## Problem 1. Testing for completely random dropouts

Let  $P_{ij}$  denote the probability that the i-th unit drops out at time  $t_j$ , j = 1, ..., m. Under the assumption of completely random dropouts, the probability  $P_{ij}$  may depend on time, treatment, or other explanatory variables, but cannot depend on the observed measurements

 $y_i = (y_{i1}, \ldots, y_{i m_i}).$ 

## Testing Method:

(a) Choose the score function  $h_k(y_1, \dots, y_k)$  so that extreme values constitute evidence against completely random dropouts. A sensible choice is

$$h_k(y_1, \dots, y_k) = \sum_{j=1}^k \omega_j y_j.$$

(b) For each of k = 1, ..., (m - 1), define

$$R_k = \{i: m_i \ge k\},$$
 
$$r_k = \{i: m_i = k\},$$

and compute the set of scores  $h_{ik} = h_k(y_{i1...,ik})$  for  $i \in R_k$ .

(c) If  $1 \le |r_k| \le |R_k|$ , test the hypothesis that the  $r_k$ 's scores so defined are a random sample from the "populations" of  $R_k$ 's scores.

Remark:

- 1. The implicit assumption that the separated p-values are mutually independent is valid precisely because once a unit drops out, it never returns.
- 2. A natural test statistics is  $\bar{h}_k = \frac{1}{|r_k|} \sum_{\{j \in r_k\}} h_{jk}$ . Under the assumption of completely random dropouts,

$$\bar{h}_k \sim N \left( \frac{|R_k| - |r_k|}{|r_k|(|R_k| - 1)} \right) \sum_{\{j \in R_k\}} (h_{jk} - \overline{H}_k)^2,$$

where

$$\overline{H}_k = \frac{1}{|R_k|} \sum_{\{j \in r_k\}} h_{jk}.$$

- When  $|R_k|$  or  $|r_k|$  is small, evaluate the randomization distribution of  $\overline{h}_k$  under the null hypothesis.
- Alternative method ...
- 3. The Final stage consists of analyzing the resulting set of p-values via
  - (a) Empirical distribution of the p-values
  - (b) Kolmogorov-Smirnov statistic  $D_{+} = \sup |\hat{F}_{n}(p) p|$

Suppose that  $\sigma^2$  is the population variance. This implies: If the random variable X is the result of

a single draw from the population, then  $\operatorname{Var}(X) = \sigma^2$ . Now consider drawing a sample of n items  $X_1, \ldots, X_n$  without replacement from the population. Since every pair  $(X_i, X_j)$  for  $i \neq j$  has the same joint distribution, the variance of the sum  $S_n := X_1 + \ldots + X_n$  is

$$Var(S_n) = nVar(X_1) + (n^2 - n) Cov(X_1, X_2) = n\sigma^2 + n(n-1)c$$
 (1),

where we write c for the covariance between the results of two distinct draws. Formula (1) applies in the case n = N as well, with the extra bonus that  $S_N$  is a constant (equal to the sum of all N values in the population). It follows that

$$0 = Var(S_N) = N\sigma^2 + N(N-1)c \quad (2).$$

Solve equation (2) for  $c=-\frac{\sigma^2}{N-1}$  (3) and plug back into (1) to obtain

$$Var(S_n) = n\sigma^2 \left( 1 - \frac{n-1}{N-1} \right) = \frac{N-n}{N-1} \cdot n\sigma^2 \quad (4)$$

and

$$\operatorname{Var}(\bar{X}_n) = \frac{N-n}{N-1} \cdot \frac{\sigma^2}{n} \quad (5).$$

Notice the difference between formulas (4) and (5) and the corresponding formulas for sampling with replacement is a factor  $\frac{N-n}{N-1}$ , which is the famous correction factor for sampling without replacement.