

# Machine Learning – Lecture 2

## Probability Density Estimation

15.10.2018

Bastian Leibe

RWTH Aachen

<http://www.vision.rwth-aachen.de>

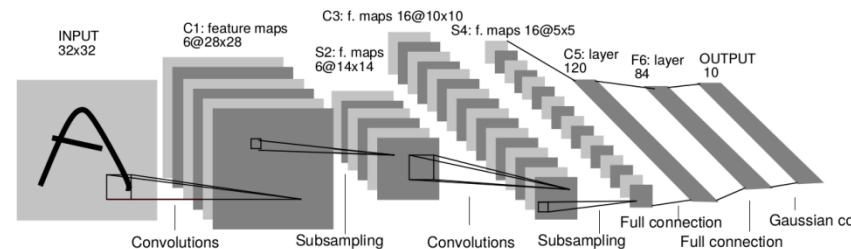
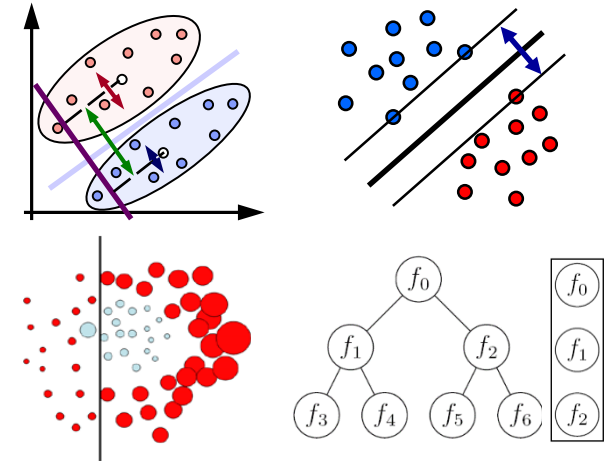
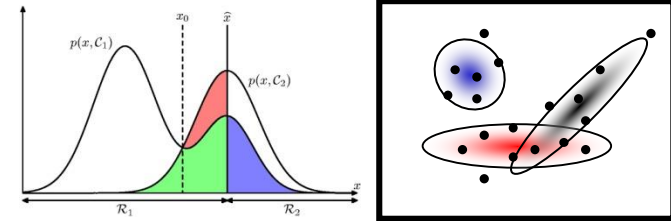
[leibe@vision.rwth-aachen.de](mailto:leibe@vision.rwth-aachen.de)

# Announcements: Reminders

- L2P electronic repository
  - Slides, exercises, and supplementary material will be made available here
  - Lecture recordings will be uploaded 2-3 days after the lecture
  - *L2P access should now be fixed for all registered participants!*
- Course webpage
  - <http://www.vision.rwth-aachen.de/courses/>
  - Slides will also be made available on the webpage
- Please subscribe to the lecture on rwth online!
  - Important to get email announcements and L2P access!

# Course Outline

- Fundamentals
  - Bayes Decision Theory
  - Probability Density Estimation
- Classification Approaches
  - Linear Discriminants
  - Support Vector Machines
  - Ensemble Methods & Boosting
  - Randomized Trees, Forests & Ferns
- Deep Learning
  - Foundations
  - Convolutional Neural Networks
  - Recurrent Neural Networks



# Topics of This Lecture

- Bayes Decision Theory
  - Basic concepts
  - Minimizing the misclassification rate
  - Minimizing the expected loss
  - Discriminant functions
- Probability Density Estimation
  - General concepts
  - Gaussian distribution
- Parametric Methods
  - Maximum Likelihood approach
  - Bayesian vs. Frequentist views on probability

# Recap: The Rules of Probability

- We have shown in the last lecture

**Sum Rule** 
$$p(X) = \sum_Y p(X, Y)$$

**Product Rule** 
$$p(X, Y) = p(Y|X)p(X)$$

- From those, we can derive

**Bayes' Theorem** 
$$p(Y|X) = \frac{p(X|Y)p(Y)}{p(X)}$$

**where** 
$$p(X) = \sum_Y p(X|Y)p(Y)$$

# Bayes Decision Theory



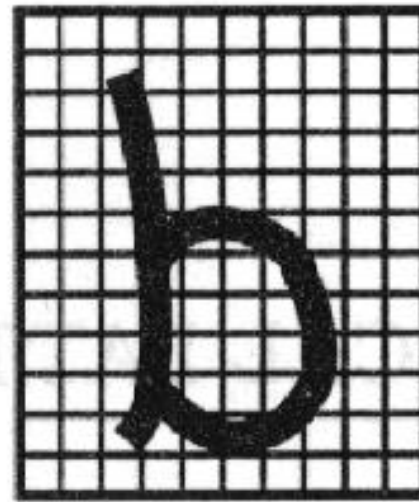
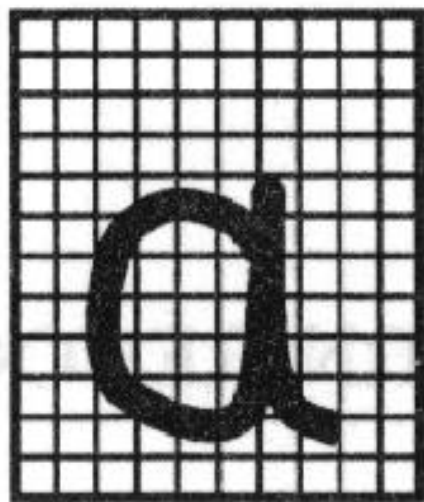
**Thomas Bayes, 1701-1761**

*“The theory of inverse probability is founded upon an error, and must be wholly rejected.”*

R.A. Fisher, 1925

# Bayes Decision Theory

- Example: handwritten character recognition



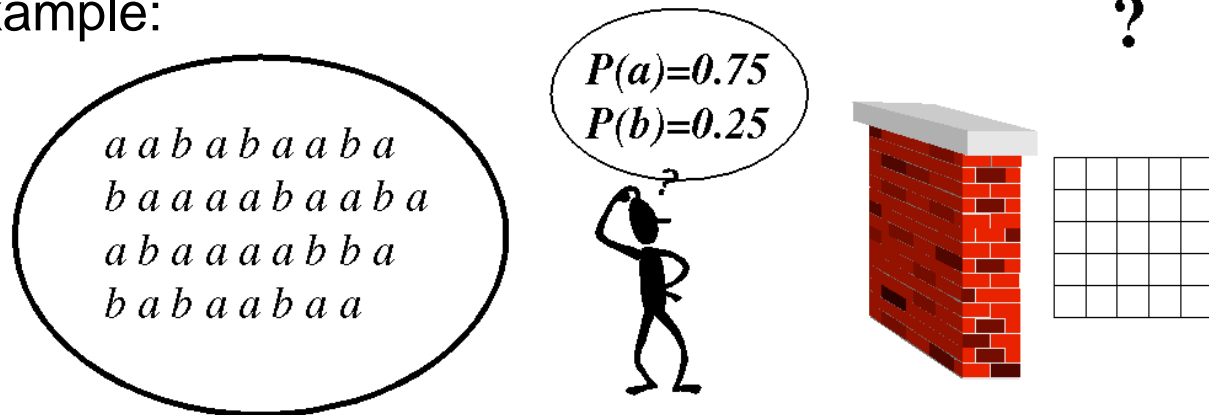
- Goal:
  - Classify a new letter such that the probability of misclassification is minimized.

# Bayes Decision Theory

- Concept 1: **Priors** (a priori probabilities)

$$p(C_k)$$

- What we can tell about the probability *before seeing the data*.
- Example:



$$C_1 = a$$

$$p(C_1) = 0.75$$

$$C_2 = b$$

$$p(C_2) = 0.25$$

- In general: 
$$\sum_k p(C_k) = 1$$



# Bayes Decision Theory

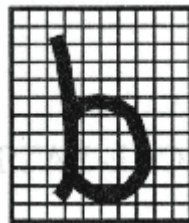
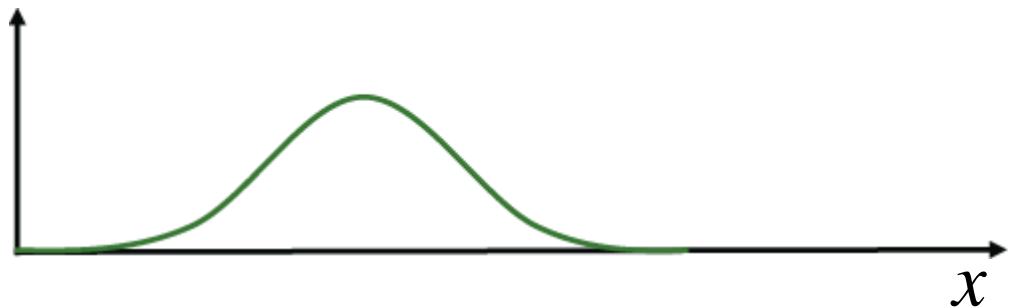
- Concept 2: **Conditional probabilities**

$$p(x | C_k)$$

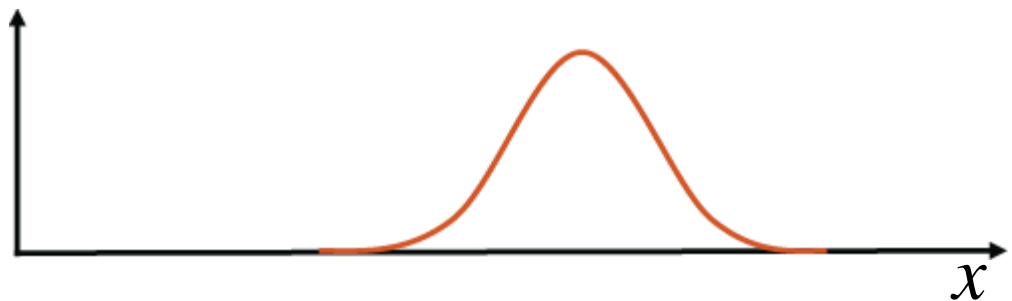
- Let  $x$  be a feature vector.
- $x$  measures/describes certain properties of the input.
  - E.g. number of black pixels, aspect ratio, ...
- $p(x|C_k)$  describes its **likelihood** for class  $C_k$ .



$$p(x | a)$$

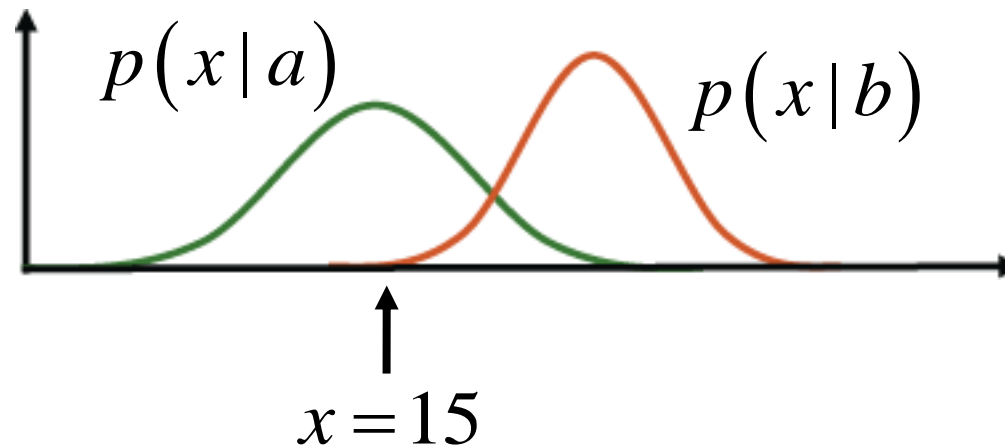


$$p(x | b)$$



# Bayes Decision Theory

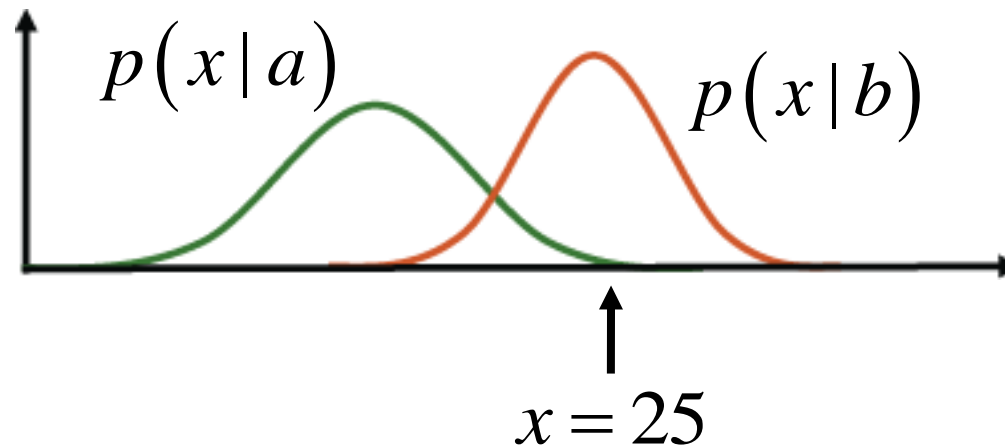
- Example:



- Question:
  - Which class?
  - Since  $p(x|b)$  is much smaller than  $p(x|a)$ , the decision should be 'a' here.

# Bayes Decision Theory

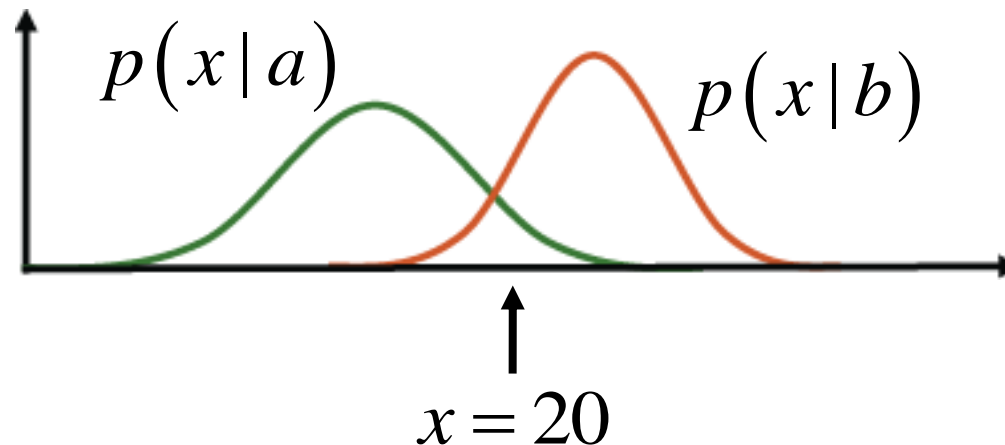
- Example:



- Question:
  - Which class?
  - Since  $p(x|a)$  is much smaller than  $p(x|b)$ , the decision should be 'b' here.

# Bayes Decision Theory

- Example:



- Question:
  - Which class?
  - Remember that  $p(a) = 0.75$  and  $p(b) = 0.25...$
  - I.e., the decision should be again 'a'.
- ⇒ How can we formalize this?

# Bayes Decision Theory

- Concept 3: **Posterior probabilities**

$$p(C_k | x)$$

- We are typically interested in the *a posteriori* probability, i.e., the probability of class  $C_k$  given the measurement vector  $x$ .

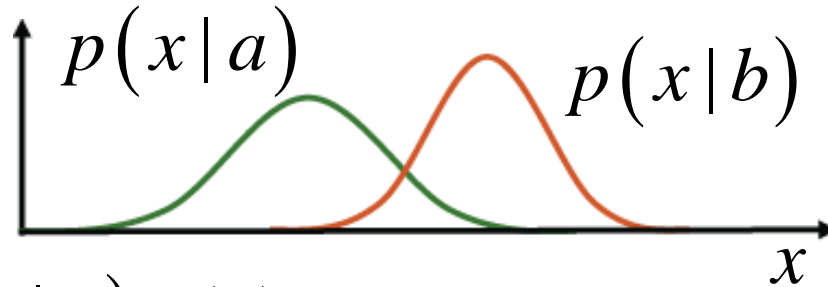
- Bayes' Theorem:

$$p(C_k | x) = \frac{p(x | C_k) p(C_k)}{p(x)} = \frac{p(x | C_k) p(C_k)}{\sum_i p(x | C_i) p(C_i)}$$

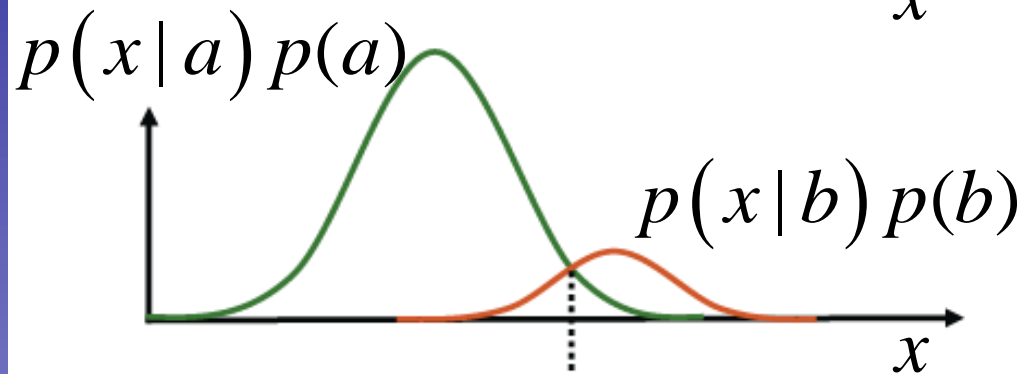
- Interpretation

$$\textit{Posterior} = \frac{\textit{Likelihood} \times \textit{Prior}}{\textit{Normalization Factor}}$$

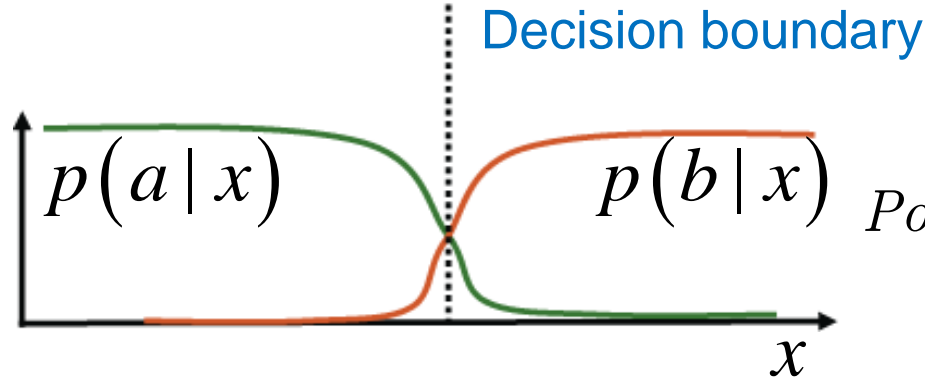
# Bayes Decision Theory



*Likelihood*



*Likelihood  $\times$  Prior*



Decision boundary

$$\text{Posterior} = \frac{\text{Likelihood} \times \text{Prior}}{\text{NormalizationFactor}}$$

# Bayesian Decision Theory

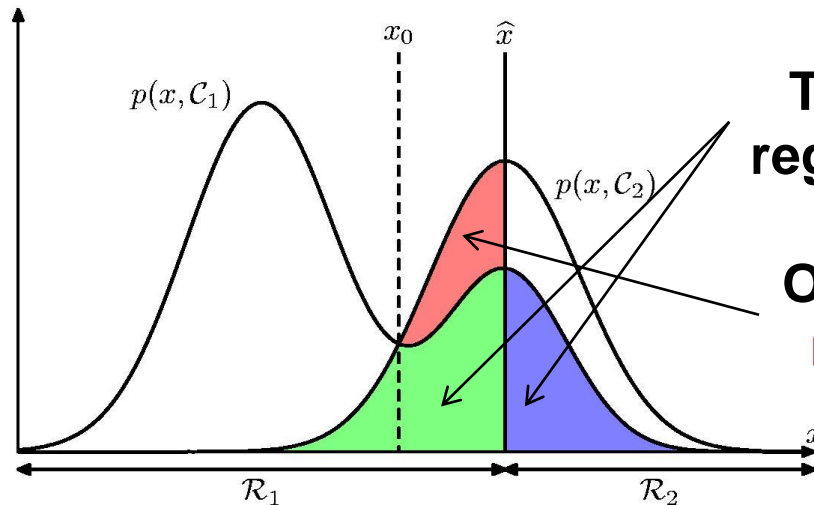
- Goal: Minimize the probability of a misclassification

Decision rule:

$$x < \hat{x} \Rightarrow \mathcal{C}_1$$

$$x \geq \hat{x} \Rightarrow \mathcal{C}_2$$

How does  
 $p(\text{mistake})$  change  
when we move  $\hat{x}$ ?



The **green** and **blue**  
regions stay constant.

Only the size of the  
**red** region varies!

$$\begin{aligned}
 p(\text{mistake}) &= p(\mathbf{x} \in \mathcal{R}_1, \mathcal{C}_2) + p(\mathbf{x} \in \mathcal{R}_2, \mathcal{C}_1) \\
 &= \int_{\mathcal{R}_1} p(\mathbf{x}, \mathcal{C}_2) d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathbf{x}, \mathcal{C}_1) d\mathbf{x} \\
 &= \int_{\mathcal{R}_1} p(\mathcal{C}_2|\mathbf{x})p(\mathbf{x})d\mathbf{x} + \int_{\mathcal{R}_2} p(\mathcal{C}_1|\mathbf{x})p(\mathbf{x})d\mathbf{x}
 \end{aligned}$$

# Bayes Decision Theory

- Optimal decision rule

- Decide for  $C_1$  if

$$p(C_1|x) > p(C_2|x)$$

- This is equivalent to

$$p(x|C_1)p(C_1) > p(x|C_2)p(C_2)$$

- Which is again equivalent to (Likelihood-Ratio test)

$$\frac{p(x|C_1)}{p(x|C_2)} > \underbrace{\frac{p(C_2)}{p(C_1)}}_{\text{Decision threshold } \theta}$$

Decision threshold  $\theta$



# Generalization to More Than 2 Classes

- Decide for class  $k$  whenever it has the greatest posterior probability of all classes:

$$p(\mathcal{C}_k|x) > p(\mathcal{C}_j|x) \quad \forall j \neq k$$

$$p(x|\mathcal{C}_k)p(\mathcal{C}_k) > p(x|\mathcal{C}_j)p(\mathcal{C}_j) \quad \forall j \neq k$$

- Likelihood-ratio test

$$\frac{p(x|\mathcal{C}_k)}{p(x|\mathcal{C}_j)} > \frac{p(\mathcal{C}_j)}{p(\mathcal{C}_k)} \quad \forall j \neq k$$

# Classifying with Loss Functions

- Generalization to decisions with a **loss function**
  - Differentiate between the possible decisions and the possible true classes.
  - Example: medical diagnosis
    - Decisions: *sick* or *healthy* (or: *further examination necessary*)
    - Classes: patient is *sick* or *healthy*
  - The cost may be asymmetric:

$$\begin{aligned} \text{loss}(\text{decision} = \text{healthy} | \text{patient} = \text{sick}) &>> \\ \text{loss}(\text{decision} = \text{sick} | \text{patient} = \text{healthy}) \end{aligned}$$

# Classifying with Loss Functions

- In general, we can formalize this by introducing a loss matrix  $L_{kj}$

$L_{kj} = \text{loss for decision } C_j \text{ if truth is } C_k.$

- Example: cancer diagnosis

$$L_{\text{cancer diagnosis}} = \begin{array}{c} \text{Truth} \\ \text{cancer} \\ \text{normal} \end{array} \begin{array}{cc} \text{Decision} & \\ \text{cancer} & \text{normal} \end{array} \begin{pmatrix} 0 & 1000 \\ 1 & 0 \end{pmatrix}$$

# Classifying with Loss Functions

- Loss functions may be different for different actors.

➤ Example:

$$L_{stocktrader}(subprime) = \begin{matrix} & \begin{matrix} \text{"invest"} & \text{"don't invest"} \end{matrix} \\ \begin{pmatrix} -\frac{1}{2}c_{gain} & 0 \\ 0 & 0 \end{pmatrix} \end{matrix}$$



$$L_{bank}(subprime) = \begin{matrix} & \begin{matrix} -\frac{1}{2}c_{gain} & 0 \end{matrix} \\ \begin{pmatrix} & \\ \text{skull and crossbones} & 0 \end{pmatrix} \end{matrix}$$



⇒ Different loss functions may lead to different Bayes optimal strategies.

# Minimizing the Expected Loss

- Optimal solution is the one that minimizes the loss.
  - But: loss function depends on the true class, which is unknown.
- Solution: Minimize the expected loss

$$\mathbb{E}[L] = \sum_k \sum_j \int_{\mathcal{R}_j} L_{kj} p(\mathbf{x}, \mathcal{C}_k) d\mathbf{x}$$

- This can be done by choosing the regions  $\mathcal{R}_j$  such that

$$\mathbb{E}[L] = \sum_k L_{kj} p(\mathcal{C}_k | \mathbf{x})$$

which is easy to do once we know the posterior class probabilities  $p(\mathcal{C}_k | \mathbf{x})$

# Minimizing the Expected Loss

- Example:

- 2 Classes:  $C_1, C_2$
- 2 Decision:  $\alpha_1, \alpha_2$
- Loss function:  $L(\alpha_j | \mathcal{C}_k) = L_{kj}$
- Expected loss (= risk  $R$ ) for the two decisions:

$$\mathbb{E}_{\alpha_1}[L] = R(\alpha_1 | \mathbf{x}) = L_{11}p(\mathcal{C}_1 | \mathbf{x}) + L_{21}p(\mathcal{C}_2 | \mathbf{x})$$

$$\mathbb{E}_{\alpha_2}[L] = R(\alpha_2 | \mathbf{x}) = L_{12}p(\mathcal{C}_1 | \mathbf{x}) + L_{22}p(\mathcal{C}_2 | \mathbf{x})$$

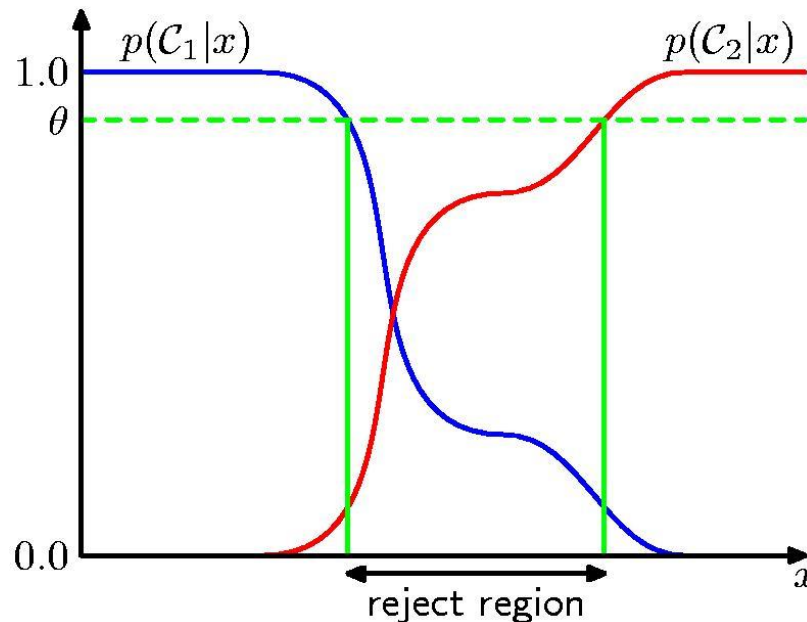
- Goal: Decide such that expected loss is minimized
  - I.e. decide  $\alpha_1$  if  $R(\alpha_2 | \mathbf{x}) > R(\alpha_1 | \mathbf{x})$

# Minimizing the Expected Loss

$$\begin{aligned}R(\alpha_2|\mathbf{x}) &> R(\alpha_1|\mathbf{x}) \\L_{12}p(\mathcal{C}_1|\mathbf{x}) + L_{22}p(\mathcal{C}_2|\mathbf{x}) &> L_{11}p(\mathcal{C}_1|\mathbf{x}) + L_{21}p(\mathcal{C}_2|\mathbf{x}) \\(L_{12} - L_{11})p(\mathcal{C}_1|\mathbf{x}) &> (L_{21} - L_{22})p(\mathcal{C}_2|\mathbf{x}) \\\frac{(L_{12} - L_{11})}{(L_{21} - L_{22})} &> \frac{p(\mathcal{C}_2|\mathbf{x})}{p(\mathcal{C}_1|\mathbf{x})} = \frac{p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)}{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)} \\\frac{p(\mathbf{x}|\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_2)} &> \frac{(L_{21} - L_{22}) p(\mathcal{C}_2)}{(L_{12} - L_{11}) p(\mathcal{C}_1)}\end{aligned}$$

⇒ Adapted decision rule taking into account the loss.

# The Reject Option



- Classification errors arise from regions where the largest posterior probability  $p(\mathcal{C}_k|\mathbf{x})$  is significantly less than 1.
  - These are the regions where we are relatively uncertain about class membership.
  - For some applications, it may be better to reject the automatic decision entirely in such a case and, e.g., consult a human expert.



# Discriminant Functions

- Formulate classification in terms of comparisons

- Discriminant functions

$$y_1(x), \dots, y_K(x)$$

- Classify  $x$  as class  $C_k$  if

$$y_k(x) > y_j(x) \quad \forall j \neq k$$

- Examples (Bayes Decision Theory)

$$y_k(x) = p(C_k|x)$$

$$y_k(x) = p(x|C_k)p(C_k)$$

$$y_k(x) = \log p(x|C_k) + \log p(C_k)$$

# Different Views on the Decision Problem

- $y_k(x) \propto p(x|\mathcal{C}_k)p(\mathcal{C}_k)$ 
  - First determine the class-conditional densities for each class individually and separately infer the prior class probabilities.
  - Then use Bayes' theorem to determine class membership.

⇒ *Generative methods*
- $y_k(x) = p(\mathcal{C}_k|x)$ 
  - First solve the inference problem of determining the posterior class probabilities.
  - Then use decision theory to assign each new  $x$  to its class.

⇒ *Discriminative methods*
- Alternative
  - Directly find a discriminant function  $y_k(x)$  which maps each input  $x$  directly onto a class label.

# Topics of This Lecture

- Bayes Decision Theory
  - Basic concepts
  - Minimizing the misclassification rate
  - Minimizing the expected loss
  - Discriminant functions
- Probability Density Estimation
  - General concepts
  - Gaussian distribution
- Parametric Methods
  - Maximum Likelihood approach
  - Bayesian vs. Frequentist views on probability
  - Bayesian Learning

# Probability Density Estimation

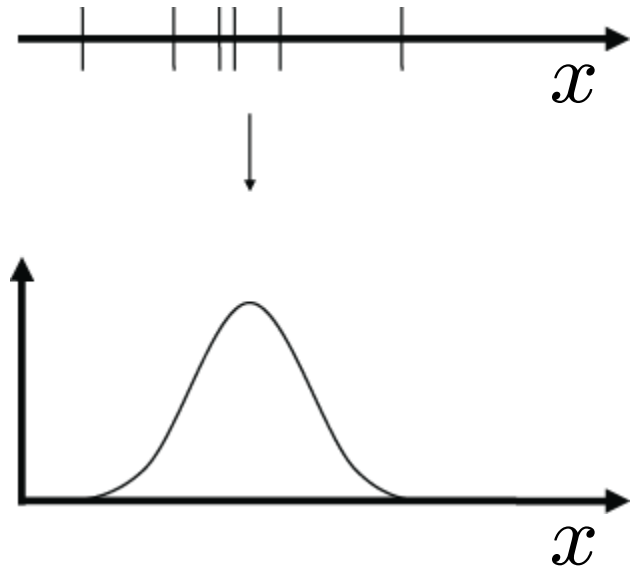
- Up to now
  - Bayes optimal classification
  - Based on the probabilities  $p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)$
- How can we estimate (= learn) those probability densities?
  - Supervised training case: data and class labels are known.
  - Estimate the probability density for each class  $\mathcal{C}_k$  separately:

$$p(\mathbf{x}|\mathcal{C}_k)$$

- (For simplicity of notation, we will drop the class label  $\mathcal{C}_k$  in the following.)

# Probability Density Estimation

- Data:  $x_1, x_2, x_3, x_4, \dots$
- Estimate:  $p(x)$
- Methods
  - Parametric representations
  - Non-parametric representations
  - Mixture models



(today)

(lecture 3)

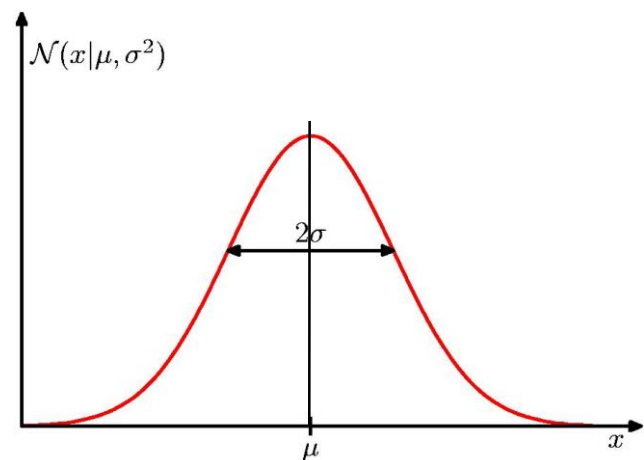
(lecture 4)

# The Gaussian (or Normal) Distribution

- One-dimensional case

- Mean  $\mu$
- Variance  $\sigma^2$

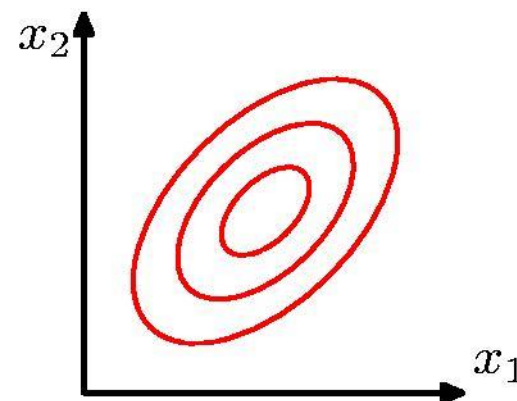
$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp \left\{ -\frac{(x - \mu)^2}{2\sigma^2} \right\}$$



- Multi-dimensional case

- Mean  $\mu$
- Covariance  $\Sigma$

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}|\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \right\}$$

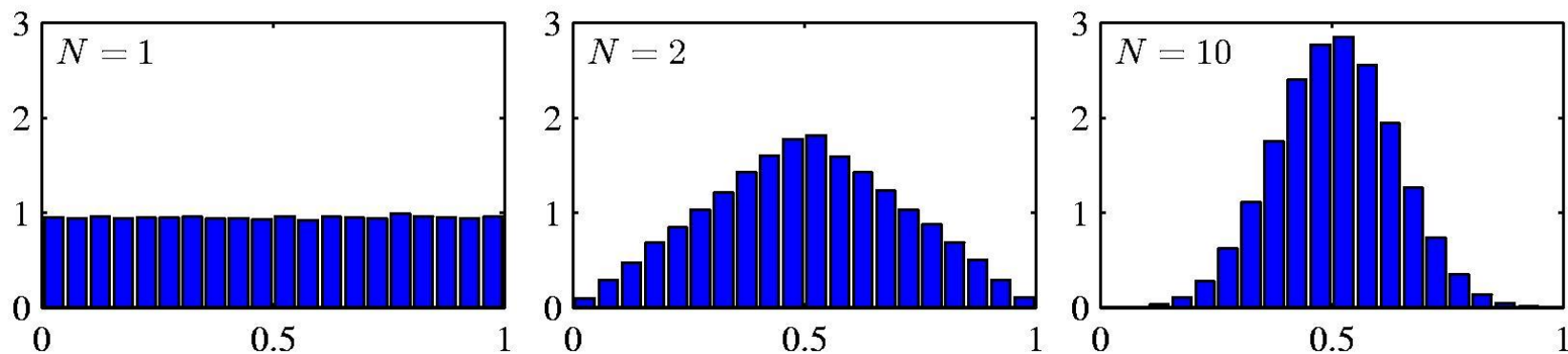


# Gaussian Distribution – Properties

- Central Limit Theorem

- “The distribution of the sum of  $N$  i.i.d. random variables becomes increasingly Gaussian as  $N$  grows.”
- In practice, the convergence to a Gaussian can be very rapid.
- This makes the Gaussian interesting for many applications.

- Example:  $N$  uniform  $[0,1]$  random variables.



# Gaussian Distribution – Properties

- Quadratic Form

- $\mathcal{N}$  depends on  $\mathbf{x}$  through the exponent

$$\Delta^2 = (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})$$

- Here,  $\Delta$  is often called the **Mahalanobis distance** from  $\mathbf{x}$  to  $\boldsymbol{\mu}$ .

- Shape of the Gaussian

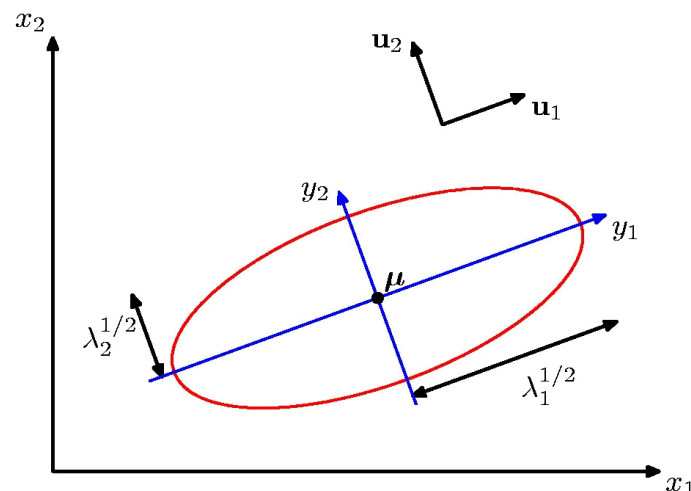
- $\boldsymbol{\Sigma}$  is a real, symmetric matrix.
  - We can therefore decompose it into its eigenvectors

$$\boldsymbol{\Sigma} = \sum_{i=1}^D \lambda_i \mathbf{u}_i \mathbf{u}_i^T$$

$$\boldsymbol{\Sigma}^{-1} = \sum_{i=1}^D \frac{1}{\lambda_i} \mathbf{u}_i \mathbf{u}_i^T$$

and thus obtain  $\Delta^2 = \sum_{i=1}^D \frac{y_i^2}{\lambda_i}$  with  $y_i = \mathbf{u}_i^T (\mathbf{x} - \boldsymbol{\mu})$

⇒ **Constant density on ellipsoids** with main directions along the eigenvectors  $\mathbf{u}_i$  and scaling factors  $\sqrt{\lambda_i}$





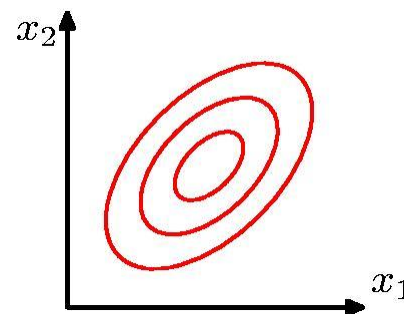
# Gaussian Distribution – Properties

- Special cases

- Full covariance matrix

$$\Sigma = [\sigma_{ij}]$$

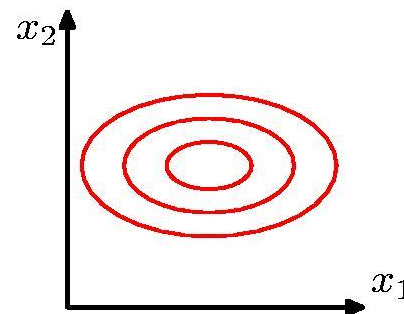
⇒ General ellipsoid shape



- Diagonal covariance matrix

$$\Sigma = \text{diag}\{\sigma_i\}$$

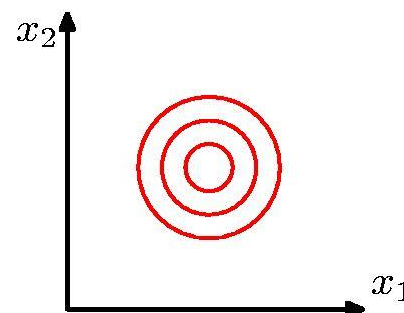
⇒ Axis-aligned ellipsoid



- Uniform variance

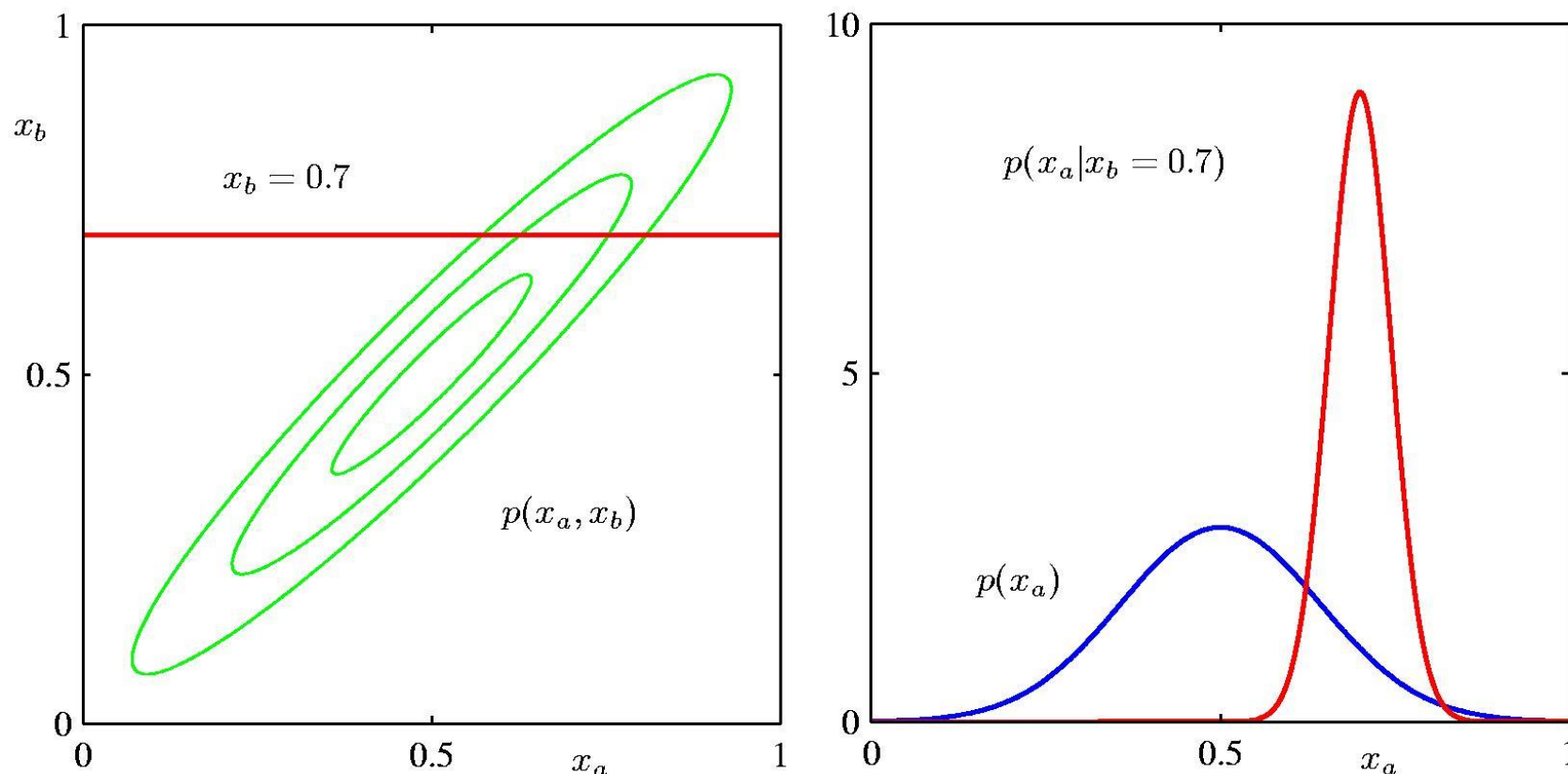
$$\Sigma = \sigma^2 \mathbf{I}$$

⇒ Hypersphere



# Gaussian Distribution – Properties

- The marginals of a Gaussian are again Gaussians:



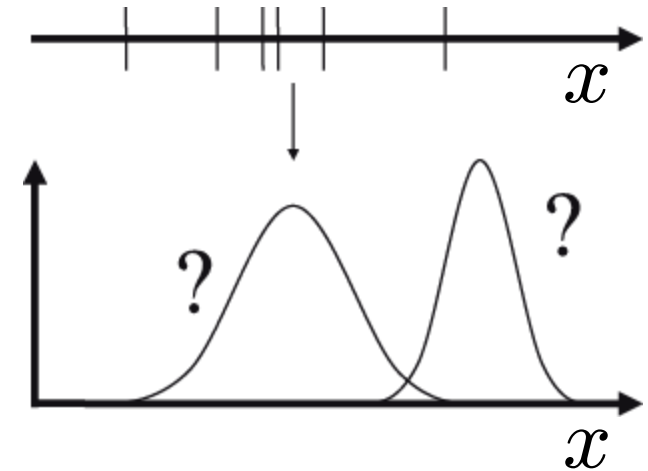
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  - Basic concepts
  - Minimizing the misclassification rate
  - Minimizing the expected loss
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- Probability Density Estimation
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- Parametric Methods
  - Maximum Likelihood approach
  - Bayesian vs. Frequentist views on probability

# Parametric Methods

- Given

- Data  $X = \{x_1, x_2, \dots, x_N\}$
- Parametric form of the distribution with parameters  $\theta$
- E.g. for Gaussian distrib.:  $\theta = (\mu, \sigma)$



- Learning

- Estimation of the parameters  $\theta$

- Likelihood of  $\theta$

- Probability that the data  $X$  have indeed been generated from a probability density with parameters  $\theta$

$$L(\theta) = p(X|\theta)$$

# Maximum Likelihood Approach

- Computation of the likelihood

- Single data point:  $p(x_n|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp \left\{ -\frac{(x - \mu)^2}{2\sigma^2} \right\}$

- Assumption: all data points are independent

$$L(\theta) = p(X|\theta) = \prod_{n=1}^N p(x_n|\theta)$$

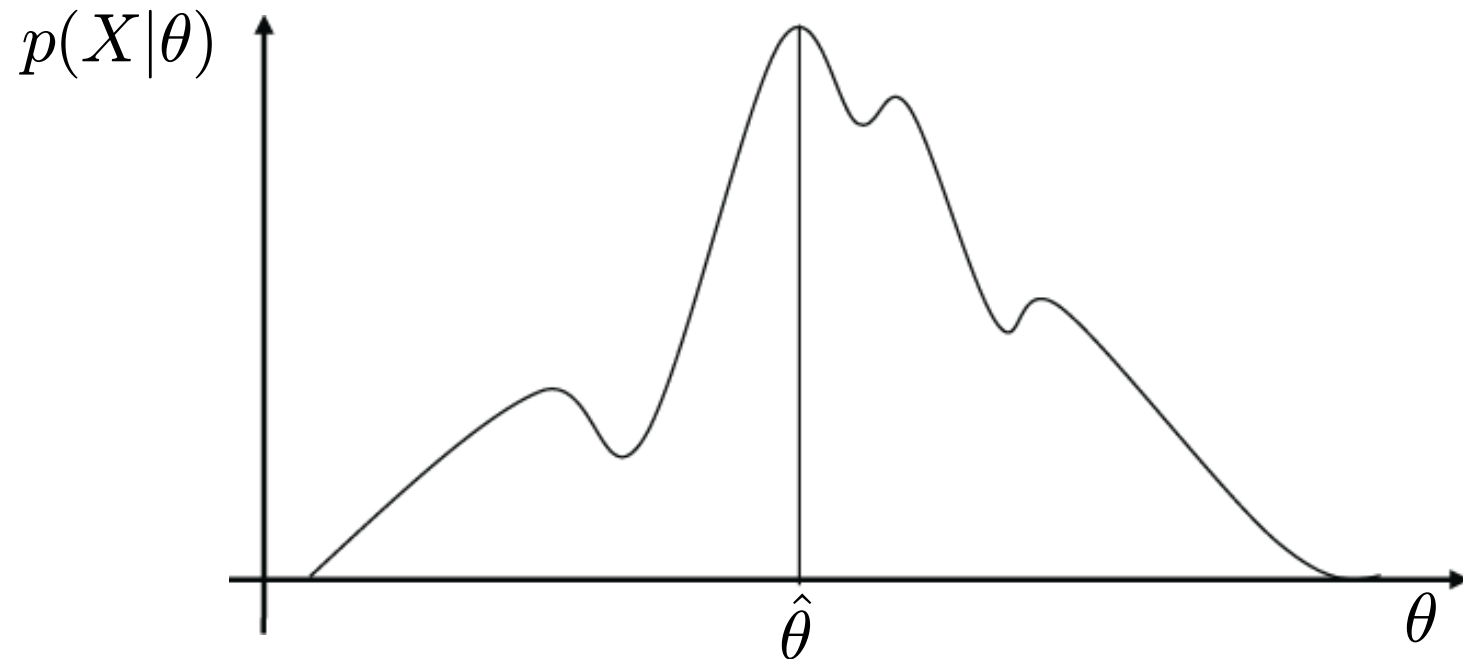
- Log-likelihood

$$E(\theta) = -\ln L(\theta) = -\sum_{n=1}^N \ln p(x_n|\theta)$$

- Estimation of the parameters  $\theta$  (Learning)
    - Maximize the likelihood
    - Minimize the negative log-likelihood

# Maximum Likelihood Approach

- Likelihood:  $L(\theta) = p(X|\theta) = \prod_{n=1}^N p(x_n|\theta)$
- We want to obtain  $\hat{\theta}$  such that  $L(\hat{\theta})$  is maximized.



# Maximum Likelihood Approach

- Minimizing the log-likelihood

- How do we minimize a function?

- ⇒ Take the derivative and set it to zero.

$$\frac{\partial}{\partial \theta} E(\theta) = -\frac{\partial}{\partial \theta} \sum_{n=1}^N \ln p(x_n | \theta) = -\sum_{n=1}^N \frac{\frac{\partial}{\partial \theta} p(x_n | \theta)}{p(x_n | \theta)} \stackrel{!}{=} 0$$

- Log-likelihood for Normal distribution (1D case)

$$\begin{aligned} E(\theta) &= -\sum_{n=1}^N \ln p(x_n | \mu, \sigma) \\ &= -\sum_{n=1}^N \ln \left( \frac{1}{\sqrt{2\pi}\sigma} \exp \left\{ -\frac{\|x_n - \mu\|^2}{2\sigma^2} \right\} \right) \end{aligned}$$

# Maximum Likelihood Approach

- Minimizing the log-likelihood

$$\frac{\partial}{\partial \mu} E(\mu, \sigma) = - \sum_{n=1}^N \frac{\frac{\partial}{\partial \mu} p(x_n | \mu, \sigma)}{p(x_n | \mu, \sigma)}$$

$$= - \sum_{n=1}^N - \frac{2(x_n - \mu)}{2\sigma^2}$$

$$= \frac{1}{\sigma^2} \sum_{n=1}^N (x_n - \mu)$$

$$= \frac{1}{\sigma^2} \left( \sum_{n=1}^N x_n - N\mu \right)$$

$$\frac{\partial}{\partial \mu} E(\mu, \sigma) \stackrel{!}{=} 0 \quad \Leftrightarrow \quad \hat{\mu} = \frac{1}{N} \sum_{n=1}^N x_n$$

$$p(x_n | \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{||x_n - \mu||^2}{2\sigma^2}}$$



# Maximum Likelihood Approach

- We thus obtain

$$\hat{\mu} = \frac{1}{N} \sum_{n=1}^N x_n$$

*“sample mean”*

- In a similar fashion, we get

$$\hat{\sigma}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \hat{\mu})^2$$

*“sample variance”*

- $\hat{\theta} = (\hat{\mu}, \hat{\sigma})$  is the **Maximum Likelihood estimate** for the parameters of a Gaussian distribution.
- This is a very important result.
- Unfortunately, it is wrong...

# Maximum Likelihood Approach

- Or not wrong, but rather **biased**...
- Assume the samples  $x_1, x_2, \dots, x_N$  come from a true Gaussian distribution with mean  $\mu$  and variance  $\sigma^2$ 
  - We can now compute the expectations of the ML estimates with respect to the data set values. It can be shown that

$$\begin{aligned}\mathbb{E}(\mu_{\text{ML}}) &= \mu \\ \mathbb{E}(\sigma_{\text{ML}}^2) &= \left(\frac{N-1}{N}\right) \sigma^2\end{aligned}$$

⇒ The ML estimate will underestimate the true variance.

- Corrected estimate:

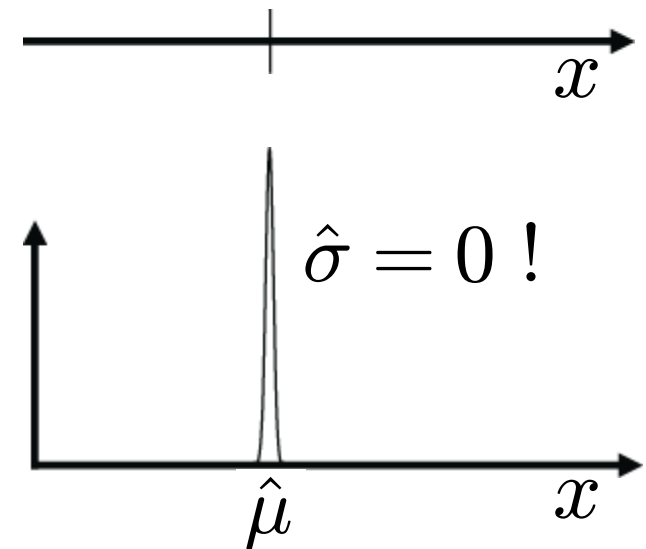
$$\tilde{\sigma}^2 = \frac{N}{N-1} \sigma_{\text{ML}}^2 = \frac{1}{N-1} \sum_{n=1}^N (x_n - \hat{\mu})^2$$

# Maximum Likelihood – Limitations

- Maximum Likelihood has several significant limitations
  - It systematically underestimates the variance of the distribution!
  - E.g. consider the case

$$N = 1, X = \{x_1\}$$

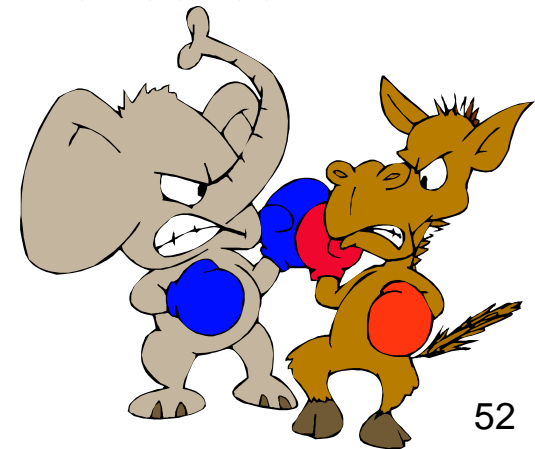
⇒ Maximum-likelihood estimate:



- We say ML *overfits to the observed data*.
- We will still often use ML, but it is important to know about this effect.

# Deeper Reason

- Maximum Likelihood is a **Frequentist** concept
  - In the **Frequentist view**, probabilities are the frequencies of random, repeatable events.
  - These frequencies are fixed, but can be estimated more precisely when more data is available.
- This is in contrast to the **Bayesian** interpretation
  - In the **Bayesian view**, probabilities quantify the uncertainty about certain states or events.
  - This uncertainty can be revised in the light of new evidence.
- Bayesians and Frequentists do not like each other too well...



# Bayesian vs. Frequentist View

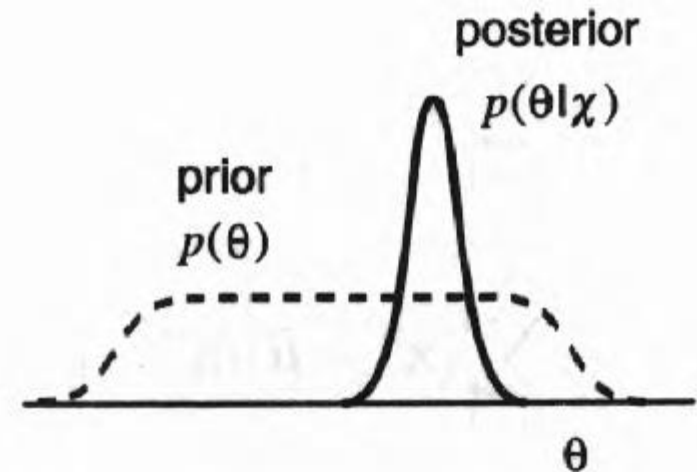
- To see the difference...
  - Suppose we want to estimate the uncertainty whether the Arctic ice cap will have disappeared by the end of the century.
  - This question makes no sense in a Frequentist view, since the event cannot be repeated numerous times.
  - In the Bayesian view, we generally have a prior, e.g., from calculations how fast the polar ice is melting.
  - If we now get fresh evidence, e.g., from a new satellite, we may revise our opinion and update the uncertainty from the prior.

$$Posterior \propto Likelihood \times Prior$$

- This generally allows to get better uncertainty estimates for many situations.
- Main Frequentist criticism
  - The prior has to come from somewhere and if it is wrong, the result will be worse.

# Bayesian Approach to Parameter Learning

- Conceptual shift
  - Maximum Likelihood views the true parameter vector  $\theta$  to be unknown, but fixed.
  - In Bayesian learning, we consider  $\theta$  to be a random variable.
- This allows us to use knowledge about the parameters  $\theta$ 
  - i.e. to use a prior for  $\theta$
  - Training data then converts this prior distribution on  $\theta$  into a posterior probability density.



# Bayesian Learning

- Bayesian Learning is an important concept
  - However, it would lead to far here.
  - ⇒ I will introduce it in more detail in the [Advanced ML lecture](#).

# References and Further Reading

- More information in Bishop's book
  - Gaussian distribution and ML: Ch. 1.2.4 and 2.3.1-2.3.4.
  - Bayesian Learning: Ch. 1.2.3 and 2.3.6.
  - Nonparametric methods: Ch. 2.5.

Christopher M. Bishop  
Pattern Recognition and Machine Learning  
Springer, 2006

