#### Problem 1.

 $y_t$  is a is a stationary process with the autocovariance function  $\gamma_k$ .

$$\bar{y} := \frac{1}{n} \sum_{t=1}^{n} y_t.$$

Show that

$$\operatorname{Var}[\bar{y}] = \frac{\gamma_0}{n} + \frac{2}{n} \sum_{k=1}^{n-1} \left( 1 - \frac{k}{n} \right) \gamma_k = \frac{1}{n} \sum_{k=-n+1}^{n-1} \left( 1 - \frac{|k|}{n} \right) \gamma_k.$$

Solution.

$$\begin{aligned} \operatorname{Var}[\bar{y}] &= \operatorname{Cov}[\bar{y}, \bar{y}] \\ &= \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \operatorname{Cov}[y_i, y_j] \\ &= \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \gamma_{|i-j|} \end{aligned}$$

$$\begin{aligned} \operatorname{Var}[\bar{y}] &= \frac{1}{n^2} \left[ n\gamma_0 + 2(n-1)\gamma_1 + 2(n-2)\gamma_2 + \dots + 2(n-(n-1))\gamma_{n-1} \right] \\ &= \frac{1}{n^2} \left[ n\gamma_0 + 2\sum_{k=1}^{n-1} (n-k)\gamma_k \right] \\ &= \frac{1}{n} \left[ \gamma_0 + 2\sum_{k=1}^{n-1} \left( 1 - \frac{k}{n} \right) \gamma_k \right] \quad = \frac{\gamma_0}{n} + \frac{2}{n} \sum_{k=1}^{n-1} \left( 1 - \frac{k}{n} \right) \gamma_k \\ &= \frac{1}{n} \left[ \sum_{k=-(n-1)}^{-1} \left( 1 - \frac{|k|}{n} \right) \gamma_k + \left( 1 - \frac{|0|}{n} \right) \gamma_0 + \sum_{k=1}^{n-1} \left( 1 - \frac{|k|}{n} \right) \gamma_k \right] \\ &= \frac{1}{n} \sum_{k=-n+1}^{n-1} \left( 1 - \frac{|k|}{n} \right) \gamma_k \end{aligned}$$

## Problem 2.

Assume  $x_t$  is a stationary process and define  $y_t = \begin{cases} x_t & \text{for odd } t \\ x_t + 3 & \text{for even } t \end{cases}$ 

- (a) Show that  $Cov[y_t, y_{t-k}]$  is independent of t for all lags k.
- (b) Is  $y_t$  stationary?

Solution.

(a)

For stationary process  $x_t$ , let  $\gamma_k = \text{Cov}[x_t, x_{t-k}]$ 

$$\operatorname{Cov}[y_t,y_{t-k}] = \left\{ \begin{array}{ll} \operatorname{Cov}[x_t,x_{t-k}] & t \text{ odd, } k \text{ even} \\ \operatorname{Cov}[x_t,x_{t-k}+3] & t \text{ odd, } k \text{ odd} \\ \operatorname{Cov}[x_t+3,x_{t-k}+3] & t \text{ even, } k \text{ even} \\ \operatorname{Cov}[x_t+3,x_{t-k}] & t \text{ even, } k \text{ odd} \end{array} \right\} = \gamma_k$$

Therefore,  $Cov[y_t, y_{t-k}]$  is independent of t for all lags k.

#### Stationarity

- Strict Stationarity
  - Too strong for most applications
  - The whole probability structure must depend only on time differences
  - Second-order stationarity and an assumption of normality are sufficient to produce strict stationarity.
- Second-order Stationarity (Weak Stationarity)
  - Imposes conditions only on the first two moments of the series
  - The mean  $\mathbf{E}[x_t] = \mu$  is a fixed constant for all t
  - The autocovariances  $\mathrm{Cov}[x_t,x_{t-k}]=\gamma_k$  depend only on the time difference or time lag k for all t.

(b)

$$\mathbf{E}[y_t] = \begin{cases} x_t & t \text{ odd} \\ x_t + 3 & t \text{ even} \end{cases}$$

The mean of the process  $\{y_t\}$  is not a constant. Therefore, the process  $\{y_t\}$  is nonstationary.

## Problem 3.

Let  $\{y_t\}$  be a stationary process with an autocovariance function  $\gamma_k$ .

(a) Show that  $z_t = \nabla y_t = y_t - y_{t-1}$  is stationary by finding the mean and autocovariance function for  $z_t$ .

(b) Show that  $w_t = \nabla_2 y_t = z_t - z_{t-1} = y_t - 2y_{t-1} + y_{t-2}$  is stationary.

Solution.

(a)

Since  $\{y_t\}$  is a stationary process, it has constant mean  $\mu$  over time and autocovariance function  $\gamma_k$  which does not depend on time t. We have  $z_t = \nabla y_t = y_t - y_{t-1}$ 

$$\begin{aligned} \mathbf{E}[z_t] &= \mathbf{E}[y_t - y_{t-1}] \\ &= \mathbf{E}[y_t] - \mathbf{E}[y_{t-1}] \\ &= \mu - \mu = 0 \end{aligned}$$

The mean function of  $\{z_t\}$  is constant over time t.

$$\begin{split} &\operatorname{Cov}[z_t, z_{t-k}] = \operatorname{Cov}[y_t - y_{t-1}, y_{t-k} - y_{t-k-1}] \\ &= \operatorname{Cov}[y_t, y_{t-k}] + \operatorname{Cov}[y_t, -y_{t-k-1}] + \operatorname{Cov}[-y_{t-1}, y_{t-k}] + \operatorname{Cov}[-y_{t-1}, -y_{t-k-1}] \\ &= \operatorname{Cov}[y_t, y_{t-k}] - \operatorname{Cov}[y_t, y_{t-k-1}] - \operatorname{Cov}[y_{t-1}, y_{t-k}] + \operatorname{Cov}[y_{t-1}, y_{t-k-1}] \\ &= \gamma_k - \gamma_{k+1} - \gamma_{k-1} + \gamma_k \\ &= -\gamma_{k+1} + 2\gamma_k - \gamma_{k-1} \quad := \gamma_k' \end{split}$$

The autocovariance function of  $\{z_t\}$  only depends on time lag k. So we conclude that  $\{z_t\}$  is stationary.

(b)

$$\begin{split} \mathbf{E}[w_t] &= \mathbf{E}[z_t - z_{t-1}] \\ &= \mathbf{E}[z_t] - \mathbf{E}[z_{t-1}] \\ &= 0 - 0 = 0 \end{split}$$

$$\begin{split} \operatorname{Cov}[w_t, w_{t-k}] &= \operatorname{Cov}[z_t - z_{t-1}, z_{t-k} - z_{t-k-1}] \\ &= \operatorname{Cov}[z_t, z_{t-k}] + \operatorname{Cov}[z_t, -z_{t-k-1}] + \operatorname{Cov}[-z_{t-1}, z_{t-k}] + \operatorname{Cov}[-z_{t-1}, -z_{t-k-1}] \\ &= \operatorname{Cov}[z_t, z_{t-k}] - \operatorname{Cov}[z_t, z_{t-k-1}] - \operatorname{Cov}[z_{t-1}, z_{t-k}] + \operatorname{Cov}[z_{t-1}, z_{t-k-1}] \\ &= \gamma_k' - \gamma_{k+1}' - \gamma_{k-1}' + \gamma_k' \\ &= -\gamma_{k+1}' + 2\gamma_k' - \gamma_{k-1}' \end{split}$$

The mean function of  $\{w_t\}$  is constant over time t. The autocovariance function of  $\{w_t\}$  only depends on time lag k. So we conclude that  $\{w_t\}$  is stationary.

## Problem 4.

Let  $x_t$  be stationary with  $\mathbf{E}[x_t]=0$ ,  $\mathrm{Var}[x_t]=1$ , autocorrelation function  $\rho_k$ . Define that  $\mu_t$  is a nonconstant function and  $\sigma_t$  is a positively nonconstant function (that is to say:  $\mu_t$  and  $\sigma_t$  are both deterministic and in function of t). Now we observe a time series formulated as

$$y_t = \mu_t + \sigma_t x_t.$$

- (a) Find the mean and autocovariance function of  $y_t$ .
- (b) Show that the autocorrelation of  $y_t$  depends only on the lag k. Is  $y_t$  stationary?
- (c) Let  $\mu_t = \mu_0$  be a constant value. Justify that  $y_t$  is still nonstationary.

Solution.

(a)

$$\begin{split} \mathbf{E}[y_t] &= \mathbf{E}[\mu_t + \sigma_t x_t] \\ &= \mathbf{E}[\mu_t] + \mathbf{E}[\sigma_t x_t] \\ &= \mu_t + 0 = \mu_t \end{split}$$

$$\begin{split} \operatorname{Cov}[y_t, y_{t-k}] = & \operatorname{Cov}[\mu_t + \sigma_t x_t, \mu_{t-k} + \sigma_{t-k} x_{t-k}] \\ = & \operatorname{Cov}[\sigma_t x_t, \sigma_{t-k} x_{t-k}] \\ = & \sigma_t \sigma_{t-k} \operatorname{Cov}[x_t, x_{t-k}] \\ = & \sigma_t \sigma_{t-k} \operatorname{Corr}[x_t, x_{t-k}] \sqrt{\operatorname{Var}[x_t] \operatorname{Var}[x_{t-k}]} \\ = & \sigma_t \sigma_{t-k} \operatorname{Corr}[x_t, x_{t-k}] \operatorname{Var}[x_t] \\ = & \sigma_t \sigma_{t-k} \rho_k \end{split}$$

(b)

$$\operatorname{Var}[y_x] = \sigma_t^2 \underbrace{\rho_0}_{1} = \sigma_t^2$$

$$\begin{split} \operatorname{Corr}[y_t, y_{t-k}] &= \frac{\operatorname{Cov}[y_t, y_{t-k}]}{\sqrt{\operatorname{Var}[y_t]\operatorname{Var}[y_{t-k}]}} \\ &= \frac{\sigma_t \sigma_{t-k} \rho_k}{\sqrt{(\sigma_{t-k}^2 \rho_k)}} \\ &= \frac{\sigma_t \sigma_{t-k} \rho_k}{\sigma_t \sigma_{t-k}} = \rho_k \end{split}$$

The autocorrelation of  $y_t$  depends only on the lag k.

 $y_t$  is not stationary because the mean function  $\mu_t$  is not a constant over time t.

(c)

Let  $\mu_t=\mu_0$  be a constant value.  $y_t$  is still nonstationary because the autocovariance function depend on time t.

## Problem 5.

Let  $x_t$  be the series of the "expected" measurements during the production process. Because the measuring tool itself won't be perfect, we observe  $y_t = x_t + e_t$ , assuming  $x_t$  and  $e_t$  are independent. In general, we call  $x_t$  the signal and  $e_t$  the measurement (white) noise. If  $x_t$  is stationary with the autocorrelation function  $\rho_k$ , show that  $y_t$  is also a stationary process with

$$\operatorname{Corr}[y_t, y_{t-k}] = \frac{\rho_k}{1 + \frac{\sigma_e^2}{\sigma_z^2}}, \quad \text{ for } k \ge 1.$$

 $\frac{\sigma_e^2}{\sigma_x^2}$  is usually referred to as the **signal-to-noise ratio**, or **SNR** for short. The larger the SNR, the closer the autocorrelation function of the observed series  $y_t$  is to the autocorrelation function of the desired signal  $x_t$ .

#### White Noise Process

The most fundamental example of a stationary process is a sequence of **independent** and identically distributed random variables, denoted as  $\alpha_1, \ldots, \alpha_t, \ldots$ , which we also assume to have **mean zero** and variance  $\sigma_{\alpha}^2$ . This process is strictly stationary and is referred to as a **white noise process**. Because independence implies that the  $\alpha_t$  are uncorrelated, its autocovariance function is simply

$$\gamma_k = \mathbf{E}[x_t, x_{t-k}] = \begin{cases} \sigma_\alpha^2 & k = 0 \\ 0 & k \neq 0 \end{cases}$$

Solution.

$$\begin{split} \mathbf{E}[y_t] &= \mathbf{E}[x_t + e_t] \\ &= \mathbf{E}[x_t] + \mathbf{E}[e_t] \\ &= \mu \quad \text{is a fixed constant} \end{split}$$

$$\begin{split} \operatorname{Cov}[y_t,y_{t-k}] &= \operatorname{Cov}[x_t + e_t, x_{t-k} + e_{t-k}] \\ &= \operatorname{Cov}[x_t, x_{t-k}] + \operatorname{Cov}[x_t, e_{t-k}] + \operatorname{Cov}[e_t, x_{t-k}] + \operatorname{Cov}[e_t, e_{t-k}] \\ &= \operatorname{Corr}[x_t, x_{t-k}] \sigma_x^2 + \operatorname{Cov}[e_t, e_{t-k}] \quad x_t \text{ and } e_t \text{ are independent} \\ &= \begin{cases} \rho_0 \sigma_x^2 + \sigma_e^2 & k = 0 \\ \rho_k \sigma_x^2 + 0 & k \geq 1 \end{cases} \\ &= \begin{cases} \sigma_x^2 + \sigma_e^2 = \operatorname{Var}[y_t] & k = 0 \\ \rho_k \sigma_x^2 & k \geq 1 \end{cases} \end{split}$$

 $y_t$  is a stationary process with constant mean and autocovariances depending only on the time difference or time lag k for all t.

$$\begin{split} \operatorname{Corr}[y_t, y_{t-k}] &= \frac{\operatorname{Cov}[y_t, y_{t-k}]}{\sqrt{\operatorname{Var}[y_t]\operatorname{Var}[y_{t-k}]}} \\ &= \frac{\rho_k \sigma_x^2}{\operatorname{Var}[y_t]} \\ &= \frac{\rho_k \sigma_x^2}{\sigma_x^2 + \sigma_e^2} \\ &= \frac{\rho_k}{1 + \frac{\sigma_e^2}{\sigma_x^2}} \quad \text{for } k \geq 1. \end{split}$$

#### Problem 6. Cyclical Behavior and Periodicity

Mixtures of periodic series with multiple frequencies and amplitudes:

Suppose

$$y_t = \alpha_0 + \sum_{i=1}^{q} \left[ \alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t) \right],$$

where  $\alpha_0, f_1, f_2, \ldots, f_q$  are constants and  $\alpha_1, \alpha_2, \ldots, \alpha_q, \beta_1, \beta_2, \ldots, \beta_q$  are independent random variables with zero means and variances

$$\operatorname{Var}[\alpha_i] = \operatorname{Var}[\beta_i] = \sigma_i^2$$
.

Show that  $y_t$  is stationary and find its autocovariance function.

(Hint: show  $Cov[y_t, y_s]$  depends only on t - s.

Solution.

$$\begin{split} \mathbf{E}[y_t] &= \mathbf{E}\left[\alpha_0 + \sum_{i=1}^q \left[\alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t)\right]\right] \\ &= \mathbf{E}[\alpha_0] + \mathbf{E}\left[\sum_{i=1}^q \left[\alpha_i \cos(2\pi f_i t) + \beta_i \sin(2\pi f_i t)\right]\right] \\ &= \mathbf{E}[\alpha_0] + \sum_{i=1}^q \left\{\cos(2\pi f_i t) \underbrace{\mathbf{E}[\alpha_i]}_{0} + \sin(2\pi f_i t) \underbrace{\mathbf{E}[\beta_i]}_{0}\right\} \\ &= \mathbf{E}[\alpha_0] \quad \text{constant} \end{split}$$

Let

$$s = t - k$$

$$u_i(t) = \alpha_i \cos(2\pi f_i t) = \alpha_i c_{it}$$

$$v_i(t) = \beta_i \sin(2\pi f_i t) = \beta_i s_{it}$$

$$\begin{split} &\operatorname{Cov}[y_t,y_s] = \operatorname{Cov}\left[\left\{\alpha_0 + \sum_{i=1}^q \left[u_i(t) + v_i(t)\right]\right\}, \left\{\alpha_0 + \sum_{i=1}^q \left[u_i(s) + v_i(s)\right]\right\}\right] \\ &= \operatorname{Cov}\left[\left\{\sum_{i=1}^q \left[u_i(t) + v_i(t)\right]\right\}, \left\{\sum_{i=1}^q \left[u_i(s) + v_i(s)\right]\right\}\right] \\ &= \sum_{i=1}^q \sum_{j=1}^q \left\{\operatorname{Cov}[u_i(t), u_j(s)] + \operatorname{Cov}[u_i(t), v_j(s)] + \operatorname{Cov}[v_i(t), u_j(s)] + \operatorname{Cov}[v_i(t), v_j(s)]\right\} \\ &= \sum_{i=1}^q \sum_{j=1}^q \left\{\operatorname{Cov}[\alpha_i c_{it}, \alpha_j c_{js}] + \operatorname{Cov}[\alpha_i c_{it}, \beta_j s_{js}] + \operatorname{Cov}[\beta_i s_{it}, \alpha_j c_{js}] + \operatorname{Cov}[\beta_i s_{it}, \beta_j s_{js}]\right\} \\ &= \sum_{i=1}^q \left\{c_{it} c_{is} \operatorname{Cov}[\alpha_i, \alpha_i] + s_{it} s_{is} \operatorname{Cov}[\beta_i, \beta_i]\right\} \quad \alpha_1, \alpha_2, \dots, \alpha_q, \beta_1, \beta_2, \dots, \beta_q \text{ are independent r.v.} \end{split}$$

$$\begin{split} &= \sum_{i=1}^{q} \left\{ c_{it} c_{is} \mathrm{Var}[\alpha_i] + s_{it} s_{is} \mathrm{Var}[\beta_i] \right\} \\ &= \sum_{i=1}^{q} \left\{ c_{it} c_{is} + s_{it} s_{is} \right\} \sigma_i^2 \\ &= \sum_{i=1}^{q} \left\{ \cos(2\pi f_i t) \cos(2\pi f_i s) + \sin(2\pi f_i t) \sin(2\pi f_i s) \right\} \sigma_i^2 \\ &= \sum_{i=1}^{q} \left\{ \cos(2\pi f_i t - 2\pi f_i s) \right\} \sigma_i^2 \\ &= \sum_{i=1}^{q} \left\{ \cos(2\pi f_i k) \right\} \sigma_i^2 \end{split}$$

 $\gamma_k = \operatorname{Cov}[y_t, y_{t-k}] = \operatorname{Cov}[y_t, y_s] \text{ depends only on } k = t - s.$ 

 $y_t$  is a stationary process with constant mean and autocovariances depending only on the time difference or time lag k for all t.

# Trigonometric Identities (Sum & Difference Identities)

$$\begin{aligned} \cos(\alpha \pm \beta) &= \cos(\alpha)\cos(\beta) \mp \sin(\alpha)\sin(\beta) \\ \sin(\alpha \pm \beta) &= \sin(\alpha)\cos(\beta) \pm \cos(\alpha)\sin(\beta) \\ \tan(\alpha \pm \beta) &= \frac{\tan(\alpha) \pm \tan(\beta)}{1 \mp \tan(\alpha)\tan(\beta)} \end{aligned}$$