Probability and Statistics with R

Assignment 2

Submission Nov 16-2022 (Wednesday)

Note: Below I explain how you collaborate on GitHub.

- 1. It will be group assignment.
- 2. A group would be of size at most 3. If you want to create a group size more than 3, you must take permission.
- 3. Decide among yourself and one of you create a GitHub repository for Probability Statistics Assignments.
- 4. In that repository add your group members as collaborator
- 5. Once you add your collaborator (or group members), create a folder and name it as Assignment_2
- 6. In that folder you should have 2 folders code and report. And one README.md file. Write a brief report in README.md file.
- 7. For each problem, you should create a separate GitHub issue. All your discussion should be documented in the issue.
- 8. In the issue mention clearly, which group member is taking ownership of what problem?
- 9. The other member should fork the repository in their GitHub account.
- 10. Once you have your forked the main repository in your GitHub account you should clone the repository in you local laptop or just download it as zip.
- 11. Once you develop the code you should commit the code first in your repository and then push it.
- 12. Finally you make the pull-request in the final repository.
- 13. Once a member make a pull request, the other members have to review the code.
- 14. While reviewing the code the reviewer may have to download the code and run the code in his or her system and reproduce the result.
- 15. If the result is reproduced then she or he would accept and merge the code in final repository.
- 16. At the end you submit the link of the repository in the moodle.
- 17. The entire process will be evaluated.

Problem 1

Suppose X denote the number of goals scored by home team in premier league. We can assume X is a random variable. Then we have to build the probability distribution to model the probability of number of goals. Since X takes value in $\mathbb{N} = \{0, 1, 2, \dots\}$, we can consider the geometric progression sequence as possible candidate model, i.e.,

$$S = \{a, ar, ar^2, ar^3, \cdots\}.$$

But we have to be careful and put proper conditions in place and modify S in such a way so that it becomes proper probability distributions.

- 1. Figure out the necesary conditions and define the probability distribution model using S. We need the following conditions for S to be a pdf:
 - 1. $0 \le ar^n \le 1 \forall n$.
 - 2. $\sum_{n=0} \mathbb{P}(X=n)$ converges to 1, hence r < 1.
 - 3. Since $\sum_{n=0} ar^n = \frac{a}{1-r}$ and using the above, we have $\mathbb{P}(X=n) = (1-r)r^n$. (It is clear that $X \sim Geom(1-r)$)

- 2. Check if mean and variance exists for the probability model. Yes, they exist, since we know the mean and variance of Geometric random variable.
- 3. Can you find the analytically expression of mean and variance.
 - 1. Mean

$$\begin{split} E[X] &= \sum_{n=0} n \mathbb{P}(X=n) \\ &= \sum_{n=0} n (1-r) r^n \\ &= r \sum_{n=0} n (1-r) r^{n-1} \\ &= r \sum_{n=1} (n-1) (1-r) r^{n-1} + r \sum_{n=1} (1-r) r^{n-1} \\ &= r E[X] + r (1-r) (\frac{1}{1-r}) \\ &= r E[X] + r \end{split}$$

2. Variance

$$\begin{split} E[X^2] &= \sum_{n=0} n^2 \mathbb{P}(X=n) \\ &= \sum_{n=0} n^2 (1-r) r^n \\ &= r \sum_{n=0} n^2 (1-r) r^{n-1} \\ &= r \sum_{n=1} (n-1)^2 (1-r) r^{n-1} + \sum_{n=1} 2n (1-r) r^n - r \sum_{n=1} (1-r) r^{n-1} \\ &= r E[X^2] + 2 E[X] - r (1-r) \frac{1}{1-r} \\ &= r E[X^2] + \frac{2r}{1-r} - r \\ &\Rightarrow E[X^2] &= \frac{r^2 + r}{(1-r)^2}. \end{split}$$
 Therefore, $Var(X) = E[X^2] - (E[X])^2 \\ &= \frac{r^2 + r}{(1-r)^2} - \frac{r^2}{(1-r)^2} \\ &= \frac{r}{(1-r)^2} \end{split}$

4. From historical data we found the following summary statistics

mean	median	variance	total number of matches	
1.5	1	2.25	380	

Using the summary statistics and your newly defined probability distribution model find the following:

- The above summary statistics do not seem to follow from a Geimetric distribution since if we assume mean is true and solve for r, the variance is incorrect, and vice versa. But for this question, we will model our parameters using the mean. So if $\frac{r}{1-r} = 1.5$, then r = 0.6, meaning that our model has a variance of 3.5
- a. What is the probability that home team will score at least one goal?

Ans.

$$\mathbb{P}(\text{at least one goal}) = 1 - \mathbb{P}(\text{no goal})$$

$$= 1 - (1 - r)$$

$$= r$$

$$= 0.6$$

b. What is the probability that home team will score at least one goal but less than four goal?

Ans.

$$\mathbb{P}(1 <= X < 4) = \sum_{i=1}^{3} \mathbb{P}(X = i)$$

$$= \sum_{i=1}^{3} (1 - r)r^{i}$$

$$= r(1 - r^{3})$$

$$= 0.47$$

- 5. Suppose on another thought you want to model it with off-the shelf Poisson probability models. Under the assumption that underlying distribution is Poisson probability find the above probabilities, i.e.,
- Once again, we will model it using the mean as parameter. The given summary statistics cannot be that of a Poisson distribution, since it has unequal mean and variance, but like in the previous question, we will assume mean is correct and hence, $\lambda = 1.5$. So, $\mathbb{P}(X = n) = \frac{e^{-\lambda}\lambda^n}{n!}$.
- a. What is the probability that home team will score at least one goal?

Ans.

$$\mathbb{P}(\text{at least one goal}) = 1 - \mathbb{P}(\text{no goal})$$

$$= 1 - e^{-\lambda}$$

$$= 0.77$$

b. What is the probability that home team will score at least one goal but less than four goal? Ans.

$$\mathbb{P}(1 <= X < 4) = \sum_{i=1}^{3} \mathbb{P}(X = i)$$

$$= \sum_{i=1}^{3} \frac{e^{-\lambda} \lambda^{i}}{i!}$$

$$= 0.758$$

6. Which probability model you would prefer over another?

Ans.

Model	Mean	Median	Variance
Expected	1.5	1	2.25
Geometric	1.5	2	3.5
Poisson	1.5	1	1.5

- Poisson has the least variance and also seems to have central tendencies closer to the expected values, hence it could be a better fit.
- 7. Write down the likelihood functions of your newly defined probability models and Poisson models. Clearly mention all the assumptions that you are making.

Ans.

(a)
$$\mathcal{L}(r|(x_1, x_2, ..., x_n)) = \Pi_i(1-r)r^{x_i} = (1-r)r^{\sum x_i}$$

(b)
$$\mathcal{L}(\lambda|(x_1, x_2, ..., x_n)) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} = \frac{e^{-n\lambda} \lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!}$$

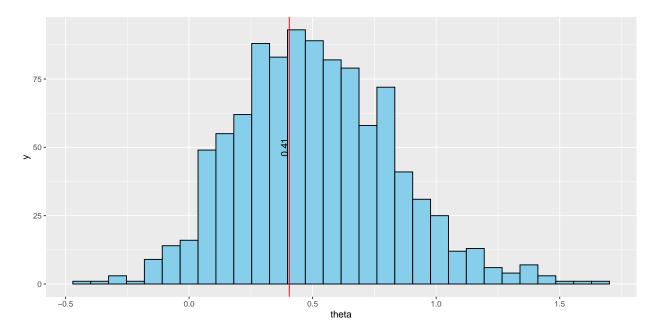
Problem 2: Simulation Study to Understand Sampling Distribution

Part A Suppose $X_1, X_2, \cdots, X_n \stackrel{iid}{\sim} Gamma(\alpha, \sigma)$, with pdf as

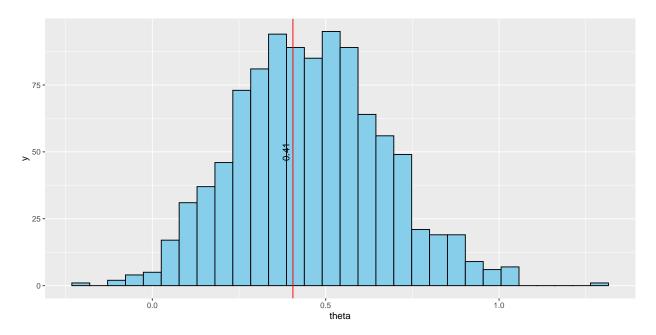
$$f(x|\alpha,\sigma) = \frac{1}{\sigma^{\alpha}\Gamma(\alpha)}e^{-x/\sigma}x^{\alpha-1}, \quad 0 < x < \infty,$$

The mean and variance are $E(X) = \alpha \sigma$ and $Var(X) = \alpha \sigma^2$. Note that shape = α and scale = σ .

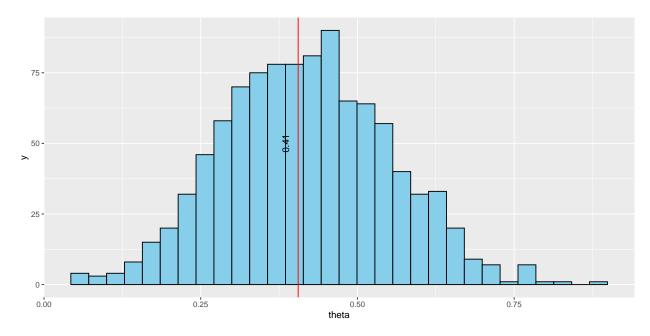
- 1. Write a function in R which will compute the MLE of $\theta = \log(\alpha)$ using optim function in R. You can name it MyMLE
- 2. Choose n=20, and alpha=1.5 and sigma=2.2
 - (i) Simulate $\{X_1, X_2, \cdots, X_n\}$ from rgamma(n=20,shape=1.5,scale=2.2)
 - (ii) Apply the MyMLE to estimate θ and append the value in a vector
 - (iii) Repeat the step (i) and (ii) 1000 times
 - (iv) Draw histogram of the estimated MLEs of θ .
 - (v) Draw a vertical line using abline function at the true value of θ .
 - (vi) Use quantile function on estimated θ 's to find the 2.5 and 97.5-percentile points.



- ## The 2.5 th percentile is -0.05863449
- ## The 97.5th percentile is 1.176011
- ## The gap between 2.5 and 97.5 percentile points is 1.234645
 - 3. Choose n=40, and alpha=1.5 and repeat the (2).



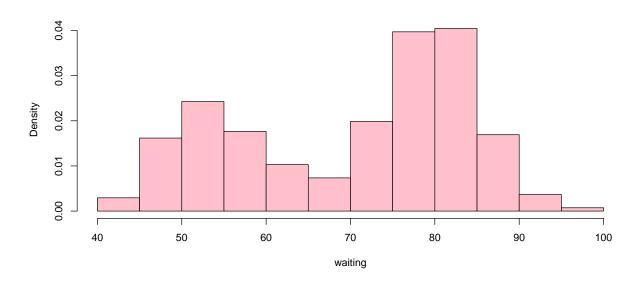
- ## The 2.5 th percentile is 0.06994037
- ## The 97.5th percentile is 0.8917431
- ## The gap between 2.5 and 97.5 percentile points is 0.8218028
 - 4. Choose n=100, and alpha=1.5 and repeat the (2).



- ## The 2.5 th percentile is 0.1718295
- ## The 97.5th percentile is 0.6742208
- ## The gap between 2.5 and 97.5 percentile points is 0.5023913
 - 5. Check if the gap between 2.5 and 97.5-percentile points are shrinking as sample size n is increasing? **Ans**. The gap between 2.5 and 97.5 percentile points for n=20, n=40 and n=100 are 1.177962, 0.819691, 0.5194647 respectively. The gap is decreasing as n is increasing.

Problem 3: Analysis of faithful datasets.

Consider the faithful datasets:



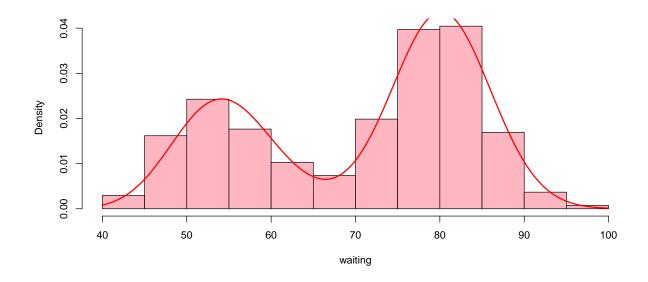
Fit following three models using MLE method and calculate **Akaike information criterion** (aka., AIC) for each fitted model. Based on AIC decides which model is the best model? Based on the best model calculate the following probability

$$\mathbb{P}(60 < \mathtt{waiting} < 70)$$

(i) Model 1:

$$f(x) = p*Gamma(x|\alpha,\sigma_1) + (1-p)N(x|\mu,\sigma_2^2), \quad 0$$

The optimal values of p, alpha, sigma_1, mu, sigma_2 is 0.3652577 82.78556 1.510993 80.16598 5.81013 ## AIC = 2076.18



(ii) **Model 2**:

$$f(x) = p * Gamma(x|\alpha_1, \sigma_1) + (1-p)Gamma(x|\alpha_2, \sigma_2), \quad 0$$

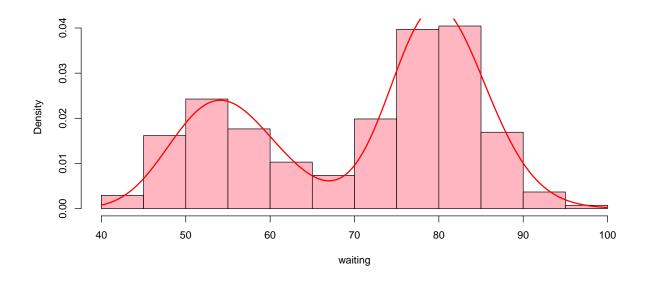
The optimal values of p, alpha, sigma_1, mu. sigma_2 is 0.3709333 79.66891 1.449311 199.7652 2.48809 ## AIC = 2076.116



(iii) Model 3:

$$f(x) = p * logNormal(x|\mu_1, \sigma_1^2) + (1-p)logNormal(x|\mu_1, \sigma_1^2), \quad 0$$

The optimal values of p, alpha, sigma_1, mu. sigma_2 is $0.6238638\ 4.384306\ 0.06972831\ 4.003846\ 0.114$ ## AIC = 2075.42



The AIC values for the three models are 2076.18, 2076.116 and 2075.42 respectively. Since model 3 has least value of AIC, it is considered as the best fit.

$$P(60 < waiting < 70) = P(X < 70) - P(X < 60) \label{eq:proposition}$$
 Where X has pdf in Model 3

[1] 0.09080635

Problem 4: Modelling Insurance Claims

Consider the Insurance datasets in the MASS package. The data given in data frame Insurance consist of the numbers of policyholders of an insurance company who were exposed to risk, and the numbers of car insurance claims made by those policyholders in the third quarter of 1973.

This data frame contains the following columns:

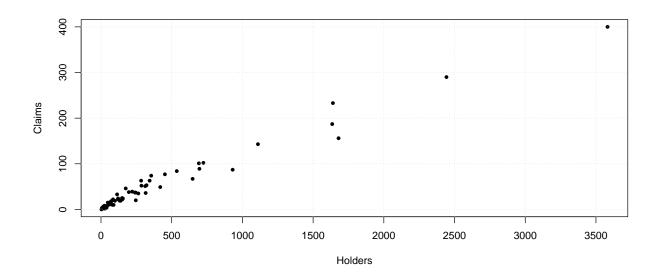
District (factor): district of residence of policyholder (1 to 4): 4 is major cities.

Group (an ordered factor): group of car with levels <1 litre, 1–1.5 litre, 1.5–2 litre, >2 litre.

Age (an ordered factor): the age of the insured in 4 groups labelled <25, 25–29, 30–35, >35.

Holders: numbers of policyholders.

Claims: numbers of claims



Note: If you use built-in function like 1m or any packages then no points will be awarded.

Part A: We want to predict the Claims as function of Holders. So we want to fit the following models:

$$\mathtt{Claims}_i = \beta_0 + \beta_1 \ \mathtt{Holders}_i + \varepsilon_i, \quad i = 1, 2, \cdots, n$$

Assume: $\varepsilon_i \sim N(0, \sigma^2)$. Note that $\beta_0, \beta_1 \in \mathbb{R}$ and $\sigma \in \mathbb{R}^+$.

The above model can also be re-expressed as,

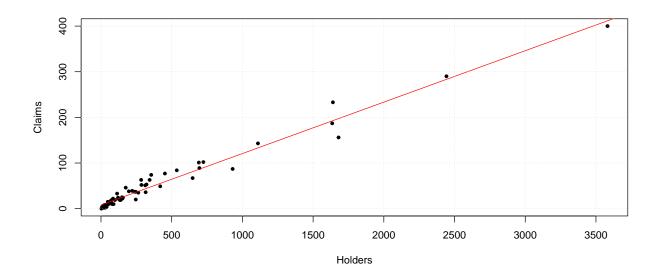
$${\tt Claims}_i \sim N(\mu_i, \sigma^2), \ where$$

$$\mu_i = \beta_0 + \beta_1 \; \text{Holders}_i + \varepsilon_i, \quad i = 1, 2, \cdots, n$$

- (i) Clearly write down the negative-log-likelihood function in R. Then use optim function to estimate MLE of $\theta = (\beta_0, \beta_1, \sigma)$
- (ii) Calculate Bayesian Information Criterion (BIC) for the model.

MLE= 8.121281 0.1126417 11.86943

BIC = 510.7587



Part B: Now we want to fit the same model with change in distribution:

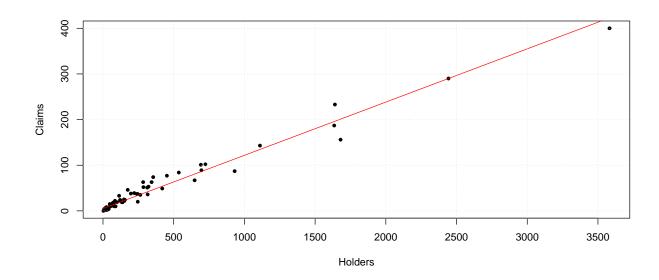
$$\mathtt{Claims}_i = \beta_0 + \beta_1 \; \mathtt{Holders}_i + \varepsilon_i, \quad i = 1, 2, \cdots, n$$

Assume : $\varepsilon_i \sim Laplace(0, \sigma^2)$. Note that $\beta_0, \beta_1 \in \mathbb{R}$ and $\sigma \in \mathbb{R}^+$.

- (i) Clearly write down the negative-log-likelihood function in R. Then use optim function to estimate MLE of $\theta = (\beta_0, \beta_1, \sigma)$
- (ii) Calculate Bayesian Information Criterion (BIC) for the model.

MLE= 5.084407 0.1166253 8.213428

BIC = 498.6869



Part C: We want to fit the following models:

$$Claims_i \sim LogNormal(\mu_i, \sigma^2), where$$

$$\mu_i = \beta_0 + \beta_1 \log(\text{Holders}_i), \quad i = 1, 2, ..., n$$

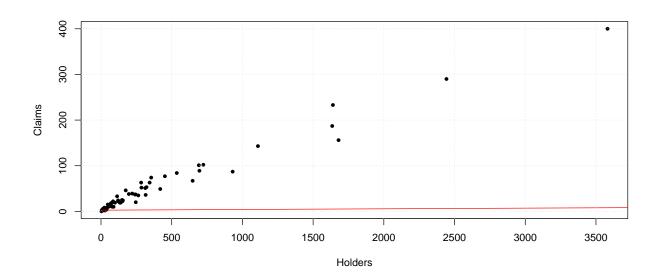
Note that $\beta_0, \beta_1 \in \mathbb{R}$ and $\sigma \in \mathbb{R}^+$.

- (i) Clearly write down the negative-log-likelihood function in R. Then use optim function to estimate MLE of $\theta = (\alpha, \beta, \sigma)$
- (ii) Calculate Bayesian Information Criterion (BIC) for the model.

Note that we have one point in our data set where Claims=0. Since the support of Lognormal distribution is $(0,\infty)$, we remove this point.

Optimal parameters are 2.639779 0.001472126 0.8229641

The BIC for this model 568.0199



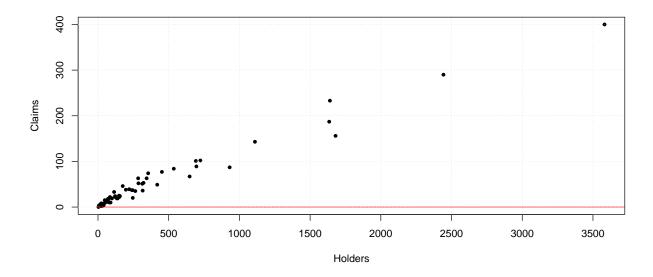
Part D: We want to fit the following models:

$$Claims_i \sim Gamma(\alpha_i, \sigma), where$$

$$log(\alpha_i) = \beta_0 + \beta_1 \log(\text{Holders}_i), \quad i = 1, 2, ..., n$$

Optimal parameters are 0 0 49.23438

The BIC for this model 639.2404



(iii) Compare the BIC of all three models

Ans. The BIC of the models are 510.76, 498.69, 568.02 and 639.24 respectively. Since the number of parameters being estimated are the same in all cases (3), lower the BIC, higher the likelihood and hence better the model. Based on this remark, the second model, i.e., $\epsilon_i \sim \text{Laplace}(0, \sigma^2)$ is the best fit for the data. The Gauss-Markov assumptions $(\epsilon_i \sim \mathcal{N}(0, \sigma^2))$ follow closely as the second best model. The 4th model $(\epsilon_i \sim \Gamma(0, \sigma^2))$ fits the worst.

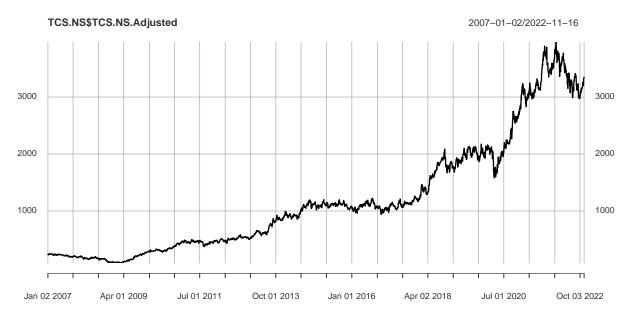
Problem 5: Computational Finance - Modelling Stock prices

Following piece of code download the prices of TCS since 2007

```
## [1] "TCS.NS"
```

##		TCS.NS.Open	TCS.NS.High	TCS.NS.Low	TCS.NS.Close	TCS.NS.Volume
##	2022-11-09	3249.8	3249.80	3201.65	3216.05	1162267
##	2022-11-10	3170.0	3225.00	3170.00	3205.65	1573092
##	2022-11-11	3269.6	3341.60	3255.05	3315.95	3265394
##	2022-11-14	3324.0	3349.00	3309.00	3335.50	1342074
##	2022-11-15	3321.0	3339.95	3292.00	3332.60	1400708
##	2022-11-16	3338.9	3367.90	3321.45	3355.35	1747982
##		TCS.NS.Adjus	ted			
##	2022-11-09	3216	.05			
##	2022-11-10	3205	.65			
##	2022-11-11	3315	.95			
##	2022-11-14	3335	.50			
##	2022-11-15	3332	.60			
##	2022-11-16	3355	.35			

Plot the adjusted close prices of TCS



Download the data of market index Nifty50. The Nifty 50 index indicates how the over all market has done over the similar period.

```
## [1] "^NSEI"
```

##		NSEI.Open	NSEI.High	NSEI.Low	NSEI.Close	NSEI.Volume	NSEI.Adjusted
##	2022-11-09	18288.25	18296.40	18117.50	18157.00	307200	18157.00
##	2022-11-10	18044.35	18103.10	17969.40	18028.20	256500	18028.20
##	2022-11-11	18272.35	18362.30	18259.35	18349.70	378500	18349.70
##	2022-11-14	18376.40	18399.45	18311.40	18329.15	301400	18329.15
##	2022-11-15	18362.75	18427.95	18282.00	18403.40	250900	18403.40
##	2022-11-16	18398.25	18442.15	18344.15	18409.65	0	18409.65

Plot the adjusted close value of Nifty50

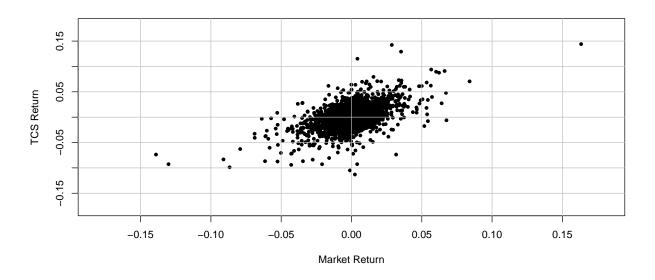


Log-Return

We calculate the daily log-return, where log-return is defined as

$$r_t = \log(P_t) - \log(P_{t-1}) = \Delta \log(P_t),$$

where P_t is the closing price of the stock on t^{th} day.



• Consider the following model:

$$r_t^{TCS} = \alpha + \beta r_t^{Nifty} + \varepsilon,$$

where $\mathbb{E}(\varepsilon) = 0$ and $\mathbb{V}ar(\varepsilon) = \sigma^2$.

- 1. Estimate the parameters of the models $\theta = (\alpha, \beta, \sigma)$ using the method of moments type plug-in estimator discussed in the class.
- ## Method of moments plug in estimator of theta is: (0.0004633056 , 0.7436614 , 0.01618073)
 - 2. Estimate the parameters using the 1m built-in function of R. Note that 1m using the OLS method.
- ## Estimator of theta using lm function is: (0.0004633056 , 0.7436614 , 0.01618293)
 - 3. Fill-up the following table

Parameters	Method of Moments	OLS
α	0.0004611203	0.0004611203
β	0.7436965	0.7436965
σ	0.01618653	0.01618873

- 4. If the current value of Nifty is 18000 and it goes up to 18200. The current value of TCS is Rs. 3200/-. How much you can expect TCS price to go up?
- ## We can expect the TCS price to go up to 3227.899