

Enhancing Inflation Forecasting Models through Bayesian Hierarchical Approaches and Stochastic Seasonality

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Outline

- 1 Introduction
- 2 Data
- 3 Methodology
- 4 Results
- 5 Conclusion

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Introduction

- Central banks have clearly defined inflation targets
- Inflation affects real income and thus consumer spending decisions
- Inflation is of interest to the financial world
- Literature leaves room for improved estimation methods

- Stock and Watson (2007)
 - Inflation has become less volatile but harder to forecast
 - Introduce the unobserved component stochastic volatility model
- Ciccarelli and Mojon (2010)
 - 70% of inflation in OECD countries can be explained by a common factor
- Lis and Porqueddu (2018)
 - Inflation is seasonal, even when correcting for food and energy prices

Research Questions

- 1 *Do seasonality features improve the forecasting performance of existing inflation models?*
- 2 *Does imposing similarity in inflation patterns across countries improve the forecasting performance of inflation models?*

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Inflation dataset

- Global database from the World Bank (Ha et al., 2023)
- Quarterly data
- Headline Consumer Price Index (CPI) inflation
- Core CPI inflation as robustness check
- Subset of 20 countries from Q1 1977 - Q1 2023

Inflation seasonality (1/2)

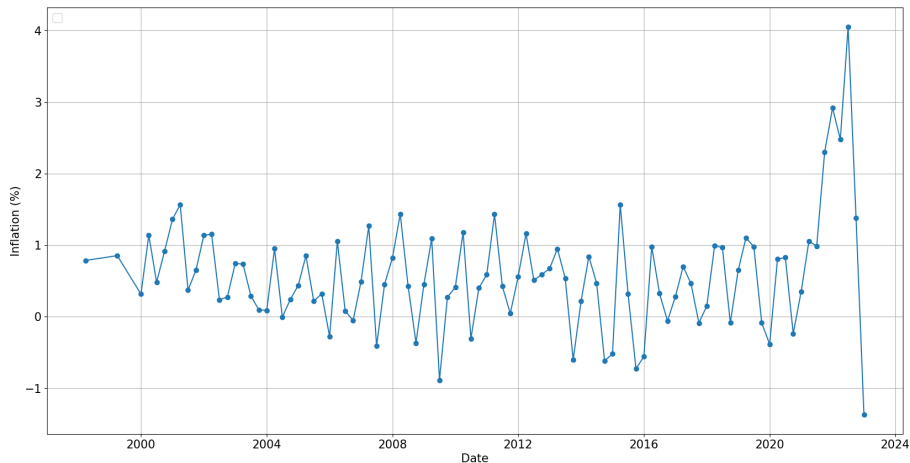


Figure: Quarterly inflation in The Netherlands

Inflation seasonality (2/2)

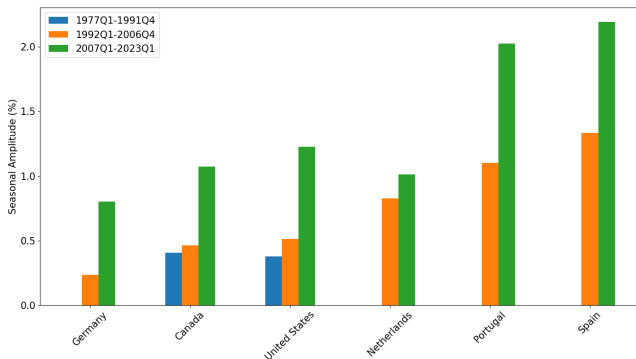


Figure: Range of the seasonal components across countries and sub-samples

- Seasonality stronger for countries with high inflation
- Seasonality increased over time

Inflation correlation and distances

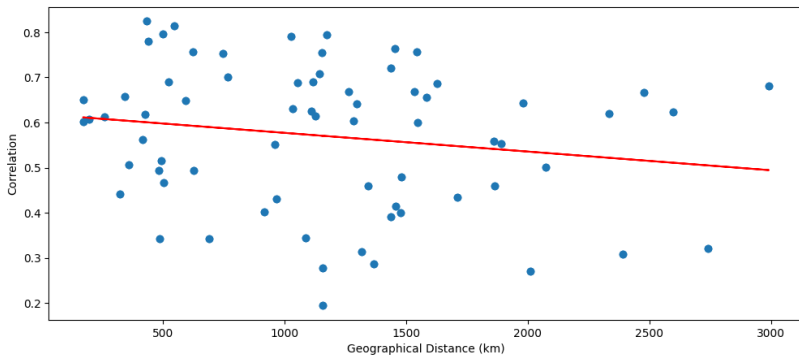


Figure: Correlation of countries and geographical distance

- 1 Inflation correlation decreasing with distance
- 2 Distance data from Mayer and Zignago (2011)

- Commodity prices (Ha et al., 2023)
- Unemployment rates (ILO, 2024)
- GDP growth (OECD, 2024)
- Interest rates (Bank for International Settlements, 2024)

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Random Effects Bayesian Hierarchical Model

$$y_{t,i} = \alpha_i + \beta_i x_{t-1,i} + \gamma_i w_{t-1} + \epsilon_{t,i}$$

Random Effects Bayesian Hierarchical Model

$$y_{t,i} = \alpha_i + \beta_i x_{t-1,i} + \gamma_i w_{t-1} + \epsilon_{t,i}$$

$$\alpha_i \sim N(\alpha_0, \tau_\alpha^{-1})$$

$$\alpha_0 \sim N(0, \tau_{\alpha_0}^{-1})$$

$$\beta_i \sim N(\beta_0, \tau_\beta^{-1})$$

$$\beta_0 \sim N(0, \tau_{\beta_0}^{-1})$$

$$\gamma_i \sim N(\gamma_0, \tau_\gamma^{-1})$$

$$\gamma_0 \sim N(0, \tau_{\gamma_0}^{-1})$$

$$\epsilon_{t,i} \sim N(0, \tau_\epsilon^{-1})$$

$$\tau_{\alpha_0}, \tau_{\beta_0}, \tau_{\gamma_0}, \tau_\alpha, \tau_\beta, \tau_\gamma, \tau_\epsilon \sim \Gamma(1, 1)$$

$$\theta_s \sim \mathcal{N}_n(\mathbf{0}, \Sigma),$$

- Covariance based Bayesian hierarchical model
 - Use sample covariance matrix: $\hat{\Sigma}$
- Distance based Bayesian hierarchical model
 - Matérn kernel to convert distances into covariances: $\Sigma_{\text{Matérn,dist}}$

- Unobserved Component Stochastic Volatility model (Stock and Watson, 2007)
- Split time-series into components
- Variances of the disturbances follow a logarithmic random walk
- Type of Bayesian model

$$y_{t,i} = \tau_{t,i} + \eta_{t,i}$$

$$\tau_{t,i} = \tau_{t-1,i} + \epsilon_{t,i}$$

$$\ln(\sigma_{\eta,t,i}^2) = \ln(\sigma_{\eta,t-1,i}^2) + v_{\eta,t,i}$$

$$\ln(\sigma_{\epsilon,t,i}^2) = \ln(\sigma_{\epsilon,t-1,i}^2) + v_{\epsilon,t,i}$$

$$\eta_{t,i} \sim N(0, \sigma_{\eta,t,i}^2)$$

$$\epsilon_{t,i} \sim N(0, \sigma_{\epsilon,t,i}^2)$$

$$v_{\eta,t,i} \sim N(0, \gamma_i)$$

$$v_{\epsilon,t,i} \sim N(0, \gamma_i)$$

- Add a stochastic seasonality component to the UC-SV model
- seasonal components follow a stochastic process

$$y_{t,i} = \tau_{t,i} + \delta_{t,i,1} \cdot \text{seas}_1(t) + \dots + \delta_{t,i,4} \cdot \text{seas}_4(t) + \eta_{t,i},$$

$$\tau_{t,i} = \tau_{t-1,i} + \epsilon_{t,i},$$

$$\delta_{t,i,j} = \delta_{t-1,i,j} + \text{seas}_j(t) \cdot \xi_{t,i}, \quad j = 1, \dots, 4,$$

$$\ln(\sigma_{\eta,t,i}^2) = \ln(\sigma_{\eta,t-1,i}^2) + v_{\eta,t,i},$$

$$\ln(\sigma_{\epsilon,t,i}^2) = \ln(\sigma_{\epsilon,t-1,i}^2) + v_{\epsilon,t,i}.$$

- Estimate using particle filters

- AO-RW model
 - Introduced by Atkeson and Ohanian (2001), extended by Stock and Watson (2007)
 - Forecast inflation as a four-quarter average
- ARMA(X) model
- PCA-VAR model
 - Apply PCA to the correlation matrix of inflation
 - Apply a VAR(1) to the principal components
 - Transform forecasts back to inflation using eigenvectors

Training and Testing Windows

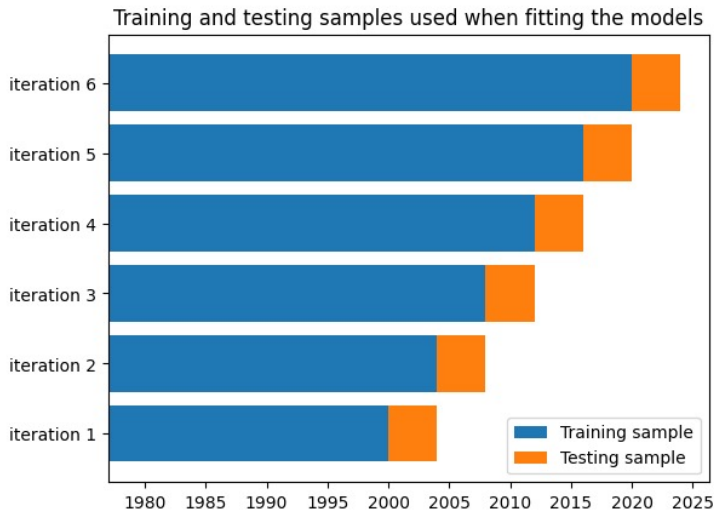
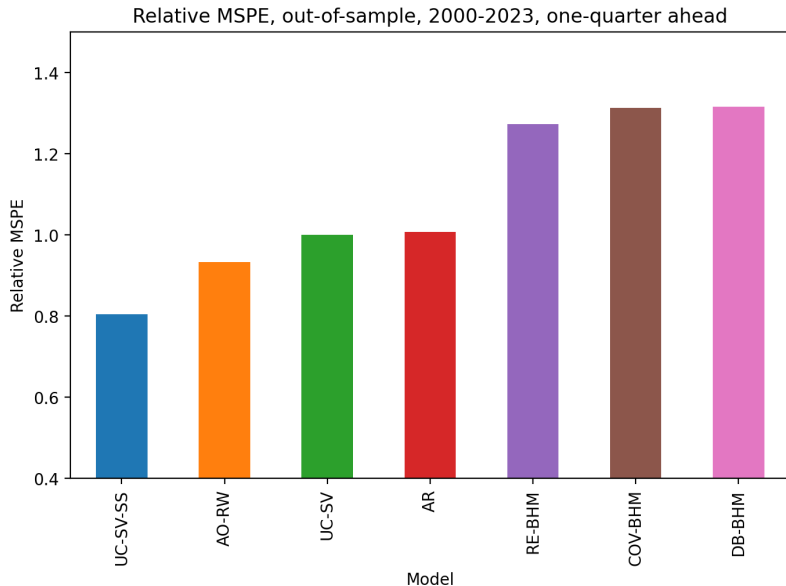


Figure: Training and testing windows used when fitting the models

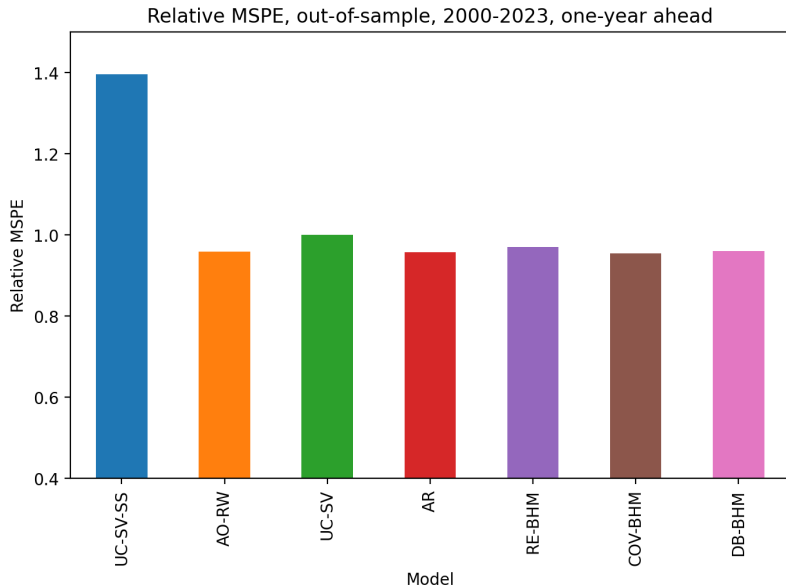
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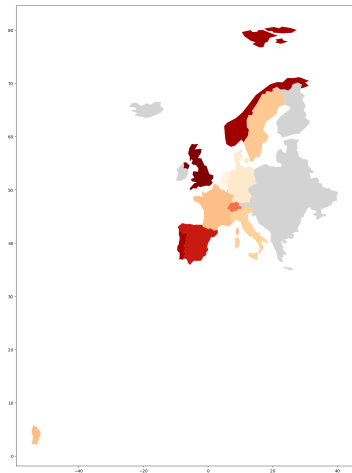
Model Performance



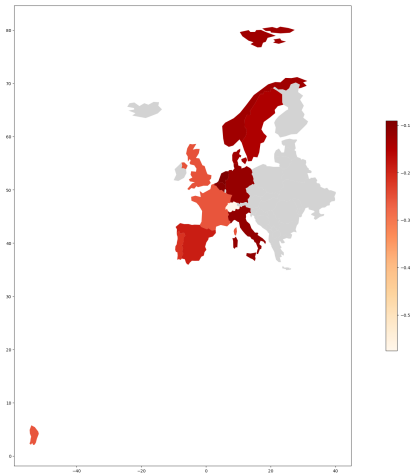
Model Performance



Coefficient Comparison



(a) AR Model



(b) Distance Based Model

Figure: Intercepts of the models per country, 1977-2000

Density Forecasts

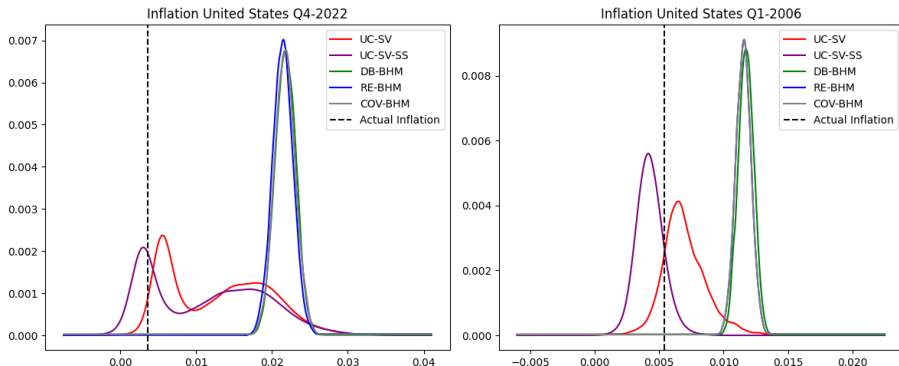


Figure: Density forecasts

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Research Questions:

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Further Research:

- Proxies that better measure the political and economic distance
- Weaker priors for the random effects model

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Fitting a BHM

To obtain the posterior distribution of $\theta|x, y$, we can use Bayes' theorem:

$$P(\theta | x, y) = \frac{P(x, y | \theta) \cdot P(\theta)}{P(x, y)} \propto P(x, y | \theta) \cdot P(\theta)$$

However, $P(x, y)$ is difficult to calculate.

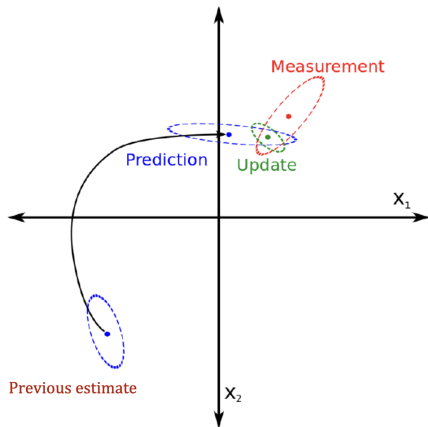
Solution: Markov Chain Monte Carlo (MCMC), e.g. Metropolis-Hastings

Particle filters

- Bayesian method: update priors

Sequential Importance Sampling with Resampling (SISR) algorithm:

- 1 Predict new hidden state of all points
- 2 Resample points using measurement likelihood as weights



AO-RW (Baseline model)

- Atkeson–Ohanian Random Walk model
- Forecast inflation as the average of the past four quarters
- Univariate and deterministic model

$$y_{t,i} = \frac{1}{4}(y_{t-1,i} + y_{t-2,i} + y_{t-3,i} + y_{t-4,i}) + \epsilon_{t,i}$$

ARMA(X) (Baseline model)

- Autoregressive Moving Average model with exogenous variables
- Univariate and flexible time-series model

$$y_{t,i} = c + \sum_{p=1}^P \alpha_p y_{t-p,i} + \sum_{q=1}^Q \theta_q \epsilon_{t-q,i} + \sum_{m=1}^M \beta_m x_{m,t-1,i} + \epsilon_{t,i}$$

- Lag length based on AIC

PCA-VAR (Baseline model)

- apply PCA to correlation matrix of inflation
- Vector Autoregression model of order 1

$$f_t = c + A_1 f_{t-1} + \epsilon_t,$$

- transform back to the original data space using the eigenvectors
- lead to parsimonious forecasts

- Evaluate accuracy of point forecasts

$$MSPE_{j,i} = \frac{\sum_{t=1}^T (\hat{y}_{j,t,i} - y_{t,i})^2}{T}$$

$$MAE_{j,i} = \frac{\sum_{t=1}^T |\hat{y}_{j,t,i} - y_{t,i}|}{T}$$

Mincer-Zarnowitz regression

- test efficiency and bias of forecast
- joint hypothesis of $\alpha_i = 0, \beta_i = 1$

$$y_{t+1,i} = \alpha_i + \beta_i \hat{y}_{t+1,i} + \epsilon_{t+1,i}$$

Country-Specific MSPE

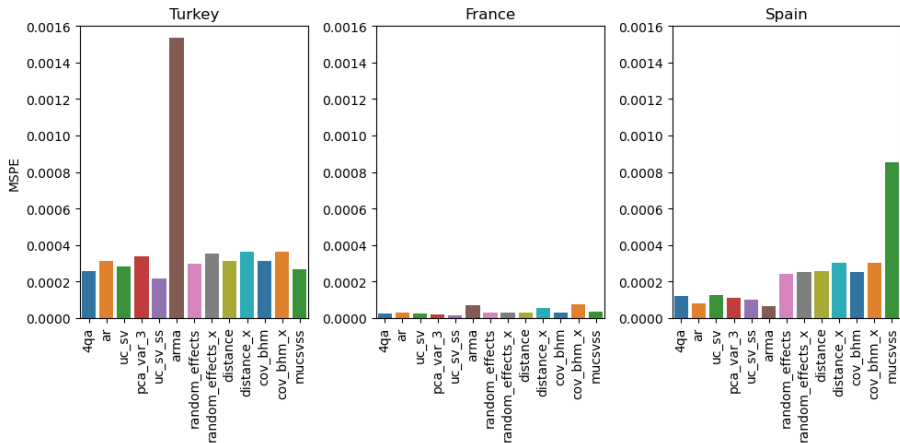


Figure: MSPE of the different models, one-quarter ahead, 2000-2023

Coefficient Comparison

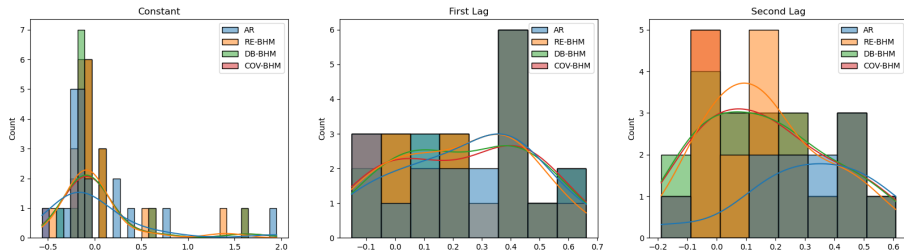


Figure: Coefficients of the AR model, RE-BHM, DB-BHM, and COV-BHM, 1977-2000

- Introduce a simple model that forecasts inflation in the next four quarters to be equal to that of the previous four quarters (Atkeson and Ohanian, 2001)
- Phillips curve based models outperform simple AR models (Stock and Watson, 1999)
- Bayesian model averaging produces better forecasts than UC-SV model (Medeiros et al., 2021)
- Machine learning techniques can outperform simpler inflation benchmarks (Groen et al., 2013)
- Per sector estimates of inflation can be used to improve the estimate of trend inflation (Stock and Watson, 2016)