

STAT 308 – Chapter 6

Background Information

We have previously discussed how we can model observed values of Y by our knowledge of the **independent variable** X . We discussed methods to assess if $\beta_1 > 0$ and create confidence intervals for β_1 and $\mu_{Y|X}$ and prediction intervals for Y . We also discussed **graphical methods** to assess the goodness of fit of our linear model. We will now discuss **numerical methods** to make inference on our linear model.

Important Definitions

Correlation Coefficient: A number describing the *strength* and *direction* of the *linear association* between X and Y .

The sample correlation coefficient is defined as

$$r = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{(\sum_{i=1}^n (x_i - \bar{x})^2 \sum_{i=1}^n (y_i - \bar{y})^2)^{1/2}} = \frac{s_{xy}}{s_x s_y},$$

where s_{xy} is the sample **covariance** between the observed x and y .

It can be noted that r is directly related to the estimated regression slope $\hat{\beta}_1$,

$$r = \frac{s_x}{s_y} \hat{\beta}_1.$$

Example:

Recall the **bloodpressure** dataset we have previously used. Calculate the correlation coefficient between Age and Systolic Blood Pressure.

```
library(tidyverse)
```

```
## -- Attaching packages ----- tidyverse 1.3.1 --
```

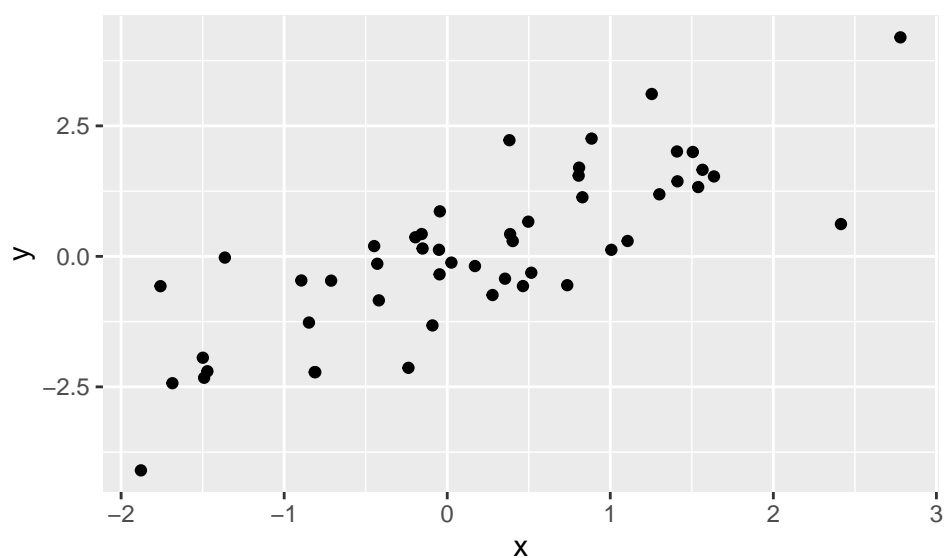
```
## v ggplot2 3.3.6      v purrr  0.3.4
## v tibble  3.1.7      v dplyr  1.0.9
## v tidyr   1.2.0      v stringr 1.4.0
## v readr   2.1.2      v forcats 0.5.1
```

```
## -- Conflicts ----- tidyverse_conflicts() --
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()     masks stats::lag()
```

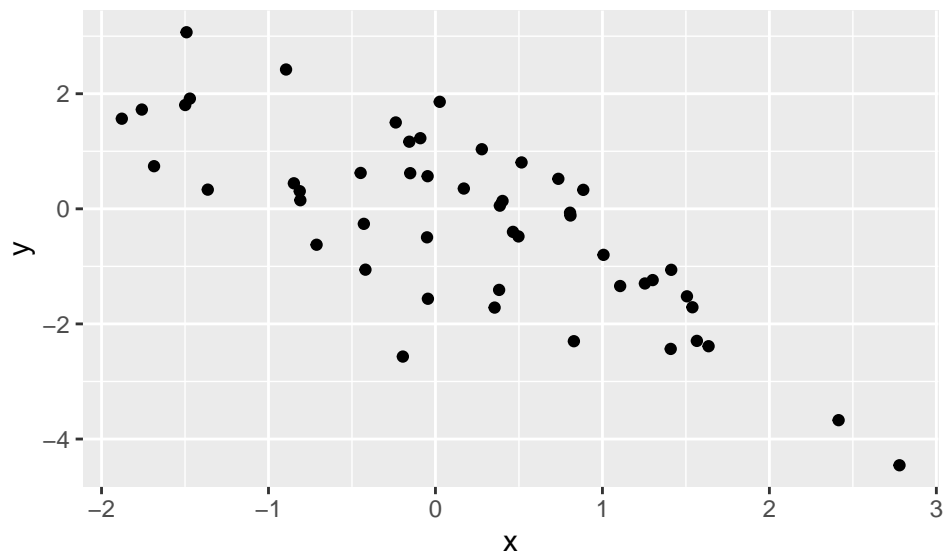
Properties of the sample correlation coefficient

- r is a value between -1 and 1.
- $r = 1$ means there is a direct positive linear relationship between X and Y .
- $r = -1$ means there is a direct negative linear relationship between X and Y .
- r is a *unitless* measure
 - r has the same sign as $\hat{\beta}_1$. That is
$$r > 0 \iff \hat{\beta}_1 > 0$$
and
$$r < 0 \iff \hat{\beta}_1 < 0$$
 - $r = 0$ means there is no **linear** association between X and Y . That does not mean there is no pattern at all that can be made out by the graph of X and Y .

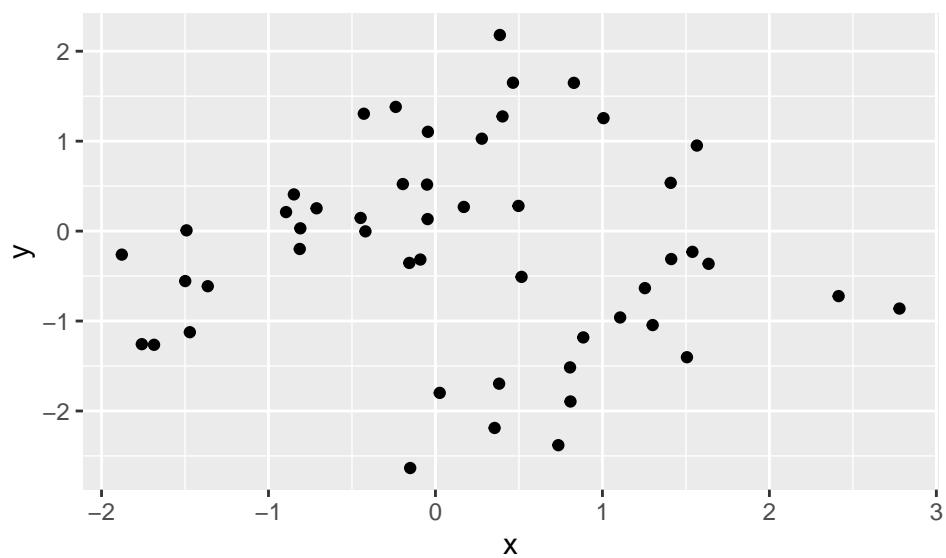
Examples of scatterplots and the correlation coefficient



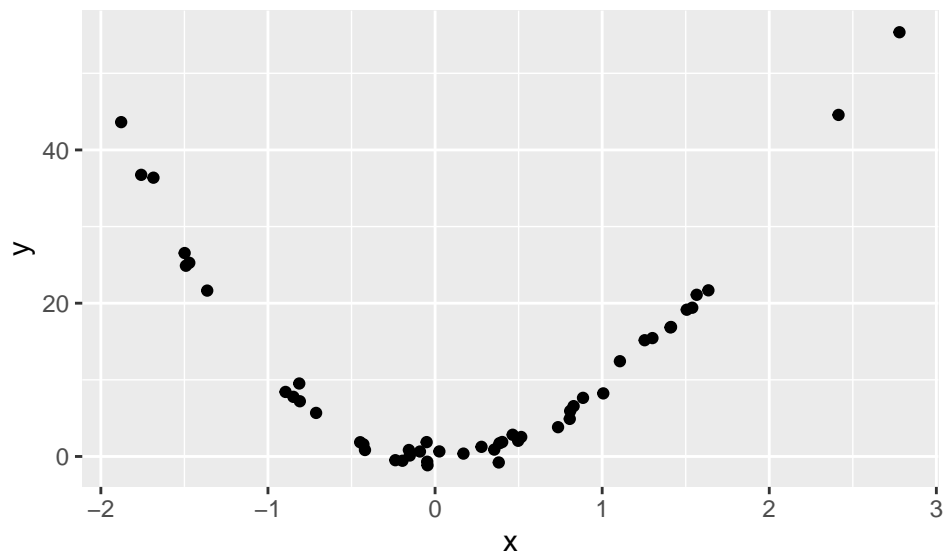
```
## [1] "Correlation between X and Y is 0.81."
```



```
## [1] "Correlation between X and Y is -0.77."
```



```
## [1] "Correlation between X and Y is -0.03."
```



```
## [1] "Correlation between X and Y is 0.05."
```

Example

What is the correlation coefficient of the blood pressure dataset?

Bivariate Normal Distribution: A distribution that describes the joint relationship between two different normally distributed random variables X and Y

Parameters of Bivariate Normal Distribution:

- μ_X : univariate mean of X
- μ_Y : univariate mean of Y
- σ_X^2 : univariate variance of X
- σ_Y^2 : univariate variance of Y
- ρ_{XY} : correlation between X and Y

A nice property of the bivariate normal distribution is that we can slice the distribution at a fixed value of X to obtain the *conditional distribution* of Y at a given value of X . This distribution is also normally distributed with

- $\mu_{Y|X} = \mu_Y + \frac{\rho_{XY}\sigma_Y}{\sigma_X}(X - \mu_X)$ and
- $\sigma_{Y|X}^2 = \sigma_Y^2(1 - \rho_{XY}^2)$.

Recall from the Chapter 5 and 6 notes, we can say that

- $\hat{\beta}_1 = \frac{rs_y}{s_x}$ and

- $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$.

If we substitute in the parameters for their respective estimates, we have

- $\beta_1 = \frac{\rho_{XY}\sigma_Y}{\sigma_X}$ and
- $\beta_0 = \mu_y - \beta_1\mu_x$.

Then, using some substitution, we have

- $\mu_{Y|X} = \beta_0 + \beta_1 X$,

showing the relationship between the least squares regression line and the bivariate normal distribution!

Now, if we were to take the formula for $\sigma_{Y|X}^2$ and solve for ρ_{XY}^2 , we would get

$$\rho_{XY}^2 = \frac{\sigma_Y^2 - \sigma_{Y|X}^2}{\sigma_Y^2}.$$

In other words, ρ_{XY}^2 is the

R-squared

R-Squared(r^2): the percent of variation in the response variable Y that can be explained through its linear relationship with the explanatory variable X

Formally,

$$r^2 = \frac{SSY - SSE}{SSY},$$

where $SSY = \sum_{i=1}^n (y_i - \bar{y})^2$ and $SSE = \sum_{i=1}^n (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i))^2$.

Example

Find the r^2 of the systolic blood pressure dataset. Interpret this value in the context of the given problem.

About R-squared

- Naturally, r^2 is the square of the sample correlation coefficient, so $0 \leq r^2 \leq 1$
- The larger the value of r^2 , the more variance in Y we can explain through its linear relationship with X , and thus, the stronger the linear relationship between the two variables
- If $r^2 = 1$, all of the variation in Y can be explained linearly by X (in other words, $SSE = 0$)
- If $r^2 = 0$, no variation in Y can be explained linearly by X
- r^2 does NOT measure the magnitude of $\hat{\beta}_1$ (i.e. r^2 can be close to one, but $\hat{\beta}_1$ may still be close to zero, or r^2 can be close to zero, but $\hat{\beta}_1$ may be large.)
- r^2 is NOT a measure of the appropriateness of the linear model.