iterations = 1000:5000
## Thinning interval = 1
## Number of chains = 3
## Sample size per chain = 4001
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##          Mean        SD  Naive SE Time-series SE
## beta[1] -2.111e+02 2.202e+00 2.010e-02      2.005e-02

```

```

## beta[2] -1.571e-01 4.409e-02 4.025e-04      3.993e-04
## beta[3]  3.224e-05 2.197e-04 2.006e-06      1.964e-06
## beta[4]  3.244e-07 2.895e-07 2.642e-09      2.628e-09
## prec     2.916e-02 5.642e-03 5.150e-05      5.393e-05
##
## 2. Quantiles for each variable:
##
##          2.5%       25%       50%       75%      97.5%
## beta[1] -2.155e+02 -2.126e+02 -2.112e+02 -2.097e+02 -2.067e+02
## beta[2] -2.436e-01 -1.867e-01 -1.569e-01 -1.280e-01 -6.883e-02
## beta[3] -4.036e-04 -1.135e-04  3.375e-05  1.800e-04  4.618e-04
## beta[4] -2.355e-07  1.312e-07  3.252e-07  5.168e-07  8.993e-07
## prec    1.937e-02  2.512e-02  2.886e-02  3.267e-02  4.114e-02

effectiveSize(cubic.burn)

## beta[1]  beta[2]  beta[3]  beta[4]      prec
## 12095.29 12216.77 12527.04 12136.51 10994.64

gelman.diag(cubic.burn)

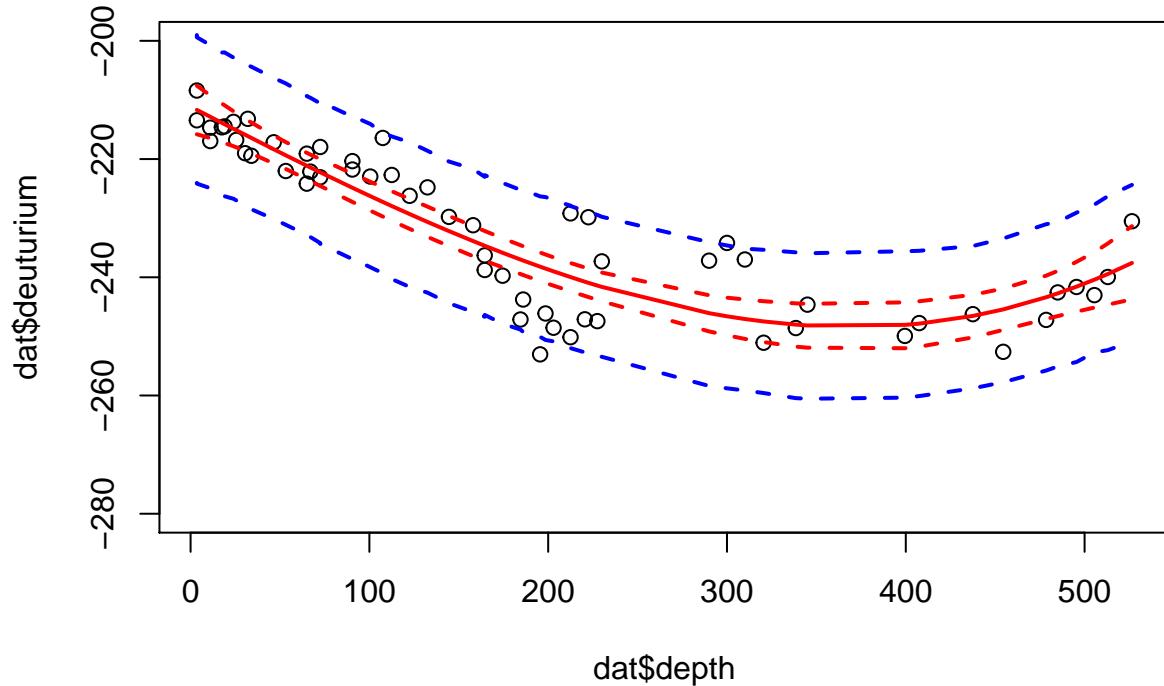
## Potential scale reduction factors:
##
##          Point est. Upper C.I.
## beta[1]        1        1
## beta[2]        1        1
## beta[3]        1        1
## beta[4]        1        1
## prec          1        1
##
## Multivariate psrf
##
## 1

#predictive and credible intervals
#initial values
cubic.mat <- as.matrix(cubic.burn)
nsamp <- 10000
samp <- sample.int(nrow(cubic.mat), nsamp)
xpred <- dat$depth
npred <- length(xpred)
ycred <- matrix(NA, nrow=nsamp, ncol=npred)
ypred <- matrix(NA, nrow=nsamp, ncol=npred)

#calculate intervals
for(g in seq_len(nsamp)){
  theta = cubic.mat[samp[g],]
  ycred[g,] <- theta["beta[1]"] + theta["beta[2]"]*xpred + theta["beta[3]"]*xpred^2 + theta["beta[4]"]*
    ypred[g,] <- rnorm(npred, ycred[g,], 1/sqrt(theta["prec"]))
}
ci.cubic <- apply(ycred, 2, quantile, c(0.025, 0.5, 0.975)) ## credible interval and median
pi.cubic <- apply(ypred, 2, quantile, c(0.025, 0.975))      ## prediction interval

```

```
#plot it
plot(dat$depth, dat$deuturium, ylim= c(-280, -200))
lines(xpred, ci.cubic[2,], type="l", col = "red", lwd= 2) #median model
lines(xpred, ci.cubic[1,], type="l", lty = 2, col = "red", lwd = 2) #CI
lines(xpred, ci.cubic[3,], type="l", lty = 2, col = "red", lwd = 2)
lines(xpred, pi.cubic[1,], type="l", lty = 2, col = "blue", lwd = 2) #PI
lines(xpred, pi.cubic[2,], type="l", lty = 2, col = "blue", lwd = 2)
```



And a quartic. JAGS crashes when trying to sample this model, the MCMC has been hardcoded in the next chunk

```
##quad reg model
quartic_regression<- "
model{

#priors
beta ~ dmnorm(b0,Vb)      ## prior regression params
prec ~ dgamma(s1,s2)    ## prior precision

#process and data model
for(i in 1:n){
  mu[i] <- beta[1] + beta[2]*x[i,1] + beta[3]*x[i,2] + beta[4]*x[i,3] +beta[5]*x[i,4] ## process
  Isotope[i] ~ dnorm(mu[i],prec)      ## data model
}
```

```

}

#specify data values
data.quartic<- list(Isotope = dat$deuturium, x = cbind(dat$depth, dat$depth^2, dat$depth^3, dat$depth^4))

##priors
# regression beta means
data.quartic$b0 <- as.vector(c(0,0,0,0,0))
# regression beta precisions
data.quartic$Vb <- solve(diag(10000,5))

#uninformative error prior
data.quartic$s1 <- .1
data.quartic$s2 <- .1

#initial conditions
inits.quartic<- list()
for (i in 1:nchain){
  inits.quartic[[i]]<- list(beta = rnorm(5,0,5), prec = runif(1,1/1000,1/100))
}

##fit and sample from model
#jags object
quartic.model <- jags.model(file = textConnection(quartic_regression),
                               data = data.quartic,
                               inits = inits.quartic,
                               n.chains = nchain)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph information:
##   Observed stochastic nodes: 57
##   Unobserved stochastic nodes: 2
##   Total graph size: 577
##
## Initializing model

#sample from model will crash JAGS
#quartic.out<- coda.samples (model = quartic.model, variable.names = c("beta", "prec"), n.iter = 5000)

```

Now lets do the MCMC the hard way... Beta[5] appears to be non identifiable

```

##manual sampling of quartic reg
#define data
x<- cbind(rep(1,nrow(dat)),dat$depth, dat$depth^2, dat$depth^3, dat$depth^4)
y<- dat$deuturium

## specify priors
#number of itterations
n.g<- 50000

```

```


#prior on the mean


bprior <- as.vector(c(0,0,0,0,0))


#inverse of prior variance


vinvert <- solve(diag(1000,5))


#inverse gamma prior on variance so 1/s1 from the precision prior


s1 <- 10
s2 <- 10
n <- nrow(dat) #sample size, should this be my sample or prior sample size? right now its sample

##precompute frequently used quantities
XX <- t(x) %*% x
XY <- t(x) %*% y
VbB <- vinvert %*% bprior

##load libraries
library(coda)
library(mvtnorm)

## Gibbs loop
gibbs_loop<- function(itt){
  ## initial conditions
  sg <- 50
  sinv <- 1/sg

  #mcmc storage
  ngibbs<- length(itt)
  bgibbs <- matrix(0.0,nrow=ngibbs,ncol=5) ## storage for beta
  sgibbs <- numeric(ngibbs) ## storage for sigma2
  dgibbs <- numeric(ngibbs) ## storage for deviance at each itteration

  for(g in itt){

    ## sample regression parameters
    bigV <- solve(sinv*XX + v invert, tol=1e-25) ## Covariance matrix
    littlev <- sinv*XY + VbB
    b = t(rmvnorm(1,bigV %*% littlev,bigV)) ## Vu is the mean vector

    ## sample variance
    u1 <- s1 + n/2
    u2 <- s2 + 0.5*crossprod(y-x%*%b)
    sinv <- rgamma(1,u1,u2)
    sg <- 1/sinv

    ## storage
    bgibbs[g,] <- b ## store the current value of beta vector
    sgibbs[g] <- sg ## store the current value of the variance

    ##deviance
    l_lik <- sum(dnorm(y, x%*%b, sg*1, log=T)) #log liklihood
    dgibbs[g] <- -2*l_lik #deviance
  }
}

```

```

#create mcmc outputs
allgibbs<- cbind(bgibbs,sgibbs)
colnames(allgibbs)<- c("beta1", "beta2", "beta3", "beta4", "beta5", "Variance")
all_mcmc<- mcmc(allgibbs)

return(list(all_mcmc, dgibbs))

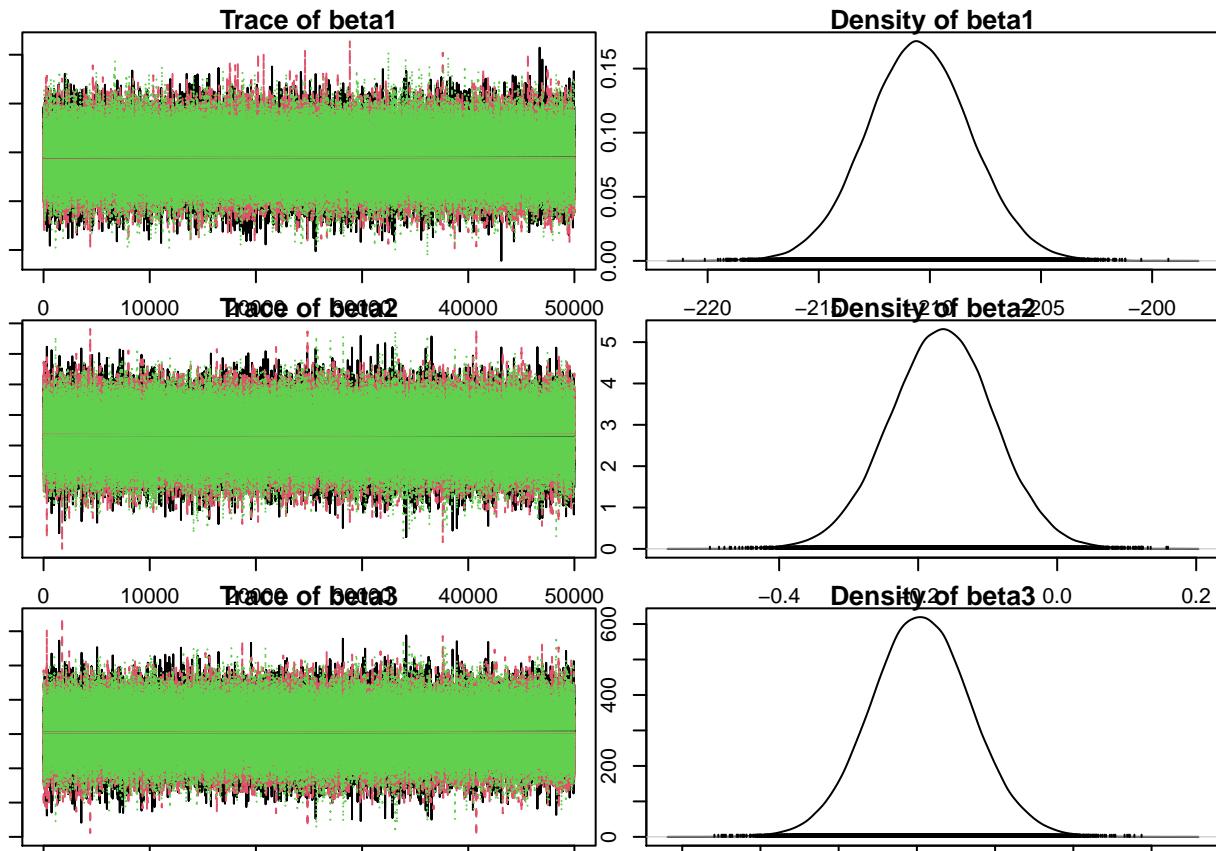
}

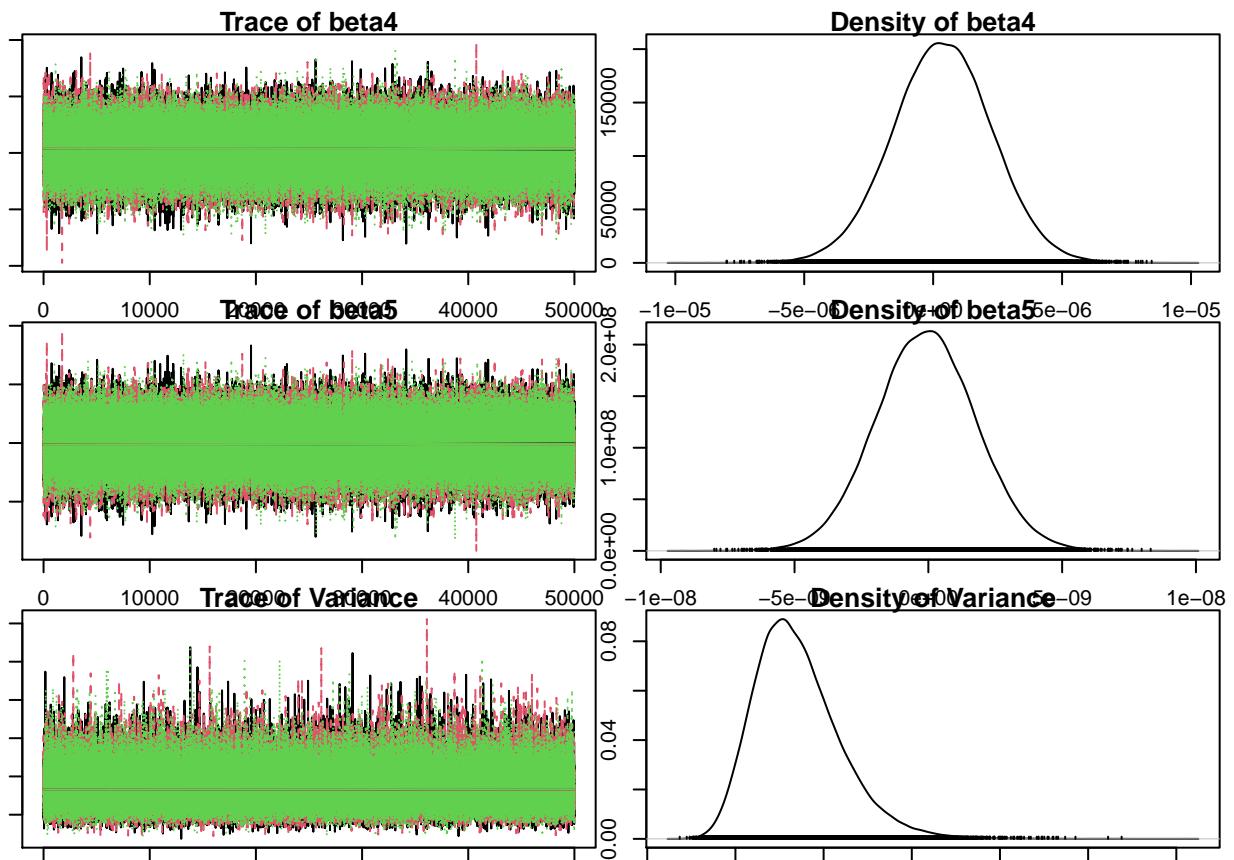
#run mcmc, extract outputs and deviance
gibbs_list<- lapply(list(1:n.g, 1:n.g, 1:n.g), gibbs_loop) #run three chains
gibbs_mcmc <- mcmc.list(list(gibbs_list[[1]][[1]], gibbs_list[[2]][[1]], gibbs_list[[3]][[1]])) #mcmc
gibbs_dev <- list(gibbs_list[[1]][[2]], gibbs_list[[2]][[2]], gibbs_list[[3]][[2]]) #deviance vector

#calculate DIC
gibbs_sum<- summary(gibbs_mcmc)
beta_bar<- gibbs_sum$statistics[1:5,1] #mean beta
var_bar<- gibbs_sum$statistics[6,1] #mean variance
D_thetabar <- -2*sum(dnorm(y, x%*%beta_bar, var_bar*1, log=T))
Dtheta_bar <- mean(c(gibbs_dev[[1]], gibbs_dev[[2]], gibbs_dev[[3]]))
quartic.DIC <- 2*Dtheta_bar - D_thetabar

#outputs
par(mar=c(1,1,1,1))
plot(gibbs_mcmc)

```





```
effectiveSize(gibbs_mcmc)
```

```
##      beta1     beta2     beta3     beta4     beta5 Variance
## 149018.2 150000.0 150000.0 150000.0        0.0 125462.9
```

```
gelman.diag(gibbs_mcmc)
```

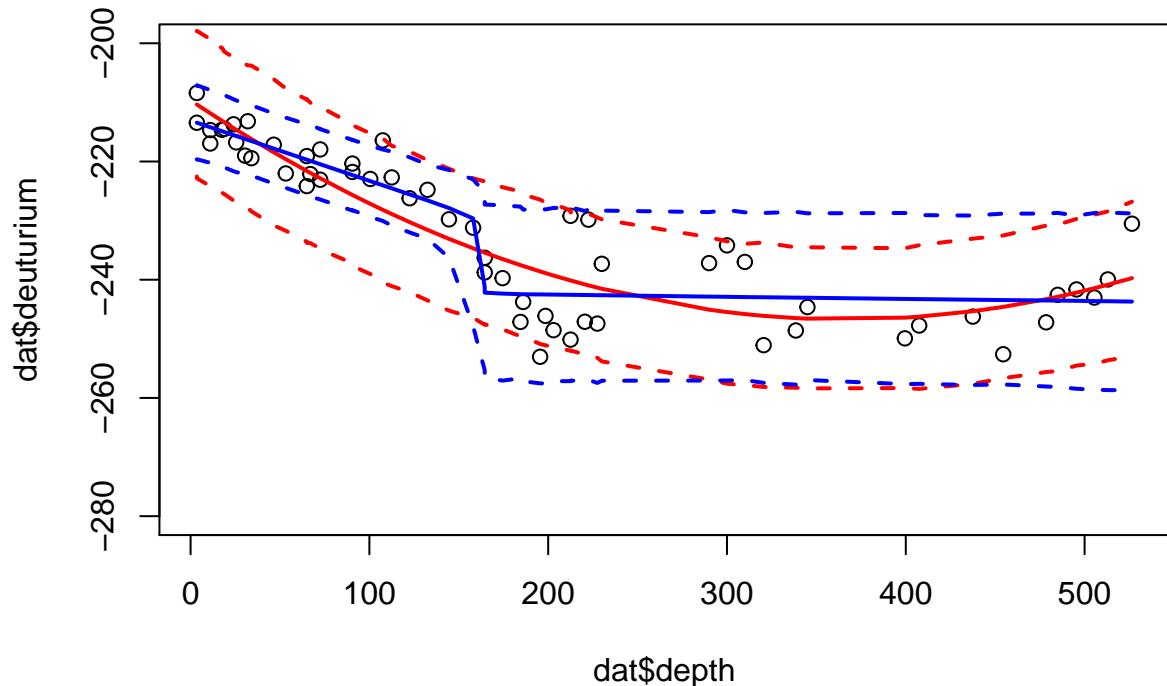
```
## Potential scale reduction factors:
##
##          Point est. Upper C.I.
## beta1          1         1
## beta2          1         1
## beta3          1         1
## beta4          1         1
## beta5          1         1
## Variance       1         1
##
## Multivariate psrf
##
## 1
```

```
#compare DICs
#quad.DIC
#cubic.DIC
quartic.DIC
```

```
## [1] 480.6342
```

lets compare the predictive intervals between best quad and best BP

```
#plot to compare model intervals between quad and breakpoint
plot(dat$depth, dat$deuturium, ylim=c(-280, -200))
lines(xpred, ci.quad[2,], type="l", col = "red", lwd= 2)
lines(xpred, pi.quad[1,], type="l", lty = 2, col = "red", lwd = 2)
lines(xpred, pi.quad[2,], type="l", lty = 2, col = "red", lwd = 2)
lines(xpred, ci.bp[2,], type="l", col = "blue", lwd= 2) #median model
lines(xpred, pi.bp[1,], type="l", lty = 2, col = "blue", lwd = 2) #PI
lines(xpred, pi.bp[2,], type="l", lty = 2, col = "blue", lwd = 2)
```



```
#hahahahahaha looks like shit with too many lines
```