

# Causal Inference with Noisy and Missing Covariates

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- **Causal Inference with Noisy and Missing Covariates** by Nathan Kallus, Xiaojie Mao, Madeleine Udell (2018)

## Summary

- Observational vs. experimental
- Samples contain Treated  $T_i = 1$  and Control  $T_i = 0$
- Covariates  $U_i$  (either observed or unobserved)
- Outcome of interest  $Y_i$
- Confounders are covariates that effect both the outcome and treatment

## Assumptions

- Stable Unit Treatment Value Assignment
- Consistency
- Ignorability
- Positivity

To estimate

- Average Treatment Effect, ATE:  $\mathbb{E}[Y^1] - \mathbb{E}[Y^0]$

# Causal Inference from observational data

## Potential Outcome framework

- Matching (Mahalanobis-distance, calipers etc.)
- Propensity score based inverse probability weighting (IPW)
- Doubly Robust methods

## Techniques

- Assume a causal relationship (a causal graph)
- Identify set of confounders to be controlled (or measured)
- Use backdoor criteria or frontdoor criteria etc. to assist discovery of valid adjustment set

# Problem Setting

## Notation

- Treatment  $T \in \{0, 1\}^N$
- Unobserved Covariates  $U \in \mathbb{R}^{N \times r}$
- Outcome  $Y \in \mathbb{R}^{N \times 1}$
- Observed noisy and missing covariates  $X \in \mathbb{R}^{N \times p}$

## Assumptions

- Linear  $Y_i = U_i \alpha + \tau T_i + \epsilon_i$
- Low rank matrix factorization of the observed  $X$  to yield confounders  $X = UV^T + W$
- Exponential Family Matrix Completion preprocessing

## Claim

- Preprocessing to augment wide variety of causal inferences
- Matrix factorization based preprocessing is a general framework
- Seamlessly integrated into regression adjustment, propensity score reweighting, matching etc.
- Bounded error on the induced average treatment effect estimator



## Problem Formulation

- Estimate  $\tau = \mathbb{E}[Y_i(1) - Y_i(0)]$
- **Unconfoundedness assumption**  $\mathbb{P}(Y_i | T_i, U_i) = \mathbb{P}(Y_i | U_i) \quad \forall i$
- $X \in \mathbb{R}^{N \times p}$  and observed over  $\Omega \subset [N] \times [p], p < N$
- Generative assumption  $X_{ij} \sim \mathcal{N}(U_i^T V_j, 1)$
- $W \in \mathbb{R}^{N \times p}$  independent with (mean, variance)  $= (0, \sigma_w^2)$
- Linear  $Y_i = U_i^T \alpha + \tau T_i + \epsilon_i$
- Additive noise model  $X = UV^T + W$

## Measurement Noise and Bias

Asymptotic bias of least squares estimator in linear regression of  $Y_i$  on  $X_i, T_i$

$$\frac{\mathbb{E}(T_i U_i) \mathbb{E}(U_i^T U_i)^{-1} [\frac{1}{\sigma_w^2} V^T V + \mathbb{E}(U_i^T U_i)^{-1}]^{-1} \alpha}{\mathbb{E}(T_i^2) - \mathbb{E}(T_i U_i) [\frac{1}{\sigma_w^2} V^T V + \mathbb{E}(U_i^T U_i)]^{-1} \mathbb{E}(U_i^T U_i)}$$

This asymptotically diminishes to 0, when  $\|V\| \rightarrow \infty$

# Inference with low rank matrix decomposition

## Low rank matrix factorization preprocessing

**Low rank Assumption** Observed  $X$  is a noisy realization of a low rank matrix  $\Phi \in \mathbb{R}^{N \times p}$  with rank  $r \ll \min\{N, p\}$

## Missing Completely At Random Assumption

$\forall i, j \in \Omega, i \sim \text{Unif}([N]), j \sim \text{Unif}([p])$

## Natural Exponential Family Assumption

$\mathbb{P}(X_{ij}|\Phi_{ij}) = h(X_{ij})\exp(X_{ij}\Phi_{ij} - G(\Phi_{ij}))$ , where  $G : \mathbb{R} \mapsto \mathbb{R}$  is the log-partition and strictly convex,  $\nabla^2 G \geq e^{-\eta|u|}$ , for  $\eta > 0$ , and  $u \in \mathbb{R}$

EFMC estimates using regularized M-estimator:

$$\hat{\Phi} = \min_{\Phi} \frac{Np}{|\Omega|} [\sum_{(i,j) \in \Omega} \log \mathbb{P}(X_{ij}|\Phi_{ij})] + \lambda \|\Phi\|_*$$

Left singular matrix of  $\hat{\Phi}$  is an estimate of the confounder  $U$

## Theoretical guarantee (sufficient conditions)

**Definition** Principal angle between column spaces of two matrices  $M, \hat{M}$  is defined as  $\angle(\hat{M}, M) = \sqrt{1 - \sigma_{\min r, k}^2(\hat{M}^T M)}$

**Theorem** There exists a constant  $c > 0$  such that with probability at least  $1 - 2\exp(-c\sqrt{N})$ ,

$$|\hat{\tau} - \tau^*| \leq \frac{\frac{2A}{\sqrt{N}} \|T\| (\frac{1}{\sqrt{Nr}} \|U\|) (r\angle(U, \hat{U})) - \frac{\sigma}{N^{1/4}}}{\frac{1}{N} T^T (I - P_U) T - \frac{2}{N} \|T\|^2 \angle(U, \hat{U})}, \text{ which } \rightarrow 0 \text{ as } N \rightarrow \infty, \text{ where}$$
$$\|\alpha\|_{\max} \leq A$$

## Theoretical guarantee (sufficient conditions)

- $\|\alpha\|_{\max} \leq A$
- $\frac{1}{\sqrt{Nr}}\|U\|$
- $\frac{1}{N}T^T(I - P_U)T$  is bounded away from 0
- $r\angle\hat{U}, U \rightarrow 0$  as  $N \rightarrow 0$
- Unconfoundedness

## Theoretical guarantee(Confounders and Covariates Loadings)

- $\underline{v}, \bar{v}, c_V, c_L > 0$
- For  $i \in [N]$ ,  $U_i$  are i.i.d Gaussian samples with covariance  $\Sigma_{r \times r} = LL^T$ , full rank  $L \in \mathbb{R}^{r \times r}$ , such that  $\frac{1}{\sqrt{r}}\|L\| < c_L$  (Gaussian random design)
- $\underline{v}p \leq \sigma_r^2(VL^T) \leq \sigma_1^2(VL^T) \leq \bar{v}p$
- $\frac{\max_j \|V_j\|}{\|V\|_F} \leq \frac{c_V}{\sqrt{p}}, j \in [p]$  (excludes degenerate case)

## Theoretical guarantee

- Let  $X_{ij}$  be sub-exponential on  $U_i$  with parameter  $\sigma'$  for  $\forall i, j$
- $T_i$  is almost certainly not a linear combination of  $U_i$
- Suppose EFMC is used as the preprocessing step with  $\lambda = 2c_0\sigma'\sqrt{Np}\sqrt{\frac{r\bar{N}\log\bar{N}}{|\Omega|}}$ , where  $\bar{N} = \max(N, p)$ , and  $|\Omega| > c_1 r \bar{N} \log \bar{N}$  for  $c_0, c_1 > 0$
- $\exists \delta$ , s.t.  $p^{1+\delta}/N \rightarrow 0$

## Theoretical guarantee (Consistent Estimator)

**Theorem** There exist constants  $c_2, c_3, c_{\sigma', \eta}$  such that with probability at least  $1 - c_2 \exp(-c_3 N^{1/2}) - c_2 N^{-1/2} - 2 \exp(-c_3 p^\delta)$ ,

$$|\hat{\tau} - \tau| \leq \frac{A_{CL} c_{\sigma', \eta} c_V \sqrt{\frac{r^5 \bar{r} \bar{N} \log \bar{N}}{|\Omega|}} - \frac{\sigma}{N^{1/4}} [\sqrt{\frac{\underline{v}}{\underline{v} + 2\bar{v}}} - \Lambda(r, \bar{N}, |\Omega|)]}{[\sqrt{\frac{\underline{v}}{\underline{v} + 2\bar{v}}} - \Lambda(r, \bar{N}, |\Omega|)] [\frac{1}{N} T^T (I - P_U) T - 2\Lambda(r, \bar{N}, |\Omega|)]}, \text{ where}$$

$$\Lambda(r, \bar{N}, |\Omega|) = c_{\sigma', \eta} c_V \sqrt{\frac{\bar{r} r^3 \bar{N} \log \bar{N}}{|\Omega|}}, \text{ and } \bar{r} = \max r, \log \bar{N}$$



## Numerical Results (Consistent Estimator)

- $r^5 \bar{r} \bar{N} \log \bar{N} / |\Omega| \rightarrow 0$  vs.  $r \bar{N} \log \bar{N} / |\Omega| \rightarrow 0$
- Numerical results:  $N = 1500, p = 1450, r = 5$ , that is  $r^6 \gg N$

# Questions

Thank you