

1 The Model

1.1 Generalized Least Squares (GLS) Regression

Consider the model of data

$$\begin{aligned} y &= X\beta + \epsilon \\ \epsilon &\sim \mathcal{N}(\epsilon|0, \Sigma) \end{aligned} \tag{1}$$

Assuming X is given, we have $y|X \sim \mathcal{N}(y|X, \beta, \Sigma)$ and conditional log-likelihood is:

$$l_{y|X}(\beta) = \log f(y|X; \beta) \propto -\frac{1}{2}(y - X\beta)^T \Sigma^{-1}(y - X\beta)$$

To find $\hat{\beta}^{MLE} = \operatorname{argmax}_{\beta} l_{y|X}(\beta)$ we solve

$$\nabla_{\beta} l_{y|X}(\beta) = -X^T \Sigma^{-1} X \beta + X^T \Sigma^{-1} y = 0$$

Hence

$$\hat{\beta}^{MLE} = (X^T \Sigma^{-1} X)^{-1} X^T \Sigma^{-1} y \tag{2}$$

2 Implementation for multi-variate return time series

Consider the problem

$$\begin{aligned} R_{t,i} &= \sum_{k=1}^K \beta_k A_{t,i}^{(k)} + \epsilon_{t,i}, \\ (t,i) &\in \{(1, i_{1,1}), \dots, (1, i_{1,N_1}), \\ &\quad \dots \\ &\quad (T, i_{T,1}), \dots, (T, i_{T,N_T})\} \end{aligned} \tag{3}$$

In the above:

- $R_{t,i}$ are returns for stock i over period t ;
- $A_{t,i}^{(k)}$ are features, usually alphas or their lagged versions.

Note that indices are 2-dimensional, but not necessarily rectangular, i.e. set of stocks can be different on different days. For notation convenience we introduce daily (random) vectors where we put all variables with same index t in a column¹:

$$R_{t,\cdot} = \begin{bmatrix} R_{t,i_{t,1}} \\ \vdots \\ R_{t,i_{t,N_t}} \end{bmatrix}, A_{t,\cdot}^{(k)} = \begin{bmatrix} A_{t,i_{t,1}}^{(k)} \\ \vdots \\ A_{t,i_{t,N_t}}^{(k)} \end{bmatrix}, \epsilon_{t,\cdot} = \begin{bmatrix} \epsilon_{t,i_{t,1}} \\ \vdots \\ \epsilon_{t,i_{t,N_t}} \end{bmatrix} \tag{4}$$

and we can further stack daily vectors:

¹note that dimensions for vectors with different index t might be different

$$R = \begin{bmatrix} R_{1,\cdot} \\ \vdots \\ R_{T,\cdot} \end{bmatrix}, A^{(k)} = \begin{bmatrix} A_{1,\cdot}^{(k)} \\ \vdots \\ A_{T,\cdot}^{(k)} \end{bmatrix}, \epsilon = \begin{bmatrix} \epsilon_{1,\cdot} \\ \vdots \\ \epsilon_{T,\cdot} \end{bmatrix} \quad (5)$$

so we can rewrite Equation 3 in vector form:

$$R = A\beta + \epsilon, \epsilon \sim \mathcal{N}(\epsilon|0, \Sigma) \quad (6)$$

$$A = [A^{(1)} \dots A^{(k)}] \quad (7)$$

3 Exploting special structure of errors ϵ

In practice, we can assume special structure of Σ that let us simplify the implementation:

- errors for different time periods are uncorrelated $cov(\epsilon_{t_1,i}, \epsilon_{t_2,j}) = 0$ if $t_1 \neq t_2$.

$$\Sigma = \begin{bmatrix} \Sigma_1 & 0 & \dots & 0 \\ 0 & \Sigma_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \Sigma_T \end{bmatrix} \quad (8)$$

- errors in the same period t have covariance matrix Σ_t that has a "low-rank plus diagonal" structure

$$\Sigma_t = L_t \Sigma_t^f L_t^T + \Lambda_t. \quad (9)$$

We can exploit these properties to simplify Equation 2:

- inverse of Σ is

$$\Sigma^{-1} = \begin{bmatrix} \Sigma_1^{-1} & 0 & \dots & 0 \\ 0 & \Sigma_2^{-1} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \Sigma_T^{-1} \end{bmatrix} \quad (10)$$

- each Σ_t can be inverted using Woodbury inversion formula:

$$\Sigma_t^{-1} = (L_t \Sigma_t^f L_t^T + \Lambda_t)^{-1} = \Lambda_t^{-1} + \Lambda_t^{-1} L_t ((\Sigma_t^f)^{-1} + L^T \Lambda_t^{-1} L) L_t^T \Lambda_t^{-1} \quad (11)$$

Thus computation Equation 2 reduces to:

$$\beta^{GLS} = \left(\sum_{t=1}^T \right)^{-1} \quad (12)$$