Mushahid Hussain

Report









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TIME SERIRES ANALYSIS:

CRYPTOCURRENCY INFLUENCE
ON TRADITIONAL STOCK
MARKET INDICES: A
COMPARATIVE MULTIPLE
REGRESSION ANALYSIS OF
VOLATILITY AND
CORRELATION





BASIC ECONOMETRICS

Presented by:

PRESENTED TO: SIR MANSOOR MUSHTAQ

///////

MUSSAB BIN UMAIR (221-9809) RAED ALV (221-7460)I

SEMESTER PROJECT





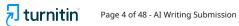


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Abstract

By examining the cross-correlations between the two most valuable cryptocurrencies, Bitcoin (BTC) and Ethereum (ETH), this research paper focuses on how cryptocurrencies affect traditional stock market indices. However, Litecoin (LTC) has also been a significant factor in the study. By examining the increasing potential for institutional involvement in the cryptocurrency region, we hope to analyse the impact of the cryptocurrency field's growing capitalization on the stock market as a whole.

This study looks at the long-term relationship between cryptocurrencies and the traditional stock market. As exchange rates have been rising and falling over time, and as institutional involvement in this industry has grown, we have looked at various models that explain and describe the relationships between these variables that help us with our research. Since more and more individuals and businesses are investing in cryptocurrencies, institutional involvement has been a significant factor. In the near future, we can also anticipate cryptocurrencies playing a significant role in the economy due to their rapid, exponential growth. Our analysis looks at how cryptocurrencies have the potential to be a significant turning point and how institutional investment in this space will impact the conventional stock market as well.

INTRODUCTION

The data that we have sourced from investing.com, is of the following:

- 1. NASDAQ 100 (NDX) Traditional stock market representation
- 2. BITCOIN (BTC) The largest cryptocurrency in the market
- 3. ETHEREUM (ETH) A leading blockchain platform with smart contracts
- 4. LITECOIN (LTC) One of the earliest Bitcoin alternatives
- 5. Grayscale Bitcoin ETF (GBTC) One of the biggest crypto ETFs in the world.
- 6. EUR/USD The exchange rate of Euros and US Dollar.

With a timeframe spanning from May 2017 to October 2024, the data offers a solid framework for examining the potential effects of cryptocurrencies on the conventional stock





market over time. The data's weekly frequency guarantees that long-term trends are prioritized while short-term market noise is eliminated.

Understanding the relationship between cryptocurrencies and the conventional stock market is the aim of this investigation, which also looks at whether cryptocurrencies have a discernible and substantial influence on stock market indices such as the NASDAQ 100. This study investigates the interactions and influences that have occurred over time between the cryptocurrency environment and the conventional stock market.

VARIABLE EXPLANATION:

Functional Form of the Model:

The econometric model can be expressed as:

 $NASDAQ = \beta 0 + \beta 1(BTC) + \beta 2(ETH) + \beta 3(LTC) + \beta 4(DOGE) + U$

Where:

- NASDAQ: Dependent variable, representing the stock market index.
- β0:Constant term (intercept), capturing the baseline level of the NASDAQ index when all other variables are zero.
- **BTC:** Explanatory variable for Bitcoin prices.
- **ETH:** Explanatory variable for Ethereum prices.
- LTC: Explanatory variable for Litecoin prices.
- **DOGE**: Explanatory variable for Dogecoin prices.





- β1,β2,β3,β4: Coefficients of the explanatory variables, representing the marginal impact of each cryptocurrency on the NASDAQ index.
- U: Error term, accounting for unobserved factors affecting NASDAQ.

Hypothesis

- **Null Hypothesis** (Ho: The explanatory variables (BTC, ETH, LTC, DOGE) have no significant impact on the NASDAQ index.
- **H0**: βBTC=βETH=βLTC=βDOGE=0
- Alternative Hypothesis H1: At least one of the explanatory variables has a significant impact on the NASDAQ index.
- **H1**:At least one $\beta \neq 0$

DESCRIPTIVE STATISTICS:

In this section of the report, we discuss the descriptive statistics of the variables taken in our study and analyse all of them to see if they are acceptable or meet set standard which is, Normality.

Normality refers to whether the data follows a normal distribution (bell shaped curve). It also involves checking if the data has properties like symmetry around the mean and specific kurtosis and skewness values. Normality is important to further proceed with our analysis and moving on to liner regression and other models.

For better understanding and results being consistent, we take logs of the variables and use all the information in that order so we can achieve optimal skewness and kurtosis, which infers that the variables are normal and are a fit for this model for further analysis.





| Name | Mean | Median | Std. Deviation | Skewness | Kurtosis |
|---------|----------|----------|----------------|-----------|----------|
| NDX | 9.285207 | 9.370836 | 0.370466 | -0.113446 | 1.690794 |
| BTC | 9.745074 | 9.853644 | 0.925030 | -0.160306 | 1.856700 |
| ETH | 6.693582 | 7.105831 | 1.144797 | -0.224778 | 1.551545 |
| LTC | 4.389632 | 4.288105 | 0.511719 | 0.502912 | 2.832180 |
| GBTC | 2.803711 | 2.656405 | 0.733260 | 0.133763 | 2.064442 |
| EUR/USD | 0.116854 | 0.114266 | 0.050394 | -0.238765 | 2.840106 |

Interpretation of Mean vs. Median:

A near mean and median indicate that the data is reasonably symmetrical. Skewness is indicated by larger discrepancies.

The mean and median for all of the variables (NDX, BTC, ETH, GBTC, EUR/USD and LTC) are rather close, suggesting moderate symmetry.

For Skewness:

- Skewness = 0: A distribution that is perfectly symmetrical.
- Skewness < 0: Left-skewed, meaning the left tail is longer.
- Skewness > 0: Right-skewed, meaning the right tail is longer

Here,

- NDX: Shows a slight inclination toward lower values, with a slightly left-skewed value of -0.113.
- BTC: Shows a slight asymmetry, with a slightly left-skewed value of -0.160.
- ETH: Displays a longer left tail and is moderately left-skewed (-0.224).
- LTC: More extreme high values are present because it is right-skewed (0.503).
- GBTC: Near symmetry, as indicated by being quite close to 0 (0.133).
- EUR/USD: Displays a longer left tail and is moderately left-skewed (-0.238765).

For Kurtosis:





- Kurtosis = 3: Normal distribution (mesokurtic).
- Kurtosis > 3: Leptokurtic (sharper peak, heavier tails).
- Kurtosis < 3: Platykurtic (flatter peak, lighter tails)

Here,

- NDX: 1.69: Platykurtic, which denotes a flatter peak and lighter tails.
- Bitcoin: Platykurtic, 1.86, with comparable traits.
- ETH: 1.55: Platykurtic, meaning tails are lighter.
- LTC: 2.83: This value is close to normal but leptokurtic, indicating slightly heavier tails.
- GBTC: 2.840106: This value is close to normal but leptokurtic, indicating slightly heavier tails.

So far, these are the descriptive statistics for every variable in this model.

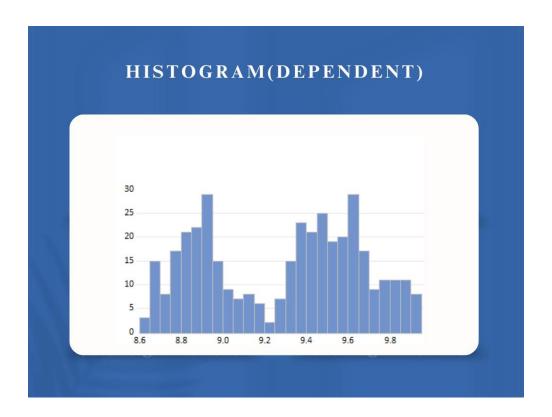
Moving on we have the histogram tables for all the variables in the next figure.

HISTOGRAM TABLES:

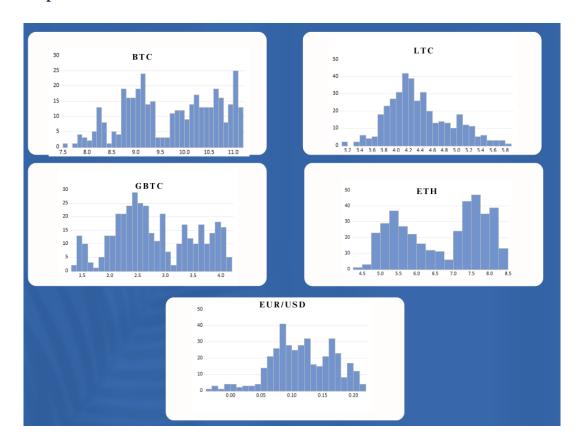
Dependent Variable







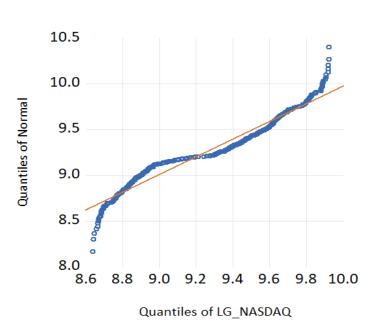
Independent Variables:



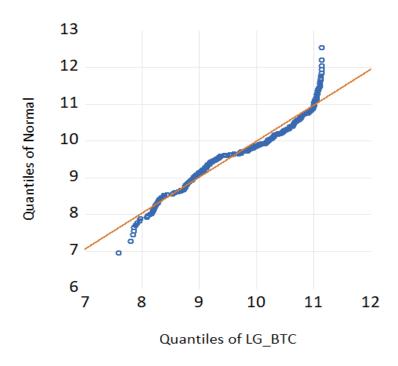


Q-Q PLOTS:

NASDAQ

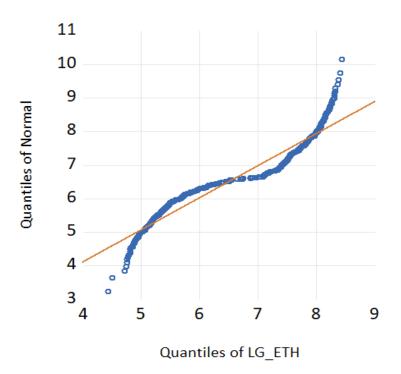


BITCOIN



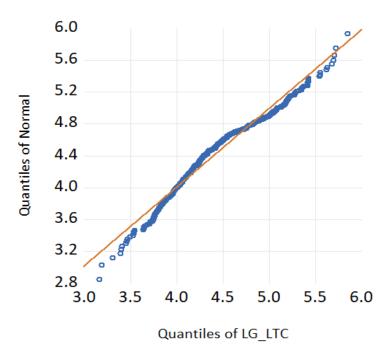


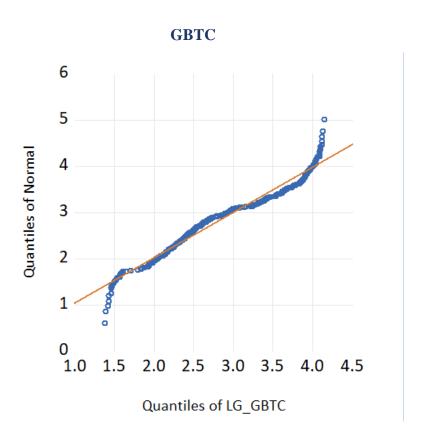
ETHEREUM





LITECOIN

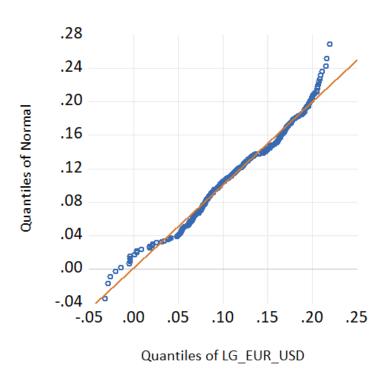








EUR/USD

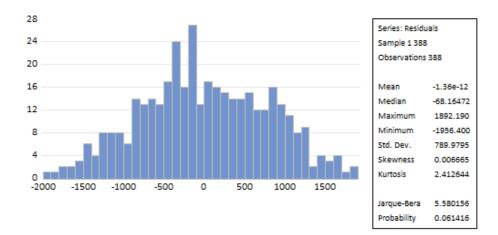


JARQUE-BERA TEST:

Now we move on to the Jarque-Bera test, for this, we take estimate our model and then take the residuals and check their normality to see if they are normally distributed.

We go into E-views and estimate our model.





After estimating our model, we go into the descriptive stats and look the Jarque-Bera tabulated value and the probability of the residuals.

Here, the following two things need to happen to so that we can assess the following hypothesis.

Ho: The residuals are normally distributed.

H₁: The residuals are not normally distributed.

So, we look into the values given and compare them with what is a set standard, for the Jarque-Bera test we need the Jarque-Bera factor to be less than 1.76 and the probability value needs to be greater than 0.05.

In this case, we come to the conclusion that the residuals are in fact normal and we fail to reject the null hypothesis, which states, the residuals are normally distributed, and they follow a bell-shaped curve.

CORRELATION ANALYSIS:





CORRELATION MATRIX:

| | NASDAQ | LG_BTC | ETH | LTC | GBTC | EUR_USD |
|--------|-----------|-----------|-----------|-----------|-----------|-----------|
| NASDAQ | 1.000000 | 0.925418 | 0.859858 | 0.163167 | 0.856291 | -0.329473 |
| LG_BTC | 0.925418 | 1.000000 | 0.895691 | 0.437440 | 0.724788 | -0.219723 |
| ETH | 0.859858 | 0.895691 | 1.000000 | 0.448356 | 0.641422 | -0.192364 |
| LTC | 0.163167 | 0.437440 | 0.448356 | 1.000000 | -0.135629 | 0.496422 |
| GBTC | 0.856291 | 0.724788 | 0.641422 | -0.135629 | 1.000000 | -0.576778 |
| EUR | -0.329473 | -0.219723 | -0.192364 | 0.496422 | -0.576778 | 1.000000 |

CORRELATION MATRIX EXPLAINED:

For correlation, it explains the linear relationship between the variables and this is indicated by the values of the coefficients and it is ranged between -1 to 1 where:

- +1: Perfect positive correlation (as one variable increases, the other increases proportionally).
- **0**: No correlation.
- -1: Perfect negative correlation (as one variable increases, the other decreases proportionally).

CORRELATION BETWEEN VARIABLES:

Now, after getting out results we will be explaining the correlation matrix and the values obtained from it as follows:

NASDAQ vs. CRYPTOCURRENCIES CORRELATION:

NASDAQ & BTC (0.925418):

There is a substantial positive link between the movements of the NASDAQ 100 and the weekly logarithmic returns of Bitcoin. This implies that there is a





substantial correlation between the price trends of Bitcoin and the success of the stock market.

• NASDAQ & ETH (0.859858):

Although not as strong as Bitcoin's, a substantial positive correlation indicates Ethereum's close link with the NASDAQ.

• NASDAQ & LTC (0.163167):

A weak positive correlation, indicating Litecoin's movements are less synchronized with NASDAQ compared to other cryptocurrencies.

INTER-CRYPTOCURRENCY CORRELATION:

• BTC & ETH (0.895691):

Bitcoin and Ethereum share a **very strong positive correlation**, indicating that their price trends move in tandem most of the time. This is expected as they are the two largest cryptocurrencies by market capitalization and are often influenced by similar market dynamics, news, and investor sentiment.

• BTC & LTC (0.4347440):

Bitcoin and Litecoin have a **moderate positive correlation**, meaning they share some similarities in price movements but are not as closely aligned as BTC with ETH or DOGE. Litecoin, being an earlier altcoin, may have its price influenced by unique factors (e.g., its use as a testbed for Bitcoin innovations).

• ETH & LTC (0.448356):

Ethereum and Litecoin have a **moderate positive correlation**, indicating less price dependency compared to ETH's relationship with BTC and DOGE. This suggests that Litecoin's market behavior is less influenced by Ethereum's trends.





STATIONARITY OF VARIABLES:

Stationarity at I (1) via ADF (Augmented Dicky Fuller) Testing:

| Time Series | ADF stat | at 1% | at 5% | at 10% | p-value |
|-----------------------|-----------|-----------|-----------|-----------|---------|
| Nasdaq (NDX) | -20.35719 | -3.982011 | -3.421508 | -3.133535 | 0.0000 |
| Bitcoin (BTC) | -18.31172 | -3.982011 | -3.421508 | -3.133535 | 0.0000 |
| Ethereum (ETH) | -15.47034 | -3.982074 | -3.421539 | -3.133553 | 0.0000 |
| Litecoin (LTC) | -17.59367 | -3.982074 | -3.421539 | -3.133553 | 0.0000 |
| Grayscale Bitcoin ETF | | | | | |
| (GBTC) | -17.46934 | -3.982011 | -3.421508 | -3.133535 | 0.0000 |
| EUR/USD | -22.17754 | -3.982011 | -3.421508 | -3.133535 | 0.0000 |

After assessing these, we come to the conclusion that all the variables are stationary at first difference $I\left(1\right)$.

MULTIPLE REGRESSION ANALYSIS:

Dependent Variable: NASDAQ Method: Least Squares Date: 12/11/24 Time: 10:30

Sample: 1 388

Included observations: 388

Huber-White-Hinkley (HC1) heteroskedasticity consistent standard errors and covariance

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|---|--|--|--|
| LG_BTC ETH LTC GBTC EUR_USD C | 3005.882 0.843489 -21.30266 67.76639 11638.16 -31505.43 | 136.2210 0.077499 1.501086 5.525758 1029.026 1621.120 | 22.06621 10.88385 -14.19149 12.26373 11.30987 -19.43436 | 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) Prob(Wald F-statistic) | 0.961671 0.961169 811.0105 2.51E+08 -3146.458 1916.871 0.000000 0.000000 | Mean depen S.D. depend Akaike info c Schwarz cri Hannan-Qui Durbin-Wats Wald F-stati | lent var riterion terion nn criter. son stat | 11525.89 4115.657 16.24978 16.31103 16.27407 0.301517 2163.703 |





Interpretation:

Coefficients:

- Intercept (C): The constant term, -31505.43, represents the NASDAQ's predicted value when all independent variables are zero.
- **BTC** (**Bitcoin**): A coefficient of 3005.882 indicates that a 1-unit increase in BTC leads to a 3005.882 -unit increase in NASDAQ, holding other variables constant.
- ETH (Ethereum): A coefficient of 0.843489 suggests that a 1-unit increase in ETH results in a 0.843489 -unit increase in NASDAQ, holding other variables constant.
- LTC (Litecoin): A coefficient of -21.30266 indicates that a 1-unit increase in LTC leads to a 21.30266-unit decrease in NASDAQ, holding other variables constant.
- **GBTC** (**Grayscale Bitcoin**): A coefficient of 67.76639 suggests that a 1-unit increase in DOGE results in a 67.76639-unit increase in NASDAQ.
- **EUR/USD**: A coefficient of 11638.16 suggests that a 1-unit increase in DOGE results in a 11638.16-unit increase in NASDAQ.

"A key thing to note here would be to understand that in order to get significant results, log of GBTC has been taken to make sure it does not give us problems later on like multi-collinearity and other problems."

Significance:

Individual Hypothesis

- Null Hypothesis (H₀):
 - β1=0
 - \circ $\beta 2=0$





- \circ $\beta 3=0$
- β4=0

• Alternative Hypothesis (H₁):

- o β1≠0
- o β2≠0
- o β3≠0
- o β4≠0

a. T-values:

• BTC:

- o tcal=22.06621>2 (benchmark).
- We reject H0 and conclude that there is a significant relationship between NASDAQ and BTC.

• ETH:

- o tcal=10.88385>2 (benchmark).
- We reject H0 and conclude that there is a significant relationship between NASDAQ and ETH.

• LTC:

- o |tcal|=-14.19149>2 (benchmark).
- We reject H0 and conclude that there is a significant relationship between NASDAQ and LTC.

• GBTC:

- o tcal=12.26373>2 (benchmark).
- We reject H0 and conclude that there is a significant relationship between NASDAQ and GBTC.





• EUR/USD:

- \circ tcal=11.30987>2 (benchmark).
- We reject H0 and conclude that there is a significant relationship between NASDAQ and EUR/USD.

b. P-values at 5% level of significance:

- **BTC:** p=0.0000<0.05, so we reject H0 and conclude that there is a significant relationship between NASDAQ and BTC.
- ETH:p=0.0000<0.05, so we reject H0 and conclude that there is a significant relationship between NASDAQ and ETH.
- LTC: p=0.0000<0.05, so we reject H0 and conclude that there is a significant relationship between NASDAQ and LTC.
- **GBTC:** p=0.0000<0.05, so we reject H0 and conclude that there is a significant relationship between NASDAQ and GBTC.
- **EUR/USD:** p=0.0000<0.05, so we reject H0 and conclude that there is a significant relationship between NASDAQ and EUR/USD.

Model Fit:

• $R^2 = 0.961671$ or 96.1671%

This indicates that **96.1671**% of the variation in NASDAQ is explained by BTC, ETH, LTC, and DOGE, while the remaining is explained by other factors not included in the model.

• Adjusted $R^2 = 0.961169$

This adjusted value accounts for the number of predictors in the model and confirms a strong model fit.





DIAGNOSTIC TESTS

MULTICOLLINEARITY

When independent variables in your regression model have a high degree of correlation with one another, this is known as multicollinearity, and it can skew the estimates of your regression coefficients.

VARIANCE INFLATION FACTOR (VIF)

TABLE FOR VIF

| Variable | Coefficient | Uncentered | Centered |
|----------|-------------|------------|----------|
| | Variance | VIF | VIF |
| LG_BTC | 18556.16 | 1122.689 | 9.943675 |
| ETH | 0.006006 | 13.10171 | 5.319776 |
| LTC | 2.253260 | 14.29306 | 3.533292 |
| GBTC | 30.53400 | 16.45677 | 5.146750 |
| EUR_USD | 1058895. | 817.7363 | 1.958466 |
| C | 2628030. | 1631.990 | NA |

EXPLANATION

Interpreting VIF Values

- **VIF < 10**: Indicates there is low multicollinearity and it's not a concern.
- **VIF > 10**: Suggests high multicollinearity and may require action to address it.

Here,





- **BTC**: The Centered VIF is 9.790880, which is below 10, indicating that BTC is not experiencing severe multicollinearity.
- **ETH**: The Centered VIF is 8.031357, which is below 10, suggesting that ETH has low multicollinearity and it is not a concern.
- LTC: The Centered VIF is 2.364753, which is well below 10, indicating no multicollinearity issue.
- **GBTC**: The Centered VIF is 5.452328, below 10, so there is no concern of multicollinearity.
- **EUR/USD:** The Centered VIF is 2.971264, below 10, so there is no concern of multicollinearity.

HETEROSCEDASTICITY

When the variance of the residuals varies across all levels of the independent variables, this is known as heteroskedasticity. Hypothesis tests may become unreliable as a result of this violation of the homoscedasticity assumption.

Heteroskedasticity causes inefficient coefficient estimates and invalid standard errors, leading to unreliable t- and F-tests.

WHITE TEST





Heteroskedasticity Test: White Null hypothesis: Homoskedasticity

| F-statistic | 3.503942 | Prob. F(20,367) | 0.0000 |
|---------------------|----------|----------------------|--------|
| Obs*R-squared | 62.20981 | Prob. Chi-Square(20) | 0.0000 |
| Scaled explained SS | 44.43614 | Prob. Chi-Square(20) | 0.0013 |

Test Equation:

Dependent Variable: RESID^2 Method: Least Squares Date: 12/11/24 Time: 10:05

Sample: 1 388

Included observations: 388

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|---------------------|-------------|--------------------|-------------|----------|
| С | -63641473 | 47648204 | -1.335653 | 0.1825 |
| LG_BTC ² | -8781.434 | 438743.2 | -0.020015 | 0.9840 |
| LG_BTC*ETH | -310.2896 | 429.9922 | -0.721617 | 0.4710 |
| LG_BTC*LTC | -989.4735 | 6703.438 | -0.147607 | 0.8827 |
| LG_BTC*GBTC | -56551.21 | 59411.02 | -0.951864 | 0.3418 |
| LG_BTC*EUR_USD | -1865401. | 4899125. | -0.380762 | 0.7036 |
| LG_BTC | 3534179. | 6579734. | 0.537131 | 0.5915 |
| ETH ² | -0.306026 | 0.151167 | -2.024421 | 0.0437 |
| ETH*LTC | 9.664104 | 2.872438 | 3.364426 | 0.0008 |
| ETH*GBTC | 39.17647 | 17.31756 | 2.262240 | 0.0243 |
| ETH*EUR_USD | -6908.152 | 3076.324 | -2.245587 | 0.0253 |
| ETH | 10481.07 | 4248.904 | 2.466771 | 0.0141 |
| LTC ² | -24.56818 | 30.33287 | -0.809952 | 0.4185 |
| LTC*GBTC | -15.78533 | 538.3052 | -0.029324 | 0.9766 |
| LTC*EUR_USD | -9700.116 | 45577.58 | -0.212826 | 0.8316 |
| LTC | 16389.13 | 78149.03 | 0.209716 | 0.8340 |
| GBTC ² | 2259.857 | 1636.614 | 1.380813 | 0.1682 |
| GBTC*EUR_USD | 26920.36 | 238948.9 | 0.112662 | 0.9104 |
| GBTC | 324651.6 | 485401.7 | 0.668831 | 0.5040 |
| EUR_USD^2 | -21134190 | 22355769 | -0.945357 | 0.3451 |
| EUR_USD | 71467985 | 60325498 | 1.184706 | 0.2369 |
| R-squared | 0.160335 | Mean depen | dent var | 647566.8 |
| Adjusted R-squared | 0.114576 | S.D. dependent var | | 787167.3 |
| S.E. of regression | 740700.5 | Akaike info o | | 29.92118 |
| Sum squared resid | 2.01E+14 | Schwarz cri | terion | 30.13557 |
| Log likelihood | -5783.710 | Hannan-Qui | nn criter. | 30.00618 |
| F-statistic | 3.503942 | Durbin-Wats | son stat | 0.854311 |

Since the p-value is less than 0.05, that means we accept the Null Hypothesis stating that there is heterokedasticity present in our model.

This infers that there is no constant mean and constant variance across the model.

However, in order to tackle this problem we will have used robust mean method to cater the issue of heteroscedasticity here.

It will be shown in an MLS output below after all the tests.





BREUSCH-PAGAN-GODFREY TEST

Heteroskedasticity Test: Breusch-Pagan-Godfrey Null hypothesis: Homoskedasticity

| F-statistic Obs*R-squared | 13.19943 | Prob. F(5,382) Prob. Chi-Square(5) | 0.0209 0.0216 |
|---------------------------|----------|---------------------------------------|------------------|
| Scaled explained SS | 9.428282 | Prob. Chi-Square(5) | 0.0932 |

Test Equation:

Dependent Variable: RESID² Method: Least Squares Date: 12/11/24 Time: 10:06

Sample: 1 388

Included observations: 388

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|---|--|--|--|
| C LG_BTC ETH LTC GBTC EUR_USD | -1042650. -42403.37 37.08359 -128.0006 -3500.148 1901222. | 1551854. 124877.6 76.01843 1309.537 5475.500 1012803. | -0.671874 -0.339559 0.487824 -0.097745 -0.639238 1.877189 | 0.5021 0.7344 0.6260 0.9222 0.5231 0.0613 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.034019 0.021375 778708.8 2.32E+14 -5810.897 2.690594 0.020949 | Mean depen S.D. depend Akaike info d Schwarz cri Hannan-Qui Durbin-Wats | lent var riterion terion nn criter. | 647566.8 787167.3 29.98401 30.04526 30.00829 0.795120 |

As we can see that the Chi-Squared Probability is greater than 0.05, this tells us that we should be accepting the Null Hypothesis and claiming that homoskedastcity exists in this model.

This infers that there is constant mean and constant variance across the model.

The functional form for this test will be:

 $Resid^2 = C + LG_BTC + ETH + LTC + GBTC + EUR_USD$



HARVEY TEST

Heteroskedasticity Test: Harvey Null hypothesis: Homoskedasticity

| F-statistic | 2.891751 | Prob. F(5,382) | 0.0141 |
|---------------------|----------|---------------------|--------|
| Obs*R-squared | 14.15027 | Prob. Chi-Square(5) | 0.0147 |
| Scaled explained SS | 7.697453 | Prob. Chi-Square(5) | 0.1737 |

Test Equation: Dependent Variable: LRESID2 Method: Least Squares Sample: 1.388 Included observations: 388

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|---|--|--|--|
| C LG_BTC ETH LTC GBTC EUR USD | 6.637280 -0.163310 0.000137 -0.001675 0.003346 6.577825 | 3.230121 0.259928 0.000158 0.002726 0.011397 2.108108 | 2.054808 -0.628288 0.868174 -0.614526 0.293612 3.120251 | 0.0406 0.5302 0.3858 0.5392 0.7692 0.0019 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.036470 0.023858 1.620850 1003.574 -734.9097 2.891751 0.014113 | Mean depen S.D. depend Akaike info d Schwarz cri Hannan-Qui Durbin-Wats | dent var lent var riterion terion nn criter. | 12.55387 1.640539 3.819122 3.880375 3.843408 1.096867 |

Here we also see that the Chi-Square Probability is greater than 0.05 which leads us to the conclusion that Heteroskedasticity does not exist in the model.

Harvey test is distinct as it tests Heteroskedasticity by estimating the model by taking logarithmic residual^2 as the dependent variable and keeping the other independent variables the same as the regular OLS ones.

The function form for this becomes:

As for the concern of White test and it not proving significant for homoscedasticity, we apply robust mean methods which makes an exception for the heteroskedasticity scenario.

Since, the white test is more robust and provides a non-parametric idea for heteroskedasticity, we apply robust mean method and use the following as our model, this will be used as our model which has its own VIFs (controlled) and the residuals are normal as well and the model is also consistent (will be proven in Ramsey RESET test)





We run the regular OLS model and make a few changes.

Dependent Variable: NASDAQ Method: Least Squares Date: 12/11/24 Time: 10:39

Sample: 1 388

Included observations: 388

Huber-White-Hinkley (HC1) heteroskedasticity consistent standard errors and covariance

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|------------------------|-------------|----------------------|-------------|----------|
| LG_BTC | 3005.882 | 136.2210 | 22.06621 | 0.0000 |
| ETH | 0.843489 | 0.077499 | 10.88385 | 0.0000 |
| LTC | -21.30266 | 1.501086 | -14.19149 | 0.0000 |
| GBTC | 67.76639 | 5.525758 | 12.26373 | 0.0000 |
| EUR_USD | 11638.16 | 1029.026 | 11.30987 | 0.0000 |
| Ċ | -31505.43 | 1621.120 | -19.43436 | 0.0000 |
| R-squared | 0.961671 | Mean dependent var | | 11525.89 |
| Adjusted R-squared | 0.961169 | S.D. depend | lent var | 4115.657 |
| S.E. of regression | 811.0105 | Akaike info o | riterion | 16.24978 |
| Sum squared resid | 2.51E+08 | Schwarz cri | terion | 16.31103 |
| Log likelihood | -3146.458 | Hannan-Quinn criter. | | 16.27407 |
| F-statistic | 1916.871 | Durbin-Watson stat | | 0.301517 |
| Prob(F-statistic) | 0.000000 | Wald F-stati | stic | 2163.703 |
| Prob(Wald F-statistic) | 0.000000 | | | |

Here, it clearly explains that it caters for the heteroskedasticity consistent standard errors and covariance.

AUTOCORRELATION

When the residuals are not independent of one another, autocorrelation arises, which goes against the presumption that errors are random.

We can detect Autocorrelation in our model by running the following tests on our model.

DURBIN WATSON TEST





Dependent Variable: NASDAQ Method: Least Squares Date: 12/11/24 Time: 10:39 Sample: 1 388

Included observations: 388

Huber-White-Hinkley (HC1) heteroskedasticity consistent standard errors and covariance

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|---|---|--|--|--|
| LG_BTC ETH LTC GBTC EUR_USD C | 3005.882 0.843489 -21.30266 67.76639 11638.16 -31505.43 | 136.2210 0.077499 1.501086 5.525758 1029.026 1621.120 | 22.06621 10.88385 -14.19149 12.26373 11.30987 -19.43436 | 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) Prob(Wald F-statistic) | 0.961671 0.961169 811.0105 2.51E+08 -3146.458 1916.871 0.000000 0.000000 | Mean depen S.D. depend Akaike info d Schwarz cri Hannan-Qui Durbin-Wats Wald F-stati | dent var criterion terion nn criter. son stat | 11525.89 4115.657 16.24978 16.31103 16.27407 0.301517 2163.703 |

Here, we look at the Durbin Watson statistic and as we can see that it is close to zero which makes it near positive autocorrelation according to the following threshold:

DW \approx **2**: No autocorrelation.

DW = 2: Positive autocorrelation.

DW = **4**: Negative autocorrelation.

Since, Heteroskedasticity exists here we go a step further to make our results more concise and consistent.

DURBIN WATSON -h TEST

To apply Durbin h test, we use to following formula:

$$h = (1 - d/2) \sqrt{(n-1)/(1-n*var)}$$

Here,

d= regular Durbin Watson stat (0.301517)





n= number of observations (388)

var = is the estimated variance of the coefficient of the lagged dependent variable (0.000731323849)

And to achieve these values we run the serial correlation LM test which gives us this answer.

Breusch-Godfrey Serial Correlation LM Test: Null hypothesis: No serial correlation at up to 1 lag

| F-statistic | Prob. F(1,381) | 0.0000 |
|---------------|---------------------|--------|
| Obs*R-squared | Prob. Chi-Square(1) | 0.0000 |

Test Equation:

Dependent Variable: RESID Method: Least Squares Date: 12/11/24 Time: 11:40

Sample: 1 388

Included observations: 388

Presample missing value lagged residuals set to zero.

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|--|--|--|---|
| LG_BTC ETH LTC GBTC EUR_USD C RESID(-1) | -124.5307 0.020637 1.386062 3.992011 -358.3802 1375.811 0.853748 | 68.59779 0.041694 0.719505 3.005486 555.5483 852.1693 0.027043 | -1.815374 0.494964 1.926409 1.328241 -0.645093 1.614481 31.57050 | 0.0703 0.6209 0.0548 0.1849 0.5193 0.1073 0.0000 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.723451 0.719096 427.0526 69484457 -2897.096 166.1161 0.000000 | Mean dependent var S.D. dependent var Akaike info criterion Schwarz criterion Hannan-Quinn criter. Durbin-Watson stat | | -1.79E-12 805.7543 14.96957 15.04103 14.99790 1.880047 |

Now we take the values and get the h stat:

h= 14.138974417

Now, as the threshold gives us the reasoning that the lagged variable's probability should be less than 0.005 and it is 0.0000,

AND

The h stat value is greater than 1.96 which gives us proof that serial correlation exists and we reject the Null Hypothesis saying:

H0: No serial correlation at up to 1 lag.





Ramsey RESET Test

Now we apply the Ramsey RESET test in order to check any misspecifications in our model.

After application of the test, we get:

Ramsey RESET Test Equation: UNTITLED

Omitted Variables: Squares of fitted values

Specification: NASDAQ LG_BTC ETH LTC GBTC EUR_USD C

| | Value | df | Probability |
|------------------|----------|----------|-------------|
| t-statistic | 0.105633 | 381 | 0.9159 |
| F-statistic | 0.011158 | (1, 381) | 0.9159 |
| Likelihood ratio | 0.011363 | 1 | 0.9151 |

This table shows a p-value of 0.9151 which means that there is no misspecification in our model and this model is optimal.





CORRELOGRAM OF RESIDUALS

| Correlogram of Residuals | | | | | | |
|---|---------------------|------|---------|--------|--------|-------|
| | Correlogram | OI F | Cesidua | ZIS | | |
| Date: 12/11/24 Tim Sample: 1 388 Included observation | | | | | | |
| Autocorrelation | Partial Correlation | | AC | PAC | Q-Stat | Prob |
| 1 | | 1 | 0.847 | 0.847 | 280.77 | 0.000 |
| ı | 101 | 2 | | -0.048 | 475.39 | 0.000 |
| I | ib | 3 | 0.610 | 0.088 | 621.59 | 0.000 |
| ı | 101 | 4 | | -0.036 | 728.08 | 0.000 |
| ı | I | 5 | 0.455 | 0.055 | 809.95 | 0.000 |
| ı | I | 6 | 0.424 | 0.082 | 881.17 | 0.000 |
| ı | | 7 | 0.406 | 0.052 | 946.52 | 0.000 |
| ı | I | 8 | 0.399 | 0.066 | 1009.9 | 0.000 |
| 1 | 1 1 | 9 | 0.383 | -0.009 | 1068.4 | 0.000 |
| ı <u>—</u> | I <u>d</u> I | 10 | 0.338 | -0.072 | 1114.2 | 0.000 |
| ı | 1 1 | 11 | 0.299 | 0.007 | 1150.0 | 0.000 |
| ı 🗀 | [| 12 | 0.257 | -0.034 | 1176.5 | 0.000 |
| 1 | 1 1 | 13 | | -0.006 | 1195.4 | 0.000 |
| ı 🗀 | [| 14 | 0.182 | -0.021 | 1208.7 | 0.000 |
| ı 🗖 | 1 1 | 15 | 0.153 | -0.015 | 1218.3 | 0.000 |
| ı | I <u>I</u> I | 16 | 0.112 | -0.078 | 1223.4 | 0.000 |
| ı j p | 1 1 | 17 | 0.082 | 0.001 | 1226.1 | 0.000 |
| ı j jı | 1 1 | 18 | 0.065 | 0.010 | 1227.9 | 0.000 |
| 1 j ji | | 19 | 0.059 | 0.035 | 1229.3 | 0.000 |
| 1 j ji | | 20 | 0.065 | 0.043 | 1231.0 | 0.000 |
| ١ 🌓 | [] | 21 | | -0.089 | 1231.8 | 0.000 |
| 1 1 | | 22 | 0.023 | 0.005 | 1232.0 | 0.000 |
| 1 1 | | 23 | 0.015 | 0.014 | 1232.1 | 0.000 |
| 1 1 | '[' | 24 | 0.001 | -0.006 | 1232.1 | 0.000 |
| 1 1 | | 25 | 0.007 | 0.082 | 1232.1 | 0.000 |
| 1 1 | ']' | 26 | 0.018 | 0.005 | 1232.3 | 0.000 |
| 1] 1 | | 27 | 0.021 | -0.007 | 1232.5 | 0.000 |

The result of the correlogram shows us that autocorrelation exists in our model, this explains that the residuals are not independent of each other which leads us to believe that the lagged variables are connected and it is by one another.

We look at the p-value here, it is visible that they are less than 0.05 which indicates the presence of Autocorrelation.

Generally, autocorrelation can never be zero in a model as dependency of events across the model is something which is inevitable.

As autocorrelation is present in our residuals:

1. Include Lagged Terms:

o Add lagged dependent or independent variables to the model.





This would be our go to remedy to decrease the autocorrelation problem.

As this can create this problem:

"Autocorrelation invalidates standard errors and makes coefficient estimates inefficient."

After including lagged term AR (1) in our model we found out that the autocorrelation controlled and there is no fluctuated spike in the ACF PACF.

| | Correlogram | of R | esidua | als | | |
|---|---------------------|--|---|--|--|--|
| Date: 12/11/24 Time: 12:13 Sample: 1 388 Q-statistic probabilities adjusted for 1 ARMA term | | | | | | |
| Autocorrelation | Partial Correlation | | AC | PAC | Q-Stat | Prob* |
| | | 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 | 0.004 0.013 -0.060 -0.062 0.050 -0.060 -0.024 0.049 0.013 0.005 -0.082 0.005 0.033 0.115 -0.046 0.010 0.032 -0.044 0.145 0.034 -0.029 | -0.046 0.001 0.013 -0.059 -0.068 0.045 -0.055 -0.032 0.038 0.020 0.005 -0.097 0.004 0.118 -0.046 -0.005 0.046 -0.005 0.046 -0.005 0.046 -0.005 - | 0.8119 0.8167 0.8818 2.2807 3.8227 4.8230 6.2709 6.4909 7.4306 7.4937 7.5024 10.233 10.242 10.678 15.997 16.858 16.898 17.320 18.131 26.719 27.185 27.527 28.582 | 0.366 0.643 0.516 0.431 0.438 0.394 0.484 0.491 0.586 0.677 0.510 0.595 0.638 0.314 0.327 0.392 0.433 0.447 0.111 0.130 0.154 0.157 |
| 1]1 | | | 0.026 | 0.008 | 28.853 | 0.185 |

JOHANSEN CO-INTEGRATION TEST

As the variables are all stationary at first difference I (1), we use the Johansen Cointegration test and see the long-term relationships in our model. As the Ramsey RESET test also showed that there are no misspecifications in our model, this model is sound to be tested by the Johansen Cointegration Test.





TRACE TEST

| | | | 0.05 Critical | |
|------------|------------|------------|---------------|--------|
| No of CE's | Eigenvalue | Trace Stat | Value | Prob |
| None* | 0.237912 | 195.155 | 103.8473 | 0.0000 |
| At most 1* | 0.105792 | 90.55312 | 76.97277 | 0.0032 |
| At most 2 | 0.054954 | 47.50341 | 54.07904 | 0.1692 |
| At most 3 | 0.034845 | 25.74259 | 35.199275 | 0.3565 |
| At most 4 | 0.019229 | 12.08813 | 20.26184 | 0.441 |
| At most 5 | 0.01191 | 4.612992 | 9.164546 | 0.3285 |

After the Johansen Cointegration test is applied, we look at Trace Test here which tells us how many cointegrating equations exist.

According to the trace test, we can see that there are at most 2 co-integrating equations.

This output from E-views gives the trace test table. Trace test indicates 2 cointegrating eqn(s) at the 0.05 level.

RANK TEST

In the rank test, we look at the maximum Eigen Value and look at our table and see for cointegrating equations.

| No of CE's | Eigenvalue | Trace Stat | 0.05 Critical Value | Prob |
|------------|------------|------------|------------------------|--------|
| None* | 0.237912 | 104.6019 | 40.9568 | 0.0000 |
| None. | 0.237912 | 104.0019 | 40.9308 | |
| At most 1* | 0.105792 | 43.04971 | 34.80587 | 0.0042 |
| At most 2 | 0.054954 | 21.76083 | 28.58808 | 0.2897 |
| At most 3 | 0.034845 | 13.65445 | 22.29962 | 0.4941 |
| At most 4 | 0.019229 | 7.47514 | 15.8921 | 0.6113 |
| At most 5 | 0.01191 | 4.612992 | 9.164546 | 0.3285 |

Max-eigenvalue test indicates 2 cointegrating equations at the 0.05 level.

"*" denotes the rejection of the hypothesis at the 0.05 level.

Now, we look at the cointegrating equation:





| 1 Cointegrating | Equation(s): | Log likelihood | -5190.758 | | | |
|-----------------|--------------------|------------------|-------------------|-----------|-----------|-----------|
| Normalized coir | ntegrating coeffic | cients (standard | error in parenthe | eses) | | |
| NASDAQ | LG BTC | ÉTH | GBTC | EUR USD | LTC | С |
| 1.000000 | -3094.991 | -1.005364 | -211.9446 | -9197.858 | 22.24454 | 29326.57 |
| | (870.956) | (0.52571) | (38.7634) | (7086.09) | (9.50925) | (11028.0) |

Here, we can obtain the t-stat values from this table and see if the variables are significant or not which in this case are significant, which tells us that there exists a long-term relationship between our variables in the model one way or another.

Vector Error Correction Model (VECM)

In order to see our model exhibits long-term relationship, we need to see if the ECT (Error Correction Term) is stationary at level I (0) and all the other variables are stationary at first difference. So, we run the stationarity test on our error term and we find out that:

| Augmented Dickey-Fuller Unit Root Test on VECM | | | | | |
|---|---|--|--------|--|--|
| Null Hypothesis: VECM has a unit root Exogenous: Constant, Linear Trend Lag Length: 0 (Automatic - based on SIC, maxlag=16) | | | | | |
| | | t-Statistic | Prob.* | | |
| Augmented Dickey-Fu Test critical values: | uller test statistic 1% level 5% level 10% level | -5.574405 -3.981949 -3.421478 -3.133517 | 0.0000 | | |

Here the value is significant and we reject the null hypothesis which states, VECM has a unit root, this means that our error correction term is stationary at level and therefore exhibits a long-term relationship in our model.



Granger Casualty Test

Now, we move onto the next thing, which is Granger Casualty Test which we obtained from E-Views as follows.

> Pairwise Granger Causality Tests Date: 12/11/24 Time: 13:56 Sample: 1 388

Lags: 1

| Null Hypothesis: | Obs | F-Statistic | Prob. |
|---|-----|--------------------|------------------|
| LG_BTC does not Granger Cause NASDAQ | 387 | 4.07745 | 0.0442 |
| NASDAQ does not Granger Cause LG_BTC | | 3.11340 | 0.0784 |
| ETH does not Granger Cause NASDAQ | 387 | 5.09070 | 0.0246 |
| NASDAQ does not Granger Cause ETH | | 2.64422 | 0.1047 |
| GBTC does not Granger Cause NASDAQ | 387 | 0.49558 | 0.4819 |
| NASDAQ does not Granger Cause GBTC | | 0.34821 | 0.5555 |
| EUR_USD does not Granger Cause NASDAQ | 387 | 2.92783 | 0.0879 |
| NASDAQ does not Granger Cause EUR_USD | | 0.07626 | 0.7826 |
| LTC does not Granger Cause NASDAQ | 387 | 0.05119 | 0.8211 |
| NASDAQ does not Granger Cause LTC | | 0.52973 | 0.4672 |
| ETH does not Granger Cause LG_BTC | 387 | 6.05274 | 0.0143 |
| LG_BTC does not Granger Cause ETH | | 0.04968 | 0.8237 |
| GBTC does not Granger Cause LG_BTC | 387 | 1.06556 | 0.3026 |
| LG_BTC does not Granger Cause GBTC | | 0.25454 | 0.6142 |
| EUR_USD does not Granger Cause LG_BTC LG_BTC does not Granger Cause EUR_USD | 387 | 2.10026 0.11747 | 0.1481 0.7320 |
| LTC does not Granger Cause LG_BTC | 387 | 0.48898 | 0.4848 |
| LG_BTC does not Granger Cause LTC | | 0.75395 | 0.3858 |
| GBTC does not Granger Cause ETH | 387 | 2.44957 | 0.1184 |
| ETH does not Granger Cause GBTC | | 0.23872 | 0.6254 |
| EUR_USD does not Granger Cause ETH ETH does not Granger Cause EUR_USD | 387 | 3.32589 0.31322 | 0.0690 0.5760 |
| LTC does not Granger Cause ETH | 387 | 10.2805 | 0.0015 |
| ETH does not Granger Cause LTC | | 2.66191 | 0.1036 |
| EUR_USD does not Granger Cause GBTC GBTC does not Granger Cause EUR_USD | 387 | 0.27377 1.87094 | 0.6011 0.1722 |
| LTC does not Granger Cause GBTC | 387 | 0.89704 | 0.3442 |
| GBTC does not Granger Cause LTC | | 0.01895 | 0.8906 |
| LTC does not Granger Cause EUR_USD EUR_USD does not Granger Cause LTC | 387 | 1.30218 2.22288 | 0.2545 0.1368 |
| | | | |

This table explains how the variables have predictive causality between one another, these results show if the causal relationship is bidirectional, no causality and unidirectional causality and it can be measured based on the following criteria.

H0: There is no Granger Cause.

H1: There is Granger Cause.





If the p-value is less than 0.05 then we reject Ho and if it is greater than 0.05, there is a granger cause.

MODEL SELECTION:

ARMA/ARIMA Modelling

So, for the selection of ARMA/ARIMA models, there is a procedure that needs to be followed to get there, we incorporated the following steps for the model selection.

We will be assessing Nasdaq (NDX) for the modelling basis. Since we used the log values in the normality procedure, we will take log of NDX.

Before the procedure, the rule of thumb for ARMA/ARIMA model is as follows:

INSTRUCTIONS:

Model Selection (ARMA vs. ARIMA):

- **ARMA** (**p**, **q**): Use when the time series is stationary (no trends or seasonality).
- ARIMA (p, d, q): Use when the time series is non-stationary (trend or seasonality present). The "d" (differencing) removes non-stationarity.

Checking Stationarity:

- Use the Augmented Dickey-Fuller (ADF) test or visual inspection.
 - o If p>0.05, the series is non-stationary \rightarrow Apply differencing (d).
 - o If p≤0.05, the series is stationary \rightarrow No differencing needed (d=0).





Parameter Selection:

- Autoregressive (AR, p): Determine from the Partial Autocorrelation Function (PACF).
 - 1. Significant spikes at lag k in the PACF suggest AR terms up to p=k.
- Moving Average (MA, q): Determine from the Autocorrelation Function (ACF)
 - 2. Significant spikes at lag k in the ACF suggest MA terms up to q=k.
- **Differencing (d)**: Number of times the series is differenced to achieve stationarity.
 - 3. Start with d=1 and test stationarity after each differencing.

APPLICATION:

Nasdaq (NDX):

We took log (NDX) first, and we checked the stationarity:

At level I (0):





Augmented Dickey-Fuller Unit Root Test on LG_NASDAQ

Null Hypothesis: LG_NASDAQ has a unit root

Exogenous: Constant, Linear Trend

Lag Length: 0 (Automatic - based on SIC, maxlag=16)

| | | t-Statistic | Prob.* |
|--|---|--|--------|
| Augmented Dickey-Fu Test critical values: | uller test statistic 1% level 5% level 10% level | -2.226742 -3.981949 -3.421478 -3.133517 | 0.4728 |

^{*}MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation Dependent Variable: D(LG_NASDAQ)

Method: Least Squares Date: 12/01/24 Time: 18:24 Sample (adjusted): 2 388

Included observations: 387 after adjustments

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|--|--|------------------------------------|--|
| LG_NASDAQ(-1) C @TREND("1") | -0.025466 0.248521 -7.82E-05 | 0.011437 0.113115 3.78E-05 | -2.226742 2.197065 -2.069194 | 0.0265 0.0286 0.0392 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.012752 0.007610 0.029348 0.330751 817.9126 2.479964 0.085087 | Mean dependent var S.D. dependent var Akaike info criterion Schwarz criterion Hannan-Quinn criter. Durbin-Watson stat | | -0.003167 0.029461 -4.211435 -4.180749 -4.199267 2.086152 |

Since it proves to be insignificant as the p-value is greater than 0.05 and the tabulated value is also less than the critical values.

At First Difference I (1):





Augmented Dickey-Fuller Unit Root Test on D(LG_NASDAQ)

Null Hypothesis: D(LG_NASDAQ) has a unit root

Exogenous: Constant, Linear Trend

Lag Length: 0 (Automatic - based on SIC, maxlag=16)

| | | t-Statistic | Prob.* |
|--|---|--|--------|
| Augmented Dickey-Fu Test critical values: | ıller test statistic 1% level 5% level 10% level | -20.73082 -3.982011 -3.421508 -3.133535 | 0.0000 |

^{*}MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation Dependent Variable: D(LG_NASDAQ,2)

Method: Least Squares Date: 12/01/24 Time: 18:28 Sample (adjusted): 3 388

Included observations: 386 after adjustments

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|--|--|------------------------------------|---|
| D(LG_NASDAQ(-1)) C @TREND("1") | -1.058099 -0.003650 1.26E-06 | 0.051040 0.003026 1.35E-05 | -20.73082 -1.206178 0.093648 | 0.0000 0.2285 0.9254 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.528781 0.526321 0.029510 0.333526 813.6871 214.8930 0.000000 | Mean dependent var S.D. dependent var Akaike info criterion Schwarz criterion Hannan-Quinn criter. Durbin-Watson stat | | 2.13E-05 0.042877 -4.200451 -4.169706 -4.188259 1.997063 |

As the p-value is significant, being less than 0.05 and the tabulated value is also greater than the critical values, this shows us that the series is stationary at first difference.



By having our series stationary at first difference, we come to the conclusion that we will have an ARIMA model which is going to be ARIMA (p, 1, q) model.

Now,

We take the differenced series and run a correlogram on it.

Date: 12/01/24 Time: 18:35 Sample (adjusted): 2 388 Included observations: 387 after adjustments Autocorrelation Partial Correlation AC PAC Q-Stat Prob ığı 1 -0.058 -0.058 1.3109 0.252 1 | 1 ı 0.008 1.3639 0.012 0.506 1 | 0.034 0.035 1.8131 0.612 10 3.6044 0.462 -0.068 -0.064 7.9387 -0.105 -0.114 0.1601 | 0.041 9.0502 0.053 0.171-0.032 -0.0199.4538 101 -0.057 -0.061 10.747 0.216 0.049 0.026 11.704 0.231 10 -0.021 -0.019 11.882 0.293 0.045 0.052 12.688 0.314 12 -0.064 -0.078 14.333 1 0.280 13 -0.010 -0.024 14.372 0.348 0.041 0.051 15.047 0.050 0.058 16.075 16 -0.043 -0.040 16.822 0.397 0.021 -0.007 17.004 17 0.45417.350 18 0.029 0.041 0.49919 -0.059 -0.032 18.752 0.47320.805 20 0.071 0.053 0.40921 0.034 0.037 21.276 22 -0.041 -0.022 21.965 0.46223 -0.074 -0.079 24.204 0.393ı 24 0.025 0.001 24.453 0.43625 0.074 0.111 26.704 1 🗖 0.371 26 -0.033 -0.024 27.168 0.401 27 -0.019 -0.046 27.314 0.44728 0.016 0.006 27.427 0.010 27.486 -0.0120.545 0.003 0.018 27.491 0.597 29.330 31 0.0660.034 0.55232.175 0.082 0.106 0.458 32.659 -0.034 -0.0060.48434 0.017 -0.007 32.778 0.527 35 0.116 0.108 38.537 0.313 36 -0.085 -0.049 41.658 0.238

Looking at significant spikes we see that,





AR (5) & MA (5)

AR (35) & MA (35)

These will turn into:

ARIMA (5,1,5)

ARIMA (35,1,35)

Now, we check these models by running them on E-views by estimating them one by one.

ARIMA (5,1,5)

By estimating:

d (lg_nasdaq) ar (5) ma (5) c

we get:

Dependent Variable: D(LG_NASDAQ)

Method: ARMA Maximum Likelihood (OPG - BHHH)

Date: 12/01/24 Time: 18:43 Sample: 2 388

Included observations: 387

Convergence achieved after 14 iterations

Coefficient covariance computed using outer product of gradients

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|--|--|--|--|
| C AR(5) MA(5) SIGMASQ | -0.003186 0.027365 -0.134535 0.000856 | 0.001375 0.454562 0.452110 4.65E-05 | -2.317600 0.060200 -0.297571 18.38723 | 0.0210 0.9520 0.7662 0.0000 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.011572 0.003830 0.029404 0.331146 817.6524 1.494666 0.215546 | Mean depen S.D. depend Akaike info d Schwarz cri Hannan-Qui Durbin-Wats | dent var criterion terion nn criter. | -0.003167 0.029461 -4.204922 -4.164008 -4.188699 2.116280 |
| Inverted AR Roots | .49 39+.29i | .15+.46i | | 3929i |
| Inverted MA Roots | .67 54+.39i | .21+.64i | .2164i | 5439i |

Findings:

From this estimation we see that the AR and MA variables are not significant at all so we cannot be using this model.





ARIMA (35,1,35)

By estimating:

d(lg_nasdaq) ar(35) ma(35) c

we get:

Dependent Variable: D(LG_NASDAQ)

Method: ARMA Maximum Likelihood (OPG - BHHH)

Date: 12/01/24 Time: 18:46

Sample: 2 388

Included observations: 387

Convergence achieved after 10 iterations

Coefficient covariance computed using outer product of gradients

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|--|--|--|--|
| C AR(35) MA(35) SIGMASQ | -0.003174 -0.456646 0.585331 0.000846 | 0.001643 0.313620 0.295993 4.84E-05 | -1.931920 -1.456049 1.977519 17.49502 | 0.0541 0.1462 0.0487 0.0000 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.022721 0.015066 0.029238 0.327411 819.3281 2.968104 0.031871 | Mean dependent var S.D. dependent var Akaike info criterion Schwarz criterion Hannan-Quinn criter. Durbin-Watson stat | | -0.003167 0.029461 -4.213582 -4.172668 -4.197359 2.079116 |
| Inverted AR Roots Inverted MA Roots | .9709i .88+.42i .68+.71i .38+.90i .0498i 30+.93i 61+.76i 84+.50i 96+.17i .9809i .8943i .6871i .39+.91i .0498i 30+.94i 61+.77i 85+.51i 97+.18i | .97+.09i .8842i .6871i .3890i .04+.98i 3093i 6176i 8450i 9617i .98+.09i .89+.43i .68+.71i .3991i .04+.98i 3094i 6177i 8551i 9718i | 4686i 7464i 9234i 98 .9526i .80+.58i .54+.82i .2296i 13+.98i 4787i 7465i | .94+.26i .7957i .54+.82i .22+.95i 1397i 46+.86i 74+.64i 92+.34i .95+.26i .8058i .5482i .22+.96i 1398i 47+.87i 74+.65i 92+.35i |

Findings:

We see here that only the MA variable is significant and we AR is not.





To move forward, we make sure to use a combination of both to see how it plays out, by keeping the MA (35) and replacing the AR (35) with AR (5).

ARIMA (5,1,35)

By estimating:

d(lg_nasdaq) ar(5) ma(35) c

we get:

Dependent Variable: D(LG_NASDAQ)
Method: ARMA Maximum Likelihood (OPG - BHHH)
Date: 12/01/24 Time: 18:50
Sample: 2 388

Included observations: 387

Convergence achieved after 7 iterations

Coefficient covariance computed using outer product of gradients

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|--|--|---|--|
| C AR(5) MA(35) SIGMASQ | -0.003190 -0.109781 0.139900 0.000840 | 0.001545 0.052801 0.052114 4.74E-05 | -2.065261 -2.079141 2.684499 17.73196 | 0.0396 0.0383 0.0076 0.0000 |
| R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) | 0.029552 0.021950 0.029136 0.325122 820.8575 3.887675 0.009291 | Mean depe S.D. depen Akaike info Schwarz ci Hannan-Qu Durbin-Wa | dent var criterion riterion iinn criter. | -0.003167 0.029461 -4.221486 -4.180572 -4.205263 2.100193 |
| Inverted AR Roots | .5238i 64 | .52+.38i | 20+.61i | 2061i |
| Inverted MA Roots | 04 .94+.08i .8541i .65+.68i .3787i .0494i 2990i 5974i 81+.48i 93+.17i | .9408i .85+.41i .6568i .37+.87i .04+.94i 29+.90i 59+.74i 8148i 9317i | 45+.83i 71+.62i | .9125i .7656i .5279i .21+.92i 13+.94i 4583i 7162i 89+.33i |

Findings:





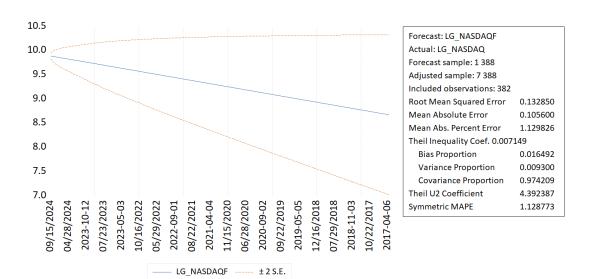
Here, we get both AR and MA to be significant and hence we can use this as our model for forecasting and moving on to the next steps.

FORECASTING:

Moving on to the next step, we will be forecasting another period for the model we selected.

We select the forecasting option from e-views and then:

We get these results:



The forecast gives us the following the insights:

GRAPH EXPLANATION





The ARMA/ARIMA model's predicted values (blue line) and 95% confidence interval (dashed orange lines) for the variable LG_NASDAQ are displayed in the chart. Using historical data up to the most recent observation, a forecast is created for the future (starting in September 2024).

The forecasted values stretch from the most recent observation, displaying the anticipated values for the future, while the blue line represents the actual data.

CONFIDENCE INTERVAL:

The range that the actual future numbers are expected to fall within is represented by the dashed lines surrounding the forecast, which are the 95% confidence interval (± 2 Standard Errors). As the confidence interval around the predicted line gets smaller, the forecast's dependability rises.

TABLE EXPLANATION

- RMSE, or root mean squared error 0.132850 means that, on average, there is a 0.1328-unit difference between the predicted and actual values. The better the model fits the data, the lower the RMSE.
- Mean Absolute Error 0.105600 indicates that there is an average absolute error of 0.1056 units between the predicted and actual values.
- Mean Absolute Percentage Error 1.129826% indicates that, on average, there is a 1.13% discrepancy between the predicted and actual values. Good forecasting performance is indicated by this comparatively low percentage.
- Theil U2 Coefficient





A relative indicator of predicted performance is 4.392387. Although the model's performance might yet be enhanced, a result nearer 1 would signify perfect predictions.

CONCLUSION:

This study examined the effect of cryptocurrency markets on traditional stock market indices using a comprehensive multiple regression analysis of volatility and correlation. By using GBTC ETF data to reflect institutional engagement, the study provided deep insights into the growing relationship between these two financial sectors.

Significant findings revealed a robust correlation between cryptocurrency volatility and shifts in traditional stock indices, pointing to a shifting dynamic fueled by increased institutional participation. The results demonstrate how cryptocurrencies are altering investment strategies and market dynamics.

Future research could expand on these findings by utilizing larger datasets, looking into alternative institutional proxies, and analyzing how specific macroeconomic events affect the relationships between cryptocurrencies and the stock market. Examining geographic variations in these relationships may also yield a deeper understanding of regional market dynamics.

This study concludes by emphasizing the importance of understanding the integration of new financial products into current markets, which paves the way for more informed frameworks for investment and regulatory decisions.





REFERENCES:

Following were the links used to gather data for our study:

https://www.investing.com/indices/nq-100

https://www.investing.com/crypto/bitcoin/btc-usd-historical-data

https://www.investing.com/crypto/ethereum/eth-usd

https://www.investing.com/crypto/litecoin/ltc-usd

https://www.investing.com/crypto/dogecoin/doge-usd

