Limited evidence of test-retest reliability in infant-directed speech preference in a large pre-registered infant sample 2 Melanie S. Schreiner^{1,2}, Christina Bergmann³, Michael C. Frank⁴, Tom Fritzsche⁵, Nayeli Gonzalez-Gomez⁶, Kiley Hamlin⁷, Natalia Kartushina⁸, Danielle J. Kellier⁴, Nivedita Mani¹, Julien Mayor⁸, Jenny Saffran⁹, Melanie Soderstrom¹⁰, Mohinish Shukla¹¹, Priya Silverstein^{12,13}, Martin Zettersten^{9,14}, & Matthias Lippold^{1,2} ¹ University of Goettingen ² Leibniz Science Campus PrimateCognition ³ Max Planck Insitute for Psycholinguistics ⁴ Stanford University 10 ⁵ University of Potsdam 11 ⁶ Oxford Brookes University 12 ⁷ University of British Columbia 13 ⁸ University of Oslo ⁹ University of Wisconsin-Madison ¹⁰ University of Manitoba ¹¹ Università di Padova 17 ¹² Lancaster University 18 ¹³ Center for Open Science 19

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37 Abstract

Test-retest reliability — establishing that measurements remain consistent across multiple testing sessions — is critical to measuring, understanding, and predicting individual 39 differences in infant language development. However, previous attempts to establish 40 measurement reliability in infant speech perception tasks are limited, and reliability of 41 frequently-used infant measures is largely unknown. The current study investigated the 42 test-retest reliability of infants' preference for infant-directed speech (hereafter, IDS) over adult-directed speech (hereafter, ADS) in a large sample (N=158) in the context of the ManyBabies1 collaborative research project (hereafter, MB1; Frank et al., 2017; ManyBabies Consortium, 2020). Labs of the original MB1 study were asked to bring in participating infants for a second appointment retesting infants on their IDS preference. This approach allows us to estimate test-retest reliability across three different methods used to investigate preferential listening in infancy: the head-turn preference procedure, 49 central fixation, and eye-tracking. Overall, we find no consistent evidence of test-retest 50 reliability in measures of infants' speech preference (overall r = .09, 95% CI [-.06,.25]). 51 While increasing the number of trials that infants needed to contribute for inclusion in the 52 analysis revealed a numeric growth in test-retest reliability, it also considerably reduced the 53 study's effective sample size. Therefore, future research on infant development should take into account that not all experimental measures may be appropriate for assessing 55 individual differences between infants.

Keywords: language acquisition; speech perception; infant-directed speech; adult-directed speech; test-retest reliability

59 Word count: 3998

Limited evidence of test-retest reliability in infant-directed speech preference in a large

pre-registered infant sample

Obtaining a quantitative measure of infants' cognitive abilities is an extraordinarily difficult endeavor. The most frequent way to assess what infants know or prefer is to track overt behavior. However, measuring overt behavior at early ages presents many challenges: participants' attention span is short, they do not follow instructions, their mood can change instantly, and their behavior is often inconsistent. Therefore, most measurements are noisy and the typical sample size of an infant study is small (around 20 infants per group), resulting in low power (Oakes, 2017). In addition, there is individual and environmental variation that may add even more noise to the data (e.g., Johnson & Zamuner, 2010). Despite these demanding conditions, reliable and robust methods for assessing infants' behavior are critical to understanding development.

In order to address these challenges, the ManyBabies collaborative research 72 consortium was formed to conduct large-scale, conceptual, consensus-based replications of 73 seminal findings to identify sources of variability and establish best practices for experimental studies in infancy (Frank et al., 2017). The first ManyBabies collaborative research project (hereafter, MB1, ManyBabies Consortium, 2020) explored the reproducibility of the well-studied phenomenon that infants prefer infant-directed speech 77 (hereafter, IDS) over adult-directed speech (hereafter, ADS, Cooper & Aslin, 1990). Across 78 many different cultures, infants are commonly addressed in IDS, which typically is characterized by higher pitch, greater pitch range, and shorter utterances, compared to the language used between interacting adults (Fernald et al., 1989). A large body of behavioral studies finds that infants show increased looking times when hearing IDS compared to ADS stimuli across ages and methods (Cooper & Aslin, 1990; see Dunst, Gorman, & Hamby, 2012 for a meta-analysis). This attentional enhancement is also documented in neurophysiological studies showing increased neural activation during IDS compared to

ADS exposure (Naoi et al., 2012; Zangl & Mills, 2007). IDS has also been identified as facilitating early word learning. In particular, infants' word segmentation abilities (Floccia et al., 2016; Schreiner & Mani, 2017; Singh, Nestor, Parikh, & Yull, 2009; Thiessen, Hill, & Saffran, 2005) and their learning of word-object associations (Graf Estes & Hurley, 2013; Ma, Golinkoff, Houston, & Hirsh-Pasek, 2011) are enhanced in the context of IDS. In sum, several lines of evidence suggest that IDS is beneficial for early language development.

Within MB1, 67 labs contributed data from 2,329 infants showing that babies 92 generally prefer to listen to IDS over ADS. Nevertheless, the overall effect size of d=0.3593 was smaller than a previously reported meta-analytic effect size of d = 0.67 (Dunst et al., 2012). The results revealed several additional factors that influenced the effect size. First, 95 older infants showed a larger preference of IDS over ADS. Second, the stimulus language was linked to IDS preference, with North American English learning infants showing a 97 larger IDS preference than infants learning other languages. Third, comparing the different methods employed, the head-turn preference procedure yielded the highest effect size, while the central fixation paradigm and eye-tracking methods revealed smaller effects. Finally, 100 exploratory analyses assessed the effect of different inclusion criteria. Across methods, 101 using stricter inclusion criteria led to an increase in effect sizes despite the larger 102 proportion of excluded participants (see also Byers-Heinlein, Bergmann, & Savalei, 2021). 103

However, there is a difference between a result being reliable in a large sample of 104 infants and the measurement of an individual infant being reliable. In studies tracking 105 individual differences, the measured behavior during an experimental setting is often used 106 to predict a cognitive function or specific skill later in life. Individual differences research of this kind often has substantial implications for theoretical and applied work. For example, 108 research showing that infants' behavior in speech perception tasks can be linked to later 109 language development (see Cristia, Seidl, Junge, Soderstrom, & Hagoort, 2014 for a 110 meta-analysis) has the potential to identify infants at risk for later language delays or 111 disorders. However, a necessary precondition for this link to be observable is that 112

individual differences between infants can be measured with high reliability at these earlier stages, in order to ensure that measured inter-individual variation mainly reflects differences in children's abilities rather than measurement error. How reliable are the measures used in infancy research?

Previous attempts to address the reliability of measurements have typically been 117 limited to adult populations (Hedge, Powell, & Sumner, 2018), or have been conducted 118 with small sample sizes (e.g., Houston, Horn, Qi, Ting, & Gao, 2007). For example, 119 Colombo, Mitchell, and Horowitz (1988) used a paired-comparison task, in which infants 120 were familiarized with a stimulus and presented with the familiarized and a novel stimulus 121 side-by-side at test. Results indicated that infants' novelty preference was extremely 122 variable from task to task. Assessing infants' performance from one week to another 123 revealed that infants' attention measures were moderately reliable. However, reliability 124 seemed to increase with the number of tasks infants completed in the younger age group, 125 suggesting that reliability is influenced by the number of assessments. In addition, infants' 126 performance from 4 to 7 months was longitudinally stable but somewhat smaller than 127 week-to-week reliability. Cristia, Seidl, Singh, and Houston (2016) also retested infant 128 populations by independently conducting 12 different experiments on infant speech perception at three different labs with different implementations of the individual studies. Hence, it was only after completed data collection that the data was pooled together by the different labs revealing potential confounds. Nevertheless, the results showed that 132 reliability was extremely variable across the different experiments and labs and low overall 133 (meta-analytic r = .07).

Against this background, the current study investigates test-retest reliability of infants' performance in a speech preference task. Within MB1, a multi-lab collaboration, we examine whether infants' preferential listening behavior to IDS and ADS is reliable across two different test sessions. We also investigate the influence of various moderators on the reliability of IDS preference (e.g., time between test and retest; infants' language

background).

Our study was faced with a critical design choice: what stimuli to use to assess 141 test-retest reliability. One constraint on our study was that, since it was a follow-on to 142 MB1, any stimulus we used would always be presented after the MB1 stimuli. One option 143 would be simply to bring back infants and have them hear exactly the same stimulus 144 materials. A weakness of this design would be the potential for stimulus familiarity effects, 145 however, since infants would have heard the materials before. Further complicating 146 matters, infants might show a preference for or against a familiar stimulus depending on 147 their age (Hunter & Ames, 1988). The ideal solution then would be to create a brand new 148 stimulus set with the same characteristics. Unfortunately, because of the process of how 149 MB1 stimuli were created, we did not have enough normed raw recordings available to 150 make brand new stimulus items that conformed to the same standards as the MB1 stimuli. 151 We therefore chose an intermediate path: we reversed the ordering of MB1 stimuli. 152 Average looking times in MB1 were always lower than 9s per trial, even for the youngest 153 children on the earliest trials (the group who looked the longest on average), so most 154 children in MB1 did not hear the second half of most trials. Thus, by reversing the order, 155 we had a perfectly matched stimulus set that was relatively unfamiliar to most infants. The disadvantage of this design was that infants who looked longer might be more likely to 157 hear a familiar clip heard in the previous study. If infants then showed a familiarity 158 preference — an assumption which might not be true — the end result could be to inflate 159 our estimates of test-retest reliability slightly, since longer lookers would on average look 160 longer at retest due to their familiarity preference. We view this risk as relatively low, but 161 do note that it is a limitation of our design. 162

The current study also explores whether there are any differences in test-retest reliability between three widely used methods: central fixation (CF), eye-tracking (ET), and the head-turn preference procedure (HPP). Exploring differences in CF, ET, and HPP, Junge et al. (2020) provide experimental and meta-analytic evidence in favor of using the

HPP in speech segmentation tasks. Similarly, the MB1 project reported an increase in the effect size for HPP compared to CF and ET (ManyBabies Consortium, 2020). HPP 168 requires gross motor movements relative to other methods, such as CF and ET paradigms, 169 for which subtle eye movements towards a monitor located in front of the child are 170 sufficient. One possible explanation for the stronger effects with HPP may be a higher 171 sensitivity to the contingency of the presentation of auditory stimuli and infants' head 172 turns away from the typical forward-facing position. While these findings suggest that 173 HPP may be a more sensitive index of infant preference, they do not necessarily imply 174 higher reliability for individual infants' performance using HPP. For example, Marimon 175 and Höhle (2022) found no evidence for test-retest reliability when testing infants' prosodic 176 preferences using the HPP method. It remains an open question whether the same 177 measures that produce larger effect sizes at the group-level also have higher test-retest reliability for individual infants (Byers-Heinlein, Bergmann, et al., 2021). Therefore, 179 assessing the test-retest reliability of the different preference measures is crucial, so that 180 researchers can make informed decisions about the appropriate methods for their particular 181 research question. Critically, only measures with high test-retest reliability should be used 182 for studies of individual differences.

184 Method

185 Preregistration

Prior to the start of data collection, we preregistered the current study on the Open Science Framework (https://osf.io/v5f8t; see S1 in the Supplementary Materials for details).

Data Collection

A call was issued to all labs participating in the original MB1 study on January 24th,
2018 (ManyBabies Consortium, 2020). The collection of retest session data was initially set
to end on May 31st, 2018, one month after the end date of the original MB1 project. Due
to the fact that the original MB1 project extended the time frame for data collection and
the late start of data collection for the MB1 test-retest study, we also allowed participating
labs to continue data collection past the scheduled end date.

196 Participants

Contributing labs were asked to re-recruit their monolingual participants between the ages of 6 to 12 months who had already participated in the MB1 project. If participating labs had not committed to testing either of these age groups, they were also allowed to re-recruit participants from the youngest age group of 3- to 6-month-olds and/or the oldest age group of 12- to 15-month-olds. Labs were asked to contribute half (n=16) or full samples (n=32); however, a lab's data was included in the study regardless of the number of included infants. The study was approved by each lab's respective ethics committee and parental consent was obtained for each infant prior to participation in the study.

Our final sample consisted of 158 monolingual infants from 7 different labs (Table 1).

In order to be included in the study, infants needed a minimum of 90% first language

exposure, to be born full term with no known developmental disorders, and normal hearing

and vision. We excluded 11 participants due to session errors and 11 participants who did

not have at least one valid trial per condition (IDS and ADS) at their first or second

session. The mean age of infants included in the study was 245 days (range: 108 – 373

days).

Materials

Visual stimuli. The visual stimuli and instructions were identical to MB1. For the
CF paradigm and ET, labs used a multicolored static checkerboard as the fixation stimulus
as well as a multicolored moving circle with a ringing sound as an attention-getter between
trials. For the HPP method, labs used their standard procedure, as in MB1.

Speech stimuli. We used the identical training stimuli of piano music from MB1.

A second set of naturalistic IDS and ADS recordings of mothers either talking to their

infant or to an experimenter was created for the retest session by reversing the order of

clips within each sequence of the original study. This resulted in eight new sequences of

natural IDS and eight new sequences of natural ADS with a length of 18 seconds each.

Procedure. Infants were retested using the identical procedure as during the first testing day: CF, HPP, or ET. Participating labs were asked to schedule test and retest sessions 7 days apart with a minimum number of 1 day and a maximum number of 31 days. However, infants whose time between test and retest exceeded 31 days were still included in the analyses (n = 3). The mean number of days between test and retest was 10 (range: 1 - 49).

A total of 18 trials, including two training, eight IDS, and eight ADS trials, were
presented in one of four pseudo-randomized orders. Trial length was either infant-controlled
or fixed depending on the lab's standard procedure: a trial stopped either if the infant
looked away for 2 seconds or after the total trial duration of 18 seconds. The online coding
experimenter and the parent listened to music masked with the stimuli of the study via
noise-cancelling headphones. If the experimenter was in an adjacent room separate from
the testing location, listening to masking music was optional for the experimenter.

Data exclusion. A child was excluded if they had a session error, i.e., an
experimenter error (e.g., inaccurate coding, or presentation of retest stimuli on the first
test session) or equipment failure (visual stimuli continued to play after the end of a trial).

Table 1
Statistics of the included labs. n refers to the number of infants included in the final analysis.

Lab	Method	Language	Mean age (days)	N
babylab-potsdam	HPP	German	227	22
babyling-oslo	eye-tracking	Norwegian	249	10
brookes-babylab	central fixation	English	267	18
InfantCog-UBC	central fixation	English	147	7
infantll-madison	HPP	English	230	30
lancslab	eye-tracking	English	236	16
wsi-goettingen	central fixation	German	280	39
wsi-goettingen	HPP	German	242	16

Trials were excluded if they were marked as trial errors, i.e., if the infant was reported as
fussy, an experimental or equipment error occurred, or there was parental interference
during the task (e.g., if the parent spoke with the infant during the trial). Trials were also
excluded if the minimum looking time of 2 s was not met. If a participant was unable to
contribute at least one IDS and one ADS trial for either test or retest, all data of that
participant was excluded from the test-retest analyses.

Results Results

IDS preference

First, we examined infants' preference for IDS in both sessions. Two-samples t-tests comparing the difference in average looking time between IDS and ADS to zero revealed that infants showed a preference of IDS over ADS in Session 1, t(157) = 6.47, p < .001, and

Trial type	Session 1 Mean	Session 1 SD	Session 2 Mean	Session $2 SD$
ADS	7.72	2.77	6.96	2.92
IDS	8.76	2.85	7.75	2.75

Table 2 Average looking times (in seconds) for each session and condition

Session 2, t(157) = 4.19, p < .001, replicating the main finding from MB1 (Table 2). 68.35% of infants in Session 1 and 63.29% of infants in Session 2 showed a preference for IDS. In order to test whether there was a difference in the strength of the preference effect across sessions, we fit a linear mixed-effects model predicting infants' average difference in 252 looking time between IDS and ADS from test session (1 vs. 2), including by-lab and 253 by-participant random intercepts. There was no significant difference in the magnitude of 254 infants' preference between the two sessions, β =-0.30, SE=0.24, p=.208.

Reliability

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We assessed test-retest reliability in two ways. First, we fit a linear mixed-effects 257 model predicting IDS preference in Session 2 from IDS preference in Session 1, including a 258 by-lab random intercept. The results revealed no significant relationship between IDS 259 preference in Session 1 and 2 (Table 3). Second, we calculated the Pearson correlation 260 coefficient. While a simple correlation coefficient might overestimate the test-retest 261 reliability in our sample because it does not control for the differences between different labs and methods (HPP, CF, and ET), we felt it was important to also conduct a Pearson correlation as it is commonly used to assess reliability. The size of the correlation 264 coefficient was not statistically different from zero and the estimate was small, r = .09, 95%265 CI [-.06, .25], t(156) = 1.19, p = .237. Moreover, no significant correlations emerged in 266 each sample considered separately (Figure 1; see Supplementary Materials S3 for a 267

Table 3

Coefficient estimates from a linear mixed effects model predicting IDS preference in Session 2.

	Estimate	SE	t	р
Intercept	0.87	0.46	1.92	0.10
IDS Preference Session 1	0.04	0.09	0.41	0.68

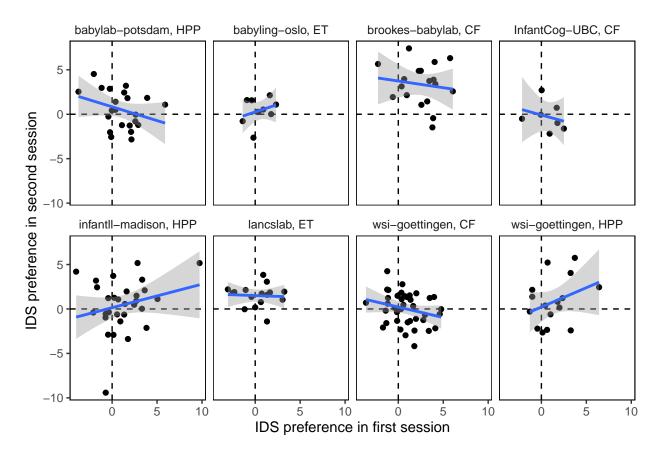


Figure 1. Correlation between IDS Preference in Session 1 and Session 2 in each lab and method. Dots indicate individual participants. Error bands represent 95 percent confidence intervals. The dashed line indicates no preference (i.e., a value of zero) for the first and second session, respectively.

Table 4

Coefficient estimates from a linear mixed effects model predicting IDS preference in Session 2 and Pearson correlation coefficient for each method separately.

Method	beta	SE	p	Pearson r
HPP	0.15	0.14	0.28	0.13
ET	0.03	0.16	0.84	0.02
CF	-0.20	0.12	0.12	0.08

meta-analytic approach). 41.77% of the infants reversed their direction of preference for IDS versus ADS from the test to the retest session.

To investigate the test-retest reliability of each specific method, we computed Pearson 270 correlation coefficients and the same mixed-effects model described above for HPP, CF, 271 and ET separately (Table 4). None of the three methods showed evidence of test-retest 272 reliability. Neither the Pearson correlation coefficients nor the coefficients of the multilevel 273 analysis were significant, all p-values > 0.12. In planned secondary analyses, we found that 274 time between test sessions, participant age, and language background did not moderate the 275 relationship between IDS preference in session 1 and session 2 (see Supplementary 276 Materials S2). Taken together, we find no significant evidence of test-retest reliability 277 across our preregistered analyses. 278

79 Results with different inclusion criteria

To this point, all analyses were performed using the inclusion criteria from MB1,
which required only that infants contribute at least one trial per condition for inclusion
(i.e., one IDS and one ADS trial). However, more stringent inclusion criteria yielded larger

effect sizes in MB1. We therefore conducted exploratory analyses assessing test-retest 283 reliability after applying progressively stricter inclusion criteria, requiring two, four, six, 284 and eight valid trials per condition. Applying stricter criteria — and thereby increasing the 285 number of test trials — increased reliability numerically from r = 0.07 to r = 0.34 (Figure 286 2). In part due to a decrease in sample size, only one of these correlations was statistically 287 significant (when requiring six trial pairs): two valid trial pairs, t(152) = 0.90, p = .367; 288 four valid trial pairs, t(143) = 1.03, p = .306; six valid trial pairs, t(98) = 2.23, p = .028; 280 eight valid trial pairs — all trials in both sessions — t(22) = 1.68, p = .108. The analyses 290 provide tentative evidence that stricter inclusion criteria may lead to higher test-retest 291 reliability, but at the cost of substantial decreases in sample size (see Supplementary 292 Materials S5 for additional analyses). 293

General Discussion

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The current study investigated the test-retest reliability of infants' preference for IDS 295 over ADS. We retested the IDS preference of infants participating in the original MB1 296 project to assess the extent to which their pattern of preference would remain consistent 297 across multiple testing sessions. While we replicated the original effect of infants' speech 298 preference for IDS over ADS for both the test and retest session on the group-level, we 299 found that infants' speech preference measures showed no evidence of test-retest reliability. 300 In other words, we were unable to detect stable individual differences in infants' preference 301 for IDS. This finding is consistent with past research suggesting low test-retest reliability in 302 other infant paradigms (Cristia et al., 2016). Given that most experimental procedures conducted in infant research are interested in the comparison of groups, individual differences between participants within a specific condition are usually minimized by the experimental procedure while differences between conditions are maximized. Therefore, 306 infant preference measures may be a good approach for capturing group-level phenomena, 307 but may be less appropriate for examining individual differences in development. 308

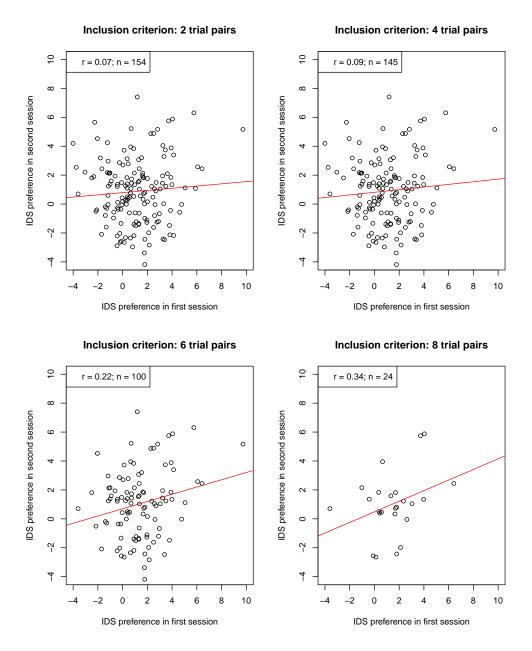


Figure 2. IDS preferences of both sessions plotted against each other for each inclusion criterion. n indicates the number of included infants, r is the Pearson correlation coefficient as the indicator for reliability.

Consistent with general psychometric theory (e.g., DeBolt, Rhemtulla, & Oakes, 309 2020), stricter inclusion criteria — and consequently a larger number of included test trials 310 per participant — tended to increase the magnitude of the correlation between test 311 sessions. However, this association was based on exploratory analyses and was in part only 312 observed descriptively, and hence should be interpreted with caution. A similar effect on 313 the group-level was found in the MB1 project, where a stricter inclusion criterion led to 314 bigger effect sizes (ManyBabies Consortium, 2020). As in MB1, higher reliability through 315 strict exclusions came at a high cost. In particular, with the strictest criterion, only a small 316 portion of the original sample size (24 out of 158 infants) could be included in the final 317 sample. In other words, applying stricter criteria leads to a higher drop-out rate and can 318 dramatically reduce the sample size. In the case of studies in the field of developmental 319 science, where there are many practical restrictions in collecting large samples of infants (e.g., birth rate in the area, restricted lab capacities, budget restrictions), a strict drop-out 321 criterion may often be difficult to implement. Note that studies in developmental science already have above-average drop-out rates (Miller, 2017). In addition, drop out may not be 323 random, and so having high drop-out rates can further limit the generalizability of a study. 324 In fact, the number of trials individual infants contributed was highly correlated between 325 test sessions in the current study (see Supplementary Materials S6). Particularly in the 326 context of turning individual differences measures into diagnostic tools, high drop-out rates 327 have an additional limitation of not being broadly usable. 328

An alternative approach to increasing the number of valid trials is to increase the
number of experimental trials. This approach seeks to increase the likelihood that
participants will contribute sufficient trials (after trial-level exclusions) to allow for precise
individual-level estimates (DeBolt et al., 2020; see also Silverstein, Feng, Westermann,
Parise, & Twomey, 2021). While this approach is promising, it may not always be feasible,
because the attention span of a typical infant participant is limited. Therefore, prolonging
the experimental procedure to maximize the absolute number of trials is often challenging

in practice. Other avenues for obtaining higher numbers of valid trials may include changes in the procedure (e.g., Egger, Rowland, & Bergmann, 2020) or implementing multi-day test sessions (Fernald & Marchman, 2012).

As our results are only based on the phenomenon of IDS preference (albeit, with 339 three widely used methods: HPP, CF, ET) it is essential to further assess the underlying 340 reliability of preferential looking measures within other areas of speech perception (Marimon & Höhle, 2022). While most infants prefer IDS over ADS (Dunst et al., 2012), patterns of preferential looking in other tasks (e.g., speech segmentation) are often inconsistent and difficult to predict (Bergmann & Cristia, 2016). These inconsistencies in looking behavior are especially important to consider in the context of relating a direction of preference to later language development, and can sometimes lead to seemingly contradictory findings. That is, both familiarity and novelty responses have been suggested 347 to be predictive of infants' later linguistic abilities (DePaolis, Vihman, & Keren-Portnoy, 348 2014; Newman, Ratner, Jusczyk, Jusczyk, & Dow, 2006; Newman, Rowe, & Ratner, 2016). 349 In light of our findings, researchers conducting longitudinal studies with experimental data 350 from young infants predicting future outcomes should be cautious, as there may be large 351 intra-individual variability affecting preference measurement. 352

353 Limitations

While we had an above-average sample size for a study in infant research, we were
unable to approach the number of participants collected within the original MB1 study. In
addition to a delayed call, the extra effort of having to schedule a second lab visit for each
participant and the fact that there were already other collaborative studies taking place
simultaneously (MB1B, Byers-Heinlein, Tsui, Bergmann, et al., 2021; MB1G,
Byers-Heinlein, Tsui, Van Renswoude, et al., 2021), might have contributed to a low
participation rate. A higher sample size and a larger number of participating labs from
different countries would have enabled us to conduct a more highly-powered test of

differences in test-retest reliability across different methods, language backgrounds, and participant age.

A further limitation concerns the stimuli. While the order of the audio recording clips 364 presented to infants within a given trial differed between the first and second session, the 365 exact same stimulus material as in MB1 was used in both sessions. In particular, all 366 children heard the exact same voices in Session 1 and in Session 2. From a practical point of view, this was the most straightforward solution for coordinating the experiment within the larger MB1 project. However, familiarity effects might have influenced infants' looking behavior. Infants with longer looking times in their first session might have had more 370 opportunity to recognize familiar audio clips in their second session. For infants with short looking times, familiar audio clips would only occur towards the end of second-session 372 trials, thus offering infants less opportunity to recognize voices from their first session. 373 Therefore, inconsistent familiarity with the stimulus material in the second session across 374 infants might have artificially lowered test-retest reliability. 375

376 Conclusion

Following the MB1 protocol, the current study could not detect test-retest reliability
in measures of infants' preference for IDS over ADS. Subsequent analyses provided
tentative evidence that stricter criteria for the inclusion of participants may enhance
test-retest reliability at the cost of high drop-out rates. Developmental studies relying on
stable individual differences between their participants need to consider the underlying
reliability of their measures, and we recommend a broader assessment of test-retest
reliability in infant research.

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