Lemma 1 (Constant optimizations and duplication effects for Lemma 4.1 of [hmp01]). Suppose (f,g) are each functions from [n] to $\{0,1\}^r$. Let $F_x = \{i \in [n] : f(i) = x\}$ and $G_x = \{i \in [n] : g(i) = x\}$. Let $\mu : [n] \to [n]$ be a "deduplication" map, so that for all $x, y \in \{0,1\}^r$, μ maps all elements of $U_{xy} := \{i \in [n] : f(i) = x \land g(i) = y\}$ to a single arbitrary element of U_{xy} . Then in $O(n \log n)$ deterministic time and $O(n \log n)$ bits of space, one can construct $d : \{0,1\}^r \to \{0,1\}^r$ for which, with function $h(x) = g(x) \oplus d(f(x))$, and $H_x = \{i \in [n] : h(i) = x\}$, we have:

1.
$$\sum_{x \in \{0,1\}^r} {|H_x| \choose 2} - \sum_{i \in [n]} \frac{1}{|\mu^{-1}(i)|} {|\mu^{-1}(i)| \choose 2} \le \frac{1}{2^r} {n \choose 2}$$

2.
$$\sum_{x \in \{0,1\}^r} {|H_x| \choose 2} - \sum_{i \in [n]} {|\mu^{-1}(i)| \choose 2} \le n \left| \frac{1}{2^r} \left(\max\left(n, \sum_{x:|F_x| \ge 2} |F_x|^2\right) - 1 \right) \right|.$$

Proof. This is derived from the proofs in Section 4 of [hmp01]. To construct d, select a permutation $\pi: \{0,1\}^r \to \{0,1\}^r$ for which $|F_{\pi(1)}| \geq |F_{\pi(2)}| \geq \ldots \geq |F_{\pi(2^r)}|$. (The last sets in the sequence will all be empty if $n < 2^r$.) Then in order, for each $i \in [2^r]$, choose $d(\pi(i))$ to have value $a \in \{0,1\}^r$ so that the multiset $S_{a,i} := (a \oplus g(j) : h \in F_{\pi(i)})$ has no more than the average number of collisions with preceding multisets $\{S_{d(\pi(j)),j}\}_{j < i}$. The number of collisions c(A,B) between two multisets $A = (a_1,\ldots,a_{|A|})$ and $B = (b_1,\ldots,b_{|B|})$ is defined as $|\{x \in [|A|], y \in [|B|] : a_x = b_y\}$. Specifically, we want:

$$\sum_{j < i} c\left(S_{a,i}, S_{d(\pi(j)),j}\right) \le \left[\frac{1}{2^r} \sum_{b \in \{0,1\}} \sum_{j < i} c\left(S_{b,i}, S_{d(\pi(j)),j}\right)\right]$$

$$= \left[\frac{1}{2^r} \sum_{j < i} \sum_{b \in \{0,1\}} c\left(S_{b,i}, S_{d(\pi(j)),j}\right)\right]$$

$$= \left[\frac{1}{2^r} \sum_{j < i} \left|F_{\pi(i)}\right| \left|F_{\pi(j)}\right|\right]$$

where the last step follows because for each $x \in F_{\pi(i)}$, $y \in F_{\pi(j)}$, there is exactly one value of $b \in \{0,1\}^r$ for which $b \oplus g(x) = d(\pi(j)) \oplus g(y)$. This can be done (even with multi-sets!) using a dynamic search table structure as described in Section 4.3 of [hmp01].

The quantity $\sum_{j < i} c\left(S_{a,i}, S_{d(\pi(j)),j}\right)$ counts the total number of colliding pairs $a, b \in [n]$ where $f\left(a\right) \neq f\left(b\right)$ and $h\left(a\right) = h\left(b\right)$. Since $g\left(i\right) = h\left(i\right) \oplus d\left(f\left(i\right)\right)$, the number of colliding pairs where $a, b \in [n]$ satisfy $f\left(a\right) = f\left(b\right)$ and $h\left(a\right) = h\left(b\right)$ is equal to $\sum_{i \in [n]} {|\mu^{-1}(i)| \choose 2}$ (the number of collisions that (f, g) have.)

Consequently,

$$\sum_{x \in \{0,1\}^r} {|H_x| \choose 2} - \sum_{i \in [n]} \frac{1}{|\mu^{-1}(i)|} {|\mu^{-1}(i)| \choose 2} = \sum_{i \in \{0,1\}^r} \sum_{j < i} c\left(S_{a,i}, S_{d(f(j)),j}\right)$$

$$\leq \sum_{i \in \{0,1\}^r} \left| \frac{1}{2^r} \sum_{j < i} |F_{\pi(i)}| \left|F_{\pi(j)}\right| \right|$$

There are two ways to bound this. First,

$$\begin{split} \sum_{i \in \{0,1\}^r} \left[\frac{1}{2^r} \sum_{j < i} \left| F_{\pi(i)} \right| \left| F_{\pi(j)} \right| \right] &\leq \frac{1}{2^r} \sum_{\{i,j\} \in {\binom{\{0,1\}^r}{2}}} \left| F_{\pi(i)} \right| \left| F_{\pi(j)} \right| \\ &\leq \frac{1}{2^r} \cdot \frac{1}{2} \left(\sum_{i,j \in \{0,1\}^r} \left| F_{\pi(i)} \right| \left| F_{\pi(j)} \right| - \sum_{i \in [n]} \left| F_{\pi(i)} \right|^2 \right) \\ &\leq \frac{1}{2^r} \cdot \frac{n^2 - n}{2} = \frac{1}{2^r} \binom{n}{2} \end{split}$$

This bound does *not* use the permutation sort order; the following one does (and needs it, when $(F_{\pi(i)})_{i \in [n]}$ looks like $\sqrt{n}, \sqrt{n}, 1, 1, 1, \dots, 1$). Specifically:

$$\begin{split} \sum_{i \in \{0,1\}^r} \left\lfloor \frac{1}{2^r} \sum_{j < i} \left| F_{\pi(i)} \right| \left| F_{\pi(j)} \right| \right] &\leq n \max_{i \in \{0,1\}^r} \left\lfloor \frac{1}{2^r} \sum_{j < i} \left| F_{\pi(i)} \right| \left| F_{\pi(j)} \right| \right] \\ &\leq n \max_{i \in \{0,1\}^r} \left\{ \left\lfloor \frac{1}{2^r} \sum_{j < i} \left| F_{\pi(j)} \right|^2 \right] & \text{if } \left| F_{\pi(i)} \right| \geq 2 \\ \left\lfloor \frac{1}{2^r} \sum_{j < i} \left| F_{\pi(j)} \right| \right] & \text{if } \left| F_{\pi(i)} \right| \leq 1 \\ &\leq n \max_{i \in \{0,1\}^r} \left\{ \left\lfloor \frac{1}{2^r} \left(\sum_{x : |F_x| \geq 2} |F_x|^2 - \min_{|F_x| \geq 2} |F_x|^2 \right) \right] & \text{if } \left| F_{\pi(i)} \right| \geq 2 \\ \left\lfloor \frac{1}{2^r} \left(n - 1 \right) \right\rfloor & \text{if } \left| F_{\pi(i)} \right| \leq 1 \\ &\leq n \left\lfloor \frac{1}{2^r} \left(\max \left(n, \sum_{x : |F_x| \geq 2} |F_x|^2 \right) - 1 \right) \right\rfloor \end{split}$$

Lemma 2 (Deterministic double displacement.). Applying Lemma 1 twice, with $r = \lceil \log_2(\alpha n) \rceil$, gives a perfect hash function mapping n unique pairs $(f_i, g_i)_{i=1}^n$ to values $\lambda_i \in \{0,1\}^r$, when $\alpha \geq \sqrt{2}$.

Proof. First, apply Lemma 1 to $(f_i, g_i)_{i=1}^n$, producing $(h_i)_{i=1}^n$ with each $h_i \in \{0,1\}^r$ satisfying $h_i = g_i \oplus d_{1,i}(f_i)$ for some displacement function d_1 from $\{0,1\}^r \to \{0,1\}^r$. Then with $H_x := \{i \in [n] : h_i = x\}$ as defined in Lemma 1, we have $\sum_{x \in \{0,1\}^r} \binom{|H_x|}{2} \le \frac{1}{2^r} \binom{n}{2}$. Next, apply to Lemma 1 to $(h_i, f_i)_{i=1}^n$,

producing $(\lambda_i)_{i=1}^n$ with each $\lambda_i \in \{0,1\}^r$ satisfying $\lambda_i = f_i \oplus d_{2,i}$ (h_i) for some displacement function $d_2 : \{0,1\}^r \to \{0,1\}^r$. Define $\Lambda_x := \{i \in [n] : \lambda_i = x\}$.

$$\frac{1}{2^r} \binom{n}{2} \ge \sum_{x \in \{0,1\}^r} \binom{|H_x|}{2} \ge \frac{1}{4} \sum_{x \in \{0,1\}^r : |H_x| \ge 2} |H_x|^2 \tag{1}$$

so by the second bound in Lemma 1

$$\sum_{x \in \{0,1\}^r} {\binom{|\Lambda_x|}{2}} \le n \left\lfloor \frac{1}{2^r} \max \left(n - 1, 4\frac{1}{2^r} \binom{n}{2}\right) \right\rfloor$$

$$= n \left\lfloor \frac{2}{2^{2r}} n (n - 1) \right\rfloor \qquad \text{if } 2^r \ge n$$

$$= 0 \qquad \qquad \text{if } 2^r \ge n\sqrt{2}$$

Note 3. The first and second bounds of Lemma 1 do not fit together when $i\mapsto (f\left(i\right),g\left(i\right))$ is not one-to-one. It is possible that, when $\sum_{x\in\{0,1\}^r} {|F_x|\choose 2} - \sum_{i\in[n]} \frac{1}{|\mu^{-1}(i)|} {|\mu^{-1}(i)|\choose 2} \le n$, the bound $\sum_{x\in\{0,1\}^r} {|H_x|\choose 2} - \sum_{i\in[n]} \frac{1}{|\mu^{-1}(i)|} {|\mu^{-1}(i)|\choose 2} = 0$ holds, but proving or disproving this may require going into the details of the search table procedure. (If there is a hard instance, it might have each nonempty multiset F_x contain two distinct g values (possibly with duplicates) structured to trick the search procedure into using a small branch of the table.)

Say that for some x, F_x has $k = |\mu(F_x)|$ equivalence classes by μ , of sizes a_1, \ldots, a_k , with all $a_j \ge 1$. Because $\sum_{j \in [k]} (a_j - 1)^2 \le \left(\sum_{j \in [k]} (a_j - 1)\right)^2$:

$$\begin{aligned} \begin{pmatrix} |F_x| \\ 2 \end{pmatrix} &- \sum_{i \in F_x} \frac{1}{|\mu^{-1}(i)|} \binom{|\mu^{-1}(i)|}{2} \\ &= \binom{|F_x|}{2} - \sum_{j \in [k]} \binom{a_j}{2} \\ &= \binom{|F_x|}{2} - \frac{1}{2} \sum_{j \in [k]} \left[(a_j - 1)^2 + (a_j - 1) \right] \\ &= \binom{|F_x|}{2} - \frac{1}{2} (|F_x| - k) - \frac{1}{2} \sum_{j \in [k]} (a_j - 1)^2 \\ &\geq \binom{|F_x|}{2} - \frac{1}{2} (|F_x| - k) - \frac{1}{2} (|F_x| - k)^2 \\ &= \binom{|F_x|}{2} - \binom{|F_x| - k}{2} \\ &= \frac{1}{2} \left(|F_x|^2 - |F_x| - (|F_x| - k)^2 + (|F_x| - k) \right) \\ &= \frac{k}{2} \left(2|F_x| - k - 1 \right) \end{aligned}$$

Therefore, the nontrivial collision count κ satisfies:

$$\kappa := \sum_{x \in \{0,1\}^r} \binom{|F_x|}{2} - \sum_{i \in [n]} \frac{1}{|\mu^{-1}(i)|} \binom{|\mu^{-1}(i)|}{2}$$

$$= \sum_{x \in \{0,1\}^r} \binom{|F_x|}{2} - \sum_{i \in F_x} \frac{1}{|\mu^{-1}(i)|} \binom{|\mu^{-1}(i)|}{2}$$

$$\geq \sum_{x \in \{0,1\}^r} \frac{|\mu(F_x)|}{2} (2|F_x| - |\mu(F_x)| - 1)$$

This can be used to bound the cost of delayed deduplication for the second displacement round. For example, for each F_x , one can by a variant of insertion sort construct a sorted list of unique elements in $O(|F_x| |\mu(F_x)|)$ time, which summed over all $x \in \{0,1\}^r$ is O(n). Or the search table design can be modified so that, when it is time to update table frequencies after choosing $d(\pi(i))$, the leaves are updated to indicate just whether an element has been used, not how many times it has been used, and the remaining values derived from the leaves. Then the key quantity to bound would be:

$$\max_{i \in \{0,1\}^r} |F_{\pi(i)}| \sum_{j < i} |\mu(F_{\pi(j)})| \le \max \left(n - 1, \sum_{x \in \{0,1\}^r : |F_x| \ge 2} |F_x| |\mu(F_x)| \right)$$

$$\le \max(n - 1, 4\kappa)$$

The last inequality follows because if $|F_x| > |\mu(F_x)|$, then $|\mu(F_x)| (2|F_x| - |\mu(F_x)| - 1) \ge |F_x| |\mu(F_x)|$, while if $|F_x| = |\mu(F_x)|$ and $|F_x| \ge 2$ then $|\mu(F_x)| (2|F_x| - |\mu(F_x)| - 1) \ge \frac{1}{2} |\mu(F_x)| |F_x|$.

Corollary 4. For deterministic double displacement, f and g both map to $\{0,1\}^r$. The displacement procedure can also be used, just once, to convert a hash function with few collisions to one with none. Then we may have $f:[n] \to \{0,1\}^t$ and $g:[n] \to \{0,1\}^r$. We require $x \mapsto (f(x),g(x))$ to be a bijection. Define F_x and G_x as in 1; the total number of collisions for f is $\sum_x {|F_x| \choose 2}$. Say this is f c. Then one can construct a displacement function f construct f

$$\max\left(n, \sum_{x:|F_x|\geq 2} |F_x|^2\right) \leq 2^r \qquad where \qquad \sum_{x:|F_x|\geq 2} |F_x|^2 \leq 4c.$$

Proof. Modify the proof of Lemma 1; the bound on $\sum_{x:|F_x|\geq 2} |F_x|^2$ follows from Eq. 1.

Definition 5. Odd-multiply-shift (OMS) hashing. Used by [r95], proved 2-approximately universal by [dhkp97], with the hash function type known at

latest since TAOCP Vol 3 in 1973. The hash family $\mathcal{H}_{b,s}$ maps $\{0,\ldots,2^u-1\} \equiv \{0,1\}^u$ to $\{0,\ldots,2^s-1\} \equiv \{0,1\}^s$ and is parameterized by odd integers $a \in \{0,\ldots,2^u-1\}$; individual functions are given by

$$h_{\alpha}(x) = (ax \mod 2^u) \operatorname{div} 2^{u-s}$$
.

Lemma 6. For any distinct $x, y \in \{0, 1\}^u$, with α drawn uniformly at random from the odd integers in $\{0, \dots, 2^u - 1\}$, we have the following two bounds:

$$\Pr_{a} [h_{a}(x) = h_{a}(y)] \leq \begin{cases}
0 & \text{if } (x - y \mod 2^{u - s}) = 0 \\
1 & \text{otherwise}
\end{cases}$$

$$\Pr_{a} [h_{a}(x) = h_{a}(y)] \leq \Pr_{a} [a(x - y) \mod 2^{u} \operatorname{div} 2^{u - s} = 0]$$

$$+ \Pr_{a} [a(y - x) \mod 2^{u} \operatorname{div} 2^{u - s} = 0] = \frac{2}{2^{s}}$$

Proof. We have

$$\Pr_{a}\left[h_{a}\left(x\right)=h_{a}\left(y\right)\right]=\Pr_{a}\left[\left(ax\mod 2^{u}\right)\operatorname{div}2^{u-s}=\left(ay\mod 2^{u}\right)\operatorname{div}2^{u-s}\right]$$

Let i be the largest integer for which $x \mod 2^i = y \mod 2^i$. Decompose $x = x_h 2^i + c$ and $y = y_h 2^i + c$ where $c = x \mod 2^i = y \mod 2^i$. If $i \ge u - s$, then

$$(ax \mod 2^u) \operatorname{div} 2^{u-s} = (ax_h 2^i + ac) \mod 2^u \operatorname{div} 2^{u-s}$$

$$= (((ax_h \mod 2^{u-i}) 2^i + (ac \mod 2^u)) \mod 2^u) \operatorname{div} 2^{u-s}$$

$$= ((ax_h \mod 2^{u-i}) 2^{i-(u-s)} + (ac \mod 2^u \operatorname{div} 2^{u-s})) \mod 2^{u-s}$$

$$= ((ax_h \mod 2^{u-i}) 2^{i-(u-s)} + d) \mod 2^{u-s}$$

where $d = (ac \mod 2^u \operatorname{div} 2^{u-s})$. Since a is odd, for $z \in \{0, \dots 2^{u-i}\}$, the map $z \to (az \mod 2^{u-i})$ is a bijection; and since $x_h \neq y_h$, it follows $h_a(x) \neq h_a(y)$. On the other hand, if i < u - s, then

$$\begin{split} &\Pr_{a} \left[(ax \mod 2^u) \ \text{div} \ 2^{u-s} = (ay \mod 2^u) \ \text{div} \ 2^{u-s} \right] \\ &\leq \Pr_{a} \left[((ax \mod 2^u) - (ay \mod 2^u)) \ \text{div} \ 2^{u-s} = 0 \right] \\ &\quad + \Pr_{a} \left[((ay \mod 2^u) - (ax \mod 2^u)) \ \text{div} \ 2^{u-s} = 0 \right] \\ &= \Pr_{a} \left[(a \ (x-y) \mod 2^u) \ \text{div} \ 2^{u-s} = 0 \right] \\ &\quad + \Pr_{a} \left[(a \ (y-x) \mod 2^u) \ \text{div} \ 2^{u-s} = 0 \right] \end{split}$$

The first inequality follows since $w \operatorname{div} 2^{u-s} = z \operatorname{div} 2^{u-s}$ implies both w and z lie in the same set $\{2^{u-s}k, \ldots, 2^{u-s}(k+1)-1\}$; if $w \leq z$, then $(z-w) \mod 2^u < 2^{u-s}$ so $(z-w) \mod 2^u \operatorname{div} 2^{u-s} = 0$, while if $w \geq z$, then $(w-z) \mod 2^u \operatorname{div} 2^{u-s} = 0$. Note that $(x-y) \mod 2^u = ((x_h-y_h)2^i) \mod 2^u$ and

and $x_h - y_h \mod 2 = 1$. Multiplying by $x_h - y_h$ under $\mod 2^{u-i}$ permutes odd values in $\{0, \dots, 2^{u-i} - 1\}$, so $\left(a\left(x_h - y_h\right)2^i\right) \mod 2^u$ is uniformly distributed over odd multiples of 2^i in 2^u . Since i < u - s, and $\{0, \dots, 2^{u-s}\}$ contains $2^{u-s-i-1}$ odd multiples of 2^i (out of 2^{u-i-1} possible values).

$$\Pr[(a(x-y) \mod 2^u) \in \{0, \dots, 2^{u-s}\}] = \frac{1}{2^s}.$$

The $\Pr_a\left[h_a\left(x\right)=h_a\left(y\right)\right] \leq \frac{1}{2^{s-1}}$ upper bound is tight. For example, with u=4,s=3, $\Pr\left[h_a\left(1\right)=h_a\left(y\right)\right] = \frac{1}{4}$ holds y=3 and y=11.

Note 7. For comparison, the upper bound from [r95] on $Pr_a[h_a(x) = h_a(y)]$ is

$$\Pr_{a} \left[a \left(x - y \right) \mod 2^{u} \in \left\{ 0, \dots 2^{u - s} - 1 \right\} \cup \left\{ 2^{u} - 2^{u - s} + 1, \dots 2^{u - 1} \right\} \right]$$

for which a $\leq \frac{1}{2^{s-1}}$ upper bound also holds. The bound of Lemma 6 has the advantage of being slightly easier to compute in the incremental a-finding hash construction approach of [r95]. Also, either bound can be used to find that the expected number of collisions when uniformly randomly hashing an n-element set n from $\{0,1\}^s$ to $\{0,1\}^b$ is $\leq \frac{1}{2^{s-1}}\binom{n}{2}$. Consequently, for a perfect hash function from $\mathcal{H}_{b,s}$ to exist for all sets of size n, one needs $\frac{1}{2^{s-1}}\binom{n}{2} < 1$ which holds iff $s \geq \lceil \log (n(n-1)+1) \rceil$.

However, this is not a necessary condition, and one can do slightly better in some cases. When n=1, $2^{\lceil \log(n(n-1)+1) \rceil}=1$, but when n=2, $2^{\lceil \log(n(n-1)+1) \rceil}=4$. However, given any $x,y \in \{0,1\}^u$, let $j \in \{0,\dots,u-1\}$ be the highest bit for which $x_j \neq y_j$. If j=u-1, then $h_1(x)=x_j \neq y_j=h_1(y)$ and choosing a=1 separates the two. Otherwise, let $a=2^{u-1-j}$ (which is admittedly even); then $ax=2^{u-1-j}x+x$; the multiplication does not overflow 2^u so $ax \operatorname{div} 2^{u-1}=x_j$, and similarly $ay \operatorname{div} 2^{u-1}=y_j$, so $h_a(x) \neq h_a(y)$. Thus a perfect hash function for n=2 with s=1 always exists. It may be possible that similar minimal perfect hashing can be done with odd multiply-shift hashes for $n=2,3,4,\dots$ up to some small set size threshold, although space lower bounds imply the threshold is at most $O(\log u)$.

Definition 8. A hash function h from A to B is c-colliding on a set $S \subseteq A$ of size n if $\sum_{\{x,y\}\in\binom{S}{2}} 1_{h(x)=h(y)} \leq c$.

Lemma 9. Say that $h: A \to B$ is c-colliding on the set S. Let b_1, \ldots, b_k be the sizes of the equivalence classes of S under h, so that $\sum_i b_i = n = |S|$ and $\sum_i {b_i \choose 2} \le c$. Let

$$s(b) = \begin{cases} b & \text{if } b \le 2\\ 2^{\lceil \log(b_i(b_i-1)+1) \rceil} & \text{if } b \ge 3 \end{cases}$$

Then

$$\sum_{i} s(b_i) < n + 4c$$

$$\max\left(n, \sum_{b_i > 2} s(b_i)\right) \le \max(n, 4c)$$

(These give bounds on the total output sizes for the FKS two-level hashing scheme ([fks84]) if placing leaf hash tables disjointly using offset values, and if packing size-1 tables into the gaps of the other tables. Size 2-tables are also contiguous, so one may be able to pack those into the $\max\left(\frac{1}{2}s\left(b\right)-b,0\right)$ gaps of length 2 left when hashing b-sized equivalence classes; to what extent this works depends on the exact equivalence class size distribution.)

Proof. Let $r(b) = \frac{s(b)}{b(b-1)/2}$ (i.e, the ratio of table size to collision count for an equivalence class). We have $r(1) = \infty$, r(2) = 2, $r(3) = r(4) = \frac{8}{3}$, $r(5) = \frac{16}{5}$, and $\sup_{b \ge 6} r(b) = 4$ (although *exactly* 4 is never reached; as $b \to \infty$, r(b) oscillates between 2 and 4 depending on how much the $\lceil \cdot \rceil$ operation in s(b) adds.)

$$\sum_{b_{i} \ge 2} s(b_{i}) < 4 \sum_{b_{i} \ge 2} \frac{b_{i}(b_{i} - 1)}{2} \le 4c$$

Since s(1) = 1, $\sum_{b_i=1} s(b_i) = |\{i : b_i = 1\}| \le n$, so

$$\sum_{i} s(b_i) = \sum_{b_i=1} s(b_i) + \sum_{b_i=2} s(b_i) \le n + 4c$$

(This is roughly tight; if all collisions are concentrated in a single equivalence class of size $d = O(\sqrt{c})$, then the large class will contribute $s(d) = \Theta(c)$ while the small ones add n - d for the singleton equivalence classes.

Remark 10. If the FKS hashing scheme ([fks84]) is implemented using OMS hashing (Definition 5), with hash functions picked to have a sub-average number of collisions, using size-2 leaf tables for size-2 leaves, tightly packing size-1 subtables in the free space of leaf tables, using α and β bits per primary and secondary table entries, and with main hash output bit count s, the total space usage is

$$\alpha 2^s + \beta \max\left(n, \frac{4}{2^{s-1}} \binom{n}{2}\right)$$

which is minimized when

$$s = \min \left(\lceil \log (n-1) \rceil, \lceil \frac{1}{2} \log \left(\frac{4\beta \binom{n}{2}}{\alpha} \right) \rceil \right)$$

(Choosing $s = \left\lceil \frac{1}{2} \log \left(\frac{x}{2} \right) \right\rceil$ minimizes $2^s + \frac{x}{2^s}$ with value $\leq \frac{3}{\sqrt{2}} \sqrt{x}$, so the total space usage for the table is $\leq \max \left(\frac{6\sqrt{\beta\alpha}}{\sqrt{2}} n, \beta n + 2\alpha n \right)$.)

Lemma 11 (Parameter selection for an odd-multiply-shift + displacement table perfect hash construction.). Let n and s be integers, and α, β nonnegative reals. Under the constraints $r+t \geq s$ and

$$2^r \ge \max(n, 4c)$$
 where $c = \frac{n(n-1)}{2^t}$.

we have that the minimum value of $\alpha 2^r + \beta 2^t$ occurs when:

$$r = \min\left(\max\left(\left\lceil\frac{1}{2}\left(\log\left(\frac{\beta}{2\alpha}\right) + w\right)\right\rceil, \lceil\log n\rceil\right), w\right)$$

$$where \qquad w = \max\left(s, \lceil\log\left(4\left(n\left(n-1\right)\right)\right)\right\rceil\right)$$

Proof. First, note that $2^r \geq 4\frac{n(n-1)}{2^t}$ is true iff $r+t \geq \lceil \log \left(4\left(n\left(n-1\right)\right)\right) \rceil$. Since reducing either of r or t non-increases $\alpha 2^r + \beta 2^t$, the optimal value of

$$r + t = \max(s, \lceil \log(4(n(n-1))) \rceil) =: w.$$

The quantity $2^r + \frac{\frac{\beta}{\alpha}2^w}{2^t}$ is minimized when $r = \left\lceil \frac{1}{2}\log\left(\frac{\left(\frac{\beta}{\alpha}2^w\right)}{2}\right) \right\rceil = \left\lceil \frac{1}{2}\log\left(\frac{\beta}{2\alpha}2^w\right) \right\rceil$, but r is also subject to the constraints $r \in \left[\left\lceil \log n \right\rceil, w\right]$. The space usage then is

$$\begin{split} & \min\left(\max\left(\alpha 2^{\lceil\log n\rceil} + \beta 2^{w - \lceil\log n\rceil}, \frac{3}{\sqrt{2}}\sqrt{\beta\alpha 2^w}\right), \alpha 2^w + \beta\right) \\ & \leq \min\left(\max\left(2\alpha n + \beta\max\left(\frac{2^s}{n}, 8\left(n-1\right)\right), \frac{3}{\sqrt{2}}\sqrt{\beta\alpha\max\left(2^s, 8n\left(n-1\right)\right)}\right), \alpha 2^w + \beta\right) \\ & \leq \min\left(\max\left(2\alpha n + \beta\frac{2^s}{n}, 2\alpha n + 8\beta n, \frac{3}{\sqrt{2}}\sqrt{\beta\alpha}2^{s/2}, 6\sqrt{\beta\alpha}n\right), \alpha\max\left(8n^2, 2^s\right) + \beta\right) \end{split}$$

Lemma 12 (Evaluating conditional expectation for multiply-shift collision probability.). (This is similar to [r95, Lemma 4], but instead optimizing the simpler bounds of Lemma 6.) Let $x, y \in \{0, \dots, 2^u - 1\}$ be distinct with $(x - y) \mod 2^u = z2^i$ for z odd and i a nonnegative integer. Fix $\alpha \in \{0, 1\}^*$ of length α , and let $E(\alpha)$ be the set of integers $a \in \{0, \dots, 2^u - 1\}$ whose $|\alpha|$ least significant bits match α . Then for integer $s \in [0, \dots, u]$ with i < u - s:

$$\begin{split} &\Pr_{a \sim E(\alpha)} \left[a \left(x - y \right) \mod 2^u \operatorname{div} 2^{u - s} = 0 \right] \\ &= \begin{cases} 2^{-s} & i + |\alpha| \leq u - s \\ 1_{\left\{ (\alpha(x - y) + 2^{u - s} - 1) \mod \min\left(2^u, 2^{|\alpha| - i}\right) < 2^{u - s}\right\}} \frac{1}{2^{u - |\alpha| - i}} & i + |\alpha| > u - s \end{cases} \end{split}$$

Proof. As in the proof of [r95, Lemma 4], we observe that a can be written as $a'2^{|\alpha|} + \alpha$ where α is interpreted as an integer in $\{0, \ldots, 2^{|\alpha|} - 1\}$ and a' is an

integer in $\{0,\ldots,2^{u-|\alpha|}-1\}$. Then:

$$\begin{split} p &:= \Pr_{a \sim E(\alpha)} \left[a \left(x - y \right) \mod 2^u \operatorname{div} 2^{u-s} = 0 \right] \\ &= \Pr_{a' \sim \left\{ 0, \dots, 2^{u-|\alpha|} - 1 \right\}} \left[a' z 2^{|\alpha|+i} + \alpha z 2^i \mod 2^u \operatorname{div} 2^{u-s} = 0 \right] \end{split}$$

Since z is odd, $a'z \mod 2^{u-|\alpha|-i}$ is uniformly distributed over $\{0,\ldots,2^{u-|\alpha|-i}\}$. Let b=a'z. Consequently,

$$p = \Pr_{b \sim \left\{0, \dots, 2^{u - |\alpha| - i} - 1\right\}} \left[b2^{|\alpha| + i} \mod 2^u \in \left\{\alpha z 2^i, \dots \alpha z 2^i + 2^{u - s} - 1\right\} \mod 2^u \right]$$

We now have two cases: if $|\alpha|+i \leq u-s$, then the set $\left\{\alpha z 2^i, \ldots \alpha z 2^i + 2^{u-s} - 1\right\}$ contains exactly $\frac{2^{u-s}}{2^{|\alpha|+i}}$ distinct values of $b2^{|\alpha|+i}$, in which case $p=\frac{2^{u-s}}{2^{|\alpha|+i}}\cdot \frac{1}{2^{u-|\alpha|-i}}=\frac{1}{2^s}$. On the other hand, if $|\alpha|+i>u-s$, then $\left\{\alpha z 2^i, \ldots \alpha z 2^i + 2^{u-s} - 1\right\}$ contains at most one multiple of $b2^{|\alpha|+i}$. Iff this occurs, then for some $c\in\{0,\ldots,2^{u-s}-1\}$, we have need $\left(\alpha z 2^i + c\right) \mod \min\left(2^u,2^{|\alpha|-i}\right)=0$. This occurs iff

$$(\alpha z 2^i + 2^{u-s} - 1) \mod \min (2^u, 2^{|\alpha|-i}) < 2^{u-s}$$

In that case, $p = \frac{1}{2^{u-|\alpha|-i}}$.

Definition 13. Odd-Multiply-add-shift (OMA) hashing. A *variant* of this was proven by proven by [thorup15] to be 2-independent/strongly universal, but that is stronger than needed for universe reduction or the finding of low-collision hash functions: for both tasks approximate universality suffices. Define the hash family $\mathcal{G}_{b,s}$ mapping $\{0,\ldots,2^u-1\}\equiv\{0,1\}^u$ to $\{0,\ldots,2^s-1\}\equiv\{0,1\}^s$ to contain, over all odd integers $a\in\{1,3,\ldots,2^u-1\}$, and integers $b\in\{0,1,2,\ldots,2^{u-s}-1\}$, the functions

$$h_{a,b} = (ax + b) \mod 2^u \operatorname{div} 2^{u-s}.$$

We will show in the following lemma that this family is (1-approximately) universal, a factor two better than the Odd-multiply-shift family.

Lemma 14 (Universality of OMA hash functions.). Fix $u \ge s > 0$. For any distinct $x, y \in \{0, 1\}^u$, and $h_{a,b}$ drawn uniformly at random from $\mathcal{G}_{b,s}$ (see Definition 13),

$$\Pr\left[h_{a,b}\left(x\right) = h_{a,b}\left(y\right)\right] = \begin{cases} 0 & \text{if } (x-y) \mod 2^{u-s} = 0\\ \frac{1}{2^{s}} & \text{otherwise} \end{cases}$$

Proof. Define for brevity $p := \Pr[h_{a,b}(x) = h_{a,b}(y)]$. Let i be the integer for which $(x-y) \mod 2^u = z2^i$ for z odd. Decompose $x = x_h2^i + x_l$, $y = y_h2^i + y_l$, where $x_h, y_h \in \{0, \dots 2^{u-i} - 1\}$ and $x_l, y_l \in \{0, \dots 2^i - 1\}$. There are two cases:

• If $i \ge u - s$, then

$$(ax + b) - (ay + b) \mod 2^u = (a(x_h - y_h) \mod 2^{u-i}) 2^i$$

which is always nonzero because both a and $x_h - y_h = z$ are odd. Consequently, $h_{a,b}(x)$ and $h_{a,b}(y)$ differ by at least $2^{i-(u-s)}$, so p=0.

• If i < u - s, let $c = x_l = y_l$, and observe:

$$p = \Pr_{a,b} \left[\left(ax_h 2^i + ac + b \mod 2^u \right) \operatorname{div} 2^{u-s} = \left(ay_h 2^i + ac + b \mod 2^u \right) \operatorname{div} 2^{u-s} \right]$$

For any $u, v \in \{0, \dots, 2^u - 1\}$, consider

$$\Pr_b \left[(u+b) \mod 2^u \operatorname{div} 2^{u-s} = (v+b) \mod 2^u \operatorname{div} 2^{u-s} \right]$$

Let d be the signed difference $u-v \bmod 2^u$ so that $d \in \left\{-2^{u-1}+1, \ldots -1, 0, 1 \ldots, 2^{u-1}-1\right\}$, with the special case $u=\left(v+2^{u-1}\right) \bmod 2^u$, mapped arbitrarily to either $+2^{u-1}$ or -2^{u-1} (either case will lead to the same calculated value because $s \geq 1$). If $|d| > 2^{u-s}$, then $(u+b) \bmod 2^u \operatorname{div} 2^{u-s}$ and $(v+b) \bmod 2^u \operatorname{div} 2^{u-s}$ always differ by at least one – for them to fall in the same set $\{k2^{u-s}, \ldots, k2^{u-s} + 2^{u-s} - 1\}$ for some k requires $|d| < 2^{u-s}$. If $|d| \leq 2^{u-s}$, then exactly $2^{u-s} - |d|$ of the possible values for $b \in \{0, \ldots, 2^{u-s} - 1\}$ will make u+b and v+b fall in the same set $\{k2^{u-s}, \ldots, k2^{u-s} + 2^{u-s} - 1\}$ for some k. (If $d \geq 0$, then all d values of b which for which $(v+b) \bmod 2^{u-s} < 2^{u-s} - (u-v)$ work, and a symmetric condition holds if $d \leq 0$.) Thus

$$\Pr_b \left[(u+b) \mod 2^u \operatorname{div} 2^{u-s} = (v+b) \mod 2^u \operatorname{div} 2^{u-s} \right] = \max \left(0, \frac{2^{u-s} - |d|}{2^{u-s}} \right)$$

Since a is chosen uniformly at random from odd values in $\{0, \ldots, 2^u - 1\}$, and $x_h - y_h$ is odd, the product $a(x_h - y_h) \mod 2^{u-i}$ is distributed uniformly at random over odd values in $\{0, \ldots, 2^{u-i} - 1\}$, and the signed difference $ax_h2^i - ay_h2^i$ is distributed uniformly over the odd multiples of 2^i in $\{-2^{u-1} + 1, \ldots, 0, \ldots, 2^{u-1} - 1\}$. (Since i < u - s, these odd multiples never contain the value $\pm 2^{u-1}$.) By the law of total probability,

$$\begin{split} p &:= \frac{1}{2^{u-i-1}} \sum_{k \in \{+1,-1\}} \sum_{j=0}^{2^{u-i-1}} \max \left(0, \frac{2^{u-s} - \left| k \left(2j+1 \right) 2^i \right|}{2^{u-s}} \right) \\ &= \frac{1}{2^{u-i-1}} \cdot 2 \sum_{j=0}^{2^{u-s-i-1}} \frac{2^{u-s-i} - \left(2j+1 \right)}{2^{u-s-i}} \\ &= \frac{1}{2^{u-i-1}} \cdot 2 \cdot 2^{u-s-i-1} \cdot \frac{1}{2} = \frac{1}{2^s} \,. \end{split}$$

Lemma 15 (Conditional probabilities for bit-by-bit selection of OMAS hash functions.). Let $x, y \in \{0, \ldots, 2^u - 1\}$ be distinct with $(x - y) \mod 2^u = z2^i$ for z odd and i a nonnegative integer. Fix $\alpha \in \{0, 1\}^*$ of length α , and let $E(\alpha)$ be the set of integers $a \in \{0, \ldots, 2^u - 1\}$ whose $|\alpha|$ least significant bits match α . Let $s \in [0, \ldots, u]$. As argued in the proof of Lemma 6, if $i \geq u - s$, then any $h_{a,b} \in \mathcal{G}_{b,s}$ separates x and y. Assume i < u - s. Then if $|\alpha| + i \leq u - s$, we have:

$$\Pr_{h_{a,b} \in \mathcal{G}_{b,s}: a \in E(\alpha)} \left[h_{a,b} \left(x \right) = h_{a,b} \left(y \right) \right] = \frac{1}{2^{s}}$$

while if $|\alpha| + i \ge u$, we have:

$$\Pr_{h_{a,b} \in \mathcal{G}_{b,s}: a \in E(\alpha)} \left[h_{a,b} \left(x \right) = h_{a,b} \left(y \right) \right] = \max \left(0, \frac{2^{u-s} - \left(2^{u-1} - \left| 2^{u-1} - \left| 2^{u-1} - \left(\alpha \left(x - y \right) \mod 2^{u} \right) \right| \right)}{2^{u-s}} \right)$$

and if both $|\alpha| + i > u - s$ and $|\alpha| + i \le u$, we have:

$$\Pr_{h_{a,b} \in \mathcal{G}_{b,s}: a \in E(\alpha)} \left[h_{a,b} \left(x \right) = h_{a,b} \left(y \right) \right] = \frac{1}{2^{u-|\alpha|-i}} \cdot \max \left(0, \frac{2^{u-s} - \left| 2^{|\alpha|+i-1} - \left(\alpha \left(x - y \right) \mod 2^{|\alpha|+i} \right) \right|}{2^{u-s}} \right).$$

Furthermore, for fixed $a, \beta \in \{0,1\}^*$, and $F(\beta)$ the set of integers $b \in \{0,\ldots,2^u-1\}$ whose $|\beta|$ most significant bits match β , $\Pr_{b\in F(\beta)}[h_{a,b}(x)=h_{a,b}(y)]$ equals, where $w=ax+\beta 2^{u-s-|\beta|}$ and $v=ay+\beta 2^{u-s-|\beta|}$:

$$\begin{cases} \frac{\iota(w) - \iota(v)}{2^{u-s-|\beta|}} & v-w \mod 2^u < 2^{u-s} \wedge w \mod 2^{u-s} > v \mod 2^{u-s} \\ 1 + \frac{\iota(w) - \iota(v)}{2^{u-s-|\beta|}} & v-w \mod 2^u < 2^{u-s} \wedge w \mod 2^{u-s} \leq v \mod 2^{u-s} \\ \frac{\iota(v) - \iota(w)}{2^{u-s-|\beta|}} & w-v \mod 2^u < 2^{u-s} \wedge v \mod 2^{u-s} > w \mod 2^{u-s} \\ 1 + \frac{\iota(v) - \iota(w)}{2^{u-s-|\beta|}} & w-v \mod 2^u < 2^{u-s} \wedge v \mod 2^{u-s} \leq w \mod 2^{u-s} \\ 0 & otherwise \end{cases}$$

where $\iota(w) := \max(0, w \mod 2^{u-s} - (2^{u-s} - 2^{u-s-|\beta|})).$

Proof. Let $p := \Pr_{h_{a,b} \in \mathcal{G}_{b,s}: a \in E(\alpha)} [h_{a,b}(x) = h_{a,b}(y)]$. Partition random variable a into $a'2^{|\alpha|} + \alpha$ where α is interpreted as an integer in $\{0, \ldots, 2^{|\alpha|} - 1\}$; the random variable b is uniform over $\{0, \ldots, 2^{u-s} - 1\}$. Decompose $x = x_h 2^i + x_l$ and $y = y_h 2^i + y_l$, and $c = x_l = y_l$. Then

$$p = \Pr \left[\left(ax_h 2^i + ac + b \right) \mod 2^u \operatorname{div} 2^{u-s} = \left(ay_h 2^i + ac + b \right) \mod 2^u \operatorname{div} 2^{u-s} \right]$$

$$= \Pr \left[\left(a'x_h 2^{|\alpha|+i} + \alpha x_h 2^i + ac + b \right) \mod 2^u \operatorname{div} 2^{u-s} \right]$$

$$= \left(a'y_h 2^{|\alpha|+i} + \alpha y_h 2^i + ac + b \right) \mod 2^u \operatorname{div} 2^{u-s} \right]$$

First, we address the easy case, when $|\alpha| + i \leq u - s$. Then since $x_h - y_h$ is odd, $a'(x_h - y_h) \mod 2^{u - |\alpha| - i}$ is uniformly distributed over $\{0, \dots, 2^{u - |\alpha| - i} - 1\}$, and $a'y_h 2^{|\alpha| + i}$ over all multiples of $2^{|\alpha| + i}$ in $\{0, \dots, 2^u - 1\}$. Let

$$w = a'x_h 2^{|\alpha|+i} + \alpha x_h 2^i + ac \mod 2^u$$

 $v = a'y_h 2^{|\alpha|+i} + \alpha y_h 2^i + ac \mod 2^u$

Then the distribution of w-v matches that of $\delta 2^{|\alpha|+i} + \alpha (x_h - y_h) 2^i \pmod{2^u}$, where δ ranges over the integers $\{0, \ldots, 2^{u-|\alpha|-i}-1\}$. As proven in Lemma 14, for any given signed difference $w-v \in \{-2^{u-1}, \ldots, 2^{u-1}\}$,

$$\Pr\left[w + b \mod 2^u \operatorname{div} 2^{u-s} = v + b \mod 2^u \operatorname{div} 2^{u-s}\right] = \max\left(0, \frac{2^{u-s} - |w - v|}{2^{u-s}}\right)$$

Since $u-s \geq |\alpha|+i$, we can pair up intersections between w-v and $\{-2^{u-s},\dots,2^{u-s}-1\}$, matching value $z_1 < 0$ with $z_2 = z_1 + 2^{u-s} \geq 0$ so that $|z_1| + |z_2| = 2^{u-s}$. (If $w-v=\pm 2^{u-s}$, the contribution to the probability is zero and so it does no matter whether we match intersections in $\{-2^{u-s},\dots,2^{u-s}-1\}$ or $\{-2^{u-s}+1,\dots,2^{u-s}\}$.) Since on average each pair has collision probability $\frac{1}{2}$, and since there are $\frac{2^{u-s+1}}{2|\alpha|+i}$ intersecting values out of $2^{u-|\alpha|-i}$ candidate values of δ ,

$$p = \frac{1}{2} \cdot \frac{2^{u-s+1}}{2^{|\alpha|+i}} \cdot \frac{1}{2^{u-|\alpha|-i}} = \frac{1}{2^s}$$
.

In the hard case, $|\alpha|+i>u-s$, and there will be at most one intersection between u-v and $\{-2^{u-s}+1,\ldots,2^{u-s}-1\}$. The logic is essentially the same between the sub-cases $|\alpha|+i\geq u$ (for which w-v is fixed and unaffected by a') and $|\alpha|+i< u$ (where w-v may take $2^{u-|\alpha|-i}$ possible values, depending on a'). As before, the distribution of w-v matches $\delta 2^{|\alpha|+i}+\alpha \left(x_h-y_h\right)2^i\pmod{2^u}$, where δ ranges over integers in $\{0,\ldots,2^{\max(u-|\alpha|-i,0)}-1\}$. We have an intersection if $\alpha \left(x_h-y_h\right)2^i$ is within $\{-2^{u-s},\ldots,2^{u-s}\}$ of any multiple of $2^{|\alpha|+i}$. Consider the signed difference $d'=\left(\alpha \left(x_h-y_h\right)2^i\right)$ (mod $\min\left(2^{|\alpha|+i},2^u\right)$); our final collision probability will be, explicitly calculating the absolute value of the signed difference:

$$\begin{split} p &= \frac{1}{2^{\max(u-|\alpha|-i,0)}} \max \left(0, \frac{2^{u-s}-|d'|}{2^{u-s}}\right) \\ \text{where} \quad d' &= \min \left(2^{|\alpha|+i-1}, 2^{u-1}\right) - \\ &\left|\min \left(2^{|\alpha|+i-1}, 2^{u-1}\right) - \left(\alpha \left(x-y\right) \mod \min \left(2^{|\alpha|+i-1}, 2^{u-1}\right)\right)\right|. \end{split}$$

(Compared to the odd-multiply-shift hash, we evaluate a triangular function instead of the sum of two step functions.)

Finally, we consider the case of fixed a, where the offset b is to be determined. Write $b = \beta 2^{u-s-|\beta|} + b'$ where $b' \in \{0, \dots, 2^{u-s-|\beta|} - 1\}$, $w = ax + \beta 2^{u-s-|\beta|}$

and $v = ay + \beta 2^{u-s-|\beta|}$, so that

$$p := \Pr \left[(ax + b) \mod 2^u \operatorname{div} 2^{u-s} = (ay + b) \mod 2^u \operatorname{div} 2^{u-s} \right]$$

= $\Pr \left[(w + b') \mod 2^u \operatorname{div} 2^{u-s} = (v + b') \mod 2^u \operatorname{div} 2^{u-s} \right]$

Consider the case where $d=v-w \mod 2^u < 2^{u-s}$. (If $v-w \mod 2^u \ge 2^{u-s}$ and $w-v \mod 2^u \ge 2^{u-s}$, then p=0, so this is one half of the nontrivial case.) Define $\iota(z)=\max\left(0,z \mod 2^{u-s}-\left(2^{u-s}-2^{u-s-|\beta|}\right)\right)$ to count the $b'\in\left\{0,\ldots,2^{u-s-|\beta|}-1\right\}$ for which $(z+b')\operatorname{div}2^{u-s}>z\operatorname{div}2^{u-s}$. Then we have a two cases:

- $p = \frac{\iota(w) \iota(v)}{2^{u-s} |\beta|}$: holds if $w \mod 2^{u-s} > v \mod 2^{u-s}$ at b' = 0 we have no collision, and increasing $w \operatorname{div} 2^u$ but not $v \operatorname{div} 2^u$ will create a collision
- $p = 1 + \frac{\iota(w) \iota(v)}{2^{u-s-|\beta|}}$: holds if $w \mod 2^{u-s} \le v \mod 2^{u-s}$ at b' = 0 we have a collision, and increasing $v \operatorname{div} 2^u$ but not $w \operatorname{div} 2^u$ will prevent a collision

Note: if $w-v \mod 2^u \ge 2^{u-s}$, then the cases have probabilities $\frac{\iota(v)-\iota(w)}{2^{u-s-|\beta|}}$ and $1+\frac{\iota(v)-\iota(w)}{2^{u-s-|\beta|}}$.

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