

Problem 1.

Show that for an MA(1) process

- $\max_{-\infty < \theta < \infty} \rho_1 = 0.5$
- $\min_{-\infty < \theta < \infty} \rho_1 = -0.5$

White Noise Process

The most fundamental example of a stationary process is a sequence of **independent and identically distributed** random variables, denoted as $\alpha_1, \dots, \alpha_t, \dots$, which we also assume to have **mean zero** and variance σ_α^2 . This process is strictly stationary and is referred to as a **white noise process**. Because independence implies that the α_t are uncorrelated, its autocovariance function is simply

$$\gamma_k = \mathbf{E}[x_t, x_{t-k}] = \begin{cases} \sigma_\alpha^2 & k = 0 \\ 0 & k \neq 0 \end{cases}.$$

Thus, the autocorrelation function of white noise has a particularly simple form

$$\rho_k = \begin{cases} 1 & k = 0 \\ 0 & k \neq 0 \end{cases}.$$

Solution.

Consider the MA(1) model $y_t = a_t + \theta a_{t-1}$.

$$\mathbf{E}[y_t] = \mathbf{E}[a_t + \theta a_{t-1}] = 0 \quad (1)$$

$$\begin{aligned} \gamma_0 &= \text{Var}[y_t] \\ &= \text{Var}[a_t + \theta a_{t-1}] \\ &= \text{Var}[a_t] + \theta^2 \text{Var}[a_{t-1}] \\ &= \text{Var}[a_t] + \theta^2 \text{Var}[a_{t-1}] \\ &= (1 + \theta^2) \sigma_a^2 \end{aligned} \quad (2)$$

$$\begin{aligned} \gamma_1 &= \text{Cov}[y_t, y_{t+1}] \\ &= \text{Cov}[a_t + \theta a_{t-1}, a_{t+1} + \theta a_t] \\ &= \text{Cov}[a_t, a_{t+1}] + \text{Cov}[a_t, \theta a_t] + \text{Cov}[\theta a_{t-1}, a_{t+1}] + \text{Cov}[\theta a_{t-1}, \theta a_t] \\ &= 0 + \theta \sigma_a^2 + 0 + 0 \\ &= \theta \sigma_a^2 \end{aligned} \quad (3)$$

$$\begin{aligned}\rho_1 &= \frac{\gamma_1}{\gamma_0} \\ &= \frac{\theta \sigma_a^2}{(1 + \theta^2) \sigma_a^2} \\ &= \frac{\theta}{1 + \theta^2}\end{aligned}\tag{4}$$

$$\begin{aligned}\frac{d}{d\theta} \rho_1 &= \frac{(1 + \theta^2) - \theta(2\theta)}{(1 + \theta^2)^2} \\ &= \frac{1 - \theta^2}{(1 + \theta^2)^2}\end{aligned}\tag{5}$$

Set the first derivative to zero

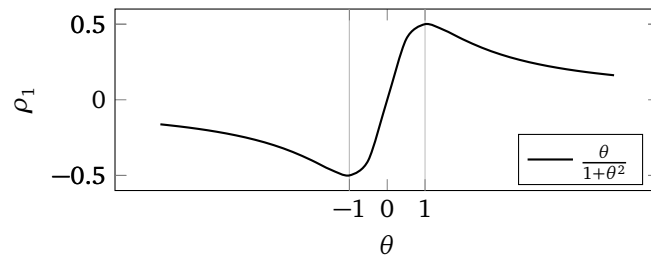
$$\frac{d}{d\theta} \rho_1 = \frac{1 - \theta^2}{(1 + \theta^2)^2} \stackrel{\text{set}}{=} 0\tag{6}$$

Solving the equation we have

$$\theta = \pm 1\tag{7}$$

So we have

$$\begin{aligned}\max_{\theta} \rho_1 &= \rho_1(\theta = -1) \\ &= -0.5 \\ \min_{\theta} \rho_1 &= \rho_1(\theta = 1) \\ &= 0.5\end{aligned}\tag{8}$$



■

Problem 2.

For an AR(2) process $y_t - 1.0y_{t-1} + 0.5y_{t-2} = a_t$:

1. Calculate ρ_1 .
2. Using ρ_0 and ρ_1 as starting values and the difference equation form for the autocorrelation function, calculate the values for ρ_k for $k = 2, \dots, 15$.

Solution.

The ACF of an AR(2) Process

For an AR(2) process $y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + a_t$:

$$\begin{aligned}
 \gamma_k &= \text{Cov}[y_t, y_{t+k}] \\
 &= \text{Cov}[y_t, (\phi_1 y_{t-1+k} + \phi_2 y_{t-2+k} + a_{t+k})] \\
 &= \text{Cov}[y_t, \phi_1 y_{t-1+k}] + \text{Cov}[y_t, \phi_2 y_{t-2+k}] + \text{Cov}[y_t, a_{t+k}] \\
 &= \text{Cov}[y_t, \phi_1 y_{t-1+k}] + \text{Cov}[y_t, \phi_2 y_{t-2+k}] + \text{Cov}[(\phi_1 y_{t-1} + \phi_2 y_{t-2} + a_t), a_{t+k}] \\
 &= \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2} + \text{Cov}[a_t, a_{t+k}] \\
 &= \begin{cases} \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2} + \sigma_a^2 & k = 0 \\ \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2} & k \neq 0 \end{cases}
 \end{aligned} \tag{9}$$

For $k = 0$

$$\begin{aligned}
 \text{Var}[y_t] = \gamma_0 &= \phi_1 \gamma_{-1} + \phi_2 \gamma_{-2} + \sigma_a^2 \\
 &= \phi_1 \gamma_1 + \phi_2 \gamma_2 + \sigma_a^2
 \end{aligned} \tag{10}$$

For $k = 1, 2, \dots$

$$\gamma_k = \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2} \tag{11}$$

Divide (11) through by γ_0 to obtain the difference equation for the ACF of the process:

$$\rho_k = \phi_1 \rho_{k-1} + \phi_2 \rho_{k-2} \tag{12}$$

For $k = 1$

$$\begin{aligned}
 \rho_1 &= \phi_1 \rho_0 + \phi_2 \rho_{-1} \\
 \rho_1 &= \phi_1 \rho_0 + \phi_2 \rho_1 \\
 (1 - \phi_2) \rho_1 &= \phi_1 \\
 \rho_1 &= \frac{\phi_1}{1 - \phi_2}
 \end{aligned} \tag{13}$$

using the initial condition $\rho_0 = 1$.

For an AR(2) process $y_t - 1.0y_{t-1} + 0.5y_{t-2} = a_t$:

- $\phi_1 = 1$
- $\phi_2 = -0.5$

For $k = 1, 2, \dots, 15$, the difference equation for the ACF of the process:

$$\rho_k = \rho_{k-1} + (-0.5)\rho_{k-2} \quad (14)$$

For $k = 1$

$$\rho_1 = \frac{1}{1 + 0.5} = 2/3 \quad (15)$$

k	ρ_k
0	1
1	0.666666667
2	0.166666667
3	-0.166666667
4	-0.25
5	-0.166666667
6	-0.041666667
7	0.041666667
8	0.0625
9	0.041666667
10	0.010416667
11	-0.010416667
12	-0.015625
13	-0.010416667
14	-0.002604167
15	0.002604167

■

Problem 3.

Put the following four models in B notation, and check whether it is stationary and invertible.

1. $y_t = a_t - 1.3a_{t-1} + 0.4a_{t-2}$.
2. $y_t - 0.5y_{t-1} = a_t - 1.3a_{t-1} + 0.4a_{t-2}$
3. $y_t - 1.5y_{t-1} + 0.6y_{t-2} = a_t$
4. $y_t - y_{t-1} = a_t - 0.5a_{t-1}$

Problem 4.

For each of the models of Problem 3, obtain:

- (a) The first three ψ_j weights of the model form: $y_t = a_t + \psi_1 a_{t-1} + \psi_2 a_{t-2} + \dots$
- (b) The first three π_j weights of the model form: $y_t = \pi_1 y_{t-1} + \pi_2 y_{t-2} + \dots + a_t$
- (c) $\text{Var}[y_t]$, assuming that $\sigma_a^2 = 1.0$

Solution.

1. MA(2)

For MA(q) model

$$y_t = a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2} - \dots - \theta_q a_{t-q} \quad (16)$$

Using the backshift operator

$$(1 - \theta_1 B - \theta_2 B^2 - \dots - \theta_q B^q) a_t = y_t \quad (17)$$

We define the **MA characteristic polynomial(moving average operator)**

$$\theta(B) = 1 - \theta_1 B - \theta_2 B^2 - \dots - \theta_q B^q \quad (18)$$

and the corresponding **MA characteristic equation**

$$\theta(B) = 1 - \theta_1 B - \theta_2 B^2 - \dots - \theta_q B^q = 0 \quad (19)$$

It can be shown that MA(q) model is **invertible**; that is, there are coefficients π_i such that

$$y + t = \pi_1 y_{t-1} + \pi_2 y_{t-2} + \pi_3 y_{t-3} + \dots + a_t \quad (20)$$

if and only if the roots of the MA characteristic equation exceed 1 in modulus.

For the MA(2) process:

$$y_t = a_t - 1.3a_{t-1} + 0.4a_{t-2} \quad (21)$$

B Operator Form

$$\begin{aligned} y_t &= \theta(B)a_t \\ &= (1 - 1.3B + 0.4B^2)a_t \end{aligned} \quad (22)$$

Stationarity

$$\begin{aligned} E[y_t] &= E[a_t - 1.3a_{t-1} + 0.4a_{t-2}] \\ &= 0 \quad \text{constant} \end{aligned} \quad (23)$$

$$\begin{aligned} \gamma_k &= \text{Cov}[y_t, y_{t-k}] \\ &= \text{Cov}[(a_t - 1.3a_{t-1} + 0.4a_{t-2}), (a_{t-k} - 1.3a_{t-k-1} + 0.4a_{t-k-2})] \\ &= \begin{cases} (1 + 1.3^2 + 0.4^2)\sigma_a^2 & k = 0 \\ (-1.3 + (-1.3)0.4)\sigma_a^2 & k = 1 \\ (-1.3)\sigma_a^2 & k = 2 \\ 0 & k > 2 \end{cases} \end{aligned} \quad (24)$$

The mean function of y_t is constant over time t . The autocovariance function of y_t only depends on time lag k . So we conclude that y_t is stationary.

Moving average processes are always stationary.

Invertibility

We can obtain the MA characteristic equation

$$\begin{aligned} \theta(B) &= 1 - 1.3B + 0.4B^2 = 0 \\ (1 - 0.5B)(1 - 0.8B) &= 0 \end{aligned} \quad (25)$$

The roots of the MA characteristic equation exceed 1 in modulus. Thus, The MA(2) process is invertible.

ψ weights

$$y_t = a_t - 1.3a_{t-1} + 0.4a_{t-2} \quad (26)$$

$$\begin{aligned} \psi_0 &= 1 \\ \psi_1 &= -1.3 \\ \psi_2 &= 0.4 \\ \psi_3 &= 0 \end{aligned} \quad (27)$$

π weights

Infinite autoregressive representation

$$\pi(B)y_t = y_t - \sum_{j=1}^{\infty} \pi_j y_{t-j} = a_t \quad (28)$$

where

$$\pi(B) = \theta^{-1}(B) \quad (29)$$

The weights π are determined from the relation $\theta(B)\pi(B) = 1$ to satisfy

$$\pi_j = \theta_1 \pi_{j-1} + \theta_2 \pi_{j-2} + \cdots + \theta_q \pi_{j-q} \quad j > 0 \quad (30)$$

with $\pi_0 = -1, \pi_j = 0$ for $j < 0$, from which the weights π_j can easily be computed recursively in terms of θ_i

$$\begin{aligned} \pi_0 &= -1 \\ \pi_1 &= \theta_1 \pi_0 + \theta_2 \pi_{-1} \\ &= (1.3)(-1) + (-0.4)(0) \\ &= -1.7 \quad -1.3 \\ \pi_2 &= \theta_1 \pi_1 + \theta_2 \pi_0 \\ &= (1.3)(-1.7) + (-0.4)(-1) \\ &= -1.81 \quad -1.29 \\ \pi_3 &= \theta_1 \pi_2 + \theta_2 \pi_1 \\ &= (1.3)(-1.81) + (-0.4)(-1.7) \\ &= -1.673 \quad -1.157 \end{aligned} \quad (31)$$

Variance

$$\begin{aligned} \text{Var}[y_t] &= \gamma_0 = (1 + 1.3^2 + 0.4^2) \sigma_a^2 \\ &= 1 + 1.3^2 + 0.4^2 \\ &= 2.85 \end{aligned} \quad (32)$$

2. ARMA(1,2) / MA(1)

$$y_t - 0.5y_{t-1} = a_t - 1.3a_{t-1} + 0.4a_{t-2} \quad (33)$$

B Operator Form

In operator form

$$\begin{aligned} (1 - 0.5B)y_t &= (1 - 1.3B + 0.4B^2)a_t \\ (1 - 0.5B)y_t &= (1 - 0.8B)(1 - 0.5B)a_t \\ \phi(B)y_t &= \theta(B)a_t \\ y_t &= \frac{\theta(B)}{\phi(B)}a_t \end{aligned} \quad (34)$$

The ARMA(1,2) process can be reduced to an MA(1) process.

$$\begin{aligned} y_t &= (1 - 0.8B)a_t \\ y_t &= \theta(B)a_t \end{aligned} \quad (35)$$

Stationarity

A stationary solution to (33) exists if and only if all the roots of the AR characteristic equation $\phi(B) = 0$ exceed unity in modulus.

$$\begin{aligned} \phi(B) &= 1 - 0.5B = 0 \\ B &= 2 \end{aligned} \quad (36)$$

The root of the AR characteristic equation $\phi(B) = 0$ exceed unity in modulus. Thus, the process is stationary.

An MA process is always stationary.

Invertibility

The roots of $\theta(B) = 0$ must lie outside the unity circle if the process is to be invertible.

$$\begin{aligned} \theta(B) &= 1 - 1.3B + 0.4B^2 = 0 \\ (1 - 0.8B)(1 - 0.5B) &= 0 \end{aligned} \quad (37)$$

$$\begin{aligned} B_1 &= 2.0 \\ B_2 &= 1.25 \end{aligned} \quad (38)$$

The root of the MA characteristic equation $\theta(B) = 0$ exceed unity in modulus. Thus, the process is invertible. ψ weights

Moving average representation

$$y_t = \psi(B)a_t = \sum_{j=0}^{\infty} \psi_j a_{t-j} \quad (39)$$

where

$$\psi(B) = \frac{\theta(B)}{\phi(B)} \quad (40)$$

The weights ψ are determined from the relation $\psi(B)\phi(B) = \theta(B)$ to satisfy

$$\psi_j = \phi_1\psi_{j-1} + \phi_2\psi_{j-2} + \cdots + \phi_p\psi_{j-p} - \theta_j \quad j > 0 \quad (41)$$

with $\psi_0 = 1, \psi_j = 0$ for $j < 0$, and $\theta_j = 0$ for $j > q$

$$\begin{aligned} \psi_0 &= 1 \\ \psi_1 &= \phi_1\psi_0 - \theta_1 \\ &= (0.5)(1) - (1.3) \\ &= -0.8 \\ \psi_2 &= \phi_1\psi_1 - \theta_2 \\ &= (0.5)(-0.8) - (-0.4) \\ &= 0 \\ \psi_3 &= \phi_1\psi_2 - \theta_3 \\ &= (0.5)(0) - (0) \\ &= 0 \end{aligned} \quad (42)$$

π weights

Infinite autoregressive representation

$$\pi(B)y_t = y_t - \sum_{j=1}^{\infty} \pi_j y_{t-j} = a_t \quad (43)$$

where

$$\pi(B) = \frac{\phi(B)}{\theta(B)} \quad (44)$$

The weights π are determined from the relation $\theta(B)\pi(B) = \phi(B)$ to satisfy

$$\pi_j = \theta_1\pi_{j-1} + \theta_2\pi_{j-2} + \cdots + \theta_q\pi_{j-q} + \phi_j \quad j > 0 \quad (45)$$

with $\pi_0 = -1$, $\pi_j = 0$ for $j < 0$, and $\phi_j = 0$ for $j > p$

$$\begin{aligned}
 \pi_0 &= -1 \\
 \pi_1 &= \theta_1 \pi_0 + \theta_2 \pi_{-1} + \phi_1 \\
 &= (1.3)(-1) + (-0.4)(0) + (0.5) \\
 &= -0.8 \\
 \pi_2 &= \theta_1 \pi_1 + \theta_2 \pi_0 + \phi_2 \\
 &= (1.3)(-0.8) + (-0.4)(-1) + (0) \\
 &= -0.64 \\
 \pi_3 &= \theta_1 \pi_2 + \theta_2 \pi_1 + \phi_3 \\
 &= (1.3)(-0.64) + (-0.4)(-0.8) + (0) \\
 &= -0.512
 \end{aligned} \tag{46}$$

Variance

$$\begin{aligned}
 \text{Var}[y_t] &= \text{Var}\left[\sum_{j=0}^{\infty} \psi_j a_{t-j}\right] \\
 &= (\psi_0^2 + \psi_1^2 + \cdots) \text{Var}[a_t] \\
 &= 1^2 + (-0.8)^2 \\
 &= 1.64
 \end{aligned} \tag{47}$$

3. AR(2)

An autoregressive model of order p , abbreviated AR(p), is of the form

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + a_t. \quad (48)$$

Using the backshift operator

$$(1 - \phi_1 B - \phi_2 B^2 - \cdots - \phi_p B^p) y_t = a_t \quad (49)$$

We define the **AR characteristic polynomial (autoregressive operator)**

$$\phi(B) = 1 - \phi_1 B - \phi_2 B^2 - \cdots - \phi_p B^p \quad (50)$$

and the corresponding **AR characteristic equation**

$$\phi(B) = 1 - \phi_1 B - \phi_2 B^2 - \cdots - \phi_p B^p = 0 \quad (51)$$

The process $\phi(B)y_t = a_t$ can be written as

$$y_t = \phi^{-1}(B)a_t \equiv \psi(B)a_t = \sum_{j=0}^{\infty} \psi_j a_{t-j} \quad (52)$$

provided that the right-side expression is convergent. Using the factorization

$$\phi(B) = (1 - G_1 B)(1 - G_2 B) \cdots (1 - G_p B) \quad (53)$$

where $G_1^{-1}, G_2^{-1}, \dots, G_p^{-1}$ are the roots of equation $\phi(B) = 0$, and expanding $\phi^{-1}(B)$ in partial fractions yields

$$y_t = \phi^{-1}(B)a_t = \sum_{i=1}^p \frac{K_i}{1 - G_i B} a_t \quad (54)$$

Hence, if $\psi(B) = \phi^{-1}(B)$ is to be convergent series for $|B| < 1$, that is if the weights $\psi_j = \sum_{i=1}^p K_i G_i^j$ are to be absolutely summable so that AR(p) process is stationary, we must have $|G_i| < 1$, for $i = 1, \dots, p$. Equivalently, the roots of the **AR characteristic equation** $\phi(B) = 0$ must lie outside the unity circle.

For the AR(2) process:

$$y_t - 1.5y_{t-1} + 0.6y_{t-2} = a_t \quad (55)$$

B Operator Form

$$\phi(B)y_t = (1 - 1.5B + 0.6B^2)y_t = a_t \quad (56)$$

Stationarity

For stationarity, the roots of

$$\phi(B) = 1 - 1.5B + 0.6B^2 = 0 \quad (57)$$

must lie outside the unity circle, which implies that the parameters ϕ_1, ϕ_2 must lie in the triangular region

$$\begin{aligned} \phi_2 + \phi_1 &< 1 \\ \phi_2 - \phi_1 &< 1 \\ -1 &< \phi_2 < 1 \end{aligned} \quad (58)$$

Check

$$\begin{aligned} -0.6 + 1.5 &= 0.9 < 1 \\ -0.6 - 1.5 &= -2.1 < 1 \\ -1 &< -0.6 < 1 \end{aligned} \quad (59)$$

Therefore, the process is stationary

Invertibility

Pure AR models are always invertible (since they contain no MA terms).
Thus, the process is invertible.

ψ weights

Infinite moving average representation

$$y_t = \psi(B)a_t = \sum_{j=0}^{\infty} \psi_j a_{t-j} \quad (60)$$

where

$$\psi(B) = \frac{1}{\phi(B)} \quad (61)$$

The weights ψ are determined from the relation $\psi(B)\phi(B) = 1$ to satisfy

$$\psi_j = \phi_1 \psi_{j-1} + \phi_2 \psi_{j-2} + \cdots + \phi_p \psi_{j-p} \quad j > 0 \quad (62)$$

with $\psi_0 = 1, \psi_j = 0$ for $j < 0$, from which the weights ψ_j can be computed recur-

sively in terms of the ϕ_i .

$$\psi_0 = 1$$

$$\begin{aligned}\psi_1 &= \phi_1\psi_0 + \phi_2\psi_{-1} \\ &= (0.5)(1) + (-0.6)(0) \\ &= 0.5\end{aligned}$$

$$\begin{aligned}\psi_2 &= \phi_1\psi_1 + \phi_2\psi_0 \\ &= (0.5)(0.5) + (-0.6)(1) \\ &= -0.35\end{aligned}\tag{63}$$

$$\begin{aligned}\psi_3 &= \phi_1\psi_2 + \phi_2\psi_1 \\ &= (0.5)(-0.35) + (-0.6)(0.5) \\ &= -0.475\end{aligned}$$

π weights

$$y_t = 1.5y_{t-1} - 0.6y_{t-2} + a_t\tag{64}$$

$$\pi_0 = -1$$

$$\pi_1 = 1.5$$

$$\pi_2 = -0.6\tag{65}$$

$$\pi_3 = 0$$

Variance

The variance of AR(2) process is

$$\begin{aligned}\text{Var}[y_t] &= \frac{\sigma_a^2}{1 - \rho_1\phi_1 - \rho_2\phi_2} \\ &= \frac{1 - \phi_2}{1 + \phi_2} \frac{\sigma_a^2}{(1 - \phi_2)^2 - \phi_1^2}\end{aligned}\tag{66}$$

So we have

$$\text{Var}[y_t] = \frac{1 + 0.6}{1 - 0.6} \frac{1}{(1 + 0.6)^2 - (-1.5)^2}\tag{67}$$

4. ARMA(1,1)

$$y_t - y_{t-1} = a_t - 0.5a_{t-1} \quad (68)$$

B Operator Form

$$\begin{aligned} (1-B)y_t &= (1-0.5B)a_t \\ \phi(B)y_t &= \theta(B)a_t \end{aligned} \quad (69)$$

Stationarity

$$\phi(B) = 1 - B = 0 \quad (70)$$

The root of the AR characteristic equation does not lie outside the unity circle. Thus the process is not stationary.

Invertibility

$$\theta(B) = 1 - 0.5B = 0 \quad (71)$$

The root of the MA characteristic equation $B = 2$ lie outside the unity circle. Thus the process is invertible.

 ψ weights

Infinite moving average representation

$$y_t = \psi(B)a_t = \sum_{j=0}^{\infty} \psi_j a_{t-j} \quad (72)$$

$$\begin{aligned} y_t &= y_{t-1} + a_t - 0.5a_{t-1} \\ &= y_{t-2} + a_{t-1} - 0.5a_{t-2} + a_t - 0.5a_{t-1} \\ &= y_{t-3} + a_{t-2} - 0.5a_{t-3} + a_{t-1} - 0.5a_{t-2} + a_t - 0.5a_{t-1} \\ &\vdots \\ &= a_t + \cdots + 0.5a_{t-1} + 0.5a_{t-2} + \cdots + 0.5a_1 + 0.5a_0 + \cdots - 0.5a_{t-\infty} \end{aligned} \quad (73)$$

$$\begin{aligned} \psi_0 &= 1 \\ \psi_1 &= 0.5 \\ \psi_2 &= 0.5 \\ \psi_3 &= 0.5 \end{aligned} \quad (74)$$

 π weights

Infinite autoregressive representation

$$\pi(B)y_t = y_t - \sum_{j=1}^{\infty} \pi_j y_{t-j} = a_t \quad (75)$$

where

$$\pi(B) = \frac{\phi(B)}{\theta(B)} \quad (76)$$

The weights π are determined from the relation $\theta(B)\pi(B) = \phi(B)$ to satisfy

$$\pi_j = \theta_1 \pi_{j-1} + \theta_2 \pi_{j-2} + \cdots + \theta_q \pi_{j-q} + \phi_j \quad j > 0 \quad (77)$$

with $\pi_0 = -1$, $\pi_j = 0$ for $j < 0$, and $\phi_j = 0$ for $j > p$

$$\begin{aligned} \pi_0 &= -1 \\ \pi_1 &= \theta_1 \pi_0 + \phi_1 \\ &= (0.5)(-1) + (1) \\ &= 0.5 \\ \pi_2 &= \theta_1 \pi_1 + \phi_2 \\ &= (0.5)(0.5) + (0) \\ &= 0.25 \\ \pi_3 &= \theta_1 \pi_2 + \phi_3 \\ &= (0.5)(0.25) + (0) \\ &= 0.125 \end{aligned} \quad (78)$$

Variance

$$\begin{aligned} \text{Var}[y_t] &= \gamma_0 = \text{Var}[y_{t-1} + a_t - 0.5a_{t-1}] \\ &= \text{Var}[y_{t-1}] + 0.25\sigma_a^2 \end{aligned} \quad (79)$$

$$\begin{aligned} y_t &= y_{t-1} + a_t - 0.5a_{t-1} \\ &= y_{t-2} + a_{t-1} - 0.5a_{t-2} + a_t - 0.5a_{t-1} \\ &= y_{t-3} + a_{t-2} - 0.5a_{t-3} + a_{t-1} - 0.5a_{t-2} + a_t - 0.5a_{t-1} \\ &\vdots \end{aligned} \quad (80)$$

$$\begin{aligned} &= y_0 + a_t + 0.5a_{t-1} + 0.5a_{t-2} + \cdots + 0.5a_1 - 0.5a_0 \\ \text{Var}[y_t] &= \text{Var}[y_0 + a_t + 0.5a_{t-1} + 0.5a_{t-2} + \cdots + 0.5a_1 - 0.5a_0] \\ &= (1 + 0.25t)\sigma_a^2 + \text{Var}(y_0) \\ &= 1 + 0.25t \end{aligned} \quad (81)$$

$$\begin{aligned} y_t &= y_{t-1} + a_t - 0.5a_{t-1} \\ &= y_{t-2} + a_{t-1} - 0.5a_{t-2} + a_t - 0.5a_{t-1} \\ &= y_{t-3} + a_{t-2} - 0.5a_{t-3} + a_{t-1} - 0.5a_{t-2} + a_t - 0.5a_{t-1} \\ &\vdots \\ &= a_t + \cdots + 0.5a_{t-1} + 0.5a_{t-2} + \cdots + 0.5a_1 + 0.5a_0 + \cdots - 0.5a_{-\infty} \end{aligned} \quad (82)$$

$$\begin{aligned}\text{Var}[y_t] &= \text{Var}[a_t + \cdots + 0.5a_{t-1} + 0.5a_{t-2} + \cdots + 0.5a_1 + 0.5a_0 + \cdots - 0.5a_{-\infty}] \\ &= (1 + \sum_{i=1}^{\infty} 0.25)\sigma_a^2 \\ &= 1 + \sum_{i=1}^{\infty} 0.25 \\ &= 1 + (0.25)\infty\end{aligned}\tag{83}$$

■

Problem 5.

Consider y_t a stationary process. Show that if $\rho_1 < 0.5$, $(1-B)y_t$ has a larger variance than does y_t .

Solution.

$$\begin{aligned}\text{Var}[(1-B)y_t] &= \text{Var}[y_t - y_{t-1}] \\ &= \text{Var}[y_t] + \text{Var}[y_{t-1}] - 2\text{Cov}[y_t, y_{t-1}] \\ &= \gamma_0 + \gamma_0 - 2\gamma_1 \\ &= (2 - 2\rho_1)\gamma_0 \\ &= \underbrace{2(1 - \rho_1)}_{>1}\gamma_0 \\ &> \gamma_0\end{aligned}\tag{84}$$

We then conclude that $(1-B)y_t$ has a larger variance than does y_t . ■

Problem 6.

Consider an AR(1) process satisfying $y_t = \phi y_{t-1} + e_t$, where ϕ can be **any** number and e_t is a white noise process such that e_t is independent of the past y_{t-1}, y_{t-2}, \dots . Let y_0 be a random variable with mean μ_0 and variance σ_0^2 .

(a) For $t > 0$, show that

$$y_t = e_t + \phi^2 e_{t-2} + \phi^3 e_{t-3} + \dots + \phi^{t-1} e_1 + \phi^t y_0$$

(b) Show that $E[y_t] = \phi^t \mu_0$, for $t > 0$.

(c) Show that for $t > 0$, we have

$$\text{Var}[y_t] = \begin{cases} \frac{1-\phi^{2t}}{1-\phi^2} \sigma_e^2 + \phi^{2t} \sigma_0^2 & \phi \neq 1 \\ t \sigma_e^2 + \sigma_0^2 & \phi = 1. \end{cases}$$

(d) Assuming $\mu_0 = 0$, show that, we must have $\phi \neq 1$ to make y_t stationary.

(e) following (d) and supposing that $\mu_0 = 0$ and y_t is stationary, show that $\text{Var}[y_t] = \frac{\sigma_e^2}{1-\phi^2}$ and we must have $|\phi| < 1$.

Solution.

(a)

$$\begin{aligned} y_t &= \phi y_{t-1} + e_t \\ &= \phi(\phi y_{t-2} + e_{t-1}) + e_t \\ &= \phi^2 y_{t-2} + \phi e_{t-1} + e_t \\ &= \phi^2(\phi y_{t-3} + e_{t-2}) + \phi e_{t-1} + e_t \\ &= \phi^3 y_{t-3} + \phi^2 e_{t-2} + \phi e_{t-1} + e_t \\ &\vdots \\ &= \phi^t y_0 + \phi^{t-1} e_1 + \dots + \phi e_{t-1} + e_t \end{aligned} \tag{85}$$

as required.

(b)

$$\begin{aligned} E[y_t] &= E[\phi^t y_0 + \phi^{t-1} e_1 + \dots + \phi e_{t-1} + e_t] \\ &= E[\phi^t y_0] \\ &= \phi^t E[y_0] \\ &= \phi^t \mu_0 \end{aligned} \tag{86}$$

(c)

$$\begin{aligned}
\text{Var}[y_t] &= \text{Var}[\phi^t y_0 + \phi^{t-1} e_1 + \cdots + \phi e_{t-1} + e_t] \\
&= \phi^{2t} \sigma_0^2 + (\phi^2 + \phi^4 + \phi^6 + \cdots + \phi^{2(t-1)}) \sigma_e^2 \\
&= \phi^{2t} \sigma_0^2 + \sigma_e^2 \sum_{k=1}^{t-1} \phi^{2k} \\
&= \begin{cases} \phi^{2t} \sigma_0^2 + \frac{1-\phi^{2t}}{1-\phi^2} \sigma_e^2 & \phi \neq 1 \\ \sigma_0^2 + t \sigma_e^2 & \phi = 1. \end{cases} \tag{87}
\end{aligned}$$

Geometric Series

The sum of a n -term (finite) geometric series is given by:

$$S_n = \begin{cases} \frac{a_1(1-r^n)}{1-r} & r \neq 1 \\ a_1 n & r = 1 \end{cases}$$

with initial value $a = a_1$ and common ratio r .

(d) If $\phi = 1$, we have

$$\begin{aligned}
\text{Var}[y_t] &= \text{Var}[y_{t-1} + e_t] \\
&= \text{Var}[y_{t-1}] + \sigma_e^2
\end{aligned}$$

which is against the stationarity. Therefore, we must have $\phi \neq 1$ to make y_t stationary.

(e)

$$\begin{aligned}
\text{Var}[y_t] &= \text{Var}[\phi y_{t-1} + e_t] \\
&= \phi^2 \text{Var}[y_{t-1}] + \sigma_e^2
\end{aligned}$$

Due to the requirement of stationarity, we have

$$\begin{aligned}
\text{Var}[y_t] &= \phi^2 \text{Var}[y_t] + \sigma_e^2 \\
(1 - \phi^2) \text{Var}[y_t] &= \sigma_e^2 \\
\text{Var}[y_t] &= \frac{\sigma_e^2}{1 - \phi^2}
\end{aligned}$$

Since the variance $\text{Var}[y_t]$ must be positive, we must have $|\phi| < 1$.

■