

Detecting and quantifying parasite-induced host mortality from intensity data: method comparisons and limitations

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Abstract

Parasites can often have significant impacts on host populations through parasite-induced host mortality. Detecting and quantifying the magnitude of parasite-induced mortality, particularly for macroparasites in which pathology is linked to the intensity of infection, can provide important insight into parasite dynamics and host population management. This is, however, notoriously difficult in practice as the only data often available are measures of parasite intensities. In this study we provide a novel method for detecting and quantifying parasite-induced mortality in wildlife populations from intensity data and, using simulations, we show that our method is more reliable for detecting and quantifying parasite-induced host mortality than the methods that are currently available. We also show that this method provides quantitative estimates of parasite-induced mortality in empirical data that are consistent with previously published qualitative estimates. However, we stress that this method, along with all methods for estimating parasite-induced mortality in host populations from intensity data alone, has a number of critical assumptions that limit their applicability in the real world. We conclude that methods for estimating and detecting parasite-induced mortality from intensity data should be used only as an exploratory tool for informing more rigorous studies of parasite-induced host mortality.

1 Introduction

Infectious agents can have major impacts on animal populations through changing population dynamics and stability (Dobson & Hudson 1992), altering predator-prey interactions (Joly & Messier 2004), and even causing species' decline and extinction (De Castro & Bolker 2005; McCallum 2012). Accurately estimating the impact of these infectious agents in wildlife is critical to understanding what regulates host and

7 parasite populations, making predictions about disease transmission, and managing
8 disease outbreaks (Langwig *et al.* 2015). The impact of microparasite pathogens, such
9 as rabies (Coyne *et al.* 1989), bovine TB (Cox *et al.* 2005), and rinderpest (Tillé *et al.*
10 1991), is typically quantified based on the presence or absence of disease, and does not
11 account for the number of infectious agents present. This method is sufficient for many
12 bacterial and viral agents that reproduce within a host, however for macroparasites,
13 pathology is linked to the intensity of infection and hosts cannot be simply categorized
14 as infected and uninfected (Anderson & May 1979; Lafferty & Kuris 2002). Helminths
15 exhibiting this intensity dependent pathology have significant impacts on human health
16 (Brooker *et al.* 2004), domestic livestock economics (Roeber *et al.* 2013), and wildlife
17 survival (Kirk 2003; Logiudice 2003). While it is generally assumed that some fraction
18 of wild host populations must succumb to parasitic infections, it is notoriously difficult
19 to actually quantify parasite-induced host mortality (PIHM) in wild animal populations
20 (McCallum 2000).

21 Ideally, parasite-induced host mortality is quantified by experimentally infecting
22 and tracking individual hosts in the wild population; however, for logistical and ethical
23 reasons this method is rarely feasible (McCallum 2000). Data on parasite intensity is much
24 easier to collect and has often been used to identify the presence of PIHM (Crofton 1971;
25 Lester 1977, 1984; Lanciani & Boyett 1989; Royce & Rossignol 1990; Ferguson *et al.*
26 2011) and to quantify the relationship between infection intensity and host mortality
27 (Adjei *et al.* 1986).

28 Crofton (1971) first proposed that PIHM could be identified by comparing the
29 observed parasite distribution in the host population to the distribution predicted in the
30 absence of parasite-induced mortality. We briefly introduce the Crofton Method here and
31 provide a more detailed explanation of its implementation in *Supplementary Material*
32 (*SI*) 1. This method assumes that, prior to host mortality, infection intensity in the
33 host population follows a negative binomial distribution and the tail of the distribution
34 is truncated as intensity dependent pathology removes the most heavily infected hosts.

35 Assuming mortality occurs only in heavily infected hosts, evidence of this parasite-induced
36 mortality should then be detectable by iteratively fitting a negative binomial distribution
37 to hosts with lower and lower parasite loads, and comparing these truncated predicted
38 distributions to the corresponding truncated observed parasite data. [FIGURE]

39 The Crofton Method may be able to detect the presence of PIHM, but it does
40 not quantify the relationship between infection intensity and host survival probability.
41 Adjei *et al.* (1986) suggested that this relationship could be calculated by first using
42 the Crofton Method to estimate the pre-mortality parasite distribution and then using
43 this distribution to calculate the probability of host survival with increasing parasite
44 intensity. To do this, Adjei *et al.* (1986) modeled host survival as a logistic function and
45 then used a generalized linear model (GLM) to estimate the logistic parameters (see SI 2
46 for a technical description of the Adjei Method). Adjei *et al.* suggested that this method
47 could provide an estimate for the parasite intensity at which a host has a 50% chance
48 of suffering parasite-induced mortality (LD_{50}). However, to implement this method the
49 observed data must be modified to fit the GLM framework and subjectively binned when
50 mean infection intensity is high or sample sizes are small (see SI 2 for details).

51 After 30 years, and despite clear limitations (McCallum 2000), these methods
52 (particularly the Crofton Method) are still discussed among parasitologists and are the
53 primary techniques for examining population level impacts of parasitism using parasite
54 intensity data. In these methods, PIHM can only be identified by visually examining plots
55 and, with no clear decision rule, it can be difficult to determine the significance of PIHM
56 across different host-parasite systems. The survival function given by the Adjei Method
57 offers one solution; however, this method requires manipulating the original data and its
58 accuracy has never been validated.

59 Intensity data should be used to estimate parasite impacts on host populations
60 only if unbiased and accurate methods exist. In this study, we first propose a novel method
61 for detecting and quantifying PIHM. We next use simulations to compare our method with
62 the previous Adjei Method to test the ability of both to (1) detect occurrence of PIHM

and (2) estimate the lethal parasite load (LD_{50}) and the associated survival function. We then apply both methods to real datasets previously used in PIHM analyses and compare the results. Finally, we discuss the limitations of inferring PIHM from intensity data and whether any method for inferring PIHM has a place in quantitative parasitology.

Methods

A novel, likelihood-based method for estimating PIHM

Here we propose a novel, likelihood-based method (henceforth Likelihood Method) that does not require binning or data alteration, reduces the number of parameters to be estimated, is highly generalizable, and uses standard statistical techniques to determine PIHM significance. The Likelihood Method first assumes that prior to mortality the parasite distribution can be described by the distribution $g(x; \phi)$, which specifies the probability of a host having x parasites when it is observed. ϕ is a vector of parameters that described the shape and scale of this distribution.

The Likelihood Method then assumes that the probability of a host surviving for t units given it has some parasite intensity x is described by the function $h(\text{survival}; x, t, \theta)$ where θ is a vector of parameters of the survival function. For most parasite-intensity datasets, the observer has no knowledge of how long a host has been infected with x . One way to account for this would be to integrate out time such that we are left with the function $h(\text{survival}; x, \theta)$ specifying the probability of survival given x parasites and some additional parameters θ . While eliminating the functions dependence on time is far from ideal, this is an inherent limitation of attempting inference on cross-sectional data (jargon).

With these two assumptions, we can define a probability distribution that gives the probability of having a parasite load of x parasites conditional on host survival. Using standard rules of conditional probability this distribution can be written as

$$P(x|\text{survival}) = \frac{P(\text{survival}|x) * P(x)}{P(\text{survival})} \quad (1)$$

One can see that $P(\text{survival}|x)$ is the survival function $h(\text{survival}; x, \theta)$, $P(x)$ is the pre-mortality parasite distribution $g(x; \phi)$ and $P(\text{survival}) = \sum_{x=0}^{\infty} P(\text{survival}|x) * P(x) = \sum_{x=0}^{\infty} h(\text{survival}; x, \theta) * g(x; \phi)$. Therefore equation 1 can be written as

$$P(x|\text{survival}) = \frac{h(\text{survival}; x, \theta) * g(x; \phi)}{\sum_{x=0}^{\infty} h(\text{survival}; x, \theta) * g(x; \phi)} \quad (2)$$

Using this probability distribution, one can then find the parameters θ and ϕ that maximize the likelihood of an observed host-parasite dataset. To estimate the significance of PIHM in a host-parasite system, a likelihood ratio test can be used in which the full model is given by equation 2 and the reduced model is given by the pre-mortality distribution $g(x; \phi)$. If PIHM is not significant in the system, the resulting likelihood ratio statistic should approximately follow a χ^2 distribution with two degrees of freedom.

Equation 2 could be parameterized in many different ways depending on the parasite system of interest. In this study, we will follow the precedent set by all previous methods for estimating PIHM and assume that the pre-mortality parasite distribution $g(x; \phi)$ follows a negative binomial distribution with the parameters mean parasite intensity (μ_p) and aggregation (k_p) before mortality, respectively (smaller k_p indicates more aggregation). The negative binomial distribution is not just a phenomenological assumption and can arise as the equilibrium parasite distribution under a variety of different biological assumptions (Calabrese *et al.* 2011) [MORE]. However, it is also an incredibly flexible distribution that fits many host-parasite systems regardless of whether the underlying mechanisms lead to an exact negative binomial distribution (Shaw *et al.* 1998).

Choosing a function for $h(\text{survival}; x, \theta)$

The equation $\exp(a/b)$ can then be used to calculate the parasite LD_{50} , here defined as the infection intensity at which 50% of hosts experience PIHM. All parameters

111 are defined in Table 1.

112 Evaluating the Adjei and Likelihood Methods

113 *Question 1: Can we detect PIHM?*

114 We tested the ability of the Adjei and the Likelihood Methods to identify the
115 presence of PIHM on simulated data with known pre-mortality parameters. First, we
116 created a pre-mortality host population by drawing N_p randomly infected hosts from a
117 negative binomial distribution with parameters μ_p and k_p . Second, we chose values of a
118 and b and calculated the probability of survival for all N_p hosts using equation ???. Then,
119 for each host, we drew a random number from a uniform distribution between 0 and 1
120 and if the calculated host survival probability was less than this random number, the
121 host experienced parasite-induced mortality. The parasite distribution in these simulated
122 surviving hosts is equivalent to the observed parasite distribution in a wild host population
123 that has undergone parasite-induced host mortality.

124 We used these simulated pre-mortality and post-mortality datasets to test the
125 ability of both methods to correctly determine whether or not PIHM was occurring when
126 the parameters N_p , μ_p and k_p were known. Although the parameters N_p , μ_p , and k_p
127 are always unknown in real systems, a method that fails under these ideal simulation
128 conditions will certainly also fail using less ideal, empirical data. In practice, for the
129 Adjei Method, N_p , μ_p , and k_p are estimated using the Crofton Method (Adjei *et al.*
130 1986), while μ_p and k_p in the Likelihood Method can be estimated jointly with a and b
131 or via the Crofton Method.

132 We used three different values of μ_p (10, 50, 100) and for each μ_p we examined
133 three different survival functions that had graduate, moderate, and steep decreases in
134 the host survival with increasing parasite intensity (Figure 1A). For a given μ_p , each
135 survival function had the same LD_{50} ($[\mu_p = 10, LD_{50} = 7.39]$, $[\mu_p = 50, LD_{50} = 35.57]$,
136 $[\mu_p = 100, LD_{50} = 77.3]$), but different values of a and b . We examined each μ_p -survival

function pair at three levels of parasite aggregation, $k_p = 0.1, 0.5$, and 1 — realistic values of parasite aggregation in natural populations (Shaw *et al.* 1998). For each of these parameter combinations we simulated 150 datasets and tested the probability of each method correctly identifying PIHM in the post-mortality dataset (power) and incorrectly identifying PIHM in the pre-mortality dataset (Type I error). For each method, we used a likelihood ratio test to determine whether the full model with PIHM provided a significantly better fit than the reduced model without PIHM at significance level of 0.05 . We tested each parameter combinations for pre-mortality population sizes of $N_p = [50, 100, 200, 300, 400, 500]$. N_p is not technically the sample size on which the methods are being tested on the post-mortality data because PIHM reduces N_p for each simulated dataset. We therefore used the average number of surviving hosts over all 150 simulations for a given parameter combination as our measure of sample size in the power simulations.

Question 2: Can we estimate fatal parasite intensity and the host survival function?

To compare the ability of the Adjei Method and the Likelihood Method to recover the LD_{50} and the parameters a and b of the survival function, we used the same simulation procedure and parameter combinations described above. For each parameter combination we simulated 150 datasets, estimated a , b , and LD_{50} and calculated the standardized bias and precision (Walther & Moore 2005) for these estimates over pre-mortality host population sizes of $N_p = [300, 500, 1000, 2000, 5000, 7500, 10000]$. We used the average number of surviving hosts over all 150 simulations for a given parameter combination as our measure of sample size. Because parameters a and b showed similar patterns of bias and precision, we only show the results for a .

Efficacy of the Likelihood Method with unknown pre-mortality parameters

In the final simulation, we tested the ability of the Likelihood Method to correctly identify PIHM and estimate LD_{50} when the pre-mortality parameters are unknown. The previous

163 simulations showed that the Likelihood Method effectively identified PIHM when μ_p and
 164 k_p were known with values of 10 and 1, respectively. As a best-case scenario, we simulated
 165 host- parasite systems with these pre-mortality parameters and tested the power of the
 166 Likelihood Method to identify PIHM for gradual, moderate and steep survival functions
 167 when the pre-mortality parameters μ_p and k_p also needed to be estimated. We perform
 168 500 simulations over a range of different samples sizes following the simulation procedure
 169 described above.

170 **Application to real data**

171 We tested the ability of the Adjei Method and the Likelihood Method to identify PIHM in
 172 6 host-parasite datasets given in Crofton (1971) and 4 datasets given in Adjei *et al.* (1986)
 173 (Table 2). Crofton analyzed infection patterns in the snail *Gammarus pulex* infected with
 174 the acanthocephalan *Polmorphus minutus*. Adjei *et al.* analyzed males and females of two
 175 species of lizard fish *Saurida tumbil* and *S. undosquamis* that were infected by the cestode
 176 *Callitetrarhynchus gracilis*.

177 In both earlier studies, the authors reported PIHM in some of the datasets and we
 178 tested whether the Adjei Method and/or the Likelihood Method also predicted PIHM.
 179 For the 6 datasets from Crofton (1971), we truncated the data at 4 parasites, applied the
 180 Crofton Method to estimate the pre-mortality distribution, and then ran the Likelihood
 181 Method and Adjei Method using these pre-mortality parameters. For the Adjei *et al.*
 182 (1986) datasets, we followed the same procedure as the authors and first truncated
 183 the data at 2 parasites and then fit the Crofton Method for the female fish of both
 184 species. Then, following the original authors' methods, we parameterized the male pre-
 185 mortality distributions for each species with the results from the females. Finally, we
 186 applied the Adjei Method and the Likelihood Method to determine whether or not PIHM
 187 was significant for these species and compared our results to those given by the authors.
 188 All fitting to data was done with the code provided in *SI* 4.

189 Results

190 Question 1: Detecting presence of PIHM

191 The power of the Adjei Method to detect PIHM in a system was close to unity for larger
192 sample sizes and tended to decrease as sample size decreased (Figure 1C; *SI* 3 Figs 1-3).
193 The Likelihood Method had a power close to unity for all parameter combinations and
194 sample sizes considered. With gradual survival functions, the power of the Likelihood
195 Method decreased slightly for small samples sizes (Fig. 1C, *SI* 3 Figs 1-3).

196 The Adjei Method showed highly inflated Type I error rates (i.e. falsely detected
197 PIHM) for all parameter combinations that we considered (Fig. 1B; *SI* 3 Figs 1-3).
198 This method also showed the unintuitive pattern of Type I error rate decreasing as
199 sample size decreased. This pattern was due to the issue of binning discussed in the
200 *Introduction* and *SI* 2. For small samples sizes, the applicability of the Adjei Method is
201 compromised without binning the observed data in some way. In contrast, the Likelihood
202 Method showed a Type I error rate at or near the pre-set level of 0.05 for all parameter
203 combinations and sample sizes considered (Fig. 1B; *SI* 3 Figs 1-3).

204 Question 2: Estimating the LD_{50} and survival function

205 The Likelihood Method gave asymptotically unbiased estimates of the LD_{50} for all
206 combinations of parameters examined in this study (Fig. 2, *SI* 3 Figs 4-6). Even for
207 the smallest sample sizes we considered, the Likelihood Method's estimate of LD_{50} was
208 largely unbiased, with small biases occurring for host survival functions that were gradual.
209 The precision of the Likelihood Method's LD_{50} estimates decreased (increasing coefficient
210 of variation) as sample size decreased for all parameter combinations we examined (Fig
211 2, *SI* 3 Figs 4-6).

212 The Adjei Method produced biased estimates of the LD_{50} across nearly all
213 parameter combinations (Fig 2, *SI* 3 Figs 4-6). For $\mu_p = 10$, the LD_{50} estimates from the

214 Adjei Method were largely unbiased for large samples sizes, but as μ_p increased, the Adjei
 215 Method produced biased estimates of LD_{50} across all sample sizes, with bias increasing
 216 as sample size decreased (Figure 2, *SI* 3 Fig 4-6). The LD_{50} estimates from the Adjei
 217 Method showed large decreases in precision with the steepest survival function across all
 218 values of μ_p (Figure 2, *SI* 3 Fig 4-6).

219 In terms of the host survival function, the Likelihood Method gave unbiased
 220 estimates of survival function parameters when sample sizes were large, however as sample
 221 size decreased these estimates became severely biased (Fig. 2, *SI* Fig 7 - 9) The Adjei
 222 Method produced biased estimates of the host survival function across all sample sizes,
 223 with the bias consistently being larger when the survival function was steeper and μ_p was
 224 larger (Fig 2, *SI* 3 Figs 7-9).

225 **Detecting PIHM with unknown pre-mortality parameters**

226 When all parameters were jointly estimated, the Likelihood Method showed highly
 227 context-dependent results when detecting PIHM in the best-case scenario of $\mu_p = 10$ and
 228 $k_p = 1$. The Likelihood Method's power of detecting PIHM was greater than 0.8 when
 229 when host sample sizes were 424 and 83 for survival functions that were moderate and
 230 steep, respectively (Fig 3). When the host survival function was gradual, the Likelihood
 231 Method never had a power greater than 0.8 for any post-mortality samples sizes we
 232 considered (Fig 3).

233 **Application to real data**

234 Of the 10 datasets we considered, the previous authors qualitatively detected PIHM
 235 in 7 of them (Table 2). The Likelihood Method parameterized from the pre-mortality
 236 parameters of the Crofton Method detected significant PIHM in 6 of these 7 datasets
 237 at a significance level of 0.05. The only dataset in which the Likelihood Method did
 238 not detect a significant effect of PIHM was the Adjei dataset for female *S. tumbil*. For

239 this dataset there was a marginally significant effect of PIHM ($\chi^2_{df=2} = 5.34; p = 0.069$).
240 The Adjei Method detected PIHM in 9 of the 10 datasets (Table 2), consistent with our
241 simulation results that the Adjei Method has a high Type I error rate.

242 Discussion

243 Quantifying the impact of parasitism on wild host populations is critical for managing
244 wildlife populations and understanding parasite transmission. Ideally the relationship
245 between infection intensity and host survival would be measured experimentally, but
246 for logistical and ethical reasons, this is often impossible (McCallum 2000). Looking
247 for evidence of mortality in parasite distribution data requires the least amount of
248 information, but is notoriously difficult to implement. The methodological flaws in the
249 Adjei Method limit its utility, so here we propose an alternative, likelihood-based, method
250 to estimate host survival and the LD_{50} from observed parasite intensity data. This
251 method is a significant improvement over the previous methods because it requires fewer
252 parameters, provides a statistical decision rule for identifying PIHM and does not require
253 any data manipulation.

254 Using simulated data, we found that the Likelihood Method always out performed
255 the Adjei Method. For simply detecting the presence of PIHM, the Likelihood Method
256 was both more powerful and had fewer false detection events (Type I errors). When
257 both methods were applied to published datasets previously used in PIHM analyses,
258 the Adjei Method tended to detect PIHM where it had not previously been reported,
259 consistent with the high Type I error rate observed in our simulations. The Likelihood
260 Method was also more precise and less biased in calculations of both the parasite LD_{50}
261 and host survival curve over the parameter values we considered. However, while only the
262 Likelihood Method produced precise and unbiased LD_{50} estimates, neither method could
263 provide unbiased estimates of the host survival function at realistic sample sizes. These
264 simulations demonstrate that the Likelihood Method is more powerful and precise than

265 the previously propose Adjei Method.

266 Although superior to the Adjei Method, the Likelihood Method is not universally
267 applicable to real data. Our simulations showed when the when pre- mortality parameters
268 were estimated directly, the Likelihood Method needed at least 83-424 samples to have
269 80% power and for steep to moderate survival functions and even more as the survival
270 function became more gradual. While some of these sample sizes are reasonable for hosts
271 such as invertebrates or small fish, even the smallest sample sizes are completely unfeasible
272 for many vertebrates, particularly the species of conservation concern where addressing
273 the impact of parasitism would be most important. An even larger sample size would be
274 required to identify PIHM when parasites are highly aggregated, mean infection intensity
275 is high, or parasite prevalence is low, all of which are common in many parasitic helminths.
276 Moreover, our results are in agreement with previous work that has shown that as host-
277 survival functions become progressively more linear, PIHM becomes all but impossible
278 to detect (Lanciani & Boyett 1989). This result, however, does not preclude the use of
279 this method as non-linear survival functions are not uncommon in empirical host-parasite
280 systems (Benesh 2011). Finally, while linear functions make PIHM undetectable, at the
281 other extreme, steep, non-linear survival curves produce severely biased estimates of the
282 survival function. Give the interaction between all of these different factors, the Likelihood
283 Method is probably limited to detecting PIHM in systems where greater than 100 hosts
284 can be collected, parasites are common and only moderately aggregated, and substantial
285 host mortality occurs at relatively low parasite intensity.

286 While we have improved on the existing methods for quantifying PIHM from
287 parasite intensity data, all such methods require several fundamental, and potentially
288 problematic assumptions. Nearly all current methods derive from Crofton (1971) (but
289 see Ferguson *et al.* 2011) and assume that, prior to any PIHM, parasites are distributed
290 in the host population following a negative binomial distribution. But, it is fundamentally
291 impossible to know what the pre-mortality parasite distribution was in a wild host
292 population and it is widely recognized that different processes can lead to a variety of

293 parasite distributions in hosts (Anderson & Gordon 1982; Duerr *et al.* 2003). However, the
294 negative binomial is extremely flexible and there is substantial empirical and theoretical
295 evidence to support the assumption that, prior to any PIHM, parasite distributions can
296 be fit by a negative binomial distribution (Shaw & Dobson 1995; Shaw *et al.* 1998; Wilson
297 *et al.* 2002).

298 Unfortunately, this flexibility in the distribution may also reduce our ability to
299 detect PIHM. If a negative binomial can be fit to the observed post-mortality parasite
300 distribution then, regardless of how lethal the parasite was, it will be impossible to detect
301 PIHM because there is no need for a more complex model. Most observed parasite
302 distributions are well fit by the negative binomial distribution (Shaw *et al.* 1998),
303 suggesting that systems where these methods are applicable may be more the exception
304 than the rule. Furthermore, even when truncation of the negative binomial distribution is
305 detected, it may be caused by other processes such as within host density dependence, age
306 dependent variation in host resistance and/or heterogeneous infection rates (McCallum
307 2000; Anderson & Gordon 1982; Rousset *et al.* 1996). This means that in the event that
308 PIHM is detected, it may actually not be the result of PIHM.

309 Given these numerous caveats, is there a place in parasitology for methods that
310 estimate PIHM from intensity data? We are in agreement with Lester (1984) that, at the
311 very least, methods for estimating PIHM can provide preliminary insight into whether
312 or not PIHM is worth further exploration. However, we stress that these methods should
313 only be used as an exploratory tool when assessing the role of PIHM in a system and
314 potential users should critically evaluate whether they think they have a large enough
315 sample size and an appropriate host survival function/post-mortality distribution for the
316 methods developed in this paper to be applicable. Even if they are applicable, inferring
317 PIHM from distributional data is no substitute for field experiments and an in depth
318 understanding of the natural history of the host-parasite system under consideration.

319 Acknowledgments

320 TODO

321 References

322 1.

323 Adjei, E. L., Barnes, A. & Lester, R. J. G. (1986). A method for estimating possible
324 parasite-related host mortality, illustrated using data from *Callitetrarhynchus gracilis*
325 (Cestoda: Trypanorhyncha) in lizardfish (*Saurida* spp.). *Parasitology*, 92, 227–243.

326 2.

327 Anderson, R. M. & Gordon, D. M. (1982). Processes influencing the distribution of
328 parasite numbers within host populations with special emphasis on parasite-induced host
329 mortalities. *Parasitology*, 85, 373–398.

330 3.

331 Anderson, R. M. & May, R. M. (1979). Population biology of infectious diseases: Part I.
332 *Nature*, 280, 361 – 367.

333 4.

334 Benesh, D. P. (2011). Intensity-dependent host mortality: what can it tell us about
335 larval growth strategies in complex life cycle helminths? *Parasitology*, 138, 913–25. URL
336 <http://www.ncbi.nlm.nih.gov/pubmed/21554844>.

337 5.

338 Brooker, S., Bethony, J. & Hotez, P. J. (2004). Human hookworm hnfection in the 21st
339 century. *Advances in Parasitology*, 58, 197–288.

340 6.

341 Calabrese, J. M., Brunner, J. L. & Ostfeld, R. S. (2011). Partitioning the aggregation of
342 parasites on hosts into intrinsic and extrinsic components via an extended Poisson-gamma
343 mixture model. *PloS one*, 6, e29215.

344 7.

345 Cox, D. R., Donnelly, C. a., Bourne, F. J., Gettinby, G., McInerney, J. P., Morrison, W. I.
346 & Woodroffe, R. (2005). Simple model for tuberculosis in cattle and badgers. *Proceedings*
347 *of the National Academy of Sciences of the United States of America*, 102, 17588–17593.

348 8.

349 Coyne, M. J., Smith, G. & McAllister, Fiona, E. (1989). Mathematic model for the
350 population biology of rabies in raccoons in the mid-Atlantic states. *The American Journal*
351 *of Veterinary Research*, 50, 2148–2154.

352 9.

353 Crofton, H. D. (1971). A quantitative approach to parasitism. *Parasitology*, 62, 179–193.

- 354 10.
355 De Castro, F. & Bolker, B. (2005). Mechanisms of disease-induced extinction. *Ecology Letters*, 8, 117–126. URL <http://doi.wiley.com/10.1111/j.1461-0248.2004.00693.x>.
356
- 357 11.
358 Dobson, A. P. & Hudson, P. J. (1992). Regulation and stability of a free-living host-
359 parasite system: *Trichostrongylus tenuis* in red grouse. II. Population models. *Journal of*
360 *Animal Ecology*, 61, 487–498.
- 361 12.
362 Duerr, H. P., Dietz, K. & Eichner, M. (2003). On the interpretation of age-intensity
363 profiles and dispersion patterns in parasitological surveys. *Parasitology*, 126, 87–101.
- 364 13.
365 Ferguson, J. a., Koketsu, W., Ninomiya, I., Rossignol, P. a., Jacobson, K. C. & Kent,
366 M. L. (2011). Mortality of coho salmon (*Oncorhynchus kisutch*) associated with burdens
367 of multiple parasite species. *International journal for parasitology*, 41, 1197–205. URL
368 <http://www.ncbi.nlm.nih.gov/pubmed/21855547>.
- 369 14.
370 Joly, D. O. & Messier, F. (2004). The distribution of *Echinococcus granulosus* in moose:
371 Evidence for parasite-induced vulnerability to predation by wolves? *Oecologia*, 140, 586–
372 590.
- 373 15.
374 Kirk, R. S. (2003). The impact of *Anguillicola crassus* on European eels. *Fisheries*
375 *Management and Ecology*, 10, 385–394.
- 376 16.
377 Lafferty, K. D. & Kuris, A. M. (2002). Trophic strategies, animal diversity and body size.
378 *Trends in Ecology and Evolution*, 17, 507–513.
- 379 17.
380 Lanciani, C. A. & Boyett, J. M. (1989). Demonstrating parasitic water mite-induced
381 mortality in natural host populations. *Parasitology*, 81, 465–475.
- 382 18.
383 Langwig, K. E., Voyles, J., Wilber, M. Q., Frick, W. F., Murray, K. a., Bolker, B. M.,
384 Collins, J. P., Cheng, T. L., Fisher, M. C., Hoyt, J. R., Lindner, D. L., McCallum,
385 H. I., Puschendorf, R., Rosenblum, E. B., Toothman, M., Willis, C. K., Briggs, C. J.
386 & Kilpatrick, a. M. (2015). Context-dependent conservation responses to emerging
387 wildlife diseases. *Frontiers in Ecology and the Environment*, 13, 195–202. URL
388 <http://www.esajournals.org/doi/10.1890/140241>.
- 389 19.
390 Lester, R. J. G. (1977). An estimate of mortality in a population of *Perca flavescens* owing
391 to the trematode *Diplostomum adamsi*. *Canadian Journal of Zoology*, 55, 288–292.

392 20.

393 Lester, R. J. G. (1984). A review of methods for estimating mortality due to
394 parasites in wild fish populations. *Helgoländer Meeresuntersuchungen*, 37, 53–64. URL
395 <http://link.springer.com/10.1007/BF01989295>.

396 21.

397 Logiudice, K. (2003). Trophically Transmitted Parasites and the Conservation of Small
398 Populations: Raccoon Roundworm and the Imperiled Allegheny Woodrat \rParásitos
399 Transmitidos Tróficamente y la Conservación de Poblaciones Pequeñas: el Ascárido
400 de los Mapaches y la Rata de la . *Conservation Biology*, 17, 258–266. URL
401 <http://dx.doi.org/10.1046/j.1523-1739.2003.01293.x>.

402 22.

403 McCallum, H. (2012). Disease and the dynamics of extinction. *Philosophical transactions*
404 *of the Royal Society of London. Series B, Biological sciences*, 367, 2828–39. URL
405 <http://www.ncbi.nlm.nih.gov/pubmed/22966138>.

406 23.

407 McCallum, H. I. (2000). Host-pathogen and host-parasite models. In: *Population*
408 *Parameters: Estimation for Ecological Models* (eds. Lawton, J. H. & Likens, G. E.), chap.
409 Chapter 10. Blackwell Science Ltd., pp. 284–312.

410 24.

411 Roeber, F., Jex, A. R. & Gasser, R. B. (2013). Impact of gastrointestinal parasitic
412 nematodes of sheep, and the role of advanced molecular tools for exploring epidemiology
413 and drug resistance - an Australian perspective. *Parasites & vectors*, 6, 153. URL
414 <http://www.pubmedcentral.nih.gov/articlerender.fcgi?artid=3679956&tool=pmcentrez&rend>

415 25.

416 Rousset, F., Thomas, F., Meeûs, T. D. & Renaud, F. (1996). Inference of parasite-induced
417 host mortality from distributions of parasite loads. *Ecology*, 77, 2203–2211.

418 26.

419 Royce, L. A. & Rossignol, P. (1990). Epidemiology of honey bee parasites. *Parasitology*
420 *Today*, 6, 348–353.

421 27.

422 Shaw, D. J. & Dobson, A. P. (1995). Patterns of macroparasite abundance and
423 aggregation in wildlife populations: a quantitative review. *Parasitology*, 111 Suppl, S111–
424 27.

425 28.

426 Shaw, D. J., Grenfell, B. T. & Dobson, a. P. (1998). Patterns of macroparasite aggregation
427 in wildlife host populations. *Parasitology*, 117 (Pt 6, 597–610.

428 29.

429 Tillé, a., Lefèvre, C., Pastoret, P. P. & Thiry, E. (1991). A mathematical model of
430 rinderpest infection in cattle populations. *Epidemiology and infection*, 107, 441–452.

431 30.

432 Walther, B. a. & Moore, J. L. (2005). The concepts of bias, precision and accuracy,
433 and their use in testing the performance of species richness estimators, with a literature
434 review of estimator performance. *Ecography*, 28, 815–829.

435 31.

436 Wilson, K., Bjørnstad, O. N., Dobson, A. P., Merler, S., Pogliayen, G., Read, A. F. &
437 Skorpning, A. (2002). Heterogeneities in macroparasite infections: patterns and processes.
438 In: *The Ecology of Wildlife Diseases* (eds. Hudson, P. J., Rizzoli, A., Grenfell, B.,
439 Heesterbeek, H. & Dobson, A.), chap. 2. Oxford University Press, Oxford, pp. 6–44.

Table 1: Definition of parameters and functions used in the main text

Parameter	Definition
μ_p	Pre-mortality mean parasite intensity
k_p	Pre-mortality parasite aggregation
N_p	Pre-mortality host population size
x	Number of parasites in a given host
$g(x; \mu_p, k_p)$	Pre-mortality negative binomial parasite distribution
a	Parameter of the logistic host survival function
b	Parameter of the logistic host survival function
$h(x; a, b)$	The host survival function
LD_{50}	$\exp(a/b)$, parasite intensity at which 50% of hosts die

Table 2: Comparison of the PIHM predictions of previously used host-parasite datasets to those given by the Adjei Method and the Likelihood Method. The first column specifies the identity of the dataset, the second column specifies whether or not the authors indicated that PIHM was occurring in the system based on a qualitative assessment, the third column indicates whether or not the Likelihood Method with pre-mortality parameters estimated from the Crofton Method detects significant PIHM, and the final column indicates whether the Adjei Method with pre-mortality parameters estimated from the Crofton Method detects PIHM. If a method detected significant PIHM the predicted LD_{50} is given in parentheses.

Data Set (sample size)	Author detected PIHM?	Likelihood Method?	Adjei Method?
Crofton, Station 1 ($n = 538$)	Yes	Yes (7.27)	Yes (9.33)
Crofton, Station 2 ($n = 507$)	Yes	Yes (6.92)	Yes (14.95)
Crofton, Station 3 ($n = 633$)	Yes	Yes (5.93)	Yes (5.98)
Crofton, Station 4 ($n = 486$)	No	No	Yes (7.99)
Crofton, Station 5 ($n = 276$)	No	No	Yes (10.58)
Crofton, Station 6 ($n = 191$)	No	No	No
Adjei, <i>S. tumbil</i> female ($n = 446$)	Yes (5.7)	No	Yes (6.37)
Adjei, <i>S. tumbil</i> male ($n = 452$)	Yes (3.4)	Yes (3.42)	Yes (3.66)
Adjei, <i>S. undosquamis</i> female ($n = 2573$)	Yes (3.2)	Yes (3.04)	Yes (3.11)
Adjei, <i>S. undosquamis</i> male ($n = 2440$)	Yes (1.8)	Yes (1.83)	Yes (1.78)

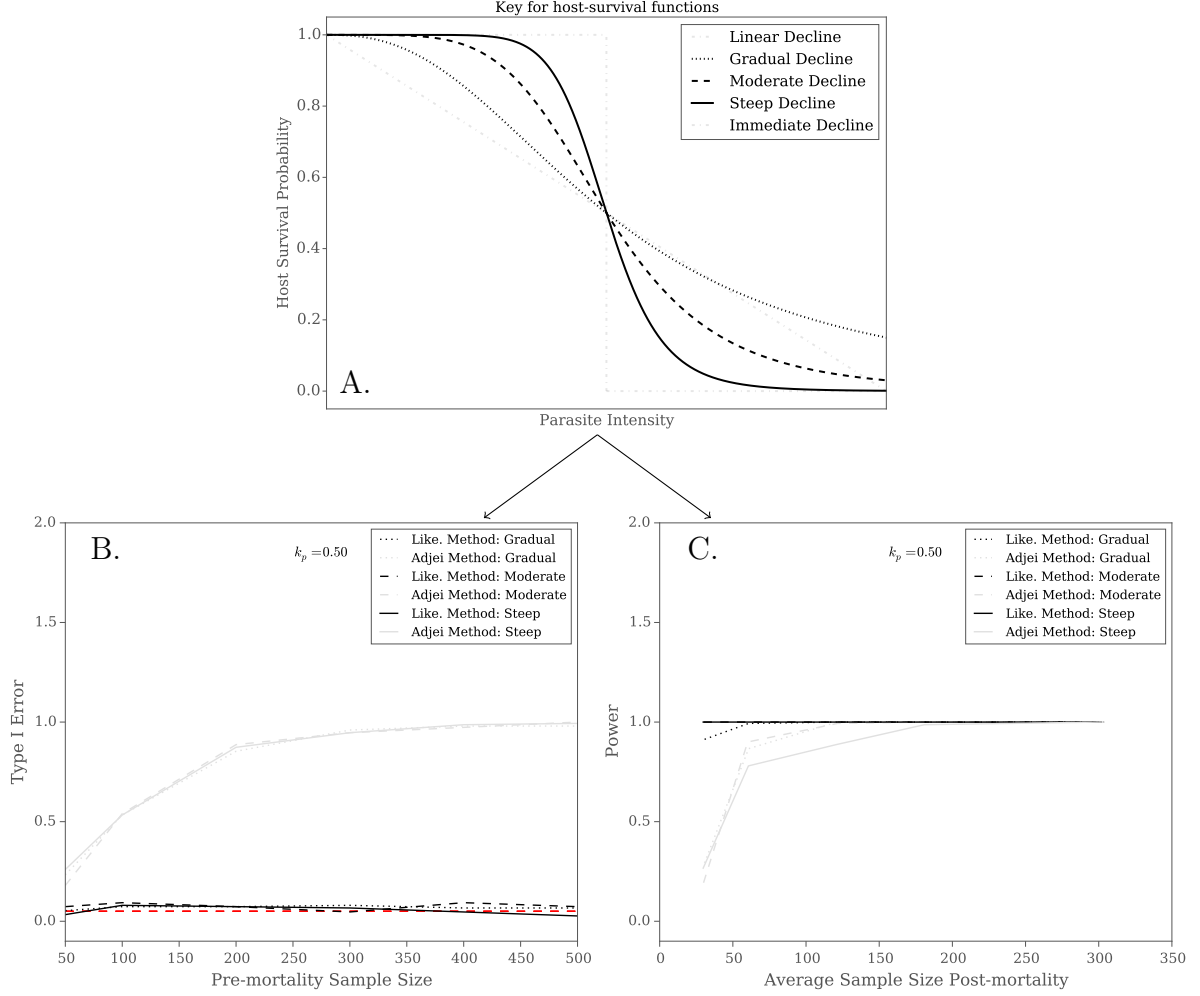


Figure 1: A) Five potential shapes for a host-survival functions. In the simulations we used a gradual survival function (dotted line), and moderate survival function (dashed line), and a steep survival function (solid line). The linear and immediate survival functions represent two potential extremes that we do not include in the simulations. For each of these survival functions and the parameter combinations described in the main text, we tested the Type I error and power of the Likelihood Method and Adjei Method. B) Gives the Type I error of each method over a range of pre-mortality sample sizes with a pre-mortality mean parasite intensity (μ_p) of 50 and pre-mortality parasite aggregation (k_p) at 0.5. The red line shows the pre-set significance level of 0.05. C) Gives the Power of each method for detecting PIHM over a range of post-mortality sample sizes for $\mu_p = 50$ and $k_p = 0.5$. In general, the Likelihood Method has higher power and lower Type I error than the Adjei Method. See the *SI 3* Fig 1 - 3 for Type I Error and power results for all parameter combinations.

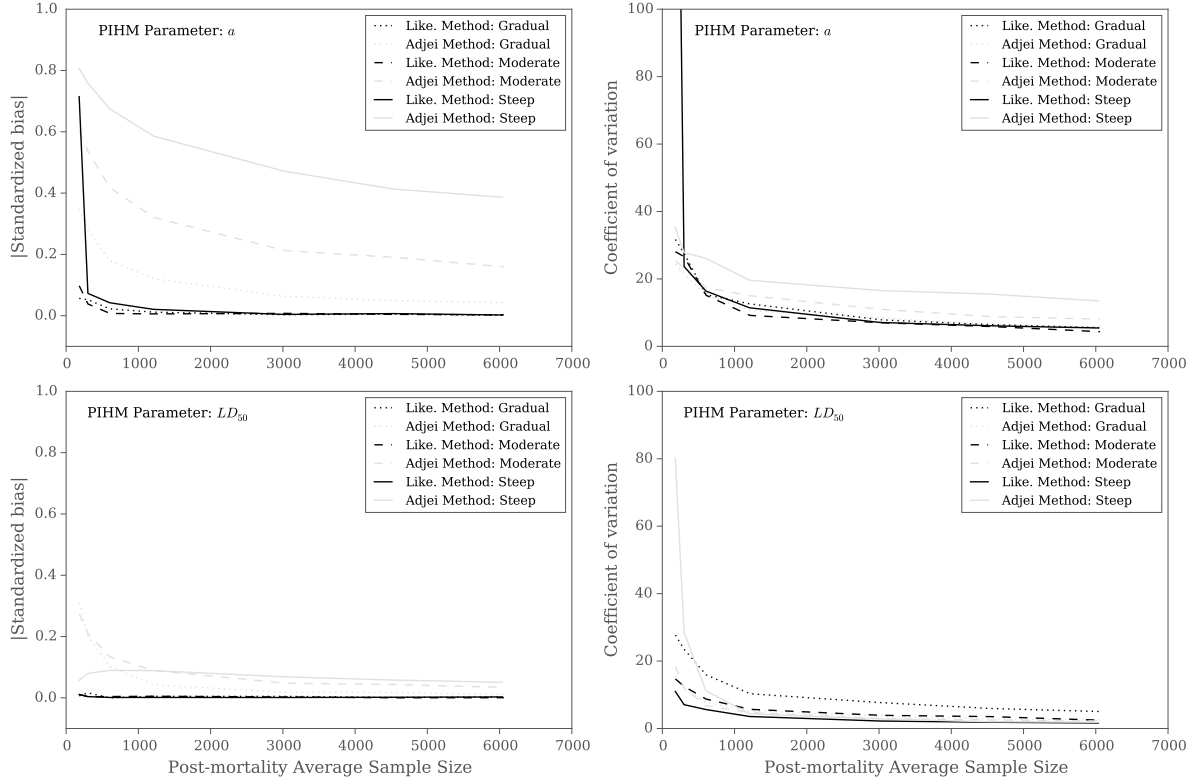


Figure 2: Bias and precision (coefficient of variation) for the Likelihood Method and Adjei Method estimates of the a parameter and the LD_{50} of the host survival function based on simulated PIHM data over a range of post-mortality sample sizes. As the coefficient of variation increases, precision decreases. The pre-mortality parameters for this simulation were $\mu_p = 50$ and $k_p = 0.5$. The figure shows the simulations for three different host survival functions (gradual, moderate, and steep), each with the same LD_{50} . Bias and precision results of LD_{50} and a for all other parameter combinations can be found in SI 3 Fig 4 - 9.

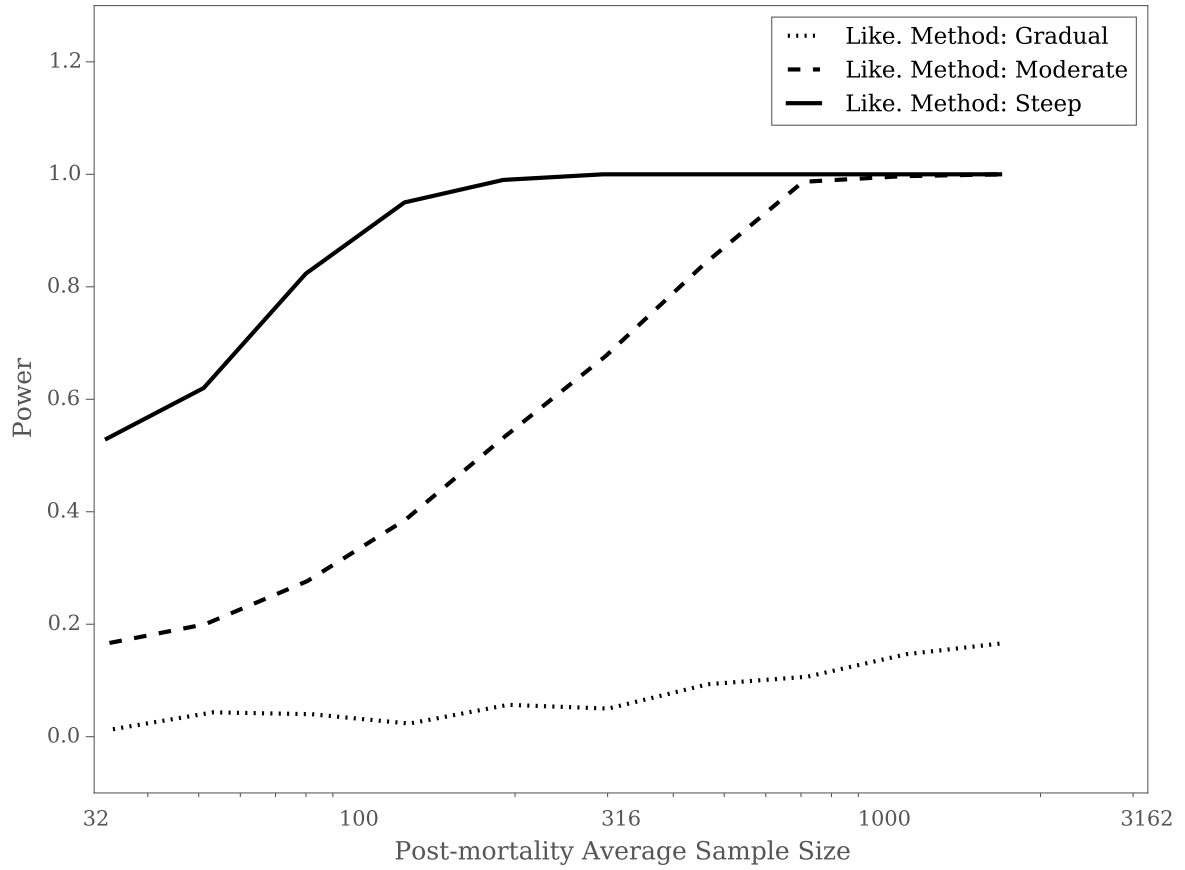


Figure 3: The power of the Likelihood Method to detect PIHM for gradual, moderate, and steep survival functions when all four parameters μ_p , k_p , a , and b were jointly estimated. The curves were generated from 500 simulations for 10 pre-mortality population sizes, N_p .