

Efficient simulation of the SABR model

Applied Stochastic Processes (FIN 514)

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2023-24 Module 3 (Spring 2024)

Stochastic-alpha-beta-rho (SABR) Model

$$\frac{dF_t}{F_t^\beta} = \sigma_t dW_t \quad \text{and} \quad \frac{d\sigma_t}{\sigma_t} = \xi dZ_t \quad (\sigma_0 = \alpha) \quad \text{for} \quad dW_t dZ_t = \rho h$$

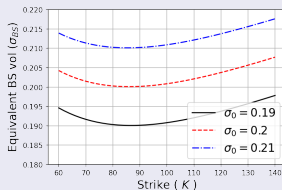
or $\frac{dF_t}{F_t^\beta} = \sigma_t (\rho dZ_t + \rho_* dX_t) \quad \text{for} \quad \rho_* = \sqrt{1 - \rho^2}, \quad dX_t dZ_t = 0$

- One of the most popular stochastic volatility (SV) models used in financial engineering ([Hagan et al., 2002](#)).
- A flexible choice of *backbone* (volatility v.s. spot change): $0 \leq \beta \leq 1$
- Parsimonious and intuitive parameters: $\sigma_0 = \alpha$, ρ and ξ (and β)
- Asymptotic approximation (HKLW formula) for the equivalent BS volatility ($\xi\sqrt{T} \ll 1$)

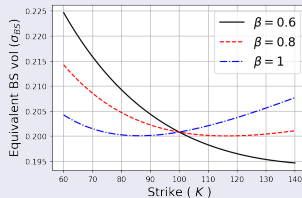
The impact of parameters

Base parameters: $\sigma_0 = 0.2$, $\xi = 0.2$, $\rho = 0.1$, $\beta = 1$.

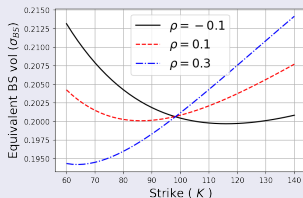
Initial vol σ_0 (level)



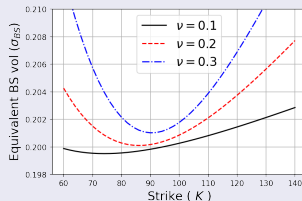
Backbone β (skew)



Correlation ρ (skew)



Vol-of-vol ν (smile)



Black-Scholes (BS) Model

- Black and Scholes (1973); Merton (1973).
- Dynamics (geometric BM):

$$\frac{dF_t}{F_t} = \sigma_{\text{BS}} dW_t$$

- Distribution:

$$F_T \sim \text{LN}(F_0, \sigma_{\text{BS}}^2 T).$$

- Option Price (Black, 1976):

$$C_{\text{BS}} = F_0 N(d_+) - K N(d_-) \quad \text{where} \quad d_{\pm} = \frac{\log(F_0/K)}{\sigma_{\text{BS}} \sqrt{T}} \pm \frac{1}{2} \sigma_{\text{BS}} \sqrt{T}$$

- Asset class: equity, FX, etc.
- As $\xi \downarrow 0$, SABR \rightarrow BS model.

Bachelier (Normal) Model

- The birth of Brownian motion ([Bachelier, 1900](#))
- Dynamics (arithmetic BM):

$$dF_t = \sigma_N dW_t$$

- Distribution:

$$F_T \sim N(F_0, \sigma_N^2 T).$$

- Option Price:

$$C_N = (F_0 - K)N(d_N) + \sigma\sqrt{T}n(d_N) \quad \text{where} \quad d_N = \frac{(F_0 - K)}{\sigma_N\sqrt{T}}$$

- Daily price changes may not be proportional to the price level. Underlying price can be negative.
- Asset class: interest rate, inflation, spread ($X - Y$), WTI futures
- Fast implied volatility inversion ([Choi et al., 2009](#)) and review ([Choi et al., 2022](#)).
- As $\xi \downarrow 0$, SABR \rightarrow Bachelier model.

Constant-Elasticity-of-Variance (CEV) Model

- Dynamics (CEV):

$$\frac{dF_t}{F_t^\beta} = \sigma_{\text{CEV}} dW_t \quad (\beta_* = 1 - \beta)$$

- Distribution:

$$F_T : \sim \text{CEV}(\beta, F_0, \sigma_{\text{CEV}}^2 T).$$

- Option Price (Schroder, 1989):

$$C_{\text{CEV}} = F_0 \bar{F}_{\chi^2} \left(\frac{q_K^2}{\sigma_{\text{CEV}}^2 T}; 2 + \frac{1}{\beta_*}, \frac{q_0^2}{\sigma_{\text{CEV}}^2 T} \right) - K \underbrace{F_{\chi^2} \left(\frac{q_0^2}{\sigma_{\text{CEV}}^2 T}; \frac{1}{\beta_*}, \frac{q^2(K)}{\sigma_{\text{CEV}}^2 T} \right)}_{=\text{Prob}(F_T > K) = 1 - F_{\text{CEV}}(K)},$$

where $F_{\chi^2}(x; r, x_0)$ (\bar{F}_{χ^2}) is the CDF (1-CDF) of noncentral χ^2 with degrees of freedom r and noncentrality x_0 .

- F_T is **not** a χ^2 RV.
- Mass at zero exists:

$$\text{Prob}(F_T = 0) = \bar{F}_{\chi^2} \left(\frac{q_0^2}{\sigma_{\text{CEV}}^2 T}; \frac{1}{\beta_*}, 0 \right) = \bar{F}_{\Gamma} \left(\frac{q_0^2}{2\sigma_{\text{CEV}}^2 T}; \frac{1}{2\beta_*} \right)$$

- As $\xi \downarrow 0$, SABR \rightarrow CEV (β). (Motivation for this study and Choi and Wu (2021a))

Normal SABR ($\beta = 0$)

- **Closed-form Monte-Carlo simulation of the normal SABR:**
[Choi, Liu and Seo \(2019\)](#). Hyperbolic normal stochastic volatility model. *Journal of Futures Markets*, 39:186–204. [arXiv:1809.04035](#)
- Choi and Seo (2024). Option pricing under the normal SABR model with quadratic approximation. *Working paper*

CEV SABR ($0 < \beta < 1$)

- [Choi and Wu \(2021a\)](#). The equivalent CEV volatility of the SABR model. *Journal of Economic Dynamics and Control*, 128:104143. [arXiv:1911.13123](#)
- **Option pricing for the uncorrelated ($\rho = 0$) case:**
[Choi and Wu \(2021b\)](#) A note on the option price and ‘Mass at zero in the uncorrelated SABR model and implied volatility asymptotics.’ *Quantitative Finance*, 21:1083–1086. [arXiv:2011.00557](#)
- Choi, Hu, and Kwok (2024). Efficient simulation of the SABR model. *Working paper*

The equivalent CEV volatility of the SABR model

Choi, J., & Wu, L. (2021). *Journal of Economic Dynamics and Control*, 128:104143. [arXiv:1911.13123](#).

Equivalent CEV volatility ($0 < \beta < 1$)

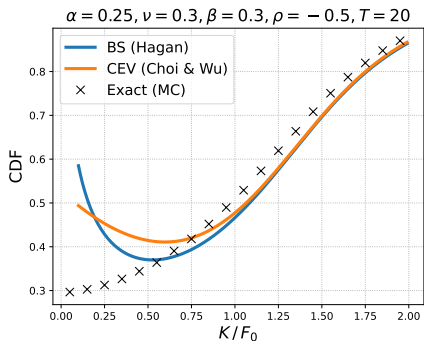
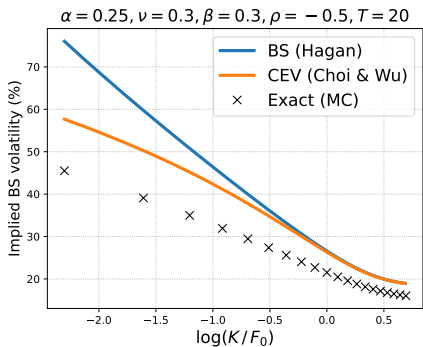
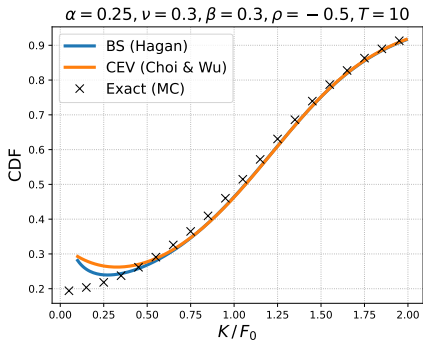
- SABR is the CEV model (backbone) with stochastic vol.
- When $\xi \downarrow 0$, SABR \rightarrow CEV (β) model with the fixed vol $\sigma_0 = \alpha$:

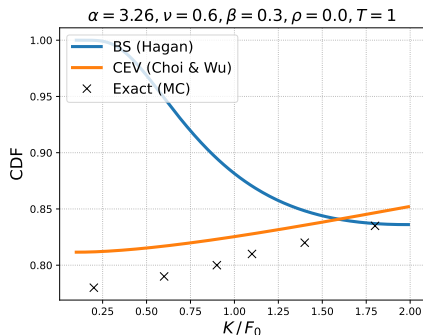
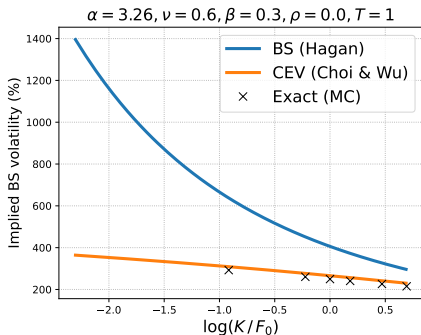
$$\sigma_{BS} = \alpha \frac{H(0) \rightarrow 1}{K^{\beta_*/2}} \frac{1 + \left(\frac{\beta_*^2}{24 K^{\beta_*}} \alpha^2 + \frac{\rho \beta}{4 K^{\beta_*/2}} \alpha \cdot 0 + \frac{2-3\rho^2}{24} 0^2 \right) T}{1 + \frac{\beta_*^2}{24} \log^2 K + \frac{\beta_*^4}{1920} \log^4 K} \quad (F_0 = 1)$$

- The blue terms are for the approximate CEV-to-BS vol conversion, which has nothing to do with the stochastic vol (ξ).
- Instead, derive an equivalent CEV vol; (i) the conversion error is avoided, (ii) $\sigma_{CEV} \rightarrow \alpha$ as $\xi \downarrow 0$, and (iii) expressions are simpler.

$$\sigma_{CEV} = \alpha H(z) \left(1 + \left[\frac{\rho \beta}{4 K^{\beta_*/2}} \alpha \xi + \frac{2-3\rho^2}{24} \xi^2 \right] T \right) \quad \text{for} \quad z = \frac{\xi}{\alpha \beta_*} (K^{\beta_*} - 1)$$

- The BS vol can be converted from the analytic CEV option price.
- Equivalent CEV vol allows less arbitrage in low strike.





Option pricing under the normal SABR model with quadratic approximation.

Choi & Seo (2024). Working Paper.

Normal SABR model ($\beta = 0$)

- The integral representations for option value has been investigated by many ([Henry-Labordère, 2008](#); [Korn and Tang, 2013](#); [Antonov et al., 2015](#)).
- [Antonov et al. \(2015\)](#)'s representation is the simplest: single integral with the elliptic integral of the 2nd kind:

$$C_{\text{SABR}} = [F_0 - K]^+ + \frac{\sigma_0 \sqrt{T}}{\pi} \int_{u_0}^{\infty} F(u) w(u) du,$$

where

$$u_0 = \frac{1}{2\xi} \text{acosh} \left(\frac{V - \rho k}{\rho_*^2} \right) \quad \text{where} \quad \xi = \frac{\xi \sqrt{T}}{2}, \quad k = z_N + \rho, \quad V = \sqrt{k^2 + \rho_*^2},$$

$$w(u) = e^{-\frac{\xi^2}{2}} n(u) [R_{+\xi}(u) + R_{-\xi}(u)] \quad \text{with} \quad R_{\Delta}(u) := \frac{N(-u - \Delta)}{n(u + \Delta)}$$

$$F(u) = \sqrt{\alpha_+(u)} \text{EllipticE} \left(\frac{\alpha_+(u) - \alpha_-(u)}{\alpha_+(u)} \right) \quad (F(u_0) = 0),$$

$$\alpha_{\pm}(u) = \rho k + \rho_*^2 \cosh(2\xi u) \pm \rho_* \sqrt{\sinh^2(2\xi u) - (k - \rho \cosh(2\xi u))^2}.$$

Approximation for option value

- Analytic integrals are available for

$$\int_{u_0}^{\infty} (u - u_0)^n w(u) du \quad \text{for } n = 0, 1, \text{ and } 2.$$

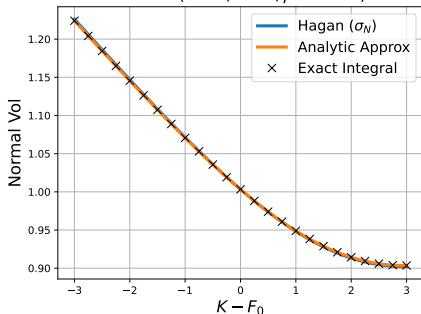
- Approximate $F(u)$ as a quadratic function near $u = u_0$:

$$F(u) \approx \frac{\pi}{2} \sqrt{V} [1 + A \xi(u - u_0) + B \xi^2(u - u_0)^2],$$
$$A = \frac{3}{4} \left| \rho - \frac{k}{V} \right| \quad \text{and} \quad B = \frac{3}{4} \left(1 - \frac{\rho k}{V} \right) - \frac{15}{64} \left| \rho - \frac{k}{V} \right|^2$$

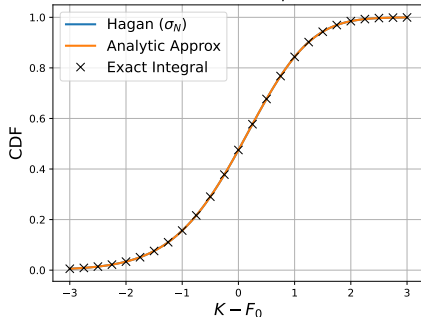
Approximate option value under the normal SABR

$$C \approx [F_0 - K]^+ + \frac{\sigma_0 \sqrt{V} n(u_0)}{\xi e^{\frac{\xi^2}{2}}} \left[\begin{array}{c} [R_{-\xi}(u_0) - R_{+\xi}(u_0)] \\ + A [R_{+\xi}(u_0) + R_{-\xi}(u_0) - 2R_0(u_0)] \\ + 2B [R_{-\xi}(u_0) - R_{+\xi}(u_0) - 2\xi(1 - u_0 R_0(u_0))] \end{array} \right]$$

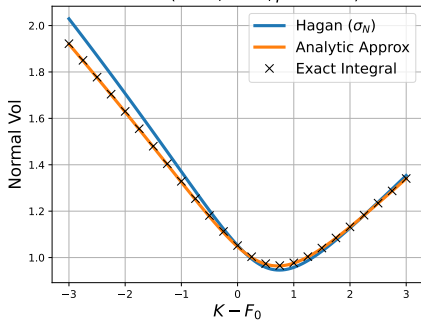
$\nu=0.25$ ($\alpha=1, T=1, \rho=-0.5$)



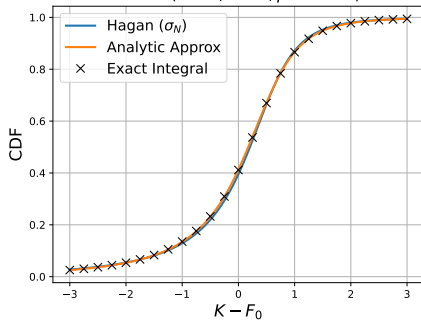
$\nu=0.25$ ($\alpha=1, T=1, \rho=-0.5$)



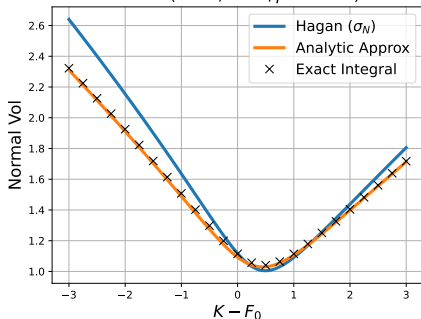
$\nu=1.0$ ($\alpha=1, T=1, \rho=-0.5$)



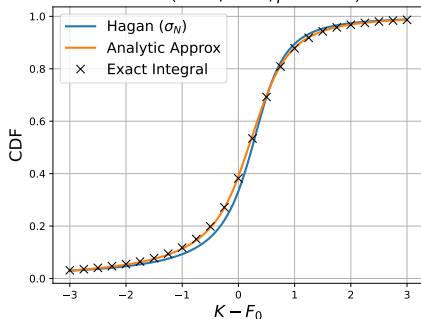
$\nu=1.0$ ($\alpha=1, T=1, \rho=-0.5$)



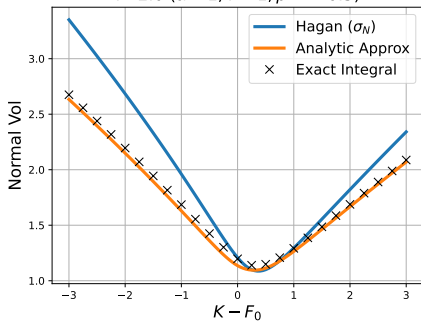
$\nu=1.5$ ($\alpha=1, T=1, \rho=-0.5$)



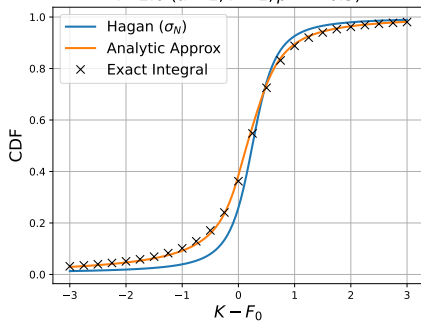
$\nu=1.5$ ($\alpha=1, T=1, \rho=-0.5$)



$\nu=2.0$ ($\alpha=1, T=1, \rho=-0.5$)



$\nu=2.0$ ($\alpha=1, T=1, \rho=-0.5$)



Stochastic Integral

The volatility process is a geometric BM:

$$\frac{d\sigma_t}{\sigma_t} = \xi dZ_t, \quad \sigma_T = \sigma_0 \exp\left(\xi Z_T - \frac{\xi^2 T}{2}\right) \Rightarrow \int_0^T \sigma_t dZ_t = \frac{1}{\xi}(\sigma_T - \sigma_0).$$

The price process is integrated to

$$\begin{aligned} \int_0^T \frac{dF_t}{F_t^\beta} &= \int_0^T \sigma_t (\rho dZ_t + \rho_* dX_t) = \frac{\rho}{\xi}(\sigma_T - \sigma_0) + \rho_* \int_0^T \sigma_t dX_t \\ &\sim \underbrace{\frac{\rho}{\xi}(\sigma_T - \sigma_0)}_{\text{skew}} + \underbrace{\rho_* X_{V_{0,T}}}_{\text{smile}} \sim \underbrace{\frac{\rho}{\xi}(\sigma_T - \sigma_0)}_{\text{adjustment to } F_0} + \underbrace{\rho_* \sigma_0 \sqrt{TI_{0,T}}}_{\text{volatility}} X, \quad (X \sim N(0, 1)) \end{aligned}$$

where $V_{0,T}$ ($I_{0,T}$) is the (normalized) conditional integrated variance or stochastic time clock:

$$V_{0,T}(\sigma_0, \sigma_T) = \int_0^T \sigma_t^2 dt \Big|_{\sigma_0, \sigma_T} \quad \text{and} \quad I_{0,T}(\sigma_0, \sigma_T) = \frac{V_{0,T}}{\sigma_0^2 T} \left(\lim_{\xi \downarrow 0} I_{0,T} \rightarrow 1 \right).$$

The joint sampling of $(\sigma_T, I_{0,T})$ is critical for understanding the SABR model.

Simulation Steps

- 1. Simulate σ_{t+h} from σ_t (trivial):

$$\sigma_{t+h} = \sigma_t \exp \left(\xi \sqrt{h} Z - \frac{\xi^2 h}{2} \right), \quad Z \sim N(0, 1).$$

- 2. Simulate $I_{t,t+h}$ given σ_t and σ_{t+h} .
- 3. Simulate F_{t+h} given F_t , σ_{t+h} , and $I_{t,t+h}$.
 - A. For $0 < \beta < 1$, how to approximate F_{t+h} ? A CEV distribution?
 - B. How to sample F_{t+h} from the approximated distribution?
- We propose better methods for (2), (3A), and (3B).
- Our new methods independent.
- Applying all methods yield best results. However, they can be applied selectively to the existing methods.

(2) Simulation of $I_{t,t+h}$: existing studies

- Chen et al. (2012) used the lognormal sampling, but based on the short-time expansion:

$$\mu_1 = E(I_{t,t+h}) \approx 1 + \xi\sqrt{h}Z + \dots \quad \text{and} \quad \mu_2 = E(I_{t,t+h}^2) \approx \mu_1^2 + \frac{\xi^2 h}{3},$$

where Z is the normal variate used in σ_{t+h} .

- Cai et al. (2017) used the Laplace transform of $1/I_{t,t+h}$ (Matsumoto and Yor, 2005):

$$E\left(e^{-s/I_{t,t+h}}\right) = \exp\left(-\frac{\phi_x^2(\xi^2 h s) - x^2}{2\xi^2 h}\right)$$

$$\text{for } x = \log(\sigma_{t+h}/\sigma_t) \quad \text{and} \quad \phi_x(\lambda) = \text{acosh}(se^{-x} + \cosh(x)).$$

Then, they sample $1/I_{t,t+h}$ from the numerical inversion.

- Leitao et al. (2017a) use Fourier transform of $1/I_{t,t+h}$.

(2) Simulation of $I_{t,t+h}$: (shifted) lognormal approximation

- Trapezoidal rule:

$$\text{TZ} : I_{t,t+h} \approx \frac{1}{2} \left(1 + \frac{\sigma_{t+h}^2}{\sigma_t^2} \right) \quad (\text{deterministic})$$

- We sample $I_{t,t+h}$ from the lognormal approximation with moments μ_1 and μ_2 :

$$\text{LN} : I_{t,t+h} \approx \mu_1 \exp(\lambda Z - \lambda^2/2), \quad \lambda = \sqrt{\ln(\mu_2/\mu_1^2)}, \quad Z \sim N(0, 1).$$

Conditional moments of $I_{t,t+h}$ are available from [Kennedy et al. \(2012\)](#):

$$\mu_1 = E(I_{t,t+h}) = \frac{N(d_{+1}) - N(d_{-1})}{2\xi\sqrt{h} n(d_{+1})}, \quad \left(d_k = \frac{\ln(\sigma_{t+h}/\sigma_t)}{\xi\sqrt{h}} + k\xi\sqrt{h} \right)$$

$$\mu_2 = E(I_{t,t+h}^2) = \frac{1}{(\xi\sqrt{h})^{1.5}} \left[\frac{N(d_{+2}) - N(d_{-2})}{4 n(d_{+2})} - \frac{\sigma_t^2 + \sigma_{t+h}^2}{2\sigma_0^2} \frac{N(d_{+1}) - N(d_{-1})}{2 n(d_{+1})} \right].$$

- We also sample from the shifted lognormal distribution matched to μ_1 , μ_2 , and μ_3 :

$$\text{S-LN} : I_{t,t+h} \approx (1 - \eta)\mu_1 + \eta\mu_1 \exp(\lambda Z - \lambda^2/2).$$

(3AB) Conditional Simulation for $\beta = 0$ or 1

- SABR ($\beta = 0$) is simulated with conditional normal distribution:

$$F_T = F_0 + \frac{\rho}{\xi}(\sigma_T - \sigma_0) + \rho_*\sigma_0\sqrt{TI_{0,T}} X.$$

Conditional on σ_T and $I_{0,T}$,

$$F_T \sim N(\bar{F}_T, \rho_*^2\sigma_0^2 TI_{0,T}) \quad \text{where} \quad \bar{F}_T = F_0 + \frac{\rho}{\xi}(\sigma_T - \sigma_0).$$

But, closed-form simulation method is available in [Choi et al. \(2019\)](#).

- SABR ($\beta = 1$) is simulated with conditional BS distribution:

$$\log\left(\frac{F_T}{F_0}\right) \sim \frac{\rho}{\xi}(\sigma_T - \sigma_0) - \frac{\sigma_0^2 TI_{0,T}}{2} + \rho_*\sigma_0\sqrt{TI_{0,T}} X.$$

Conditional on σ_T and $I_{0,T}$,

$$F_T \sim \text{LN}(\bar{F}_T, \rho_*^2\sigma_0^2 TI_{0,T}) \quad \text{where} \quad \bar{F}_T = F_0 \exp\left(\frac{\rho}{\xi}(\sigma_T - \sigma_0) - \frac{\rho^2\sigma_0^2 TI_{0,T}}{2}\right)$$

- Note $F_0 = E(\bar{F}_T)$ for any initial $\sigma_0 > 0$.

Closed-form Simulation for $\beta = 0$

Choi, Liu and Seo (2019). Hyperbolic normal stochastic volatility model. *Journal of Futures Markets*, 39:186–204. [arXiv:1809.04035](https://arxiv.org/abs/1809.04035)

$$\sigma_T = \sigma_0 \exp(\xi Z'_T), \quad Z'_T = Z_T + \mu T$$

$$F_T \sim F_0 + \frac{\sigma_0 \rho}{\xi} \left(e^{\xi Z'_T} - e^{\lambda \xi^2 T/2} \right) + \frac{\sigma_0 \rho^*}{\xi} \cos \theta \phi \left(\xi Z'_T, \xi \sqrt{R_T^2 + (Z'_T)^2} \right),$$

- $\phi(Z, D) = e^{Z/2} \sqrt{2 \cosh D - 2 \cosh Z} \quad (Z \leq D)$
- R_T is the 2-d squared Bessel process (i.e., $R_T = X_T^2 + Y_T^2$ for two independent BMs, X_T and Y_T)
- $\theta \in [0, \pi]$ is a uniformly distributed random angle.
- 1.5 normal variates for sampling one pair of σ_T and F_T .

(3AB) Conditional Simulation for $\rho = 0$ and $0 < \beta < 1$

- F_t follows the CEV process with stochastic time:

$$\frac{dF_\tau}{F_\tau^\beta} = dW_\tau \quad \text{where} \quad \tau = \int_0^t \sigma_s^2 ds = V_{0,t}.$$

Conditional on σ_T and $I_{0,T}$,

$$F_T \sim \text{CEV}(\beta, F_0, \rho_*^2 \sigma_0^2 T I_{0,T}). \quad (\rho_* = 1)$$

- Approximate option price is available in [Choi and Wu \(2021b\)](#).
- How to sample F_T from the CEV distribution?
- However, no exact distribution for F_T if $\rho \neq 0$.

Euler/Milstein scheme

$$F_{t+h} \approx F_t + \sigma_t F_t^\beta \sqrt{h} W + \frac{\beta \sigma_t^2 F_t^{2\beta-1}}{2} h (W^2 - 1),$$

where W is normal variate with $E(WZ) = \rho$.

Option pricing for the uncorrelated ($\rho = 0$) case

Choi and Wu (2021b) A note on the option price and ‘Mass at zero in the uncorrelated SABR model and implied volatility asymptotics.’ *Quantitative Finance*, 21:1083–1086. [arXiv:2011.00557](https://arxiv.org/abs/2011.00557)

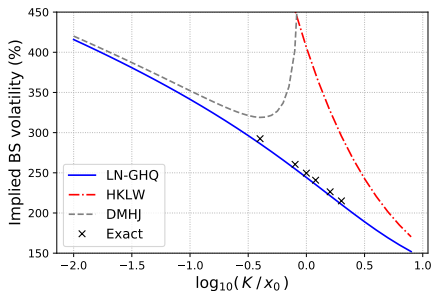
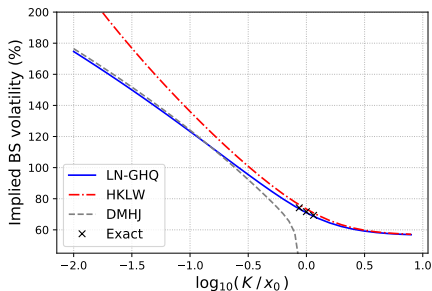
$$\begin{aligned} C_{\text{SABR}}(K) &= E \left(C_{\text{CEV}}(K, \beta, F_0, \sigma_0^2 T I_{0,T}) \right) = \int_0^\infty C_{\text{CEV}}(K, \beta, F_0, \sigma_0^2 T y) f(y) dy \\ &= \sum_{k=1}^N C_{\text{CEV}}(K, \beta, F_0, \sigma_0^2 T y_k) w_k \end{aligned}$$

- Approximate $I_{0,T}$ with moment-matched (μ_1 and μ_2) lognormal distribution.

$$y_k = \mu_1 \exp(\lambda z_k - \lambda^2/2), \quad \lambda = \sqrt{\ln(\mu_2/\mu_1^2)},$$

- (z_k, w_k) is the point and weights of the Gauss-Hermite quadrature.

Figure: The Black-Scholes volatility smile as a function of log strike price for two parameter sets: $(x_0, y_0, \xi, \beta, T) = (0.5, 0.5, 0.4, 0.5, 2)$ (left) and $(x_0, y_0, \xi, \beta, T) = (0.05, 0.4, 0.6, 0.3, 1)$ (right). The mass at zero, estimated by the method of this study, are $m_{\text{SABR}} = 0.1657$ and 0.7624 , respectively.



(3A) Approximation for CEV SABR: Islah (2009)

- All SABR MC literature ([Chen et al., 2012](#); [Leitao et al., 2017a,b](#); [Cai et al., 2017](#)) use [Islah \(2009\)](#)'s approximation.
- For $q_t = F_t^{\beta_*} / \beta_*$,

$$q_{t+h} = q_t + \frac{\rho}{\xi} (\sigma_{t+h} - \sigma_t) + \int_t^{t+h} \rho_* \sigma_s dq_s - \frac{\beta}{2\rho_*^2 \beta_*} \int_t^{t+h} \frac{(\rho_* \sigma_s)^2}{q_s} dt$$
$$\frac{\beta'}{\beta_*} = \frac{\beta}{\rho_*^2 \beta_*} \Rightarrow \beta' = \frac{\beta}{\beta + \beta_* \rho_*^2}$$
$$\text{Prob}(F_T \leq K) \approx \bar{F}_{\chi^2} \left(\frac{\left| q_0 + \frac{\rho}{\xi} (\sigma_T - \sigma_0) \right|^2}{\sigma_0^2 T I_{0,T}}; 1 + \frac{\beta}{\beta_* \rho_*^2}, \frac{q_K^2}{\sigma_0^2 T I_{0,T}} \right).$$

- As $\xi \downarrow 0$, distribution does not converge to CEV (β). $\beta' \neq \beta$.
- Only exact when $\rho = 0$.
-

$$\bar{F}_T = \left(\beta'_* \left| q_0 + \frac{\rho}{\xi} (\sigma_T - \sigma_0) \right| \right)^{1/\beta'_*}. \quad E(\bar{F}_T) \neq F_0$$

(3B) Simulation of χ^2 RV: Existing studies

It is not simple to sample F_T from [Islah \(2009\)](#)'s noncentral χ^2 approximation.

- [Cai et al. \(2017\)](#): For a uniform U , solve x with numerical root-finding:

$$\text{Prob}(F_T \leq x) = U.$$

- [Chen et al. \(2012\)](#); [Leitao et al. \(2017a,b\)](#): moment-matched quadratic Gaussian [Andersen \(2008\)](#):

$$q_T^2 \sim d(e + Z)^2 \quad \text{for } Z \sim N(0, 1).$$

- [Grzelak et al. \(2019\)](#) used stochastic collocation: efficient interpolation using orthogonal polynomial.

(3A) Approximation for CEV SABR: New method

- Conditional on F_t , σ_t , and $I_{t,t+h}$, assume F_{t+h} follows a CEV distribution with β :

$$F_{t+h} \sim \text{CEV}(\beta, \underbrace{\bar{F}_{t+h}}_{???}, \underbrace{\rho_*^2 \sigma_t^2 h I_{t,t+h}}_{\text{Same as other cases}}).$$

- Approximate dynamics with GBM with $\sigma_t := \sigma_t / F_0^{\beta_*}$:

$$\frac{dF_t}{F_t} = \frac{\sigma_t}{F_t^{\beta_*}} dW_t \approx \frac{\sigma_t}{F_0^{\beta_*}} dW_t.$$

- Borrow \bar{F}_T from the $\beta = 1$ case:

$$\bar{F}_{t+h} \approx F_t \exp \left(\frac{\rho}{\xi} \frac{(\sigma_{t+h} - \sigma_t)}{F_t^{\beta_*}} - \frac{\rho^2 \sigma_t^2 h I_{t,t+h}}{2 F_t^{2\beta_*}} \right)$$

- Exact when $\rho = 0$: $\bar{F}_T = F_t$.
- Preserves martingale: $E(F_T) = E(\bar{F}_T) = F_0$.

(3B) Simulation of the CEV process: Kang (2014)

- [Kang \(2014\)](#). Simulation of the shifted Poisson distribution with an application to the CEV model. *Management Science and Financial Engineering* 20:27–32.
[10.7737/MSFE.2014.20.1.027](#)

- Poisson variate, $N \sim \text{Pois}(\lambda)$:

$$1 = e^\lambda e^{-\lambda} = \sum_{k=0}^{\infty} \boxed{P_k := \frac{\lambda^k}{k!} e^{-\lambda}}.$$

- Let $F_\Gamma(x; \nu)$ be the CDF of standard gamma distribution (lower incomplete gamma fn):

$$F_\Gamma(x; \nu) = \frac{1}{\Gamma(\nu)} \int_0^x t^{\nu-1} e^{-t} dt = \sum_{k=0}^{\infty} \frac{x^{k+\nu} e^{-x}}{\Gamma(k + \nu + 1)}$$

- **Shifted** Poisson variate, $N \sim \text{dPois}(\lambda, \nu)$:

$$1 = \frac{F_\Gamma(\lambda; \nu)}{F_\Gamma(\lambda; \nu)} = \frac{1}{\Gamma(\nu)} \int_0^x t^{\nu-1} e^{-t} dt = \sum_{k=0}^{\infty} \boxed{P_k^\nu := \frac{1}{F_\Gamma(\lambda; \nu)} \frac{\lambda^{k+\nu} e^{-\lambda}}{\Gamma(k + \nu + 1)}},$$

- $\text{dPois}(\lambda, \nu = 0) = \text{Pois}(\lambda)$.
- [Kang \(2014\)](#) samples as Gamma-mixed Poisson variate:

$$\text{dPois}(\lambda, \nu) \sim \text{Pois}(\lambda - X) \quad \text{conditional on} \quad X \sim \Gamma(\nu) \geq \lambda.$$

- Let

$$z_t = \frac{q_t^2}{\sigma_{\text{CEV}}^2 T} = \frac{F_t^{2\beta_*}}{\beta_*^2 \sigma_{\text{CEV}}^2 T} \quad \text{and} \quad \nu = \frac{1}{2\beta_*}$$

- Kang (2014) also shows the transition probability from z_0 to z_T follows the shifted-Poisson-mixed Gamma distribution.

$$\begin{aligned} f(z_T | z_0) &= \frac{1}{2} \left(\frac{\sqrt{z_0}}{\sqrt{z_T}} \right)^\nu I_\nu(\sqrt{z_0 z_T}) e^{-\frac{z_0 + z_T}{2}} = \frac{z_0^\nu}{2^\nu} \sum_{k=0}^{\infty} \frac{(z_0/2)^k e^{-\frac{z_0}{2}}}{\Gamma(k + \nu + 1)} \frac{(z_T/2)^k}{2 k!} e^{-\frac{z_T}{2}} \\ &= \sum_{k=0}^{\infty} \frac{(z_0/2)^{k+\nu} e^{-\frac{z_0}{2}}}{\Gamma(k + \nu + 1)} f_\Gamma(z_T; k + 1, 1/2) = \sum_{k=0}^{\infty} P_k^\nu f_\Gamma(z_T; k + 1, 1/2). \end{aligned}$$

Simulation algorithm for $F_T \sim \text{CEV}(\beta, F_0, \sigma_{\text{CEV}}^2 T)$ from Kang (2014):

- 1 Sample $X \sim 2\Gamma\left(\frac{1}{2\beta_*}\right)$.
- 2 If $X \geq z_0 = \frac{F_0^{2\beta_*}}{\beta_*^2 \sigma_{\text{CEV}}^2 T}$, $F_T = 0$ and terminate.
- 3 Else, sample $z_T \sim 2\Gamma\left(\text{Pois}\left(\frac{z_0 - X}{2}\right) + 1\right)$.
- 4 Obtain F_T from $z_T = \frac{F_T^{2\beta_*}}{\beta_*^2 \sigma_{\text{CEV}}^2 T}$.

- Let

$$z_t = \frac{q_t^2}{\sigma_{\text{CEV}}^2 T} = \frac{F_t^{2\beta_*}}{\beta_*^2 \sigma_{\text{CEV}}^2 T} \quad \text{and} \quad \nu = \frac{1}{2\beta_*}$$

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Simulation algorithm for $F_{t+h} \sim \text{CEV}(\beta, F_t, \rho_*^2 \sigma_t^2 h I_{t,t+h})$ from Kang (2014):

- 1 Sample $X \sim 2\Gamma\left(\frac{1}{2\beta_*}\right)$.
- 2 If $X \geq z_t = \frac{F_t^{2\beta_*}}{\rho_*^2 \sigma_t^2 h I_{t,t+h}}$, $F_{t+dt} = 0$ and terminate.
- 3 Else, sample $z_{t+dt} \sim 2\Gamma\left(\text{Pois}\left(\frac{z_t - X}{2}\right) + 1\right)$.
- 4 Obtain F_{t+dt} from $z_{t+dt} = \frac{F_{t+dt}^{2\beta_*}}{\rho_*^2 \sigma_t^2 h I_{t,t+h}}$.

Three simulation step

- Simulate σ_{t+h} from σ_t (trivial):

$$\sigma_{t+h} = \sigma_t \exp \left(\xi \sqrt{h} Z - \xi^2 h / 2 \right), \quad Z \sim N(0, 1).$$

- (1) Simulate $I_{t,t+h}$ given σ_t and σ_{t+h} .
 - Trapezoidal:
 - Lognormal:
 - Shifted lognormal:
- Simulate F_{t+h} given F_t , σ_{t+h} , and $I_{0,T}$.
 - For $0 < \beta < 1$, approximate F_{t+h} with a martingale-preserving CEV distribution.
 - Sample F_{t+h} based on [Kang \(2014\)](#)'s algorithm.

Numerical Results: Case 4. Accuracy of $I_{t,t+h}$

Cases	F_0	σ_0	ξ	ρ	β	T	K
Case IV	0.05	0.4	0.6	0	0.3	1	[0.02,0.1]

- With $\rho = 0$, CEV approximation is accurate.
- This case validates our LN approximation of $I_{t,t+h}$.
- With one jump ($h = 1$), our method is more accurate than existing methods.

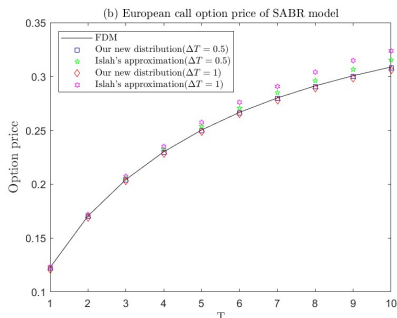
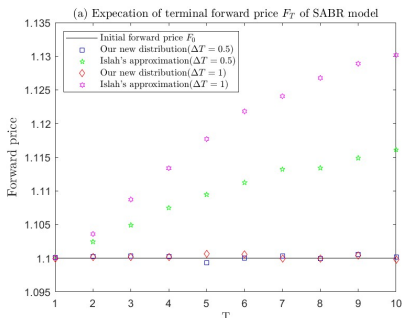
K/F_0	h	40%	80%	100%	120%	160%	200%	
Exact Price								
FDM		0.0456	0.0414	0.0394	0.0375	0.0339	0.0306	
Price Error ($\times 10^{-3}$)								CPU
Euler	1/400	1.60	1.50	1.50	1.40	1.30	1.20	49.40
	1/800	0.70	0.60	0.50	0.50	0.40	0.30	99.10
	1/1600	-0.30	-0.30	-0.30	-0.30	-0.30	-0.30	194.00
Low-bias (Chen et al., 2012)	1/4	0.50	0.50	0.50	0.40	0.40	0.40	78.40
	1/8	0.40	0.40	0.40	0.30	0.30	0.20	176.00
Exact (Cai et al., 2017)	1	0.10	0.20	0.20	0.20	0.20	0.20	98.30
LN	1	-0.01	-0.01	0.00	-0.02	-0.03	-0.04	0.03

Note: “Euler”, “Low-bias”, and “Exact” results are from other papers.

Numerical Results: Case 6. Martingale preservation

Cases	F_0	σ_0	ξ	ρ	β	T	K
Case VI	1.1	0.3	0.5	-0.8	0.4	[1,10]	1.1

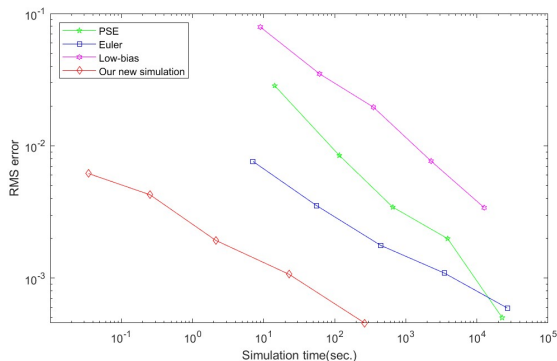
- We measure $E(F_T)$ and $E((F_T - K)^+)$ as functions of T .
- Our CEV approximation for F_{t+h} preserves martingale ($E(F_T) = F_0$).
- Naturally, option price is more accurate.



Numerical Results: Case 5. RMS v.s. CPU time

Cases	F_0	σ_0	ξ	ρ	β	T	K
Case V	1.1	0.4	0.8	-0.3	0.3	4	1.1

- We plot RMSE vs CPU time.
- For the same level of error, our new method is at least 100 times faster than existing methods.



Numerical Results: Case 1

Cases	F_0	σ_0	ξ	ρ	β	T	K
Case I	1	0.25	0.3	-0.8	0.3	10	[0.2,2]

- Our methods (LN, S-LN, TZ) accurately price challenging cases ($\xi = 0.3$ and $T = 10$.)

K	0.20	0.40	0.80	1.00	1.20	1.60	2.00
FDM	0.843	0.689	0.406	0.285	0.183	0.053	0.011
Error ($\times 10^{-3}$) for $h = 1$							
LN	-1.17	-1.45	-0.35	0.49	1.23	1.60	1.22
S-LN	-1.22	-1.49	-0.37	0.49	1.28	1.72	1.32
TZ	-3.13	-3.41	-2.15	-1.11	-0.12	0.79	0.65
Stdev	1.91	1.73	1.33	1.13	0.94	0.52	0.28
Error ($\times 10^{-3}$) for $h = 1/2$							
LN	-0.17	-0.09	0.55	0.91	1.22	1.15	0.80
S-LN	-0.01	0.10	0.78	1.16	1.43	1.36	0.98
TZ	-1.26	-1.10	-0.25	0.26	0.66	0.75	0.54
Stdev	1.83	1.65	1.26	1.06	0.85	0.50	0.27
Error ($\times 10^{-3}$) for $h = 1/4$							
LN	-0.35	-0.15	0.32	0.53	0.66	0.59	0.51
S-LN	-0.46	-0.24	0.22	0.42	0.56	0.56	0.48
TZ	-0.60	-0.36	0.12	0.31	0.41	0.41	0.33
Stdev	2.34	2.12	1.63	1.36	1.10	0.61	0.31

Numerical Results: Case 2

Cases	F_0	σ_0	ξ	ρ	β	T	K
Case II	1	0.25	0.3	-0.5	0.6	10	[0.2,2]

- Our methods (LN, S-LN, TZ) accurately price challenging cases ($\xi = 0.3$ and $T = 10$.)

K	0.20	0.40	0.80	1.00	1.20	1.60	2.00
FDM	0.829	0.670	0.398	0.291	0.207	0.100	0.050
Error ($\times 10^{-3}$) for $h = 1$							
LN	-0.04	-0.22	-0.30	-0.29	-0.30	-0.33	-0.25
S-LN	-0.14	-0.30	-0.42	-0.43	-0.43	-0.40	-0.30
TZ	-0.24	-0.38	-0.43	-0.37	-0.32	-0.30	-0.27
Stdev	2.11	1.98	1.73	1.61	1.44	1.11	0.91
Error ($\times 10^{-3}$) for $h = 1/2$							
LN	-0.34	-0.43	-0.45	-0.43	-0.39	-0.31	-0.21
S-LN	0.04	-0.05	-0.09	-0.09	-0.10	-0.11	-0.04
TZ	-0.57	-0.64	-0.59	-0.53	-0.48	-0.40	-0.27
Stdev	2.36	2.21	1.83	1.62	1.42	1.06	0.82
Error ($\times 10^{-3}$) for $h = 1/4$							
LN	0.37	0.27	0.15	0.12	0.07	-0.02	-0.02
S-LN	0.45	0.37	0.27	0.20	0.10	-0.02	0.00
TZ	0.08	-0.01	-0.05	-0.04	-0.05	-0.12	-0.11
Stdev	2.44	2.31	1.98	1.76	1.53	1.12	0.86

Numerical Results: Case 3

Cases	F_0	σ_0	ξ	ρ	β	T	K
Case III	1	0.25	0.3	-0.2	0.9	10	[0.2,2]

- Our methods (LN, S-LN, TZ) accurately price challenging cases ($\xi = 0.3$ and $T = 10$.)

K	0.20	0.40	0.80	1.00	1.20	1.60	2.00
FDM	0.816	0.652	0.393	0.303	0.236	0.151	0.105
Error ($\times 10^{-3}$) for $h = 1$							
LN	-0.83	-0.85	-0.79	-0.77	-0.75	-0.68	-0.65
S-LN	0.78	0.79	0.79	0.81	0.82	0.85	0.91
TZ	-0.63	-0.68	-0.63	-0.54	-0.50	-0.45	-0.45
Stdev	3.80	3.78	3.66	3.57	3.47	3.25	3.07
Error ($\times 10^{-3}$) for $h = 1/2$							
LN	-0.73	-0.77	-0.90	-0.89	-0.88	-0.83	-0.81
S-LN	0.50	0.48	0.42	0.39	0.38	0.44	0.47
TZ	-0.66	-0.66	-0.66	-0.62	-0.65	-0.70	-0.67
Stdev	4.94	4.89	4.72	4.61	4.50	4.30	4.17
Error ($\times 10^{-3}$) for $h = 1/4$							
LN	0.19	0.16	0.06	0.06	0.10	0.14	0.14
S-LN	-0.82	-0.85	-0.90	-0.92	-0.94	-0.99	-1.00
TZ	5.74	5.72	5.74	5.81	5.91	5.97	6.01
Stdev	4.31	4.24	4.04	3.96	3.88	3.75	3.63

Summary and Conclusion

- The SABR model is an SV model widely used in industry and academia.
- We propose an efficient simulation scheme for the SABR model.
 - Use the moment-matched lognormal sampling for the integrated variance ($I_{t,t+h}$).
 - Adopt [Kang \(2014\)](#)'s CEV simulation algorithm in the context of the SABR model.
- Our new algorithm outperforms the existing algorithm in terms of error-vs-speed tradeoff.

Python implementation: PyFENG

PyFENG (Python Financial **ENG**ineering)

- Implements standard quant finance methods: Black–Scholes, Bachelier (normal), CEV, SABR, Heston models and etc.
- **PyPI package**: `pip install pyfeng`
- Source: <https://github.com/PyFE/PyFENG/>
- Implemented in pure Python (no C/C++ extensions).

PyFengForPapers

- Inspired by [PapersWithCode](#) project.
- <https://github.com/PyFE/PyfengForPapers/>
- A collection of Jupyter Notebooks reproducing quant finance papers (mostly in derivative pricing and stochastic volatility) using the methods implemented in **PyFENG**.
- The code for this talk is also [available in PyFengForPapers](#).

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