

CH 4 Continuous Variables and Their Probability Distributions

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0.1 Intro

Continuous Random Variable: a RV that can take on any value in an interval.

0.2 The Probability Distribution for a CRV

0.2.1 Definition of CDF (distribution function)

let Y denote any RV. The DF of Y denoted by $F(y)$, is such that $F(y) = P(Y \leq y)$ for $-\infty < y < \infty$

CDF for DRV are always step functions because the CDF increases only at the finite or countable num of points w/ positive probabilities

0.2.2 Properties of a Distribution Function(CDF)

If $F(y)$ is a DF, then

1. $F(-\infty) = \lim_{y \rightarrow -\infty} F(y) = 0$
2. $F(\infty) = \lim_{y \rightarrow \infty} F(y) = 1$
3. $F(y)$ is a nondecreasing fcn of y . ($y_1 < y_2, F(y_1) < F(y_2)$)

0.2.3 Definition of a Continuous Random Variable

a RV Y w/ DF $F(y)$ is said to be continuous if $F(y)$ is continuous for $-\infty < y < \infty$

0.2.4 Derivative of CDF is PDF

let $F(Y)$ be the DF for a CRV Y . Then $f(y)$, given by

$$f(y) = \frac{dF(y)}{dy} = F'(y)$$

whenever the derivative exists, is called the probability density fcn for the RV Y .

$$F(y) = \int_{-\infty}^y f(t)dt$$

where $f(\cdot)$ is the PDF

PDF is a theoretical model

0.2.5 Properties of a Density Function (PDF)

if $f(y)$ is a PDF for a CRV then

1. $f(y) \geq 0$ for all $y, -\infty < y < \infty$

2. $\int_{-\infty}^{\infty} f(y)dy = 1$

$F(y_0)$ gives the P that $Y \leq y_0$, it's of interest to determine y of RV Y such that $P(Y \leq y) \geq$ some value

0.2.6 Definition of Quantile, Percentile HELP

Let Y denote any RV. If $0 < p < 1$, the p th quantile of Y denoted by ϕ_p is the smallest val such that $P(Y \leq \pi_q) = F(\pi_p) \geq p$. If Y is cont., π_p is the smallest val such that $F(\phi_p) = P(Y \leq \phi_p) = p$

ϕ_p is referred to the 100pth percentile of Y

0.2.7 PDF Interval Theorem

If the RV Y has a DF $f(y)$ and $a < b$ then the P that Y falls in the interval $[a, b]$ is

$$P(a \leq Y \leq b) = \int_a^b f(y)dy$$

NOT TRUE FOR DRV

0.3 Expected Values for CRV

0.3.1 Definition of EV of CRV

EV of CRV Y is

$$E(Y) = \int_{-\infty}^{\infty} yf(y)dy$$

exists if converges

0.3.2 EV of Function of Y Theorem

let $g(Y)$ be a fxn of Y ; then the EV of $g(Y)$ is

$$E[g(Y)] = \int_{-\infty}^{\infty} g(y)f(y)dy$$

0.3.3 EV of $g(y)$ properties

let g_1, \dots, g_n be fxns of CRV Y : 1. $E(c) = c$

2. $E[cg(Y)] = cE[g(Y)]$

3. $E[g_1 + g_2(Y)] = E[g_1(Y)] + E[g_2(Y)]$

some bonuses:

$$g(Y) = (Y - \mu)^2$$

$$V(Y) = E(Y - \mu)^2$$

$$V(Y) = E(Y^2) - \mu^2$$

0.4 The Uniform Probability Distribution

0.4.1 Definition of Uniform PD

If $\theta_1 < \theta_2$ a RV Y is said to have a cont. UPD on the interval (θ_1, θ_2) IFF the density fxn of Y is

$$f(y) = \begin{cases} \frac{1}{\theta_2 - \theta_1}, & \theta_1 \leq y \leq \theta_2 \\ 0, & \text{elsewhere} \end{cases}$$

0.4.2 Parameters of Density Fxn

the constants that determine form of a density fxn are called parameters of the density fxn.

0.4.3 Mean and Variance of UPD

If $\theta_1 < \theta_2$ and Y is a RV uniformly distributed on the interval θ_1, θ_2 , then

$$\mu = E(Y) = \frac{\theta_1 + \theta_2}{2} \text{ and } \sigma^2 = V(Y) = \frac{(\theta_2 - \theta_1)^2}{12}$$

0.5 The Normal Probability Distribution

The most widely used cont PD is the normal distribution

0.5.1 Definition of Normal Probability Distribution

a RV Y is said to have a NPD IFF for $\sigma > 0$ and $-\infty < \mu < \infty$ the density fcn of Y is:

$$f(y) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y-\mu)^2}{2\sigma^2}}$$

for $-\infty < y < \infty$

area is the integral of this (which doesn't exist)

0.5.2 Mean and Variance of NPD

if Y is normally distributed RV w/ param μ, σ then

$$E(Y) = \mu \text{ and } V(Y) = \sigma^2$$

NPDP is symmetric around mean

0.5.3 Z and z score

z is the distance from the mean of a normal distribution expressed in units of standard deviation

$$z = \frac{y - \mu}{\sigma}$$

Can transform a normal RV Y to a standard normal RV Z by using

$$Z = \frac{Y - \mu}{\sigma}$$

SEE TABLE 4!!!

0.6 The Gamma Probability Distribution

some PDF are skewed right where most values are near origin

0.6.1 Definition of GPD

a RV Y is said to have a gamma distribution w/ param $\alpha > 0$ and $\beta > 0$ IFF the density fcn of Y is:

$$f(y) = \begin{cases} \frac{y^{\alpha-1} e^{-\frac{y}{\beta}}}{\beta^\alpha \Gamma(\alpha)} & 0 \leq y < \infty, \\ 0, & \text{elsewhere,} \end{cases}$$

where the gamma function is

$$\Gamma(\alpha) = \int_0^\infty y^{\alpha-1} e^{-y} dy$$

α is called the shape parameter associated w/ a gamma distribution.

β is called the scale parameter

Special case: $\alpha = 1$ can be expressed as a sum of certain Poisson probabilities

0.6.2 Mean and Varaince of GPD HELP magic

if Y has a gamma distribution w/ α, β then

$$\mu = E(Y) = \alpha\beta \text{ and } \sigma^2 = V(Y) = \alpha\beta^2$$

0.6.3 Definition of chi-square distribution

Let v be a +int. A RV Y is said to have a chi-square distribution with v degrees of freedom IFF Y is a gamma distributed RV w/ parameteres $\alpha = \frac{v}{2}, \beta = 2$

A RV w/ chi-square distribution is a χ^2 RV.

0.6.4 Mean and Variance of chi-square distribution

if Y is a χ^2 RV w/ v DoF, then:

$$\mu = E(Y) = v \text{ and } \sigma^2 = V(Y) = 2v$$

SEE TABLE 6

gamma density function in which $\alpha = 1$ si called the exponential density function

0.6.5 Definition Exponential Distribution

a RV Y is said to have an exponential distribution w/ parameter $\beta > 0$ IFF the density fcn of Y is:

$$f(y) = \begin{cases} \frac{1}{\beta} e^{-y/\beta}, & 0 \leq y < \infty \\ 0, & \text{elsewhere.} \end{cases}$$

useful for modeling length of life of electronic components

0.6.6 Mean and Variance of Exponential Distribution

If Y is an exponential RV w/ param β then:

$$\mu = E(Y) = \beta \text{ and } \sigma^2 = V(Y) = \beta^2$$

see example 4.10:

exponential distribution has a property called memoryless property (geometric distribution also has)

0.7 The Beta Probability Distribution

the beta density fcn is a two param density fcn defined over $0 \leq y \leq 1$. Often used as a model for proportions

0.7.1 Definition of Beta Probability Distribution

a RV Y has a BPD w/ param $\alpha > 0, \beta > 0$ IFF density fcn of Y :

$$f(y) = \begin{cases} \frac{y^{\alpha-1}(1-y)^{\beta-1}}{B(\alpha, \beta)}. & 0 \leq y \leq 1 \\ 0, & \text{elsewhere,} \end{cases}$$

where the beta function is

$$B(\alpha, \beta) = \int_0^1 y^{\alpha-1}(1-y)^{\beta-1} dy = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$$

if $c \leq y \leq d$, then $y^* = \frac{y-c}{d-c}$ defines a new var such that $0 \leq y^* \leq 1$, thus BDF can be applied to a RV defined on int $c \leq y \leq d$ by translation

CDF for beta RV is the incomplete beta function:

$$F(y) = \int_0^y \frac{t^{\alpha-1}(1-t)^{\beta-1}}{B(\alpha, \beta)} dt = I_y(\alpha, \beta)$$

see Tables of the Incomplete Beta Function
 related to binomial function: when integration by parts

$$F(y) = \sum_{i=\alpha}^n \binom{n}{i} y^i (1-y)^{n-i}$$

where $n = \alpha + \beta - 1$

0.7.2 Mean and Varaince of Beta Probability Distribution

if Y is a beta-distributed RV w/ param $\alpha > 0, \beta$

$$\mu = E(Y) = \frac{\alpha}{\alpha + \beta} \text{ and } \sigma^2 = V(Y) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$$

0.8 General Comments

a good model is one that yields good inferences about the population of interest.
 Selecting a reasonable models is a matter of acting on theoretical considerations.
 EX: Poisson RV is appropriate when the it is characterised by random behavior of events in time.

Another way to form model is from a frequency histogram and choose a density fcn similar to the frequency curve.

Can also calculate goodness of fit.

see Weibull distribution

0.9 Other Expected Values

Moments for CRV have definitions analogous to DRV

0.9.1 Definition Moments CRV

If Y is a CRV, then the kth moment about the origin is:

$$\mu'_k = E(Y^k)$$

k1toinf.

The kth moment about the mean, or the kth central moment:

$$\mu_k = E[(Y - \mu)^k],$$

k1toinf

0.9.2 Definition of Moment Generating Function CRV

if Y is CRV, then the MGF of Y is:

$$m(t) = E(e^{tY})$$

MGF exists if there exists a constant $b > 0$ such that $m(t)$ is finite for $|t| \leq b$
binomial expansion can be used to solve

0.9.3 MGF for $g(Y)$

let Y be a RV with density fcn $f(y)$ and $g(Y)$ be a fcn of Y . then MGF for $g(Y)$

$$E[e^{tg(Y)}] = \int_{-\infty}^{\infty} e^{tg(y)} f(y) dy$$

0.10 Tchebysheff's Theorem CRV

0.10.1 Theorem

let Y be a RC w/ finite mean μ and variance σ^2 then for any $k > 0$

$$P(|Y - \mu| < k\sigma) \geq 1 - \frac{1}{k^2} \text{ or } P(|Y - \mu| \leq k\sigma) \leq \frac{1}{k^2}$$

Allows to find bounds for P easier

Can find mean and variances of RV without knowledge of distribution

0.11 Expectations of Discont Fcn and Mixed PD (HELP R2)

a RV t that has some P at discrete points and the remainder spread over intervals is a mixed distribution. EX: insurance

0.11.1 Definition of a Mixed Distribution Function

$$F(y) = c_1 F_1(y) + c_2 F_2(y)$$

where F_1 is a step distribution fcn and F_2 is a continuous distribution fcn

suppose that X_1 is a DRV w/ dist fcn $F_1(y)$

suppose that X_2 is a CRV w/ dist fcn $F_2(y)$

let $g(Y)$ denote a fcn of :

$$E[g(Y)] = c_1 E[g(X_1)] + c_2 E[g(X_2)]$$

0.12 Summary

PDF, CDF, Normal, Exponential, Gamma, χ^2 , Beta
Moments, MGF, Tchebysheff

IMPORTANT:

first moment is mean

second moment is variance

third moment is skewness (about mean)

fourth moment about mean is kurtosis