



Chapter 2 : Point Estimation

- Method of Moments (MOM)

is a method which is estimating population moment by sample moment.

Estimation of moment

$$\mu = E[X] \rightarrow \tilde{\mu}_n = \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\mu = E[X^k] \rightarrow \tilde{\mu}_n = \overline{X^k}_n = \frac{1}{n} \sum_{i=1}^n X_i^k$$

Estimation of central moment

$$\sigma^2 = \text{Var}[X] = E[(X - E[X])^2] \rightarrow \tilde{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

$$\sigma^2 = \text{Var}[X] = E[X^2] - E[X]^2 \rightarrow \tilde{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \left(\frac{1}{n} \sum_{i=1}^n X_i \right)^2$$

$$\mu_k = E[(X - E[X])^k] \rightarrow \tilde{\mu}_k = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^k$$

Ex. Suppose X_i 's be a independent sample random variables with PDF given by

$$f(x|\beta) = \frac{\beta}{x^{\beta+1}} ; x > 1$$

What is the MOM estimator for β

Soln



EX. Suppose $X_i's \stackrel{iid}{\sim} \text{Beta}(\alpha = \theta, \beta = 1)$. What is MOM estimator for θ ?

Midterm 63

Solⁿ



Ex. Suppose $X_i \sim \text{Bin}(n=i, p)$ for $i = 1, \dots, r$ and X_i 's are independent R.V.

; moreover, $\sum_{i=1}^r X_i \sim \text{Bin}\left(n = \frac{r(r+1)}{2}, p\right)$. Find MOM estimator for p

Final 66

Solⁿ



- Maximum Likelihood Estimators (MLE)

is a method that provide possible parameters which maximize our probability based on a sample

Suppose X_1, \dots, X_n are iid sample from a population with PDF or PMF

$$L(\theta|x_i's) = L(\theta_1, \dots, \theta_k|x_i's) = \prod_{i=1}^n f(x_i|\theta_1, \dots, \theta_k)$$

$$l(\theta|x_i's) = \log(L(\theta|x_i's))$$

Step of finding MLE for parameters

$$1. L(\theta|x_i's) = \prod_{i=1}^n f(x_i|\theta_1, \dots, \theta_k)$$

$$2. l(\theta|x_i's) = \log(L(\theta|x_i's))$$

$$3. \frac{\partial}{\partial x} l(\theta|x_i's)$$

$$4. \text{ set } \frac{\partial}{\partial x} l(\theta|x_i's) = 0$$

$$5. \hat{\theta}_n \text{ (MLE for } \theta \text{)}$$

Note

$$\log = \log_e \neq \log_{10}$$

Ex. Suppose $X_i's \stackrel{iid}{\sim} \text{Bin}(m, p)$. What is the MLE estimator for p ?

Soln





Ex. Suppose $X_i's \stackrel{iid}{\sim} Geo(p)$. What is the MLE estimator for p ?

Solⁿ



Ex. Suppose $X_i's \stackrel{iid}{\sim} \text{Uniform}(0, \theta)$. What is the MLE estimator for θ ?

Solⁿ



Ex. Suppose $X_i's \stackrel{iid}{\sim} \text{Uniform}(0, \theta)$. What is the MLE estimator for θ ?

Solⁿ



Ex. Suppose $X_i's \stackrel{iid}{\sim} \text{Uniform}(2\theta, 0)$ for $\theta < 0$. What is the MLE estimator for θ ?

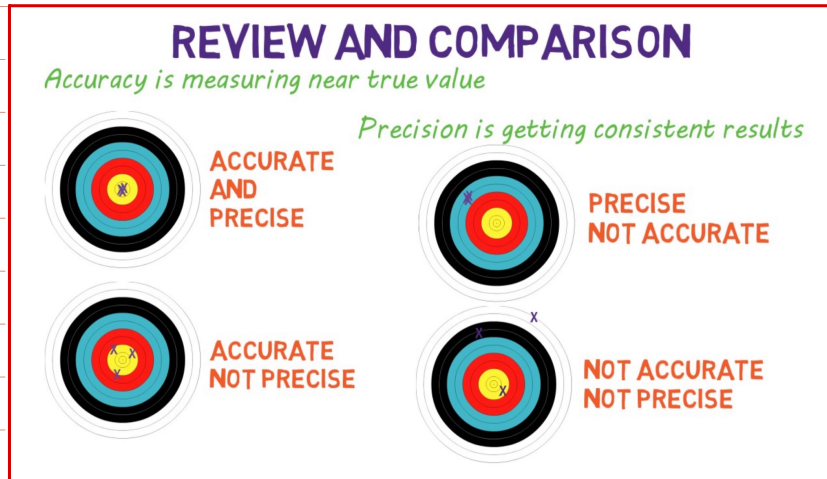
Midterm 66

Soln



- Methods of Evaluating Estimators

Accuracy vs Precision of Estimators



The bias of a point estimator W of parameter θ

$$Bias_{\theta}(W) = E[W] - \theta$$

- Bias = 0 : unbiased (W is unbiased estimator.)
- Bias > 0 : overestimate
- Bias < 0 : underestimate

The mean square error (MSE) of a point estimator W of parameter θ

$$MSE_{\theta}(W) = E[(W - \theta)^2] = Var[W] + [Bias_{\theta}(W)]^2$$

- MSE = Var if and only bias = 0

$\hat{\theta}$ is said to be more precise than $\tilde{\theta}$ if $Var[\hat{\theta}] < Var[\tilde{\theta}]$

$\hat{\theta}$ is said to be more accurate than $\tilde{\theta}$ if $|Bias_{\theta}(\hat{\theta})| < |Bias_{\theta}(\tilde{\theta})|$



EX. Suppose $X \sim \text{Exp}(\theta)$ and $Y \sim \text{Gamma}(\alpha = 2, \beta = \theta)$ which $X \perp Y$.

If $\hat{\theta}_1 = X, \hat{\theta}_2 = \frac{Y}{2}, \hat{\theta}_3 = \frac{X+Y}{3}$ and $\hat{\theta}_4 = \frac{X+2Y}{6}$ are an estimator for θ

1. Show that $\hat{\theta}_1, \hat{\theta}_2, \hat{\theta}_3$ are unbiased estimators.

2. Show that $\hat{\theta}_4$ is biased estimators.

3. Find $MSE_{\theta}(\hat{\theta}_i)$ for $i = 1, 2, 3$ and 4

Soln



True or False ?

Ex. Suppose $\hat{\theta}$ and $\tilde{\theta}$ are an estimator for parameter θ . Given that $\text{Var}[\hat{\theta}] > \text{Var}[\tilde{\theta}]$

and $\text{Bias}_{\theta}(\hat{\theta}) < \text{Bias}_{\theta}(\tilde{\theta})$, we can show that $\hat{\theta}$ is better estimator than $\tilde{\theta}$.

Solⁿ



Ex. Suppose X_i 's $\overset{iid}{\sim} \text{Uniform}(2\theta, 0)$ for $\theta < 0$. $\hat{\theta} = X_{(1)}/2$

Midterm 66

Check whether MLE estimator is a unbiased estimator or not.

Solⁿ



Relative efficiency

Given two unbiased estimators, $\hat{\theta}_1$ and $\hat{\theta}_2$,

$$RE(\hat{\theta}_1, \hat{\theta}_2) = \frac{Var[\hat{\theta}_2]}{Var[\hat{\theta}_1]}$$

$$RE(\hat{\theta}_1, \hat{\theta}_2) \begin{cases} < 1 \rightarrow \hat{\theta}_1 \text{ is less efficient than } \hat{\theta}_2 \\ = 1 \rightarrow \hat{\theta}_1 \text{ is as efficient as } \hat{\theta}_2 \\ > 1 \rightarrow \hat{\theta}_1 \text{ is more efficient than } \hat{\theta}_2 \end{cases}$$

Ex. Suppose $X_i's \stackrel{iid}{\sim} \text{Gamma}(\alpha = 2, \beta = \theta)$ for $\theta > 0$. $\tilde{\theta}_n$ and $\hat{\theta}_n$ are an estimator from MOM and MLE, respectively. Which one is more efficient?

Midterm 62

Soln





Asymptotic Evaluation

$\hat{\theta}_n$ is said to be a consistent estimator of θ if

$$\hat{\theta}_n \xrightarrow{P} \theta \text{ as } n \rightarrow \infty \text{ } (\bar{X}_n \text{ use WLLN})$$

or

$$\lim_{n \rightarrow \infty} \text{Bias}_{\theta}(\hat{\theta}_n) = 0 \text{ and } \lim_{n \rightarrow \infty} \text{Var}[\hat{\theta}_n] = 0$$

Ex. Suppose X_1, \dots, X_n be independent sample variable with given PDF,

$$f(x_i|\theta) = \begin{cases} \frac{2x_i}{\theta^2} & ; 0 < x_i \leq \theta \\ 0 & ; \text{otw.} \end{cases}$$

Show that MOM estimator, $\tilde{\theta}_n = \frac{3}{2} \bar{x}_n$, is a consistent estimator



Ex. Suppose $X_i \sim \text{Bin}(n = i, p)$ for $i = 1, \dots, r$ and X_i 's are independent R.V.

; moreover, $\sum_{i=1}^r X_i \sim \text{Bin}\left(n = \frac{r(r+1)}{2}, p\right)$. MOM estimator for p is $\frac{2 \sum_{i=1}^r X_i}{r(r+1)}$

Final 66

Show that MOM estimator is a consistent estimator.



Asymptotic Relative Efficiency

Given two consistent estimators, $\hat{\theta}_1$ and $\hat{\theta}_2$ of a parameter θ

$$ARE(\hat{\theta}_1, \hat{\theta}_2) = \lim_{n \rightarrow \infty} \left(\frac{\text{Var}(\hat{\theta}_2)}{\text{Var}(\hat{\theta}_1)} \right)$$

Ex. Suppose X_i 's $\stackrel{iid}{\sim} \text{Gamma}(\alpha = 2, \beta = \theta)$ for $\theta > 0$. $\tilde{\theta}_n$ and $\hat{\theta}_n$ are an estimator from MOM and MLE, respectively. Find $ARE(\tilde{\theta}_n, \hat{\theta}_n)$



Method of Finding Estimation II

Bayesian Estimation

Bayes Theorem

$$P(A|B) = \frac{P(AB)}{P(B)} \rightarrow P(AB) = P(A|B)P(B)$$

*In Bayesian Estimation, we treat parameter θ as a random variable;
therefore, $f(x|\theta)$ is a conditional probability.*

Joint Distribution Function

$$f(x, \theta) = f(x|\theta) \times \underline{f(\theta)}$$

└ Prior Distribution Function

Marginal of X

$$f(x) = \int_{-\infty}^{\infty} f(x, \theta) d\theta = \int_{-\infty}^{\infty} f(x|\theta)f(\theta) d\theta$$

Posterior Distribution Function

$$f(\theta|x) = \frac{f(x, \theta)}{f(x)} = \frac{f(x|\theta)f(\theta)}{f(x)}$$



Ex. Let $X \sim N\left(\mu = 0, \sigma^2 = \frac{1}{\lambda}\right)$ and $Y \sim \text{Exp}\left(\theta = \frac{1}{\lambda}\right)$ when $X \perp Y$.

Midterm 66

We assume that the prior distribution of λ is $\text{Gamma}(\alpha = 2, \beta = 3)$.

Find $f(x, y)$, $f(\lambda|x, y)$ and $E[\lambda|X, Y]$

Solⁿ



Ex. Let $X \sim \text{Geo}(p)$, $Y \sim \text{NegBin}(r = 5, p)$ and $Z \sim \text{Bin}(n = 2, p)$ when $X \perp Y \perp Z$.

We assume that the prior distribution of p is $\text{Beta}(\alpha = 2, \beta = 3)$.

Find $f(x, y, z)$, $f(p|x, y, z)$ and $E[p|X, Y, Z]$

Solⁿ





- Sufficiency

Sufficient Statistic

$T(\tilde{X})$ is said to be sufficient for θ if

$$P(X'_i s | T(\tilde{X}), \theta) = P(\tilde{X} | T(\tilde{X}))$$

Neyman's Factorization Theorem

$T(\tilde{X})$ is said to be sufficient for θ if

$$f(x'_i s | \theta) = g(T(\tilde{x}) | \theta) h(\tilde{x})$$

Ex. Suppose $X'_i s \stackrel{iid}{\sim} \text{Ber}(p)$. Show that $\sum_{i=1}^n X_i$ is a sufficient statistic for p .

Solⁿ



Ex. Suppose X_i 's $\stackrel{iid}{\sim} \text{Gamma}(\alpha = 2, \beta = \theta)$ for $\theta > 0$.

Midterm 62

Show that $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ is a sufficient statistic for θ .

Soln

Ex. Suppose X_i 's $\stackrel{iid}{\sim} \text{Uniform}(\theta, 0)$ for $\theta < 0$

Show that $\min\{X_1, X_2, \dots, X_n\}$ is a sufficient statistic for θ .

Soln




Exponential families

PDF or PMF is called an exponential family if

$$f(x|\theta) = h(x)c(\theta) \exp \left\{ \sum_{i=1}^k w_i(\theta) t_i(x) \right\}$$

where $h(x), c(\theta) > 0$

Note 

Uniform Dist.
is not an expo. fam.

Ex.

Suppose X_i 's $\stackrel{iid}{\sim} \text{Beta}(\alpha = \theta, \beta = 1)$ for $\theta > 0$.

Midterm 63

Show that $\text{Beta}(\alpha = \theta, \beta = 1)$ is an exponential family.

Solⁿ



Ex. Show that $\text{Bin}(n, p)$ is an exponential family.

Solⁿ