

Lec 15

Conditioning by a Random Variable

Conditional PMF

For any event $[Y = y]$ such that $[P_Y[y] > 0]$, the conditional PMF of $[X]$ given $Y = y$ is

- $P_{X|Y}[x|y] = P[X = x | Y = y]$
- $$\begin{cases} P_{X,Y}[x,y] = P_{X|Y}[x|y] \cdot P_Y[y] = P_{Y|X}[y|x] \cdot P_X[x] \\ P_{X|Y}[x|y] = \frac{P_{X,Y}[x,y]}{P_Y[y]} \end{cases}$$

Conditional PDF

For y such that $f_Y(y) > 0$, the conditional PDF of X given $(Y = y)$ is

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

$$f_{X,Y}(x,y) = f_{X|Y}(x|y) \cdot f_Y(y) = f_{Y|X}(y|x) \cdot f_X(x)$$

Ex. 4.19 Random variables X and Y have joint PDF

$$f_{X,Y}(x,y) = \begin{cases} 2 & 0 \leq y \leq x \leq 1 \\ 0 & \text{o.w} \end{cases}$$

Find the conditional pdf $f_{Y|X}(y|x)$ and $f_{X|Y}(x|y)$.

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

$$\text{a. } \begin{cases} \Rightarrow f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy \\ \quad = \int_0^x 2 dy \\ \quad = 2x \quad \quad \quad 0 \leq x \leq 1 \end{cases}$$

$$\text{b. } \begin{cases} f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)} \\ \quad = \frac{2}{2x} = \frac{1}{x} \quad \quad \quad 0 \leq y \leq x \end{cases}$$

$$\text{c. } \begin{cases} \Rightarrow f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx \\ \quad = \int_y^1 2 dx \\ \quad = 2(1-y) \quad \quad \quad 0 \leq y \leq 1 \end{cases}$$

$$\text{d. } \begin{cases} f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} \\ \quad = \frac{2}{2(1-y)} = \frac{1}{(1-y)} \quad \quad \quad y \leq x \leq 1 \end{cases}$$

Conditional Expected Value of a Function

For continuous r.v $[X \text{ and } Y]$ and any y such that $f_Y(y) > 0$, the conditional expected value of $g(X, Y)$ given $Y = y$ is

$$E(g(X, Y) | Y = y) = \int_{-\infty}^{\infty} g(x, y) f_{X|Y}(x|y) dx$$

$$E(X | Y = y) = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

The conditional expected value $E(X | Y)$ is a function of random variable Y s.t. if $Y = y$ then $E(X | Y) = E(X | Y = y)$

Ex 4.20) The conditional PDF of X given Y is

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)} = \begin{cases} \frac{1}{1-y} & y \leq x \leq 1 \\ 0 & \text{o.w} \end{cases}$$

Find the conditional expected values $E(X | Y = y)$ and $E(X | Y)$.

$$E(X | Y = y) = \begin{cases} \int_{-\infty}^{\infty} x \cdot f_{X|Y}(x|y) dx \\ = \int_y^1 x \cdot \frac{1}{1-y} dx \\ = \frac{x^2}{2(1-y)} \Big|_{x=y}^{x=1} \\ = \frac{1-y^2}{2(1-y)} = \frac{1+y}{2} \end{cases}$$

$$E(X | Y) = \frac{1+Y}{2}$$

Quiz 4.9) A) The probability model for random variable A is

$$P_A[a] = \begin{cases} 0.4 & a = 0 \\ 0.6 & a = 2 \\ 0 & o.w \end{cases}$$

The conditional probability model for random variable B given A is

$$P_{B|A}[b|0] = \begin{cases} 0.8 & b = 0 \\ 0.2 & b = 1 \\ 0 & o.w \end{cases} \quad P_{B|A}[b|2] = \begin{cases} 0.5 & b = 0 \\ 0.5 & b = 1 \\ 0 & o.w \end{cases}$$

- 1) What is the probability model for $[A \text{ and } B]$? Write the joint PMF $P_{A,B}[a,b]$ as a table.
- 2) The conditional expected value $E[B|A=2]$
- 3) If $[B=0]$, what is the conditional PMF $P_{A|B}[a|0]$
- 4) If $[B=0]$, what is the conditional variance $Var[A|B=0]$

1)

$P_{A,B}[a,b]$	$b = 0$	$b = 1$
$a = 0$	$P[A=0, B=0] = P_{B A}[0 0] P_A[a] = 0.8 \cdot 0.4$	$0.2 \cdot 0.4$
$a = 2$	$0.5 \cdot 0.6$	$0.5 \cdot 0.6$

$$2) E[B|A=2] = \sum_{b=0}^1 b \cdot P_{B|A}[b|2] = 0 \cdot (0.5) + 1 \cdot (0.5) = 0.5$$

$$3) P_{A|B}[a|0] = \frac{P_{A,B}[a,0]}{P_B[0]} = \begin{cases} \frac{0.8 \cdot 0.4}{0.32 + 0.3} & a = 0 \\ \frac{0.5 \cdot 0.6}{0.32 + 0.3} & a = 2 \\ 0 & o.w \end{cases}$$

$$4) \left\{ \begin{array}{l} Var[A|B=0] = E[A^2|B=0] - (E[A|B=0])^2 \\ E[A^2|B=0] = \sum_{a=0,2} a^2 P_{A|B}[a|0] = 0^2 \cdot \frac{0.32}{0.62} + 2^2 \cdot \frac{0.30}{0.62} = \frac{1.2}{0.62} \\ E[A|B=0] = \sum_{a=0,2} a \cdot P_{A|B}[a|0] = 0 \cdot \frac{0.32}{0.62} + 2 \cdot \frac{0.30}{0.62} = \frac{0.6}{0.62} \\ Var[A|B=0] = E[A^2|B=0] - (E[A|B=0])^2 = \frac{1.2}{0.62} - \left(\frac{0.6}{0.62}\right)^2 \end{array} \right.$$

B) The PDF of random variable X and the conditional PDF of random variable Y given X are

$$f_X(x) = \begin{cases} 3x^2 & 0 \leq x \leq 1 \\ 0 & \text{o.w} \end{cases} \quad f_{Y|X}(y|x) = \begin{cases} \frac{2y}{x^2} & 0 \leq y \leq x, \quad 0 \leq x \leq 1 \\ 0 & \text{o.w} \end{cases}$$

1) Find $f_{X,Y}(x, y)$

2) If $X = \frac{1}{2}$ find the conditional PDF $f_{Y|X}(y|\frac{1}{2})$

3) If $Y = \frac{1}{2}$ find the conditional PDF $f_{X|Y}(x|\frac{1}{2})$

4) If $Y = \frac{1}{2}$ find the conditional variance $\text{Var}\left(X \mid Y = \frac{1}{2}\right)$

$$1) f_{X,Y}(x, y) = f_{Y|X}(y|x) f_X(x) = \frac{2y}{x^2} \cdot 3x^2 = \begin{cases} 6y & 0 \leq y \leq x, \quad 0 \leq x \leq 1 \\ 0 & \text{o.w} \end{cases}$$

$$2) f_{Y|X}(y|\frac{1}{2}) = \begin{cases} \frac{2y}{x^2} = \frac{2y}{(\frac{1}{2})^2} = 8y & 0 \leq y \leq \frac{1}{2} \\ 0 & \text{o.w} \end{cases}$$

$$3) \left\{ \begin{array}{l} f_{X|Y}(x|\frac{1}{2}) = \frac{f_{X,Y}(x, \frac{1}{2})}{f_Y(\frac{1}{2})} \\ f_Y(\frac{1}{2}) = \int_{-\infty}^{\infty} f_{X,Y}(x, \frac{1}{2}) dx = \int_{\frac{1}{2}}^1 6 \cdot \frac{1}{2} dx = \frac{3}{2} \\ \text{for } \frac{1}{2} \leq x \leq 1 \\ f_{X|Y}(x|\frac{1}{2}) = \frac{f_{X,Y}(x, \frac{1}{2})}{f_Y(\frac{1}{2})} = \frac{6 \cdot \frac{1}{2}}{\frac{3}{2}} = 2 \end{array} \right.$$

4) If $Y = \frac{1}{2}$, the conditional PDF of X is uniform for $\frac{1}{2} \leq x \leq 1$. As we saw from theorem 3.6 of textbook page 114, the variance is $\left[\frac{1}{12}(b-a)^2 \right]$.

$$\frac{(b-a)^2}{12} = \frac{(1-\frac{1}{2})^2}{12} = \frac{1}{48}$$

4.10 Independent Random Variables

In page 21 in Ch1, it is defined that events $[A \text{ and } B]$ are independent **if and only if**

$$P[AB] = P[A] \cdot P[B].$$

And when events $[A \text{ and } B]$ have nonzero probabilities, the following formulas are equivalent to the definitions of independent events:

$$\begin{cases} P[A|B] = P[A] \\ P[B|A] = P[B] \end{cases}$$

The above theorem can be applied to the random variable case with PMF and PDF. Random variables $[A \text{ and } B]$ are independent if and only if

Discrete: $P_{X,Y}[x, y] = P_X[x] P_Y[y]$

Continuous: $f_{X,Y}(x, y) = f_X(x) f_Y(y)$

As we can see from the independent definition of conditional probability theorem, it also implies that if X and Y are independent discrete/continuous random variables, then

$$\begin{cases} P_{X|Y}[x|y] = P_X[x] \\ P_{Y|X}[y|x] = P_Y[y] \end{cases} \quad \begin{cases} f_{X|Y}(x|y) = f_X(x) \\ f_{Y|X}(y|x) = f_Y(y) \end{cases}$$

Ex 4.23) Are $[X \text{ and } Y]$ independent?

$$f_{X,Y}(x,y) = \begin{cases} 4xy & 0 \leq x \leq 1, \ 0 \leq y \leq 1 \\ 0 & o.w \end{cases}$$

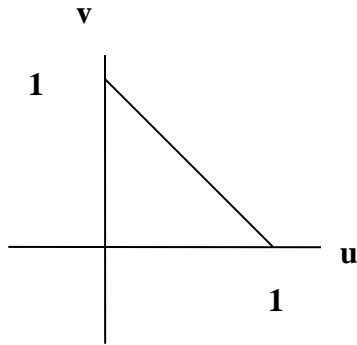
If it is true $[f_{X,Y}(x,y) = f_X(x)f_Y(y)]$, then it is independent. So we need to find out the marginal probability of $[X \text{ and } Y]$

$$f_X(x) = \begin{cases} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy \\ = \int_0^1 4xy dy \\ = \frac{4}{2} xy^2 \Big|_0^1 \\ = \begin{cases} 2x & 0 \leq x \leq 1 \\ 0 & o.w \end{cases} \end{cases} \quad f_Y(y) = \begin{cases} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx \\ = \int_0^1 4xy dx \\ = \frac{4}{2} x^2 y \Big|_0^1 \\ = \begin{cases} 2y & 0 \leq y \leq 1 \\ 0 & o.w \end{cases} \end{cases}$$

$$\begin{cases} f_X(x) \cdot f_Y(y) = f_{X,Y}(x,y) \\ 2x \cdot 2y = 4xy \end{cases} \quad \text{so it is independent}$$

Ex 4.24) Are U and V independent?

$$f_{U,V}(u,v) = \begin{cases} 24uv & 0 \leq u, 0 \leq v, u+v \leq 1 \\ 0 & \text{o.w} \end{cases}$$



We need to test $f_{U,V}(u,v) = f_U(u) f_V(v)$

$$f_U(u) = \begin{cases} \int_{-\infty}^{\infty} f_{U,V}(u,v) dv \\ = \int_0^{1-u} 24uv \cdot dv \\ = \frac{24}{2} uv^2 \Big|_0^{1-u} \\ = \begin{cases} 12u(1-u)^2 & 0 \leq u \leq 1 \\ 0 & \text{o.w} \end{cases} \end{cases} \quad f_V(v) = \begin{cases} \int_{-\infty}^{\infty} f_{U,V}(u,v) du \\ = \int_0^{1-v} 24uv du \\ = \frac{24}{2} u^2 v \Big|_0^{1-v} \\ = \begin{cases} 12v(1-v)^2 & 0 \leq v \leq 1 \\ 0 & \text{o.w} \end{cases} \end{cases}$$

$$\begin{aligned} f_{U,V}(u,v) &\neq f_U(u) \cdot f_V(v) \\ 24uv &\neq [12u(1-u)^2] \cdot [12v(1-v)^2] \end{aligned}$$

It is **not** independent

For **independent random variables** X and Y ,

a. $E(g(X)h(Y)) = E(g(X))E(h(Y))$

b. $r_{X,Y} = E(XY) = E(X)E(Y)$

c. $Cov(X,Y) = \rho_{X,Y} = 0$ independent r.v. $\xrightarrow{\quad}$ uncorrelated.
 \nwarrow

d. $Var(X+Y) = Var(X) + Var(Y)$

e. $E(X|Y=y) = E(X)$ for all $y \in S_Y$

f. $E(Y|X=x) = E(Y)$ for all $x \in S_X$

Independent random variables are uncorrelated, but the reverse is not true all the time.

Ex 4.25) Are X and Y independent? Are X and Y uncorrelated?

$P_{X,Y}[x,y]$	$y = -1$	$y = 0$	$y = 1$
$x = -1$	0	0.25	0
$x = 1$	0.25	0.25	0.25

$$\begin{cases} P_X[-1] = 0.25 \\ P_X[1] = 0.75 \end{cases} \quad \begin{cases} P_Y[-1] = 0.25 \\ P_Y[0] = 0.5 \\ P_Y[1] = 0.25 \end{cases}$$

$$\begin{cases} P_{X,Y}[-1,-1] = 0 \\ P_X[-1]P_Y[-1] = [0.25] \cdot [0.25] \end{cases} \quad \text{Not independent}$$

$$\begin{cases} E[X] = \sum_{x=-1,1} x \cdot P_X[x] = (-1) \cdot [0.25] + 1 \cdot [0.75] = 0.5 \\ E[Y] = \sum_{y=-1,0,1} y \cdot P_Y[y] = (-1) \cdot [0.25] + 0 + 1 \cdot [0.25] = 0 \\ E[XY] = \begin{cases} \sum_{x=-1,1} \sum_{y=-1,0,1} xy P_{X,Y}[x,y] \\ = [-1 \cdot -1 \cdot 0] + [-1 \cdot 0 \cdot 0.25] + [-1 \cdot 1 \cdot 0] + [1 \cdot -1 \cdot 0.25] + [1 \cdot 0 \cdot 0.25] + [1 \cdot 1 \cdot 0.25] \\ = 0 + 0 + 0 - 0.25 + 0 + 0.25 \\ = 0 \end{cases} \end{cases}$$

$$Cov[X,Y] = \begin{cases} E[XY] - E[X]E[Y] \\ = 0 - 0.5 \cdot 0 \\ = 0 \end{cases} \Rightarrow \rho_{X,Y} = 0$$

This is **uncorrelated**.

Quiz 4.10) Random variables X_1 and X_2 are *independent and identically distributed (iid) with PDF*

$$f_x(x) = \begin{cases} 1 - \frac{x}{2} & 0 \leq x \leq 2 \\ 0 & \text{o.w} \end{cases}$$

a) What is the joint PDF?

Since X_1 and X_2 are independent,

$$f_{X_1, X_2}(x_1, x_2) = \begin{cases} f_{X_1}(x_1) \cdot f_{X_2}(x_2) \\ = \left(1 - \frac{x_1}{2}\right) \cdot \left(1 - \frac{x_2}{2}\right) & 0 \leq x_1 \leq 2, 0 \leq x_2 \leq 2 \\ 0 & \text{o.w} \end{cases}$$

b) Find the CDF of $Z = \max(X_1, X_2)$

$$P(Z \leq z) = \begin{cases} P(X_1 \leq z, X_2 \leq z) \\ = P(X_1 \leq z) P(X_2 \leq z) \\ = (F_X(z))^2 \end{cases}$$

$$F_X(x) = \begin{cases} \int_{-\infty}^{\infty} f_X(x) dx \\ = \int_{-\infty}^x f_X(x) dx \\ = \int_0^x \left(1 - \frac{x}{2}\right) dx \\ = \left(x - \frac{x^2}{4}\right)_0^x \\ = x - \frac{x^2}{4} \end{cases} \quad \Rightarrow \quad F_Z(z) = \begin{cases} 0 & z < 0 \\ \left(z - \frac{z^2}{4}\right)^2 & 0 \leq z < 2 \\ 1 & 2 < z \end{cases}$$

4.11 Bivariate Gaussian Random Variables

The Bivariate Gaussian distribution is a probability model for $[X \text{ and } Y]$ with the property that X and Y are each Gaussian random variables.

Random Variables X and Y have a **Bivariate Gaussian PDF** with parameters $\mu_1, \sigma_1, \mu_2, \sigma_2$, and ρ

$$f_{X,Y}(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp \left[- \left(\frac{\left(\frac{x-\mu_1}{\sigma_1} \right)^2 - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \left(\frac{y-\mu_2}{\sigma_2} \right)^2}{2(1-\rho^2)} \right) \right]$$

where μ_1 and μ_2 can be any real numbers, $\sigma_1 > 0, \sigma_2 > 0$, and $-1 < \rho < 1$

Let $\tilde{\mu}_2(x) = \mu_2 + \rho \frac{\sigma_2}{\sigma_1}(x - \mu_1)$, $\tilde{\sigma}_2 = \sigma_2\sqrt{1-\rho^2}$ to examine properties of

$$f_{X,Y}(x,y) = \left(\frac{1}{\sigma_1\sqrt{2\pi}} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} \right) \cdot \left(\frac{1}{\tilde{\sigma}_2\sqrt{2\pi}} e^{-\frac{(y-\tilde{\mu}_2(x))^2}{2\tilde{\sigma}_2^2}} \right) \quad (4.11-1)$$

The equation (4.11-1) becomes the products of two Gaussian PDFs.

If X and Y are the Bivariate Gaussian random variables in the equation (4.11-1)

- X is the Gaussian (μ_1, σ_1) random variable
- Y is the Gaussian (μ_2, σ_2) random variable

- $f_X(x) = \frac{1}{\sigma_1\sqrt{2\pi}} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}}$

- $f_Y(y) = \frac{1}{\sigma_2\sqrt{2\pi}} e^{-\frac{(y-\mu_2)^2}{2\sigma_2^2}}$

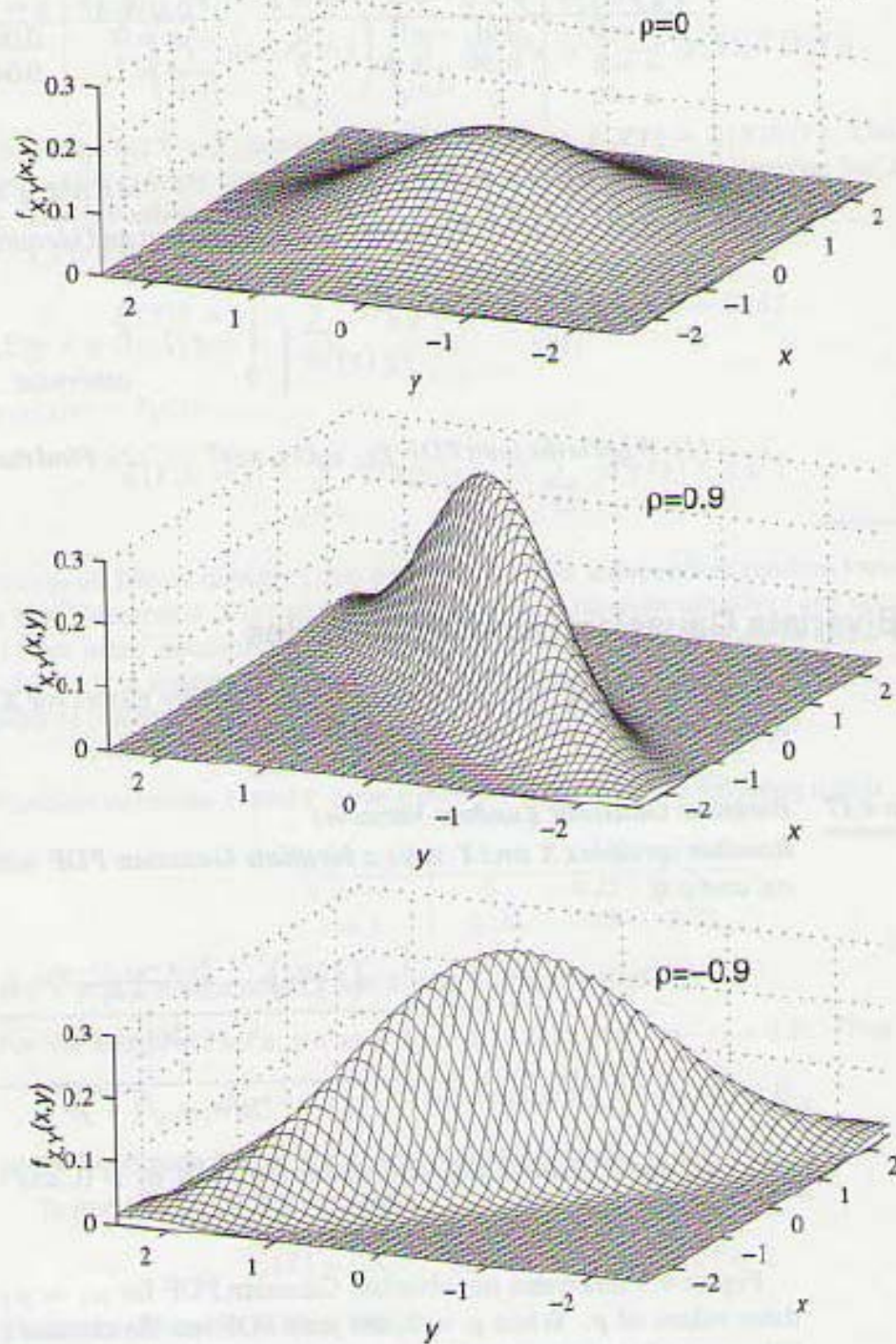


Figure 4.5 The Joint Gaussian PDF $f_{X,Y}(x,y)$ for $\mu_1 = \mu_2 = 0$, $\sigma_1 = \sigma_2 = 1$, and three values of ρ .

If X and Y are the Bivariate Gaussian random variables in the equation (4.11-1), the conditional PDF of Y given X is

$$f_{Y|X}(y|x) = \frac{1}{\tilde{\sigma}_2 \sqrt{2\pi}} e^{-\frac{(y - \tilde{\mu}_2(x))^2}{2\tilde{\sigma}_2^2}}$$

where, given $X = x$, the conditional expected value and variance of Y are

$$\tilde{\mu}_2(x) = \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (x - \mu_1), \quad \tilde{\sigma}_2^2 = \sigma_2^2 \sqrt{1 - \rho^2}$$

If X and Y are the Bivariate Gaussian random variables in the equation (4.11-1), the conditional PDF of X given Y is

$$f_{X|Y}(x|y) = \frac{1}{\tilde{\sigma}_1 \sqrt{2\pi}} e^{-\frac{(x - \tilde{\mu}_1(y))^2}{2\tilde{\sigma}_1^2}}$$

where, given $Y = y$, the conditional expected value and variance of X are

$$\tilde{\mu}_1(y) = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (y - \mu_2), \quad \tilde{\sigma}_1^2 = \sigma_1^2 \sqrt{1 - \rho^2}$$

Bivariate Gaussian random variables X and Y have correlation coefficient

$$\rho_{X,Y} = \rho$$

Bivariate Gaussian random variables X and Y are **uncorrelated if and only if they are independent**.

Quiz 4.11 Let X and Y be jointly Gaussian (0,1) random variables with correlation coefficient, $\rho = \frac{1}{2}$.

(1) What is the joint PDF of X and Y ?

$$f_{X,Y}(x,y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp \left[\frac{\left(\frac{x-\mu_1}{\sigma_1}\right)^2 - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \left(\frac{y-\mu_2}{\sigma_2}\right)^2}{2(1-\rho^2)} \right]$$

$$= \frac{1}{2\pi\sqrt{1-(\frac{1}{2})^2}} \exp \left[2 \left(\frac{x^2 - xy + y^2}{3} \right) \right]$$

where μ_1 and μ_2 can be any real numbers, $\sigma_1 > 0$, $\sigma_2 > 0$, and $-1 < \rho < 1$

$$\mu_1 = \mu_2 = 0, \sigma_1 = \sigma_2 = 1$$


```

clc; clear all;
x = -5:0.1:5;
y = -5:0.1:5;

mu1 = 0; mu2 = 0;
sig1 = 1; sig2 = 1;
ro = -0.90;%correlation coefficient

[x, y ] = meshgrid(x,y);

f = exp(-((x-mu1/sig1).^2 - (2*ro*(x-mu1).*(y-mu2))./(sig1*sig2) + ...
        (y-mu2/sig2).^2)./(2*(1-ro^2)))./(2*pi*sig1*sig2*(1-ro^2)^0.5);

meshc(x,y,f)

figure;
surf(x,y,f)
% figure;
% meshz(x,y,f)

whos

```

