Supporting Information

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Appendix 1: Implementation of the Crofton Method 2

- The algorithm for fitting the Crofton Method (Crofton 1971) proceeds as follows. 3 First, obtain a dataset with n hosts where each host has some parasite intensity 0 4 5 to p_{max} . Starting with the full dataset, guess a vector of pre-mortality parameters (N_p, μ_p, k_p) where N_p is the total number of hosts before mortality, μ_p is the 6 parasite intensity before mortality, and k_p is the parasite aggregation before mor-7 tality. Given these parameters, use a negative binomial distribution to calculate the 8 predicted number of hosts with $0, 1, 2, \dots p_{max}$ parasites. Compare the expected 9 number of hosts with $0, 1, 2, \dots p_{max}$ parasites to the observed number hosts with 10 $0,1,2,\ldots p_{max}$ parasites and calculate the χ^2 -squared statistic associated with 11 your observed and predicted vectors. In reality, one often has to bin the parasite 12 intensity data because all parasite intensities are not represented in the dataset. 13 Continue to guess (N_p, mu_p, k_p) vectors until a set of parameters is found that 14 minimizes the χ^2 - squared statistic. 15 Second, choose a truncation value (t_1) such that $t_1 < p_{max}$. Truncate the 16 data such that $data_{truncated} \ll t_1$ and repeat the above iterative procedure to 17 calculate another set of parameters $(N_{p2}, \mu_{p2}, k_{p2})$ that minimizes the χ^2 -squared 18 statistic on the truncated data. Choose a new truncated value $t_2 \ < \ t_1$ and 19 20 repeat the first two steps. Continue to truncate the dataset until it only contains hosts with 0, 1, and 2 parasites (or 3 bins). Because the method attempts to 21 estimate three parameters, at least 3 classes are needed for all 3 parameters 22 to be identifiable (Royce & Rossignol 1990). Parasite-induced host mortality is 23 traditionally identified by then plotting your truncation values t_i against your 24 different values of N_{po} and looking for a distinct kink in the resulting plot.
 - We provide an implementation and unit tests of the Crofton Method in

Supplementary Information 4. Figure 10 visually shows that our implementation of the Crofton Method agrees with results previously published by Crofton (1971).

Appendix 2: Implementation of the Adjei Method

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The Adjei Method for estimating PIHM has two steps (Adjei et al. 1986). The first 30 31 step is to estimate the parameters of the pre-mortality host-parasite distribution using the Crofton Method. The three parameters estimated are the total number of 32 hosts before mortality N_p , the mean number of parasites per host before mortality 33 μ_p , and the aggregation of parasites before mortality given by the parameter k_p 34 from a negative binomial distribution. When k_p is small, parasites are highly 35 aggregated among hosts and when k_p is large parasites are more evenly spread 36 out (Wilson et al. 2002). The implementation of the Crofton Method has been 37 discussed at length elsewhere (e.g. Royce & Rossignol 1990; Lester 1984, and 38 39 above) and we provide a tested implementation of the method in SI 4. The second step of the Adjei Method is to estimate the shape of the host 40 survival function. Adjei et al. (1986) assume that the host survival function follows 41 the logistic form 42

$$h(x|a,b) = h_x = \frac{e^{a-b\log(x)}}{1 + e^{a-b\log(x)}}$$
 (1)

where x is the parasite intensity in a given host and a and b are the two parameters of the logistic function. Generally, a larger a allows for hosts to tolerate larger parasite intensities before experiencing parasite-induced mortality and a larger bleads to a more rapid decline in the probability of host survival as parasite intensity increases. The value $\exp(a/b)$ is referred to as the LD_{50} , which gives the parasite intensity at which 50% of hosts experience mortality.

To estimate this function, the Adjei Method first calculates the expected

number of hosts with a given parasite load x by using the equation $g(x; \mu_p, k_p) * N_p$, where $g(x; \mu_p, k_p)$ is the negative binomial pre-mortality distribution. Second, the observed and predicted number of hosts with x parasites are paired as a single data point and the method then assumes that this data point follows a binomial distribution with the total number of "trials" equal to the predicted number of hosts and the total number of "successes" equal to the observed number of hosts. In some cases, the observed number of hosts is greater than the expected number of hosts and the Adjei Method alters the data so that the observed is equal to the predicted (Adjei et al. 1986). After this questionable manipulation, the (observed, predicted) pairs are fit to a standard Generalized Linear Model (McCullagh & Nelder 1989) with a binomial response variable and a logistic link function given by equation 1. This model provides estimates for parameters a, b and LD_{50} .

While not included in the original implementation of the Adjei Method, a χ^2 test with a degrees of freedom of 1 can be used to assess whether a GLM model that includes parasite intensity as a predictor of host survival probability is a "better" model than a GLM without this predictor. This allows the Adjei Method to determine whether PIHM is a significant factor in a host-parasite system.

The Adjei Method's most glaring deficiency is the need to alter the observed data in order to fit the model into the binomial GLM framework. A second more subtle problem with the Adjei Method is the potential need to bin data in order to predict greater than one host in a given parasite intensity class. For example, if the total number of hosts pre-mortality was 50, the mean number of parasites per host pre-mortality was 100 and the aggregation parameter was 1, applying the equation $g(x; \mu_p = 100, k_p = 1) * 50$ would result in less than 1 individual in all parasite intensities x. In other words, the Adjei Method cannot be applied to samples with either very high mean parasite loads, small sample sizes, or both without some sort of binning of the data. While this is not a flaw per se, it does add a certain level of subjectivity (i.e. which bins should you use?) to a method

- 78 that already has serious potential issues. In this analysis, we always assume the
- 79 Adjei Method is not binning the data, though we provide code for applying the
- 80 binning method in SI 4.

81 Appendix 3: Additional Figures

82 See Figures 1 - 9.

83 Appendix 4: Code and unit tests for estimating parasite-

84 induced host mortality

- 85 Python code, unit tests, and a help file for the Crofton Method, the Ad-
- 86 jei Method and the Likelihood Method can be found in pihm_methods.py,
- 87 test_pihm_methods.py, and help_file.txt, respectively.

88 References

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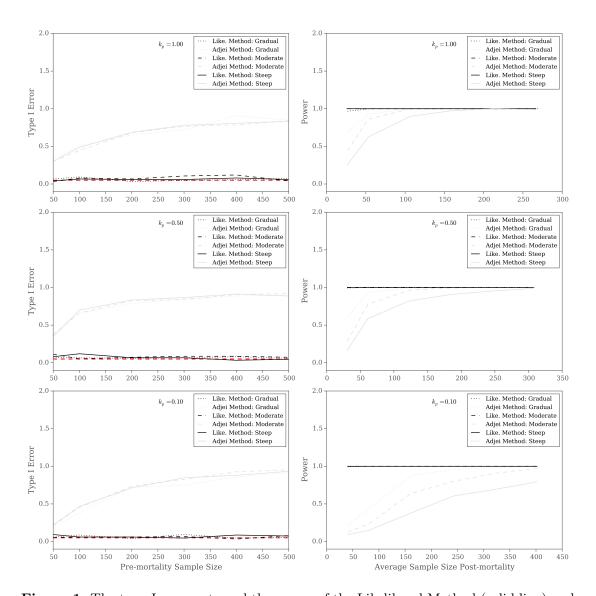


Figure 1: The type I error rate and the power of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 10$ for various shapes of the host survival function and levels of aggregation k_p . The first column gives the type I error rate of each method for falsely detecting PIHM when none is present. The red line gives the the pre-set type I error rate of $\alpha = 0.05$. The second column gives the power of a given method to detect PIHM when it is actually occurring.

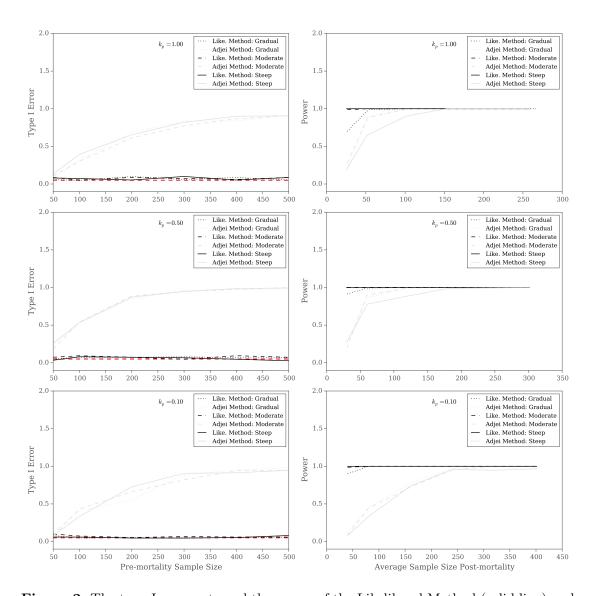


Figure 2: The type I error rate and the power of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 50$ for various shapes of the host survival function and levels of aggregation k_p . The first column gives the type I error rate of each method for falsely detecting PIHM when none is present. The red line gives the the pre-set type I error rate of $\alpha = 0.05$. The second column gives the power of a given method to detect PIHM when it is actually occurring.

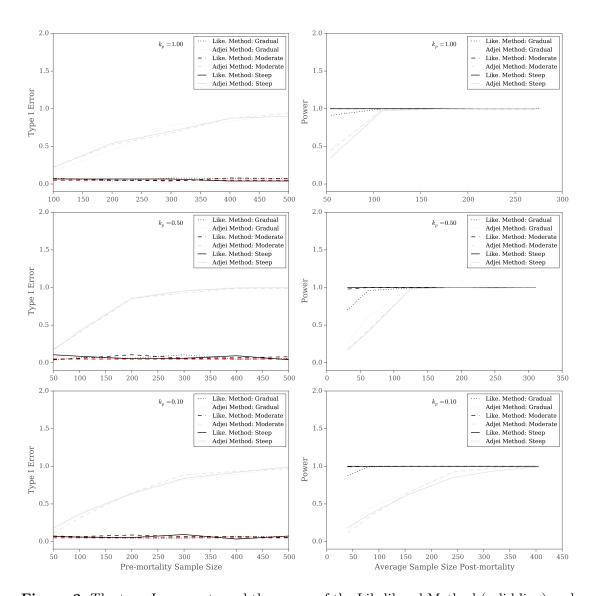


Figure 3: The type I error rate and the power of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 100$ for various shapes of the host survival function and levels of aggregation k_p . The first column gives the type I error rate of each method for falsely detecting PIHM when none is present. The red line gives the the pre-set type I error rate of $\alpha = 0.05$. The second column gives the power of a given method to detect PIHM when it is actually occurring.

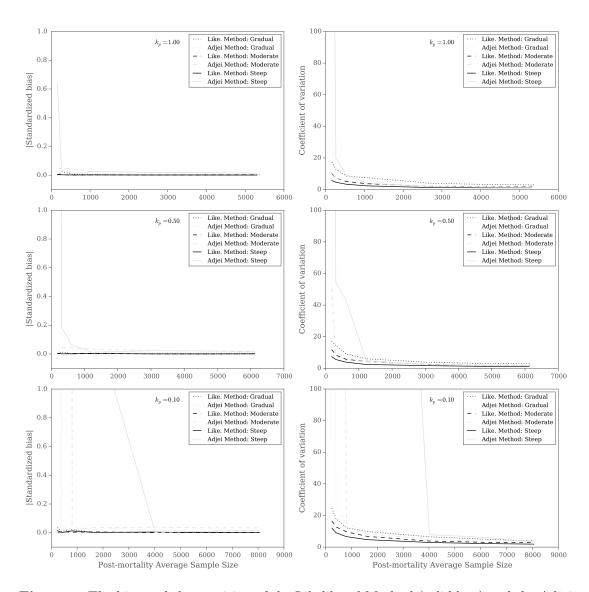


Figure 4: The bias and the precision of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 10$ for various shapes of the host survival function and levels of aggregation k_p when estimating LD_{50} . The first column gives the bias of each method's LD_{50} estimate over 150 simulations. The second column gives the precision of each method's LD_{50} estimate over 150 simulations.

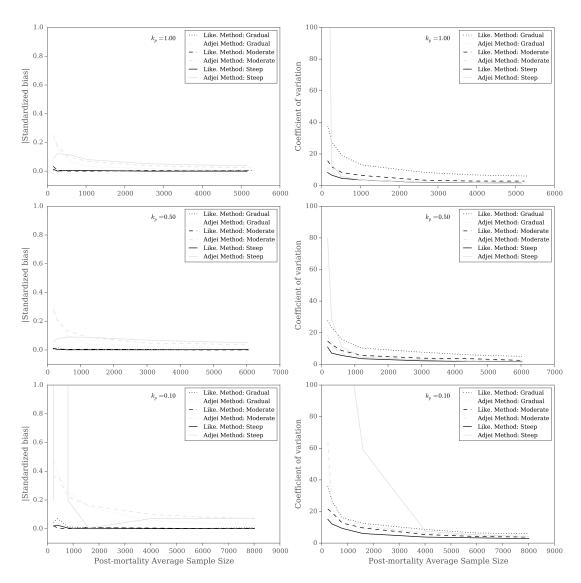


Figure 5: The bias and the precision of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p=50$ for various shapes of the host survival function and levels of aggregation k_p when estimating LD_{50} . The first column gives the bias of each method's LD_{50} estimate over 150 simulations. The second column gives the precision of each method's LD_{50} estimate over 150 simulations.

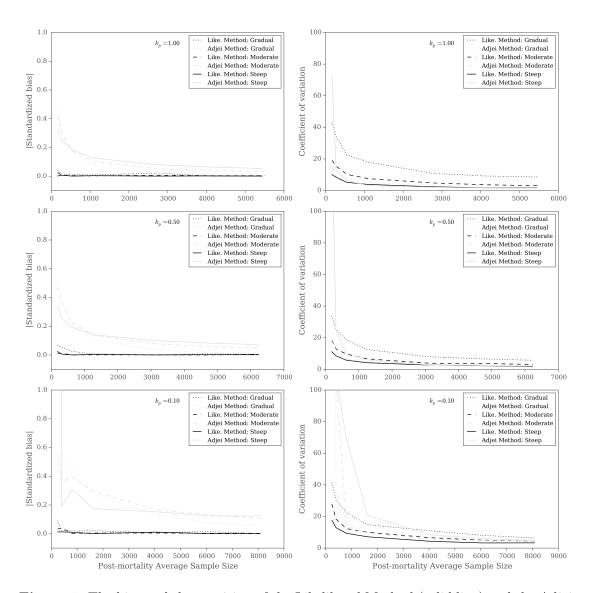


Figure 6: The bias and the precision of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 100$ for various shapes of the host survival function and levels of aggregation k_p when estimating LD_{50} . The first column gives the bias of each method's LD_{50} estimate over 150 simulations. The second column gives the precision of each method's LD_{50} estimate over 150 simulations.

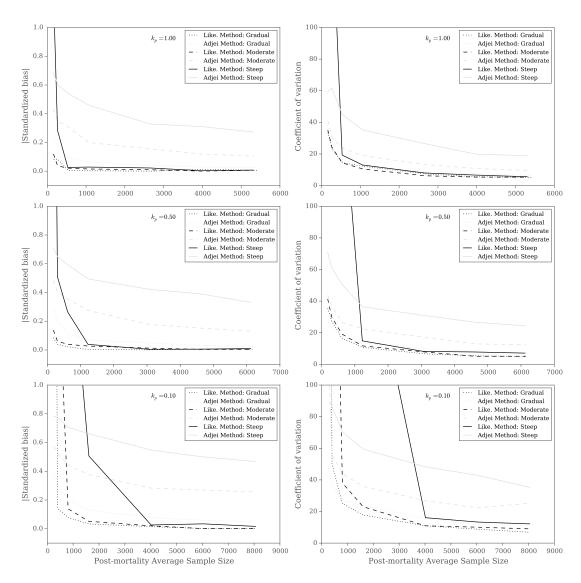


Figure 7: The bias and the precision of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p=10$ for various shapes of the host survival function and levels of aggregation k_p when estimating the a parameter of the host survival function. The first column gives the bias of each method's a estimate over 150 simulations. The second column gives the precision of each method's a estimate over 150 simulations.

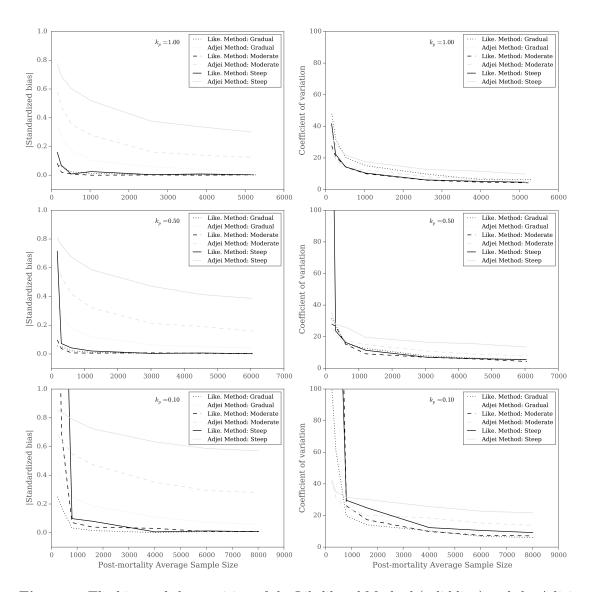


Figure 8: The bias and the precision of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 50$ for various shapes of the host survival function and levels of aggregation k_p when estimating the a parameter of the host survival function. The first column gives the bias of each method's a estimate over 150 simulations. The second column gives the precision of each method's a estimate over 150 simulations.

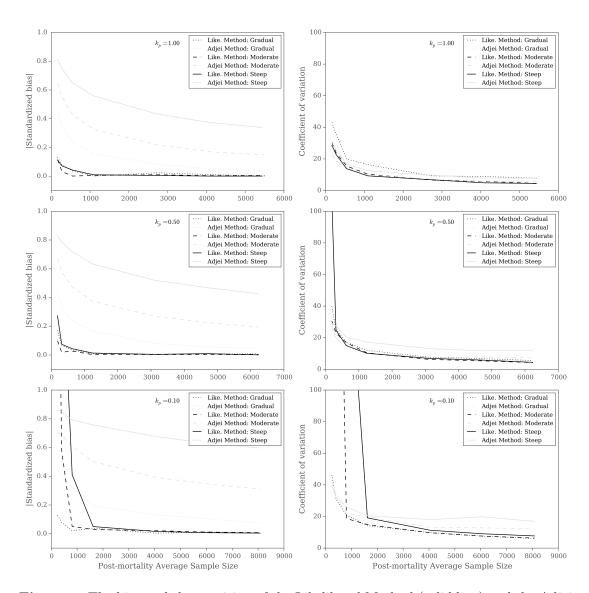


Figure 9: The bias and the precision of the Likelihood Method (solid line) and the Adjei Method (dashed lines) when $\mu_p = 100$ for various shapes of the host survival function and levels of aggregation k_p when estimating the a parameter of the host survival function. The first column gives the bias of each method's a estimate over 150 simulations. The second column gives the precision of each method's a estimate over 150 simulations.

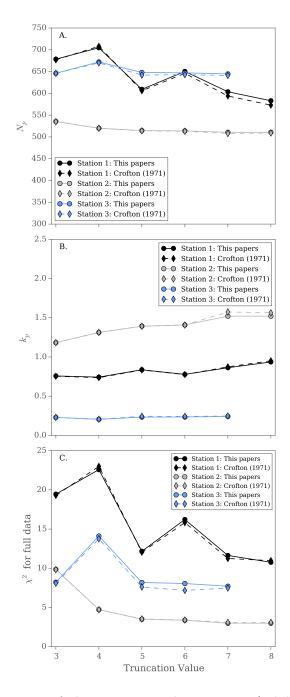


Figure 10: A comparison of this papers implementation (solid line, circles) of the Crofton Method with the results given in Crofton (1971) (dashed line, diamonds). Figure A compares the predicted number of hosts in a population pre-mortality (N_p) . Figure B compares the predicted parasite aggregation pre-mortality (k_p) . Figure C compares the χ^2 statistic for each implementation. Three of the 6 stations fit by Crofton are shown here and all show that our implementation gives very similar results to those given by Crofton.