Problem 1. Testing for completely random dropouts

Let P_{ij} denote the probability that the *i*-th unit drops out at time t_j , j = 1, ..., m.

Under the assumption of completely random dropouts, the probability P_{ij} may depend on time, treatment, or other explanatory variables, but cannot depend on the observed measurements $y_i = (y_{i1}, \ldots, y_{im_i})$.

Testing Method:

(a) Choose the score function $h_k(y_1, \dots, y_k)$ so that extreme values constitute evidence against completely random dropouts. A sensible choice is

$$h_k(y_1, \dots, y_k) = \sum_{j=1}^k \omega_j y_j.$$

(b) For each of k = 1, ..., (m - 1), define

$$R_k = \{i : m_i \ge k\},\$$
 $r_k = \{i : m_i = k\},\$

and compute the set of scores $h_{ik} = h_k(y_{i1...,y_{ik}})$ for $i \in R_k$.

(c) If $1 \le |r_k| \le |R_k|$, test the hypothesis that the r_k 's scores so defined are a random sample from the "populations" of R_k 's scores.

Remark:

- 1. The implicit assumption that the separated p-values are mutually independent is valid precisely because once a unit drops out, it never returns.
- 2. A natural test statistics is $\bar{h}_k = \frac{1}{|r_k|} \sum_{\{j \in r_k\}} h_{jk}$. Under the assumption of completely random dropouts,

$$\overline{h}_k \sim N\left(\overline{H}_k, \frac{|R_k| - |r_k|}{(|R_k| - 1)|r_k|} \sum_{\{j \in R_k\}} (h_{jk} - \overline{H}_k)^2 / |R_k|\right),$$

where

$$\overline{H}_k = \frac{1}{|R_k|} \sum_{\{j \in r_k\}} h_{jk}.$$

- When $|R_k|$ or $|r_k|$ is small, evaluate the randomization distribution of \overline{h}_k under the null hypothesis.
- Alternative method ...
- 3. The Final stage consists of analyzing the resulting set of p-values via
 - (a) Empirical distribution of the p-values
 - (b) Kolmogorov-Smirnov statistic $D_{+} = \sup |\hat{F}_{n}(p) p|$

Given a finite population of size N, with individual values $\{X_i\}_{i=1}^N$,

and a set of sample of size n, drawn from the population without replacement, with values $\{X_i\}_{i=1}^n$.

Let σ^2 be the population variance:

$$\sigma^2 = \mathbf{Var}[X_i] = \frac{1}{N} \sum_{i=1}^{N} (X_i - \mu),$$

where $\mu = \sum_{i=1}^{N} X_i$ is the population mean.

Let $\bar{X} = \frac{1}{n}S_n = \frac{1}{n}\sum_{i=1}^n X_i$ be the sample mean based on the sample set.

Since every pair (X_i, X_j) for $i \neq j$ has the same joint distribution, we have

$$\mathbf{Var}[S_n] = \sum_{i=1}^n \sum_{j=1}^n \mathbf{Cov}[X_i, X_j],$$

where

$$\mathbf{Cov}[X_i, X_j] = \begin{cases} \sigma^2 & i = j \\ c & i \neq j \end{cases}.$$

Thus,

$$\mathbf{Var}[S_n] = n\sigma^2 + n(n-1)c.$$

which applies to the case n = N as well. Notice that S_N is a constant (equal to the sum of all N values in the population). It follows that

$$0 = \mathbf{Var}[S_N] = N\sigma^2 + N(N-1)c.$$

Solve the equation above for

$$c = -\frac{\sigma^2}{N-1}.$$

Hence,

$$\mathbf{Var}[S_n] = n\sigma^2 \left(1 - \frac{n-1}{N-1}\right) = \frac{N-n}{N-1} \cdot n\sigma^2$$

and

$$\mathbf{Var}[\bar{X}] = \frac{N-n}{N-1} \cdot \frac{\sigma^2}{n}.$$

The factor $\frac{N-n}{N-1}$ is the Finite Population Correction Factor (FPC).