

```

#!/usr/bin/env python3
# -*- coding: utf-8 -*-
"""
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"""
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import pandas as pd
import numpy as np
import matplotlib.pyplot as plt
import scipy.stats as stats
from arch import arch_model, unitroot
import statsmodels.api as sm_api
from statsmodels.stats import import stattools
from datetime import datetime
import matplotlib.dates as mdates

#####
#PART 1

#load and sort the data
index_data=pd.read_excel("IndexClosingPrices.xlsx",index_col="Date").\
    sort_index().dropna()

#####
#PART 2

#Calculate return data
#return for S&P500
sp_return=(index_data["SP500"].pct_change())

#return for ASX
asx_return=(index_data["ASX"].pct_change())

#return for DAX
dax_return=(index_data["DAX"].pct_change())

combine_data=pd.concat([sp_return,asx_return,dax_return],axis=1).\
    sort_index().dropna()
combine_data.columns=["Return S&P500","Return ASX","Return DAX"]

#####
#PART 3 ~ Price Series

#plot for PRICE SERIES
# Filter data for the desired time range (up to November 29, 2023)
index_data = index_data[index_data.index <= datetime(2023, 11, 29)]

# Function to plot Price series

```

```

def plot_price_series(data, title):
    fig, ax = plt.subplots()
    data.plot(title=title, grid=True, legend=True, ax=ax)
    ax.xaxis.set_major_locator(mdates.YearLocator(4))
    ax.set_xlabel("Date")
    ax.set_ylabel("Return")

# Plotting return series for S&P500
plot_price_series(index_data["SP500"], "Price Series - S&P500")

# Plotting return series for ASX
plot_price_series(index_data["ASX"], "Price Series - ASX")

# Plotting return series for DAX
plot_price_series(index_data["DAX"], "Price Series - DAX")

plt.show()

#####
#PART 3 ~ Return Series

#plot and sort for RETURN SERIES
# Filter data for the desired time range (up to November 29, 2023)
combine_data = combine_data[combine_data.index <= datetime(2023, 11, 29)]

# Function to plot return series
def plot_return_series(data, title):
    fig, ax = plt.subplots()
    data.plot(title=title, grid=True, legend=True, ax=ax)
    ax.xaxis.set_major_locator(mdates.YearLocator(4))
    ax.set_xlabel("Date")
    ax.set_ylabel("Return")

# Plotting return series for S&P500
plot_return_series(combine_data["Return S&P500"], "Return Series - S&P500")

# Plotting return series for ASX
plot_return_series(combine_data["Return ASX"], "Return Series - ASX")

# Plotting return series for DAX
plot_return_series(combine_data["Return DAX"], "Return Series - DAX")

plt.show()

#The time series plots appear noisy as expected, aligning with our intended
#analysis. While visual inspection suggests stationarity in all three,
#it's crucial not to affirm stationarity without conducting the augmented
#Dickey-Fuller Test on the return data. We'll rely on the 5% confidence level
# benchmarks to assess data stationarity based on the test statistic values.

#####
#Part 4 ~ Dickey-Fuller test

adf_test1 = unitroot.ADF(combine_data["Return S&P500"])

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```

print(adf_test1.summary())

adf_test2 = unitroot.ADF(combine_data["Return ASX"])
print(adf_test2.summary())

adf_test3 = unitroot.ADF(combine_data["Return DAX"])
print(adf_test3.summary())

# The series is weakly stationary, which is sufficient for modelled using GARCH

#####
# PART 5 ~ Fitting Models

# Fit GARCH(1,1) models with different distributions

# Use two different models for the mean
# Use a GARCH(1,1) process for the variance

# Note: the distribution for the residuals is specified by using
# input dist = "StudentsT"

# Fit GARCH(1,1) with Student's t distribution for S&P500 constant mean
sp500_con = arch_model(100*combine_data["Return S&P500"]\
                        ,mean='constant',vol='GARCH', p=1, q=1,\
                        dist="StudentsT").fit()
print(sp500_con.summary())

# Fit GARCH(1,1) with Student's t distribution for S&P500 "AR" mean
sp500_ar = arch_model(100*combine_data["Return S&P500"],\
                      mean='AR', lags=1,vol='GARCH', p=1, q=1,\
                      dist="StudentsT").fit()
print(sp500_ar.summary())

# Fit GARCH(1,1) with Student's t distribution for ASX constant mean
asx_con = arch_model(100*combine_data["Return ASX"],\
                    mean='constant',vol='GARCH', p=1, q=1,\
                    dist="StudentsT").fit()
print(asx_con.summary())

# Fit GARCH(1,1) with Student's t distribution for ASX "AR" mean
asx_ar = arch_model(100*combine_data["Return ASX"],\
                   mean='AR', lags=1,vol='GARCH', p=1, q=1,\
                   dist="StudentsT").fit()
print(asx_ar.summary())

# Fit GARCH(1,1) with Student's t distribution for DAX constant mean
dax_con = arch_model(100*combine_data["Return DAX"],\
                    mean='constant',vol='GARCH', p=1, q=1,\
                    dist="StudentsT").fit()
print(dax_con.summary())

# Fit GARCH(1,1) with Student's t distribution for DAX "AR" mean
dax_ar = arch_model(100*combine_data["Return DAX"],\
                   mean='AR', lags=1,vol='GARCH', p=1, q=1,\

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                                dist="StudentsT").fit()
print(dax_ar.summary())
#####

# checking for serial correlation of the residuals using DURBIN WATSON test
# value below 2 indicates a positive auto correlation
# value higher than 2 indicates a negative serial correlation

# Return SP500

dw_1 = stattools.durbin_watson(sp500_con.resid.dropna())
print("\n\nFor sp500_con, the DW statistics is {0}".format(dw_1))

#The test gives us a result of around 2.108, which is not far off from 2.
#The residuals are not very serially correlated at all.

dw_2 = stattools.durbin_watson(sp500_ar.resid.dropna())
print("For sp500_ar, the DW statistics is {0}".format(dw_2))

#The test gives us a result of around 2.08, which is not far off from 2.
#The residuals are not very serially correlated at all.

# Return ASX

dw_3 = stattools.durbin_watson(asx_con.resid.dropna())
print("\n\nFor asx_con, the DW statistics is {0}".format(dw_3))

#The test gives us a result of around 1.95, which is not far off from 2.
#The residuals are not very serially correlated at all.

dw_4 = stattools.durbin_watson(asx_ar.resid.dropna())
print("For asx_ar, the DW statistics is {0}".format(dw_4))

#The test gives us a result of around 2.08, which is not far off from 2.
#The residuals are not very serially correlated at all.

# Return DAX

dw_5 = stattools.durbin_watson(dax_con.resid.dropna())
print("\n\nFor dax_con, the DW statistics is {0}".format(dw_5))

#The test gives us a result of around 2.005, which is not far off from 2.
#The residuals are not very serially correlated at all.

dw_6 = stattools.durbin_watson(dax_ar.resid.dropna())
print("For dax_ar, the DW statistics is {0}".format(dw_6))

#The test gives us a result of around 2.02, which is not far off from 2.
#The residuals are not very serially correlated at all.

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# The Durbin Watson test for all the models give a value close to 2,
# which does not show significant serial correlation,
#hence rejecting the null hypothesis

#####
# Perform log-likelihood ratio test
# and Compare the models using the likelihood ratio test

# sp500_con vs sp500_ar
LR_stat = -2*(sp500_con.loglikelihood-sp500_ar.loglikelihood)
dof = 1 # There is one less parameter in the smaller model
LR_pval = stats.chi2.sf(LR_stat, dof)

print("\nThe Likelihood ratio test statistics is {0}".format(LR_stat))
print("The corresponding p-value is {0}".format(LR_pval))

#The results seem to indicate that since the test statistic is less than the
#critical value, we have insufficient evidence to reject the null hypothesis
#so the simpler model seems to be good enough for our purposes.

# asx_con vs asx_ar
LR_stat = -2*(asx_con.loglikelihood-asx_ar.loglikelihood)
dof = 1
LR_pval = stats.chi2.sf(LR_stat, dof)

print("\nThe Likelihood ratio test statistics is {0}".format(LR_stat))
print("The corresponding p-value is {0}".format(LR_pval))

#The results seem to indicate that since the test statistic is greater than the
#critical value, we have sufficient evidence to reject the null hypothesis
#so the complex model seems to be a better fit.

# dax_con vs dax_ar
LR_stat = -2*(dax_con.loglikelihood-dax_ar.loglikelihood)
dof = 1
LR_pval = stats.chi2.sf(LR_stat, dof)

print("\nThe Likelihood ratio test statistics is {0}".format(LR_stat))
print("The corresponding p-value is {0}".format(LR_pval))

#The results seem to indicate that since the test statistic is greater than the
#critical value, we have sufficient evidence to reject the null hypothesis
#so the complex model seems to be a good fit for our use.

# Since the p-value is greater than 0.05 for SP&500 & ASX,
# with a 95% confidence the null
# hypothesis is accepted (or fails to be rejected). This implies that both
# models are equally good. Since this is the case, the simpler of the two
# models should be used.
# However, note that at a 94% confidence level, the null hypothesis of SP&500
# would be rejected, i.e. the larger model may be considered better.
# So in this case the likelihood ratio test is a little inconclusive.
# For DAX, p-value is substantially lower than 0.05, with a 95% confidence
# null hypothesis is rejected. This suggests that adding another additional

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# term (lag term), is considered better

#Based on the provided LRT results, there is evidence to suggest that the more
#complex AR-GARCH model is preferred over the constant mean GARCH model for
# each index.

# As part of any decision, the square of the (standardized) residuals should
# be inspected for GARCH effects.

#####
# plotting residuals

# Model 1 ~ sp500_con
sp500_con.plot()

sm_api.graphics.tsa.plot_acf(sp500_con.resid, lags=10)
plt.grid()
plt.title("sp500_con ACF Residuals")

sm_api.graphics.tsa.plot_acf(sp500_con.resid**2, lags=10)
plt.grid()
plt.title("sp500_con ACF Residuals Squared")

# The results_sp500_con plots of the squared residuals show very clearly that
#there are GARCH effects that the model doesn't seem to be handling correctly.

# Model 2 ~ sp500_ar
sp500_ar.plot()

sm_api.graphics.tsa.plot_acf(sp500_ar.resid, lags=10)
plt.grid()
plt.title("sp500_ar ACF Residuals")

sm_api.graphics.tsa.plot_acf(sp500_ar.resid**2, lags=10)
plt.grid()
plt.title("sp500_ar ACF Residuals Squared")

# The sp500_ar plots of the squared residuals seem to indicate that it
#deals withthe GARCH effects better than mdlA.

# Model 3 ~ asx_con
asx_con.plot()

sm_api.graphics.tsa.plot_acf(asx_con.resid, lags=10)
plt.grid()
plt.title("asx_con ACF Residuals")

sm_api.graphics.tsa.plot_acf(asx_con.resid**2, lags=10)
plt.grid()
plt.title("asx_con ACF Residuals Squared")

# The asx_con plots of the squared residuals show very clearly that
#there are GARCH effects that the model doesn't seem to be handling correctly.

# Model 4 ~ asx_ar

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```

asx_ar.plot()

sm_api.graphics.tsa.plot_acf(asx_ar.resid, lags=10)
plt.grid()
plt.title("asx_ar ACF Residuals")

sm_api.graphics.tsa.plot_acf(asx_ar.resid**2, lags=10)
plt.grid()
plt.title("asx_ar ACF Residuals Squared")

# The asx_ar plots of the squared residuals seem to indicate that it deals
# with the GARCH effects better than results_asx_con.

# Model 5 ~ dax_con
dax_con.plot()

sm_api.graphics.tsa.plot_acf(dax_con.resid, lags=10)
plt.grid()
plt.title("rdax_con ACF Residuals")

sm_api.graphics.tsa.plot_acf(dax_con.resid**2, lags=10)
plt.grid()
plt.title("dax_con ACF Residuals Squared")

# The dax_con plots of the squared residuals show very clearly that there are
# GARCH effects that the model doesn't seem to be handling correctly.

# Model 6 ~ dax_ar
dax_ar.plot()

sm_api.graphics.tsa.plot_acf(dax_ar.resid, lags=10)
plt.grid()
plt.title("dax_ar ACF Residuals")

sm_api.graphics.tsa.plot_acf(dax_ar.resid**2, lags=10)
plt.grid()
plt.title("dax_ar ACF Residuals Squared")

# The dax_ar plots of the squared residuals seem to indicate that it deals
# with the GARCH effects better than results_dax_con.

#####
# Plot each as a histogram

#to check for normality of residuals create histograms

#for sp500_con
fig, ax = plt.subplots()
plt.hist(sp500_con.resid, bins=100, axes=ax)
ax.set_title("sp500_con Residuals")
ax.grid(True)

#for sp500_ar
fig, ax = plt.subplots()
plt.hist(sp500_ar.resid, bins=100, axes=ax)

```

```

ax.set_title("sp500_ar Residuals")
ax.grid(True)

#for asx_con
fig, ax = plt.subplots()
plt.hist(asx_con.resid, bins=100, axes=ax)
ax.set_title("asx_con Residuals")
ax.grid(True)

#for asx_ar
fig, ax = plt.subplots()
plt.hist(asx_ar.resid, bins=100, axes=ax)
ax.set_title("asx_ar Residuals")
ax.grid(True)

#for dax_con
fig, ax = plt.subplots()
plt.hist(dax_con.resid, bins=100, axes=ax)
ax.set_title("dax_con Residuals")
ax.grid(True)

#for dax_ar
fig, ax = plt.subplots()
plt.hist(dax_ar.resid, bins=100, axes=ax)
ax.set_title("dax_ar Residuals")
ax.grid(True)

#####
# Plot each as a qq-plot

#to check for normality of residuals using qq-plots

#for sp500_con
fig, ax = plt.subplots()
stats.probplot(sp500_con.resid, dist=stats.norm,
               sparams=(np.mean(sp500_con.resid), np.std(sp500_con.resid)),
               plot=plt)
ax.set_title("sp500_con Residuals")
ax.grid(True)

# for sp500_ar
fig, ax = plt.subplots()
stats.probplot(sp500_ar.resid, dist=stats.norm,
               sparams=(np.mean(sp500_ar.resid), np.std(sp500_ar.resid)),
               plot=plt)
ax.set_title("sp500_ar Residuals")
ax.grid(True)

# for asx_con
fig, ax = plt.subplots()
stats.probplot(asx_con.resid, dist=stats.norm,
               sparams=(np.mean(asx_con.resid), np.std(asx_con.resid)),
               plot=plt)
ax.set_title("asx_con Residuals")

```



```

ax.grid(True)

# for asx_ar
fig, ax = plt.subplots()
stats.probplot(asx_ar.resid, dist=stats.norm,
               sparams=(np.mean(asx_ar.resid), np.std(asx_ar.resid)),
               plot=plt)
ax.set_title("asx_ar Residuals")
ax.grid(True)

# for dax_con
fig, ax = plt.subplots()
stats.probplot(dax_con.resid, dist=stats.norm,
               sparams=(np.mean(dax_con.resid), np.std(dax_con.resid)),
               plot=plt)
ax.set_title("dax_con Residuals")
ax.grid(True)

# for dax_ar
fig, ax = plt.subplots()
stats.probplot(dax_ar.resid, dist=stats.norm,
               sparams=(np.mean(dax_ar.resid), np.std(asx_ar.resid)),
               plot=plt)
ax.set_title("dax_ar Residuals")
ax.grid(True)

#NOTE ~
#The histogram and QQ plot data for all the models seems inconclusive
#since all the histogram plots seem to be showing kurtosis and mean peak
#similar to normal distributions while the QQ plots seemed to deviate at both
#ends indicating a lack of normality of all the models.

#####
# Answer ~ 1

#summary offindings for each index:

# for all three naming S&P500, ASX and DAX:
#ADF Test: The process is weakly stationary.
#GARCH Coefficients: The coefficients seem reasonable.
#LRT: The AR-GARCH model is preferred over the constant mean GARCH model.

#The GARCH(1,1) models for the ASX and DAX indices seems to be an appropriate
#fit.

#For the S&P 500 index, the GARCH(1,1) model seemed to provide some
#inconclusive findings since the p-value from the log-likelihood test just
#barely exceeded the 0.05 benchmark necessary for a 5% confidence interval
#when comparing the constant model with the complex model. However, this would
#not be the case if we changed our guidelines to a 6% confidence level.

#The findings led to the conclusion that the squared residuals from lagged
#models for all three indices (sp500_ar, asx_ar, and dax_ar) consistently
#fall within acceptable ranges. This implies the effective capture of

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#time-varying variance effects, strongly suggesting the presence and capture  
#of GARCH effects.

#####  
# Answer ~ 2

#mu at t is dependent on returns at time t-1 so mu at t+1 is dependent on  
#returns from time t  
#coefficient of return lag is irrelevant since it was established that  
#return at t is 0 in the question.

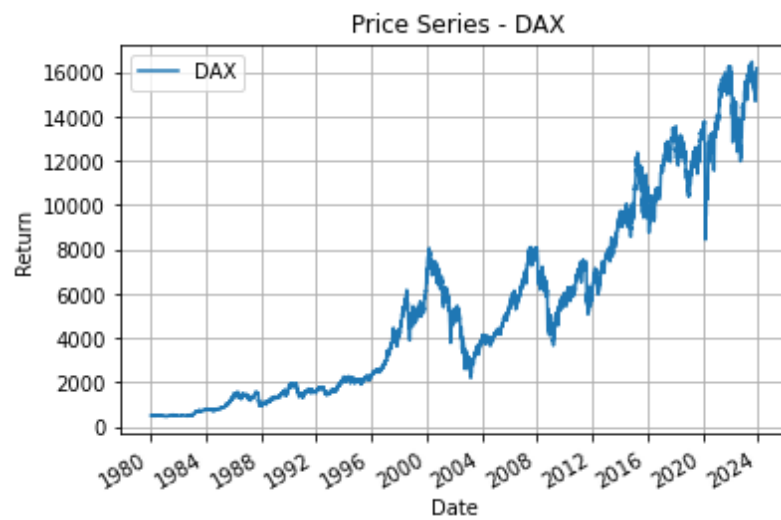
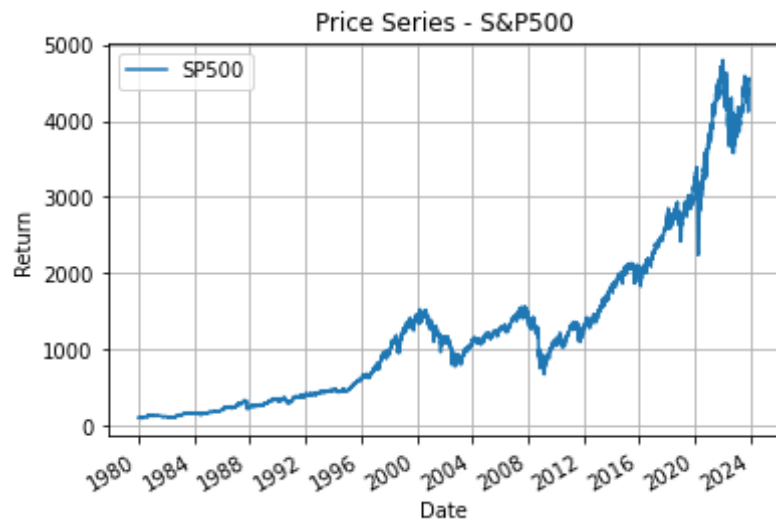
#S&P 500: Constant (mu): 0.0716  
#ASX: Constant (mu): 0.0622  
#DAX: Constant (mu): 0.0780

#  $\mu(t) = \phi(0) + \phi * R(t-1)$ ; where  $\phi(0)$  is the constant

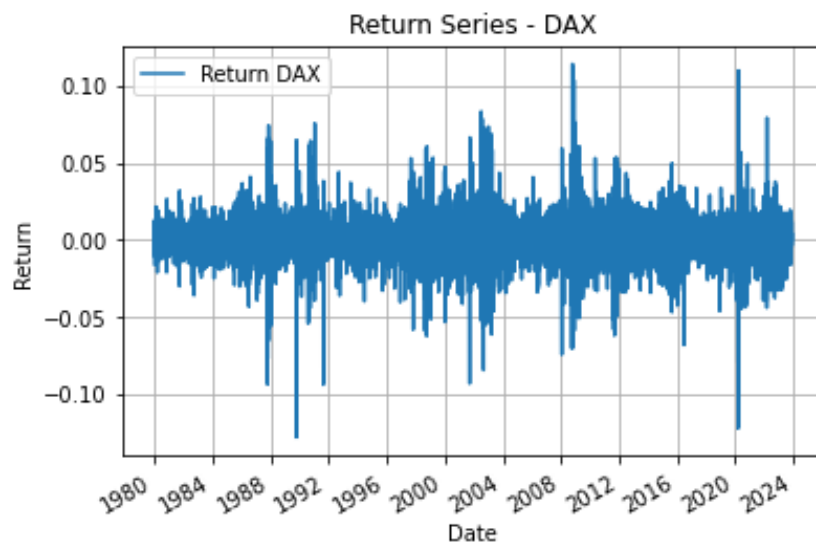
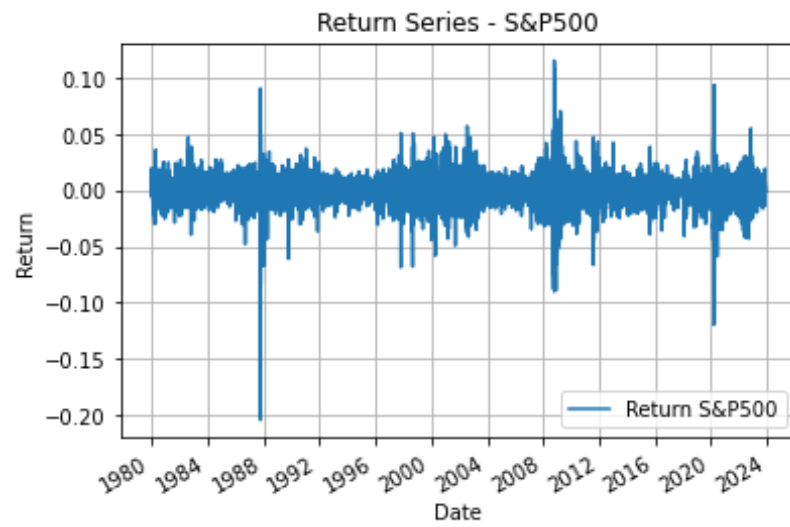
# Therefore, the index with the highest expected return at time t+1 when  
# $R_t = 0$  is the DAX.

#####

## PLOTS OF RAW DATA



## PLOTS OF COMPUTED RETURNS



## CHECKING THE STATIONARITY OF DATA USING DICKEY-FULLER TEST

### Augmented Dickey-Fuller Results

```
=====
Test Statistic      -17.774
P-value             0.000
Lags                 35
-----
```

Trend: Constant

Critical Values: -3.43 (1%), -2.86 (5%), -2.57 (10%)

Null Hypothesis: The process contains a unit root.

Alternative Hypothesis: The process is weakly stationary.

### Augmented Dickey-Fuller Results

```
=====
Test Statistic      -23.474
P-value             0.000
Lags                 17
-----
```

Trend: Constant

Critical Values: -3.43 (1%), -2.86 (5%), -2.57 (10%)

Null Hypothesis: The process contains a unit root.

Alternative Hypothesis: The process is weakly stationary.

### Augmented Dickey-Fuller Results

```
=====
Test Statistic      -39.974
P-value             0.000
Lags                 6
-----
```

Trend: Constant

Critical Values: -3.43 (1%), -2.86 (5%), -2.57 (10%)

Null Hypothesis: The process contains a unit root.

Alternative Hypothesis: The process is weakly stationary.

### MODEL~SP500\_CON

#### Constant Mean - GARCH Model Results

```
=====
Dep. Variable:      Return S&P500  R-squared:      0.000
Mean Model:         Constant Mean  Adj. R-squared:    0.000
Vol Model:          GARCH          Log-Likelihood:    -14207.0
Distribution:        Standardized Student's t  AIC:      28424.1
Method:             Maximum Likelihood  BIC:          28460.4
                   No. Observations:      10679
Date:               Mon, Dec 04 2023  Df Residuals:    10678
Time:               21:13:50  Df Model:              1
                   Mean Model
=====
```

```
=====
      coef  std err      t  P>|t|   95.0% Conf. Int.
-----
mu      0.0716  7.291e-03   9.816  9.638e-23 [5.727e-02,8.585e-02]
=====
```

#### Volatility Model

```
=====
      coef  std err      t  P>|t|   95.0% Conf. Int.
-----
omega     0.0109  2.252e-03   4.857  1.192e-06 [6.525e-03,1.535e-02]
alpha[1]   0.0838  7.619e-03  10.997  3.951e-28 [6.886e-02,9.872e-02]
beta[1]    0.9101  8.056e-03  112.963  0.000 [ 0.894, 0.926]
=====
```

## Distribution

	coef	std err	t	P> t	95.0% Conf. Int.
nu	6.2250	0.395	15.760	5.826e-56	[ 5.451, 6.999]

## MODEL~SP500\_AR

## AR - GARCH Model Results

Dep. Variable:	Return S&P500	R-squared:	0.001
Mean Model:	AR	Adj. R-squared:	0.000
Vol Model:	GARCH	Log-Likelihood:	-14205.2
Distribution:	Standardized Student's t	AIC:	28422.3
Method:	Maximum Likelihood	BIC:	28466.0
	No. Observations:	10678	
Date:	Mon, Dec 04 2023	Df Residuals:	10676
Time:	21:13:50	Df Model:	2
	Mean Model		

	coef	std err	t	P> t	95.0% Conf. Int.
Const	0.0727	7.434e-03	9.775	1.446e-22	[5.810e-02,8.724e-02]
Retu...500[1]	-0.0124	9.316e-03	-1.336	0.182	[-3.070e-02,5.815e-03]
	Volatility Model				

	coef	std err	t	P> t	95.0% Conf. Int.
omega	0.0109	2.247e-03	4.840	1.296e-06	[6.473e-03,1.528e-02]
alpha[1]	0.0836	7.614e-03	10.977	4.945e-28	[6.865e-02,9.850e-02]
beta[1]	0.9103	8.047e-03	113.130	0.000	[ 0.895, 0.926]
	Distribution				

	coef	std err	t	P> t	95.0% Conf. Int.
nu	6.2076	0.394	15.772	4.831e-56	[ 5.436, 6.979]

## MODEL~ASX\_CON

## Constant Mean - GARCH Model Results

Dep. Variable:	Return ASX	R-squared:	0.000
Mean Model:	Constant Mean	Adj. R-squared:	0.000
Vol Model:	GARCH	Log-Likelihood:	-13198.0
Distribution:	Standardized Student's t	AIC:	26406.0
Method:	Maximum Likelihood	BIC:	26442.4
	No. Observations:	10679	
Date:	Mon, Dec 04 2023	Df Residuals:	10678
Time:	21:13:50	Df Model:	1
	Mean Model		

	coef	std err	t	P> t	95.0% Conf. Int.
mu	0.0622	7.154e-03	8.701	3.295e-18	[4.822e-02,7.627e-02]
	Volatility Model				
	coef	std err	t	P> t	95.0% Conf. Int.

omega 0.0193 3.846e-03 5.009 5.474e-07 [1.173e-02,2.680e-02]  
alpha[1] 0.0919 1.034e-02 8.882 6.579e-19 [7.159e-02, 0.112]  
beta[1] 0.8868 1.308e-02 67.816 0.000 [ 0.861, 0.912]

Distribution

=====

	coef	std err	t	P> t	95.0% Conf. Int.
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nu	7.5337	0.631	11.941	7.272e-33	[ 6.297, 8.770]
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=====

#### MODEL~ASX\_AR

##### AR - GARCH Model Results

=====

Dep. Variable:	Return ASX	R-squared:	-0.002
Mean Model:	AR	Adj. R-squared:	-0.002
Vol Model:	GARCH	Log-Likelihood:	-13177.5
Distribution:	Standardized Student's t	AIC:	26366.9
Method:	Maximum Likelihood	BIC:	26410.6
	No. Observations:	10678	
Date:	Mon, Dec 04 2023	Df Residuals:	10676
Time:	21:13:50	Df Model:	2
	Mean Model		

=====

	coef	std err	t	P> t	95.0% Conf. Int.
--	------	---------	---	------	------------------

-----

Const	0.0582	7.160e-03	8.123	4.549e-16	[4.413e-02,7.219e-02]
Return ASX[1]	0.0621	1.041e-02	5.963	2.475e-09	[4.166e-02,8.244e-02]

Volatility Model

=====

	coef	std err	t	P> t	95.0% Conf. Int.
--	------	---------	---	------	------------------

-----

omega	0.0204	4.057e-03	5.026	5.012e-07	[1.244e-02,2.834e-02]
alpha[1]	0.0955	1.074e-02	8.886	6.328e-19	[7.440e-02, 0.117]
beta[1]	0.8819	1.370e-02	64.351	0.000	[ 0.855, 0.909]

Distribution

=====

	coef	std err	t	P> t	95.0% Conf. Int.
--	------	---------	---	------	------------------

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nu	7.6574	0.655	11.690	1.438e-31	[ 6.374, 8.941]
----	--------	-------	--------	-----------	-----------------

=====

#### MODEL~DAX\_CON

##### Constant Mean - GARCH Model Results

=====

Dep. Variable:	Return DAX	R-squared:	0.000
Mean Model:	Constant Mean	Adj. R-squared:	0.000
Vol Model:	GARCH	Log-Likelihood:	-16437.5
Distribution:	Standardized Student's t	AIC:	32884.9
Method:	Maximum Likelihood	BIC:	32921.3
	No. Observations:	10679	
Date:	Mon, Dec 04 2023	Df Residuals:	10678
Time:	21:13:51	Df Model:	1
	Mean Model		

=====

	coef	std err	t	P> t	95.0% Conf. Int.
--	------	---------	---	------	------------------

-----

mu	0.0780	9.214e-03	8.465	2.559e-17	[5.994e-02,9.606e-02]
----	--------	-----------	-------	-----------	-----------------------

Volatility Model

=====

	coef	std err	t	P> t	95.0% Conf. Int.
--	------	---------	---	------	------------------

-----					
omega	0.0189	3.464e-03	5.443	5.237e-08	[1.207e-02,2.564e-02]
alpha[1]	0.0854	8.157e-03	10.472	1.167e-25	[6.943e-02, 0.101]
beta[1]	0.9056	8.710e-03	103.962	0.000	[ 0.888, 0.923]
Distribution					
=====					
	coef	std err	t	P> t	95.0% Conf. Int.
-----					
nu	7.3782	0.582	12.671	8.576e-37	[ 6.237, 8.519]
=====					

MODEL~DAX\_AR

AR - GARCH Model Results					
=====					
Dep. Variable:	Return DAX	R-squared:	-0.001		
Mean Model:	AR	Adj. R-squared:	-0.001		
Vol Model:	GARCH	Log-Likelihood:	-16434.0		
Distribution:	Standardized Student's t	AIC:	32880.1		
Method:	Maximum Likelihood	BIC:	32923.7		
	No. Observations:	10678			
Date:	Mon, Dec 04 2023	Df Residuals:	10676		
Time:	21:13:51	Df Model:	2		
	Mean Model				
=====					
	coef	std err	t	P> t	95.0% Conf. Int.
-----					
Const	0.0777	9.336e-03	8.320	8.832e-17	[5.937e-02,9.597e-02]
Return DAX[1]	6.4312e-03	9.456e-03	0.680	0.496	[-1.210e-02,2.496e-02]
Volatility Model					
=====					
	coef	std err	t	P> t	95.0% Conf. Int.
-----					
omega	0.0190	3.469e-03	5.472	4.461e-08	[1.218e-02,2.578e-02]
alpha[1]	0.0857	8.173e-03	10.489	9.701e-26	[6.971e-02, 0.102]
beta[1]	0.9052	8.722e-03	103.788	0.000	[ 0.888, 0.922]
Distribution					
=====					
	coef	std err	t	P> t	95.0% Conf. Int.
-----					
nu	7.3911	0.585	12.627	1.496e-36	[ 6.244, 8.538]
=====					

Covariance estimator: robust

TESTING RESULTS FOR SERIAL CORRELATION

For sp500\_con, the DW statistics is 2.107611179288561  
For sp500\_ar, the DW statistics is 2.0829317018731257

For asx\_con, the DW statistics is 1.949987931738237  
For asx\_ar, the DW statistics is 2.075134960788002

For dax\_con, the DW statistics is 2.0048652957247675  
For dax\_ar, the DW statistics is 2.017352945329236

TESTING RESULTS FOR LIKELIHOOD RATIO

The Likelihood ratio test statistics is 3.761394375218515  
The corresponding p-value is 0.05244882325629066



The Likelihood ratio test statistics is 41.08855085916002

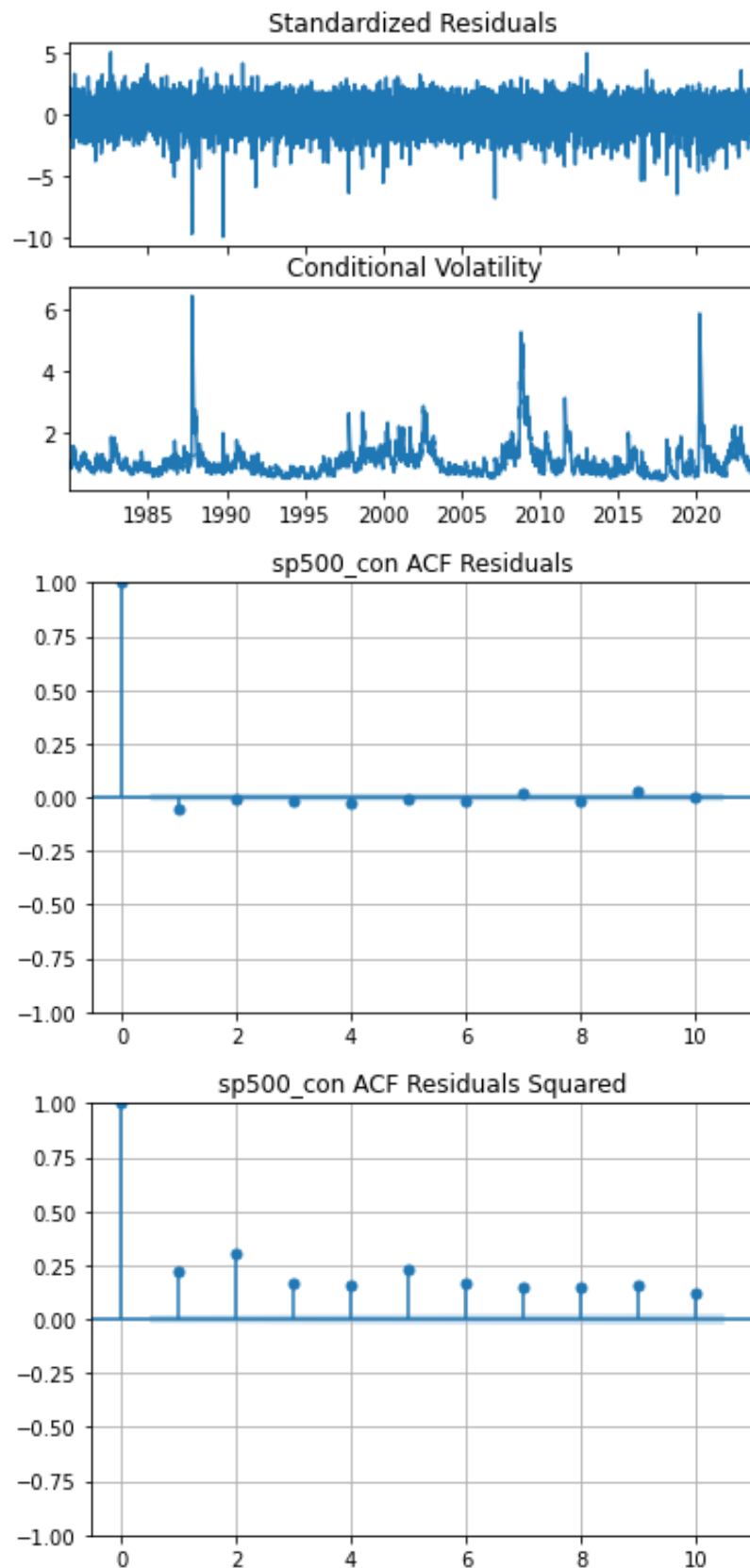
The corresponding p-value is 1.4548607233366306e-10

The Likelihood ratio test statistics is 6.85261521780194

The corresponding p-value is 0.008851161779641091

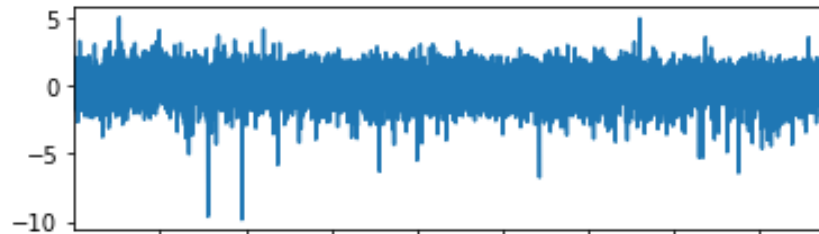
### PLOTS OF RESIDUALS AND SQUARED RESIDUALS

MODEL~ SP500\_CON

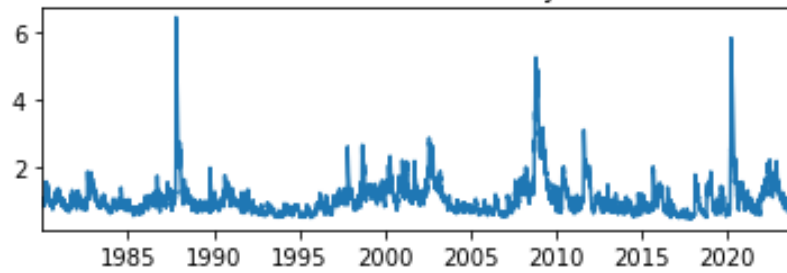


MODEL ~ SP500\_AR

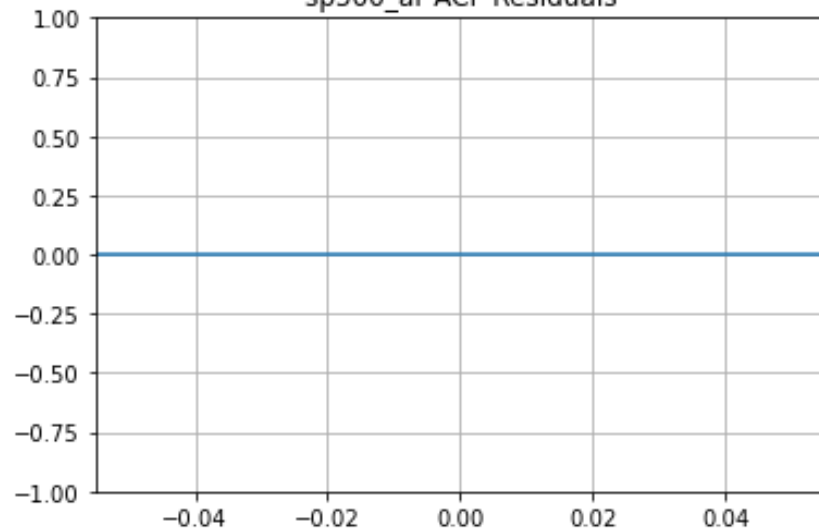
Standardized Residuals



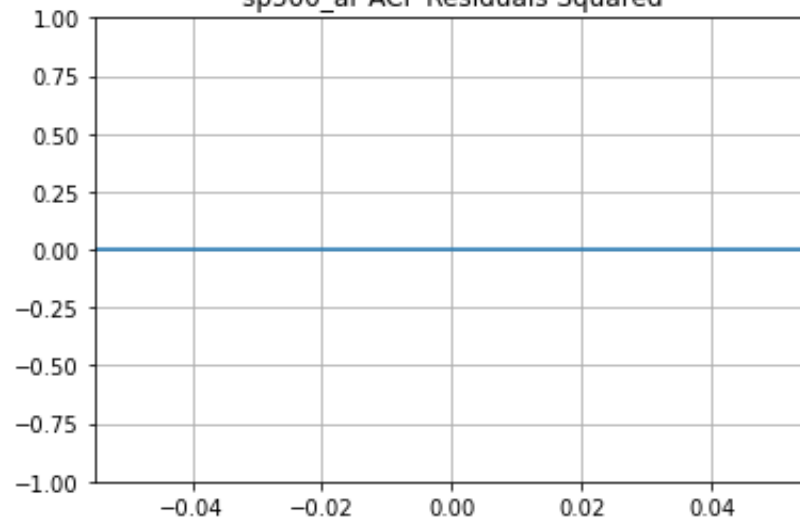
Conditional Volatility



sp500\_ar ACF Residuals

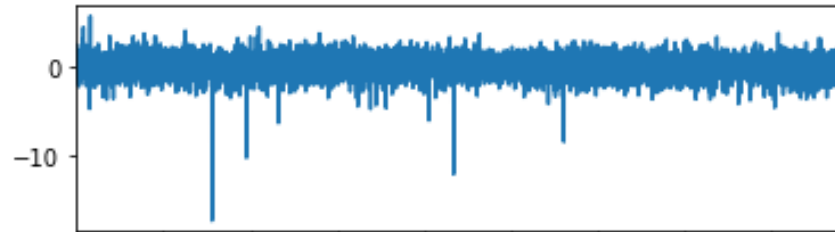


sp500\_ar ACF Residuals Squared

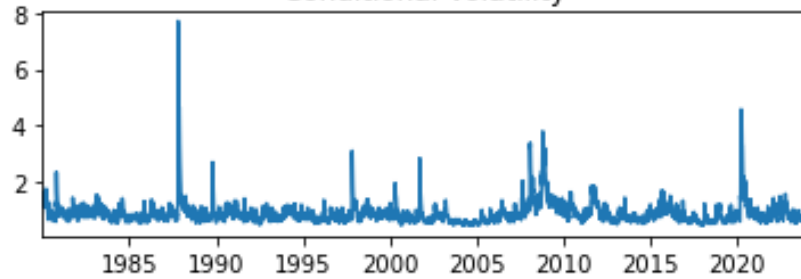


MODEL ~ ASX\_CON

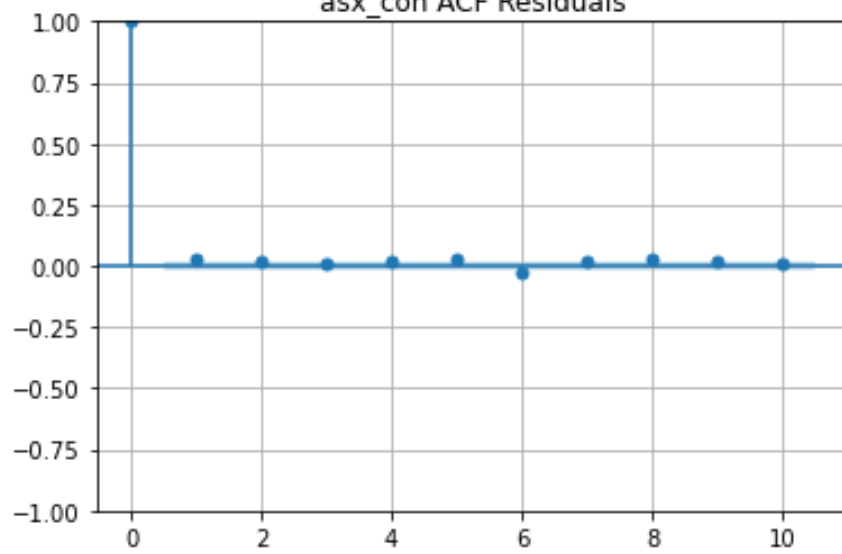
Standardized Residuals



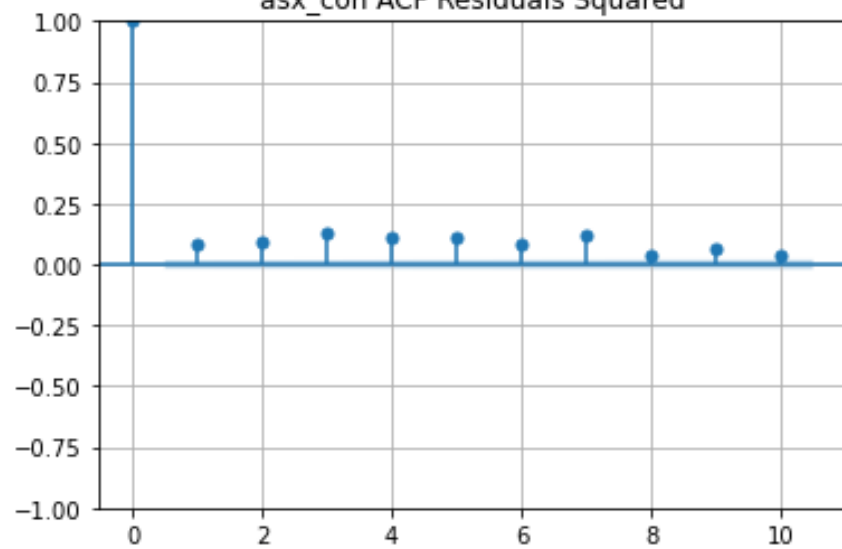
Conditional Volatility



asx\_con ACF Residuals

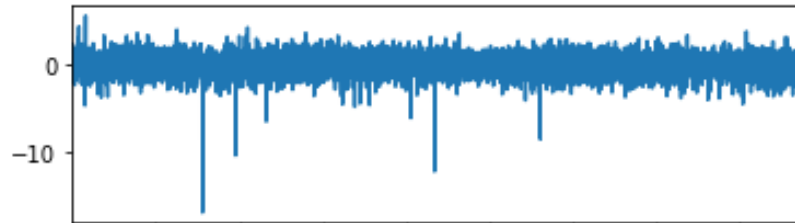


asx\_con ACF Residuals Squared

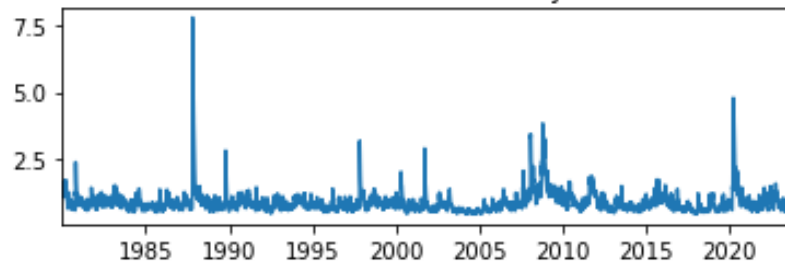


MODEL ~ ASX\_AR

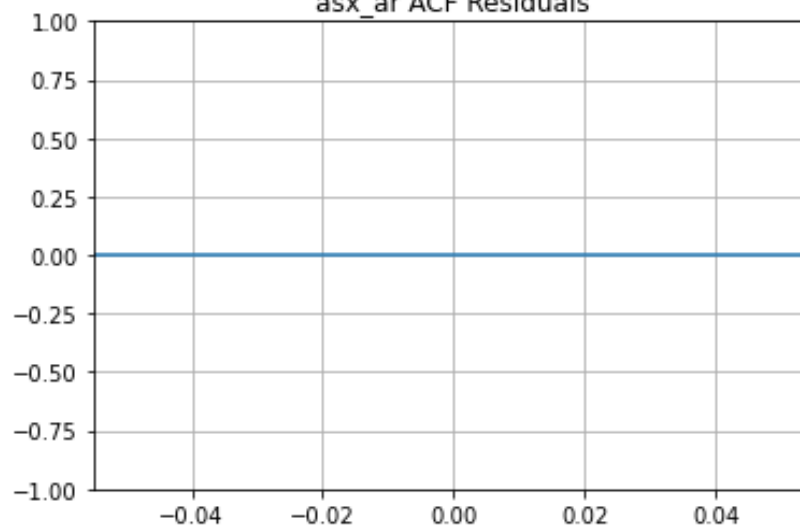
Standardized Residuals



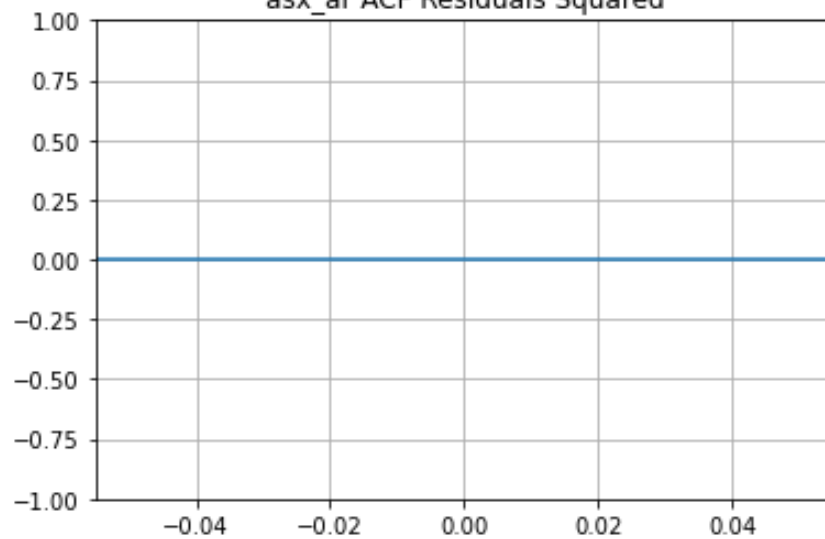
Conditional Volatility



asx\_ar ACF Residuals

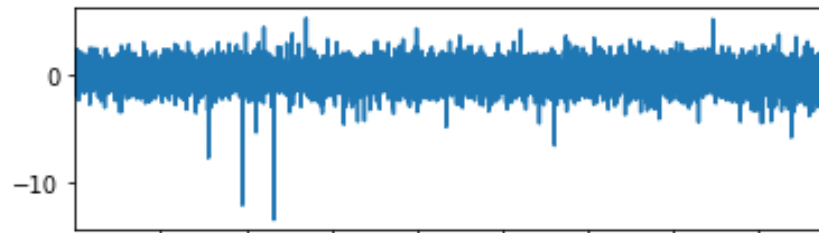


asx\_ar ACF Residuals Squared

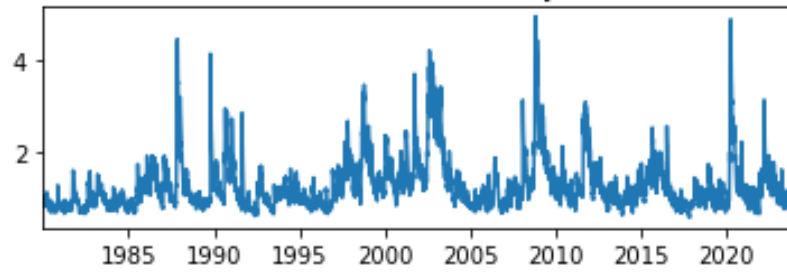


MODEL ~ DAX\_CON

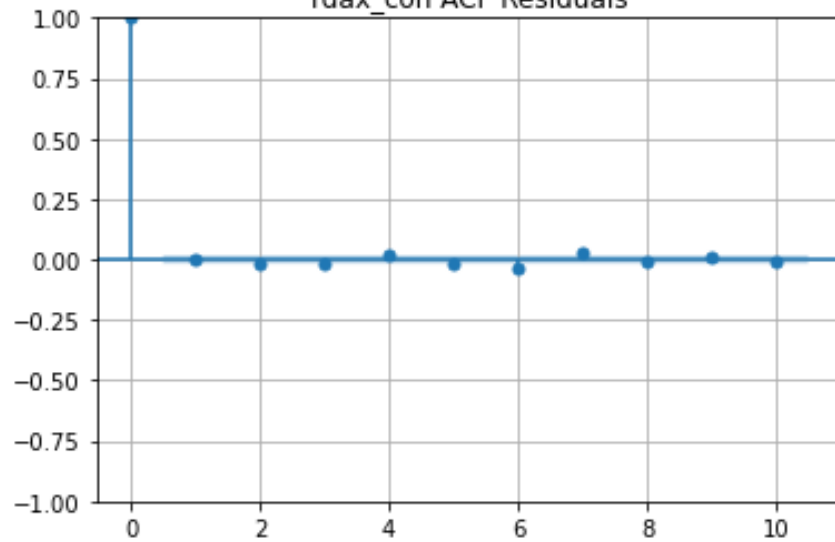
Standardized Residuals



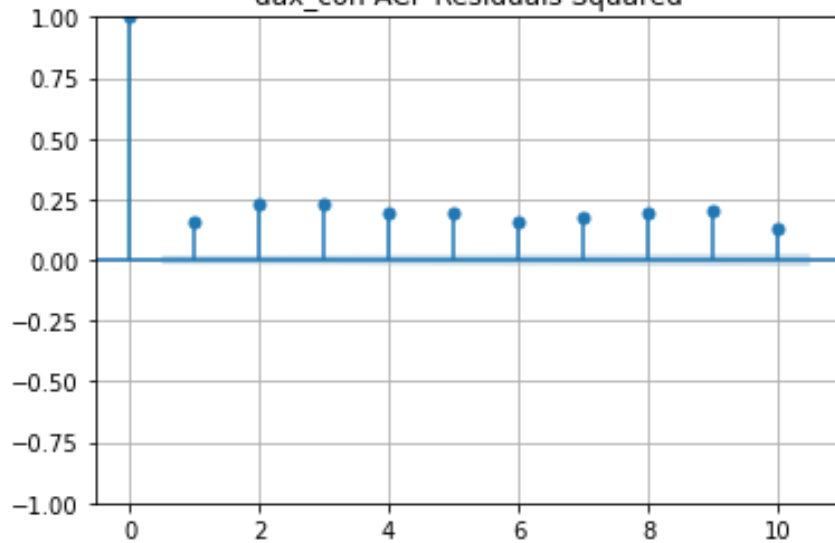
Conditional Volatility



rdax\_con ACF Residuals

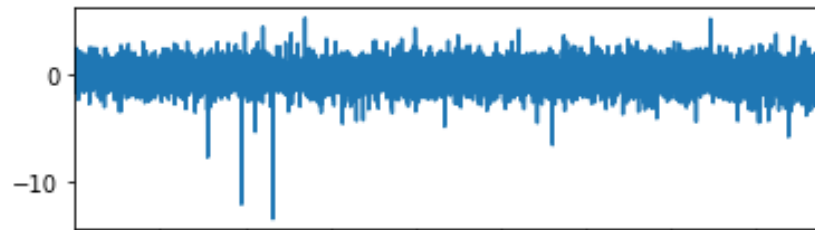


dax\_con ACF Residuals Squared

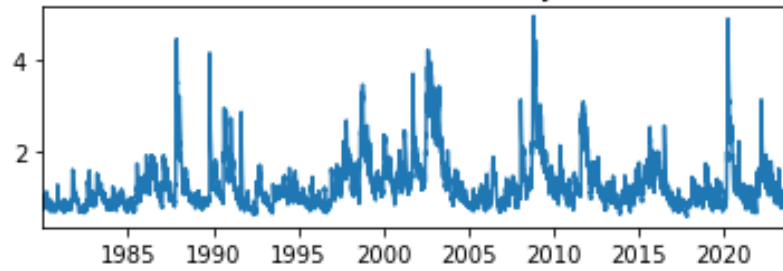


MODEL ~ DAX\_AR

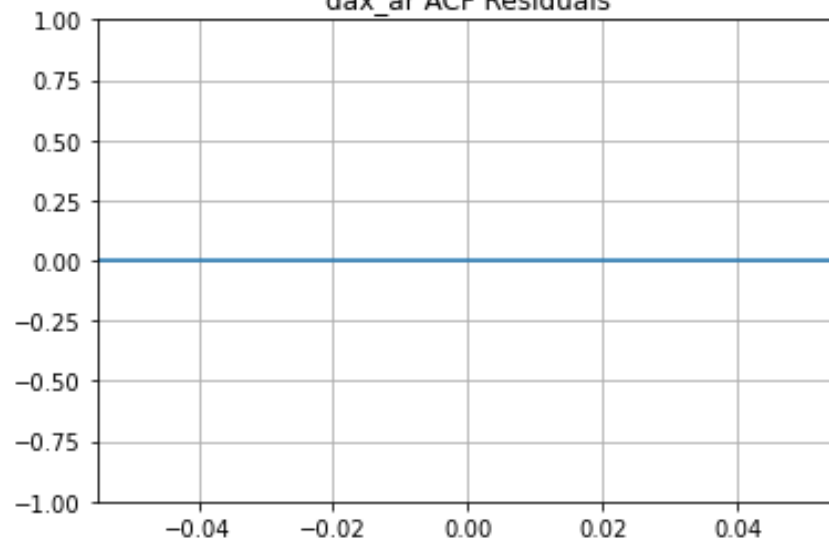
Standardized Residuals



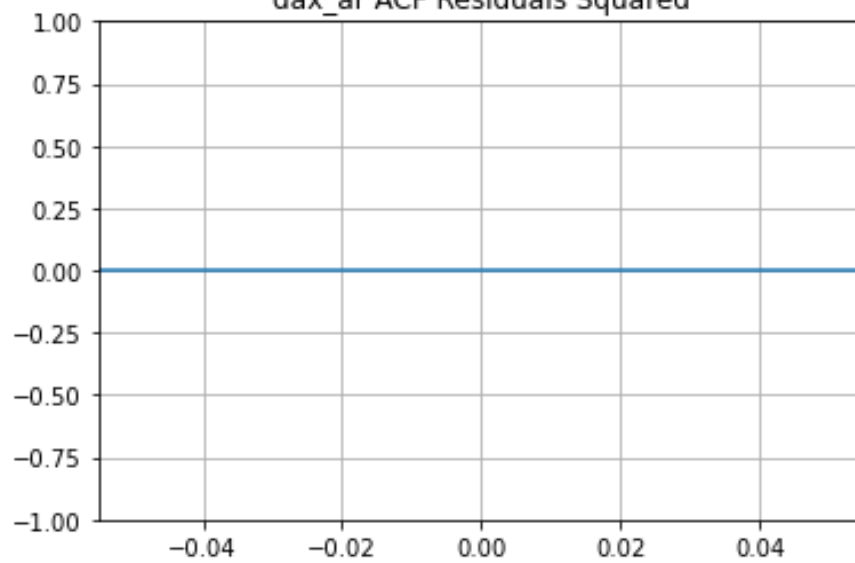
Conditional Volatility



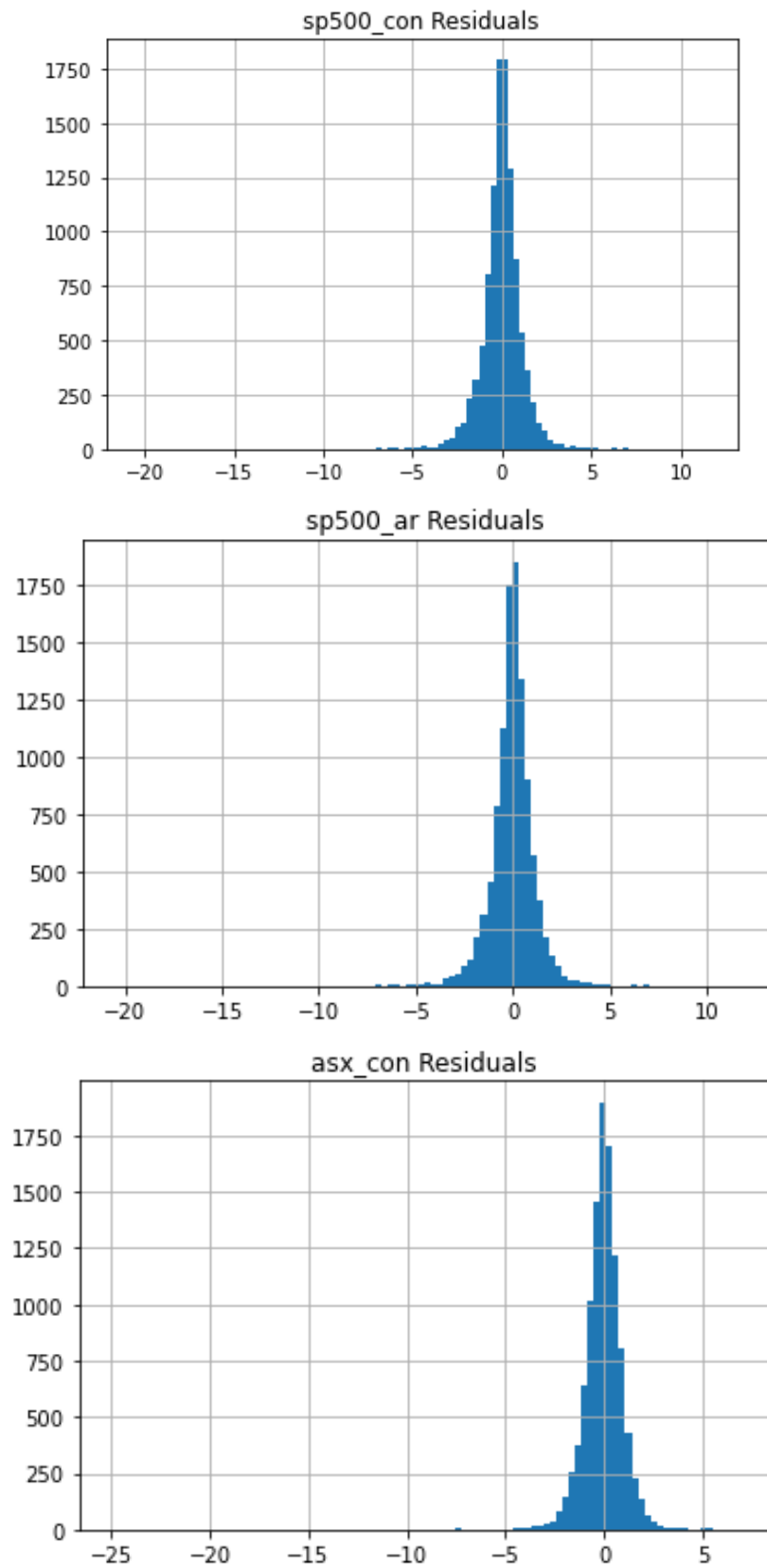
dax\_ar ACF Residuals

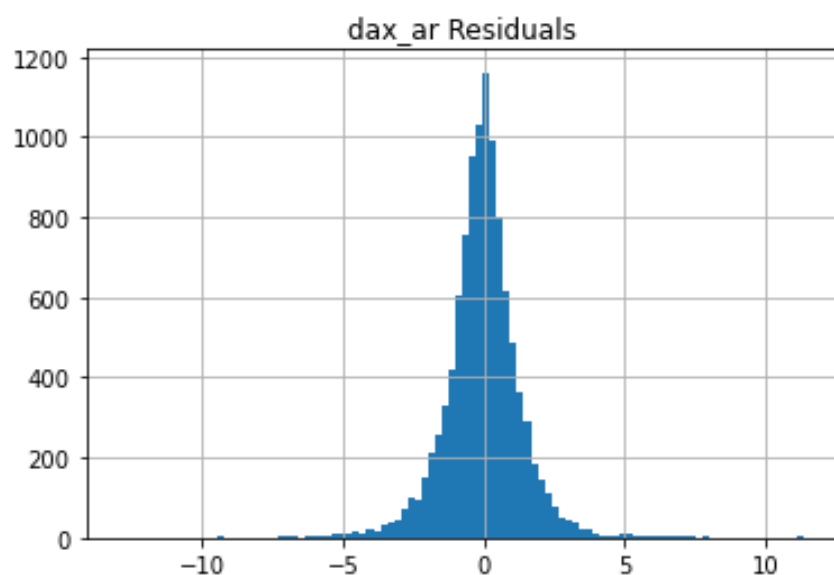
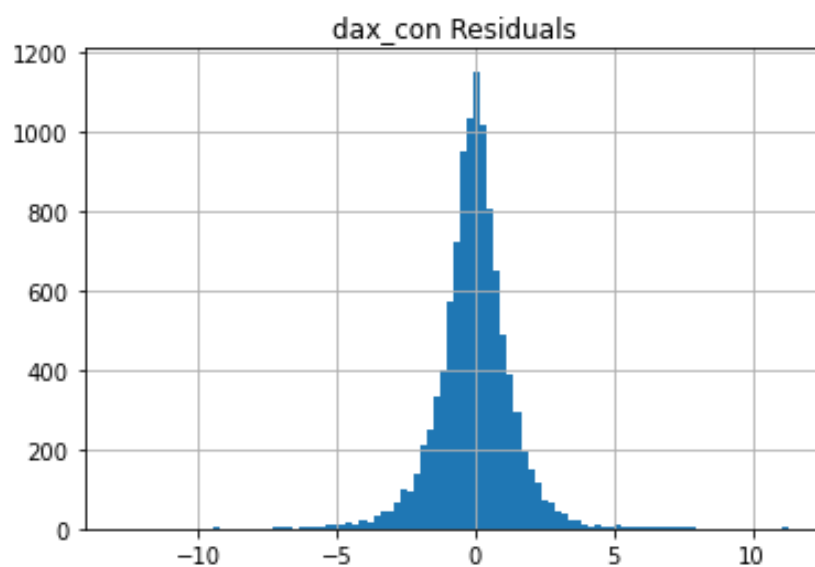
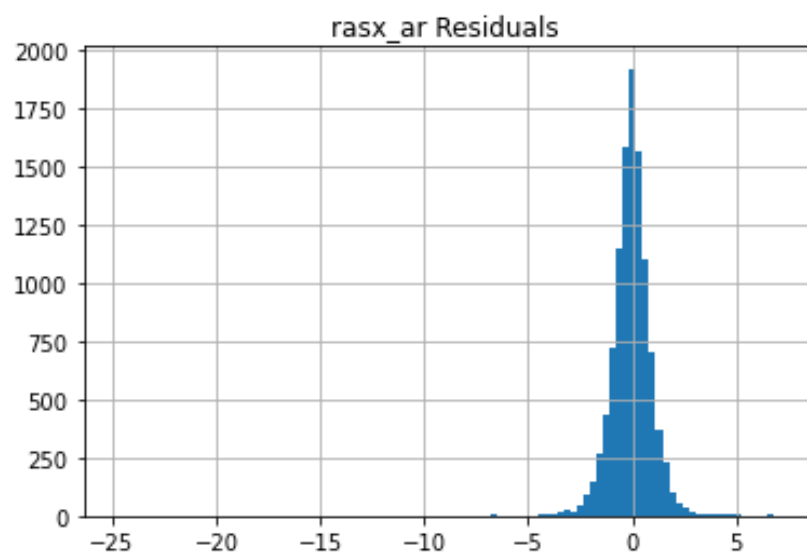


dax\_ar ACF Residuals Squared

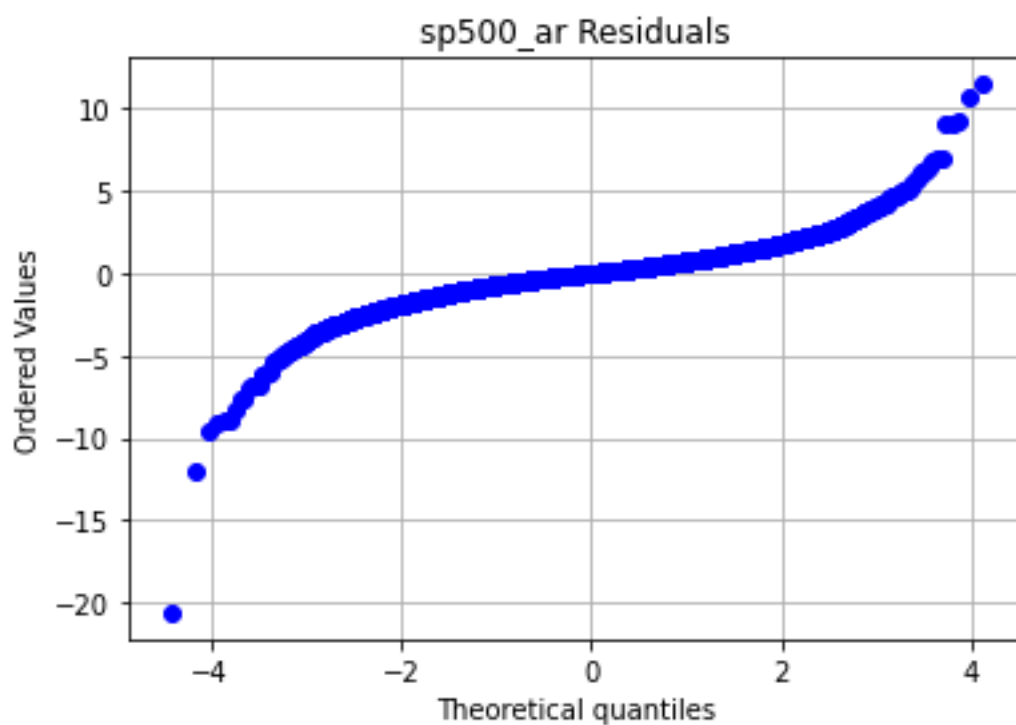
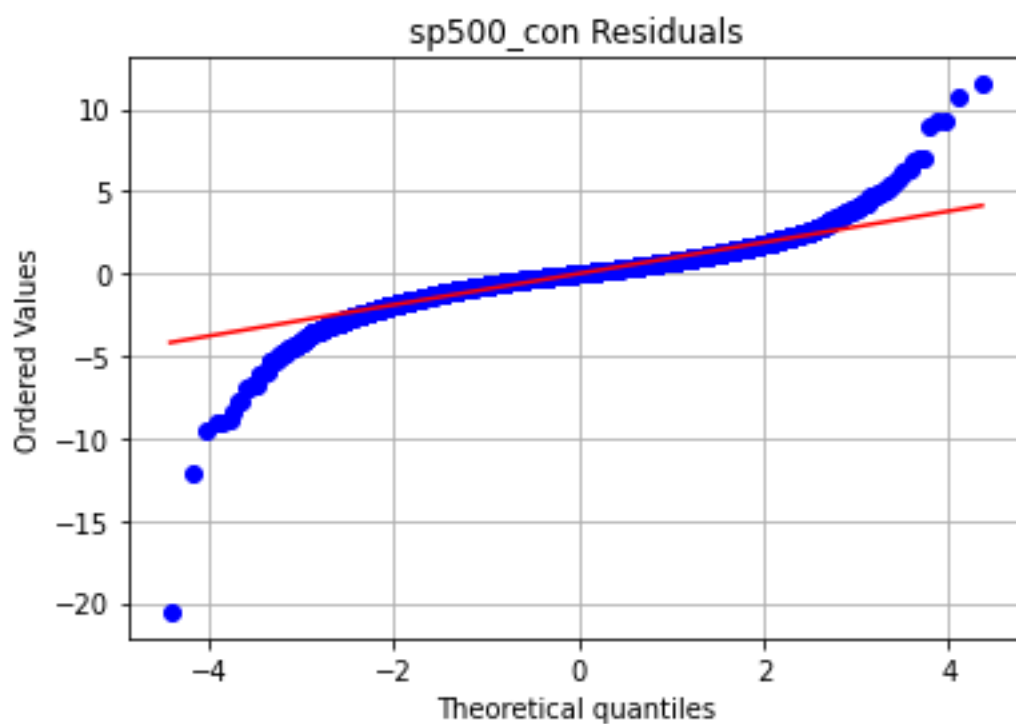


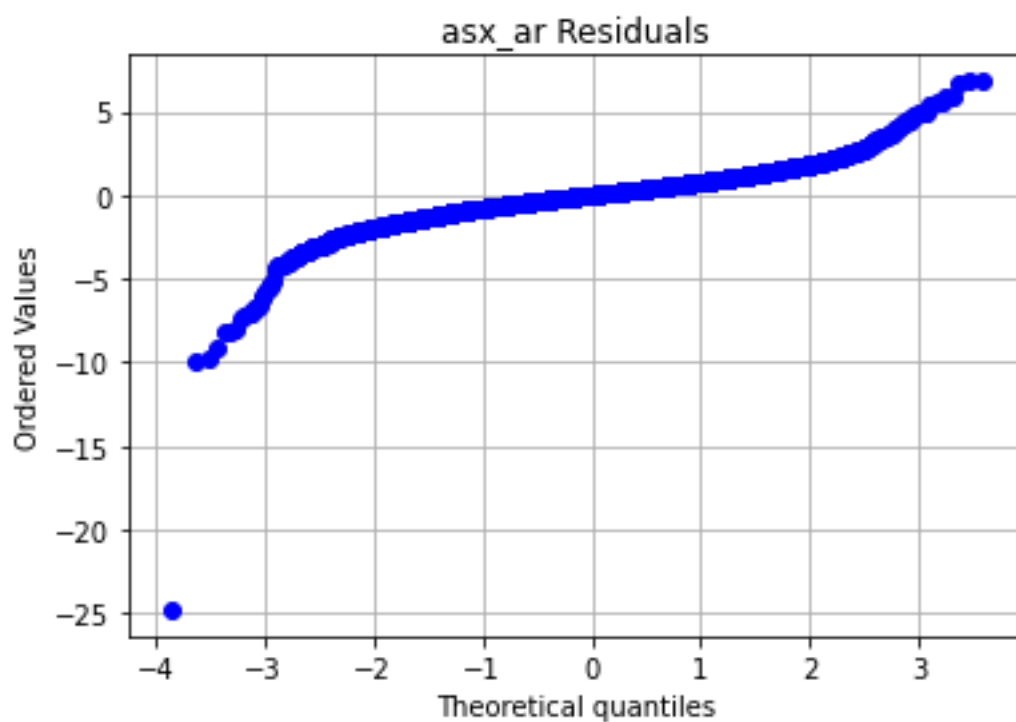
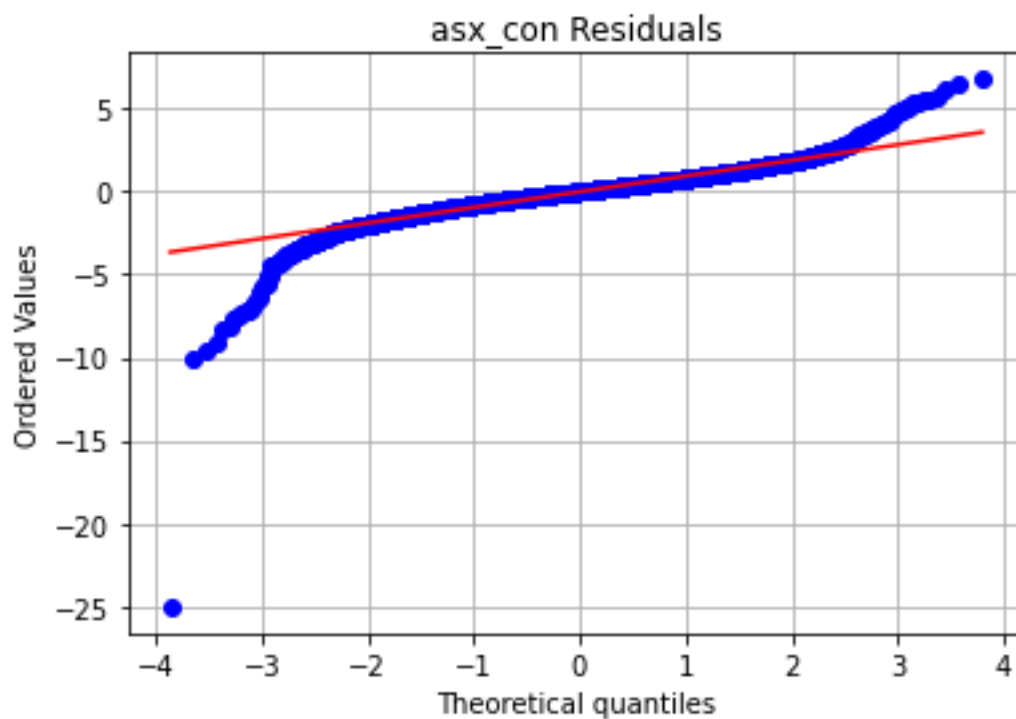
## PLOTTING HISTOGRAMS AND QQ-PLOTS FOR EACH MODEL

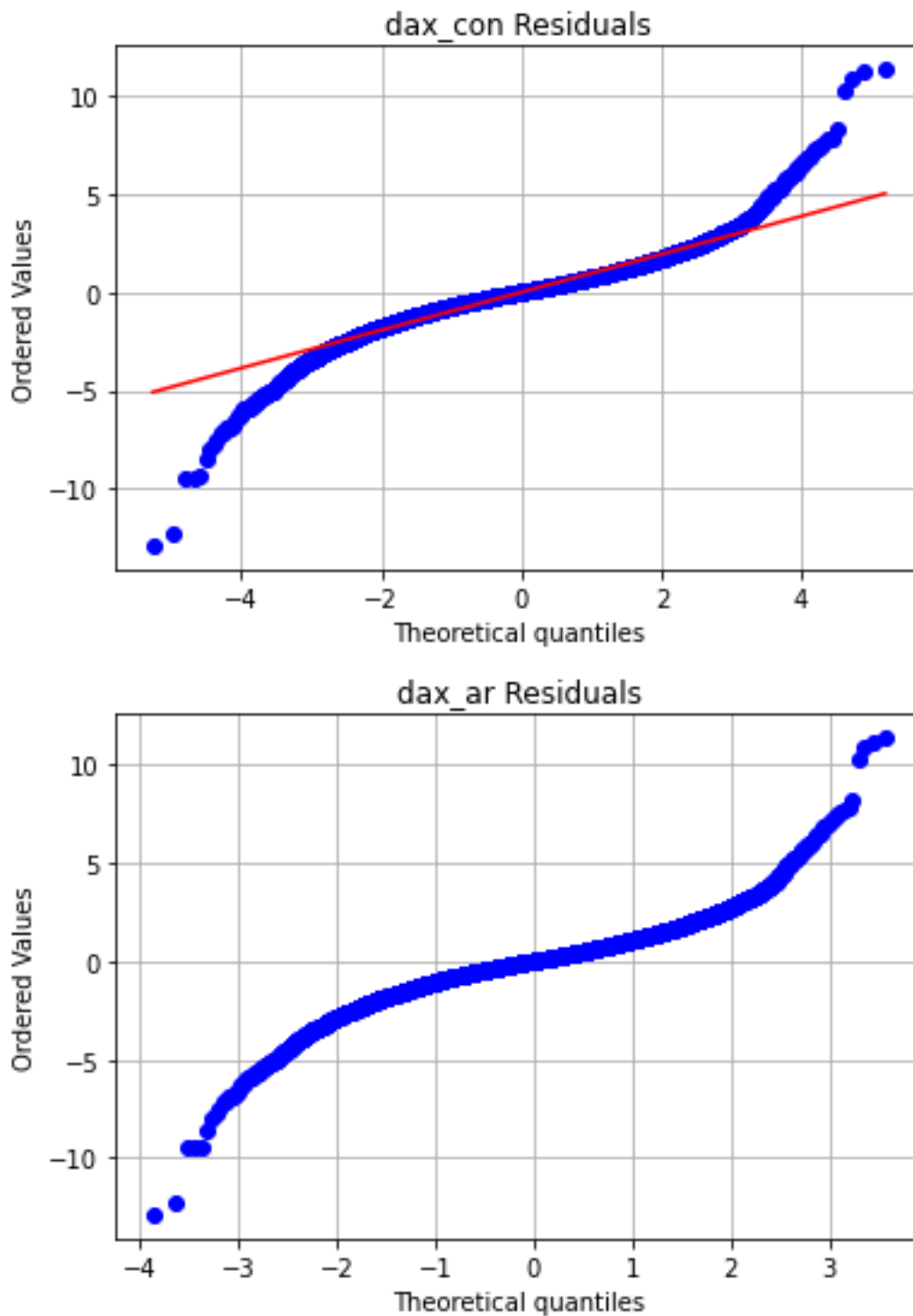












### ANSWERS

#### ANS 1:

summary Of Findings for each index:

for all three naming S&P500, ASX and DAX:

- ADF Test: The process is weakly stationary.
- GARCH Coefficients: The coefficients seem reasonable.
- LRT: The AR-GARCH model is preferred over the constant mean GARCH model.

The GARCH(1,1) models for the ASX and DAX indices seems to be an appropriate fit.

For the S&P 500 index, the GARCH(1,1) model seemed to provide some inconclusive findings since the p-value from the log-likelihood test just barely exceeded the 0.05 benchmark necessary for a 5% confidence interval when comparing the constant model with the complex model.

However, this would not be the case if we changed our guidelines to a 6% confidence level.

The findings led to the conclusion that the squared residuals from lagged models for all three indices (sp500\_ar, asx\_ar, and dax\_ar) consistently fall within acceptable ranges. This implies the effective capture of time-varying variance effects, strongly suggesting the presence and capture of GARCH effects.

**ANS 2:**

$\mu$  at  $t$  is dependent on returns at time  $t-1$  so  $\mu$  at  $t+1$  is dependent on returns from time  $t$  coefficient of return lag is irrelevant since it was established that return at  $t$  is 0 in the question.

- S&P 500: Constant ( $\mu$ ): 0.0716
- ASX: Constant ( $\mu$ ): 0.0622
- DAX: Constant ( $\mu$ ): 0.0780

$\mu(t) = \phi(0) + \phi * R(t-1)$ ; where  $\phi(0)$  is the constant

Therefore, the index with the highest expected return at time  $t+1$  when  $R_t = 0$  is the DAX.