Limited evidence of test-retest reliability in infant-directed speech preference in a large pre-registered infant sample 2 Melanie S. Schreiner^{1,2}, Christina Bergmann³, Michael C. Frank⁴, Tom Fritzsche⁵, Nayeli Gonzalez-Gomez⁶, Kiley Hamlin⁷, Natalia Kartushina⁸, Danielle J. Kellier⁴, Nivedita Mani¹, Julien Mayor⁸, Jenny Saffran⁹, Melanie Soderstrom¹⁰, Mohinish Shukla¹¹, Priya Silverstein^{12,13}, Martin Zettersten^{9,14}, & Matthias Lippold¹ ¹ University of Goettingen ² Leibniz Science Campus PrimateCognition ³ Max Planck Insitute for Psycholinguistics ⁴ Stanford University 10 ⁵ University of Potsdam 11 ⁶ Oxford Brookes University 12 ⁷ University of British Columbia 13 ⁸ University of Oslo 14 ⁹ University of Wisconsin-Madison ¹⁰ University of Manitoba ¹¹ Università di Padova 17 ¹² Lancaster University 18 ¹³ Center for Open Science

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35 Abstract

The ManyBabies1 collaborative research project (hereafter, MB1; Frank et al., 2017;

ManyBabies Consortium, 2020) explores the reproducibility of the well-studied and robust

phenomenon of infants' preference of infant-directed speech (hereafter, IDS) over

adult-directed speech (hereafter, ADS; Cooper & Aslin, 1990). The current study is a

follow-on project aiming at further investigating the test-retest reliability of infant speech

preference measures. In particular, labs of the original study were asked to bring in tested

babies for a second appointment retesting infants on their IDS preference. This allows us

to estimate test-retest reliability for the three different methods used to investigate

preferential listening in infancy: the head-turn preference procedure, central fixation, and

eye-tracking. Our results suggest that the test-retest reliability of infants' speech preference

measures is rather low. While increasing the number of trials that infants needed to

contribute for inclusion in the analysis from 2 to 8 trial pairs revealed a descriptive growth

in test-retest reliability, it also considerably reduced the study's effective sample size.

Therefore, future research on infant development should take into account that not all

50 experimental measures might be appropriate to assess individual differences between

infants, and hence, the interpretation of findings needs to be treated with caution.

Keywords: language acquisition; speech perception; infant-directed speech;

adult-directed speech, test-retest reliability

Word count: 4506

Limited evidence of test-retest reliability in infant-directed speech preference in a large pre-registered infant sample

Obtaining a quantitative measure of young infant and children's cognitive abilities is
an extraordinarily difficult endeavor; the most frequent way to assess what they know or
can do is done via tracking overt behavior. However, the attention span of this specific
group of participants is rather short, they do not follow instructions, their mood can
change instantly and their behavior can be best described as unstable and volatile.
Therefore, most measurements are noisy and the typical sample size of a study is often just
around 20 infants per group resulting in low power (Oakes, 2017). In addition, there is
individual and environmental variation that may add even more noise to the data (e.g., as
discussed by Johnson & Zamuner, 2010). Despite these demanding conditions, infancy
researchers, nevertheless, need to have reliable and robust ways to assess infants' behavior.

In order to address these challenges in infant research, the ManyBabies collaborative research project was launched (Frank et al., 2017). Within a collaborative framework, the aim of this initiative is to conduct large-scale conceptual, consensus-based replications of seminal findings to identify sources of variability and establish best practices for experimental studies in infancy.

The fact that there is a substantial amount of literature on infants' preference of infant-directed speech (hereafter, IDS) over adult-directed speech (hereafter, ADS, Cooper & Aslin, 1990) led to the decision of exploring the reproducibility of this well-studied phenomenon within the first ManyBabies collaborative research project (hereafter, MB1, ManyBabies Consortium, 2020). Across many different cultures, infants are commonly addressed in IDS, which typically is characterized by higher pitch, greater pitch range, and shorter utterances, compared to the language used between interacting adults (Fernald et al., 1989). A large body of behavioral studies finds that across ages and methods, infants show increased looking times when hearing IDS compared to ADS stimuli (Cooper & Aslin,

1990; see Dunst, Gorman, & Hamby, 2012 for a meta-analysis). This attentional
enhancement is also documented in neurophysiological studies showing increased neural
activation during IDS compared to ADS exposure (Naoi et al., 2012; Zangl & Mills, 2007).
In addition to the heightened attention, IDS has also been identified as facilitating early
word learning. In particular, infants' word segmentation abilities Thiessen, Hill, & Saffran
(2005) and their learning of word-object associations (Graf Estes & Hurley, 2013; Ma,
Golinkoff, Houston, & Hirsh-Pasek, 2011) seems to be enhanced in the context of IDS. In
sum, IDS seems to be beneficial for early language development.

Within MB1, altogether 67 labs contributed data of 2.329 infants showing that babies 89 generally prefer to listen to IDS over ADS. Nevertheless, the overall effect size d=0.35was much smaller than the meta-analytic effect size of d = 0.67 reported by Dunst et al. 91 (2012). The results revealed a number of additional factors that influenced the effect size. First, there was a developmental change with older infants showing larger preferences. Second, the stimulus language was also linked to the speech preference with North American English learning infants being more likely to prefer IDS over ADS than infants learning other languages. Third, comparing the different methods employed, the head-turn preference procedure yielded the highest effect size, followed by the central fixation paradigm, with eye-tracking revealing the smallest effect size. Finally, exploratory analyses assessed the effect of different inclusion criteria. Across methods, using stricter inclusion criteria led to an increase in effect sizes despite the larger proportion of excluded 100 participants (see also Byers-Heinlein, Bergmann, & Savalei, 2021, for the relation between 101 sample size, effect size, and exclusion criteria). 102

However, there is a difference between a result being reliable in a large sample of infants and the individual measure of an individual infant being reliable. In studies tracking individual differences, the measured behavior during an experimental setting is often used to predict a cognitive function or specific skill later in life. A precondition for this link to be observable is that inter-individual differences between infants do exist and

can be detected with a high reliability at these earlier stages. However, how reliable are the measures used in infancy research?

A salient example of this line of individual differences research is the literature 110 showing that infants' behavior in speech perception tasks can be linked to later language 111 development (see Cristia, Seidl, Junge, Soderstrom, & Hagoort, 2014 for a meta-analysis), 112 potentially identifying infants at risk for later language delays or disorders. This finding 113 thus has substantial implications for theoretical and applied work. However, to be able to 114 make such claims robustly would require a high reliability within infants across multiple 115 testing sessions. Previous attempts to address the reliability of measurements are either 116 limited to adult populations investigating various tasks (Hedge, Powell, & Sumner, 2018), 117 or have been conducted with very small sample sizes (e.g., Houston, Horn, Qi, Ting, & 118 Gao, 2007). 119

In one example of such a study, Colombo, Mitchell, and Horowitz (1988) used a 120 paired comparison task, in which infants were familiarized with a stimulus and for the test 121 trials presented with the familiarized and a novel stimulus side-by-side. Results indicated 122 that infants' novelty preference was extremely variable from task to task. Assessing infants' 123 performance from one week to another revealed that infants' attention measures were 124 moderately reliable. However, reliability seemed to increase with the number of tasks 125 infants were able to complete in the younger age group suggesting that reliability is influenced by the number of assessments. In addition, infants' performance from 4 to 7 months was longitudinally stable but somewhat smaller than the week-to-week reliability. 128

Cristia, Seidl, Singh, and Houston (2016) also retested infant populations by independently conducting 12 different experiments on infant speech perception at three different labs with different implementations of the individual studies. Hence, it was only after completed data collection that the data was pooled together by the different labs revealing potential confounds. Nevertheless, the results showed that reliability was

extremely variable across the different experiments and labs and overall low.

Against this background, the current study investigates test-retest reliability of infants' performance in an auditory preference task. Within MB1, a multi-lab collaboration, we examine whether infants' preferential listening behavior to IDS and ADS is reliable across two different test sessions. We also aim to address whether time between test and retest or infants' language background influences the reliability of the preference measure.

Our study was faced with a critical design choice: what stimulus to use to assess 141 test-retest reliability. A first constraint of our study was that, since it was a follow-on to 142 MB1, any stimulus we used would always be presented after the MB1 stimuli. One option 143 would be simply to bring back infants and have them hear exactly the same stimulus 144 materials. A weakness of this design would be the potential for stimulus familiarity effects, 145 however, since infants would have heard the materials before. Further complicating 146 matters, infants might show a preference for or against a familiar stimulus depending on 147 their age (Hunter & Ames, 1988). The ideal solution then would be to create a brand new 148 stimulus set with the same characteristics. Unfortunately, because of the process how MB1 stimuli were created, we did not have enough normed raw recordings available to make 150 brand new stimulus items that conformed to the same standards as the MB1 stimuli. 151

Given this situation, we chose an intermediate path: we reversed the ordering of MB1 stimuli. Average looking times in MB1 were always lower than 9s per trial, even for the youngest children on the earliest trials (the group who looked the longest on average). So most children in MB1 did not hear the second half of most trials. Thus, by reversing the order, we had a perfectly matched stimulus set that was relatively unfamiliar to most infants. The disadvantage of this design was that infants who looked longer might be more likely to hear a familiar clip that they had heard in the previous study. If infants then showed a familiarity preference – an assumption which might not be true – the end result

could be to inflate our estimates of test-retest reliability slightly, since longer lookers would on average look a bit longer at retest due to their familiarity preference. We view this risk as relatively low, but do note that it is a limitation of our design.

The MB1 study was also interested in effects of experimental method on 163 infant-directed speech preference. Using central fixation (hereafter, CF), eye-tracking, and 164 the head-turn preference procedure (hereafter, HPP), the current study also explores 165 whether there are any differences in test-retest reliability between the three widely used 166 methods. Exploring differences in CF, eye-tracking, and HPP, Junge et al. (2020) provide 167 some evidence in favor of using the HPP in speech segmentation tasks. In addition, their 168 re-analysis of the meta-analysis on infant speech segmentation conducted by Bergmann and 169 Cristia (2016) also suggested numerically larger effect sizes for HPP than for eye-tracking. 170 Similarly, the MB1 project reported an increase in the effect size for HPP compared to CF and eye-tracking (ManyBabies Consortium, 2020). HPP requires gross motor movements 172 relative to other methods, such as CF and eye-tracking paradigms, for which subtle eye 173 movements towards a monitor located in front of the child are sufficient. One possible 174 explanation for the stronger effects with HPP may be a higher sensitivity to the 175 contingency of the presentation of auditory stimuli and infants' head turns away from the 176 typical forward-facing position. While these findings suggest that HPP may be a more 177 sensitive index of infant preference, they do not necessarily imply higher reliability for 178 individual infants' performance using the different methods. Hence, it remains an open 179 question whether the same measures that produce larger effect sizes at the group-level also 180 have higher test-retest reliability for individual infants (Byers-Heinlein, Bergmann, et al., 181 2021). Therefore, assessing the test-retest reliability of the different preference measures is 182 crucial, so that researchers can make informed decisions about the appropriate methods for 183 their particular research question. Critically, only measures with high test-retest reliability 184 should be used for studies of individual differences. 185

Preregistration. Prior to the start of data collection, we preregistered the current

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study on the Open Science Framework (OSF; see https://osf.io/v5f8t). We also had
secondary planned analyses which were not run in the end as we had fewer participants
than originally expected.

190 Method

A call was issued to all labs participating in the original MB1 study on January 24th,
2018 (ManyBabies Consortium, 2020). The collection of retest session data was initially set
to end on May 31st, 2018, one month after the end date of the original MB1 project. Due
to the fact that the original MB1 project extended the time frame for data collection and
the late start of data collection for the MB1 test-retest study, we also allowed participating
labs to continue data collection past the scheduled end date.

97 Participants

Contributing labs were asked to re-recruit their monolingual participants between the ages of 6 to 12 months who had already participated in the MB1 project. If participating labs had not committed to testing either of these age groups, they were also allowed to re-recruit participants from the youngest age group of 3- to 6-month-olds and/or the oldest age group of 12- to 15-month-olds. Labs were asked to contribute half (n=16) or full sample (n=32), however, a lab's data was included in the study regardless of the number of included infants. The study was approved by each lab's respective ethics committee and parental consent was obtained for each infant prior to participation in the study.

Our final sample consisted of 158 monolingual infants from 7 different labs. In order to be included in the study, infants needed a minimum of 90% first language exposure, to be born full term with no known developmental disorders, and normal hearing and vision.

We excluded 11 participants due to session errors and 11 participants who did not have at least one valid trial per condition (IDS and ADS) at their first or second session. The mean

age of infants included in the study was 245 days (range: 108 – 373 days). Further information on labs and participants are provided in Table 1.

213 Materials

Visual stimuli. The visual stimuli were identical to MB1. For the central fixation
paradigm and eye-tracking, labs were asked to use a multicolored static checkerboard as
the fixation stimulus as well as a multicolored moving circle with a ringing sound as an
attention getter to reorient infants toward the screen in between trials. Labs using the
HPP method were instructed to use their lab's standard procedure, as in MB1.

Speech stimuli. We used the identical training stimuli of piano music from MB1.

A second set of naturalistic IDS and ADS recordings of mothers either talking to their infant or to an experimenter was created for the retest session by reversing the order of clips within each sequence of the original study. This resulted in eight new sequences of natural IDS and eight new sequences of natural ADS with a length of 18 seconds each.

This was in order to prevent infants who still remembered the stimuli from their first test session from easily getting bored.

Procedure. Infants were retested using the identical procedure as during the first testing day: central fixation, HPP, or eye-tracking. Participating labs were asked to ideally schedule test and retest session 7 days apart with a minimum number of 1 day and a maximum number of 31 days. Three infants whose time between test and retest exceeded 31 days were also included in the analyses. The mean number of days between test and retest was 10 (range: 1 to 49 days).

A total of 18 trials, including two training, eight IDS, and eight ADS trials, were
presented in one of four pseudo-randomized orders. Trial length was either infant-controlled
or fixed depending on the lab's standard procedure, that is a trial stopped either if the
infant looked away for 2 seconds or after the total trial duration of 18 seconds. The online

coding experimenter and the parent listened to music masked with the stimuli of the study
via noise-cancelling headphones. If the experimenter was in an adjacent room separate
from the testing location, listening to masking music was optional for the experimenter.

Data exclusion. A child was excluded if they had a session error, i.e., an
experimenter error (e.g., inaccurate coding, or presentation of retest stimuli on the first
test session), or equipment failure (visual stimuli continued to play after the end of a trial).
Trials were excluded if they were marked as trial errors, i.e., if the infant was reported as
fussy, an experimental or equipment error occurred, or there was parental interference
during the task (e.g., if the parent spoke with the infant during the trial). Trials were also
excluded if the minimum looking time of 2 s was not met. If a participant was unable to
contribute at least one IDS and one ADS trial for either test or retest, all data of that
participant was excluded from the test-retest analyses.

Results Results

1DS preference

First, we examined infants' preferences for IDS in both sessions. Two two-samples 250 t-tests revealed that the children in Session 1, t(157) = 6.47, p < .001, and in Session 2, 251 t(157) = 4.19, p < .001, showed a preference of IDS over ADS (see Table 2 for the looking 252 times in each session). In the first session, 68.35 percent of the children showed a 253 preference for IDS, and in the second session, 63.29 percent of the children showed a numerical preference for IDS. In other words, we replicated the previous finding from the 255 main MB1 study. There was no difference in the strength of the preference effect, as a multilevel analysis with a random slope and random intercept for session on the lab level 257 revealed no significant impact of session on infants' preference, with an estimate of zero 258 and very small variance, β =-0.26, SE=0.31, p=.424. 259

Table 1 $Statistics \ of \ the \ included \ labs. \ n \ refers \ to \ the \ number \ of \ infants \ included$ in the final analysis.

Lab	Method	Language	Mean age (days)	N
babylab-potsdam	HPP	German	227	22
babyling-oslo	Eyetracking	Norwegian	249	10
brookes-babylab	central fixation	English	267	18
InfantCog-UBC	central fixation	English	147	7
infantll-madison	HPP	English	230	30
lancslab	Eyetracking	English	236	16
wsi-goettingen	central fixation	German	280	39
wsi-goettingen	HPP	German	242	16

 $\label{eq:looking times} \begin{tabular}{ll} Table 2 \\ Looking times in s for each session and condition \\ \end{tabular}$

Trial type	Session 1 Mean	Session 1 SD	Session 2 Mean	Session 2 SD
ADS	7.72	2.77	6.96	2.92
IDS	8.76	2.85	7.75	2.75

Table 3

Coefficient estimates from a linear mixed effects

model predicting IDS preference in Session 2.

	Estimate	SE	t	р
Intercept	0.874	0.456	1.920	0.102
Session One	0.035	0.085	0.414	0.679

260 Reliability

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We assessed test-retest reliability in two ways. First, we conducted a multilevel 261 analysis, with Lab as random intercept, predicting the IDS preference in Session 2 based on 262 the IDS preference in Session 1. The results revealed that we could not predict the 263 preference score in Session 2 based on Session 1 (see Table 2). Second, we calculated the 264 Pearson correlation coefficient. While a simple correlation coefficient might overestimate 265 the test-retest reliability in our sample because it does not control for the differences 266 between different labs and methods (HPP, CF, and eye-tracking), we felt it was important 267 to also conduct a Pearson correlation as it is commonly used to assess reliability. Again, the size of the correlation coefficient was not statistically different from zero and the 269 estimate was, in fact, approaching nil, r = .09, 95% CI [-.06, .25], t(156) = 1.19, p = .237.Furthermore, we calculated the percentage of preference reversal between test and retest counting the number of participants for whom the preference changed between test and 272 retest, and dividing it by the number of all participants. The results revealed that 41.77 percent of the infants had a preference reversal from test to retest session. Taken together, 274 our results lead us to conclude that there is no overall test-retest reliability for the three 275 infant preference measures used within the current study. 276

To test whether the results were different for a specific method, we calculated the

Table 4

Coefficient estimates from a linear mixed effects model predicting IDS preference in Session 2 for each method separately.

Method	estimate	SE	pvalue	cor	pvalue2
НРР	0.151	0.137	0.276	0.134	0.276
Eyetracking	0.034	0.162	0.835	0.021	0.919
central fixation	-0.195	0.125	0.125	0.080	0.530

Pearson correlation coefficients and the multilevel analyses for the three different methods, 278 HPP, central fixation and eye-tracking, separately (see Table 4). Splitting the data per 279 method also did not lead to different results. Neither the Pearson correlation coefficients 280 nor the coefficients of the multilevel analysis were significant, all p-values > .286. We also 281 tested for the possibility that the Time between sessions might have an impact on the 282 reliability. The subsequent multilevel analysis, with Lab as random intercept, predicting 283 the IDS preference in Session 2 based on the IDS preference in Session 1, the number of 284 days between Session 1 and Session 2 and the interaction of these two variables, did not 285 indicate that Time between sessions had an effect. Neither the main effect of Time between 286 sessions, $\beta = 0.00$, SE = 0.03, p = .916, nor the interaction term, $\beta = -0.01$, SE = 0.02, p = .479, 287 showed significant effects. As NAE-learning infants showed greater IDS preferences than their non-NAE counterparts in the original study, we also assessed if test-retest reliability interacted with children's native language. A multilevel analysis with Lab as random intercept, predicting the IDS preference in Session 2 based on the IDS preference in Session 291 1, NAE and the interaction of these two variables, revealed no interaction, β =0.29, 292 SE=0.18, p=.115 (see Figure 1).

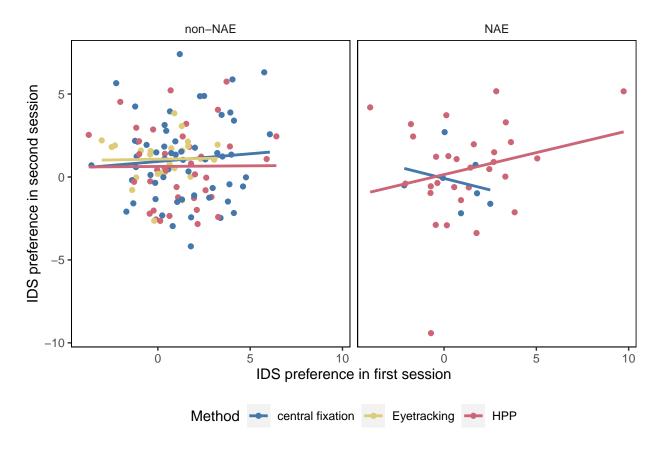


Figure 1. Infants' preference in Session 1 and Session 2 with individual data points and regression lines color-coded by method (central fixation, eye-tracking, or HPP). Results are plotted separately for North American English-learning infants (right panel) and infants learning other languages and dialects (right panel).

294 Results with different inclusion criteria

To this point, all analyses were performed on data with the inclusion criteria from MB1. For this, infants needed only 1 out of 8 valid trial pairs (i.e., any combination of an IDS and ADS trial) to be included in the analyses. Given that the use of more stringent inclusion criteria yielded larger effects sizes within the original MB1 study, we also assessed test-retest reliability by applying stricter inclusion criteria and thereby increasing test length to 2, 4, 6, and 8 included test trial pairs per condition. Applying a stricter criterion and thereby increasing test length - increased reliability numerically from r = 0.07 to r = 0.34 (Figure 2). However, in part likely due to the decrease in sample size, only one of

these correlations (when requiring a minimum of 6 trial pairs) was statistically significant:

2 valid trial pairs, t(152) = 0.90, p = .367; 4 valid trial pairs, t(143) = 1.03, p = .306; 6

valid trial pairs, t(98) = 2.23, p = .028; 8 valid trial pairs - all trials on both testing days -t(22) = 1.68, p = .108. The analyses provide tentative evidence that stricter inclusion

criteria might lead to higher test-retest reliability but at the same time comes with

tremendous decreases in sample size.

General Discussion

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The current study set out to explore the test-retest reliability of the infant speech 310 preference of IDS over ADS. Infants of the original MB1 project were retested on a 311 reversed order of stimuli in order to assess if their listening pattern would be similar to 312 that of their initial assessment. While we replicated the original effect of infants' speech 313 preference for IDS over ADS in the current MB1 follow-up study for both test and retest 314 session on the group-level using the same MB1 protocol, we found that infants' speech 315 preference measures had no test-retest reliability. In other words, we were unable to detect 316 any stable individual differences of infants' speech preference. This finding is in line with 317 other research indicating a rather low test-reliability for different developmental paradigms 318 (Cristia et al., 2016). Given that most experimental procedures conducted in 319 developmental research are interested in the comparison of groups, individual differences 320 between participants within a specific condition are usually minimized by the experimental 321 procedure while differences between conditions are maximized. Therefore, the infant 322 preference measure may be a good approach to capture universal phenomena but does not 323 seem to be appropriate for examining factors that may lead to individual differences in 324 development. 325

Consistent with general psychometric theory (e.g., DeBolt, Rhemtulla, & Oakes, 2020) a larger number of included test trials was associated with higher reliability.

However, in our dataset, this association was based on exploratory analyses and was only

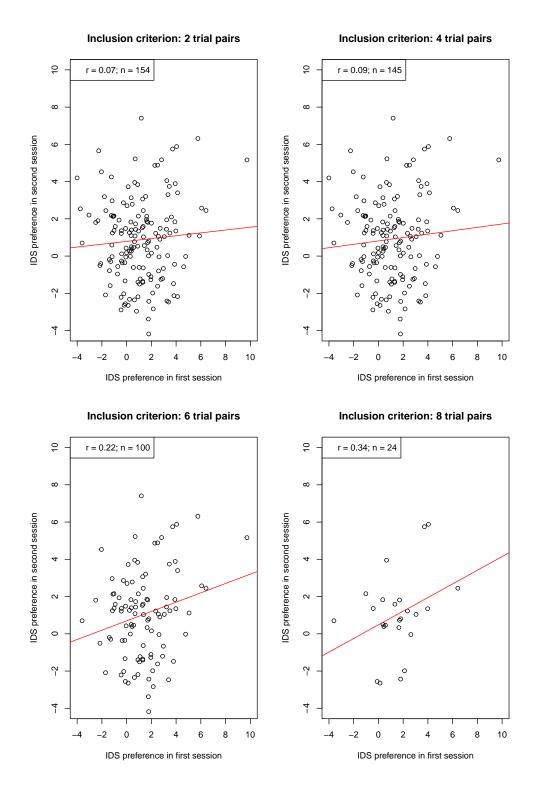


Figure 2. IDS preferences of both sessions plotted against each other for each inclusion criterion. n indicates the number of included infants, r is the Pearson correlation coefficient as the indicator for reliability.

found descriptively, hence, a replication is warranted. A similar effect on the group-level was found in the MB1 project, where a stricter inclusion criterion led to bigger effect sizes (ManyBabies Consortium, 2020).

In our study, as in the MB1 original study, higher reliability through strict exclusions 332 came at a high cost. In particular, with the strictest criterion, only a small portion of the 333 original sample size, that is 24 out of 158 infants, could be included in the final sample for 334 this particular analysis. In other words, applying a stricter criterion leads to a higher drop 335 out rate and reduces the actual sample size enormously. In the case of studies in the field 336 of developmental science, where there are many practical restrictions in collecting large 337 samples of infants (e.g., birth rate in the area, restricted lab capacities, budget 338 restrictions), a strict drop out criterion might not be easy - if even possible at all - to 339 implement. Note that studies in developmental science already have above average drop 340 out rates (Miller, 2017). In addition, drop out may not be random, and so having high drop out rates can further limit the generalisability of a study. Particularly in the context 342 of turning individual differences measures into diagnostic tools, high drop-out rates have an 343 additional limitation of not being broadly usable. 344

An alternative approach to increase the number of valid trials might be to also 345 increase the number of collected trials. In this case, a participant can have a high 346 number/proportion of invalid trials and still be included into the final sample as the 347 absolute number of trials is high and thereby decreasing trial-to-trial variability (DeBolt et 348 al., 2020; see Silverstein, Feng, Westermann, Parise, & Twomey, 2021 for an example). 349 While this approach might sound promising, it must be seen if this is realistic, because the attention span of a typical participant of a developmental study is rather short. Therefore, 351 prolonging the experimental procedure to maximize the absolute number of trials might 352 also be practically challenging. Further potential attempts in obtaining higher numbers of 353 valid trials may include changes in the procedure (e.g., Egger, Rowland, & Bergmann, 354 2020) or implementing multi-day test sessions (Fernald & Marchman, 2012). As our results 355

are only based on the particular phenomenon of IDS preference (albeit, with three widely 356 used methods: HPP, central fixation; eye-tracking) it is essential to further assess the 357 underlying reliability of these measures within other areas of speech perception. While 358 most infants prefer IDS over ADS (Dunst et al., 2012), predicting a pattern of preference, 359 for instance, within speech segmentation tasks, i.e. familiar versus novel words, seem not 360 that straightforward (Bergmann & Cristia, 2016). Especially in the context of relating a 361 direction of preference to later language development, there seem to be controversial 362 findings. That is, both familiarity and novelty responses have been suggested to be 363 predictive of infants' later linguistic abilities (DePaolis, Vihman, & Keren-Portnoy, 2014; 364 R. S. Newman, Rowe, & Ratner, 2016; R. Newman, Ratner, Jusczyk, Jusczyk, & Dow, 365 2006). In light of findings from the current study, researchers conducting longitudinal studies with experimental data from young infants predicting future outcomes should be cautious as there may be inter-individual variability affecting their preferences.

369 Limitations

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While we had an above average sample size for a study in developmental research, we 370 were unable to reach the number of participants collected within the original MB1 study. 371 In addition to a delayed call, the extra effort of having to schedule a second lab visit for 372 each participant and the fact that there were already other collaborative studies taking 373 place simultaneously (MB1B, Byers-Heinlein, Tsui, Bergmann, et al., 2021; MB1G, 374 Byers-Heinlein, Tsui, Van Renswoude, et al., 2021), might have contributed to the rather 375 low turnout. A higher sample size and a larger number of participating labs from different 376 countries might have enabled us to test for possible differences of the test-retest reliability 377 of the different methods (HPP, central fixation, eye-tracking) and NAE versus non-NAE 378 language backgrounds. Further, a larger sample size might have enabled us to conduct 379 meaningful tests of moderators such as age of the child on the test-retest reliability.

A further limitation concerns the stimuli. While the order of the clips within trials

presented to the participating children in the second session was different than in the first 382 session, the exact same stimulus material as in MB1 was used in both sessions. In 383 particular, all children heard the exact same voices in Session 1 and in Session 2. From a 384 practical point of view, it was the easiest solution. However, familiarity effects might have 385 influenced infants' looking behavior. Assuming that only infants with longer looking times 386 in Session 1 might have had the chance to recognize the voices in Session 2 from their 387 session a week ago as familiar clips would only be towards the end of trials, infants with 388 shorter looking times might not have had the opportunity to listen to the voices from their 389 first session. Therefore, for some children, familiarity with the stimulus material might 390 have led to artificially lowering test-retest reliability.

392 Conclusion

Following the MB1 protocol, the current study could not detect test-retest reliability
of infants' preference measures for IDS over ADS. Subsequent analyses showed that a
stricter criterion for the inclusion of data points may enhance the test-retest reliability at
the cost of high drop out rates. Developmental studies which rely on stable individual
differences of their participants need to consider the underlying reliability of their measures,
and we recommend a broader assessment of test-retest reliability in infant research.

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Data and materials availability statement

The data and materials that support the findings of the current study are openly available on OSF at https://osf.io/ZEQKA/.

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