

# The Voters' Veto: Local Racial Demographic Change and Exclusionary Behavior

Michael Hankinson\*

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## Abstract

As Western democracies diversify, racial threat risks activating exclusionary attitudes among members of the majority. While research has focused on support for immigration restrictions, local housing policy is the primary means of maintaining racial segregation in the US. Using referendum vote share from over 3,700 precincts in Los Angeles County, I show that an increase in the local non-white population is associated with greater behavioral support for direct democratic control over affordable housing. I argue that this “voters’ veto” is valued for excluding low-income, predominantly non-white residents from one’s community. This relationship exists across majority white, Black, and Latino precincts, demonstrating that exclusionary reactions are not limited to racial threat among the majority. Instead, a shrinking white population may signal shared concerns about neighborhood trajectory. Given local conditions can shape political beliefs, these findings provide a mechanism for the shift of even minority voters towards exclusionary political movements.

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\*Assistant Professor, Department of Political Science, George Washington University. 2115 G St NW, Washington, DC 20052. hankinson@gwu.edu. ORCID: 0000-0003-4099-2633.

The United States and other Western democracies are rapidly racially diversifying, largely due to immigration (Frey 2018). According to theories of racial threat, this increase in the multiracial minority population may present an economic, cultural, and political threat to the majority demographic group (Blalock 1967; Blumer 1958). Empirically, a growing minority population at the local and national level has been found to elicit exclusionary attitudes and behaviors among members of the majority racial group (see Craig, Rucker and Richeson (2018) and Kaufmann and Goodwin (2018) for review). However, three overlapping limitations have hampered the advancement of theory in this space.

First, researchers have largely focused on the exclusionary responses of the current majority group in these democracies — non-Hispanic whites (e.g., Abrajano and Hajnal 2015; Campbell, Wong and Citrin 2006; Fouka and Tabellini 2022; Outten et al. 2012; Reny and Newman 2018; Sahn Forthcoming). This emphasis on white voters risks oversimplifying emerging political cleavages into a majority-minority dichotomy and overlooking the motivations and behaviors of Black, Latino, Asian, and other non-white voters. While white voters may be driven by racial threat, minority residents may also respond to the attendant changes that come with local demographic change, such as concerns over the social and economic stability of their community. Still, research which has examined the attitudes of minority residents has largely relied on static distributions of group size to understand intergroup conflict rather than the dynamic process of demographic change (e.g., Bobo and Hutchings 1996; Gay 2006; McClain and Karnig 1990; Oliver and Wong 2003).

Second, research on demographic change and racial threat has focused heavily on either exclusionary attitudes or on behaviors that only indirectly shape policy outcomes. Examples of these indirect behaviors include increased voter turnout and support for conservative candidates (e.g., Enos 2016; Newman, Shah and Collingwood 2018; Hill, Hopkins and Huber 2019). But neither of these dependent variables directly captures voters' behavioral support for exclusionary policy. This limitation likely stems from the fact that citizens rarely get to vote directly on national exclusionary policy, such as immigration quotas (cf. Hainmueller

and Hangartner 2013).

Third, studies that do measure voters' policy preferences emphasize national policy (e.g., Enos 2014; Hopkins 2010). While immigration policy influences national demographics, a voter's local exposure to demographic change in the United States is heavily shaped by their municipality's zoning. Elected officials often use zoning to maintain economic segregation and therefore racial segregation (Trounstine 2018). From the enshrinement of single-family zoning to the obstruction of affordable housing, the practice of "exclusionary zoning" helps establish a local price floor that prevents the immigration of lower-income residents into a municipality (Einstein, Glick and Palmer 2020; Rothwell and Massey 2009; Trounstine 2020). Taken to an extreme, residents can stymie integration via direct democracy, where housing proposals and land use policies are placed on the local ballot (e.g., Hankinson 2018). In short, local regulatory tools give voters substantial veto power over the construction of new housing, entrenching existing patterns of segregation.

Given the importance of measuring social context using the scale at which conflict occurs (Oliver and Mendelberg 2000; Oliver and Wong 2003), the effect of local demographic change on local behavior has been under-examined. Returning to the first limitation, studies which have looked at the relationship between local change and local exclusionary policy focus on institutional (i.e., citywide rather than individual) outcomes attributable to the preferences of white homeowners. For example, as southern Blacks moved into northern cities during the Great Migration, whites not only left central cities for the surrounding suburbs, but also used zoning to limit the ability of Black migrants to follow them (Sahn Forthcoming). But while increasing diversity is expected to have the largest effect in predominantly white communities (Newman 2013), far less attention has been paid to the racially diverse communities experiencing more recent demographic changes.

In this study, I address these three limitations by measuring the relationship between local racial demographic change and local exclusionary behaviors across an array of racially diverse communities. As an exclusionary outcome, I use a California state law requiring that

any proposal for new “low-rent” housing must win the support of a majority of voters via municipal-level ballot.<sup>1</sup> Known as “Article 34,” this policy transferred decision making over affordable housing from the city council to the local electorate. Recognizing that placing this veto power in the hands of voters made it more difficult to build affordable housing, the California legislature has tried repeatedly to repeal Article 34 via statewide ballot, with failed attempts in 1974, 1980, and 1993.<sup>2</sup> These repeal efforts provide an opportunity to observe voters’ behavioral support for an exclusionary tool: direct veto power over new affordable housing.

Using returns from 3,718 precincts across Los Angeles County in the 1993 repeal effort, I find that an increase in the local non-white population from 1980 to 1993 is associated with higher precinct-level support for this “voters’ veto” over affordable housing. I use two analytical strategies to assess this relationship. First, I show that a standard deviation increase in a precinct’s non-white population is correlated with a 4.5 percentage point ( $d = .32$ ) increase in support for the voters’ veto; similar results are found when defining change as a proportion. Additionally, I show that this correlation is not explained by alternative factors which may be associated with racial turnover, such as local changes in population density, residential churn, median household income, and percent foreign born.

Second, I find that this exclusionary response emerges in not only majority white, but also majority Black and majority Latino precincts. If anything, were one to dichotomize the relationship as being among majority white versus non-white communities, the response to a decline in the local white population is greater among non-white precincts, suggesting a concern about demographic change that extends beyond simple stories of racial threat. Furthermore, I find that the association is not clearly attributable to an increase in one racial/ethnic group, suggesting that the response is most clearly related to a decrease in the local white population.

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<sup>1</sup>Though initially targeted towards public housing, “low-rent housing” has been interpreted to include any housing in which 50% or more of the units are government-subsidized.

<sup>2</sup>See Section A for the full text of Article 34 and the 1993 repeal ballot proposition.

Third, to account for the possibility that composition changes (i.e., new voters in each precinct) are driving this change, I also measure the correlates of a precinct's spatial proximity to neighborhoods in which the non-white population increased most rapidly. Using only precincts which experienced both minimal residential churn and minimal changes in their own racial demographics over this period of time, a standard deviation increase in proximity to the nearest diversifying area is linked to higher support for the voters' veto (mean +3.5 percentage points,  $d = .24$ ). Again, I find this exclusionary response to nearby demographic change in both majority white and more racially diverse precincts.

Together, these findings advance our understanding of the political effects of local racial demographic change. Exclusionary reactions are not limited to the members of the majority group, but also extend to minority voters. Furthermore, I do not find evidence that this relationship is driven by intergroup competition, as majority Black and Latino precincts do not respond to a decline in the size of their own populations. Instead, an increase in the local non-white population may come with political and socioeconomic concerns shared across a wide swath of current residents. For instance, this mechanism aligns with documented efforts by middle-class majority Black neighborhoods to separate themselves from poorer Black communities and “fortify their neighborhoods against this encroachment” (Pattillo-McCoy 1999, p. 6).

While many factors contribute to exclusionary behavior, the effect of local demographic change matters for two reasons. First, the United States faces a deepening affordability crisis primarily driven political obstacles to housing construction (Glaeser and Gyourko 2018). As minorities increasingly move to suburban communities (Badger, Bui and Gebeloff 2019; Orfield and Luce 2013), these findings suggest that voters of different racial backgrounds will be more inclined to protect existing exclusionary zoning, hampering supply. Second, there is growing evidence that sizable shares of Latino, Asian, and even Black voters may be moving away from the Democratic Party (Fraga, Velez and West Forthcoming; Geiger and Reny 2024; Kao 2023; Sommer and Franco 2023), specifically in areas with high levels of

immigration (Cai and Fessenden 2020). Given the ability of local conditions to shape social identities and even partisanship (e.g., Cramer 2016; Larsen et al. 2019; Ternullo 2024), these findings may provide a local mechanism for understanding the national shift of many minority voters away from the Democratic Party and towards more exclusionary political movements.

## Direct Democracy as an Exclusionary Institution

In the United States, direct democracy has a long history of enabling voters to bypass legislative processes and make policy via ballot measures (Gerber 2011). Within the state and local context, direct democracy frequently targets housing and has been used to perpetuate segregation. In a 1964 statewide election, California voters in predominantly white cities near diversifying areas were more likely to support a ballot measure allowing property sellers, landlords, and real estate agents to racially discriminate (Reny and Newman 2018).

Once such *de jure* segregation was prohibited, voters began constructing economic barriers to integration (Trounstine 2018). By preventing the construction of more affordable multifamily housing, single-family zoning helps set a minimum price of residency. In turn, municipalities near center cities which experienced greater levels of mid-century Black migration were more likely to adopt single-family zoning (Sahn Forthcoming). This use of policy levers to economically “defend” one’s community is collectively known as “exclusionary zoning” (Einstein, Glick and Palmer 2020).

In this study, I focus on voter support for repealing one of these exclusionary tools: direct democratic control over affordable housing. While the institution appears race-neutral, the context of its implementation underscores the tension between racial demographic change and support for voters’ veto power. Following World War II, Congress passed the Housing Act of 1949, promising “a decent home and suitable living environment for every American family” by building 810,000 units in six years. Simultaneously, the Second Great Migration was accelerating, with ultimately more than 5 million Blacks leaving the Southeast and

moving to the Northeast, Midwest, and West Coast. Backlash to this progressive housing program was swift. Residents in the Northern California coastal city of Eureka spearheaded a statewide constitutional amendment which would allow voters to stop their local housing authority from developing low-income housing in their community (Cavin 2019; Varian 2022).

Seeking a megaphone, the Eureka movement partnered with the California Real Estate Association (CREA), a precursor to the modern California Association of Realtors and the largest real estate group in the country. The CREA paid for the campaign to pass the measure, pitching the amendment as essential to countering “minority pressure groups,” preserving white neighborhoods and therefore home values (Staff 2019). In contrast, the “pro” argument in the 1950 official voter guide did not explicitly oppose affordable housing, but instead elevated local democracy:

A “Yes” vote for this proposed constitutional amendment is a vote neither for nor against public housing. It is a vote for the future right to say “yes” or “no” when the community considers a public housing project...It is an expression of confidence in the community’s future and in the democratic process of government (*Voter Information Guide for 1950, General Election 1950*).

Passing with a narrow majority, Article 34 immediately began throttling the state’s supply of new affordable housing. By 1968, voters had turned down nearly half of the public housing that had been proposed — around 15,000 units (Cavin 2019). Many public housing agencies shelved projects rather than put them to a vote. Others attempted to ameliorate voters’ concerns about aesthetics and concentrated poverty. In 1968, the San Jose Housing Authority put forward a referendum in support of small duplexes and apartments of no more than four units scattered throughout the city (Cavin 2019). Still, the measure failed.

Meanwhile, the democratic appeal of Article 34 followed it beyond the ballot box. The amendment eventually arrived before the U.S. Supreme Court, where it was upheld in a 5-3 vote. Writing for the majority, the generally liberal Justice Hugo Black emphasized the democratic nature of the law, finding no evidence that the law was racially motivated: “This

procedure ensures that all the people of a community will have a voice in a decision which may lead to large expenditures of local revenues. It gives them a voice in decisions that will affect the future development of their own community” (*James v. Valtierra*, 402 U.S. 137 (1971)). At the justices’ private conference, Chief Justice Warren Burger allegedly scoffed at the plaintiff’s claims, framing their argument as a suggestion that “too much democracy violates the Equal Protection Clause” (Cavin 2019).

Having lost at the Supreme Court, efforts to repeal Article 34 returned to the ballot box. In 1980, the pro-Article 34 campaign focused on taxes and local democracy and again defeated the repeal effort. Facing a third repeal attempt in 1993, public commentary on the 1993 repeal effort connected Article 34 to the protection of community character and quality of life. In an interview with the *Los Angeles Times*, State Assembly Member Gebe Ferguson framed support for voters’ veto power over housing as both unrelated to race and shared by a diverse constituency (Martinez 1993):

I don’t think anyone can blame the general public for feeling that way. You have a drive-by shooting in Mission Viejo and then you tell [residents] you want to move in low-income people?... It’s not a matter of race or income either, because low-income black communities don’t want low-income housing built in their communities either, because of past experiences with that.

While Assembly Member Ferguson may have extrapolated when describing the preferences of “low-income black communities,” affordable housing has been historically stigmatized due to stereotypes about its occupants (Tighe 2010). If voters view direct democracy over housing as a way to exclude poor minority residents, then the veto power granted by Article 34 would be a valuable backstop against not just affordable housing, but local racial demographic change.

# Racial Demographic Change and Exclusionary Behavior

What shapes voter support for direct democracy? From a principle-based perspective, voters may believe that some decisions should be voted on directly, others should not. In this case, support for direct democracy should remain stable or change in a way orthogonal to the expected policy outcome itself. But voters are often only “weakly principled,” caring more about policy outcomes than the process itself (e.g., Bartels and Johnston 2020; Prothro and Grigg 1960). From a policy-based perspective, voters may be more likely to protect direct democracy if they believe that doing so would lead to the policy outcomes they prefer, compared to turning control over to an administrative or legislative body.

Building on a policy-based perspective, voter support for direct democracy may be based on their self-identified ingroup’s status in comparison to relevant outgroups. For example, a primary threat to the majority group’s power is demographic change, wherein a new population threatens the status quo power structure (Blalock 1967; Blumer 1958). Threat-induced support may be less a concern about the particular policies pursued by the outgroup, but rather a general anxiety over the loss of economic, cultural, and political power (e.g., Abrajano and Hajnal 2015; Thompson 2023). Absent specific conditions found to support positive inter-group contact and limit racial bias (Allport 1954), a sudden increase in the local non-white population is expected to elicit a desire among white voters to keep direct democracy as a means to ensure exclusionary policy.

To be clear, racial demographic change may occur at the national, regional, or local level. Each level may evoke an exclusionary response through a different mechanism. National demographic change may trigger feelings about identity and culture, as well as political power over federal policy. Regional change may drive media attention and stories about crime within one’s state or the nearest center city. Local demographic change may elicit greater personal self-interest via expected changes in one’s home value and neighborhood trajectory. While change at multiple levels may occur simultaneously, local demographic change should most closely drive exclusionary behavior via control over local housing policy

as in the case of Article 34.

## **Demographic Change in Los Angeles County, 1993**

The November 1993 vote to repeal Article 34 is an opportunity to assess the importance of local racial demographic change across a racially diverse array of communities. From 1980 to 1993, Los Angeles County — the focus of this study — saw relatively little growth in its Black population. Instead, the county experienced a large increase in Latino migration and a proportionally large increase in Asian migration mainly from outside of the United States. Simultaneously, Los Angeles County's non-Hispanic white population decreased from 53 to 38 percent of the population. From an individual's perspective, the median Los Angeles County voter experienced an 14 percentage point (54 percent) increase in the non-white population share of their precinct.

For local demographic change to trigger racial threat, voters must perceive their neighborhoods as changing (Hopkins 2009; Wong 2007). A 1992 survey of Los Angeles County residents found that 56 percent of respondents felt that their neighborhood was experiencing a change in ethnic composition (Bobo and Zubrinsky 1996). Furthermore, this influx of Latino and Asian residents was concentrated in areas which had been majority Black going back to the 1950s (Bergesen and Herman 1998). As a result, Black residents were the racial group most likely (78 percent) to agree that their neighborhood was racially changing (Bobo and Zubrinsky 1996).

This local racial demographic change may lead to exclusionary behaviors through two main pathways. First, while members of many racial groups tend to show a degree of in-group preference (Charles 2006; Farley et al. 1994; Krysan et al. 2009), there may also be a preference for having a sizable share of white neighbors due to concerns for their neighborhood's political status and socioeconomic trajectory. White constituents often receive better representation, responsiveness, and public goods provision from local elected officials (e.g., Hankinson and Magazinnik 2023; Schaffner, Rhodes and La Raja 2020). At the same

time, residents' opposition to demographic change may also come from who moves in as white residents leave. Middle-class minority communities may struggle to preserve socioeconomic status, especially in segregated contexts where they are spatially proximate to poorer, majority-minority neighborhoods. Pattillo-McCoy (1999, p. 6) writes:

Middle-income black families fill the residential gap between the neighborhoods that house middle-class whites and the neighborhoods where poor African Americans live. Unlike most whites, middle-class black families must contend with the crime, dilapidated housing, and social disorder in the deteriorating poor neighborhoods that continue to grow in their direction. Residents attempt to fortify their neighborhoods against this encroachment...

Thus, turnover from white to non-white neighbors may signal a risk of declining political influence, city services, and even home values.

As evidence of these concerns in Los Angeles circa the 1993 vote to repeal Article 34, Charles (2006)'s contemporaneous research used an experimental design wherein respondents were asked their willingness to move into stylized neighborhoods of varying racial composition. Known as the Los Angeles Study of Urban Inequality (LASUI), the survey was fielded from 1992 to 1993, making its conclusions especially relevant to the November 1993 election in Los Angeles County. Charles (2006, p. 183) writes: "Across racial groups, patterns of neighborhood racial composition preferences reveal a clear and consistent racial rank-ordering of out-groups as potential neighbors. Whites are always the most preferred out-group neighbors..." Charles also finds that this preference hierarchy even extends to neighborhood change among non-white residents: "Across racial groups, blacks are indisputably the least-desired neighbors. It is equally clear that, among non-whites, integration with whites is more favorable than integration with other non-whites" (Bobo et al. 2000, p. 193).

Second, voters may be as responsive to changing local demographics due to perceived competition among racial groups for space and resources (Bobo and Hutchings 1996; Gay

2006; McClain and Karnig 1990). Drawing from her research in Los Angeles County, Charles (2006, p. 164) notes: “Concerns about relative group position are also somewhat apparent among blacks, whose preferences for both Asian and same-race neighbors are negatively influenced by the belief that this group poses a competitive threat to economic opportunities and political power.” Contemporaneous reporting reflects this conflict. A *Los Angeles Times* article written months before the 1992 Rodney King riots captures the unease among Black, Latino, and Asian residents. Clifford (1991) writes:

Cultural collisions, often violent, occasionally fatal, are occurring every day. Hostilities between black residents and Korean shop-keepers, Latinos and blacks vying for jobs at Martin Luther King Jr./Drew Medical Center, interracial fighting at Lawndale high school, and repeated charges of police brutality against minorities — all of this is disturbing the city’s racial peace in a way that has some political analysts recalling Watts.

Additional reporting at the time highlights the competition felt by long-time Black residents towards the new arrivals. Noble (1995) notes:

Marilyn Thompson, 45, a [Black woman and] telephone company manager who lived in Los Angeles for nearly 20 years before moving to Atlanta last year, also said the doors of opportunity seemed open for Hispanic and Asian residents but shut for blacks in California. “It seems like you can come into California and have nothing and end up with everything,” she said of other immigrants to the state. But, she said, blacks “can’t seem to get ahead.”

How does this perceived racial hierarchy and competition over housing and jobs shape intergroup relations? Eighteen months prior to the 1993 election, four white Los Angeles police officers were acquitted in the beating of Rodney King, a Black man. Rioting began within hours, concentrated in South Central Los Angeles and Koreatown. Researchers argue

that the riots highlighted competition-fueled resentment between racial groups, with Black and Latino rioters targeting the Korean community. Black residents viewed Latinos as having taken over their community and competing in the labor market, whereas Latinos believed they were underrepresented politically (Johnson Jr and Farrell Jr 1992). Bergesen and Herman (1998) attribute the intensity of the riots to this desegregation and neighborhood succession, with areas of racial demographic change showing the highest rates of violence.

In short, while much of the existing literature associates racial demographic change with racial threat among white voters, local demographic change can also provoke exclusionary behaviors among non-white voters. On one hand, shared concerns may be grounded in what a shrinking white population means for a community's political representation, socioeconomic status, and access to public goods. On the other hand, the response may be driven by perceived competition with the arriving minority group, animating inter-racial group tensions. Through both pathways, increases in the local non-white population may drive support for the voters' veto over affordable housing among voters broadly.

## Internal Demographic Change

To measure the relationship between racial demographic change and support for the voters' veto, I use two analytical strategies. First, I assess the link between internal precinct demographic change from 1980 to 1993 and a precinct's support for repealing Article 34 in the 1993 November election. Second, I subset to precincts which experienced minimal residential churn and were racially stable from 1980 to 1993, then measure the connection between spatial proximity to the nearest rapidly diversifying Census tract and support for repealing the voters' veto. Results from both strategies are substantively the same. With these two approaches, I also show that the connection between local demographic change and exclusionary behaviors exists beyond just white neighborhoods, but extends to majority Black and Latino precincts. I begin with the internal precinct design.

## Data

I combine precinct-level election returns with tract-level Census data to generate 3,718 precinct-level observations in Los Angeles County. Compared to citywide returns, precinct-level measurement more accurately reflects the variation in residents' local experiences. The analysis is limited to Los Angeles County as that is the only county where I have been able to find the contemporaneous voter file required for the analysis. Common resources like the California Statewide Database lack precinct returns from such early elections. County-level archives lack digitized precinct boundaries and voter files from this period. Even Los Angeles does not have digitized precinct boundaries from early votes on the repeal of Article 34.

Given these constraints, I was lucky to find the 1992 Los Angeles voter file, with it being passed down to me across two separate research teams. Still, Los Angeles County provides a uniquely dense multiracial context which allows me to observe the correlates of demographic change on a diverse array of communities. Also, as of 1993, Los Angeles County contained over 9 million residents (29 percent of California's population), making this a substantively meaningful subset of the state's voters.

Combining precinct returns with Census data presented a challenge. Los Angeles County does not have a shapefile of 1993 precinct boundaries. Thus, I use Enos, Kaufman and Sands (2019)'s geocoded Los Angeles County voter file from 1992 which includes each voter's address and precinct. I overlay the voters on shapefiles of the 1980, 1990, and 2000 Census tracts (Manson et al. 2022), generating data for the total number of voters per tract and per tract-precinct sub-unit. These quantities allow me to estimate the share of each tract's population that can be attributed to each precinct. I use this ratio to allocate counts of other Census data, e.g., the number of non-white residents, the number of manufacturing employees, etc. After allocating these tract-level counts to each tract-precinct sub-unit, I sum the counts within each precinct. Using these precinct-level counts, I then calculate the percentages and values needed for the analysis.

**Dependent Variable** My dependent variable is based on the precinct-level vote share for the repeal of Article 34 as voted on in November 1993. Within the precinct data, the weighted mean vote share was 44.6% in favor of repeal. For conceptual clarity, I define *support for the voters' veto over housing* as the complement of support for repealing Article 34. Rather than contextualize the relationship as one's opposition to repealing the amendment, I can say that 55.4% of voters supported direct democratic control over new affordable housing proposed for their community.

Figure B-1 shows the distribution of support for this voters' veto across Los Angeles County. Support for the voters' veto is concentrated outside of the central, more urbanized areas of Los Angeles County. The spatial trend largely aligns with partisanship as found using voter file party registration. Precincts with higher rates of Republican registered voters were also more likely to support the voters' veto (Figure B-2).

**Independent Variables** The treatment is local racial demographic change in the lead-up to the 1993 election. Given the election occurred prior to the creation of the American Community Survey in 2005, the data are limited to the decennial Census. To estimate values for 1993, I linearly interpolate data from the 1990 and 2000 Censuses. I operationalize this change in two ways:

- Percentage point change ( $Pct_t - Pct_{(t-1)}$ )
- Proportion change ( $\frac{Pct_t - Pct_{(t-1)}}{Pct_{(t-1)}}$ )

For reference, the population-weighted median precinct experienced an 14 percentage point (54 percent) increase in non-white population from 1980 to 1993. It is theoretically unclear which definition better captures local demographic change, so results are reported for both (e.g., Hill, Hopkins and Huber 2019). However, the correlation between the two measures of demographic change is 0.65, suggesting that the treatments are capturing different aspects of a similar phenomenon. Section C discusses in detail the ways in which these two measures differ in the context of Los Angeles County.

Given this treatment is estimated by aggregating tracts into precincts, there is inevitable measurement error in operationalizing demographic change from 1980 to 1993. Even more, because the treatment is derived from two precinct-level estimates (1980 and 1993), the two opportunities for measurement error may compound each other, producing outlier swings in demographic change which do not reflect reality. While these errors are not expected to introduce bias, I account for them by presenting all results using winsorized treatments — truncating extreme values at both ends to their 5th and 95th percentile, respectively. The results using non-winsorized data are substantively the same, statistically significant, and reported in Table H-9.

## Empirical Strategy

I regress precinct-level vote share in support of the voters' veto on the 1980-1993 change in non-white population at the precinct level. For comparability, I operationalize both treatments as a standard deviation increase in the precinct's non-white population. The identifying assumption of this strategy is selection on observables. The selection of this design is driven by the granularity of the geographic scale. In contrast, the commonly used "shift-share" instrument requires Census data on migration patterns (Borusyak, Hull and Jaravel 2025), which are only available at the county level. Consequently, the method is not viable for variation within county.

I selected theoretically motivated control variables to account for other changes which may confound the relationship between an increase in the local non-white population and increasing support for the voters' veto over affordable housing. A spurious relationship between the two may occur if the areas where voters are concerned about social and economic conditions in 1993 were also more likely to attract non-white residents from 1980 to 1993. For example, economic traits may predict both economic anxiety and an influx of immigration. Thus, I control for the precinct-level unemployment rate, vacancy rate, and household median income in 1980. Other observable traits may be associated with residents' ability to access

local government resources and organize to stabilize their neighborhood, such as the percent of residents who are homeowners, the local residential density, and the percent of residents in 1980 who were white.

To review, all models include controls for each precinct's 1980 pretreatment percent non-Hispanic white, homeownership rate, vacancy rate, log median household income, population density, and unemployment rate. In addition, Table D-1 presents the regression results including the change in each control from 1980 to 1993. While this approach accounts for how precincts were changing over time, it does risk biasing the result of demographic change as the treatment because as some of these changes may be inseparable from the demographic change. For instance, if new non-white residents tend to have lower median household incomes, then controlling for a decrease in local household income would bias the result of an increase in the local non-white population (Sen and Wasow 2016). Nevertheless, results including time-varying controls are substantively identical and statistically significant.<sup>3</sup>

I also condition for time-invariant place attributes by using a Census place-level fixed effect and weight the data by precinct population. As shown in Figures C-4 and C-5, local racial demographic change across precincts is non-random. Non-white residents may be more likely to select Census places to live, rather than selecting specific precincts. To account for this correlation across observations, I cluster Huber-White standard errors at the Census place-level.

## Results

Table 1 presents the relationship between a standard deviation increase in the non-white population and support for the voters' veto over affordable housing.<sup>4</sup> Results for a percentage point change are shown in Models 1 and results from a proportion change are in Models 2. A

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<sup>3</sup>I do not include a control for partisanship, as I do not have pretreatment, precinct-level measures of partisanship. Because precinct boundaries change each election, I cannot use 1980 electoral data without a precinct boundary shapefile.

<sup>4</sup>Figure F-8 visualizes the bivariate relationship between the two treatment variables and precinct support for the voters' veto using the raw data.

	Support for voters' veto, 1993	
	Model 1	Model 2
Pct. point $\Delta$ in non-white pop.	0.045*** (0.011)	
Prop. $\Delta$ in non-white pop.		0.034*** (0.010)
Pct. white, non-Hispanic (1980)	0.067 (0.035)	0.055 (0.040)
Pct. unemployed (1980)	-0.517*** (0.143)	-0.670** (0.223)
Pct. homeowner (1980)	0.137*** (0.039)	0.123** (0.038)
Pct. vacant (1980)	-0.303** (0.113)	-0.274* (0.113)
Log median household income (1980)	0.010 (0.014)	-0.004 (0.015)
Pop. density (1980)	-6.040*** (0.528)	-5.367*** (0.783)
City FE	Yes	Yes
R <sup>2</sup>	0.751	0.723
Adj. R <sup>2</sup>	0.740	0.711
Num. obs.	3651	3651
RMSE	2.556	2.694
N Clusters	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table 1: Effect of change in a precinct's non-white population (1980-1993) on support for voters' veto, 1993. Treatment data are winsorized.

standard deviation increase in the percent non-white population (+10.4 percentage points) is associated with a 4.5 point ( $d = .32$ ) increase in support for direct democracy over housing. A standard deviation increase in the non-white population as a proportion change (+54.2 percent) is associated with a 3.4 point ( $d = .24$ ) increase in support for the voters' veto.

While I cluster standard errors at the place-level to account for correlated treatment within places, there may be other omitted variables that contribute to the error term and are spatially correlated. Likewise, this study's proximity-based analysis (below) is premised on the belief that precinct outcomes are affected by nearby Census tracts. Both conditions violate SUTVA. As a robustness check, I use a spatial Durbin error model (SDEM) to account for potential spillover effects from neighboring precincts which may be affecting the relationship, as well as spatial autocorrelation among standard errors. The results are statistically significant and substantively similar as the non-spatial OLS models (Section E).

As further evidence of the importance of racial demographic change, I replicate Table 1 with variables capturing alternative explanations that may be correlated with racial turnover. Tables G-5 and G-6 show that support for the voters' veto is related to neither precinct-level changes in population density nor residential churn, defined as the share of the population that moved to their unit in the past 10 years. While Table G-7 shows that a decrease in the local household median income is associated with an increase in support for the voters' veto, the magnitude of the relationship is roughly half that of those from racial demographic change. Lastly, much of California's increasing diversity in the 1980's came from immigration, with undocumented immigrants alone representing 22 to 31 percent of all migrants to CA during this period (Johnson 1996). However, Table G-8 does not detect a consistent relationship from an increase in the percent foreign born within a precinct.

Two alternative explanations may also be correlated with racial turnover but are unable to be tested using my framework. First, it is possible that the observed correlation is driven by changes in the local crime rate. Unfortunately, commonly used crime databases are either limited to the city level and above (FBI Uniform Crime Reporting program) or do

not have neighborhood-level estimates for years prior to 2000 (the National Neighborhood Crime Study). Thus, crime remains a potential alternative explanation that researchers with more recent outcome data should directly explore. Second, concerns over redistribution and local fiscal policy may also drive an unwillingness to allow local affordable housing (Peterson 1981). But this mechanism is unlikely to explain my results, as variation in concern over fiscal policy would have to be within municipality, given the use of place fixed effects.

## Mechanisms by Racial Groups

Much of the literature on racial demographic change has emphasized the sensitivity of white voters to an increase in the local non-white population. To assess whether this relationship generalizes to majority Black and majority Latino communities, I adapt the interaction model to compare precincts with a majority of each group to precincts with a minority of the same racial group. For context, as of 1980, there were 401 majority Latino precincts and 372 majority Black precincts. However, because there were only 8 majority Asian precincts in 1980, I do not have the statistical power to estimate a relationship among majority Asian precincts.

As shown in Figure 1, majority Black and Latino precincts show the same increase in support for the voters' veto as do majority non-white precincts as whole. In contrast, the correlation in majority white precincts — while positive — is generally smaller than that of Black and Latino communities.

Additionally, I estimate the effect of an increase in each minority racial group to see whether the relationship is driven by a specific minority group. This approach faces empirical challenges though. An increase in one racial group may happen in conjunction with that of another group. Thus, isolating the effect of one racial group risks misrepresenting the overall treatment of a shrinking white population in these precincts.

To account for these multiple changing racial groups, I include the percentage point and proportion treatments for each minority group (non-Hispanic Black, Hispanic, Asian, and

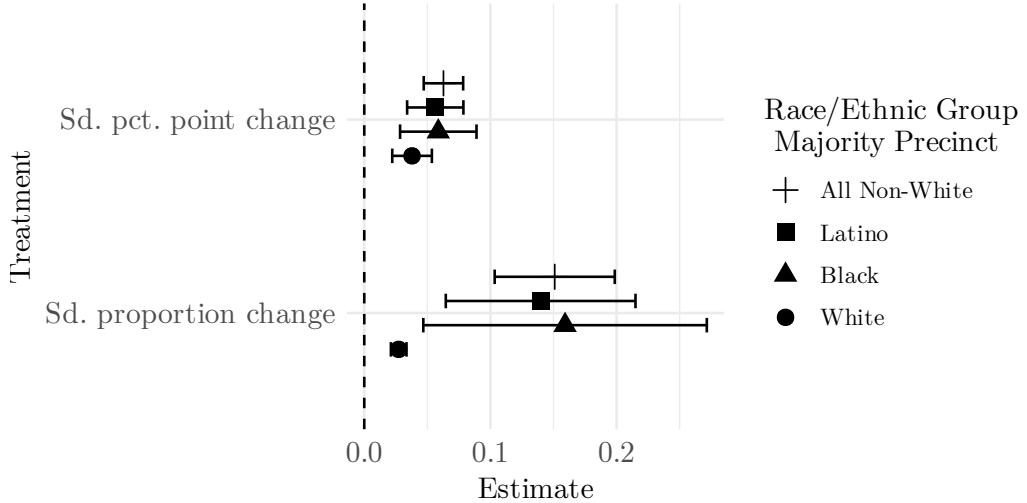


Figure 1: Variation in the effect of local racial demographic change in precincts with majority non-white, majority Black, majority Latino, and majority white residents pretreatment (1980). Point estimates shown with 95%-confidence intervals.

Other). I also include lagged levels of each racial group to help account for each precinct’s pretreatment composition. Finally, because each of these groups are different sizes, I no longer define treatment as a standard deviation increase. Here, the percentage point change model is based on moving from 0 percent to 100 percent of that ethnic group, whereas the proportion model is defined as a 100 percent increase in group size.

As shown in Figure 2, a percentage point increase in the population of any of the major non-white racial groups is associated with greater support for the voters veto. In contrast, the association with a proportion change seems to be driven by a proportion increase in the Latino population. Recall, the main mechanism of this study is a decrease in the local white population. While the declining white population may still be the main driver, it may also be the case that precincts losing white residents were experiencing the largest proportional increases in Latino residents. Still, it is difficult to identify the extent to which this is an empirical anomaly due to isolating fractions of the overall demographic turnover (an increase in the local non-white population share).

Finally, to test whether the main relationship is driven by either a concern over losing white residents or a decrease in one’s own racial group — which may stimulate group threat

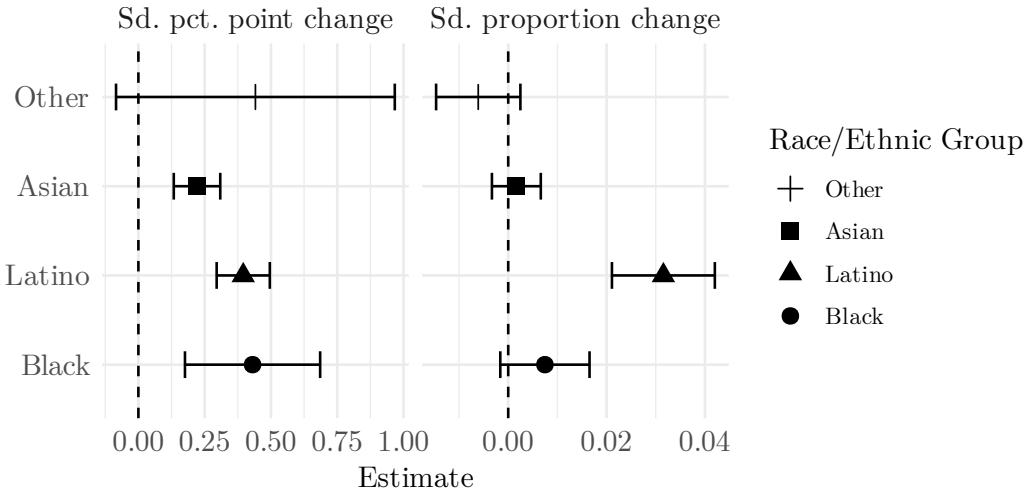


Figure 2: Association of increases by different racial/ethnic groups across all precincts. Point estimates shown with 95%-confidence intervals.

among minorities — I test whether majority Black and Latino precincts show a similar increase in support for the voters’ veto when their own racial group decreases in size. The percentage point change model is based on moving from 0 percent to 100 percent of that ethnic group, whereas the proportion model is defined as a 100 percent increase in group size.

Isolating these majority Black and majority Latino precincts, I find that a decrease in each group’s own share of the population is not associated with support for the voters’ veto (Tables D-2 and D-3). Only a decrease in the local white population spurs support for local control over affordable housing. These subgroup analyses suggest voters from an array of racial/ethnic backgrounds may feel concerns over the economic, social, and political trajectory of their neighborhoods as the local white population shrinks, rather than concerns over the size and influence of their own racial group.

## Proximity to Demographic Change

One threat to the above analysis is the possibility that voters within a diversifying precinct were not changing their behavior, but rather the composition of the precinct’s electorate was

changing. In other words, the observed increased support for the voters' veto may be coming from new voters arriving in these precincts.

This alternative explanation is unlikely for two reasons. First, precincts with more non-white voters on average show lower support for the voters' veto (Figure B-3). As a result, it is improbable that the newly arrived non-white residents would be *more* likely to vote in favor of the voters' veto compared to long-time white residents. Second, for residential churn to increase support for the voters' veto, those leaving the precinct would have to be *less* likely to support direct democracy over housing than those staying behind. This is also improbable. Exit in response to local demographic change is more likely among the least racially tolerant (Clark 1991) — those who are the most supportive of the voters' veto. If anything, I would expect to see less support for the voters' veto in diversifying areas if the least racially tolerant were being replaced with new arrivals.

Another way to account for the possibility that these associations are driven solely by residential turnover is to use an identification strategy which does not rely on racial change *within* each precinct as the treatment. Instead, I measure exposure to demographic change via the proximity of each precinct to a rapidly diversifying Census tract (e.g., Reny and Newman 2018). Across a variety of specifications, I find that racially stable, low residential churn precincts closer to diversifying neighborhoods were more supportive of the voters' veto over affordable housing compared to those farther away. This finding suggests that residential turnover is unlikely to be responsible for the results observed in the previous section.

## Data

Using the same precinct data as described above, I calculate each precinct's proximity to a “diversifying tract,” a Census tract which experienced an extreme increase in its non-white population share.

**Dependent Variable** The dependent variable remains precinct-level vote share in support of the voters' veto over affordable housing in 1993.

**Independent Variable** The treatment is the proximity of each precinct to a Census tract which experienced an extreme increase in non-white population. Defining these diversifying tracts is challenging: too low a threshold and the change may not be noticeable, too high and the test loses statistical power. For my main specification, I define diversifying tracts as those exceeding the 90th percentile of the increase in the non-white population from 1980 to 1993. For the percentage point treatment, this is a >27 percentage point increase in non-white population. For the proportion change, this a >127 percent increase in the non-white population. Results are substantively similar using cutpoints at the 85th and 95th percentiles (See Tables I-10 to I-13). I then measure proximity as the distance between the centroid of each precinct and its nearest diversifying tract.

One challenge of using Census tracts to identify diversifying areas is that they do not perfectly match known neighborhood boundaries, meaning voters' perceptions of change may not match the administrative boundaries of Census tracts (Wong 2007). On the other hand, Census demographers are careful to construct tract boundaries based on known landmarks, such as rivers, rail lines, and major streets. These boundaries tend to be very stable over time and are the most widely used unit of observation in quantitative studies of neighborhood racial change (Lee et al. 2008).

Still, some tracts have few residents, meaning a large percent change in the non-white population may be neither substantively meaningful nor perceptible to those nearby. After defining the percentile cutpoints for demographic change, I subset to tracts outside of the bottom tercile in population ( $>3,136$  residents).<sup>5</sup> In all, the set of diversifying tracts are those which experienced large increases in their non-white population share from 1980-1993 and were also large enough that their demographic change should register in voters' minds.

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<sup>5</sup>As theoretically expected, results are smaller, yet still statistically significant when removing the threshold and keeping even low population tracts (Table I-14).

Figure C-7 shows the distribution of these tracts using both the percentage point change and proportion change definitions.

## Empirical Strategy

In this design, the treatment is a precinct's proximity to a diversifying Census tract — one which diversified significantly from 1980 to 1993. As in the previous models, I control for the same pretreatment (1980) precinct-level covariates that may confound the relationship, including percent non-Hispanic white, homeownership rate, log median household income, vacancy rate, unemployment rate, and population density. Because treatment is based on proximity to a diversifying tract, I cluster Huber-White standard errors at the level of the nearest diversifying tract.

## Results

I transform “distance away from” into the more analytically useful “proximity to” by multiplying distance by -1. I then standardize proximity so that the treatment is a standard deviation increase in proximity to a diversifying Census tract, centered at the mean. For context, a standard deviation increase in proximity is approximately 3.5 kilometers (2.2 miles). I regress support for the voters’ veto in 1993 on proximity to a diversifying tract, weighting by population.

The purpose of this analysis is to test whether precincts which experienced little change in their demographics from 1980 to 1993 were responsive to demographic changes around them. If so, that would suggest that the within-precinct results are unlikely to be driven by residential turnover. To test this, I subset to precincts which are in the bottom tercile of both residential churn and change in their own non-white population from 1980 to 1993.

Table 2 shows the effect of proximity to a diversifying tract among precincts in the bottom tercile of demographic change. For Models 1 and 2, these are precincts where the non-white population increased by less than 6.5 percentage points (median change of +3 percentage

	Percentage Point $\Delta$		Proportion $\Delta$	
	Model 1	Model 2	Model 3	Model 4
Prox. to pct point. $\Delta$ in non-white	0.030*** (0.008)		0.018** (0.006)	
Prox. to prop. $\Delta$ in non-white		0.052*** (0.008)		0.040*** (0.007)
Pct. white, non-Hispanic (1980)	0.196*** (0.049)	0.171*** (0.038)	0.149** (0.047)	0.148** (0.045)
Pct. unemployed (1980)	-0.606 (0.345)	-1.010*** (0.265)	-0.799* (0.315)	-1.109*** (0.261)
Pct. homeowner (1980)	0.033 (0.057)	0.041 (0.051)	-0.018 (0.053)	-0.000 (0.052)
Pct. vacant (1980)	-0.721 (0.529)	-0.849 (0.431)	-0.922* (0.347)	-1.052** (0.312)
Log median household income (1980)	-0.072 (0.040)	-0.095** (0.028)	-0.046 (0.034)	-0.079* (0.030)
Pop. density (1980)	-5.169 (10.309)	0.599 (8.102)	-7.494 (7.052)	-4.536 (5.747)
Place Fixed Effects	Yes	Yes	Yes	Yes
R <sup>2</sup>	0.597	0.665	0.595	0.643
Adj. R <sup>2</sup>	0.552	0.628	0.545	0.598
Num. obs.	403	403	373	373
RMSE	2.333	2.125	2.434	2.286
N Clusters	37	37	33	33

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table 2: Effect of proximity to diversifying tract on support for voter's veto, among precincts which changed minimally from 1980 to 1993.

points non-white). For Models 3 and 4, the non-white population increased by less than 27 percent (median change of +4 percent non-white). I again vary whether the nearby diversifying tracts are defined based on percentage point (Models 1 and 3) or proportion changes (Models 2 and 4). All models include the same pretreatment covariates as well as Census place fixed effects.

Within precincts which saw little demographic change themselves, areas which were closer to diversifying Census tracts were on average 3.5 points ( $d = .24$ ) more supportive of the voters' veto over housing compared to those farther away. These results appear in both predominantly white and non-white precincts, though they are noisier due to the limited sample size (Table J-15). Unfortunately, the sample size also means that independent analyses on majority Latino, Black, or Asian precincts are not feasible with this design.

## Discussion

When it comes to neighborhood change, researchers often categorize the interaction of racial groups as either contact or conflict. Oliver and Wong (2003) find that racial minorities living near each other tend to hold lower levels of outgroup prejudice. At the same time, Enos (2017) finds that the salience of group identities is a function of groups' relative size, spatial proximity, and the degree of segregation. This salience may lead to conflict absent the conditions for positive contact outlined in Allport (1954).

My findings do not directly emphasize contact or conflict. Rather they propose that experiencing local racial demographic change may activate broadly shared concerns for maintaining political power and socioeconomic well-being that are attributed to racial turnover in general. A desire to protect the voters' veto is not the same as heightened racial prejudice. As a result, I would not expect to find a similar relationship between local demographic change and higher support for racially-related statewide ballot measures, such as public service access for undocumented immigrants (e.g., Campbell, Wong and Citrin 2006). But even

if not driven by general racial threat, residents' support for local exclusionary policy can manifest in similar policy outcomes, entrenching existing patterns of segregation.

From a policy perspective, these findings are timely. Los Angeles County of 1993 was a bellwether for California as a whole. In the 1990 Census, Los Angeles County was 41% non-Hispanic white, 11% non-Hispanic Black, and 10% Asian; 38% of residents identified as Latino. As of 2020, California as a state is nearly identical: 41% non-Hispanic white, 6% non-Hispanic Black, and 15% Asian; 39% of Californians identify as Latino. The housing costs and competition that Charles (2006) documented in Los Angeles have now spread across the state, with California having the highest housing costs in the nation.

Yet despite these high costs, California continues to diversify. From 2010 to 2020, California's white population decreased by 24 percent, the largest statewide decrease. Coincidentally, from 1980 to 1990, Los Angeles County's white population also decreased by 23 percent. These parallel paths may inform politics today. In November 2024, Californians were scheduled to vote again on repealing Article 34. However, the state's Democratic Party leaders pulled the ballot measure, under the guise of too many competing initiatives on the upcoming ballot. In reality, polling found the measure underwater in terms of public support (Board 2024). Given the similarities in demographic trends leading up to 1993 and 2024, advocates may have been wise to pull the measure and avoid a rebuke of the state's efforts to limit local control over affordable housing.

This strategic awareness around state preemption and limiting the voters' veto is important. Be it via ballot box or legislature, local control of land use has been a tool of segregation. Hopes to loosen local control to address both racial segregation and the supply-side of the housing affordability crisis hang on not only weakening the voters' veto, but on scaling up control to the state level. But these findings should give us pause: local demographic change is associated with greater support for local democracy in housing. And while the California state government struggles to repeal Article 34, a countermovement — “Our Neighborhood Voices” — is organizing for a competing initiative that would effectively end

state preemption of local land use policy (Brasuell 2022). With California continuing to experience demographic change, voters may seize the opportunity to solidify local control over housing and thus integration.

## Conclusion

As Western democracies diversify, our understanding of the political consequences of racial demographic change has been limited in three ways. First, research has focused heavily on the behaviors and attitudes of the majority population, emphasizing racial threat as the dominant mechanism behind observed relationships. Second, our understanding of exclusionary reactions to racial change has been most commonly observed via survey attitudes or indirect behavioral outcomes, such as voter turnout. Third, researchers have emphasized support for national policy outcomes like restrictive immigration measures, rather than the exclusionary tools available where voters may be most directly exposed to demographic change: their local context.

This study is a direct response to these three shortcomings. The attempted repeal of Article 34 is a rare case of observing voters' behavioral support for exclusionary policy, in this case the voters' veto power over the historically racialized policy of affordable housing. Using precinct-level returns and multiple analytical approaches, I show that local exposure to an increasing non-white population is associated with greater support for voter control over affordable housing. For these voters, direct democracy over affordable housing appears to be a means of defense; "fortifying" their communities in the face of potential political, economic, and social changes.

Furthermore, this exclusionary behavior was not only expressed in majority white precincts, but also in precincts with majority Black, Latino, and multiracial populations. I also show that the relationship was not clearly driven by an increase in one minority racial group nor a decline in a minority racial group's own status. Instead, support for the voter's veto is most

clearly tied to a decline in the local white population. That this impulse was shared across a range of racial contexts suggests that the effects of growing diversity in Western democracies extend beyond white racial threat. Instead, broad swaths of voters may associate an exiting of the majority population with concerns for their neighborhood's social, economic, and political trajectory.

Currently, Americans harboring exclusionary attitudes — either on national immigration policy or local zoning practices — are more likely to find their policy preferences embraced by the Republican Party. As the United States diversifies, there is growing evidence that many Latino, Asian, and to a lesser extent Black voters may be to moving away from the Democratic Party, especially in areas with high levels of immigration. The exclusionary behaviors found in this study align with and may provide a local foundation for this national shift of non-white voters towards the Republican Party. More broadly, in the context of the nation's housing affordability crisis, these findings underscore the tension between integration and local democratic control over housing in a diversifying society.

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# Online Appendix for “The Voters’ Veto: Local Racial Demographic Change and Exclusionary Behavior”

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## **A Article 34 Background**

Article 34 of the California Constitution:

"No low rent housing project shall hereafter be developed, constructed, or acquired in any manner by any state public body until, a majority of the qualified electors of the city, town or county, as the case may be, in which it is proposed to develop, construct, or acquire the same, voting upon such issue, approve such project by voting in favor thereof at an election to be held for that purpose, or at any general or special election."

1993 ballot measure to repeal Article 34 (Proposition 168):

"LOW RENT HOUSING PROJECTS. LEGISLATIVE CONSTITUTIONAL AMENDMENT. Amends state constitutional definition of low rent housing projects to include only projects owned by a governmental entity as defined. Excludes projects found to have no significant negative impact on the revenues of the affected governmental entity, and whose physical appearance is found to have no significant negative impact on the surrounding community. Requires approval by voters only upon qualification of ballot petition as specified. Exempts projects approved on or before November 3, 1992, or projects with existing contracts for federal financial assistance."

## **B Descriptive Relationships**

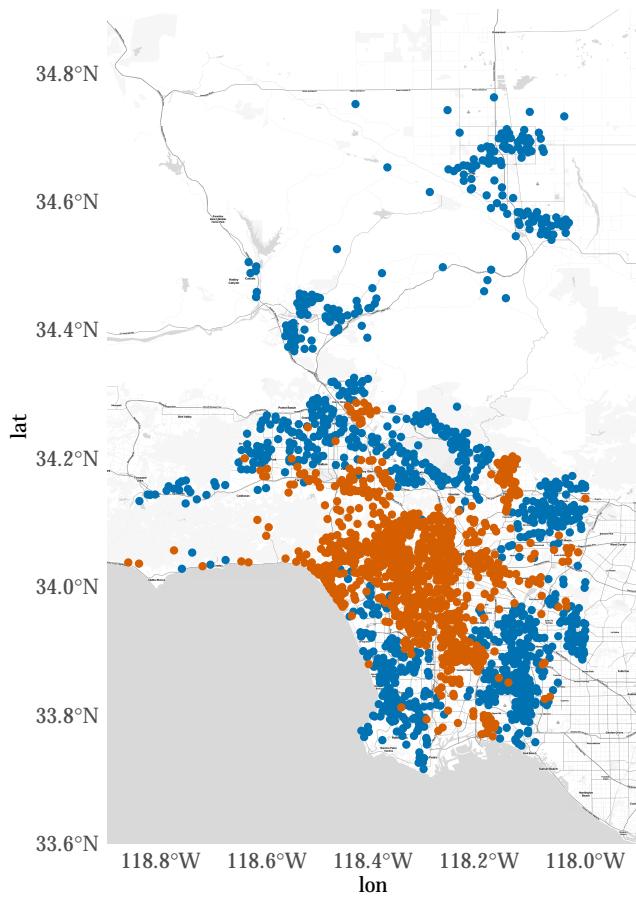


Figure B-1: Distribution of support for the voters' veto using precinct centroids. Blue as the top tercile of support ( $\geq 62\%$  in favor); orange as the bottom tercile of support ( $\leq 47\%$  in favor). The middle tercile is omitted for visualization purposes only.

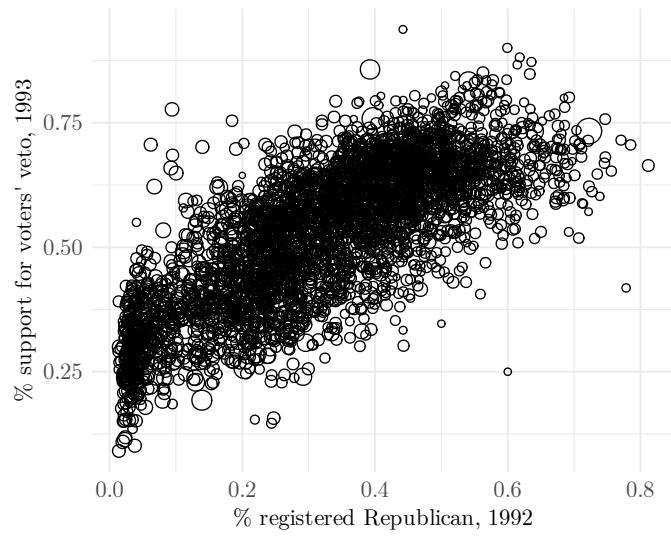


Figure B-2: Relationship between precinct-level Republican registration in 1992 and support for Article 34 repeal in 1993.

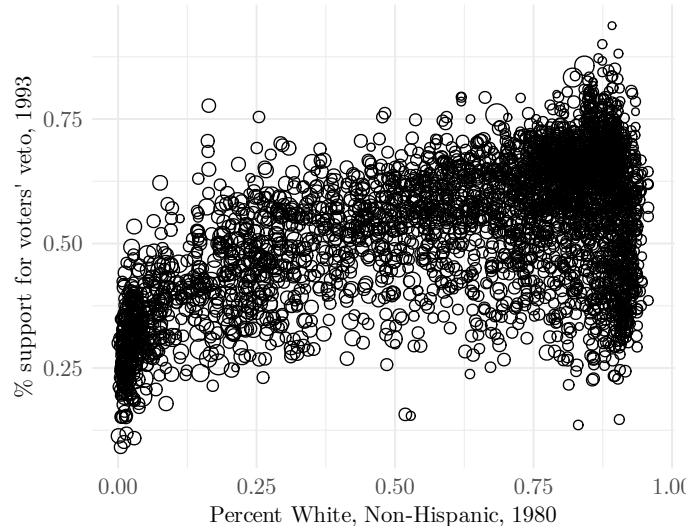


Figure B-3: Relationship between precinct-level percent white, non-Hispanic (1980) and support for Article 34 repeal (1993).

## C Defining Treatment

Figures C-4 and C-5 show maps of how precincts changed from 1980 to 1993. Green precincts experienced the smallest increase in their non-white population (bottom tercile), pink precincts experienced the greatest increase in their non-white population (top tercile), with the middle tercile omitted. Q3 precincts (pink) of both measures of racial change generally cover comparable areas, though with some notable exceptions around the Santa Monica Mountains in the upper left quadrant of each map.

To understand how areas of intense demographic change may differ, Figure C-6 shows the relationship between the 1980 percent non-Hispanic white population (i.e., pretreatment white population) on the x-axis and each treatment measure on the y-axis. In very white areas, large changes in percentage point population are unlikely, e.g., 98 percent white precincts cannot experience 20 percentage point increases in non-white population. However, a large proportion change is common in very white areas. A 98 percent white precinct could easily experience a 4 percentage point increase in the non-white population, equaling a 200 percent change in the non-white population. In short, most of the variation in the percentage point change in population occurs in areas between 50 and 80 percent white, whereas most of the variation in proportion change is in areas that are greater than 80 percent white circa 1980.

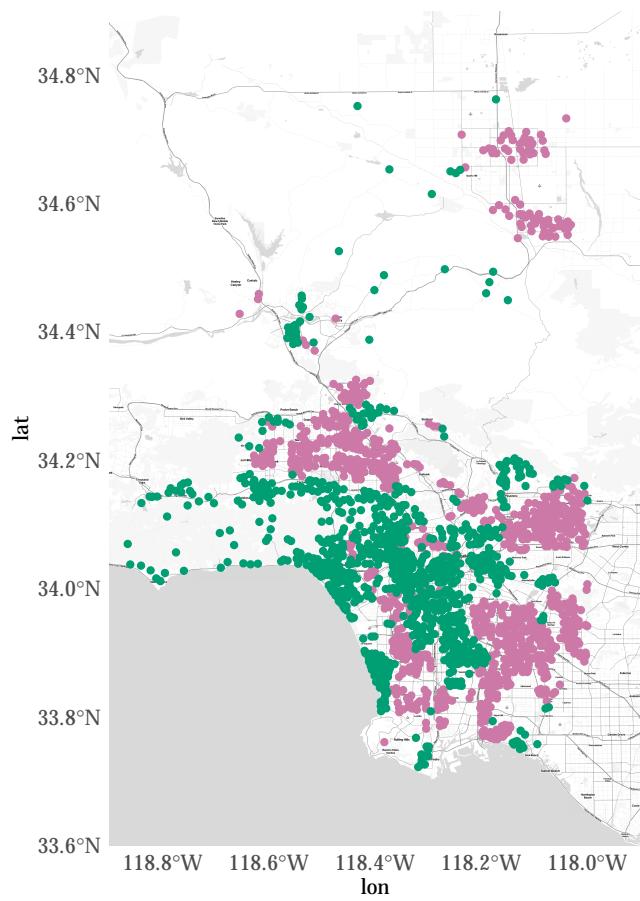


Figure C-4: Distribution of increase in non-white population as a percentage point change, from 1980 to 1993 shown using precinct centroids. Green as bottom quartile of racial change; pink as top tercile of racial change. Middle tercile omitted.

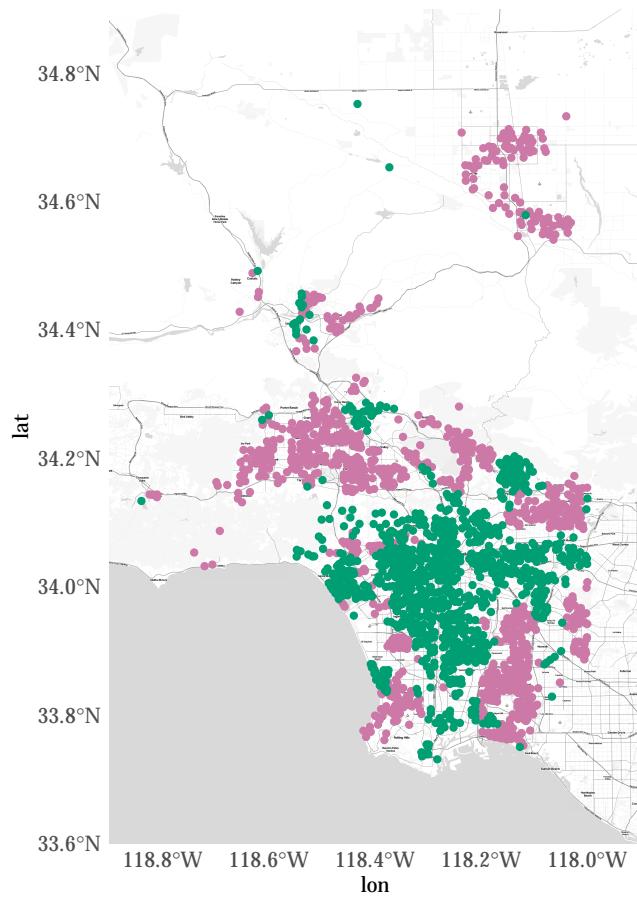


Figure C-5: Distribution of increase in non-white population as a proportion change, from 1980 to 1993 shown using precinct centroids. Green as bottom quartile of racial change; pink as top tercile of racial change. Middle tercile omitted.

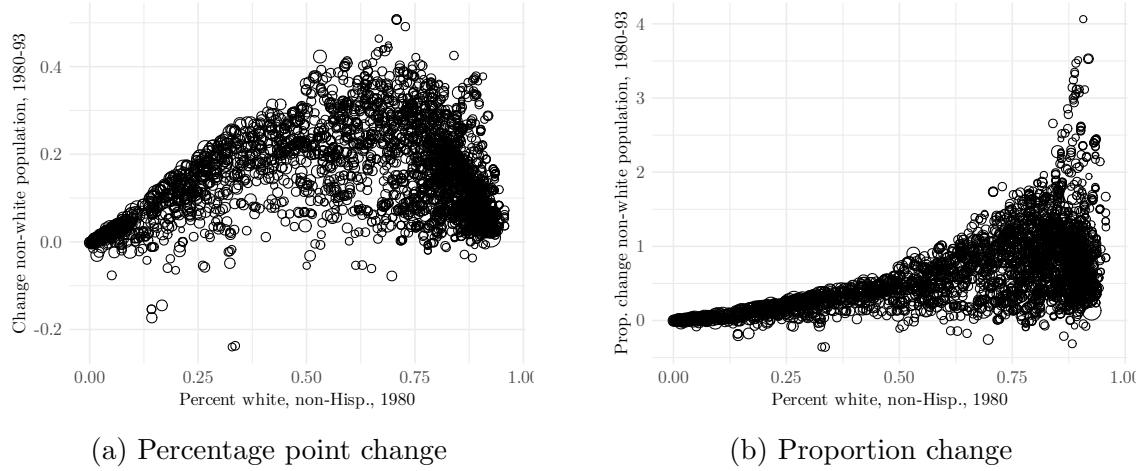
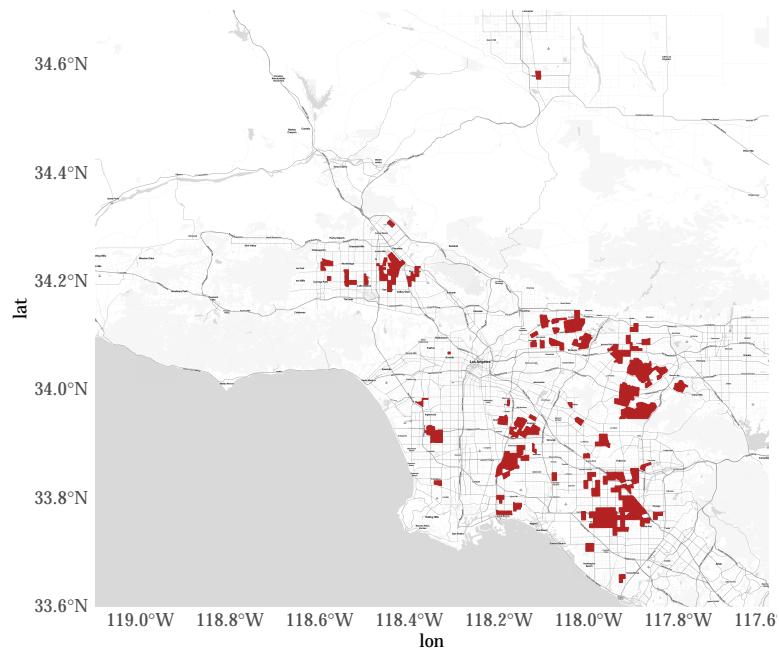
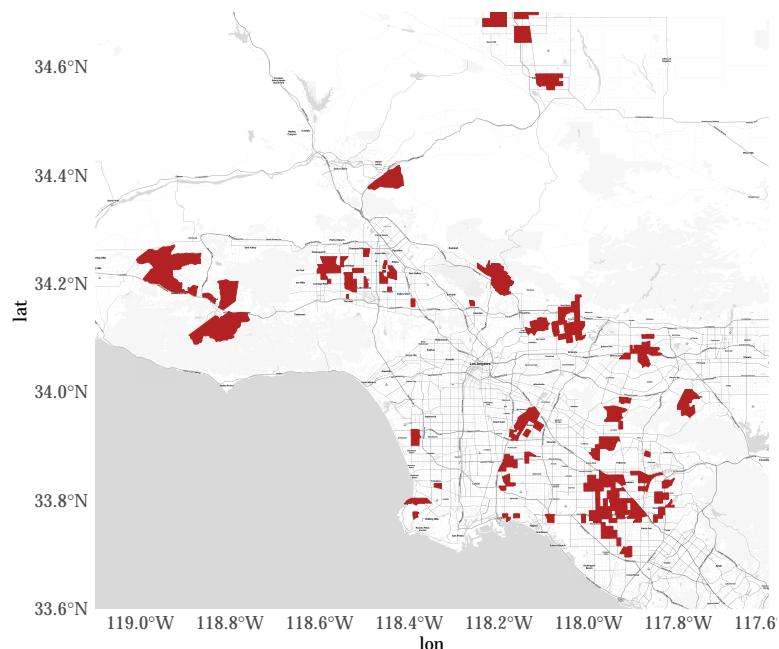


Figure C-6: Variation in the two treatment variables based on a precinct-level percent white, non-Hispanic population in 1980 (pretreatment).



(a) Percentage point change



(b) Proportion change

Figure C-7: Census tracts above the 90th percentile for their increase in non-white population share, 1980-1993. Distance to the nearest of these tracts is used as the treatment for the proximity analysis.

## D Within Precinct Results

	Support for voters' veto, 1993			
	Model 1	Model 2	Model 3	Model 4
Pct. point $\Delta$ in non-white pop.	0.045*** (0.011)	0.044*** (0.009)		
Prop. $\Delta$ in non-white pop.			0.034*** (0.010)	0.028*** (0.006)
Pct. white, non-Hispanic (1980)	0.067 (0.035)	0.073* (0.030)	0.055 (0.040)	0.085** (0.028)
Pct. unemployed (1980)	-0.517*** (0.143)	-0.784*** (0.148)	-0.670** (0.223)	-0.937*** (0.243)
Pct. homeowner (1980)	0.137*** (0.039)	0.171*** (0.035)	0.123** (0.038)	0.155*** (0.035)
Pct. vacant (1980)	-0.303** (0.113)	-0.764*** (0.074)	-0.274* (0.113)	-0.831*** (0.129)
Log median household income (1980)	0.010 (0.014)	-0.045** (0.015)	-0.004 (0.015)	-0.072*** (0.020)
Pop. density (1980)	-6.040*** (0.528)	-6.024*** (0.557)	-5.367*** (0.783)	-6.011*** (0.512)
P.p. $\Delta$ unemployed (1980)		-0.420*** (0.054)		-0.404*** (0.070)
P.p. $\Delta$ homeowner (1980)		0.143*** (0.028)		0.113*** (0.025)
P.p. $\Delta$ vacant (1980)		-0.635*** (0.063)		-0.778*** (0.102)
Prop. $\Delta$ med. household income (1980)		-0.021** (0.008)		-0.040** (0.014)
City FE	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
Time-varying Controls	No	Yes	No	Yes
R <sup>2</sup>	0.751	0.763	0.723	0.737
Adj. R <sup>2</sup>	0.740	0.753	0.711	0.726
Num. obs.	3651	3649	3651	3649
RMSE	2.556	2.495	2.694	2.627
N Clusters	142	142	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table D-1: Effect of change in a precinct's non-white population (1980-1993) on support for voters' veto, 1993. Treatment data are winsorized.

## D.1 Intergroup Competition

To test whether the main relationship is driven by either a concern over losing white residents or rather a decrease in one's own racial group which may stimulate group threat among minorities, I test whether majority Black and Latino precincts show a similar increase in support for the voters' veto when their own racial group decreases in size. These coefficients are also not standardized. The percentage point change model is based on moving from 0 percent to 100 percent of that ethnic group, whereas the proportion model is defined as a 100 percent increase in group size.

Isolating these majority Black and majority Latino precincts, I find that a rise in the non-Black and non-Latino populations, respectively does not elicit an exclusionary reaction (Tables D-2 and D-3). Only a decrease in the local white population spurs support for local control over affordable housing. These subgroup analyses suggest that voters from an array of racial and ethnic backgrounds may feel concerns over the economic, social, and political trajectory of their neighborhoods as the local white population shrinks, rather than concerns over the size and influence of their own racial group.

	Support for voters' veto, 1993	
	Model 1	Model 2
Pct. point $\Delta$ in non-Black pop.	-0.028 (0.062)	
Prop. $\Delta$ in non-Black pop.		-0.043 (0.065)
Pct. unemployed (1980)	-0.272*** (0.047)	-0.258*** (0.049)
Pct. homeowner (1980)	0.032 (0.029)	0.036 (0.031)
Pct. vacant (1980)	-0.171 (0.192)	-0.131 (0.166)
Log median household income (1980)	0.063*** (0.013)	0.056*** (0.012)
Pop. density (1980)	-2.489 (2.799)	-2.126 (2.949)
Place Fixed Effects	Yes	Yes
R <sup>2</sup>	0.585	0.587
Adj. R <sup>2</sup>	0.555	0.557
Num. obs.	369	369
RMSE	2.025	2.020
N Clusters	19	19

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table D-2: Effect of change in a precinct's non-Black population (1980-1993) on support for voters' veto, 1993. Treatment data are winsorized.

	Support for voters' veto, 1993	
	Model 1	Model 2
Pct. point $\Delta$ in non-Hisp. pop.	-0.163 (0.081)	
Prop. $\Delta$ in non-Hisp pop.		-0.092* (0.040)
Pct. Hispanic (1980)	-0.038 (0.049)	-0.029 (0.059)
Pct. unemployed (1980)	-0.838** (0.303)	-0.830* (0.313)
Pct. homeowner (1980)	0.067 (0.064)	0.068 (0.063)
Pct. vacant (1980)	0.401 (0.331)	0.393 (0.337)
Log median household income (1980)	0.156*** (0.034)	0.156*** (0.034)
Pop. density (1980)	-3.314*** (0.823)	-3.194*** (0.882)
Place Fixed Effects	Yes	Yes
R <sup>2</sup>	0.600	0.600
Adj. R <sup>2</sup>	0.541	0.541
Num. obs.	393	393
RMSE	3.492	3.493
N Clusters	44	44

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table D-3: Effect of change in a precinct's non-Hispanic population (1980-1993) on support for voters' veto, 1993. Treatment data are winsorized.

## E Spatial Durbin Error Model

Spatial autocorrelation can be a problem when using geographic data due to the data's violation of SUTVA. Neighboring observations can influence each other, violating SUTVA's independence assumption. Likewise, omitted variables may be spatially correlated. Fortunately, these relationships can be modeled.

Using a spatial model requires two steps: the definition of spatial relationships and the choice of a model to account for spatial dependence. The most common way to model spatial relationships is by using queen contiguity weights. These weights account for spatial interactions between units sharing either a border or a vertex. Contiguity weights are especially helpful for the precinct data used in this analysis because the relationships are not sensitive to absolute distance. For example, some parts of Los Angeles County are very dense, where spillovers may only extend 500 meters. In contrast, interactions between precincts in rural Los Angeles County may occur at the level of several kilometers. By defining spatial interactions at the level of shared precinct borders (i.e., a queen contiguity weights matrix), this design accounts for variation in the scale of human interaction across Los Angeles County.

To select a model, I considered how spatial dependence may be affecting the analysis. First, the treatment of demographic change is likely to be affecting neighboring precincts. In fact, the second analysis of this study is based on exposure to nearby diversifying precincts. Thus, the model would need to account for this spillover effect. Second, there are likely omitted variables which may be spatially correlated. A spatial Durbin error model (SDEM) accounts for the spatial spillover effects of covariates as well as autocorrelation in the error term (LeSage 2014).

I reproduce the results from the Table 1 using the SDEM. Due to the computational intensity of incorporating the spatial autocorrelation, adjustments were made to the model specifications. Models 1 and 3 include Census place-level fixed effects, but no covariates. Models 2 and 4 include covariates, but no fixed effects. The covariates in Models 2 and 4 are the same as those in the non-spatial results: the 1980 precinct-level percent non-white residents, homeownership rate, vacancy rate, population density, log median household income, and unemployment rate. The results are substantively similar across all four models. All models use the queen continuity weights matrix and are weighted by precinct population.

Having accounted for spatial spillovers and spatially correlated error terms, a standard deviation increase in the local non-white population increases support for direct democracy over housing by 2.0 to 2.2 percentage points. When defined as a proportion increase, the effect is a 0.8 to 2.8 point increase in support for the voters' veto over affordable housing. These effects substantively match the results from the non-spatial OLS model in Table 1.

	Support for voters' veto, 1993			
	Model 1	Model 2	Model 3	Model 4
Pct. point $\Delta$ in non-white pop.	0.022*** (0.003)	0.020*** (0.003)		
Prop. $\Delta$ in non-white pop.			0.028*** (0.003)	0.008* (0.003)
Pct. white, non-Hispanic (1980)		0.099*** (0.014)		0.110*** (0.015)
Pct. unemployed (1980)		-0.581*** (0.074)		-0.644*** (0.076)
Pct. homeowner (1980)		0.095*** (0.012)		0.084*** (0.012)
Pct. vacant (1980)		0.072 (0.068)		0.079 (0.069)
Log med. h.h. income (1980)		0.013 (0.008)		0.009 (0.009)
Pop. density (1980)		-3.178*** (0.922)		-3.197*** (0.938)
Place Fixed Effects	Yes	No	Yes	No
Controls	No	Yes	No	Yes
Num. obs.	3112	3112	3112	3112
Parameters	235	17	235	17
Log Likelihood	3592.387	3653.932	3637.647	3602.665
AIC (Linear model)	-5404.983	-6233.041	-5852.253	-5851.342
AIC (Spatial model)	-6714.774	-7273.864	-6805.294	-7171.331
LR test: statistic	1311.790	1042.823	955.041	1321.988
LR test: p-value	0.000	0.000	0.000	0.000

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table E-4: Effect of change in non-white population (1980-1993) on support for voters' veto, spatial Durbin error model. Treatment data are winsorized.

## F Visualization of Bivariate Relationships

Figure F-8 shows the bivariate relationship between my two treatment variables and precinct support for the voters' veto (Article 34 repeal) using binned box and whiskers plots. Precincts that experienced greater increases in non-white population from 1980 to 1993 showed greater support for the voters' veto over affordable housing in the 1993 November election.

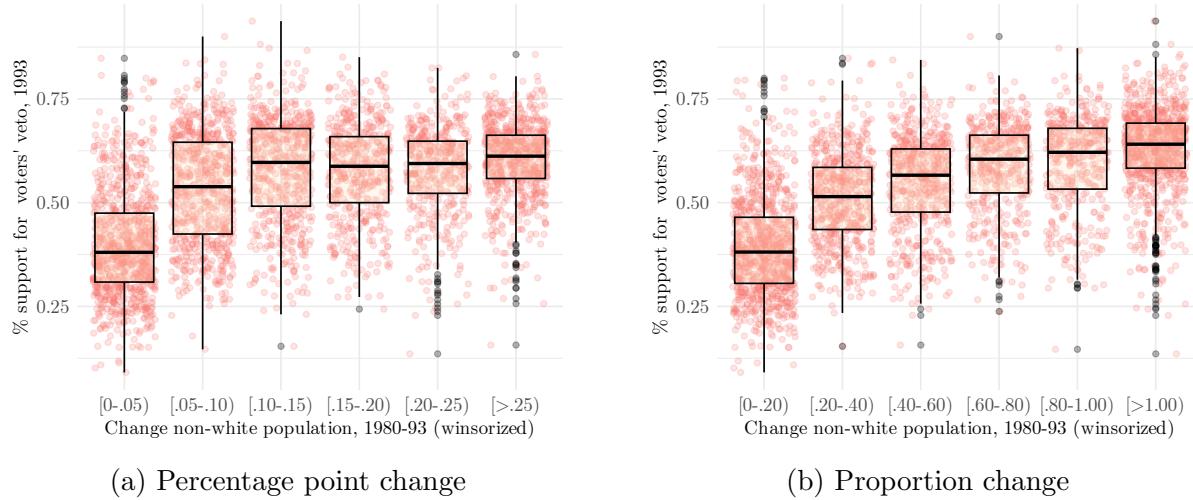


Figure F-8: Relationship between precinct-level change in non-white population (winsorized) and support for Article 34 repeal.

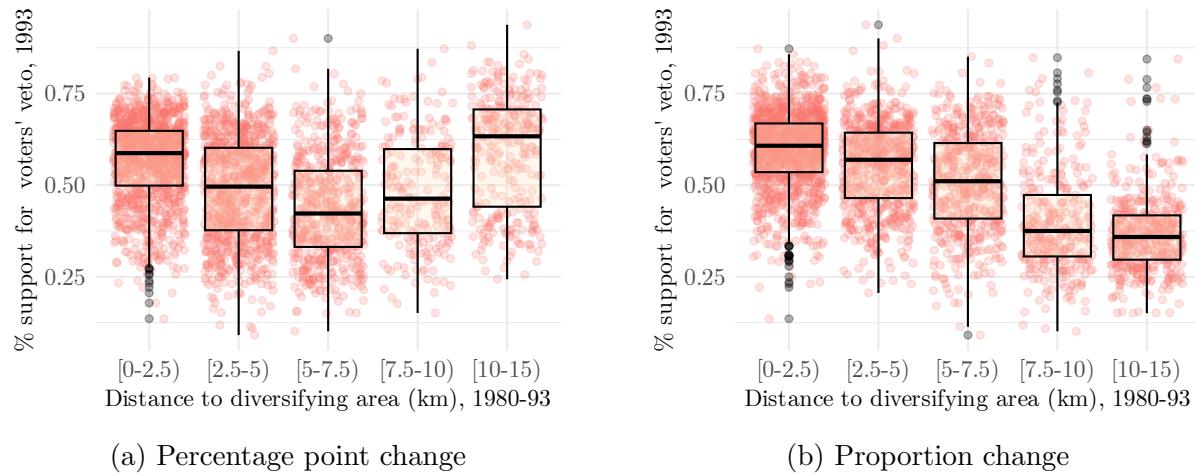


Figure F-9: Relationship between distance from a diversifying Census tract and support for Article 34 repeal in 1993. 10 to 15 kilometers binned together due to data sparseness.

Figure F-9 shows the bivariate relationship between a precinct's distance from a stimulus tract and support for the voters' veto using binned box and whiskers plots. Precincts farther from these diversifying tracts — defined using either percentage point change or proportion change — show less support for direct democracy over housing.

## G Alternative Explanations

	Support for voters' veto, 1993			
	Model 1	Model 2	Model 3	Model 4
Pct. point $\Delta$ in pop. density, 1980-93	0.010* (0.005)	0.011** (0.004)		
Prop. $\Delta$ in pop. density, 1980-93			0.015 (0.009)	0.015* (0.007)
Place Fixed Effects	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
Time-varying Controls	No	Yes	No	Yes
R <sup>2</sup>	0.713	0.732	0.713	0.732
Adj. R <sup>2</sup>	0.701	0.721	0.701	0.721
Num. obs.	3651	3649	3651	3649
RMSE	2.744	2.651	2.742	2.652
N Clusters	142	142	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table G-5: Effect of change in a precinct's population density on support for voters' veto, 1993. Treatment data are winsorized. Controls included in all models, but excluded from table due to space constraints.

	Support for voters' veto, 1993			
	Model 1	Model 2	Model 3	Model 4
Pct. point $\Delta$ in new arrivals, 1980-93	0.001 (0.002)	0.001 (0.002)		
Prop. $\Delta$ in new arrivals, 1980-93			0.002 (0.002)	0.002 (0.002)
Place Fixed Effects	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
Time-varying Controls	No	Yes	No	Yes
R <sup>2</sup>	0.710	0.729	0.710	0.729
Adj. R <sup>2</sup>	0.698	0.717	0.698	0.717
Num. obs.	3651	3649	3651	3649
RMSE	2.758	2.668	2.758	2.667
N Clusters	142	142	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table G-6: Effect of change the rate of residential churn within the past 10 years on support for voters' veto, 1993. Treatment data are winsorized. Controls included in all models, but excluded from table due to space constraints.

	Support for voters' veto, 1993			
	Model 1	Model 2	Model 3	Model 4
Pct. point $\Delta$ in med. hh. income, 1980-93	-0.030* (0.015)	-0.039*** (0.011)		
Prop. $\Delta$ in med. hh. income, 1980-93			-0.016* (0.007)	-0.029*** (0.007)
Place Fixed Effects	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
Time-varying Controls	No	Yes	No	Yes
R <sup>2</sup>	0.719	0.733	0.717	0.732
Adj. R <sup>2</sup>	0.707	0.721	0.705	0.721
Num. obs.	3649	3649	3649	3649
RMSE	2.715	2.648	2.723	2.652
N Clusters	142	142	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table G-7: Effect of change in a precinct's median household income on support for voters' veto, 1993. Treatment data are winsorized. Controls included in all models, but excluded from table due to space constraints.

	Support for voters' veto, 1993			
	Model 1	Model 2	Model 3	Model 4
Pct. point $\Delta$ in foreign born, 1980-93	0.015*** (0.003)	0.013*** (0.002)		
Prop. $\Delta$ in foreign born, 1980-93			0.002 (0.002)	-0.000 (0.002)
Place Fixed Effects	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
Time-varying Controls	No	Yes	No	Yes
R <sup>2</sup>	0.717	0.733	0.710	0.729
Adj. R <sup>2</sup>	0.705	0.722	0.698	0.717
Num. obs.	3651	3649	3651	3649
RMSE	2.725	2.647	2.757	2.668
N Clusters	142	142	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table G-8: Effect of change in a precinct's percent foreign born on support for voters' veto, 1993. Treatment data are winsorized. Controls included in all models, but excluded from table due to space constraints.

Regarding Table G-8, the null effect may be due to two sources of measurement error. First, during the period of this study, there was considerable debate among demographers over how to accurately measure the percent foreign born, let alone the share of immigrants who were non-white (Van Hook et al. 2006; Woodrow-Lafield 1995). Second, percent foreign born was measured via the long-form Census which only sampled 1 in 6 households. While this sampling approach may be sufficient for states and places, it faces limits when aggregated to the precincts with a median size of roughly 1,200 residents. In contrast, these limitations do not exist for the percent non-white population data, which comes from the short-form, full-population Census.

## H Internal Change, Non-Winsorized Data

	Support for voters' veto, 1993	
	Model 1	Model 2
Pct. point $\Delta$ in non-white pop.	0.042*** (0.010)	
Prop. $\Delta$ in non-white pop.		0.030** (0.011)
Pct. white, non-Hispanic (1980)	0.071* (0.034)	0.073 (0.041)
Pct. unemployed (1980)	-0.546*** (0.156)	-0.673** (0.228)
Pct. homeowner (1980)	0.137*** (0.039)	0.124** (0.039)
Pct. vacant (1980)	-0.307** (0.112)	-0.277* (0.114)
Log median household income (1980)	0.007 (0.015)	-0.007 (0.016)
Pop. density (1980)	-6.017*** (0.533)	-5.470*** (0.780)
City FE	Yes	Yes
R <sup>2</sup>	0.747	0.720
Adj. R <sup>2</sup>	0.736	0.708
Num. obs.	3651	3651
RMSE	2.575	2.711
N Clusters	142	142

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table H-9: Effect of change in a precinct's non-white population (1980-1993) on support for voters' veto. Treatment data are not winsorized.

# I Proximity to Change, Alternative Thresholds

	Proportion Change		
	Model 1	Model 2	Model 3
Proximity to change (95th)			0.020 (0.014)
Proximity to change (90th)		0.030*** (0.008)	
Proximity to change (85th)	0.030*** (0.008)		
Place Fixed Effects	Yes	Yes	Yes
Controls	Yes	Yes	Yes
R <sup>2</sup>	0.594	0.597	0.576
Adj. R <sup>2</sup>	0.549	0.552	0.530
Num. obs.	403	403	403
RMSE	2.340	2.333	2.391
N Clusters	41	37	27

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table I-10: Effect of proximity to diversifying tract on support for voter's veto. Controls included in all models, but excluded from table due to space constraints.

	Proportion Change		
	Model 1	Model 2	Model 3
Proximity to proportion change (95th)			0.058*** (0.008)
Proximity to proportion change (90th)		0.052*** (0.008)	
Proximity to proportion change (85th)	0.022 (0.012)		
Place Fixed Effects	Yes	Yes	Yes
Controls	Yes	Yes	Yes
R <sup>2</sup>	0.584	0.665	0.682
Adj. R <sup>2</sup>	0.538	0.628	0.647
Num. obs.	403	403	403
RMSE	2.370	2.125	2.072
N Clusters	41	37	27

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table I-11: Effect of proximity to diversifying tract on support for voter's veto. Controls included in all models, but excluded from table due to space constraints.

	Proportion Change		
	Model 1	Model 2	Model 3
Proximity to change (95th)			0.003 (0.009)
Proximity to change (90th)		0.018** (0.006)	
Proximity to change (85th)	0.017** (0.006)		
Place Fixed Effects	Yes	Yes	Yes
Controls	Yes	Yes	Yes
R <sup>2</sup>	0.591	0.595	0.581
Adj. R <sup>2</sup>	0.541	0.545	0.529
Num. obs.	373	373	373
RMSE	2.445	2.434	2.476
N Clusters	38	33	22

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table I-12: Effect of proximity to diversifying tract on support for voter's veto. Controls included in all models, but excluded from table due to space constraints.

	Proportion Change		
	Model 1	Model 2	Model 3
Proximity to proportion change (95th)			0.044*** (0.007)
Proximity to proportion change (90th)		0.040*** (0.007)	
Proximity to proportion change (85th)	0.015 (0.011)		
Place Fixed Effects	Yes	Yes	Yes
Controls	Yes	Yes	Yes
R <sup>2</sup>	0.591	0.643	0.659
Adj. R <sup>2</sup>	0.540	0.598	0.617
Num. obs.	373	373	373
RMSE	2.446	2.286	2.233
N Clusters	38	33	22

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table I-13: Effect of proximity to diversifying tract on support for voter's veto. Controls included in all models, but excluded from table due to space constraints.

	Percentage	Point $\Delta$	Proportion $\Delta$	
	Model 1	Model 2	Model 3	Model 4
Prox. to pct point. $\Delta$ in non-white	0.030*** (0.008)		0.016** (0.005)	
Prox. to prop. $\Delta$ in non-white		0.020* (0.009)		0.013 (0.009)
Pct. white, non-Hispanic (1980)	0.192*** (0.050)	0.154** (0.052)	0.148** (0.046)	0.119* (0.047)
Pct. unemployed (1980)	-0.615 (0.353)	-0.528 (0.349)	-0.774* (0.327)	-0.782* (0.320)
Pct. homeowner (1980)	0.038 (0.061)	0.059 (0.054)	-0.014 (0.059)	-0.010 (0.053)
Pct. vacant (1980)	-0.704 (0.524)	-0.756 (0.508)	-0.952* (0.425)	-1.048* (0.411)
Log median household income (1980)	-0.071 (0.038)	-0.079* (0.036)	-0.046 (0.035)	-0.055 (0.033)
Pop. density (1980)	-4.668 (10.310)	-1.800 (9.694)	-7.253 (7.638)	-6.316 (7.362)
Place Fixed Effects	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
R <sup>2</sup>	0.597	0.578	0.592	0.589
Adj. R <sup>2</sup>	0.552	0.531	0.541	0.538
Num. obs.	403	403	373	373
RMSE	2.334	2.388	2.445	2.453
N Clusters	42	42	37	37

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table I-14: Effect of proximity to diversifying tract on support for voter's veto, among precincts which changed minimally from 1980 to 1993. Population threshold for diversifying tract removed.

## J Proximity to Change, by 1980 Racial Composition

This analysis is conducted on precincts in the top tercile percent white ( $> 85\%$ ) and the bottom tercile percent white ( $< 52\%$ ). These values differ slightly from the “within precinct” analysis as the data are subset to precincts which changed only minimally between 1980 and 1993. Proximity to a diversifying tract is interacted with a binary variable for whether the precinct was in the top percent white tercile in 1980.

Results here are noisier due to the very limited sample sizes. However, precincts that were in the bottom tercile percent white still show a generally statistically significant reaction to the proximity of a diversifying tract. The lower order term in Model 1 is significant at the  $\alpha = 0.10$ .

	Percentage Point $\Delta$		Proportion $\Delta$	
	Model 1	Model 2	Model 3	Model 4
Prox. to pct. point $\Delta$ in non-white	0.028 (0.017)		0.010 (0.009)	
Prox. to prop. $\Delta$ in non-white		0.056*** (0.011)		0.043*** (0.008)
Prox. to pct. point x white prec., 1980	0.005 (0.023)		-0.002 (0.019)	
Prox. prop. x white prec., 1980		-0.008 (0.017)		-0.049* (0.019)
Pct. unemployed (1980)	-0.472 (0.255)	-0.977*** (0.210)	-0.909*** (0.198)	-1.283*** (0.191)
Pct. homeowner (1980)	0.026 (0.052)	0.034 (0.052)	-0.057 (0.047)	-0.058 (0.039)
Pct. vacant (1980)	-1.068 (0.605)	-0.964* (0.451)	-0.873* (0.349)	-0.725* (0.309)
Log median household income (1980)	-0.047 (0.043)	-0.074* (0.033)	0.023 (0.029)	-0.018 (0.024)
Pop. density (1980)	-8.927 (11.408)	-3.601 (8.387)	-13.427 (7.109)	-10.552 (5.881)
Place Fixed Effects	Yes	Yes	Yes	Yes
Controls	Yes	Yes	Yes	Yes
R <sup>2</sup>	0.599	0.664	0.596	0.654
Adj. R <sup>2</sup>	0.549	0.622	0.539	0.605
Num. obs.	374	374	340	340
RMSE	2.313	2.118	2.447	2.266
N Clusters	36	36	32	32

\*\*\* $p < 0.001$ ; \*\* $p < 0.01$ ; \* $p < 0.05$

Table J-15: Effect of proximity to diversifying tract on support for voter’s veto, among precincts which changed minimally from 1980 to 1993. Interacted with precinct percent non-white, 1980.

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