

ĐẠI HỌC BÁCH KHOA HÀ NỘI VIỆN CÔNG NGHỆ THÔNG TIN VÀ TRUYỀN THÔNG



PRESENTATION TITLE

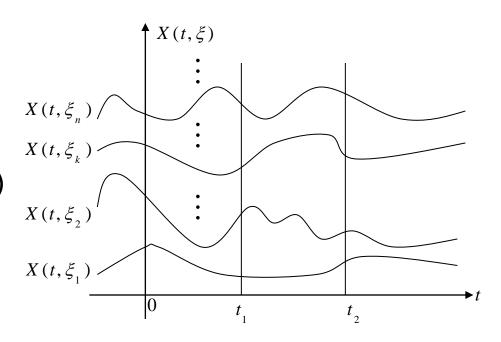
Presentation Subtitle

- Definitions
- Stationary processes
- Spectral density function
- Ergodicity of stochastic processes



Stochastic processes

- Let ξ denote the random outcome of an experiment. To every such outcome suppose a waveform X(t, ξ) is assigned.
- The collection of such waveforms form a stochastic process.



- The set of $\{\xi_k\}$ and the time index t can be continuous or discrete (countably infinite or finite) as well.
- For fixed $\xi_i \in S$ (the set of all experimental outcomes), $X(t, \xi_i)$ is a specific time function.
- For fixed t, $X_1=X(t_1, \xi_i)$ is a random variable. The ensemble of all such realizations $X(t, \xi)$ over time represents the stochastic process X(t).



- Example
 - $X(t) = A\cos(\omega_0 t + \varphi)$
 - Where φ is a uniformly distributed random variable in $(0, 2\pi)$ represents a stochastic process.
- Some stochastic processes
 - Brownian motion,
 - Stock market fluctuations,
 - Various queuing systems.



- If X(t) is a stochastic process, then for fixed t, X(t)represents a random variable.
- Its distribution function is given by

$$F_{X}(x,t) = P\{X(t) \le x\}$$

• Notice that $F_x(x, t)$ depends on t, since for a different t, we obtain a different random variable.

$$f_{x}(x,t) \triangleq \frac{dF_{x}(x,t)}{dt}$$

 $f_{_X}(x,t) \triangleq \frac{dF_{_X}(x,t)}{dt}$ • Derivative of $F_{_X}(x,t)$ represents the first-order probability density function of the process X(t).



- For $t = t_1$ and $t = t_2$,
 - X(t) represents two different random variables $X_1 = X(t_1)$ and $X_2 = X(t_2)$ respectively.
 - Their joint distribution is given by

$$F_{X}(x_{1}, x_{2}, t_{1}, t_{2}) = P\{X(t_{1}) \le x_{1}, X(t_{2}) \le x_{2}\}$$

Their joint density function is:

$$f_{X}(x_1, x_2, t_1, t_2) \triangleq \frac{\partial^2 F_{X}(x_1, x_2, t_1, t_2)}{\partial x_1 \partial x_2}$$

 And represents the second-order density function of the process X(t).



- Similarly $f_x(x_1, ..., x_n, t_1, ..., t_n)$ represents the nth order density function of the process X(t).
- Complete specification of the stochastic process X(t) requires the knowledge of $f_x(x_1, ..., x_n, t_1, ..., t_n)$ for all t_i , i = 1, ..., n and for all n. (an almost impossible task in reality).

- Characterisctics of stochastic processes
 - Mean of a stochastic process

$$\mu(t) \stackrel{\Delta}{=} E\{X(t)\} = \int_{-\infty}^{+\infty} x f_X(x,t) dx$$

- $\mu(t)$ represents the mean value of a process X(t). In general, the mean of a process can depend on the time index t.
- Autocorrelation function of a process X(t) is defined as

$$R_{XX}(t_1,t_2) \stackrel{\Delta}{=} E\{X(t_1)X^*(t_2)\} = \iint x_1x_2^* f_X(x_1,x_2,t_1,t_2) dx_1 dx_2$$
 It represents the interrelationship between the random variables X1 = $X(t1)$ and $X2 = X(t2)$ generated from the process $X(t)$.

Properties of autocorrelation function

1.
$$R_{XX}(t_1, t_2) = R_{XX}^*(t_2, t_1) = [E\{X(t_2)X^*(t_1)\}]^*$$

2.
$$R_{XX}(t,t) = E\{|X(t)|^2\} \ge 0.$$

(Average instantaneous power)

 3. R_{xx}(t₁, t₂) represents a nonnegative definite function, i.e., for any set of constants {a_i}ⁿ_{i=1}

$$\sum_{i=1}^{n} \sum_{j=1}^{n} a_i a_j^* R_{XX}(t_i, t_j) \ge 0.$$

This equation follows by noticing that: E{|Y|²}≥0 and

$$Y = \sum_{i=1}^{n} a_i X(t_i).$$



• The **autocovariance** function of the process X(t).

$$C_{XX}(t_1, t_2) = R_{XX}(t_1, t_2) - \mu_X(t_1)\mu_X^*(t_2)$$

- Examples
 - Given

$$z = \int_{-T}^{T} X(t) dt.$$

$$E[|z|^{2}] = \int_{-T}^{T} \int_{-T}^{T} E\{X(t_{1})X^{*}(t_{2})\}dt_{1}dt_{2}$$
$$= \int_{-T}^{T} \int_{-T}^{T} R_{xx}(t_{1}, t_{2})dt_{1}dt_{2}$$



Consider process X(t)

$$X(t) = a\cos(\omega_{0}t + \varphi), \quad \varphi \sim U(0,2\pi).$$

$$\mu_{x}(t) = E\{X(t)\} = aE\{\cos(\omega_{0}t + \varphi)\}$$

$$= a\cos\omega_{0}t E\{\cos\varphi\} - a\sin\omega_{0}t E\{\sin\varphi\} = 0,$$
since
$$E\{\cos\varphi\} = \frac{1}{2\pi} \int_{0}^{2\pi} \cos\varphi d\varphi = 0 = E\{\sin\varphi\}.$$

$$R_{xx}(t_{1},t_{2}) = a^{2}E\{\cos(\omega_{0}t_{1} + \varphi)\cos(\omega_{0}t_{2} + \varphi)\}$$

$$= \frac{a^{2}}{2}E\{\cos\omega_{0}(t_{1} - t_{2}) + \cos(\omega_{0}(t_{1} + t_{2}) + 2\varphi)\}$$

$$= \frac{a^{2}}{2}\cos\omega_{0}(t_{1} - t_{2}).$$



- Strict sense and wide sense stationarity
 - Stationarity
 - Stationary processes exhibit statistical properties that are invariant to shift in the time index.
 - Thus, for example, second-order stationarity implies that the statistical properties of the pairs $\{X(t_1), X(t_2)\}$ and $\{X(t_1+c), X(t_2+c)\}$ are the same for any c.
 - Similarly first-order stationarity implies that the statistical properties of $X(t_i)$ and $X(t_i+c)$ are the same for any c.



- Strict sense stationarity
 - In strict terms, the statistical properties of a stochastic process are governed by the joint probability density function.
 - A process is nth-order Strict-Sense Stationary (S.S.S) if
- for any c, where the left side represents the joint density function $f_{x}(x_{1},x_{0})$; the rankofm variables $f_{x}(x_{1},x_{0})$; $f_{x}(x_{1},x_{0})$ and the right side corresponds to the joint density function of the random variables $X'_{1}=X(t_{1}+c), ..., X'_{n}=X(t_{n}+c)$.
 - A process X(t) is said to be **strict-sense stationary** if equation above is true for all t_i , i=1, ..., n; n=1, 2, ... and any c.



- First order strict sense stationary process
 - For any c

$$f_{X}(x,t) \equiv f_{X}(x,t+c)$$

• In particular, if c = -t, then

 $f_{X}(x,t) = f_{X}(x)$ • That means, the first-order density of X(t) is independent of t. In that case

$$E[X(t)] = \int_{-\infty}^{+\infty} x f(x) dx = \mu, \quad a \ constant.$$
 (1)



- Second order strict sense stochastic process
 - From definition, we have:

$$f_{X}(x_{1}, x_{2}, t_{1}, t_{2}) \equiv f_{X}(x_{1}, x_{2}, t_{1} + c, t_{2} + c)$$

• for any c. If c is choosen so as $c = -t_2$, we get

 $f_{X}(x_1,x_2,\ t_1,t_2)\equiv f_{X}(x_1,x_2,\ t_1-t_2)$ • The second order density function of a strict sense

 The second order density function of a strict sense stationary process depends only on the difference of the time indices t₁-t₂=τ.



The autocorrelation function is given by

$$R_{xx}(t_1,t_2) \triangleq E\{X(t_1)X^*(t_2)\}$$

$$= \iint x_1 x_2^* f_x(x_1,x_2,\tau=t_1-t_2) dx_1 dx_2$$

$$= R_{xx}(t_1-t_2) \triangleq R_{xx}(\tau) = R_{xx}^*(-\tau), \tag{2}$$
 • The autocorrelation function of a second order strict-

 The autocorrelation function of a second order strictsense stationary process depends only on the difference of the time indices t₂-t₁=τ



- Wide sense stationarity
 - A process X(t) is said to be Wide-Sense Stationary if:

• $E\{X(t)\}=\mathcal{D}^d$ $E\{X(t_1)X^*(t_2)\}=R_{XX}(t_1-t_2),$ • Since these equations follow from (1) and (2), strict-sense

- Since thése équations follow from (1) and (2), strict-sensé stationarity always implies wide-sense stationarity.
- In general, the converse is not true
 - Exception: the Gaussian process (normal process).
 - This follows, since if X(t) is a Gaussian process, then by definition $X_1=X(t_1), ..., X_n=X(t_n)$ are jointly Gaussian random variables for any $t_1,..., t_n$ whose joint characteristic function is given by

$$\oint_{X} \phi_{X}(\omega_{1}, \omega_{2}, \cdots, \omega_{n}) = e^{j\sum_{k=1}^{n} \mu(t_{k})\omega_{k} - \sum_{l,k}^{n} \sum_{k=1}^{n} C_{XX}(t_{i}, t_{k})\omega_{i}\omega_{k}/2}$$
The proof of th

- Examples
 - The process:

$$X(t) = a\cos(\omega_0 t + \varphi), \quad \varphi \sim U(0,2\pi).$$

- This process is wide sense stationary but not strict sense stationary.
- If the process X(t) has zero mean,

then
$$\sigma^2_z$$
 is reduced to:

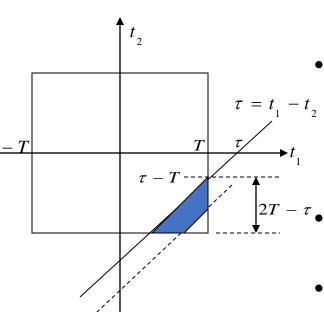
then
$$\sigma^2_z$$
 is reduced to:
$$z = \int_{-T}^T X(t) dt.$$

$$\sigma^2 = E\{|z|^2\} = \int_{-T}^T X(t) dt.$$

$$Z = \int_{-T} X(t)$$

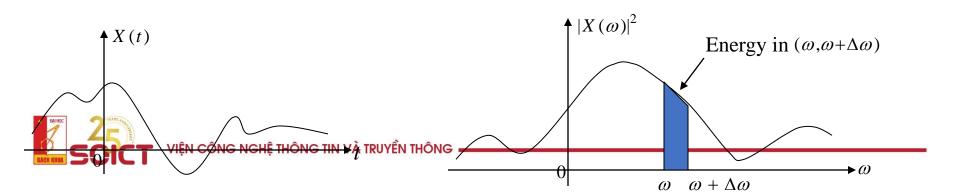
$$\sigma_z^2 = E\{|z|^2\} = \int_{-T}^T \int_{-T}^T R_{xx}(t_1 - t_2) dt_1 dt_2.$$

- $^{-\tau}$ As t₁ and t₂ varies from –T to T, so τ =t₂-t₁ varies from -2T to 2T.
 - $R_{xx}(\tau)$ is a constant over the shaded region in the figure on the left.



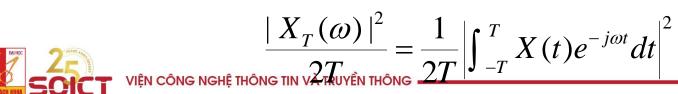
- Power spectrum
 - For a deterministic signal x(t)
 - The spectrum is well defined: If $X(\omega)$ represents its Fourier $X(\omega) = \int_{-\infty}^{+\infty} x(t)e^{-j\omega t}dt,$ transform.
 - then $|X(\omega)|^2$ represents its energy spectrum. This follows from Parseval's theorem since the signal energy is given by

$$\int_{-\infty}^{+\infty} x^2(t) dt = \frac{1}{2\pi} \int_{-\infty}^{+\infty} |X(\omega)|^2 d\omega = E.$$
• Thus $|X(\omega)|^2 \Delta \omega$ represents the signal energy in the band $(\omega, \omega + \Delta \omega)$



- For stochastic processes,
 - A direct application of Fourier transform generates a sequence of random variables for every ω
 - For a stochastic process, $E\{|X(t)|^2\}$ represents the ensemble average power (instantaneous energy) at the instant t.
 - Partial Fourier transform of a process X(t) based on (- T, T) is given by

 $X_T(\omega) = \int_{-T}^T X(t) e^{-j\omega t} dt$ • The power distribution associated with that realization based on (-T, T) is represented by





 The average power distribution based on (- T, T) is ensemble average of power distribution for ω

$$P_{T}(\omega) = E\left\{\frac{|X_{T}(\omega)|^{2}}{2T}\right\} = \frac{1}{2T}\int_{-T}^{T}\int_{-T}^{T}E\{X(t_{1})X^{*}(t_{2})\}e^{-j\omega(t_{1}-t_{2})}dt_{1}dt_{2}$$

$$= \frac{1}{2T}\int_{-T}^{T}\int_{-T}^{T}R_{xx}(t_{1},t_{2})e^{-j\omega(t_{1}-t_{2})}dt_{1}dt_{2}$$
• We have this represents the power distribution of $X(t)$

- based on (-T, T).
- If X(t) is assumed to be w.s.s, then
- and we have $R_{xx}(t_1, t_2) = R_{yy}(t_1 t_2)$



$$P(\omega) = \frac{1}{T} \int_{\text{VIEN cons nehệ thông 2NTA TRUYÊN THONG } T} R_{XX}(t_1 - t_2) e^{-j\omega(t_1 - t_2)} dt_1 dt_2.$$

• Let $\tau = t_2 - t_1$, we obtain

$$\begin{split} P_{_{T}}(\omega) &= \frac{1}{2T} \int_{-2T}^{2T} R_{_{XX}}(\tau) e^{-j\omega\tau} (2T - |\tau|) d\tau \\ &= \int_{-2T}^{2T} R_{_{XX}}(\tau) e^{-j\omega\tau} (1 - \frac{|\tau|}{2T}) d\tau \ \geq \ 0 \end{split}$$
 • This is the power distribution of the w.s.s. process X(t) based

- on (-T, T).
- Leting $T \rightarrow \infty$, we obtain

$$S_{xx}(\omega) = \lim_{T \to \infty} P_{\tau}(\omega) = \int_{-\infty}^{+\infty} R_{xx}(\tau) e^{-j\omega\tau} d\tau \ge 0$$

• $S_{xx}(\omega)$ is the power spectral density of the w.s.s process X(t).



- Khinchin-Wiener theorem
 - The autocorrelation function and the power spectrum of a w.s.s process form a Fourier transform pair.

$$R_{XX}(\omega) \xleftarrow{\text{F-T}} S_{XX}(\omega) \ge 0.$$

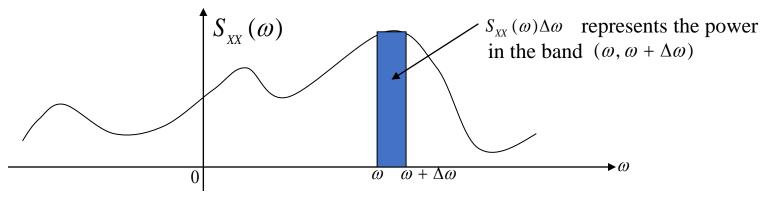
$$R_{XX}(\tau) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} S_{XX}(\omega) e^{j\omega\tau} d\omega$$

• For
$$\tau = 0_{XX}(\omega) = \lim_{T \to \infty} P_{T}(\omega) = \int_{-\infty}^{+\infty} R_{XX}(\tau) e^{-j\omega\tau} d\tau \ge 0$$

$$\frac{1}{2\pi} \int_{-\infty}^{+\infty} S_{xx}(\omega) d\omega = R_{xx}(0) = E\{|X(t)|^2\} = P, \quad the \ total \ power.$$



• The area under $S_{xx}(\omega)$ represents the total power of the process X(t), and hence $S_{xx}(\omega)$ truly represents the power spectrum.



• The nonnegative-definiteness property of the auto-correlation function translates into the "nonnegative" property for its Fourier transform (power spectrum).

$$R_{XX}(\tau)$$
 nonnegative - definite \iff $S_{XX}(\omega) \ge 0$.



If X(t) is a real w.s.s process, then

$$R_{_{XX}}(au)=R_{_{XX}}(- au)$$

So that

$$S_{XX}(\omega) = \int_{-\infty}^{+\infty} R_{XX}(\tau) e^{-j\omega\tau} d\tau$$
$$= \int_{-\infty}^{+\infty} R_{XX}(\tau) \cos \omega \tau d\tau$$

 $= 2\int_{-\infty}^{\infty} R_{xx}(\tau) \cos \omega \tau d\tau = S_{xx}(-\omega) \ge 0$ • The power spectrum is an even function, (in addition to being real and nonnegative).

4.4. Ergodicity

- Time averages
 - Given wide-sense stationary process X(t).
 - Time averages
 - Mean

$$n = \lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} x(t)dt$$

Autocorrelation

$$R(\tau) = \lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} x(t+\tau)x(t)dt$$

- These limits are random variables.
- Problems:



4.4. Ergodicity

- Ergodicity
 - X(t) is ergodic if in the most general form if all its statistics can be determined from a single function $X(t, \zeta)$ of the process.
 - X(t) is ergodic if time averages equal ensemble averages (expected values)



4.4. Ergodicity

- Ergodicity of the mean
 - Time average of a given process X(t)

$$\mathbf{n}_T = \frac{1}{2T} \int_{-T}^{T} \mathbf{x}(t) dt$$

- n_T is a random variable.
- Since E{X(t)} is a constant, we have

$$\mathsf{E}\{\mathsf{n}_\mathsf{T}\} = \mathsf{E}\{\mathsf{X}(\mathsf{t})\} = \mathsf{\eta}$$

The variance of n_T is given by:

$$\sigma_{n_T}^{\ 2} = \frac{1}{T} \int \left(1 - \frac{\tau}{2T} \right) [R(\tau) - \eta^2] d\tau$$
• R(\tau\) is the autocorrelation of X(t).

- If this variance tends to zero with $T \rightarrow \infty$, then n_T tends to its expected value.



Ergodic theorem for E{X(t)}

$$\lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} \mathbf{x}(t)dt = E\{\mathbf{x}(t)\} = \eta$$

$$iff$$

$$\lim_{T \to \infty} \frac{1}{T} \int_{0}^{2T} \left(1 - \frac{\tau}{2T}\right) [R(\tau) - \eta^{2}] d\tau = 0$$



- Ergodicity of autocorrelation
 - We form the average:

$$R_{T}(\lambda) = \frac{1}{2T} \int_{-T}^{T} x(t+\lambda)x(t)dt$$

We have

$$E\{\mathbf{R}_T(\lambda)\} = \frac{1}{2T}\int\limits_{-T}^T E\{\mathbf{x}(t+\lambda)\mathbf{x}(t)\}dt = R(\lambda)$$
 • For a given λ , $\mathbf{R}_T(\lambda)$ is the time average of the process

- $\Phi(t)=x(t+\lambda)x(t)$
- The mean of the process $\Phi(t)$ is given by

$$\mathsf{E}\{\Phi(t)\} = \mathsf{E}\{\mathsf{x}(t+\lambda)\mathsf{x}(t)\} = \mathsf{R}(\lambda)$$



Its autocorrelation

$$R_{\Phi\Phi}(\tau) = E\{x(t+\lambda+\tau)x(t+\tau)x(t+\lambda)x(t)\}$$

- Hence with $w(t) = \Phi(t)$, we have
- Ergodicity theorem for autocorrelation
 - For a given λ

$$\lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} \mathbf{x}(t+\lambda)\mathbf{x}(t)dt = E\{\mathbf{x}(t+\lambda)\mathbf{x}(t)\} = R(\lambda)$$
iff
$$\lim_{T \to \infty} \frac{1}{T} \int_{-T}^{2T} \left(1 - \frac{\tau}{2T}\right) [R_{\Phi\Phi}(\tau) - R^{2}(\lambda)]d\tau = 0$$



- Ergodicity of the distribution function
 - We determine first order distribution F(x) = E{X(t)≤x} of a given process X(t) by a suitable time average.
 - Consider the process

$$y(t) = \begin{cases} 1 & \text{if } x(t) \le x \\ 0 & \text{if } x(t) > x \end{cases}$$

- Its mean is given by
- Its autocorrelation $E\{y(t)\} = 1.P\{x(t) \le x\} = F(x)$
- Where $F(x,x;\tau)$ is the second order distribution of x(t)

$$E\{y(t+\tau)y(t)\} = 1.P\{x(t+\tau) \le x, x(t) \le x\} = F(x, x; \tau)$$



We form the time average

 $\mathbf{y}_T = \frac{1}{2T} \int_{-\infty}^{T} \mathbf{y}(t) dt$

We have

$$\mathsf{E}\{\mathsf{y}_\mathsf{T}\} = \mathsf{E}\{\mathsf{y}(\mathsf{t})\} = \mathsf{F}(\mathsf{x})$$

The variance of y_T is given by

$$\frac{1}{T}\int\limits_{0}^{2T}\left(1-\frac{\tau}{2T}\right)[R(\tau)-\eta^2]d\tau=0$$
• Where R(\tau) and \eta are replaced by F(x, x;\tau) and F(x)



- Ergodic theorem for distribution function
 - For a given x,

$$\lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} y(t)dt = F(x)$$

iff

$$\lim_{T\to\infty}\frac{1}{T}\int_{0}^{2T}\left(1-\frac{\tau}{2T}\right)[F(x,x;\tau)-F^{2}(x)]d\tau=0$$



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Thank you for your attentions!

