

On the forgetting of particle filters

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Outline



Introduction: Hidden Markov Model and filtering

Particle filter

Forgetting

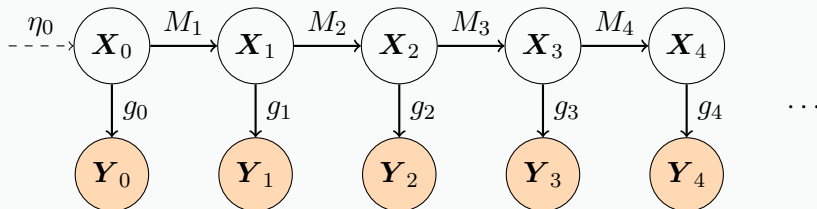
Forgetting result

Conditional particle filter

Concluding remarks

Introduction: Hidden Markov Model and filtering

Hidden Markov model (a.k.a. 'state-space model')



- Hidden (unobserved) Markov state process (X_0, X_1, X_3, \dots):
 - Initial density $\eta_0(x_0)$
 - Transition densities $M_t(x_t | x_{t-1})$
- Observations (or measurements) (Y_0, Y_1, Y_2, \dots):
 - Conditionally independent given $(X_t)_{t \geq 0}$
 - Observation densities $g_t(y_t | x_t)$
 - Observed values $y_0, y_1, \dots \rightsquigarrow$ potentials $G_t(x_t) = g_k(y_t | x_t)$

Running example: Noisy AR(1)

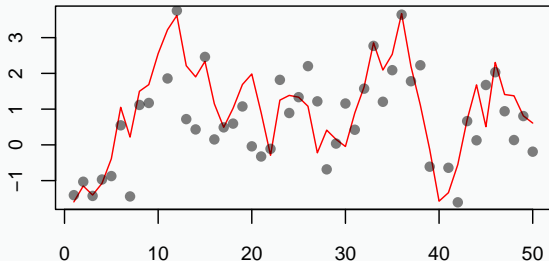


$X_{0:T}$ stationary AR(1) process:

- $X_1 \sim N(0, \sigma_\eta^2 / (1 - \phi^2))$.
- $X_k = \phi X_{k-1} + \eta_k; \eta_k \sim N(0, \sigma_\eta^2)$

$Y_{0:T-1}$ noisy observations of $X_{1:T-1}$:

- $Y_k \sim N(X_k, \sigma_Y^2)$
- $G_k(x_k) = c_Y \exp\left(-\frac{1}{2\sigma_Y^2}(x_k - y_k)^2\right)$



The filtering problem



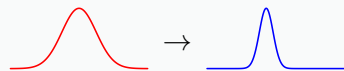
- We are interested in:
 - predictive distributions $\eta_t = \text{Law}(\mathbf{X}_t \mid \mathbf{y}_{0:t-1})$
 - filtering distributions $\pi_t = \text{Law}(\mathbf{X}_t \mid \mathbf{y}_{0:t})$
- Practical example: GPS navigation
 - varying quality GPS observations $\rightsquigarrow G_t$
 - simple model of movement such as Brownian velocity $\rightsquigarrow M_t$
- Cannot determine η_t and π_t in a closed form
(essentially unless M_t and G_t linear-Gaussian or state-space finite)
- We focus on Monte Carlo approximations, that is, sampling number of *particles* $\mathbf{X}_t^{1:N}$ “approximately from η_t ”, iteratively in $t = 0, 1, 2, \dots$

The ideal filter



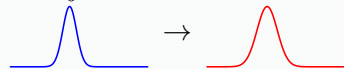
- Update $\eta_t \rightarrow \pi_t$ by weighting with G_t :

$$\pi_t(\mathbf{x}_t) = \Psi_t(\eta_t)(\mathbf{x}_t) \quad \text{where} \quad \Psi_t(\mu)(\mathbf{x}) = \frac{\mu(\mathbf{x})G_t(\mathbf{x})}{\int G_t(\mathbf{z})\mu(\mathbf{z})d\mathbf{z}}$$



- Mutate $\pi_t \rightarrow \eta_{t+1}$ by pushing through M_{t+1} :

$$\eta_{t+1}(\mathbf{x}_{t+1}) = (\pi_t M_{t+1})(\mathbf{x}_{t+1}) = \int M_{t+1}(\mathbf{x}_{t+1} | \mathbf{x}_t) \pi_t(\mathbf{x}_t) d\mathbf{x}_t$$



- We denote the composition of the above by Φ_t :

$$\eta_{t+1} = \Phi_{t+1}(\eta_t) = \Psi_t(\eta_t)M_{t+1}$$



and compositions of these by $\Phi_{t,u}$:

$$\eta_u = \Phi_{t,u}(\eta_t) = \Phi_u \circ \dots \circ \Phi_{t+1}(\eta_t)$$



Particle filter

Particle filter algorithm

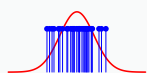


Gordon, Salmond and Smith (*IEE Proc. F*, 1993)

PF($\eta_0, (M_t)_{t \geq 1}, (G_t)_{t \geq 0}, N$)

```
1:  $\mathbf{X}_0^i \sim \eta_0(\cdot)$  for  $i \in \{1:N\}$ 
2: for  $t = 1, 2, \dots$  do
3:    $W_{t-1}^i = \frac{G_{t-1}(\mathbf{X}_{t-1}^i)}{\sum_{j=1}^N G_{t-1}(\mathbf{X}_{t-1}^j)}$  for  $i \in \{1:N\}$ 
4:    $A_{t-1}^i \sim \text{Categorical}(W_{t-1}^{1:N})$  for  $i \in \{1:N\}$ 
5:    $\mathbf{X}_t^i \sim M_t(\cdot \mid \mathbf{X}_{t-1}^{A_{t-1}^i})$  for  $i \in \{1:N\}$ 
6: end for
```

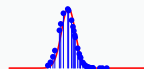
Produces empirical approximations of η_t and π_t :



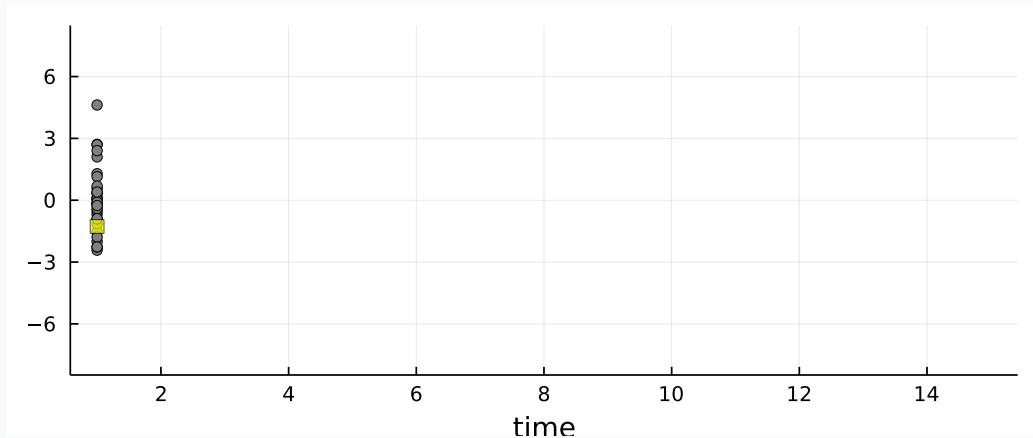
$$\eta_t^N = \frac{1}{N} \sum_{i=1}^N \delta_{\mathbf{X}_t^i}$$

and

$$\pi_t^N = \Psi_t(\eta_t^N) = \sum_{i=1}^N W_t^i \delta_{\mathbf{X}_t^i}$$

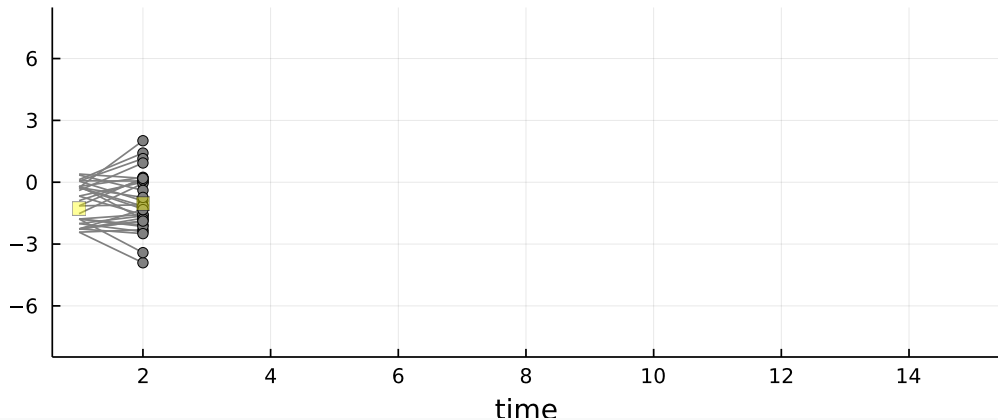


Particle filter on noisy AR(1): Initialise



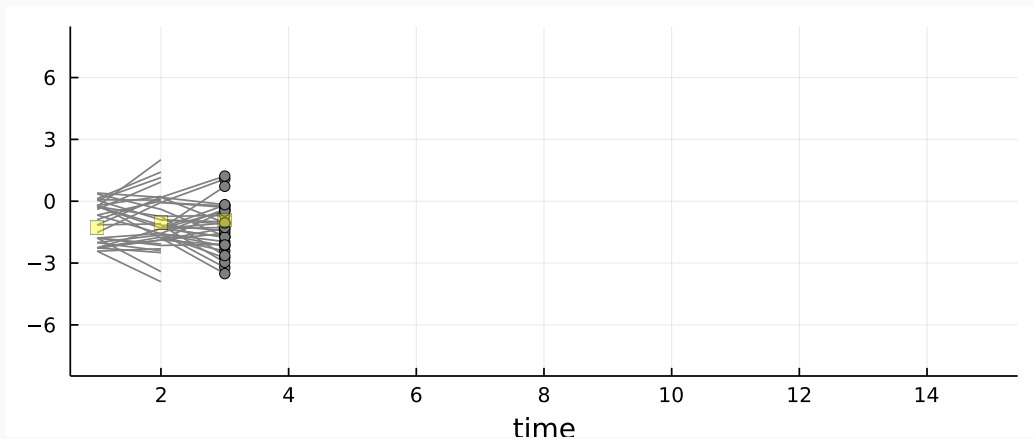
- \bullet Particles $X_1^{1:N} \sim M_1(\cdot)$
- \blacksquare Observation y_1

Particle filter on noisy AR(1): Resample and propagate



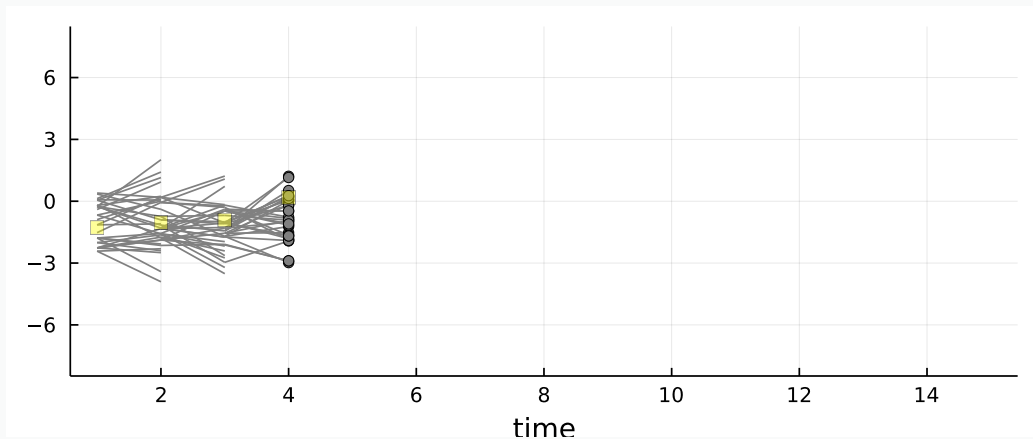
- ● Particles $X_t^i \sim M_t(X_{t-1}^{A_{t-1}^i}, \cdot)$
- ■ Observations $y_{0:t}$

Particle filter on noisy AR(1): Resample and propagate



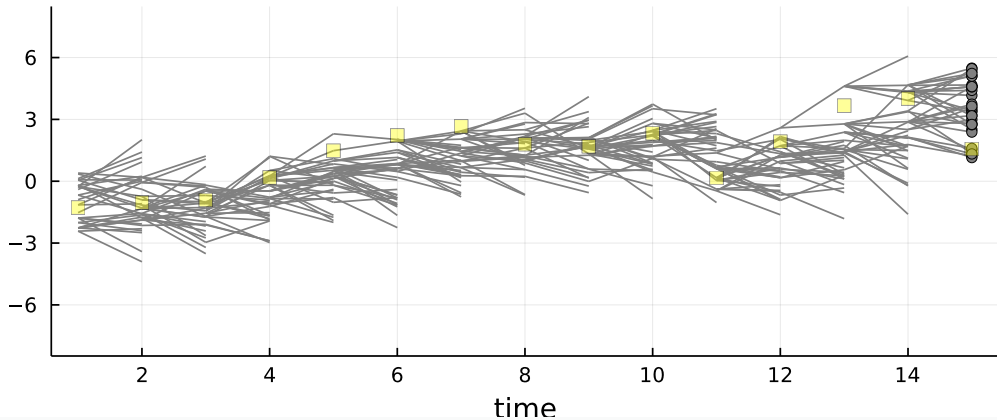
- ● Particles $X_t^i \sim M_t(X_{t-1}^{A_{t-1}^i}, \cdot)$
- ■ Observations $y_{0:t}$

Particle filter on noisy AR(1): Resample and propagate



- ● Particles $X_t^i \sim M_t(X_{t-1}^{A_t^i}, \cdot)$
- ■ Observations $y_{0:t}$

Particle filter on noisy AR(1): $t = 15$



- ● Particles $X_{15}^{1:N}$
- ■ Observations $y_{0:15}$

Forgetting

Strong mixing condition



We assume the following which is common in (quantitative) particle filter theory:

Assumption: Strong mixing

There exist $0 < \underline{M} \leq \overline{M} < \infty$ and $0 < \underline{G} \leq \overline{G} < \infty$ such that $\forall \mathbf{x}, \mathbf{x}', t$:

- $\underline{M} \leq M_t(\mathbf{x}, \mathbf{x}') \leq \overline{M}$
- $\underline{G} \leq G_t(\mathbf{x}) \leq \overline{G}$

- Typically holds only in a compact state space
- (Variants exist which e.g. extend the requirement for iterates of M_t)



Theorem (Del Moral 2004, Proposition 4.3.6)

For all probability measures μ and ν , $t \geq 0$ and $k \geq 0$:

$$\sup_{\mu, \nu} \|\Phi_{t,t+k}(\mu) - \Phi_{t,t+k}(\nu)\|_{\text{TV}} \leq \beta^k \quad \text{where} \quad \beta = 1 - \left(\frac{\underline{M}}{\overline{M}}\right)^2$$

- Ideal filter forgets at exponential rate
- NB: Φ_t is non-linear and generally **not contracting**: we might well have

$$\|\Phi_t(\mu) - \Phi_t(\nu)\|_{\text{TV}} > \|\mu - \nu\|_{\text{TV}}$$



Theorem (Del Moral 2004, Theorem 7.4.4)

For all $N \geq 2$, $n \geq 0$ and $p \geq 1$, $\text{osc}(\phi) \leq 1$:

$$\|\eta_t^N(\phi) - \eta_t(\phi)\|_p \leq \frac{c}{\sqrt{N}} \quad \text{where} \quad c_p = 2d_p^{1/p} \left(\frac{\overline{M}}{\underline{M}} \right)^3 \frac{\overline{G}}{\underline{G}}$$

where d_p has been defined in Del Moral (2004), and in particular, $d_2 = 1$.

- $\eta_t^N(\phi) = N^{-1} \sum_{i=1}^N \phi(\mathbf{X}_t^i) \approx \eta_t(\phi) = \int \phi(\mathbf{x}_t) \eta_t(\mathbf{x}_t) d\mathbf{x}_t$ for large N
uniform in t (in L^p sense)

∴ Monte Carlo errors do not accumulate \rightsquigarrow stability

Forgetting of the particle filter?



In summary:

- Ideal filter is exponentially forgetting
- Particle filter is increasingly accurate approximation of the ideal filter...
- ...in a time-uniform manner

So the particle filter must also be exponentially forgetting, at least if N is large enough?



- Unlike the ideal filter, particle filter defines a Markov chain $(\mathbf{X}_t^{1:N})_{t \geq 0}$
- Denote its Markov transition

$$\mathbf{M}_t(\mathbf{x}_{t-1}^{1:N}, \cdot) = \left(\sum_{i=1}^N \frac{G_{t-1}(\mathbf{x}_{t-1}^i)}{\sum_{j=1}^N G_{t-1}(\mathbf{x}_{t-1}^j)} M_t(\cdot \mid \mathbf{x}_{t-1}^i) \right)^{\otimes N}$$

- Is \mathbf{M}_t contracting in Dobrushin sense:

$$\beta_{\text{TV}}(\mathbf{M}_t) = \sup_{\mathbf{x}^{1:N}, \tilde{\mathbf{x}}^{1:N}} \|\mathbf{M}_t(\mathbf{x}^{1:N}, \cdot) - \mathbf{M}_t(\tilde{\mathbf{x}}^{1:N}, \cdot)\|_{\text{TV}} < 1?$$

An earlier forgetting result



Yes, \mathbf{M}_t are contracting:

Lemma (Tadić & Doucet, 2021)

For all $N \geq 1$ and $t \geq 0$:

$$\beta_{\text{TV}}(\mathbf{M}_t) \leq 1 - \epsilon^N, \quad \text{where} \quad \epsilon = \left(\frac{M}{\overline{M}} \right)^2$$

Direct corollary of the above:

$$\beta_{\text{TV}}(\mathbf{M}_{t,t+k}) \leq (1 - \epsilon^N)^k \quad \text{where} \quad \mathbf{M}_{t,t+k} = \mathbf{M}_{t+1} \mathbf{M}_{t+2} \cdots \mathbf{M}_{t+k}$$

\rightsquigarrow forgetting in $k = O(e^N)$ time 🤔

Forgetting result



Theorem (Karjalainen, Lee, Singh & V (2023))

For all $k \geq 1, t \geq 0, N \geq 2$,

$$\beta_{\text{TV}}(\mathbf{M}_{t,t+k}) \leq (1 - \varepsilon)^{\lfloor k/(c \log N) \rfloor},$$

where $\varepsilon \in (0, 1)$ and $c < \infty$ only depend on the strong mixing constants.

↪ PF forgets in $k = O(\log N)$ time 😊

- Seems like the right order: a specific example where forgetting $\Omega(\log N)$...

Proof sketch 1: Hellinger distance



Definition

The squared Hellinger distance between two probability measures P and Q having densities p and q with respect to a common dominating measure λ is

$$H^2(P, Q) = \frac{1}{2} \int (\sqrt{p(x)} - \sqrt{q(x)})^2 \lambda(dx) = 1 - \int \sqrt{p(x)q(x)} \lambda(dx)$$

Lemma (Tensorisation)

$$1 - H^2(P^{\otimes N}, Q^{\otimes N}) = (1 - H^2(P, Q))^N$$

Proof sketch 2: Total variation & random measures



Lemma (Le Cam's inequality)

For any probability measures μ and ν it holds that

$$\|\mu - \nu\|_{\text{TV}} \leq \sqrt{1 - (1 - H^2(\mu, \nu))^2}$$

Using tensorisation & Jensen's inequality:

Corollary

For random product measures

$$\|\mathbb{E}\mu^{\otimes N} - \mathbb{E}\nu^{\otimes N}\|_{\text{TV}} \leq \sqrt{1 - (1 - \mathbb{E}H^2(\mu, \nu))^{2N}}$$

Proof sketch 3: Bounding expected Hellinger with L^2



Lemma

Let μ and ν be two random probability measures, then

$$\mathbb{E}H^2(\mu M, \nu M) \leq c' \sup_{\text{osc}(\phi) \leq 1} \mathbb{E}(|\mu(\phi) - \nu(\phi)|^2) \quad \text{where} \quad c' = \frac{1}{8} \left(\frac{\overline{M}}{\underline{M}} \right)^2$$

Proof sketch 4: Putting things together



Let π_t^N and $\tilde{\pi}_t^N$ be particle filters with same M_t, G_t but with initial $\mathbf{X}_0^{1:N}$ and $\tilde{\mathbf{X}}_0^{1:N}$

$$\begin{aligned}\beta_{\text{TV}}(\mathbf{M}_{0,k}) &= \inf_{\mathbf{X}_0^{1:N}, \tilde{\mathbf{X}}_0^{1:N}} \|\mathbb{E}(\pi_{k-1}^N M_k)^{\otimes N} - \mathbb{E}(\pi_{k-1}^N M_k)^{\otimes N}\|_{\text{TV}} \\ &\leq \left(1 - \left(1 - c' \sup_{\text{osc}(\phi) \leq 1} \|\pi_{k-1}^N(\phi) - \tilde{\pi}_{k-1}^N(\phi)\|_2^2\right)^{2N}\right)^{1/2}\end{aligned}$$

and

$$\begin{aligned}\|\pi_{k-1}^N(\phi) - \tilde{\pi}_{k-1}^N(\phi)\|_2 &\leq \|\pi_{k-1}^N(\phi) - \pi_{k-1}(\phi)\|_2 && \leq cN^{-1/2} \\ &+ |\pi_{k-1}(\phi) - \tilde{\pi}_{k-1}(\phi)| && \leq \beta^{k-2} \leq cN^{-1/2} \text{ if } k \geq c \log N \\ &+ \|\tilde{\pi}_{k-1}(\phi) - \tilde{\pi}_{k-1}^N(\phi)\|_2 && \leq cN^{-1/2}\end{aligned}$$

in which case

$$\beta_{\text{TV}}(\mathbf{M}_{0,k}) \leq \left(1 - \left(1 - \frac{c}{N}\right)^{2N}\right)^{1/2} \approx (1 - e^{-2c})^{1/2} \leq 1 - \epsilon$$

□

Conditional particle filter

Conditional particle filter algorithm



Andrieu, Doucet & Holenstein (JRSS B, 2010)

CPF($\eta_0, (M_t)_{t \geq 1}, (G_t)_{t \geq 0}, (\mathbf{x}_t^*)_{t \geq 0}, N$)

```
1:  $\mathbf{X}_t^0 = \mathbf{x}_t^*$ , for  $t \geq 0$   
2:  $\mathbf{X}_0^i \sim \eta_0(\cdot)$  for  $i \in \{1:N\}$   
3: for  $t = 1, 2, \dots$  do  
4:    $W_{t-1}^i = \frac{G_{t-1}(\mathbf{X}_{t-1}^i)}{\sum_{j=1}^N G_{t-1}(\mathbf{X}_{t-1}^j)}$  for  $i \in \{0:N\}$   
5:    $A_{t-1}^i \sim \text{Categorical}(W_{t-1}^{0:N})$  for  $i \in \{1:N\}$   
6:    $\mathbf{X}_t^i \sim M_t(\cdot \mid \mathbf{X}_{t-1}^{A_{t-1}^i})$  for  $i \in \{1:N\}$   
7: end for
```

- Used for smoothing, but we think of CPF as a **perturbed** particle filter

↪ Empirical approximations of η_t and π_t :

$$\hat{\eta}_t^N = \frac{1}{N} \sum_{i=1}^N \delta_{\mathbf{X}_t^i} \quad \text{and} \quad \hat{\pi}_t^N = \sum_{i=1}^N W_t^i \delta_{\mathbf{X}_t^i}$$



Theorem (Karjalainen, Lee, Singh & V, 2023)

For every $p \geq 1$, there exists a constant $c = c(p)$ depending on strong mixing constants, such that for all $\text{osc}(\phi) \leq 1$, $t \geq 0$, $N \geq 1$ and x_t^ ,*

$$\|\hat{\pi}_t^N(\phi) - \pi_t(\phi)\|_p \leq \frac{c}{\sqrt{N}}$$

- Monte Carlo errors do not accumulate
- The effect of reference x_t^* remains limited, too



Theorem

For all $k \geq 1, t \geq 0, N \geq 2$, and references $x^* = (x_t^*)_{t \geq 0}$

$$\beta_{\text{TV}}(\mathbf{M}_{t,t+k}^{x^*}) \leq (1 - \varepsilon)^{\lfloor k/(c \log N) \rfloor},$$

where ε and c only depend on the strong mixing constants.

- Similar proof as for PF, using the time-uniform L^p error result

Concluding remarks

Maximal coupling of particle filters



- Coupled particle filters (and CPFs) have been used recently for multilevel Monte Carlo (and unbiased estimation)
- Let $\mu_k = \Phi_t(\eta_k^N)$ and $\tilde{\mu}_k = \Phi_t(\tilde{\eta}_k^N)$ stand for the conditional distributions of $X_t^{1:N}$ and $\tilde{X}_t^{1:N}$, respectively.
- Jasra & Yu (2020) suggested an ‘independent maximal coupling’ (IMC) algorithm:

$$(X_k^i, \tilde{X}_k^i) \sim \text{MAXCOUPLE}(\mu_k, \tilde{\mu}_k), \quad \text{independently for } i = 1, \dots, N$$

- Our analysis suggests another ‘joint maximal coupling’ (JMC) algorithm:

$$(X_k^{1:N}, \tilde{X}_k^{1:N}) \sim \text{MAXCOUPLE}(\mu_k^{\otimes N}, \tilde{\mu}_k^{\otimes N})$$

- Both are implementable, but have $O(N^2)$ complexity

Coupling probability when N increases



- Suppose $X_t^{1:N}$ and $\tilde{X}_t^{1:N}$ follow same dynamics, but have $\eta_0 \neq \tilde{\eta}_0$
- Suppose that $k > c \log N$
- Our analysis says that with JMC, for all $N \geq 1$:

$$\mathbb{P}(X_t^{1:N} \neq \tilde{X}_t^{1:N}) = \mathbb{E} \|\mu_k - \tilde{\mu}_k\|_{\text{TV}} \leq \left(1 - \left(1 - \frac{c}{N}\right)^{2N}\right)^{1/2} \leq \epsilon$$

\rightsquigarrow filters fully coupled after $O(\log N)$ iterations.

- In contrast, with IMC, we can only guarantee:

$$\mathbb{P}(X_t^{1:N} = \tilde{X}_t^{1:N}) = \mathbb{E}[(1 - \|\mu_k - \tilde{\mu}_k\|_{\text{TV}})^N] \geq \left(1 - \frac{c}{\sqrt{N}}\right)^N,$$

which vanishes as $N \rightarrow \infty$



- Dobrushin forgetting is a strong form of stability
 - Complementary to L^p uniform stability, which is about averages
 - Could prove useful in some PF/CPF analysis
- Forgetting entails $\log N$ 'penalty' term
 - Seems unavoidable
 - Drowned by N^{-p} where $p > 0$
- Implications to CPF out soon...



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