# Conditional Densities

A number of machine learning algorithms can be derived by using conditional exponential families of distribution (Section 2.3). Assume that the training set  $\{(x_1, y_1), \ldots, (x_m, y_m)\}$  was drawn iid from some underlying distribution. Using Bayes rule (1.15) one can write the likelihood

$$p(\theta|X,Y) \propto p(\theta)p(Y|X,\theta) = p(\theta) \prod_{i=1}^{m} p(y_i|x_i,\theta), \tag{6.1}$$

and hence the negative log-likelihood

$$-\log p(\theta|X,Y) = -\sum_{i=1}^{m} \log p(y_i|x_i,\theta) - \log p(\theta) + \text{const.}$$
 (6.2)

Because we do not have any prior knowledge about the data, we choose a zero mean unit variance isotropic normal distribution for  $p(\theta)$ . This yields

$$-\log p(\theta|X,Y) = \frac{1}{2} \|\theta\|^2 - \sum_{i=1}^{m} \log p(y_i|x_i,\theta) + \text{const.}$$
 (6.3)

Finally, if we assume a conditional exponential family model for  $p(y|x,\theta)$ , that is,

$$p(y|x,\theta) = \exp\left(\langle \phi(x,y), \theta \rangle - g(\theta|x)\right),\tag{6.4}$$

then

$$-\log p(\theta|X,Y) = \frac{1}{2} \|\theta\|^2 + \sum_{i=1}^{m} g(\theta|x_i) - \langle \phi(x_i, y_i), \theta \rangle + \text{const.}$$
 (6.5)

where

$$g(\theta|x) = \log \sum_{y \in \mathcal{Y}} \exp\left(\langle \phi(x, y), \theta \rangle\right),$$
 (6.6)

is the log-partition function. Clearly, (6.5) is a smooth convex objective function, and algorithms for unconstrained minimization from Chapter 5

can be used to obtain the maximum aposteriori (MAP) estimate for  $\theta$ . Given the optimal  $\theta$ , the class label at any given x can be predicted using

$$y^* = \operatorname*{argmax}_{y} p(y|x, \theta). \tag{6.7}$$

In this chapter we will discuss a number of these algorithms that can be derived by specializing the above setup. Our discussion unifies seemingly disparate algorithms, which are often discussed separately in literature.

#### 6.1 Logistic Regression

We begin with the simplest case namely binary classification<sup>1</sup>. The key observation here is that the labels  $y \in \{\pm 1\}$  and hence

$$g(\theta|x) = \log\left(\exp\left(\langle \phi(x, +1), \theta \rangle\right) + \exp\left(\langle \phi(x, -1), \theta \rangle\right)\right). \tag{6.8}$$

Define  $\hat{\phi}(x) := \phi(x,+1) - \phi(x,-1)$ . Plugging (6.8) into (6.4), using the definition of  $\hat{\phi}$  and rearranging

$$p(y = +1|x, \theta) = \frac{1}{1 + \exp\left(\left\langle -\hat{\phi}(x), \theta\right\rangle\right)} \text{ and}$$
$$p(y = -1|x, \theta) = \frac{1}{1 + \exp\left(\left\langle \hat{\phi}(x), \theta\right\rangle\right)},$$

or more compactly

$$p(y|x,\theta) = \frac{1}{1 + \exp\left(\left\langle -y\hat{\phi}(x), \theta\right\rangle\right)}.$$

Since  $p(y|x,\theta)$  is a logistic function, hence the name logistic regression. The classification rule (6.7) in this case specializes as follows: predict +1 whenever  $p(y=+1|x,\theta) \ge p(y=-1|x,\theta)$  otherwise predict -1. However

$$\log \frac{p(y=+1|x,\theta)}{p(y=-1|x,\theta)} = \left\langle \hat{\phi}(x), \theta \right\rangle,$$

therefore one can equivalently use sign  $\left(\left\langle \hat{\phi}(x), \theta \right\rangle\right)$  as our prediction function. Next we turn our attention to deriving the log-likelihood. After some simple algebraic manipulation one can write

$$\begin{split} g(\theta|x) - \langle \phi(x,+1), \theta \rangle &= \log \left( 1 + \exp \left( \left\langle \hat{\phi}(x), \theta \right\rangle \right) \right) - \left\langle \hat{\phi}(x), \theta \right\rangle \text{ and } \\ g(\theta|x) - \langle \phi(x,-1), \theta \rangle &= \log \left( 1 + \exp \left( \left\langle -\hat{\phi}(x), \theta \right\rangle \right) \right) + \left\langle \hat{\phi}(x), \theta \right\rangle. \end{split}$$

 $<sup>^{1}</sup>$  The name logistic regression is a misnomer!

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The log-likelihood (6.5) can now be written compactly by combining the above two equations as

$$\frac{1}{2} \|\theta\|^2 + \sum_{i=1}^m \log \left(1 + \exp\left(\left\langle y_i \hat{\phi}(x_i), \theta \right\rangle\right)\right) - y_i \left\langle \hat{\phi}(x_i), \theta \right\rangle + \text{const.}$$

To minimize the above objective function we first compute the gradient.

$$\nabla J(\theta) = \theta + \sum_{i=1}^{m} \frac{\exp\left(\left\langle y_i \hat{\phi}(x_i), \theta \right\rangle\right)}{1 + \exp\left(\left\langle y_i \hat{\phi}(x_i), \theta \right\rangle\right)} y_i \hat{\phi}(x_i) - y_i \hat{\phi}(x_i)$$
$$= \theta + \sum_{i=1}^{m} (p(y_i | x_i, \theta) - 1) y_i \hat{\phi}(x_i).$$

Notice that the second term of the gradient vanishes whenever  $p(y_i|x_i,\theta) = 1$ . Therefore, one way to interpret logistic regression is to view it as a method to maximize  $p(y_i|x_i,\theta)$  for each point  $(x_i,y_i)$  in the training set. Since the objective function of logistic regression is twice differentiable one can also compute its Hessian

$$\nabla^2 J(\theta) = -\sum_{i=1}^m p(y_i|x_i,\theta)(1 - p(y_i|x_i,\theta))\hat{\phi}(x_i)\hat{\phi}(x_i)^{\top},$$

where we used  $y_i^2 = 1$ . The Hessian can be used in the Newton method (Section 5.2.6) to obtain the optimal parameter  $\theta$ .

## 6.2 Regression

#### 6.2.1 Conditionally Normal Models

fixed variance

#### 6.2.2 Posterior Distribution

integrating out vs. Laplace approximation, efficient estimation (sparse greedy)

## 6.2.3 Heteroscedastic Estimation

explain that we have two parameters. not too many details (do that as an assignment).